```
In [1]: import pandas as pd
    import numpy as np
    import matplotlib.pyplot as plt
    import seaborn as sns
    import random as rn
    import statsmodels.api as sm
    from scipy import stats
In [2]: random_generation = np.random.default_rng(20201014)
```

Warmup: Correlation (10%)

Below function is created to generate random numbers and draw the distribution for correlation between independed and dependent variables

```
In [3]: def generate correlation plot(num sample, num iterations, draw dependent variable = False):
            # Seed for random number generation, it will ensure the results are reproduce
            random generation = np.random.default rng(20201014)
            corr = []
            for i in range(num iterations):
                xs = pd.Series(random generation.normal(size=num sample))
                ys = pd.Series(random generation.normal(size=num sample))
                if draw dependent variable:
                    zs = xs + ys
                    corr.append(xs.corr(zs))
                else:
                    corr.append(xs.corr(ys))
            coefficienct mean = round(np.mean(corr),5)
            coefficienct variance = round(np.square(np.std(corr)),5)
            sns.histplot(corr)
            plt.xlabel(f'Correlation Coefficient')
            plt.title(f'For {num sample} Samples and {num_iterations} Iterations\n Mean: {coefficienct_mean} &
```

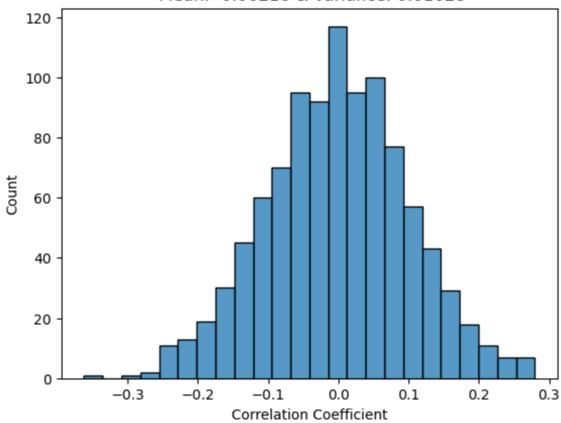
Independent Variables

- Below 2 graphs show the distribution of the correlation coefficients between 2 independent variables drawn from a normal distribution.
- Since our draw is for independent variables, we are expecting a 0 correlation coefficients and hence 0 variance and histrogram distribution should be centred around 0.
 - When I draw 100 samples in each iteration for 1000 times, the mean of correlation coefficients is not 0 however quite less (-0.00216) and the variance is (0.01028).
 - When I draw 1000 samples in each iterations for 1000 times, the mean is (-0.00052) and variance is (0.00102) which is significantly less for sample size of 100 but still not 0. However, the distribution has less variance and concentrated around mean.

From the above experiement, I can say, if we keep on drawing more samples in one iteration, then the mean of the correlation coefficient will be very much closer to 0 and will have almost insignificant variance. This aligns with definition of the independent random variables.

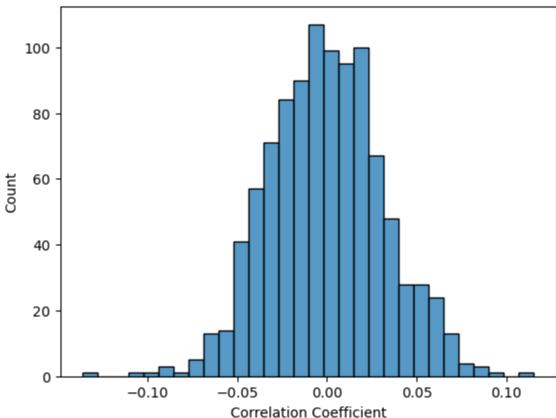
```
In [4]: NUM_ITERATIONS = 1000
sample_size = 100
generate_correlation_plot(sample_size, NUM_ITERATIONS)
```

For 100 Samples and 1000 Iterations Mean: -0.00216 & Variance: 0.01028



In [5]: sample_size = 1000
generate_correlation_plot(sample_size, NUM_ITERATIONS)





Dependent Variables

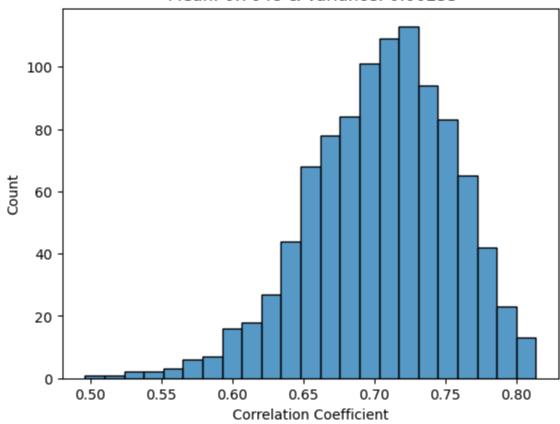
- Below I am plotting the distribution between two dependent variables xs, and zs. The zs = xs + ys. Since, zs is linearly dependent on xs and positively related i.e. change in one's value for example if it increases the dependent variable's value also increase.
- Since, the xs & zs are positively linked, so we are expecting correlatin coefficient > 0 and I am expecting a strong correlation between these 2, so mean value of the correlation coefficient should be closer to 1 but not 1 as zs is dependent on ys as well.

- Like independe variables, I draw 2 different sample sizes 100 & 1000 for 1000 iterations. In both of the distributions, the mean value is almost same and the variance is also the same.
- Also, the both distributions are centred around mean (0.70) and an less significant variance. This suggestes, even if I keep on increasing the sample size, the mean value will be around (0.70).

For the above experiement, I can say the dependent variable will have non-zero correlation coefficient and will indicate the dependency of

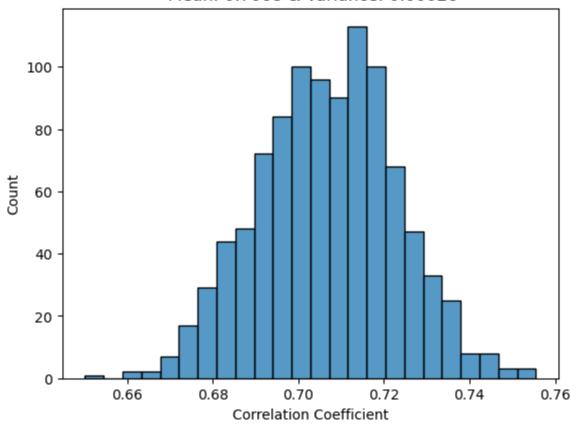
```
In [6]: sample_size = 100
generate_correlation_plot(sample_size, NUM_ITERATIONS, draw_dependent_variable = True)
```

For 100 Samples and 1000 Iterations Mean: 0.7048 & Variance: 0.00255



In [7]: sample_size = 1000
generate_correlation_plot(sample_size, NUM_ITERATIONS, draw_dependent_variable = True)

For 1000 Samples and 1000 Iterations Mean: 0.7068 & Variance: 0.00026



In [8]: # Since everywhere sample size is fixed to 1000, I am defining here
sample_size = 1000

Linear Regression (35%)

$$y = \alpha + \beta x + \epsilon$$

```
In [9]: def linear regression(num sample, slope, intercept, is single fit = True):
            random generaton = np.random.default rng(20201014)
            # Generate Independent variable samples from a standard normal distribution
            xs = random generation.standard normal(num sample)
            # Generate the random i.i.d noise
            errs = random generation.standard normal(num sample)
            # Generate the Dependent variable using Independent Variable, noise, slope, and intercept
            ys = intercept + slope * xs + errs
            df = pd.DataFrame({'X':xs, 'Y':ys})
            X = df['X']
            Y = df['Y']
            X = sm.add constant(X)
            ols model = sm.OLS(Y, X).fit()
            ols summary = ols model.summary()
            if is single fit:
                fig, axes = plt.subplots(nrows=1, ncols=2, figsize=(10, 4))
                # Plot the residual vs. fitted plot
                fitted values = ols model.fittedvalues
                residuals = ols model.resid
                axes[0].scatter(fitted values, residuals)
                axes[0].set title("Residuals vs. Fitted")
                axes[0].set_xlabel("Fitted Values")
                axes[0].set ylabel("Residuals")
                axes[0].axhline(y=0, color='r', linestyle='-')
                sm.qqplot(residuals, line='45', ax=axes[1])
                axes[1].set title("Q-Q Plot of Residuals")
                plt.show()
                print(ols summary)
                return None
            else:
                fitted slope = ols_model.params['X']
                fitted intercept = ols model.params['const']
                mdl rsquared = ols model.rsquared
```

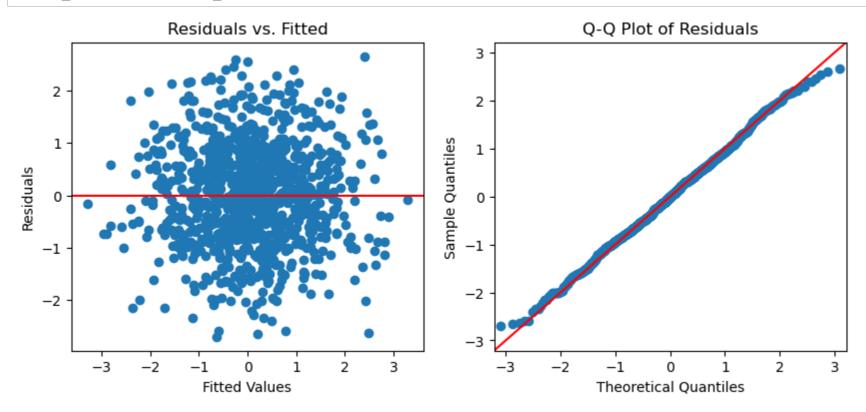
Linear Regression for (1 : slope,0 : intercept) 1,000 Samples

- Below I have fitted an OLS model for 1000 random samples generated for independent variable x.
- I have used \$Y = 1*X + 0 + \epsilon\$ to generate the dependent variable Y with i.i.d normal noise \$\epsilon\$.
- The I fitted the x on y using OLS model. The fitted model shows that:
 - 53.8% of the variance in "Y" is explained by "X" as indicated from R-squared. It may be noted here, since there is only one independent variable, hence Adj. R-squared is nearly same as R-squared
 - The low **p-value** and high **F-statistic** values indicate that this model is statistically significant which was expected as Y is dependent on X only other than the noise.
 - The model estimates the Coefficient for x (slope) is 1.0249 with a standard error of 0.030. This is closer to my given slope 1 and lowe p-value for X also indicates this coefficient is significant, which is again true. As my actual slope is 1.
 - Further as expected, the const (intercept) is not statistically significant as p-value is 0.601.
 - Confidence Interval: The 95% confidence interval for the coefficient of "X" is (0.966, 1.084).

