### Lesson 2: Simulation of stochastic dynamic models

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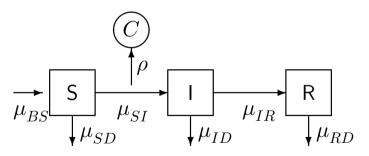
## **Objectives**

This tutorial develops some classes of dynamic models relevant to biological systems, especially for epidemiology.

- 1. Dynamic systems can often be represented in terms of *flows* between *compartments*.
- 2. We develop the concept of a *compartment model* for which we specify *rates* for the flows between compartments.
- 3. We show how deterministic and stochastic versions of a compartment model are derived and related.
- 4. We introduce Euler's method to simulate from dynamic models.
- 5. We specify deterministic and stochastic compartment models in pomp using Euler method simulation.

### A basic compartment model: The SIR model I

- We develop deterministic and stochastic representations of a susceptible-infected-recovered (SIR) system, a fundamental class of models for disease transmission dynamics.
- We set up notation applicable to general compartment models (Bretó et al. 2009).



S : susceptible I : infected and infectious

R : recovered and/or removed C : reported cases

## A basic compartment model: The SIR model II

- We suppose that each arrow has an associated rate, so here there is a rate  $\mu_{SI}(t)$  at which individuals in S transition to I, and  $\mu_{IR}$  at which individuals in I transition to R.
- To account for demography (births/deaths/migration) we allow the possibility of a source and sink compartment, which is not usually represented on the flow diagram. We write  $\mu_{BS}$  for a rate of births into S, and denote mortality rates by  $\mu_{SD}$ ,  $\mu_{ID}$ ,  $\mu_{RD}$ .
- The rates may be either constant or time-varying.
- For the simplest SIR model, ignoring demography, we set

$$\mu_{BS} = \mu_{SD} = \mu_{ID} = \mu_{RD} = 0.$$

### General notation for compartment models I

To develop a systematic notation, it turns out to be convenient to keep track of the flows between compartments as well as the number of individuals in each compartment:

- $igwedge N_{SI}(t)$ : the number of individuals who have transitioned from S to I by time t. We say that  $N_{SI}(t)$  is a *counting process*.
- $ightharpoonup N_{IR}(t)$ : the number of individuals transitioning from I to R by time t.

To include demography, we could keep track of birth and death events by the counting processes:

- $ightharpoonup N_{BS}(t)$ : the number of newborns into S by time t.
- $N_{SD}(t)$ ,  $N_{ID}(t)$ ,  $N_{RD}(t)$ : the number of deaths from S, I, and R compartments **by** time t, respectively.

### General notation for compartment models II

- For discrete population compartment models, the flow counting processes are non-decreasing and integer valued.
- For continuous population compartment models, the flow counting processes are non-decreasing and real valued.

## Compartment model from counting processes

▶ The numbers of people in each compartment can be computed via these counting processes. Ignoring demography, we have:

$$\begin{split} S(t) &= S(0) - N_{SI}(t) \\ I(t) &= I(0) + N_{SI}(t) & -N_{IR}(t) \\ R(t) &= R(0) & +N_{IR}(t) \end{split}$$

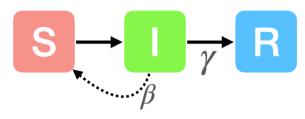
▶ These equations represent *conservation of individuals* or *what goes in must come out*.

## Ordinary differential equation interpretation

Together with initial conditions specifying S(0), I(0) and R(0), we just need to write down ordinary differential equations (ODEs) for the flow counting processes. These are:

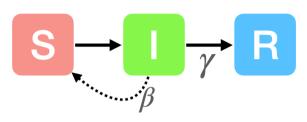
$$\begin{split} \frac{\mathrm{d}N_{SI}}{\mathrm{d}t} &= \mu_{SI}(t)\,S(t)\\ \frac{\mathrm{d}N_{IR}}{\mathrm{d}t} &= \mu_{IR}\,I(t) \end{split}$$

### Common notation for a deterministic SIR model



- $\beta$ : transmission rate, encompasses the frequency of contacts and transmission probability between individuals
- $ightharpoonup \gamma$ : recovery rate, rate that infected individuals become "uninfectious"
  - ightharpoonup Duration of infectiousness on average is  $\frac{1}{\gamma}$
- $\triangleright$  S+I+R=N

