

## 第三讲 贝叶斯决策函数

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## 内容大纲

- 引言和基本概念
- 0-1损失函数下连续变量的分类决策
- 正态分布下的线性判别分界面和二次判别分界面
- 朴素Bayes
- 离散变量的Bayes网络

## 三种主要的分类函数求解方法

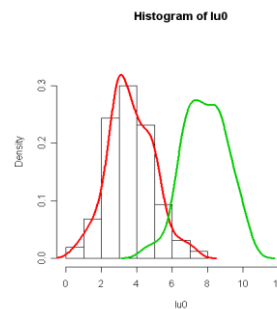
- 1. 经验风险最小化原则  
选择一个分类函数集 $\mathbf{G}$ , 搜索 $\hat{g}$ 使得经验函数最小化.
- 2. 回归: 找到函数  $\hat{\eta}(x)$  定义

$$g(x) = \begin{cases} 1 & \hat{\eta}(x) > 1/2 \\ 0 & \text{others} \end{cases}$$

- 3. 密度估计

## 第一节 基本概念

- The sea bass/salmon example
- Decision rule with only the prior information
  - Decide  $\omega_1$  if  $P(\omega_1) > P(\omega_2)$   
otherwise decide  $\omega_2$
- $P(x | \omega_1)$  and  $P(x | \omega_2)$  describe the difference in lightness between populations of sea bass and salmon



## 二分类问题

- Credit scoring: Inputs are income and savings.

Output is low-risk vs high-risk

- Input:  $\mathbf{x} = [x_1, x_2]^T$ , Output:  $C \in \{0, 1\}$

- Prediction:

choose  $\begin{cases} C = 1 & \text{if } P(C = 1 | x_1, x_2) > 0.5 \\ C = 0 & \text{otherwise} \end{cases}$

or equivalently

choose  $\begin{cases} C = 1 & \text{if } P(C = 1 | x_1, x_2) > P(C = 0 | x_1, x_2) \\ C = 0 & \text{otherwise} \end{cases}$

## Bayes' Rule

- Posterior, likelihood, evidence

$$\overset{\text{posterior}}{P(\omega_j | x)} = \overset{\text{likelihood}}{P(x | \omega_j)} \cdot \overset{\text{prior}}{P(\omega_j)} / \overset{\text{evidence}}{P(x)}$$


– Where in case of two categories

$$P(x) = \sum_{j=1}^{j=2} P(x | \omega_j) P(\omega_j)$$

– Posterior = (Likelihood. Prior) / Evidence

- Decision given the posterior probabilities

X is an observation for which:

if  $P(\omega_1 | x) > P(\omega_2 | x)$   True state of nature =  $\omega_1$

if  $P(\omega_1 | x) < P(\omega_2 | x)$   True state of nature =  $\omega_2$

Therefore:

whenever we observe a particular x, the probability of error is :

$$P(\text{error} | x) = P(\omega_1 | x) \text{ if we decide } \omega_2$$

$$P(\text{error} | x) = P(\omega_2 | x) \text{ if we decide } \omega_1$$

- Minimizing the probability of error

- Decide  $\omega_1$  if  $P(\omega_1 | x) > P(\omega_2 | x)$ ;  
otherwise decide  $\omega_2$

- Therefore:

$$P(\text{error} | x) = P(\omega_1 | x) \text{ if we decide } \omega_2$$

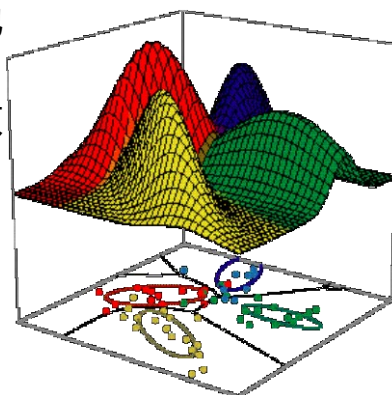
$$P(\text{error} | x) = P(\omega_2 | x) \text{ if we decide } \omega_1$$



$$P(\text{error} | x) = \min [P(\omega_1 | x), P(\omega_2 | x)]$$

## 推广到一般情况 Generalization

- 观测特征可能会多于一个;
- 判断的类别可能多于两个;
- 可能存在不同于判别的其他决策行为;
- 通过引入更一般的损失函数来替代概率误差: 损失函数测量了依照每个行为所做出每次决策的代价.