In summary, this OLS regression model is statistically significant, and the coefficient for "X" is estimated to be 1.0256, suggesting a significant relationship between "X" and "Y." The model explains 53.8% of the variance in "Y," and various diagnostic tests have been performed to assess the model's quality.

- From Residul and Q-Q plots, we can say:
 - Since the Q-Q plot closely follow a straight diagonal line, it suggests that the residuals are normally distributed.
 - From Residals vs Fitted graph, we can say the residuals are scattered around horizontal line suggesting the linearity assumption holds true.
 - Also from Residals vs Fitted, the lack of any pattern among residuals scatter plot implies that Homoscedasticity assumption holds true

In [10]: linear_regression(sample_size, 1, 0)



Dep. Variable: Model:				0.538				
			OLS	Adj.	R-squared: Adj. R-squared:			
Method:		Least Squ	ares	F-sta	_			
Date:		Fri, 13 Oct	2023	Prob	(F-statistic):	2.42e-169	
Time:		21:	19:08	Log-L	,			
No. Observation	ons:		1000	AIC:			2765.	
Df Residuals:			998	BIC:			2775.	
Df Model:			1					
Covariance Typ	e:	nonro	bust					
==========	-=====			======	========	=======		
	coe	f std err		t	P> t	[0.025	0.975]	
const	0.0160	0.031		0.523	0.601	-0.044	0.076	
X	1.0249	0.030	3	34.063	0.000	0.966	1.084	
Omnibus:	======	=======================================	===== l.190	Durbi:	======== n-Watson:	=======	 1.955	
Prob(Omnibus):	:	(0.552	Jarqu	e-Bera (JB):		1.232	
Skew:		(0.040	Prob(JВ):		0.540	
Kurtosis:		:	2.848	,	•		1.05	
=========			=====	======				

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

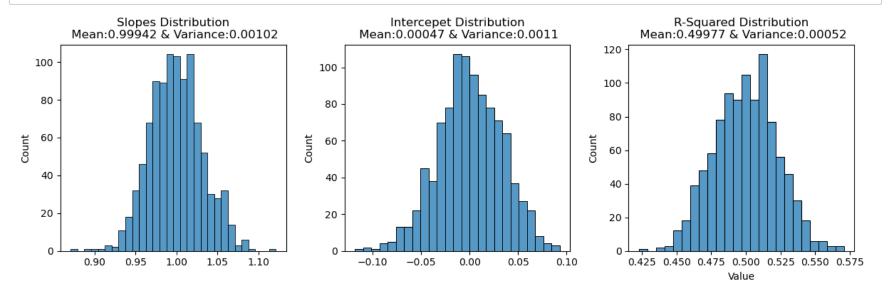
```
In [11]: def run linear regression simulation(num sample, slope, intercept, num iterations):
             slopes = []
             intercepts = []
             rsquareds = []
             for i in range(num iterations):
                 fit slope, fit intercept, rsquared = linear regression(num sample, slope, intercept, is single
                 slopes.append(fit slope)
                 intercepts.append(fit intercept)
                 rsquareds.append(rsquared)
             slope mean = round(np.mean(slopes),5)
             intercept mean = round(np.mean(intercepts),5)
             rsquared mean = round(np.mean(rsquareds),5)
             slope variance = round( np.square(np.std(slopes)),5 )
             intercept variance = round( np.square(np.std(intercepts)),5 )
             rsquared variance = round( np.square(np.std(rsquareds)),5)
             fig, axs = plt.subplots(1,3, figsize=(12, 4))
             sns.histplot(slopes, ax=axs[0])
             axs[0].set title(f'Slopes Distribution\n Mean:{slope mean} & Variance:{slope variance}')
             sns.histplot(intercepts, ax=axs[1])
             axs[1].set_title(f'Intercepet Distribution\n Mean:{intercept_mean} & Variance:{intercept variance}'
             sns.histplot(rsquareds, ax=axs[2])
             axs[2].set title(f'R-Squared Distribution\n Mean:{rsquared mean} & Variance:{rsquared variance}')
             # Add labels and a title for the entire figure
             plt.xlabel('Value')
             plt.tight layout()
             plt.show()
```

Linear Regression for (1 : slope,0 : intercept) 1,000 Samples & 1,000 simulations

- By running the above linear regression fit for 1000 times, I can that:
 - The mean coefficient is distributed around 1 with an insignificant variance aligning with the expectations.

- The mean intercept is distributed around 0 with very low variance thus aligning with the exceptations.
- ~50% of the variances in Y are explained by X

In [12]: run_linear_regression_simulation(sample_size, 1, 0, 1000)

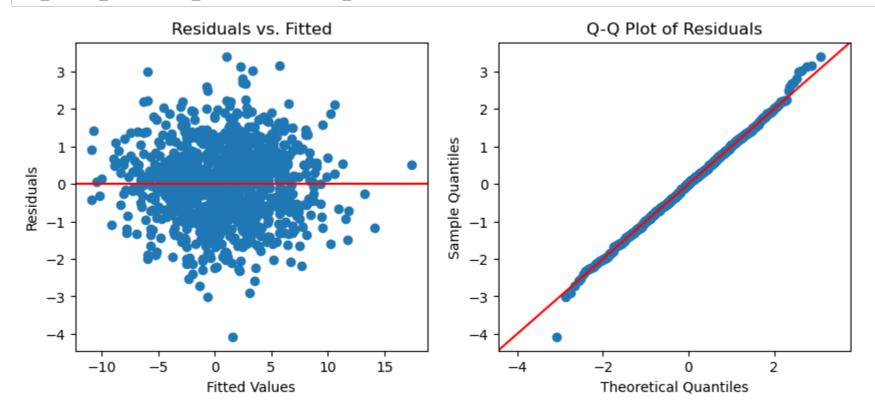


Linear Regression for (4 : slope,1 : intercept) 1,000 Samples & 1,000 simulations

- · Are the resulting model parameters what you expect?
 - Yes, the result model parameters aligns with the expectation. My **intercept** was 1 and model has predicted it to be 0.9572 with 95% confidence lying between [0.896 1.018] and this time the predicted intercept is significant (low p-value) again aligns with expectations.
 - The coefficient of x (slope) also matches the expectation as 95% confidence iterval is [3.993 4.117] with a predicted value very close to given slope = 4
- How did R² change, and why?
 - Now, ~95% of the variance in Y has been explained. This is because the coefficients for both the constant (intercept) and the independent variable X (slope) are higher, which suggests a stronger relationship between the two variables. This results in a higher R-squared value, indicating a better fit of the model to the data suggesting the choice of independent variables and their coefficients can have a significant impact on the R-squared value in regression analysis.
- Do the linear model assumptions still hold?

- Yes, the 'Q-Q plot' and 'Residuals vs Fitted' are very much similar to the previous model thus suggesting that all these assumptions still hold
- What are the distributions of the slope, intercept, and \mathbb{R}^2 if you do this 1000 times?
 - The distributions of slope, intercept, and R-Squared is also similar to previous fit model with 1,0 slope and intercept respectively.
 - The slope and intercept are very much close to the given values (4 & 1) with no significant variance in these.

```
In [13]: linear_regression(sample_size, 4, 1)
run_linear_regression_simulation(sample_size, 4, 1, 1000)
```

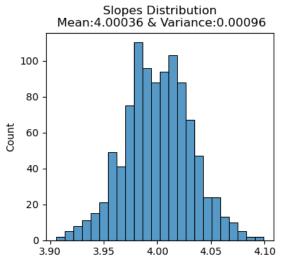


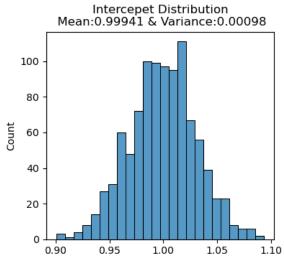
===========			=======	
Dep. Variable:	Y	R-squared:		0.944
Model:	OLS	Adj. R-squared:		0.944
Method:	Least Squares	F-statistic:		1.673e+04
Date:	Fri, 13 Oct 2023	Prob (F-statistic):		0.00
Time:	21:19:10	Log-Likelihood:		-1428.1
No. Observations:	1000	AIC:		2860.
Df Residuals:	998	BIC:		2870.
Df Model:	1			
Covariance Type:	nonrobust			
=======================================		=======================================	=======	========
_	roof ctd orr	+ D> +	rn 025	0 9751

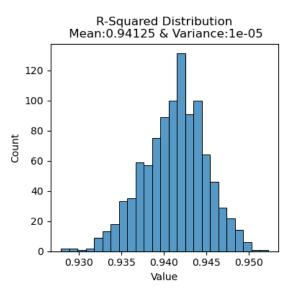
	coef	std err	t	P> t	[0.025	0.975]
const X	1.0099 4.0428	0.032 0.031	31.546 129.345	0.000	0.947 3.981	1.073 4.104
Omnibus: Prob(Omnibus) Skew: Kurtosis:	: :	0.		,	=======	1.969 1.681 0.431 1.07

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.