# Common notation for a deterministic SIR model - equations



$$\begin{split} \frac{\mathrm{d}S}{\mathrm{d}t} &= -\beta S \frac{I}{N} \\ \frac{\mathrm{d}I}{\mathrm{d}t} &= \beta S \frac{I}{N} - \gamma * I \\ \frac{\mathrm{d}R}{\mathrm{d}t} &= \gamma I \end{split}$$

# Stochastic Differential Equations (SDEs)

- ▶ By including randomness in the ODE system, we can have the stochastic differential equation (SDE) system.
- lackbox For example, for the ODE  $rac{\mathrm{d}x}{\mathrm{d}t}=h(x)$ , a natural way to add stochastic variation is

$$\frac{\mathrm{d}X}{\mathrm{d}t} = h(X) + \sigma \, \frac{\mathrm{d}B}{\mathrm{d}t}$$

where  $\{B(t)\}$  is Brownian motion and so dB/dt is Brownian noise.

# The simple counting process and the reactions I

- A deterministic SIR model has a fixed trajectory, indicating that the number of each compartment at any time is fixed with given parameters and intial states; thus the transitions between compartments are fixed at any time.
- A stochastic SIR model, in the contrary, the trajectory and the transitions between compartments at any time are stochastic.
- Recall  $N_{SI}(t)$  and  $N_{IR}(t)$  are counting processes, indicating the number of total individuals transitioned from S to I and I to R by time t, respectively.
- A *simple counting process* is one which cannot count more than one event at a time.

### The simple counting process and the reactions II

- ▶ We then can relate the counting process to the common SIR reactions with the corresponding probabilities.
- Note that we are using *little o notation* and we write  $h(\delta) = o(\delta)$  to mean  $\lim_{\delta \to 0} \frac{h(\delta)}{\delta} = 0$ .

Table 1: Relationship between the counting processes, the reactions, and the probabilities.

Counting	Reaction	Probability
$\overline{N_{SI}(t+\delta) = N_{SI}(t) + 1}$	$S \to S - 1$ $I \to I + 1$	$\beta S(t)I(t)\delta/N + o(\delta)$
$N_{SI}(t+\delta) = N_{SI}(t)$ $N_{IR}(t+\delta) = N_{IR}(t) + 1$	$I \rightarrow I - 1$	$1 - \beta S(t)I(t)\delta/N + o(\delta)$ $\gamma I(t)\delta + o(\delta)$
$N_{IR}(t+\delta) = N_{IR}(t)$	$R \to R + 1$	$1 - \gamma I(t)\delta + o(\delta)$

#### The Euler's method I

When referring the counting and its corresponding probability in Table 1, it is obvious that we can derive a continuous time Markov chain (CTMC) for the SIR model:

$$\begin{split} &\mathbb{P}\big[N_{SI}(t+\delta) = N_{SI}(t) + 1\big] &= \beta\,S(t)\,I(t)/N\,\delta + o(\delta), \\ &\mathbb{P}\big[N_{IR}(t+\delta) = N_{IR}(t) + 1\big] &= \gamma\,I(t)\delta + o(\delta). \end{split}$$

For k=1,2,..., by discretizing this CTMC with small time step  $\delta$ , we can derive a numerical solution with the state variables  $\tilde{S}(k\delta)$ ,  $\tilde{I}(k\delta)$ ,  $\tilde{R}(k\delta)$ :

$$\begin{array}{lcl} \tilde{S}(k\delta) & = & S(0) - \tilde{N}_{SI}(k\delta) \\ \tilde{I}(k\delta) & = & I(0) + \tilde{N}_{SI}(k\delta) - \tilde{N}_{IR}(k\delta) \\ \tilde{R}(k\delta) & = & R(0) + \tilde{N}_{IR}(k\delta) \end{array}$$

 $\blacktriangleright \ \tilde{N}_{SI}(t)$  and  $\tilde{N}_{IR}(t)$ : the numerical solutions for  $N_{SI}(t)$  and  $N_{IR}(t)$ 

### The Fuler's method II

- Let current  $t = k\delta$ , consider the small time interval  $t \le \tau \le t + \delta$ .
- Assume that the gradients  $\frac{\mathrm{d}N_{SI}}{\mathrm{d}t}=\mu_{SI}(t)\,S(t)$  and  $\frac{\mathrm{d}\overline{N}_{IR}}{\mathrm{d}t}=\mu_{IR}\,I(t)$  are approximately constant
- We can have  $\tilde{N}_{SI}(t+\delta)$  and  $\tilde{N}_{IR}(t+\delta)$  as:

$$\begin{array}{lcl} \tilde{N}_{SI}(t+\delta) & = & \tilde{N}_{SI}(t) + \delta\,\beta\,S(t)\,I(t)\,/N \\ \tilde{N}_{IR}(t+\delta) & = & \tilde{N}_{IR}(t) + \delta\,\gamma\,I(t) \end{array}$$

### The Euler's method III

Now we can include stochastic variation in the Euler's method.

