

Rob J Hyndman

Forecasting using



4. White noise and time series decomposition

[OTexts.com/fpp/2/6](https://otexts.com/fpp/2/6)

[OTexts.com/fpp/6/](https://otexts.com/fpp/6/)

Student award

- To students who make the most contribution to the class, as voted by their peers.
- Contributions can be on Piazza or during the webex online sessions.
- Contribute questions, answers, comments, code suggestions, etc.
- Looking for engaged learners, not experts.
- At end of course, all students to vote for best contributor.
- **\$100 to the top voted student and \$50 to the second student** (plus t-shirts).

**All sessions are at UTC 22:00
for the entire course.**

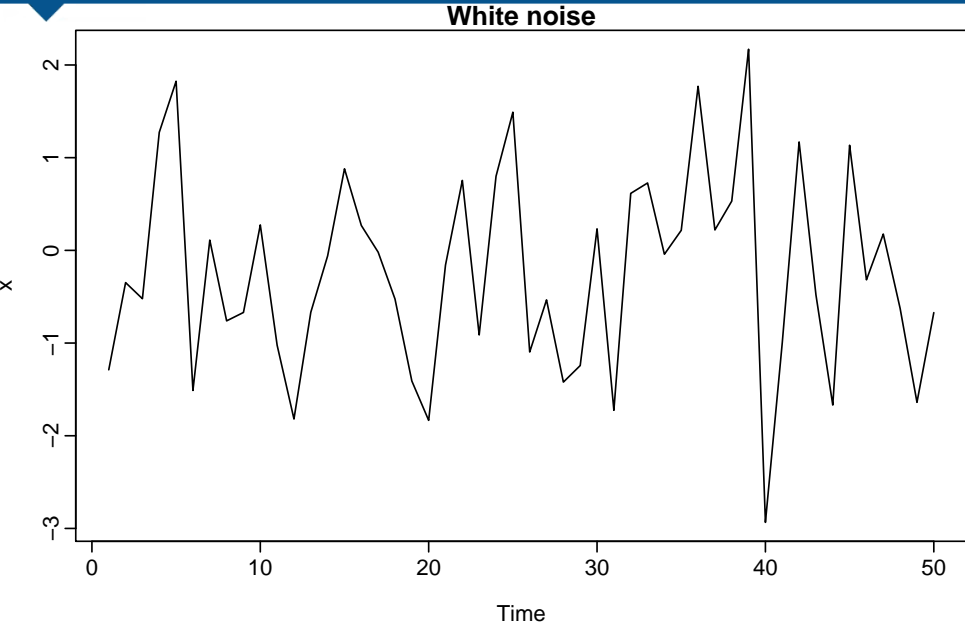
Be aware when ***your*** time zone changes.

- Most of Europe changed last Sunday. So classes are now one hour earlier for most European residents.
- Most of USA and Canada change next Sunday. So classes will be one hour earlier for most North American residents from next week.

Outline

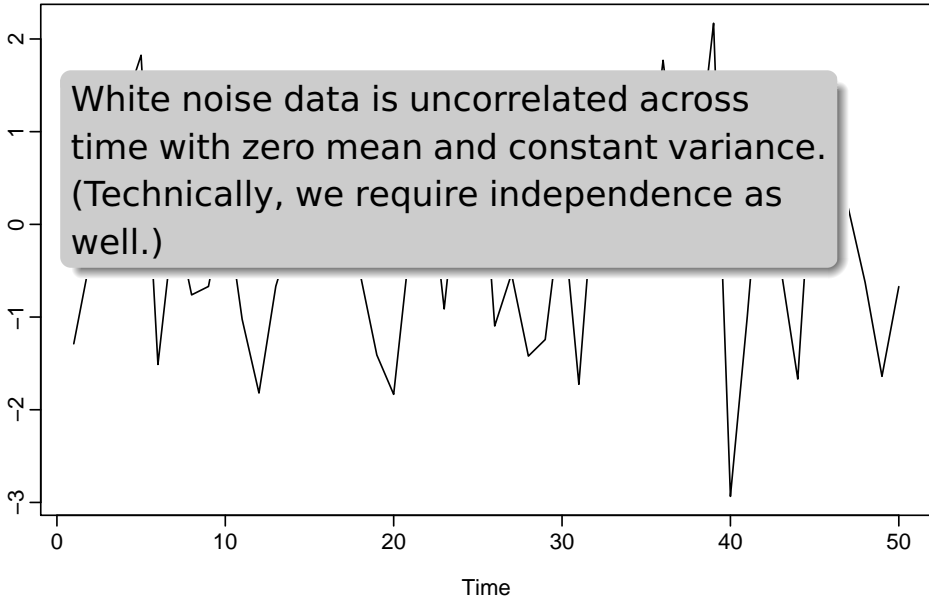
- 1 White noise**
- 2 Time series decomposition
- 3 Seasonal adjustment
- 4 Forecasting and decomposition