## 多类的情况

- **状态集:** Let  $\{\omega_1, \omega_2, \dots, \omega_c\}$  be the set of  $c$  states of nature (or “categories”)
- **决策集:** Let  $\{\alpha_1, \alpha_2, \dots, \alpha_t\}$  be the set of  $t$  possible actions
- Let  $\ell(\alpha_i, \omega_j)$  be the **loss** incurred for taking action  $\alpha_i$  when the state of nature is  $\omega_j$
- For any  $x$  define **conditional risk**:

$$R(\alpha_i | x) = \sum_{j=1}^{j=c} \ell(\alpha_i, \omega_j) P(\omega_j | x) \quad \text{for } i = 1, \dots, t$$

## 决策论(补充)

决策是对不确定的方案做决定，它的基本特点是通过计算后果来做选择。（包括状态的选择、参数的选择到模型的选择）

可能情况	自己生产 $\delta_1$	转让专利 $\delta_2$
	-80	-40

比如：新药品生产，有两种方案：自己生产，或转让专利。损失根据不同的销售前景有不同。

可能情况	自己生产 $\delta_1$	转让专利 $\delta_2$
销售好 $\theta_1$	-80	-40
销售一般 $\theta_2$	-20	-10
销售较差 $\theta_3$	10	-5

极小极大原则之下，最好的决策是：转让专利。

## 决策论的几个基本概念

- **参数集**：参数的所有自然可能的不同的状态  $\Theta = \{\theta\}$ ;
- **行动集**：所有可能的决策结果  $A = \{a\}$ ;

定义：损失函数 对于一个行动  $a$  和参数  $\theta$ ，评价不同行动或决策优劣的函数  $l(\theta, a)$ ，一般的损失函数有绝对损失，平方损失，线性损失等。

For each in a set of actions  $a \in \mathcal{A}$ , if the parameter is  $\theta$ , a **loss**  $L(a, \theta)$  is associated with choosing action  $a$ .

The **risk** is the expected loss:

$$R = \int_{\Theta} L(a, \theta) dF(\theta)$$

and one chooses the action that minimizes the risk.

# 统计决策论的基本概念

一个好的决策应该在参数空间上一致地好，  
模型选择应该在函数空间上一致的好

最小期望损失原则：

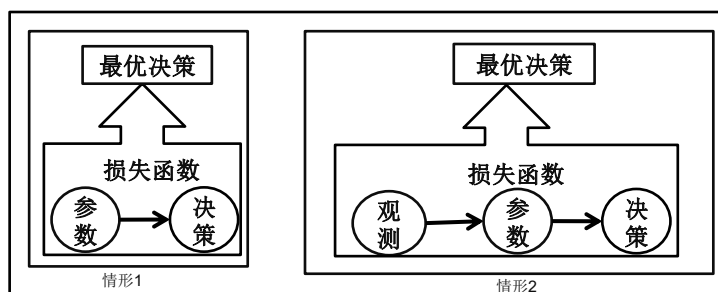
$\pi(\theta_1) = 0.2, \pi(\theta_2) = 0.5, \pi(\theta_3) = 0.3$ ，计算期望损失

$$R(\delta_1) = -80 \times 0.2 - 20 \times 0.5 + 10 \times 0.3 = -23$$

$$R(\delta_2) = -40 \times 0.2 - 10 \times 0.5 - 5 \times 0.3 = -14.5$$

例子（继续）：假如公司做了一项研究，收集了一些信息，比如调查了一些专家，他们认为公司的品牌知名度 $x$ 的取值和 $\theta$ 参数有关系，做了一些分析，根据这些分析结果如何决策？  $p(x|\theta)$