► Recall the SDE:

$$\frac{\mathrm{d}X}{\mathrm{d}t} = h(X) + \sigma \, \frac{\mathrm{d}B}{\mathrm{d}t}$$

where  $\{B(t)\}$  is Brownian motion and so dB/dt is Brownian noise.

An Euler approximation  $\tilde{X}(t)$  within the small time interval  $[t,t+\delta]$  and  $t=k\delta$  for k=0,1,2,... is

$$\tilde{X}\big(\,t+\delta\,\big) = \tilde{X}(t) + \delta\,h\big(\,\tilde{X}(t)\,\big) + \sigma\sqrt{\delta}\,Z_k$$

where  $Z_1,Z_2,\ldots$  are independent standard normal random variables, i.e.,  $Z_k \sim {\rm Normal}\,(0,1).$ 

▶ Although SDEs are often considered an advanced topic in probability, the Euler approximation doesn't demand much more than familiarity with the normal distribution.

#### The Euler's method IV

Now we can consider applying the Euler's method for a stochastic SIR model:

A binomial approximation with exponential transition probabilities.

$$\begin{split} \tilde{N}_{SI}(t+\delta) &= \tilde{N}_{SI}(t) + \text{Binomial}\big[\tilde{S}(t), 1 - \exp\big\{-\beta\,\tilde{I}(t)/N\,\delta\big\}\big], \\ \tilde{N}_{IR}(t+\delta) &= \tilde{N}_{IR}(t) + \text{Binomial}\big[\tilde{I}(t), 1 - \exp\big\{-\delta\,\gamma\big\}\big], \end{split}$$

where  $\operatorname{Binomial}(n,p)$  is a binomial random variable with mean np and variance np(1-p). Here,  $p=1-\exp\big\{-\beta\,\tilde{I}(t)/N\,\delta\big\}$  and  $p=1-\exp\big\{-\delta\,\gamma\big\}$ , respectively.

#### The Euler's method V

The following are two other ways for a stochastic SIR model with the Euler's approximation, what they are not as good as the previous one?

1. A Poisson approximation.

$$\tilde{N}_{SI}(t+\delta) = \tilde{N}_{SI}(t) + \text{Poisson}[\beta \, \tilde{S}(t) \, \tilde{I}(t)/N \, \delta],$$

where  $\mathrm{Poisson}(\mu)$  is a Poisson random variable with mean  $\mu=\beta\,\tilde{S}(t)\,\tilde{I}(t)/N\,\delta.$ 

2. A binomial approximation,

$$\tilde{N}_{SI}(t+\delta) = \tilde{N}_{SI}(t) + \text{Binomial} \big[\tilde{S}(t), \beta\,\tilde{I}(t)/N\,\delta\big].$$

### The Gillespie method I

- Numerical methods, such as the Euler's method, are approximations to the process by discretizing time using small time step  $\delta$
- ► However, the Gillespie method is the exact **Stochastic Simulation Method**, which leverages the Markov Property as well.
- In Table 1, by consider the reactions and the probabilities, we can derive the Gillespie algorithm for the stochastic SIR model.

### The Gillespie method II

With initialization, S(0), I(0), and R(0), at current time t:

- 1. Compute the total event rates:  $\lambda_1=\beta\,S(t)\,I(t)/N, \lambda_2=\gamma\,I(t), \lambda=\lambda_1+\lambda_2$
- 2. Compute the waiting time  $\Delta t \sim \operatorname{Exponential}(\lambda)$
- 3. Select the reactions by sampling from probabilities  $\left(\frac{\lambda_1}{\lambda}, \frac{\lambda_2}{\lambda}\right)$
- 4. Update the states from the selected reaction and update the time  $t \to t + \Delta t$
- 5. Repeat 1-4 till the end of the simulation time

Even though the Gillespie is an exact stochastic simulation method, it has limitations such as:

- Computational Intensity: For complex systems with many reactions, the Gillespie method can become computationally expensive.
- Rare Events: For systems where some reactions are very rare, a large number of simulation steps may be needed to capture these events, making the method slow.