Example: White noise



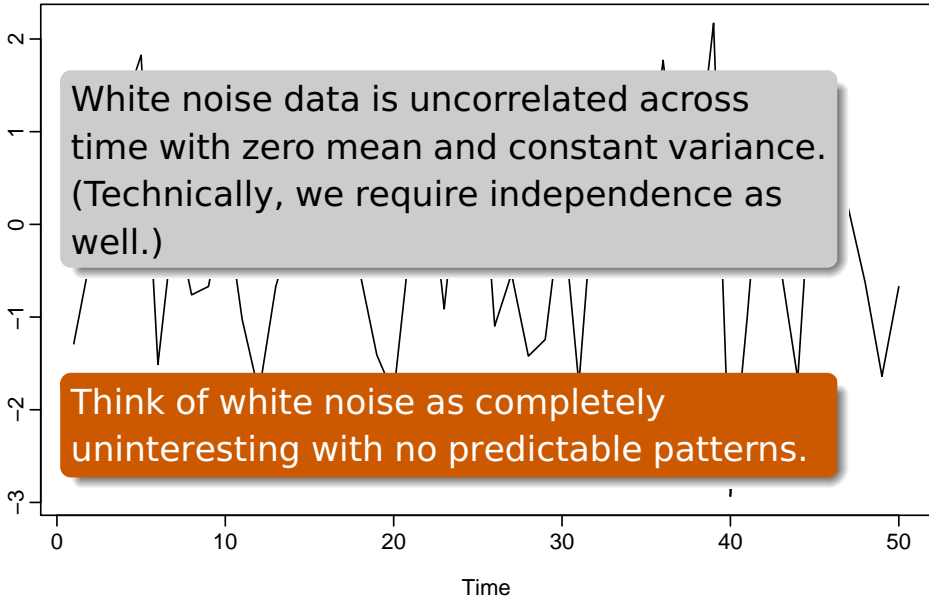
Example: White noise

White noise



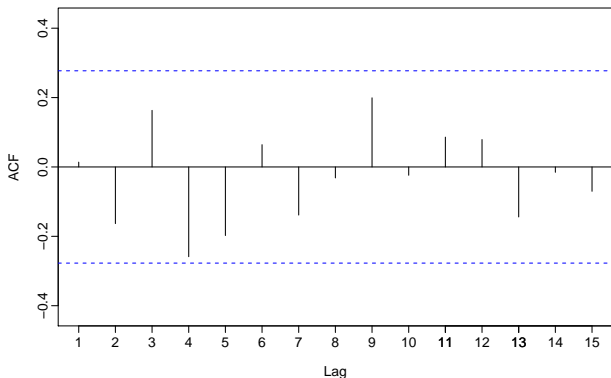
Example: White noise

White noise



Example: White noise

$r_1 = 0.013$
 $r_2 = -0.163$
 $r_3 = 0.163$
 $r_4 = -0.259$
 $r_5 = -0.198$
 $r_6 = 0.064$
 $r_7 = -0.139$
 $r_8 = -0.032$
 $r_9 = 0.199$
 $r_{10} = -0.240$



Sample autocorrelations for white noise series.
For uncorrelated data, we would expect each autocorrelation to be close to zero.

Sampling distribution of r_k for white noise data is asymptotically $N(0, 1/T)$.

- 95% of all r_k for white noise must lie within $\pm 1.96/\sqrt{T}$.
- If this is not the case, the series is probably not WN.
- Common to plot lines at $\pm 1.96/\sqrt{T}$ when plotting ACF. These are the confidence intervals.

Sampling distribution of autocorrelations

Sampling distribution of r_k for white noise data is asymptotically $N(0, 1/T)$.

- 95% of all r_k for white noise must lie within $\pm 1.96/\sqrt{T}$.
- If this is not the case, the series is probably not WN.
- Common to plot lines at $\pm 1.96/\sqrt{T}$ when plotting ACF. These are the *critical values*.