Redefined a function to fit linear model to generate sample from Normal distribution with given parameters

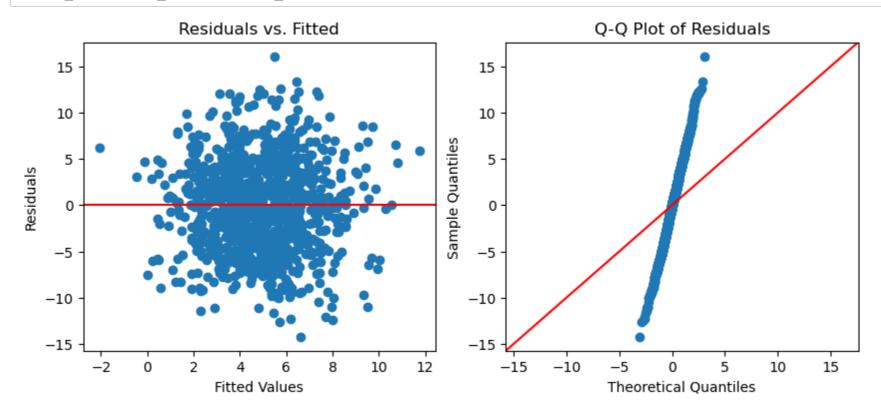
```
In [14]: def linear regression normal(num sample, slope, intercept, normal mean, normal std, normal std err):
             random generaton = np.random.default rng(20201014)
             # Generate Independent variable samples from a standard normal distribution
             xs = random generation.normal(loc= normal mean, scale= normal std, size= num sample)
             # Generate the random i.i.d noise
             errs = random generation.normal(loc= normal_mean, scale= normal_std_err, size= num_sample)
             # Generate the Dependent variable using Independent Variable, noise, slope, and intercept
             ys = intercept + slope * xs + errs
             df = pd.DataFrame({'X':xs, 'Y':ys})
             X = df['X']
             Y = df['Y']
             X = sm.add constant(X)
             ols model = sm.OLS(Y, X).fit()
             ols summary = ols_model.summary()
             fig, axes = plt.subplots(nrows=1, ncols=2, figsize=(10, 4))
             # Plot the residual vs. fitted plot
             fitted values = ols model.fittedvalues
             residuals = ols model.resid
             axes[0].scatter(fitted values, residuals)
             axes[0].set title("Residuals vs. Fitted")
             axes[0].set xlabel("Fitted Values")
             axes[0].set ylabel("Residuals")
             axes[0].axhline(y=0, color='r', linestyle='-')
             sm.qqplot(residuals, line='45', ax=axes[1])
             axes[1].set title("Q-Q Plot of Residuals")
             plt.show()
             print(ols_summary)
```

Linear Regression for (-2 : slope,5 : intercept) 1,000 Samples

• Are the resulting model parameters what you expect? What about the \mathbb{R}^2 ?

- The resulting parameters of the model are *intercept* = 4.8747 and *slope* = -2.0774 and these predicted parameters are statistically significant because of low p-value. These values are quite closer to the actual values (5 & -2) with a standard error of ~.15 in both
- The R-squared value is ~15% i.e. only 15% of the variance in Y are explained in this model and moreover F-statistic is low compared to the above models, hence this model seems not statistically significant.
- Do the linear model assumptions still hold?
 - Q-Q Plot clearly indicates deviation in Sample vs Theoretical qunantities. The deviation of residuals from the diagonal line indicates that the residuals are not normally distributed. There the assumpation of Residuals are normally distributed does not hold true.
 - However from the 'Residuals vs Fitted' plot, other assumptions related to residuals' Homoscedasticity, Linearity, and Independence assumptions still hold true

In [15]: linear_regression_normal(sample_size, -2, 5, 0, 1, 5)



Dep. Varia	======= ble:		====== Y	R-squ	======= ared:		0.132
Model:			OLS	_	R-squared:		0.131
Method:		Least Sq	uares	_	-		151.6
Date:		Fri, 13 Oct):	1.55e-32	
Time:			19:12		ikelihood:	, -	-3004.2
No. Observ	ations:		1000	AIC:			6012.
Df Residua			998	BIC:			6022.
Df Model:			1				
Covariance	Type:	nonr	obust				
=======	========		======	======	========	=======	
					P> t	[0.025	0.975]
		3 0.155			0.000	4.627	5.234
X	-1.8928	0.154	-12	2.313	0.000		
Omnibus:	=======	========	====== 2 . 933	===== Durbi	======= n-Watson:	=======	 1.960
Prob(Omnib	us):		0.231	Jarqu	e-Bera (JB):		2.749
Skew:	•		0.073	Prob(JB):		0.253
Kurtosis:			2.788	Cond.	•		1.02
========	========	========	======	======	========	========	========

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

Nonlinear Data (15%)

Generate 1000 data points with the following distributions and formula:

 $x \sim \text{Normal}(0, 1)$

 $\epsilon \sim \text{Normal}(0, 5)$

 $y = 10 + 5e^x + \epsilon$

```
In [16]: xs = random_generation.normal(loc= 0, scale= 1, size= sample_size)
    epsilon = random_generation.normal(loc= 0, scale= 5, size= sample_size)
    ys = 10 + 5*np.exp(xs) + epsilon
```

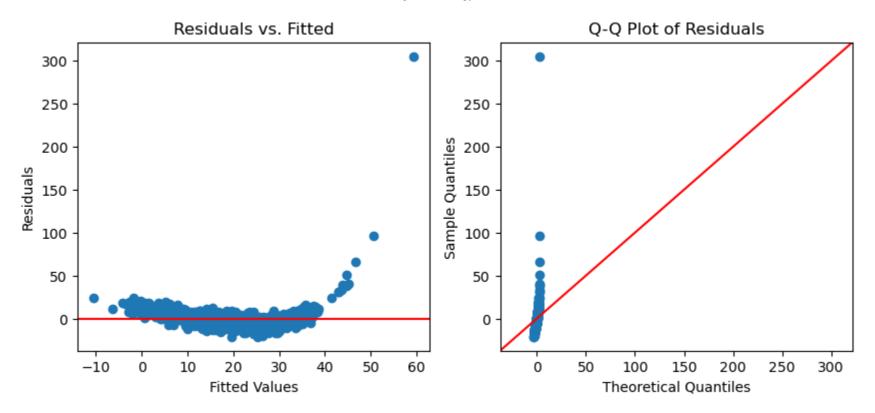
Fit a linear model predicting y with x

- How well does the model fit?
 - The R-squared value is 0.503, which indicates that approximately 50.3% of the variance in ys can be explained by xs. This suggests that the model explains only a moderate amount of the variance in the data.
 - The F-statistic is 1012 and the associated p-value is very close to zero suggesting that ys is significantly related to xs indicating that the model as a whole is statistically significant.
 - The coefficient for xs is 8.3291, and it is statistically significant and similarly the *intercept* value is 18.1970 and is also statistically significant.

This model has an R-squared value that indicates a moderate fit. While it's a linear regression model, it's possible that the relationship between ys and xs may not be perfectly linear, especially given the low p-values and relatively high coefficient values.

- Do the assumptions seem to hold?
 - From Q-Q Plot it's evident that the residuals are not normally distributed hence the assumption of Residuals are normally distributed doesn't hold true for this model.
 - Also from 'Residuals vs Fitted' plot, I can see a clear pattern between residulas around the predicted values. So the Linearity and Homoscedasticity doesn't hold true

```
In [17]: xs = sm.add constant(xs)
         ols model = sm.OLS(ys, xs).fit()
         ols summary = ols model.summary()
         fig, axes = plt.subplots(nrows=1, ncols=2, figsize=(10, 4))
         # Plot the residual vs. fitted plot
         fitted values = ols model.fittedvalues
         residuals = ols model.resid
         axes[0].scatter(fitted values, residuals)
         axes[0].set title("Residuals vs. Fitted")
         axes[0].set xlabel("Fitted Values")
         axes[0].set ylabel("Residuals")
         axes[0].axhline(y=0, color='r', linestyle='-')
         sm.qqplot(residuals, line='45', ax=axes[1])
         axes[1].set title("Q-Q Plot of Residuals")
         plt.show()
         print(ols_summary)
```