### Euler vs. Gillespie

▶ Why and When would you prefer an implementation of Gillespie's algorithm to an Euler solution?

#### Worked solution to the Exercise

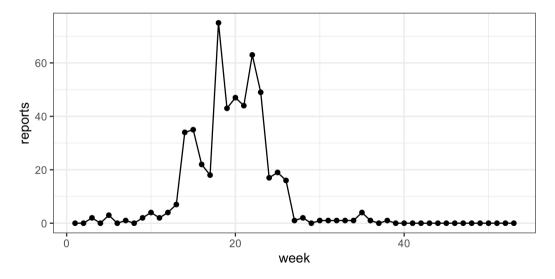
Numerically, Gillespie's algorithm is often approximated using so-called tau-leaping methods. These are closely related to Euler's approach. In this context, the Euler method has sometimes been called tau-leaping.

# Compartment models in pomp: The Consett Measles outbreaks I

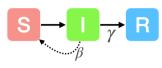
Let's look at outbreak of measles in the town of Consett in England in 1948:

- the town had population of 38820,
- with 737 births over the course of the year.

# Compartment models in pomp: The Consett Measles outbreaks II



#### The SIR as a POMP model for measles I



- The **unobserved states**: S(t), I(t), R(t), the numbers of individuals in the S, I, and R compartments, respectively.
- The constant population size: N = S(t) + I(t) + R(t), as fixed at the known population size of 38,000.
- ▶ Flows move from one compartment to another over any particular time interval are modeled as *stochastic* processes.
- ▶ Demographic stochasticity: each individual in a compartment at any given time faces the same risk of exiting the compartment; the unavoidable randomness that arises from chance events occurring in a discrete and finite population.

#### The SIR as a POMP model for measles II

Recall the application of the Euler's method to a stochastic SIR model.

 $\blacktriangleright \Delta N_{SI}$  and  $\Delta N_{IR}$ : the flows from S to I and from I to R over interval  $\Delta t$ , respectively:

$$\begin{split} & \Delta N_{SI} \sim \text{Binomial} \left( S, 1 - e^{-\beta \frac{I}{N} \Delta t} \right), \\ & \Delta N_{IR} \sim \text{Binomial} \left( I, 1 - e^{-\gamma \Delta t} \right). \end{split}$$

### The SIR as a POMP model for measles III

Implement the dynamics in pomp as an R function:

```
sir_stoch <- function (S, I, R, N, Beta, Gamma, delta.t, ...) {
  dN_SI <- rbinom(n=1,size=S,prob=1-exp(-Beta*I/N*delta.t))
  dN_IR <- rbinom(n=1,size=I,prob=1-exp(-Gamma*delta.t))
  S <- S - dN_SI
  I <- I + dN_SI - dN_IR
  R <- R + dN_IR
  c(S = S, I = I, R = R)
}</pre>
```

Note that, for a deterministic SIR model:

```
dN_SI <- Beta*S*I/N*delta.t
dN_IR <- Gamma*I*delta.t</pre>
```

#### The SIR as a POMP model for measles IV

We can implement the initialization function with the following assumptions:

- Assume the dynamics starts at week 0, t0 = 0.
- At t0, assume the initial number of infection is 1, that is I=1.
- The initial number of susceptible is unknown, so we'll treat this fraction,  $\eta$ , as a parameter to be estimated.

```
sir_rinit <- function (N, Eta, ...) {
  c(S = round(N*Eta), I = 1, R = round(N*(1-Eta)))
}</pre>
```

#### The SIR as a POMP model for measles V

With the initialization function sir\_rinit and the process function sir\_stoch, we can build a pomp object with these two components and the data:

```
library(pomp)
meas |>
  pomp(
    times="week",t0=0,
    rprocess=euler(sir_stoch,delta.t=1/7),
    rinit=sir_rinit
) -> measSIR
```

Question: what do times="week" and delta.t=1/7 indicate?

#### The SIR as a POMP model for measles VI

- Assume the **observations**, the reports, result from a process by which new infections are diagnosed in a hospital and reported with probability  $\rho$ .
- ▶ The diagnosed infections are immediately hospitalized, therefore, they have, presumably, a much lower transmission rate; let's assume each week's reports as being related to the number of individuals who have moved from I to R over the course of that week.
- We then define a new variable, H, that tracks these daily counts.