Sampling distribution of autocorrelations

Sampling distribution of r_k for white noise data is asymptotically $N(0, 1/T)$.

- 95% of all r_k for white noise must lie within $\pm 1.96/\sqrt{T}$.
- If this is not the case, the series is probably not WN.
- Common to plot lines at $\pm 1.96/\sqrt{T}$ when plotting ACF. These are the *critical values*.

Sampling distribution of autocorrelations

Sampling distribution of r_k for white noise data is asymptotically $N(0, 1/T)$.

- 95% of all r_k for white noise must lie within $\pm 1.96/\sqrt{T}$.
- If this is not the case, the series is probably not WN.
- Common to plot lines at $\pm 1.96/\sqrt{T}$ when plotting ACF. These are the **critical values**.

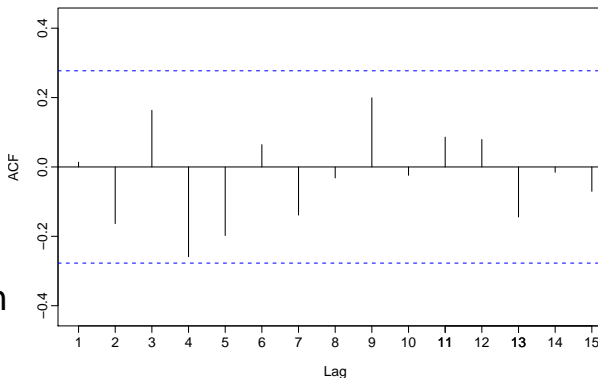
Autocorrelation

Example:

$T = 50$ and so
critical values at
 $\pm 1.96 / \sqrt{50} =$
 ± 0.28 .

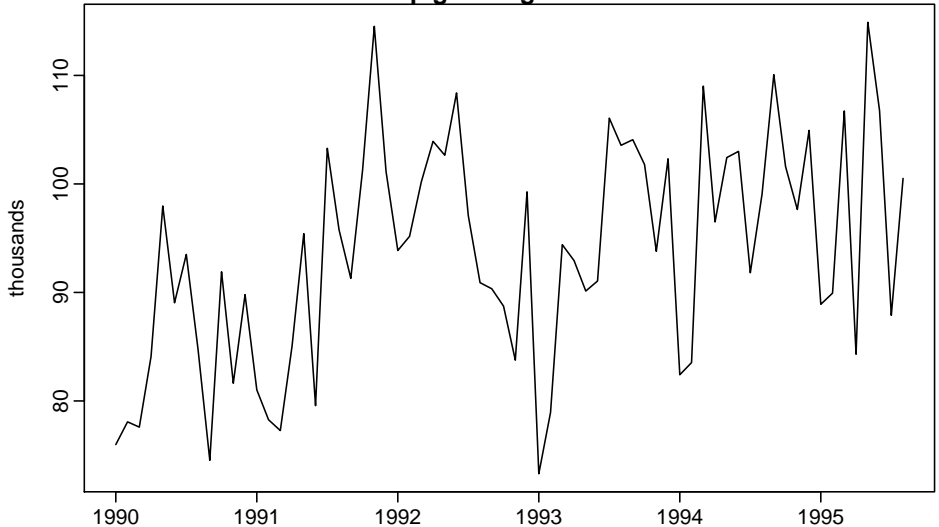
All autocorrelation
coefficients lie within
these limits,
confirming that the
data are white noise.

(More precisely, the data cannot be
distinguished from white noise.)