=========	=======		====	======	.========	======		
Dep. Variable:		У		R-squ	R-squared:		0.355	
Model:			OLS		Adj. R-squared:		0.355	
Method:		Least Squa	Least Squares		tistic:		550.1	
Date:		Fri, 13 Oct 2	ri, 13 Oct 2023		<pre>Prob (F-statistic):</pre>		3.01e-97	
Time:		21:19	:12	Log-I	Log-Likelihood:		-3970.8	
No. Observa	tions:	1	.000	AIC:			7946.	
Df Residual	s:		998	BIC:			7955.	
Df Model:			1					
Covariance '	Type:	nonrob	ust					
========	=======	-=======		======	=========	======	========	
	coef	std err		t	P> t	[0.025	0.975]	
const	18.4022	2 0.406	4	5.303	0.000	17.605	19.199	
x1	9.6125	0.410	2	3.455	0.000	8.808	10.417	
Omnibus:	=======	1952.	325	===== Durbi	========= .n-Watson:	======	1.942	
Prob(Omnibus):		0.000		Jarque-Bera (JB):		4357458.298		
Skew:	•	14.	225	Prob(JB):		0.00	
Kurtosis:		325.	133	Cond.	No.		1.02	
========	========		====	======	=========	======	========	

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

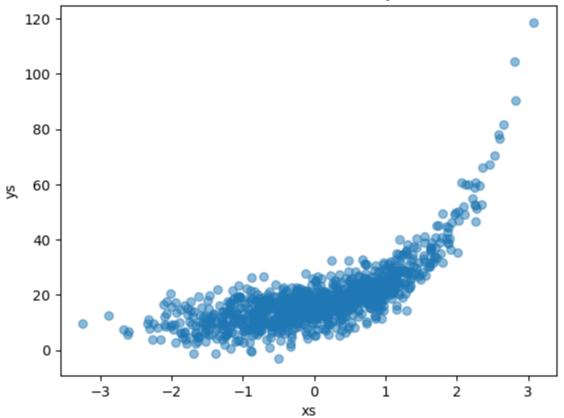
Draw a scatter plot of x and y.

```
In [18]: xs = random_generation.normal(loc= 0, scale= 1, size= sample_size)
    epsilon = random_generation.normal(loc= 0, scale= 5, size= sample_size)
    ys = 10 + 5*np.exp(xs) + epsilon

plt.scatter(xs, ys, alpha=0.5)
    plt.title('Scatter Plot of xs vs ys')
    plt.xlabel('xs')
    plt.ylabel('ys')

# Display the plot
    plt.show()
```

Scatter Plot of xs vs ys



Repeat with
$$y = -2 + 3x^3 + \epsilon$$

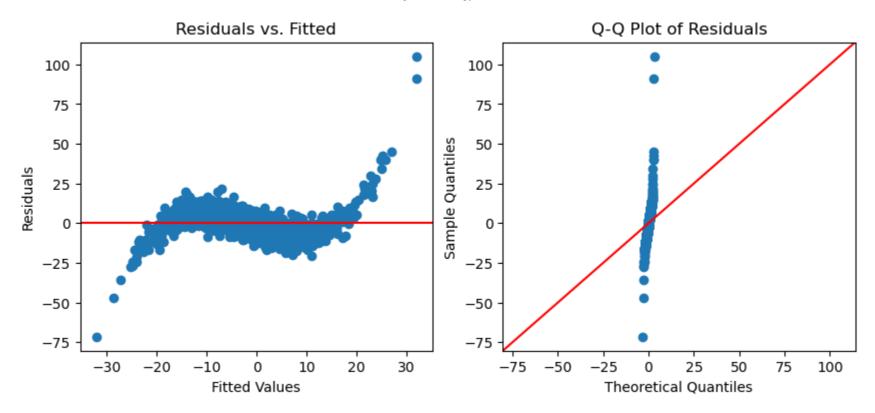
Similar to the above equation $y = 10 + 5 e^x + epsilon$, this model also doesn't hold all the assumption true and shows some non-linear relationship between ys and xs which is also evident from the scatter plotted below.

- How well does the model fit?
 - The R-squared value is 0.495, which indicates that approximately 49.5% of the variance in ys can be explained by xs. This suggests that the model explains only a moderate amount of the variance in the data.
 - The F-statistic is 978.4 and the associated p-value is very close to zero suggesting that ys is significantly related to xs indicating that the model as a whole is statistically significant.
 - The coefficient for xs is 8.861, and it is statistically significant and similarly the *intercept* value is -2.0738 and is also statistically significant.

This model has an R-squared value that indicates a moderate fit. While it's a linear regression model, it's possible that the relationship between ys and xs may not be perfectly linear, especially given the low p-values and relatively high coefficient values.

- · Do the assumptions seem to hold?
 - From Q-Q Plot it's evident that the residuals are not normally distributed hence the assumption of Residuals are normally distributed doesn't hold true for this model.
 - Also from 'Residuals vs Fitted' plot, I can see a clear pattern between residulas around the predicted values. So the Linearity and Homoscedasticity doesn't hold true

```
In [19]: xs = random generation.normal(loc= 0, scale= 1, size= sample size)
         epsilon = random generation.normal(loc= 0, scale= 5, size= sample size)
        ys = -2 + 3*np.power(xs, 3) + epsilon
         xs = sm.add constant(xs)
        ols model = sm.OLS(ys, xs).fit()
         ols summary = ols model.summary()
         fig, axes = plt.subplots(nrows=1, ncols=2, figsize=(10, 4))
         # Plot the residual vs. fitted plot
         fitted values = ols model.fittedvalues
        residuals = ols model.resid
         axes[0].scatter(fitted values, residuals)
         axes[0].set title("Residuals vs. Fitted")
         axes[0].set xlabel("Fitted Values")
         axes[0].set ylabel("Residuals")
         axes[0].axhline(y=0, color='r', linestyle='-')
         sm.qqplot(residuals, line='45', ax=axes[1])
         axes[1].set title("Q-Q Plot of Residuals")
         plt.show()
        print(ols_summary)
```

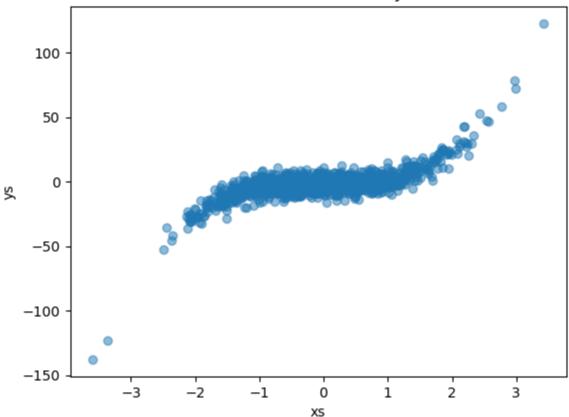


Dep. Variable: Model:			у	R-squared:			0.514	
			OLS	Adj. I	0.514			
Method:		Least Sq	uares	F-stat	F-statistic:			
Date:		Fri, 13 Oct	2023	Prob	(F-statistic):	1.07e-158	
Time:		21:	19:13	Log-Li	kelihood:		-3684.0 7372.	
No. Observa	ations:		1000	AIC:				
Df Residual	Ls:		998	BIC:			7382.	
Df Model:			1					
Covariance	Type:	nonro	obust					
========			=====	======	-=======	=======		
	coei	f std err		t	P> t	[0.025	0.975]	
const	-1.4855	0.305		 4.872	0.000	-2.084	-0.887	
x1	9.5652	0.294	32	2.510	0.000	8.988	10.143	
Omnibus:	=======	 57(====== 6.354	===== Durbir	========= n-Watson:	=======	 1.929	
Prob(Omnibu	ıs):	(0.000	Jarque	e-Bera (JB):		28318.738	
Skew:	•		1.922	_	, ,		0.00	
Kurtosis:		28	3.785	,	•		1.04	
========		=========		======		========		

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

Scatter Plot of xs vs ys



Non-Normal Covariates (15%)

Generate 1000 data points with the model:

```
y = -10 + 5x + \epsilon

\epsilon \sim \text{Normal}(0, 30)

x \sim \text{Uniform}(0, 100)
```

```
In [21]: xs = random_generation.uniform(0,100,sample_size)
    epsilon = random_generation.normal(loc= 0, scale= 30, size= sample_size)
    ys = -10 + 5*xs + epsilon
```

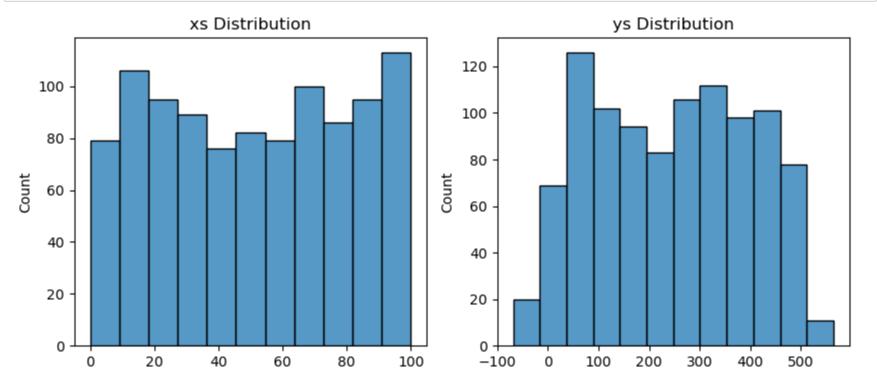
Plot the distributions of X and Y

```
In [22]: fig, axes = plt.subplots(nrows=1, ncols=2, figsize=(10, 4))

# Plot the residual vs. fitted plot
sns.histplot(xs, ax=axes[0])
axes[0].set_title("xs Distribution")

sns.histplot(ys, ax=axes[1])
axes[1].set_title("ys Distribution")

plt.show()
```