	$x = 0$	$x = 1$
$\theta_1$	0.3	0.7
$\theta_2$	0.6	0.4
$\theta_3$	0.7	0.3



## 两种情形下的最优决策

情形 2. 状态不可直接观测,但其可能性可借助其他更易观测的变量推断出来,将  $P(\theta)$  由分布  $P(\theta|x)$  代替就产生了条件风险的概念。

【条件风险的定义】: 参数状态集设为  $\Theta = \{\theta_1, \theta_2, \dots\}$ , 由观测集  $X$  决定,  $D = \{\delta\}$  是决策集, 损失函数是  $L(\theta, \delta)$ , 条件风险  $R(\delta|x)$  定义为

$$R(\delta|x) = \int_{\Theta} L(\theta, \delta) dP(\theta|x) = \int_{(\Theta, X)} \frac{1}{p(x)} L(\theta, \delta) dP(x|\theta)P(\theta)$$

最优决策如下:

$$\delta^*|x = \arg \min_{\delta \in \Delta} R(\delta|x)$$

•[例]: 两类分类问题:

- State:  $\{\omega_1, \omega_2\}$ ,
- Action :

$\alpha_1$  : deciding  $\omega_1$

$\alpha_2$  : deciding  $\omega_2$

- Loss:  $\ell_{ij} = \ell(\alpha_i, \omega_j)$ : loss incurred for deciding  $\alpha_i$  when the true state of nature is  $\omega_j$

- Conditional risk:

$$R(\alpha_1|x) = \ell_{11}P(\omega_1|x) + \ell_{12}P(\omega_2|x)$$

$$R(\alpha_2|x) = \ell_{21}P(\omega_1|x) + \ell_{22}P(\omega_2|x)$$



Our decision rule is the following:

if  $R(\alpha_1 | x) < R(\alpha_2 | x)$

action  $\alpha_1$ : “decide  $\omega_1$ ” is taken

*Remember:*

$$R(\alpha_1 | x) = \ell_{11}P(\omega_1 | x) + \ell_{12}P(\omega_2 | x)$$

$$R(\alpha_2 | x) = \ell_{21}P(\omega_1 | x) + \ell_{22}P(\omega_2 | x)$$

This results in the equivalent rule :

decide  $\omega_1$  if:

$$(\ell_{21} - \ell_{11}) P(x | \omega_1) P(\omega_1) >$$

$$(\ell_{12} - \ell_{22}) P(x | \omega_2) P(\omega_2)$$

and decide  $\omega_2$  otherwise

The preceding rule is equivalent to the following rule:

$$\text{if } \frac{P(x | \omega_1)}{P(x | \omega_2)} > \frac{\ell_{12} - \ell_{22}}{\ell_{21} - \ell_{11}} \cdot \frac{P(\omega_2)}{P(\omega_1)}$$

Then take action  $\alpha_1$  (decide  $\omega_1$ )

Otherwise take action  $\alpha_2$  (decide  $\omega_2$ )

结论: 贝叶斯决策规则可以解释成如果似然比超过某个不依赖于观测值 $x$ 的阈值, 那么判断为 $\omega_1$ .

## 课堂案例

Select the optimal decision where:

$$= \{\omega_1, \omega_2\}$$

$$P(x | \omega_1) \longrightarrow N(2, 0.5^2) \text{ (Normal distribution)}$$

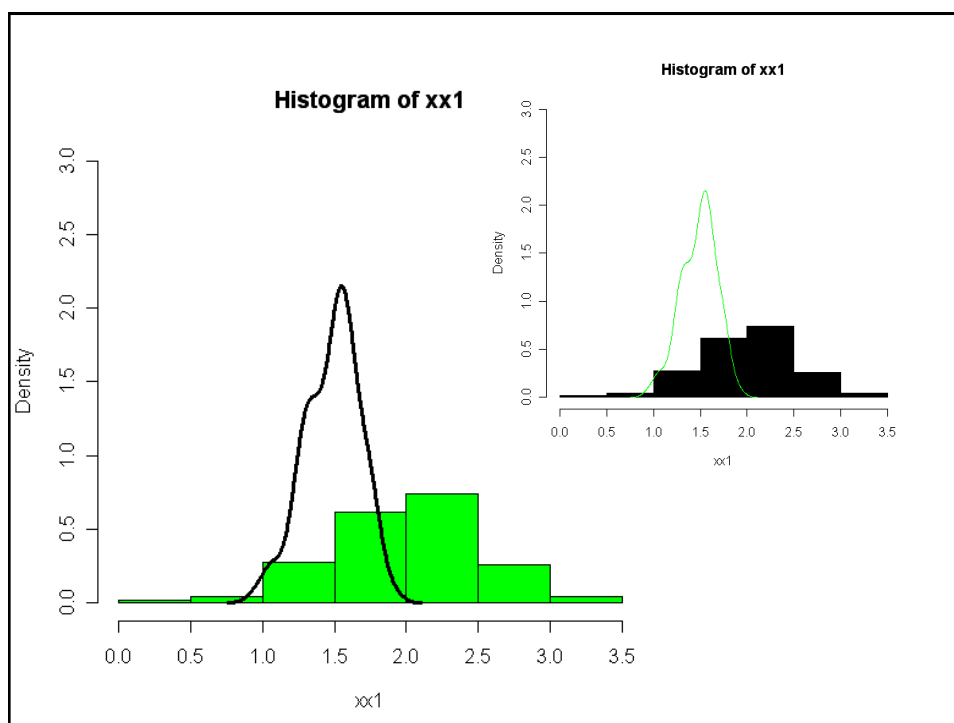
$$P(x | \omega_2) \longrightarrow N(1.5, 0.2^2)$$

$$P(\omega_1) = 2/3$$

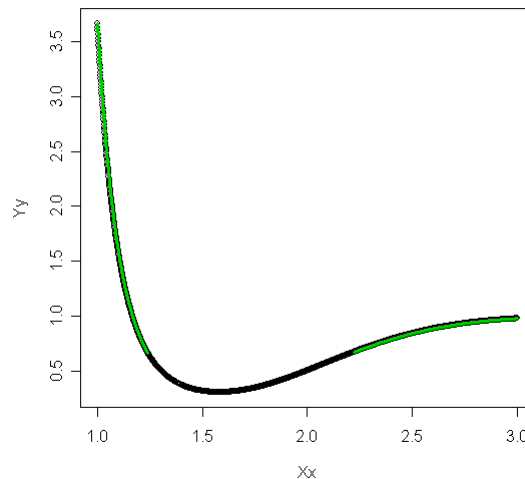
$$P(\omega_2) = 1/3$$

$$l = \begin{bmatrix} -1 & 2 \\ 3 & -4 \end{bmatrix}$$

$$l_2 = \begin{bmatrix} 0 & 1 \\ 3 & 0 \end{bmatrix}$$



# Decision Boundary



## 20201113课下作业1

在R的library(ISLR)中有一个棒球球员薪资与技术能力的数据Hitters，  
根据薪水高低标记为 $\{\omega_1, \omega_2\}$ ，高于3/4的标记为 $\omega_1$ ，低于3/4的标记为 $\omega_2$ ，x选为Hits

根据题意设置如下分布，Select the optimal decision where:

$$\begin{array}{ll} P(x | \omega_1) & \Rightarrow N(\mu_1, \sigma_1^2) \text{ (Normal distribution)} \\ P(x | \omega_2) & \Rightarrow N(\mu_2, \sigma_2^2) \end{array}$$

$P(\omega_1), P(\omega_2)$ ，随机选出 2/3 比例的数据作为训练数据，其余数据作为测试数据，结合如下损失函数和参数的极大似然估计给出决策区域，给出决策边界，讨论错误率。根据实际决策场景自定义损失函数的合理代价，根据自定义损失函数重新讨论决策。

$$l_2 = \begin{bmatrix} 0 & 1 \\ 3 & 0 \end{bmatrix}$$

$$l_2 = \begin{bmatrix} & \\ & \end{bmatrix}$$

## 20201113课下作业2

在R的library(ISLR)中有一个棒球球员薪资与技术能力的数据Hitters，  
根据薪水高低标记为 $\{\omega_1, \omega_2\}$ ，高于3/4的标记为 $\omega_1$ ，低于3/4的标记为 $\omega_2$ ，x选为Hits