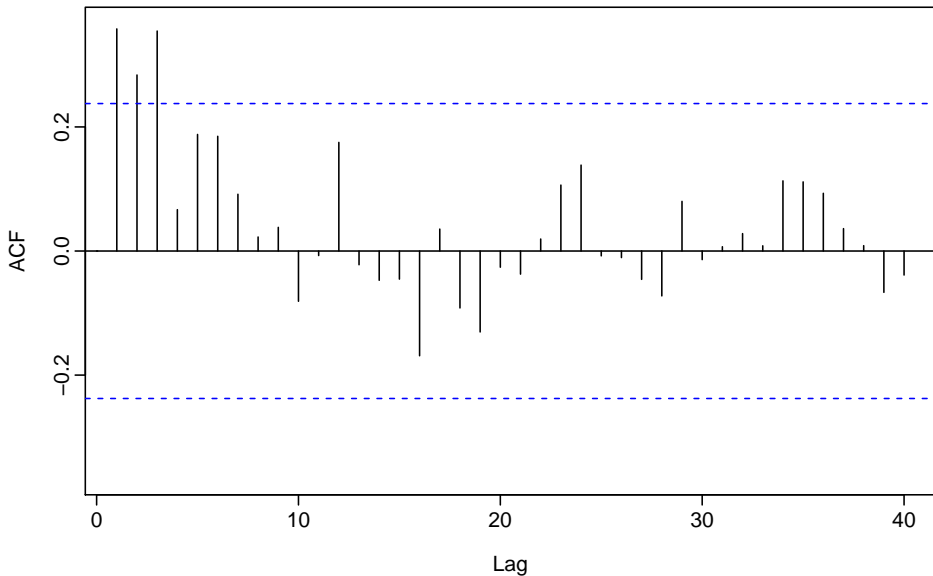


Example: Pigs slaughtered

Number of pigs slaughtered in Victoria



Example: Pigs slaughtered



Example: Pigs slaughtered

Monthly total number of pigs slaughtered in the state of Victoria, Australia, from January 1990 through August 1995. (Source: Australian Bureau of Statistics.)

- Difficult to detect pattern in time plot.
- ACF shows some significant autocorrelation at lags 1, 2, and 3.
- r_{12} relatively large although not significant. This may indicate some slight seasonality.

These show the series is **not a white noise series**.

Example: Pigs slaughtered

Monthly total number of pigs slaughtered in the state of Victoria, Australia, from January 1990 through August 1995. (Source: Australian Bureau of Statistics.)

- Difficult to detect pattern in time plot.
- ACF shows some significant autocorrelation at lags 1, 2, and 3.
- r_{12} relatively large although not significant. This may indicate some slight seasonality.

These show the series is **not a white noise series**.

Example: Pigs slaughtered

Monthly total number of pigs slaughtered in the state of Victoria, Australia, from January 1990 through August 1995. (Source: Australian Bureau of Statistics.)

- Difficult to detect pattern in time plot.
- ACF shows some significant autocorrelation at lags 1, 2, and 3.
- r_{12} relatively large although not significant. This may indicate some slight seasonality.

These show the series is **not a white noise series**.

Example: Pigs slaughtered

Monthly total number of pigs slaughtered in the state of Victoria, Australia, from January 1990 through August 1995. (Source: Australian Bureau of Statistics.)

- Difficult to detect pattern in time plot.
- ACF shows some significant autocorrelation at lags 1, 2, and 3.
- r_{12} relatively large although not significant. This may indicate some slight seasonality.

These show the series is **not a white noise series**.

Example: Pigs slaughtered

Monthly total number of pigs slaughtered in the state of Victoria, Australia, from January 1990 through August 1995. (Source: Australian Bureau of Statistics.)

- Difficult to detect pattern in time plot.
- ACF shows some significant autocorrelation at lags 1, 2, and 3.
- r_{12} relatively large although not significant. This may indicate some slight seasonality.

These show the series is **not a white noise series**.

Example: Pigs slaughtered

Monthly total number of pigs slaughtered in the state of Victoria, Australia, from January 1990 through August 1995. (Source: Australian Bureau of Statistics.)

- Difficult to detect pattern in time plot.
- ACF shows some significant autocorrelation at lags 1, 2, and 3.
- r_{12} relatively large although not significant. This may indicate some slight seasonality.

These show the series is **not a white noise series**.

ACF of residuals

- We assume that the residuals are white noise (uncorrelated, mean zero, constant variance). If they aren't, then there is information left in the residuals that should be used in computing forecasts.
- So a standard residual diagnostic is to check the ACF of the residuals of a forecasting method.
- We *expect* these to look like white noise.

Dow-Jones naive forecasts revisited

$$\hat{y}_{t-1} = y_{t-1}$$
$$e_t = y_t - \hat{y}_{t-1}$$

ACF of residuals

- We assume that the residuals are white noise (uncorrelated, mean zero, constant variance). If they aren't, then there is information left in the residuals that should be used in computing forecasts.
- So a standard residual diagnostic is to check the ACF of the residuals of a forecasting method.
- We expect these to look like white noise.