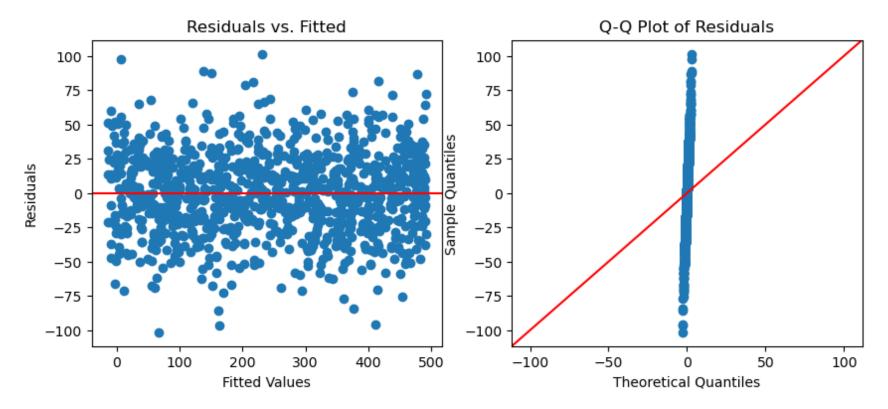


Fit a linear model predicting y with x

- How well does this model fit? How much of the variance does it explain?
 - The models explains the 95.7% variance in ys with high statistical significance of the model i.e. dependency of ys on xs.
 - The coefficient of xs (slope) is predicted to be 5.0413 quite closer to actual coefficient with a standard error of 0.034. However the standard error in the **intercept** is quite high (1.895) though the value is quite near to the actual one. Both the predicted parameters are statistically significant.

- Do the assumptions seem to hold?
 - From the Q-Q Plot there is a clear violation of assumption that Residuals are normally distributed.
 - Also, Residual vs Fitted plot exhibit some kind of linear trend between the residuals and the predicted values hence other assumptions doesn't seem hold true
- Does the linear regression seem appropriate to the data?
 - No, this linear regression doesn't seem appropriate as the Residuals are not normally distributed and also show a linear trend with the predicted value.

```
In [23]: xs = sm.add constant(xs)
         ols model = sm.OLS(ys, xs).fit()
         ols summary = ols model.summary()
         fig, axes = plt.subplots(nrows=1, ncols=2, figsize=(10, 4))
         # Plot the residual vs. fitted plot
         fitted values = ols model.fittedvalues
         residuals = ols model.resid
         axes[0].scatter(fitted values, residuals)
         axes[0].set title("Residuals vs. Fitted")
         axes[0].set xlabel("Fitted Values")
         axes[0].set ylabel("Residuals")
         axes[0].axhline(y=0, color='r', linestyle='-')
         sm.qqplot(residuals, line='45', ax=axes[1])
         axes[1].set title("Q-Q Plot of Residuals")
         plt.show()
         print(ols_summary)
```



========	:======:	=========	=====	=====	=========	=======	=======
Dep. Variab	ole:		У	R-squ	ared:		0.962
Model:			OLS	Adj.	R-squared:		0.962
Method:		Least Squ	ares	F-sta	tistic:		2.498e+04
Date:		Fri, 13 Oct	2023	Prob	(F-statistic)	:	0.00
Time:		21:1	9:13	Log-L	ikelihood:		-4817.7
No. Observa	ations:		1000	AIC:			9639.
Df Residual	.S:		998	BIC:			9649.
Df Model:			1				
Covariance	Type:	nonro	bust				
========	-=======		=====	=====	=========	=======	=======
	coei	f std err		t	P> t	[0.025	0.975]
const	-14.3934	1.892	-7	.608	0.000	-18.106	-10.681
x1	5.0600	0.032	158	.042	0.000	4.997	5.123
Omnibus:	=======	======================================	-==== .•694	===== Durbi	======== n-Watson:	=======	======= 1.959
Prob(Omnibu	ıs):	(.429	Jarqu	e-Bera (JB):		1.644
Skew:		-0	.023	Prob(JB):		0.439
Kurtosis:		3	3.193	Cond.	No.		118.
========		=========	=====	=====	=========	=======	=======

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

Generate 1000 data points with the model:

```
y = 10 + 2x + \epsilon

\epsilon \sim \text{Normal}(0, 1)

x \sim \text{Exponential}(5)
```

```
In [24]: xs = random_generation.exponential(scale= 5, size = sample_size)
    epsilon = random_generation.normal(loc= 0, scale= 1, size= sample_size)
    ys = 10 + 2*xs + epsilon
```

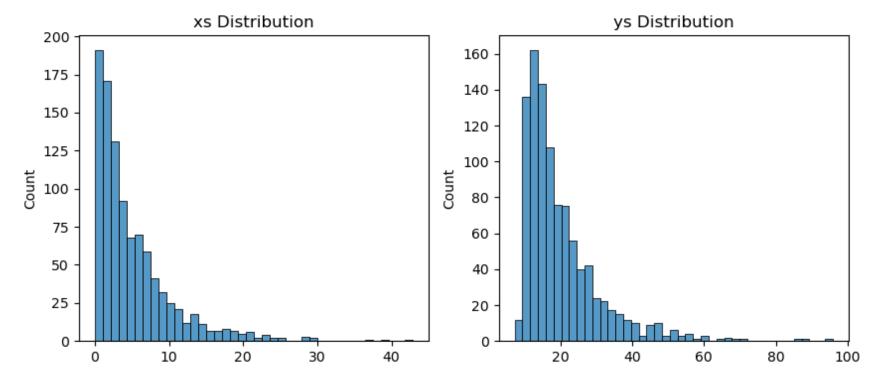
Plot the distributions of *X* and *Y*

```
In [25]: fig, axes = plt.subplots(nrows=1, ncols=2, figsize=(10, 4))

# Plot the residual vs. fitted plot
sns.histplot(xs, ax=axes[0])
axes[0].set_title("xs Distribution")

sns.histplot(ys, ax=axes[1])
axes[1].set_title("ys Distribution")

plt.show()
```

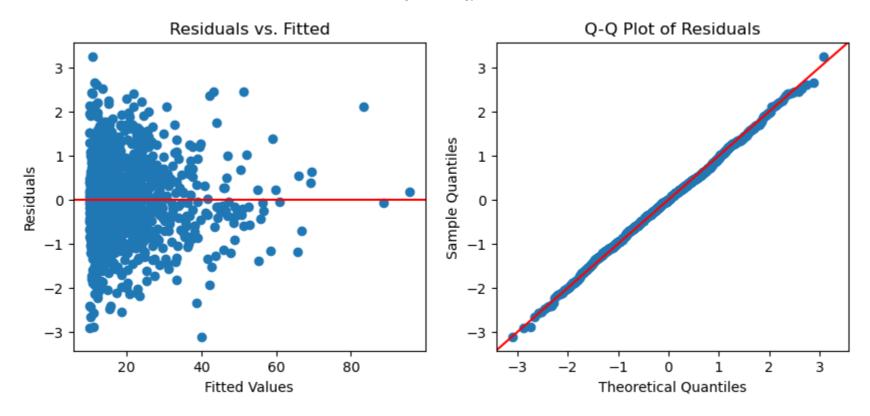


Fit a linear model predicting y with x

- How well does this model fit? How much of the variance does it explain?
 - The models explains the 98.9% variance in ys with high statistical significance of the model i.e. dependency of ys on xs.
 - The coefficient of xs (slope) is predicted to be 1.9896 quite closer to actual coefficient with a standard error of 0.007 which is quite low. The **intercept** also is close to the actual value with substantially a low standard error. Both of the predicted parameters are statistically significant

- Do the assumptions seem to hold?
 - The Q-Q Plot doesn't show much deviation from the diagnoal line, so it seems like the residuals are normally distributed.
 - However the Residual vs Fitted plot exhibit a clear trend in residuals (all residuals are skewed to the left) thus exhibiting the
 relationship between residuals. Hence, Homoscedasticity assumption doesn't hold true and also the residuals doesn't seem
 independent as well
- Does the linear regression seem appropriate to the data?
 - No, this linear regression doesn't seem appropriate as the Residuals are not normally distributed and also show a linear trend with the predicted value. Although the fitted model explains a significant variance in the ys

```
In [26]: xs = sm.add constant(xs)
         ols model = sm.OLS(ys, xs).fit()
         ols summary = ols model.summary()
         fig, axes = plt.subplots(nrows=1, ncols=2, figsize=(10, 4))
         # Plot the residual vs. fitted plot
         fitted values = ols model.fittedvalues
         residuals = ols model.resid
         axes[0].scatter(fitted values, residuals)
         axes[0].set title("Residuals vs. Fitted")
         axes[0].set xlabel("Fitted Values")
         axes[0].set ylabel("Residuals")
         axes[0].axhline(y=0, color='r', linestyle='-')
         sm.qqplot(residuals, line='45', ax=axes[1])
         axes[1].set title("Q-Q Plot of Residuals")
         plt.show()
         print(ols_summary)
```