根据题意设置如下分布，Select the optimal decision where:

$$P(x | \omega_1) \quad \longrightarrow \quad p1(x)$$

$$P(x | \omega_2) \quad \longrightarrow \quad p2(x)$$

$P(\omega_1), P(\omega_2)$ ，不做正态分布假设，用非参数密度估计各自的密度，  
选择合适的带宽，结合鲜艳分布，随机选出 2/3比例的数据作为训练  
数据，其余数据作为测试数据，结合如下损失函数和参数的极大似  
然估计给出决策区域，给出决策边界，讨论错误率。与1的结果进行  
比较

$$l_2 = \begin{bmatrix} 0 & 1 \\ 3 & 0 \end{bmatrix}$$

## 20201113课下作业3

根据上一个作业中薪水高低的标记方式  $\{\omega_1, \omega_2\}$ ，高于3/4的标记为 $\omega_1$ ，  
低于3/4的标记为 $\omega_2$ ，x选为Hits，CAAtBat,CHits

根据题意设置  $P(\omega_1), P(\omega_2)$ ，随机选出 2/3比例的数据作为训练数据，  
其余数据作为测试数据，给出LDA的判别结果，画出判别图

## 第二节 0-1损失函数下的分类决策

- Introduction of the zero-one loss function:

$$l(\alpha_i, \omega_j) = \begin{cases} 0 & i = j \\ 1 & i \neq j \end{cases} \quad i, j = 1, \dots, c$$

- Therefore, the conditional risk is:

$$\begin{aligned} R(\alpha_i | x) &= \sum_{j=1}^{j=c} l(\alpha_i, \omega_j) P(\omega_j | x) \\ &= \sum_{j \neq i} P(\omega_j | x) = 1 - P(\omega_i | x) \end{aligned}$$

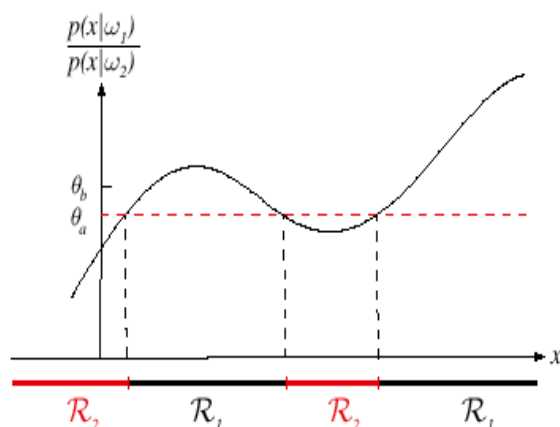
- Minimize the risk requires maximize  $P(\omega_i | x)$   
(since  $R(\alpha_i | x) = 1 - P(\omega_i | x)$ )

- Regions of decision and zero-one loss function, therefore:

$$\text{Let } \frac{l_{12} - l_{22}}{l_{21} - l_{11}} \cdot \frac{P(\omega_2)}{P(\omega_1)} > \theta_l \text{ then decide } \omega_1 \text{ if: } \frac{P(x | \omega_1)}{P(x | \omega_2)} > \theta_l$$

- If  $\ell$  is the zero-one loss function which means:

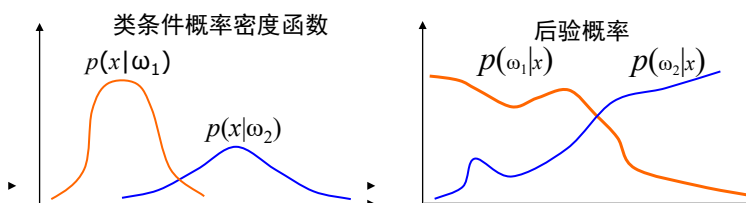
$$\begin{aligned} l &= \begin{pmatrix} 0 & 1 \\ 1 & 0 \end{pmatrix} \\ \text{then } \theta_l &= \frac{P(\omega_2)}{P(\omega_1)} = \theta_a \\ \text{if } l &= \begin{pmatrix} 0 & 2 \\ 1 & 0 \end{pmatrix} \text{ then } \theta_l = \frac{2P(\omega_2)}{P(\omega_1)} = \theta_b \end{aligned}$$