Dow-Jones naive forecasts revisited

$$\hat{y}_{t|t-1} = y_{t-1}$$
$$e_t = y_t - \hat{y}_{t|t-1}$$

ACF of residuals

- We assume that the residuals are white noise (uncorrelated, mean zero, constant variance). If they aren't, then there is information left in the residuals that should be used in computing forecasts.
- So a standard residual diagnostic is to check the ACF of the residuals of a forecasting method.
- *We expect these to look like white noise.*

Dow-Jones naive forecasts revisited

$$\hat{y}_{t|t-1} = y_{t-1}$$
$$e_t = y_t - y_{t-1}$$

ACF of residuals

- We assume that the residuals are white noise (uncorrelated, mean zero, constant variance). If they aren't, then there is information left in the residuals that should be used in computing forecasts.
- So a standard residual diagnostic is to check the ACF of the residuals of a forecasting method.
- We *expect* these to look like white noise.

Dow-Jones naive forecasts revisited

$$\hat{y}_{t|t-1} = y_{t-1}$$
$$e_t = y_t - y_{t-1}$$

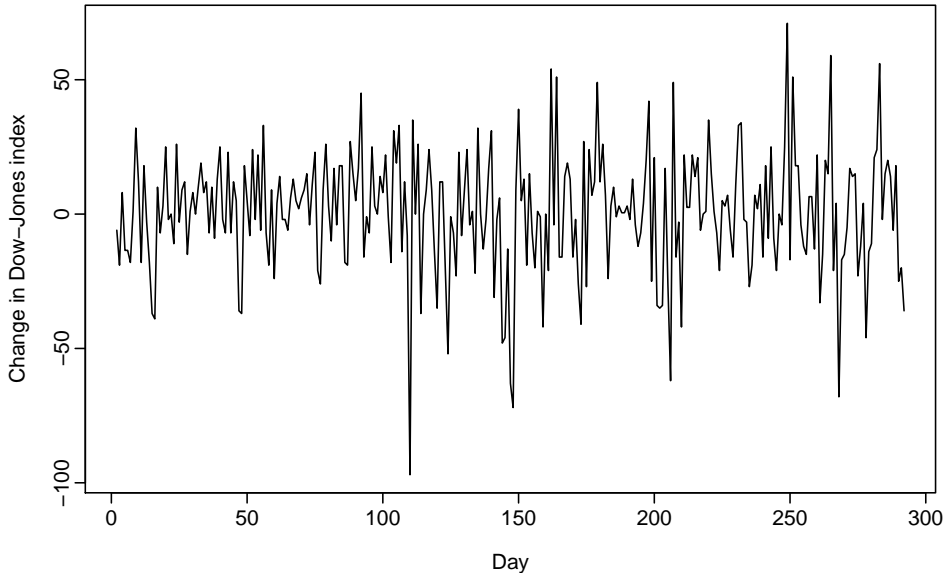
ACF of residuals

- We assume that the residuals are white noise (uncorrelated, mean zero, constant variance). If they aren't, then there is information left in the residuals that should be used in computing forecasts.
- So a standard residual diagnostic is to check the ACF of the residuals of a forecasting method.
- We *expect* these to look like white noise.