#### The SIR as a POMP model for measles VII

We now can modify the R functions to incorporate the new variable H:

```
sir stoch <- function (S, I, R, N, Beta, Gamma, delta.t, H, ...) {
  dN SI <- rbinom(n=1, size=S, prob=1-exp(-Beta*I/N*delta.t))
  dN IR <- rbinom(n=1, size=I, prob=1-exp(-Gamma*delta.t))
  S <- S - dN SI
  I \leftarrow I + dN SI - dN IR
  R \leftarrow R + dN IR
  H \leftarrow H + dN IR
  c(S = S, I = I, R = R, H = H)
sir rinit <- function (N, Eta, ...) {
  c(S = round(N*Eta), I = 1, R = round(N*(1-Eta)), H = 0)
```

### The SIR as a POMP model for measles VIII

Note that, we are so far accounting for the *flows* between compartments by days, while the reports are weekly cases. Since we want H to tally only the incidence over the week, we'll need to reset it to zero at the beginning of each week. Thus, in pomp terminology, H is an **accumulator variable**. We accomplish this using the accumvars argument to pomp when build the object:

```
measSIR |>
  pomp(
    rprocess=euler(sir_stoch,delta.t=1/7),
    rinit=sir_rinit,
    accumvars="H"
) -> measSIR
```

▶ Question: what does that mean by running a pomp function with the pomp object measSTR?

#### The SIR as a POMP model for measles IX

Last but not least, we need to define a **measurement model** to relate the **observations**, reports, to the **unobserved** accumulative state, H.

We will model the data by a negative binomial variable,

$$\operatorname{reports}_t \sim \operatorname{NegBin}\left(\rho H(t), k\right).$$

with mean  $\rho\,H(t)$  and variance  $\rho H(t) + \left(\rho H(t)\right)^2/k$ . The binomial distribution does not have a separate variance parameter.

#### The SIR as a POMP model for measles X

To include the observations in the model, we must write either a dmeasure or an rmeasure component, or both:

```
sir_dmeas <- function (reports, H, Rho, k, log, ...) {
  dnbinom(x=reports, size=k, mu=Rho*H, log=log)
}
sir_rmeas <- function (H, Rho, k, ...) {
  c(reports=rnbinom(n=1, size=k, mu=Rho*H))
}</pre>
```

### The SIR as a POMP model for measles XI

Eventually, we can add these two components to the previous measSIR object to update the dmeasure and rmeasure arguments:

```
measSIR |>
  pomp(
    rmeasure=sir_rmeas,
    dmeasure=sir_dmeas
) -> measSIR
```

# Specifying model components using C snippets I

- ▶ Although we can always specify basic model components using R functions, as above, we'll typically want the computational speed-up that we can obtain only by using compiled native code.
- pomp provides a facility for doing so with ease, using *C snippets*.
- C snippets are small pieces of C code used to specify basic model components.

# Specifying model components using C snippets II

▶ For example, a C snippet encoding the rprocess for an sir model is as follows.

```
sir_stoch <- Csnippet("
  double dN_SI = rbinom(S,1-exp(-Beta*I/N*dt));
  double dN_IR = rbinom(I,1-exp(-Gamma*dt));
  S -= dN_SI;
  I += dN_SI - dN_IR;
  R += dN_IR;
  H += dN_IR;
  ")</pre>
```

# Specifying model components using C snippets III

#### Note:

- ▶ It is necessary to define the data type for the real values dN\_SI and dN\_IR as double. The data type for states does not need to be defined at this stage and will be addressed later.
- rbinom is a built-in function used to generate random values from a binomial distribution. For additional built-in distributions in R, you can refer to this Rmath.h document.
- ▶ Remember to add a semicolon (;) after each line to ensure proper syntax.
- C snippets for the initializer and measurement model are:

# Specifying model components using C snippets IV

```
sir_rinit <- Csnippet("</pre>
  S = nearbyint(Eta*N);
  I = 1:
  R = nearbyint((1-Eta)*N);
 H = 0;
sir_dmeas <- Csnippet("</pre>
  lik = dnbinom_mu(reports,k,Rho*H,give_log);
")
sir rmeas <- Csnippet("</pre>
  reports = rnbinom_mu(k,Rho*H);
")
```

# Specifying model components using C snippets V

- No need to define the type for likelihood (lik) here, as it is already predefined.
- ▶ nearbyint is a built-in function used to find the closest integer to a given value.
- reports is the variable name specified in your dataset.

