

5/3/21

Lec 20

$$P(x_* | x) \propto \int_0^\infty \int_{\mathbb{R}} k(x_* | \theta, \sigma^2) k(\theta, \sigma^2 | x) d\theta d\sigma^2$$

Assume Jeffreys

Prior $P(\theta, \sigma^2) \propto \frac{1}{\sigma^2}$

$$= \int_0^\infty \int_{\mathbb{R}} \frac{1}{\sqrt{\sigma^2}} e^{-\frac{1}{2\sigma^2}(x_* - \theta)^2} (\sigma^2)^{-\frac{n-1}{2}} e^{-\frac{(n-1)s^2/2}{\sigma^2}} e^{-\frac{n}{2\sigma^2}(\theta - \bar{x})^2} d\theta d\sigma^2$$

$$= \int_0^\infty (\sigma^2)^{\frac{1}{2}} (\sigma^2)^{-\frac{n-1}{2}} e^{-\frac{(n-1)s^2/2}{\sigma^2}} \int_{\mathbb{R}} e^{-\frac{1}{2\sigma^2} \left(\underbrace{(x_* - \theta)^2 + n(\theta - \bar{x})^2}_{\substack{x_*^2 - 2x_*\theta + \theta^2 + n\theta^2 \\ -2n\theta\bar{x} + n\bar{x}^2}} \right)} d\theta d\sigma^2$$

$$= \int_0^\infty (\sigma^2)^{-\frac{n+1}{2}} e^{-\frac{(n-1)s^2/2 + x_*^2/2 + n\bar{x}^2/2}{\sigma^2}} \int_{\mathbb{R}} e^{-\frac{x_* + n\bar{x}}{\sigma^2} \theta - \frac{n+1}{2\sigma^2} \theta^2} d\theta d\sigma^2$$

Side note:

$$\int_{\mathbb{R}} e^{a\theta + b\theta^2} d\theta = \sqrt{\frac{\pi}{b}} e^{a^2/4b}$$

$$= \int_0^\infty (\sigma^2)^{-\frac{n+1}{2}} e^{-\frac{(n-1)s^2/2 + x_*^2/2 + n\bar{x}^2/2}{\sigma^2}} \sqrt{\frac{\pi}{\frac{n+1}{2\sigma^2}}} e^{\frac{(x_* + n\bar{x})^2}{4 \cdot \frac{n+1}{2\sigma^2}}} d\sigma^2$$

$$= \int_0^\infty (\sigma^2)^{-\frac{n+1}{2}} e^{-\frac{(n-1)s^2/2 + x_*^2/2 + n\bar{x}^2/2}{\sigma^2}} \sqrt{\frac{\pi}{\frac{n+1}{2\sigma^2}}} e^{\frac{(x_* + n\bar{x})^2}{2\sigma^2(n+1)}} d\sigma^2$$

$$\propto \int_0^\infty (\sigma^2)^{-\frac{n}{2}} e^{-\frac{\beta}{\sigma^2}} d\sigma^2 = \Gamma(\alpha) \beta^{-\alpha} = \Gamma\left(\frac{n}{2}\right) \beta^{-\alpha} \propto \beta^{-\alpha}$$

$$= \left[\frac{(n-1)s^2}{2} + \frac{x_*^2}{2} + \frac{n\bar{x}^2}{2} - \frac{(x_* + n\bar{x})^2}{2(n+1)} \right]^{-\frac{n}{2}} = (ax_*^2 + bx_* + c)^{-\frac{n}{2}}$$

$$a = \frac{1}{2} - \frac{1}{2n+2} = \frac{1}{2} \left(1 - \frac{1}{n+1} \right) = \frac{1}{2} \frac{n}{n+1}$$

$$b = \frac{2n\bar{x}}{2n+2} = -\frac{n\bar{x}}{n+1}$$

$$c = \frac{(n-1)s^2}{2} + \frac{n\bar{x}^2}{2} - \frac{n^2\bar{x}^2}{2(n+1)} = \frac{1}{2} \left((n-1)s^2 + n\bar{x}^2 - \frac{n^2\bar{x}^2}{n+1} \right)$$