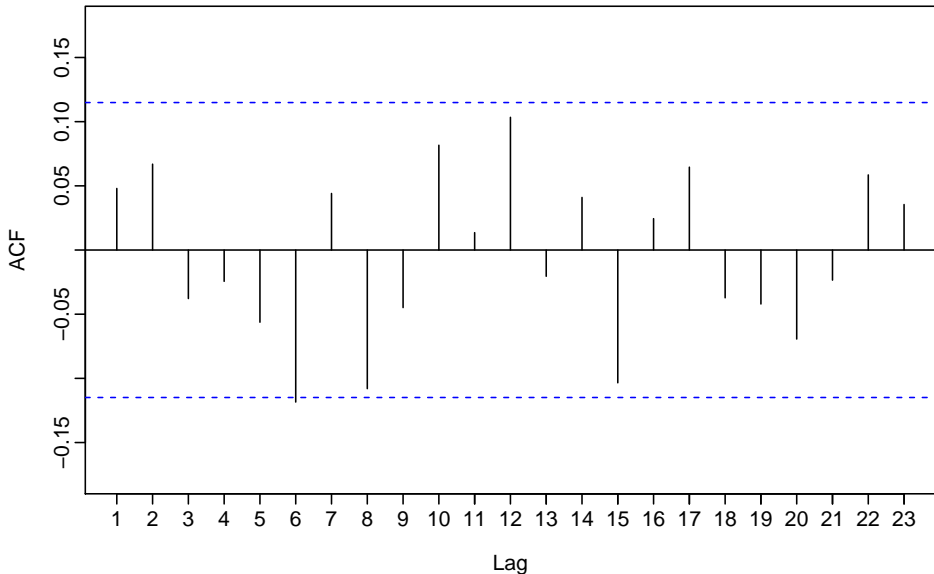
Dow-Jones naive forecasts revisited

$$\hat{y}_{t|t-1} = y_{t-1}$$
$$e_t = y_t - y_{t-1}$$

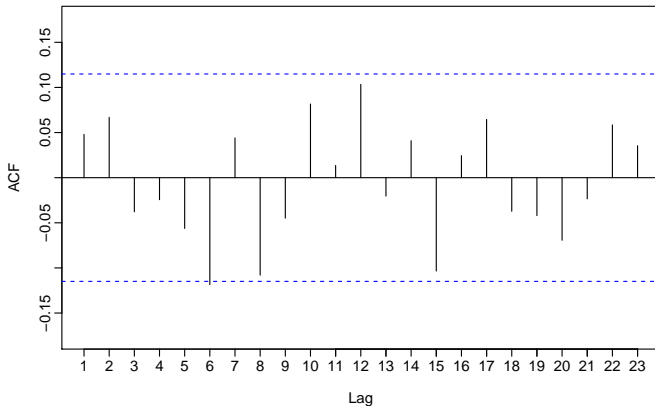
Forecasting Dow-Jones index



Forecasting Dow-Jones index

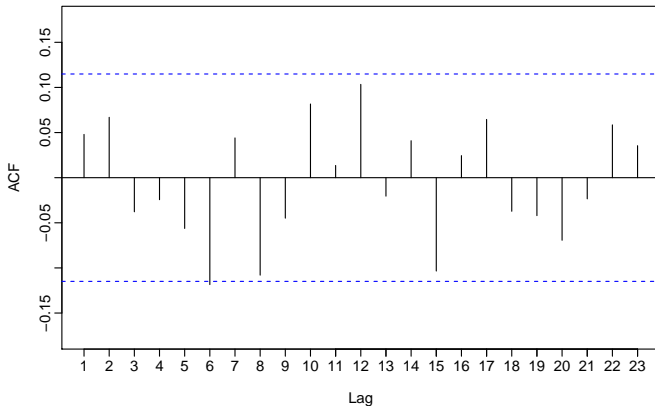


Example: Dow-Jones residuals



- These look like white noise.
- But the ACF is a multiple testing problem.

Example: Dow-Jones residuals



- These look like white noise.
- But the ACF is a multiple testing problem.

Portmanteau tests

Consider a *whole set* of r_k values, and develop a test to see whether the set is significantly different from a zero set.

Portmanteau tests

Consider a *whole* set of r_k values, and develop a test to see whether the set is significantly different from a zero set.

Box-Pierce test

$$Q = T \sum_{k=1}^h r_k^2$$

where h is max lag being considered and T is number of observations.

- My preferences: $h = 10$ for non-seasonal data, $h = 2m$ for seasonal data.
- If each r_k close to zero, Q will be small.
- If some r_k values large (positive or negative), Q will be large.

Portmanteau tests

Consider a *whole* set of r_k values, and develop a test to see whether the set is significantly different from a zero set.

Box-Pierce test

$$Q = T \sum_{k=1}^h r_k^2$$

where h is max lag being considered and T is number of observations.

- My preferences: $h = 10$ for non-seasonal data, $h = 2m$ for seasonal data.
- If each r_k close to zero, Q will be **small**.
- If some r_k values large (positive or negative), Q will be **large**.

Portmanteau tests

Consider a *whole* set of r_k values, and develop a test to see whether the set is significantly different from a zero set.

Box-Pierce test

$$Q = T \sum_{k=1}^h r_k^2$$

where h is max lag being considered and T is number of observations.

- My preferences: $h = 10$ for non-seasonal data, $h = 2m$ for seasonal data.
- If each r_k close to zero, Q will be **small**.
- If some r_k values large (positive or negative), Q will be **large**.

Portmanteau tests

Consider a *whole* set of r_k values, and develop a test to see whether the set is significantly different from a zero set.

Box-Pierce test

$$Q = T \sum_{k=1}^h r_k^2$$

where h is max lag being considered and T is number of observations.