OLS Regression Results

=========			=====	======			
Dep. Variab	le:		У	R-squ	ared:		0.992
Model:			OLS	Adj.	R-squared:		0.992
Method:		Least Squ	ares	F-sta	tistic:		1.257e+05
Date:		Fri, 13 Oct	2023	Prob	(F-statistic)	:	0.00
Time:		21:	19:13	Log-I	ikelihood:		-1380.4
No. Observat	tions:		1000	AIC:			2765.
Df Residuals	5:		998	BIC:			2775.
Df Model:			1				
Covariance Type:		nonro	bust				
========				======			
	coet	f std err		t	P> t	[0.025	0.975]
const	10.0064	0.042	23	 9 . 508	0.000	9.924	10.088
x1	1.9973	0.006	35	4.606	0.000	1.986	2.008
Omnibus:	=======	:=======)	===== 0.801	===== Durbi	======== .n-Watson:	=======	 2.020
Prob(Omnibus	s):	(0.670	Jarqu	ue-Bera (JB):		0.672
Skew:	•	_(0.019	Prob(JB):		0.715
Kurtosis:		;	3.121	Cond.	No.		10.3
=========	========		=====	======			

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

Multiple Regression (10%)

$$x_1 \sim \text{Normal}(10, 2)$$

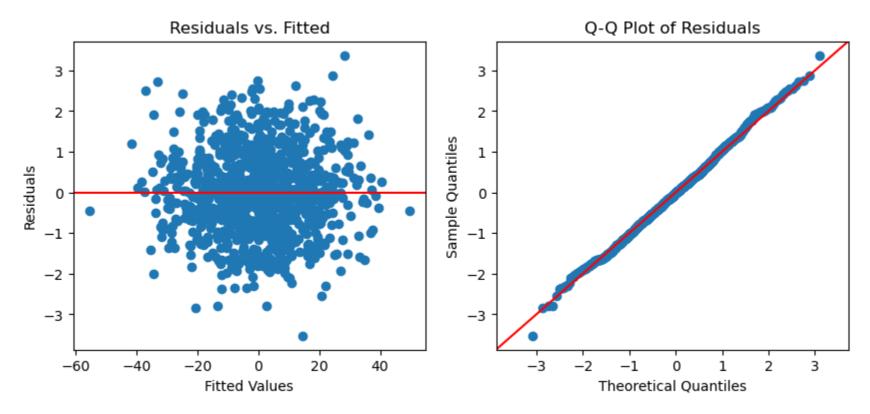
 $x_2 \sim \text{Normal}(-2, 5)$
 $\epsilon \sim \text{Normal}(0, 1)$
 $y = 1 + 0.5x_1 + 3x_2 + \epsilon$

- What are the intercept and coefficients from the model? Are they what you expect?
 - The model explains 99.6% variance in the ys variable and the model is statistically significant as well.
 - The predicted values for the parameters are as below:
 - x1 --> 0.5195 with standard error 0.016. The actual value is .5
 - $\circ \quad x2 \ \mbox{-->} 2.9928$ with standard error 0.016. The actual value is 3
 - intercept --> 0.8192 with standard error 0.162. The actual value is 1

The predicted values for x1 and x2 are quite close to the actual values and all the predicted parameters are statistically significant. The intecept predicted value is quite far from the actual value with significantly higher standard error as well. This might be an indication for that model shows dependency on x1 & x2 more rather than the intecept

- Check the model assumptions do they hold?
 - All the assumptions for this model seems to hold true as *Q-Q Plot* shows no deviation and *Residual vs Fitted* plot also doens't show any sign that might indicate any assumption being violated.

```
In [27]: xs1 = random generation.normal(loc= 10, scale= 2, size= sample size)
         xs2 = random generation.normal(loc= -2, scale= 5, size= sample size)
         epsilon = random generation.normal(loc= 0, scale= 1, size= sample size)
         ys = 1 + 0.5*xs1 + 3*xs2 + epsilon
         X = sm.add constant(np.column stack((xs1, xs2)))
         ols model = sm.OLS(ys, X).fit()
         ols summary = ols model.summary()
         fig, axes = plt.subplots(nrows=1, ncols=2, figsize=(10, 4))
         # Plot the residual vs. fitted plot
         fitted values = ols model.fittedvalues
         residuals = ols model.resid
         axes[0].scatter(fitted values, residuals)
         axes[0].set title("Residuals vs. Fitted")
         axes[0].set xlabel("Fitted Values")
         axes[0].set ylabel("Residuals")
         axes[0].axhline(y=0, color='r', linestyle='-')
         sm.qqplot(residuals, line='45', ax=axes[1])
         axes[1].set title("Q-Q Plot of Residuals")
         plt.show()
         print(ols summary)
```



OLS Regression Results

=========	======	========	======	=====		=======	=======
Dep. Variable	:		У	R-sq	uared:		0.995
Model:			OLS	Adj.	R-squared:		0.995
Method:		Least Sq	uares	F-sta	atistic:		1.089e+05
Date:		Fri, 13 Oct	2023	Prob	(F-statistic)	:	0.00
Time:		21:	19:13	Log-	Likelihood:		-1440.5
No. Observati	ons:		1000	AIC:			2887.
Df Residuals:			997	BIC:			2902.
Df Model:			2				
Covariance Ty	pe:	nonre	obust				
=========	======	========	======	=====		=======	=======
	coe	f std err		t	P> t	[0.025	0.975]
const	1.0962	2 0.167	 6	5.577	0.000	0.769	1.423
x1	0.4842	0.016	29	.608	0.000	0.452	0.516
x2	2.9910	0.006	466	.590	0.000	2.978	3.004
Omnibus:	======	 :	===== 2.069	Durb:	========= in-Watson:	=======	======== 2.167
Prob(Omnibus)	:		0.355		ue-Bera (JB):		2.142
Skew:			0.103	Prob	, ,		0.343
Kurtosis:			2.904	Cond	` '		54.4
=========	======	========	======	:====:		=======	=======

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

Correlated Predictors (10%)

```
In [28]: def correlated predictors(x1 mean, x2 mean, cov, num sample, only plot pairs = False, is single run = Fal
                                  random generaton = np.random.default rng(20201014)
                                  cov matrix = np.array([[1, cov], [cov, 1]])
                                  means = np.array([x1 mean, x2 mean])
                                  xs = random generation.multivariate normal(means, cov matrix, num sample)
                                  df = pd.DataFrame(xs, columns=['X1', 'X2'])
                                 df['epsilon'] = random generation.normal(loc= 0, scale= 2, size= sample size)
                                 df['Y'] = 3 + 2*df['X1'] + 3*df['X2'] + df['epsilon']
                                  if only plot pairs:
                                            sns.pairplot(df, vars=['X1', 'X2', 'Y'], diag kind='kde')
                                            plt.show()
                                            return None
                                 Y = df['Y']
                                  df = sm.add constant(df[['X1','X2']])
                                  ols model = sm.OLS(Y, df).fit()
                                  ols summary = ols model.summary()
                                  if is single run:
                                            print(ols summary)
                                            fig, axes = plt.subplots(nrows=1, ncols=2, figsize=(10, 4))
                                            # Plot the residual vs. fitted plot
                                            fitted values = ols model.fittedvalues
                                            residuals = ols model.resid
                                            axes[0].scatter(fitted values, residuals)
                                            axes[0].set title("Residuals vs. Fitted")
                                            axes[0].set xlabel("Fitted Values")
                                            axes[0].set ylabel("Residuals")
                                            axes[0].axhline(y=0, color='r', linestyle='-')
                                            sm.qqplot(residuals, line='45', ax=axes[1])
                                            axes[1].set title("Q-Q Plot of Residuals")
                                            plt.show()
                                            return None
                                  else:
                                            x1 coeff pred = ols model.params['X1']
```

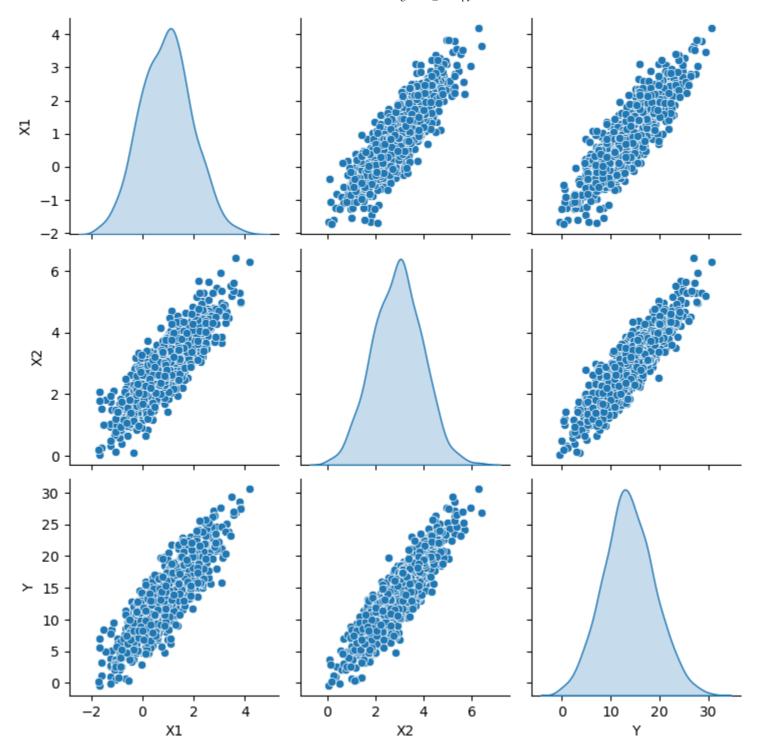
```
x2_coeff_pred = ols_model.params['X2']
intercept_pred = ols_model.params['const']
return x1_coeff_pred, x2_coeff_pred, intercept_pred
```

$$y = 3 + 2x_1 + 3x_2 + \epsilon$$

Show a pairplot of our variables X_1 , X_2 , and Y. What do we see about their distributions and relationships?