**FIGURE 2.3.** The likelihood ratio  $p(x|\omega_1)/p(x|\omega_2)$  for the distributions shown in Fig. 2.1. If we employ a zero-one or classification loss, our decision boundaries are determined by the threshold  $\theta_a$ . If our loss function penalizes miscategorizing  $\omega_2$  as  $\omega_1$  patterns more than the converse, we get the larger threshold  $\theta_b$ , and hence  $\mathcal{R}_1$  becomes smaller. From: Richard O. Duda, Peter E. Hart, and David G. Stork, *Pattern Classification*. Copyright © 2001 by John Wiley & Sons, Inc.

## 最优分类器

- 最优分类器?
- 已有: 类条件概率密度函数
  - This is called the class-conditional probability describing the probability of occurrence of the features on category.
- 欲求: 后验概率
  - make a decision that maximize the conditional probability of the object, given certain feature measurements.
  - Also called posterior probability function.



- Let  $g_i(x) = -R(\alpha_i | x)$   
(max. discriminant corresponds to min. risk!)

- For the minimum error rate, we take

$$g_i(x) = P(\omega_i | x)$$

(max. discrimination corresponds to max. posterior!)

$$g_i(x) \equiv P(x | \omega_i) P(\omega_i)$$

判别函数可以替换为任意一个单调增函数 $f(g(.))$

$$g_i(x) = \ln P(x | \omega_i) + \ln P(\omega_i)$$

(ln: natural logarithm!)

- Feature space divided into  $c$  decision regions(判决空间被分成 $c$ 个判决区域)

$$\text{if } g_i(x) > g_j(x) \quad \forall j \neq i \text{ then } x \text{ is in } \mathcal{R}_i$$

( $\mathcal{R}_i$  means assign  $x$  to  $\omega_i$ )

二分类问题

- A classifier is a “dichotomizer” that has two discriminant functions  $g_1$  and  $g_2$

$$\text{Let } g(x) \equiv g_1(x) - g_2(x)$$

Decide  $\omega_1$  if  $g(x) > 0$  ; Otherwise decide  $\omega_2$

- Multivariate density

- Multivariate normal density in d dimensions is:

$$P(\mathbf{x}) = \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp \left[ -\frac{1}{2} (\mathbf{x} - \mu)^t \Sigma^{-1} (\mathbf{x} - \mu) \right]$$

where:

$\mathbf{x} = (x_1, x_2, \dots, x_d)^t$  (t stands for the transpose vector form)

$\mu = (\mu_1, \mu_2, \dots, \mu_d)^t$  mean vector

$\Sigma = d \times d$  covariance matrix

$|\Sigma|$  and  $\Sigma^{-1}$  are determinant and inverse respectively

## Discriminant Functions for the Normal Density 二次判别函数

$$P(x | w_i) = \frac{1}{(2\pi)^{d/2} |\Sigma_i|^{1/2}} \exp \left[ -\frac{1}{2} (x - \mu_i)^t \Sigma_i^{-1} (x - \mu_i) \right]$$

- The minimum error-rate classification can be achieved by the discriminant function

Quadratic Discriminant Analysis

$$g_i(x) = \ln P(x | w_i) + \ln P(w_i)$$

QDA

- In Multinormal case, If  $p(x | w_i) \sim N(\mu_i, \Sigma_i)$

$$g_i(x) = -\frac{1}{2} (x - \mu_i)^t \Sigma_i^{-1} (x - \mu_i) - \frac{d}{2} \ln 2\pi - \frac{1}{2} \ln |\Sigma_i| + \ln P(w_i)$$



$$r_i^2(x) = \frac{1}{2}(x - \mu_i)^t \sum_i^{-1} (x - \mu_i)$$

$$f^*(x) = \begin{cases} 1 & r_1^2(x) < r_0^2(x) + 2 \ln \frac{\pi_1}{\pi_0} + \ln \frac{|\Sigma_0|}{|\Sigma_1|} \\ 0 & \text{others} \end{cases}$$