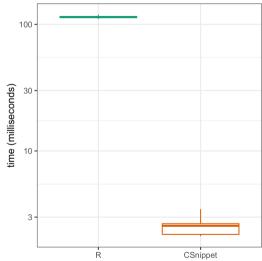
### Specifying model components using C snippets VI

▶ A call to pomp replaces the basic model components with these, much faster, implementations:

```
measSTR |>
  pomp(
    rprocess=euler(sir stoch, delta.t=1/7),
    rinit=sir rinit,
    rmeasure=sir rmeas,
    dmeasure=sir dmeas.
    accumvars="H".
    statenames=c("S","I","R","H"),
    paramnames=c("Beta", "Gamma", "N", "Eta", "Rho", "k")
  ) -> measSIR C
```

# Specifying model components using C snippets VII

- Note that, when using C snippets, one has to tell pomp which of the variables referenced in the C snippets are state variables and which are parameters. This is accomplished using the statenames and paramnames arguments.
- ➤ We can tell from the summary table that CSnippet is approximate 50 times faster than R.



### Guessing plausible parameter values I

- ▶ To check the code is working properly, we simulate. This requires us to assign parameters. A little thought will get us some ballpark estimates.
- Recall that  $\mathfrak{R}_0$  is the expected number of secondary infections resulting from one primary infection introduced into a fully susceptible population. For an SIR infection, one has that  $\mathfrak{R}_0 \approx \frac{L}{A}$ , where L is the lifespan of a host and A is the mean age of infection. Analysis of age-stratified serology data establish that the mean age of infection for measles during this period was around 4–5yr (Anderson and May 1991). Assuming a lifespan of 60–70yr, we have  $\mathfrak{R}_0 \approx 15$ .
- ▶ The basic theory of SIR epidemics gives the final-size equation,

$$\mathfrak{R}_0 = -\frac{\log\left(1 - f\right)}{f},$$

where f is the final size of the epidemic—the fraction of those susceptible at the beginning of the outbreak who ultimately become infected. For  $\Re_0 > 5$ , this equation predicts that f > 0.99.

### Guessing plausible parameter values II

- In the data, it looks like there were a total of 521 infections. Assuming 50% reporting, we have that  $S_0 \approx 1042$ , so that  $\eta = \frac{S_0}{N} \approx 0.027$ .
- If the infectious period is roughly 2 weeks, then  $1/\mu_{IR}\approx 2$  wk and  $\beta=\mu_{IR}\,\Re_0\approx 7.5$  wk $^{-1}$ .

### Guessing plausible parameter values III

```
measSIR |>
  simulate(params=c(Beta=7.5, Gamma=0.5, Rho=0.5, k=10, Eta=0.03, N=38000),
    nsim=20,format="data.frame",include.data=TRUE) |>
  ggplot(aes(x=week,y=reports,group=.id,color=.id=="data")) +
  geom line() + guides(color="none")
  60 -
reports
  40
  20
   0 -
                               20
                                                       40
                                      week
```

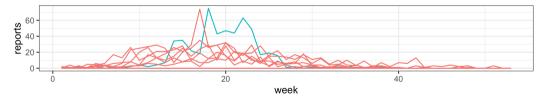
#### Exercise I: Explore the SIR model

Fiddle with the parameters to see if you can't find a model for which the data are a more plausible realization.

### Worked solutions I: Explore the SIR model I

In the simulated outbreaks, the overall incidence is much too low, and the outbreak dies out immediately. We might try increasing the force of infection:

```
measSIR |>
simulate(params=c(Beta=25,Gamma=0.5,Rho=0.5,k=10,Eta=0.03,N=38000),
    nsim=20,format="data.frame",include.data=TRUE) |>
ggplot(aes(x=week,y=reports,group=.id,color=.id=="data")) +
geom_line() + guides(color="none")
```



### Worked solutions I: Explore the SIR model II

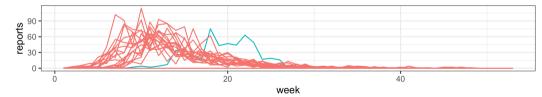
Taking it farther...

```
measSTR |>
  simulate(params=c(Beta=40,Gamma=0.5,Rho=0.5,k=10,Eta=0.03,N=38000),
    nsim=20,format="data.frame",include.data=TRUE) |>
  ggplot(aes(x=week,y=reports,group=.id,color=.id=="data")) +
  geom line() + guides(color="none")
  90
reports
  60 -
                             20
                                     week
```

# Worked solutions I: Explore the SIR model III

While this increases the overall incidence, the epidemic is now peaking too quickly. To counteract this, we might try reducing the recovery rate.