- Below graphs are the pairplot between (X1,X2,&Y).
- The pair of the variables (X1,X2), (X1,Y), (X2,Y) seems to be linearly correlated. These pairs seem to be positively correlated i.e. increase in one will increase in the other variable value.
- The individual variable as expected is normally distributed.

In [29]: correlated_predictors(1,3,.85, sample_size, only_plot_pairs=True)



Fit a linear regression for $y \sim x1 + x2$. How well does it fit? Do its assumptions hold?

- What are the intercept and coefficients from the model? Are they what you expect?
 - The model explains 86.7% variance in the ys variable and the model is statistically significant as well.
 - The predicted values for the parameters are as below:
 - x1 --> 1.9533 with standard error 0.120. The actual value is 2
 - x2 --> 3.0535 with standard error 0.120. The actual value is 3
 - intercept --> 2.9186 with standard error 0.272. The actual value is 3

The predicted values for x1 and x2 are quite close to the actual values and all the predicted parameters are statistically significant. The intecept predicted value is quite far from the actual value with significantly higher standard error as well. This might be an indication for that model shows dependency on x1 & x2 more rather than the intecept

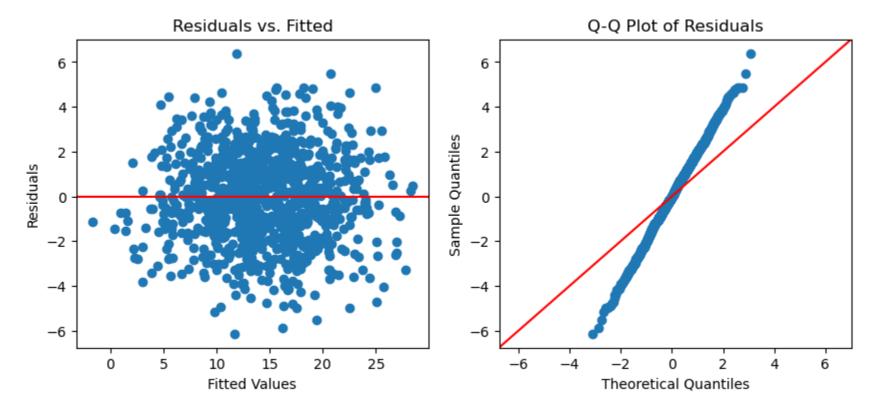
- Check the model assumptions do they hold?
 - As the residuals in 'Q-Q Plot' are deviated from the diagonal, it seems like the residuals are not normally distributed. Hence, the assumption of Residuals are Normally distributed fails to hold.
 - There are some outliers in 'Residuals vs Fitted' plot however other than this, no other observable trend is seen. It seems like the Homoscedasticity assumption holds well

In [30]: correlated_predictors(1,3,.85, sample_size, is_single_run=True)

		OLS R	egress	sion Re	esults 		
Dep. Variable Model: Method: Date: Time: No. Observati Df Residuals: Df Model: Covariance Ty	ons:		2023 9:14 1000 997 2	Adj. F-sta Prob	uared: R-squared: atistic: (F-statistic Likelihood:):	
========	coef	std err	=====	t	P> t	[0.025	0.975]
const X1 X2	2.9186 1.9533 3.0535	0.120	16	720 3.331 5.522	0.000 0.000 0.000	2.384 1.719 2.819	3.453 2.188 3.288
Omnibus: Prob(Omnibus) Skew: Kurtosis:	:	0 -0	.807 .668 .048	Jarqı Prob	in-Watson: ie-Bera (JB): (JB): . No.	=======	2.057 0.891 0.640 17.8

Notes:

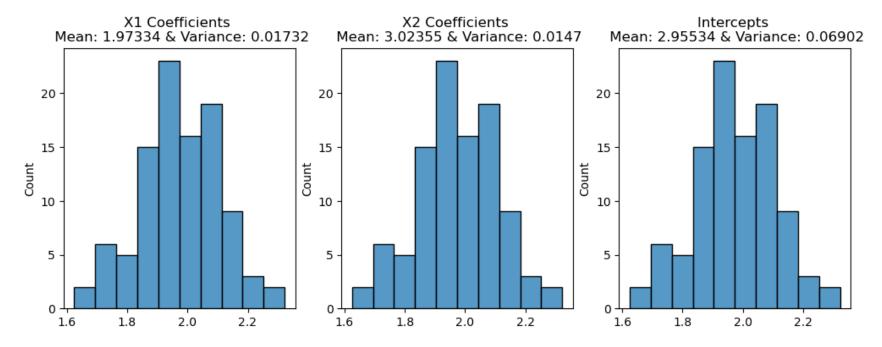
[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.



Run this simulation (drawing 1000 variables and fitting a linear model) 100 times. Show the mean, variance, and appropriate distribution plots of the estimated intercepts and coefficients (for x1 and x2)

• Below graphs show the distribution of the coefficients of X1 & X2 and the intercept. The distribution shows significantly higher variance in the intercept as compare to X1 & X2

```
In [31]: x1 coeffs = []
         x2 coeffs = []
         intercepts = []
         for i in range(100):
             x1_coeff_pred, x2_coeff_pred, intercept_pred = correlated_predictors(1,3,.85, sample_size)
             x1 coeffs.append(x1 coeff pred)
             x2_coeffs.append(x2 coeff pred)
             intercepts.append(intercept pred)
         x1 pred mean = round(np.mean(x1 coeffs), 5)
         x1 pred var = round(np.square(np.std(x1 coeffs)), 5)
         x2 \text{ pred mean} = round(np.mean(x2 coeffs), 5)
         x2 pred var = round(np.square(np.std(x2 coeffs)), 5)
         intercept mean = round(np.mean(intercepts), 5)
         intercept var = round(np.square(np.std(intercepts)), 5)
         fig, axes = plt.subplots(nrows=1, ncols=3, figsize=(12, 4))
         sns.histplot(x1 coeffs, ax=axes[0])
         axes[0].set title(f'X1 Coefficients \n Mean: {x1 pred mean} & Variance: {x1 pred var}')
         sns.histplot(x1 coeffs, ax=axes[1])
         axes[1].set title(f'X2 Coefficients \n Mean: {x2 pred mean} & Variance: {x2 pred var}')
         sns.histplot(x1 coeffs, ax=axes[2])
         axes[2].set title(f'Intercepts \n Mean: {intercept mean} & Variance: {intercept var}')
         plt.show()
```



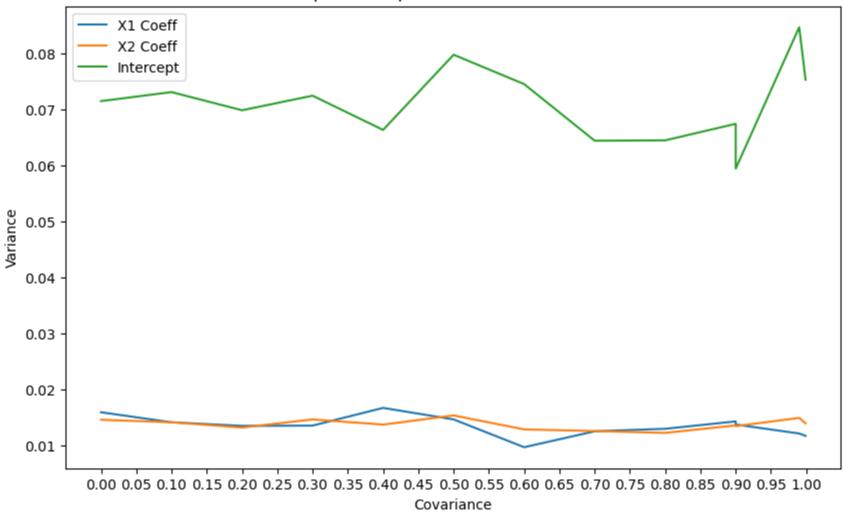
Repeat the repeated simulation for a variety of different covariances from 0 to 1 (including at least 0, 0.9, 0.99, and 0.999). Create line plots (or a single line plot with multiple colors) that show how the variance of the estimated regression parameters (intercept and x_1 and x_2 coefficients) change as you increase the correlation (covariance) between X_1 and X_2 .