推广到多类

$Y = \{1, \dots, c\}$ , 若  $p(x|y=j)$  是正态的, 则 *Bayes decision function*:

$$f^*(x) = \arg \max f_i(x)$$

$$f_i(x) = -\frac{1}{2}(x - \mu_i)^t \sum_i^{-1} (x - \mu_i) - \frac{1}{2} \ln |\Sigma_i| + \ln(\pi_i)$$

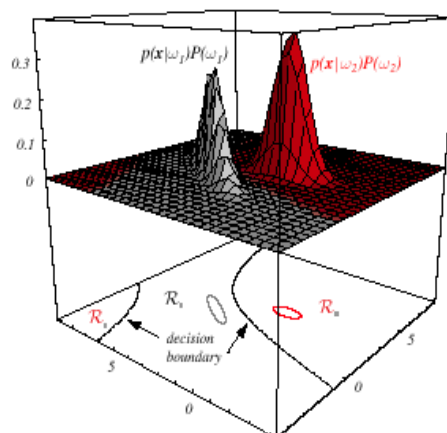
用估计表示

$$g_i(x) = -\frac{1}{2}(x - \mu_i)^t \sum_i^{-1} (x - \mu_i) - \frac{d}{2} \ln 2\pi - \frac{1}{2} \ln |\Sigma_i| + \ln P(\omega_i)$$

- 定理: 若  $p(x|\omega_i) \sim N(\mu_i, \Sigma_i)$

$$\hat{g}_i(x) = -\frac{1}{2}(x - \hat{\mu}_i)^t (\hat{\Sigma}_i^{-1})(x - \hat{\mu}_i) - \frac{d}{2} \ln 2\pi - \frac{1}{2} \ln |\hat{\Sigma}_i| + \ln P(\omega_i)$$

$$\hat{\mu}_i = \frac{1}{n_i} \sum_{j=1}^{n_i} X_{1,j}, \hat{\Sigma}_i = \frac{1}{n_i} \sum_{j=1}^{n_i} (X_{1,j} - \hat{\mu}_i)(X_{1,j} - \hat{\mu}_i)^T$$



**FIGURE 2.6.** In this two-dimensional two-category classifier, the probability densities are Gaussian, the decision boundary consists of two hyperbolas, and thus the decision region  $\mathcal{R}_2$  is not simply connected. The ellipses mark where the density is  $1/e$  times that at the peak of the distribution. From: Richard O. Duda, Peter E. Hart, and David G. Stork, *Pattern Classification*. Copyright © 2001 by John Wiley & Sons, Inc.

情形1.  $\Sigma_i = \sigma^2 I$   
 ( $I$  stands for the identity matrix)

$$g_i(x) = -\frac{\|x - \mu_i\|^2}{2\sigma^2} + \ln P(\omega_i)$$

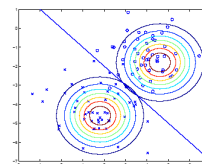
$$\|x - \mu_i\|^2 = (x - \mu_i)^t (x - \mu_i) = x^t x - 2\mu_i^t x + \mu_i^t \mu_i$$

$$g_i(x) = w_i^t x + w_{i0} \text{ (linear discriminant function)}$$

where :

$$w_i = \frac{\mu_i}{\sigma^2}; \quad w_{i0} = -\frac{1}{2\sigma^2} \mu_i^t \mu_i + \ln P(\omega_i)$$

( $w_{i0}$  is called the threshold for the  $i$ th category!)