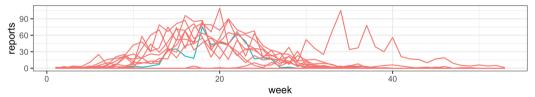
```
measSIR |>
simulate(params=c(Beta=40,Gamma=0.2,Rho=0.5,k=10,Eta=0.03,N=38000),
    nsim=20,format="data.frame",include.data=TRUE) |>
ggplot(aes(x=week,y=reports,group=.id,color=.id=="data"))+
geom_line() + guides(color="none")
```



## Worked solutions I: Explore the SIR model IV

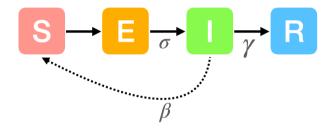
Additionally, we might have a look at the effects of changing the initial susceptible fraction,  $\eta$ :

```
measSIR |>
  simulate(params=c(Beta=15,Gamma=0.5,Rho=0.5,k=10,Eta=0.06,N=38000),
    nsim=20,format="data.frame",include.data=TRUE) |>
  ggplot(aes(x=week,y=reports,group=.id,color=.id=="data")) +
  geom_line() + guides(color="none")
```



#### Exercise II: Extend the SIR model to SEIR model

Below is a diagram of the so-called SEIR model. This differs from the SIR model in that infected individuals must pass a period of latency before becoming infectious.



Modify the codes above to construct a pomp object containing the Consett measles data and an SEIR model. Perform simulations as above and adjust parameters to get a sense of whether improvement is possible by including a latent period.

#### Worked solutions II: The SEIR model I

The existing code may be modified as follows:

```
seir stoch <- Csnippet("</pre>
 double dN SE = rbinom(S,1-exp(-Beta*I/N*dt));
 double dN EI = rbinom(E,1-exp(-Sigma*dt));
 double dN_IR = rbinom(I,1-exp(-Gamma*dt));
  S -= dN SE;
  E += dN SE - dN EI;
  I += dN EI - dN IR;
  R += dN IR:
 H += dN IR:
```

#### Worked solutions II: The SEIR model II

```
seir_init <- Csnippet("</pre>
  S = nearbyint(Eta*N);
  E = 0: I = 1:
  R = nearbyint((1-Eta)*N);
 H = 0;
measSIR |>
  pomp(
    rprocess=euler(seir_stoch,delta.t=1/7),
    rinit=seir init.
    paramnames=c("N","Beta","Sigma","Gamma","Rho","Eta","k"),
    statenames=c("S","E","I","R","H")
  ) -> measSEIR
```

#### Worked solutions II: The SEIR model III

Using the rough estimate that the latent period in measles is 8–10da, we take  $\sigma \sim 0.8 \text{wk}^{-1}$  and  $\gamma \sim 1.3 \text{wk}^{-1}$  (roughly the same generation time as before).

```
measSETR |>
   simulate(params=c(Beta=30,Sigma=0.8,Gamma=1.3,
                      Rho=0.5, k=10, Eta=0.06, N=38000)
     nsim=20,format="data.frame",include.data=TRUE) |>
   ggplot(aes(x=week,y=reports,group=.id,color=.id=="data")) +
   geom line() + guides(color="none")
reports
                                     week
```

#### Worked solutions II: The SEIR model IV

Again one can increase the force of infection:

#### References

Anderson, R. M., and R. M. May. 1991. *Infectious Diseases of Humans*. Oxford: Oxford University Press.

Bretó, Carles, Daihai He, Edward L. Ionides, and Aaron A. King. 2009. "Time Series Analysis via Mechanistic Models." *Ann Appl Stat* 3 (1): 319–48. https://doi.org/10.1214/08-AOAS201.

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- ▶ The materials build on previous versions of this course and related courses.
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- Produced with R version 4.4.1 and pomp version 5.10.
- Compiled on 2024-07-24.

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