$$\begin{aligned}
 &= \left(\frac{1}{a}\right)^{\frac{n}{2}} \left(\frac{1}{a}\right)^{-\frac{n}{2}} (ax_*^2 + bx_* + c)^{-\frac{n}{2}} = \left(\frac{1}{a}\right)^{-\frac{n}{2}} \left(x_*^2 + \frac{b}{a}x_* + \frac{c}{a}\right)^{-\frac{n}{2}} \\
 &\propto \left(x_*^2 + \frac{b}{a}x_* + \frac{c}{a}\right)^{-\frac{n}{2}} \propto \left(\left(x_* + \frac{b}{2a}\right)^2 + \frac{c}{a} - \frac{b^2}{4a^2}\right)^{-\frac{n}{2}} \left(\frac{1}{\frac{c}{a} - \frac{b^2}{4a^2}}\right)^{-n/2} \\
 &\propto \left(1 + \frac{\left(x_* + \frac{b}{2a}\right)^2}{\frac{c}{a} - \frac{b^2}{4a^2}}\right)^{-\frac{(n-1)+1}{2}} \\
 &= \left(1 + \frac{1}{\frac{(n-1)}{2S^2}} \frac{\left(x_* - \frac{b}{2a}\right)^2}{\left(\frac{c}{a} - \frac{b^2}{4a^2}\right)/(n-1)}\right)^{-\frac{(n-1)+1}{2}} \propto T_2(\mu, S^2) \\
 &= T_{n-1}\left(\bar{x}, \frac{n+1}{n}S^2\right) \stackrel{n \text{ large}}{\approx} N(\bar{x}, S^2)
 \end{aligned}$$

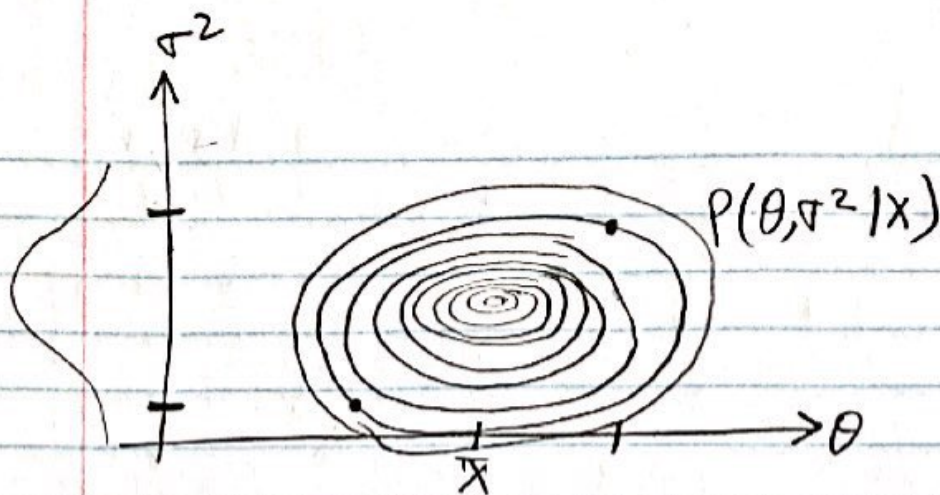
$$\begin{aligned}
 \frac{-b}{2a} &= \frac{\frac{n\bar{x}}{n+1}}{\frac{n}{n+1}} = \bar{x} \\
 \frac{c}{a} - \frac{b^2}{4a^2} &= \frac{c}{a} - \frac{b^2}{(2a)^2} = \frac{\frac{1}{2}((n-1)S^2 + n\bar{x}^2 - \frac{n^2\bar{x}^2}{n+1})}{\frac{1}{2}\frac{n}{n+1}} = \frac{(n-1)(n+1)S^2}{n} \\
 &= \frac{(n-1)(n+1)}{n}S^2 + \frac{(n+1)\bar{x}^2}{n} - n\bar{x}^2 \\
 -\left(\frac{b^2}{4a^2}\right) &= \left(\frac{-b}{2a}\right)^2 = -\bar{x}^2 \Rightarrow \frac{c}{a} - \frac{b^2}{4a^2} \\
 S_1^2 &= \frac{(n-1)(n+1)S^2}{n} = \frac{n+1}{n}S^2
 \end{aligned}$$