- A classifier that uses linear discriminant functions is called “a linear machine”  
(使用线性判别函数的分类模型,称为线性机)
- The decision surfaces for a linear machine are pieces of **hyperplanes** defined by:  
(线性判别机是超平面)

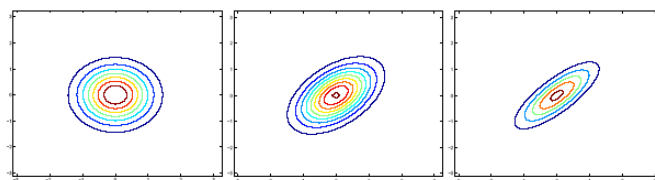
$$g_i(x) = g_j(x)$$

$$w^T(x - x_0) = 0$$

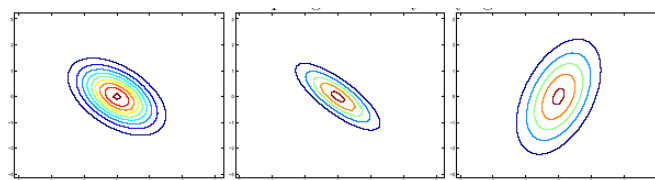
where :

$$w = \mu_i - \mu_j$$

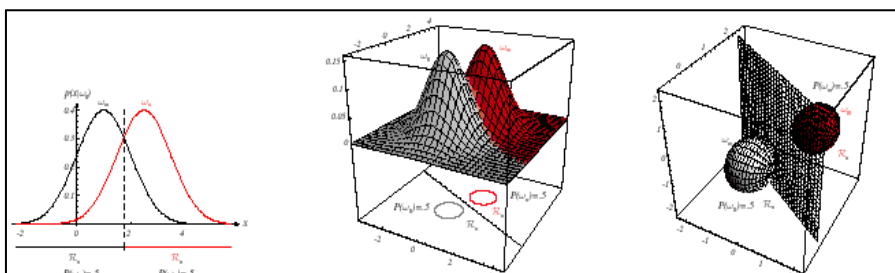
$$x_0 = \frac{1}{2}(\mu_i + \mu_j) - \frac{\sigma^2}{\|\mu_i - \mu_j\|^2} \ln \frac{P(\omega_i)}{P(\omega_j)} (\mu_i - \mu_j)$$



$$\Sigma = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}; \Sigma = \begin{bmatrix} 1 & 0.5 \\ 0.5 & 1 \end{bmatrix}; \Sigma = \begin{bmatrix} 1 & 0.8 \\ 0.8 & 1 \end{bmatrix}$$



$$\Sigma = \begin{bmatrix} 1 & -0.5 \\ -0.5 & 1 \end{bmatrix}; \Sigma = \begin{bmatrix} 1 & -0.8 \\ -0.8 & 1 \end{bmatrix}; \Sigma = \begin{bmatrix} 3 & 0.8 \\ 0.8 & 1 \end{bmatrix}$$



**FIGURE 2.10.** If the covariance matrices for two distributions are equal and proportional to the identity matrix, then the distributions are spherical in  $d$  dimensions, and the boundary is a generalized hyperplane of  $d - 1$  dimensions, perpendicular to the line separating the means. In these one-, two-, and three-dimensional examples, we indicate  $p(x|\omega_i)$  and the boundaries for the case  $P(\omega_1) = P(\omega_2)$ . In the three-dimensional case, the grid plane separates  $\mathcal{R}_1$  from  $\mathcal{R}_2$ . From: Richard O. Duda, Peter E. Hart, and David G. Stork, *Pattern Classification*. Copyright © 2001 by John Wiley & Sons, Inc.

– The hyperplane separating  $\mathcal{R}_i$  and  $\mathcal{R}_j$

$$x_0 = \frac{1}{2}(\mu_i + \mu_j) - \frac{\sigma^2}{\|\mu_i - \mu_j\|^2} \ln \frac{P(\omega_i)}{P(\omega_j)} (\mu_i - \mu_j)$$

always orthogonal to the line linking the means!

$$\text{if } P(\omega_i) = P(\omega_j) \text{ then } x_0 = \frac{1}{2}(\mu_i + \mu_j)$$

# 先验概率相等和不等的分界面