- From the below graph, it's also clearly evident that the intercept has the highest variance compared X1 & X2 and is quite sensitive to the covariance between X1 & X2.
- It seems like the variance in all the predicted parameters is less for covariance values between [0.20 0.80]

```
In [32]: | cov plot dict = {}
         cov plot dict['covariance'] = []
         cov plot dict['x1 variance'] = []
         cov plot dict['x2 variance'] = []
         cov plot dict['intercept'] = []
         base case = [0,0.9,.99,0.999]
         for cov in base case:
             cov plot dict['covariance'].append(cov)
             temp x1 = []
             temp x2 = []
             temp intercept = []
             for i in range(100):
                 x1 coeff pred, x2 coeff pred, intercept pred = correlated predictors(1,3,.85, sample size)
                 temp x1.append(x1 coeff pred)
                 temp x2.append(x2 coeff pred)
                 temp intercept.append(intercept pred)
             cov plot dict['x1 variance'].append( round(np.square(np.std(temp x1)),5) )
             cov plot dict['x2 variance'].append( round(np.square(np.std(temp x2)),5) )
             cov plot dict['intercept'].append( round(np.square(np.std(temp intercept)),5) )
         start = 0.10
         end = .95
         step = 0.10
         next covs = []
         current = start
         while current <= end:</pre>
             next covs.append(current)
             current += step
         for cov in next covs:
             cov plot dict['covariance'].append(cov)
             temp x1 = []
             temp x2 = []
             temp intercept = []
             for i in range(100):
                 x1 coeff pred, x2 coeff pred, intercept pred = correlated predictors(1,3,.85, sample size)
                 temp x1.append(x1 coeff pred)
```

```
temp x2.append(x2 coeff pred)
        temp intercept.append(intercept pred)
    cov plot dict['x1 variance'].append( round(np.square(np.std(temp x1)),5) )
    cov plot dict['x2 variance'].append( round(np.square(np.std(temp x2)),5) )
    cov plot dict['intercept'].append( round(np.square(np.std(temp intercept)),5) )
df plot = pd.DataFrame(cov plot dict).sort values(by='covariance')
plt.figure(figsize=(10, 6))
sns.lineplot(x="covariance", y="x1 variance", data=df plot, label="X1 Coeff")
sns.lineplot(x="covariance", y="x2 variance", data=df plot, label="X2 Coeff")
sns.lineplot(x="covariance", y="intercept", data=df plot, label="Intercept")
plt.xlabel("Covariance")
plt.ylabel("Variance")
plt.title("Variance in predicted parameters at various Covariances")
plt.legend()
plt.xticks([0.00, 0.05, 0.10, 0.15, 0.20, 0.25, 0.30, 0.35, 0.40, 0.45, 0.50, 0.55, 0.60, 0.65, 0.70, 0
plt.show()
```

Variance in predicted parameters at various Covariances

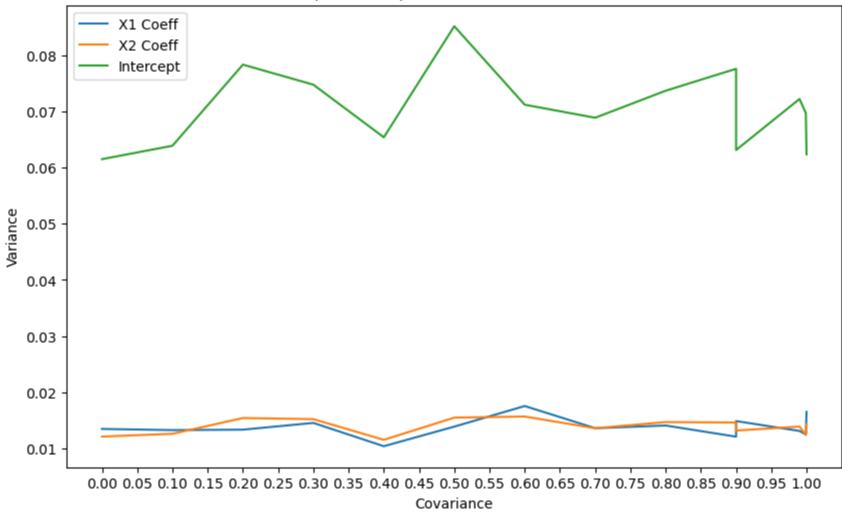


- If I include covariance = 1
 - The variance for the all the 3 parameters was decreasing till 0.999 however at 1, I see a sudden rise.
 - The covariance 1 means the model is still significant, however the predicted parameters are not significant at all. It seems like, if the model will fit Y with either X1 or X2 only but both makes the parameter values insignificant

```
In [33]: | cov plot dict = {}
         cov plot dict['covariance'] = []
         cov plot dict['x1 variance'] = []
         cov plot dict['x2 variance'] = []
         cov plot dict['intercept'] = []
         base case = [0,0.9,.99,0.999]
         for cov in base case:
             cov plot dict['covariance'].append(cov)
             temp x1 = []
             temp x2 = []
             temp intercept = []
             for i in range(100):
                 x1 coeff pred, x2 coeff pred, intercept pred = correlated predictors(1,3,.85, sample size)
                 temp x1.append(x1 coeff pred)
                 temp x2.append(x2 coeff pred)
                 temp intercept.append(intercept pred)
             cov plot dict['x1 variance'].append( round(np.square(np.std(temp x1)),5) )
             cov plot dict['x2 variance'].append( round(np.square(np.std(temp x2)),5) )
             cov plot dict['intercept'].append( round(np.square(np.std(temp intercept)),5) )
         start = 0.10
         end = 1
         step = 0.10
         next covs = []
         current = start
         while current <= end:</pre>
             next covs.append(current)
             current += step
         for cov in next covs:
             cov plot dict['covariance'].append(cov)
             temp x1 = []
             temp x2 = []
             temp intercept = []
             for i in range(100):
                 x1 coeff pred, x2 coeff pred, intercept pred = correlated predictors(1,3,.85, sample size)
                 temp x1.append(x1 coeff pred)
```

```
temp x2.append(x2 coeff pred)
        temp intercept.append(intercept pred)
    cov plot dict['x1 variance'].append( round(np.square(np.std(temp x1)),5) )
    cov plot dict['x2 variance'].append( round(np.square(np.std(temp x2)),5) )
    cov plot dict['intercept'].append( round(np.square(np.std(temp intercept)),5) )
df plot = pd.DataFrame(cov plot dict).sort values(by='covariance')
plt.figure(figsize=(10, 6))
sns.lineplot(x="covariance", y="x1 variance", data=df plot, label="X1 Coeff")
sns.lineplot(x="covariance", y="x2 variance", data=df plot, label="X2 Coeff")
sns.lineplot(x="covariance", y="intercept", data=df plot, label="Intercept")
plt.xlabel("Covariance")
plt.ylabel("Variance")
plt.title("Variance in predicted parameters at various Covariances")
plt.legend()
plt.xticks([0.00, 0.05, 0.10, 0.15, 0.20, 0.25, 0.30, 0.35, 0.40, 0.45, 0.50, 0.55, 0.60, 0.65, 0.70, 0
plt.show()
```

Variance in predicted parameters at various Covariances

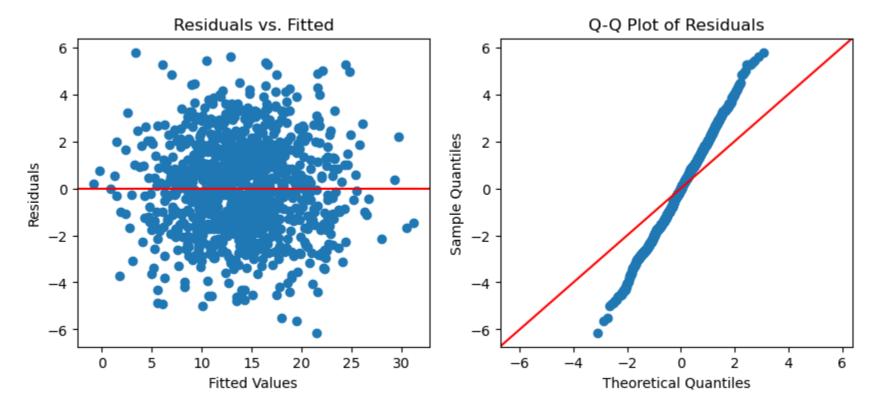


In [34]: correlated_predictors(1,3,1, sample_size, is_single_run=True)

OLS Regression Results								
Dep. Variable:	Y	R-squ	uared:		0.859			
Model:	OLS	Adj.	R-squared:		0.859			
Method:	Least Squares		atistic:		3038.			
Date:	Fri, 13 Oct 2023	Prob	(F-statist	ic):	0.00			
Time:	21:19:20	Log-I	Likelihood:	,	-2112.3			
No. Observations:	1000	AIC:			4231.			
Df Residuals:	997	BIC:			4245.			
Df Model:	2							
Covariance Type:	nonrobust							
=======================================	=========	======	-=======	========	========			
coe	f std err	t	P> t	[0.025	0.975]			
const 4.754e+0	6 1.54e+07	0.308	0.758	-2.55e+07	3.5e+07			
X1 2.377e+0	6 7.72e+06	0.308	0.758	-1.28e+07	1.75e+07			
X2 -2.377e+0	6 7.72e+06	-0.308	0.758	-1.75e+07	1.28e+07			
Omnibus:		===== Durbi	======= in-Watson:	========	2.055			
Prob(Omnibus):	0.426	Jarqu	ıe-Bera (JB):	1.707			
Skew:	0.056	Prob((JB):	•	0.426			
Kurtosis:	2.832	Cond.	No.		1.05e+09			

Notes:

- [1] Standard Errors assume that the covariance matrix of the errors is correctly specified.
- [2] The smallest eigenvalue is 1.12e-14. This might indicate that there are strong multicollinearity problems or that the design matrix is singular.



Reflection (5%)

This assignment was quite exhaustive about Linear Regression and validating whether the assumptions of the linear regression holds or not for the fitted model. From the various model fitted in this assignment, I learnt that even if the model seemes to fit perfectly and explains majority of the variances in the dependent variable, still the Linear Regression might not be suitable model as the model may be violating the fundamental assumptions of the Linear Regression. So, it becomes imperative to look at the various output graphs/output such as *Q-Q Plot*, *Residuals vs Fitted Plot* to validate whether we can use the Linear Regression or not.

From this work, I also learnt how to use and generate the random variables from a given distribution and the importance of the seed. If I need to make the results reproducible, it's very important to provide a seed.

Also, I thought I would take maybe slightly more time as mentioned in the assignment, however I think I end up spending ~20 hours in this assignment almost double the said time.