under Jeffreys Prior

$$P(\theta, \sigma^2 | x) = P(\theta | x, \sigma^2) P(\sigma^2 | x) \quad \begin{array}{l} \text{By def of Cond.} \\ \text{Probability} \end{array}$$

$$\text{Norm Inv Gamma}(\dots) \stackrel{\downarrow}{=} N\left(\bar{x}, \frac{\sigma^2}{n}\right) \cdot \stackrel{\downarrow}{\text{Inv Gamma}}\left(\frac{n-1}{2}, \frac{(n-1)S^2}{2}\right)$$

You can think of a normal inverse gamma as first sampling from an Inverse gamma $\left(\frac{n-1}{2}, \frac{(n-1)S^2}{2}\right)$ to get a σ^2 value and then you see that value of σ^2 to draw a θ from $N\left(\bar{x}, \frac{\sigma^2}{n}\right)$ and return the two-dimensional point $[\theta, \sigma^2]$.



This can also be done another way...

$$P(\theta, \sigma^2 | x) = P(\sigma^2 | x, \theta) P(\theta | x) \downarrow \text{InvGamma}\left(\frac{n}{2}, \frac{n \hat{\sigma}_{MLE}^2}{2}\right) T_{n-1}\left(\bar{x}, \frac{s^2}{2}\right)$$

If we decompose the first way, we draw θ from $N(\bar{x}, \frac{\sigma^2}{n})$ and thus σ^2 must be known. What if we break this by instead using **Jeffrey's** prior, using

$$P(\theta) = N(\mu_0, \tau^2) \text{ and } P(\sigma^2) = \text{InvGamma}\left(\frac{n_0}{2}, \frac{n_0 \sigma_0^2}{2}\right)$$

These were the two priors we began with when we started investigating the normal likelihood model. However, it's important to note we are not allowing $\tau^2 = \sigma^2/n_0$

What happens? The two priors are disconnected completely.

$$P(\theta, \sigma^2) = P(\theta)P(\sigma^2) \text{ not } P(\theta | \sigma^2)P(\sigma^2)$$

Let's derive the posterior under this two-Dimen. Prior.

$$\begin{aligned} P(\theta, \sigma^2 | x) &\propto P(x | \theta, \sigma^2) P(\theta, \sigma^2) \downarrow P(x | \theta, \sigma^2) P(\theta) P(\sigma^2) \\ &\propto k(x | \theta, \sigma^2) k(\theta) k(\sigma^2) \sigma^2 \exp\left(-\frac{x^2}{2\sigma^2}\right) \\ &= (\sigma^2)^{-\frac{n}{2}} e^{-\frac{1}{2\sigma^2}((n-1)s^2 + n(\bar{x} - \theta)^2)} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2\tau^2}(\theta - \mu_0)^2} \left(\frac{n_0}{\sigma^2}\right)^{\frac{n_0}{2}-1} e^{-\frac{n_0 \sigma_0^2}{2\sigma^2}} \\ &= (\sigma^2)^{-\frac{n}{2} - \frac{n_0}{2} - 1} e^{-\frac{1}{2\sigma^2}((n-1)s^2 + n_0 \sigma_0^2 + n\bar{x}^2)} e^{\left(\frac{n\bar{x}}{\sigma^2} + \frac{\mu_0}{\tau^2}\right)\theta - \left(\frac{n}{2\sigma^2} + \frac{1}{2\tau^2}\right)\theta^2} \end{aligned}$$