- My preferences: $h = 10$ for non-seasonal data, $h = 2m$ for seasonal data.
- If each r_k close to zero, Q will be **small**.
- If some r_k values large (positive or negative), Q will be **large**.

Portmanteau tests

Consider a *whole* set of r_k values, and develop a test to see whether the set is significantly different from a zero set.

Ljung-Box test

$$Q^* = T(T+2) \sum_{k=1}^h (T-k)^{-1} r_k^2$$

where h is max lag being considered and T is number of observations.

- My preferences: $h = 10$ for non-seasonal data, $h = 2m$ for seasonal data.
- Better performance, especially in small samples.

Portmanteau tests

- If data are WN, Q^* has χ^2 distribution with $(h - K)$ degrees of freedom where $K = \text{no. parameters in model}$.
- When applied to raw data, set $K = 0$.
- For the Dow-Jones example,

```
res <- residuals(naive(dj))
```

```
# lag=h and fitdf=K
```

```
> Box.test(res, lag=10, fitdf=0)
```

```
Box-Pierce test
```

```
X-squared = 14.0451, df = 10, p-value = 0.1709
```

```
> Box.test(res, lag=10, fitdf=0, type="Lj")
```

```
Box-Ljung test
```

```
X-squared = 14.4615, df = 10, p-value = 0.153
```

Portmanteau tests

- If data are WN, Q^* has χ^2 distribution with $(h - K)$ degrees of freedom where $K = \text{no. parameters in model}$.
- When applied to raw data, set $K = 0$.
- For the Dow-Jones example,

```
res <- residuals(naive(dj))
```

```
# lag=h and fitdf=K
```

```
> Box.test(res, lag=10, fitdf=0)
```

```
Box-Pierce test
```

```
X-squared = 14.0451, df = 10, p-value = 0.1709
```

```
> Box.test(res, lag=10, fitdf=0, type="Lj")
```

```
Box-Ljung test
```

```
X-squared = 14.4615, df = 10, p-value = 0.153
```

Portmanteau tests

- If data are WN, Q^* has χ^2 distribution with $(h - K)$ degrees of freedom where $K = \text{no. parameters in model}$.
- When applied to raw data, set $K = 0$.
- For the Dow-Jones example,

```
res <- residuals(naive(dj))
```

```
# lag=h and fitdf=K
```

```
> Box.test(res, lag=10, fitdf=0)
```

```
Box-Pierce test
```

```
X-squared = 14.0451, df = 10, p-value = 0.1709
```

```
> Box.test(res, lag=10, fitdf=0, type="Ljung")
```

```
Box-Ljung test
```

```
X-squared = 14.4615, df = 10, p-value = 0.153
```

Exercise

- 1 Calculate the residuals from a seasonal naive forecast applied to the quarterly Australian beer production data from 1992.
- 2 Test if the residuals are white noise.

Exercise

- 1 Calculate the residuals from a seasonal naive forecast applied to the quarterly Australian beer production data from 1992.
- 2 Test if the residuals are white noise.

```
beer <- window(ausbeer, start=1992)
fc <- snaive(beer)
res <- residuals(fc)
Acf(res)
Box.test(res, lag=8, fitdf=0, type="Lj")
```

Outline

1 White noise

2 Time series decomposition

3 Seasonal adjustment

4 Forecasting and decomposition

Time series decomposition

$$Y_t = f(S_t, T_t, E_t)$$

where $Y_t =$ data at period t

$S_t =$ seasonal component at period t

$T_t =$ trend-cycle component at period t

$E_t =$ remainder (or irregular or error)
component at period t

Additive decomposition: $Y_t = S_t + T_t + E_t$.

Multiplicative decomposition: $Y_t = S_t \times T_t \times E_t$.

Time series decomposition

$$Y_t = f(S_t, T_t, E_t)$$

where $Y_t =$ data at period t

$S_t =$ seasonal component at period t

$T_t =$ trend-cycle component at period t

$E_t =$ remainder (or irregular or error)
component at period t

Additive decomposition: $Y_t = S_t + T_t + E_t$.

Multiplicative decomposition: $Y_t = S_t \times T_t \times E_t$.

Time series decomposition

$$Y_t = f(S_t, T_t, E_t)$$

where $Y_t =$ data at period t

$S_t =$ seasonal component at period t

$T_t =$ trend-cycle component at period t

$E_t =$ remainder (or irregular or error)
component at period t

Additive decomposition: $Y_t = S_t + T_t + E_t$.

Multiplicative decomposition: $Y_t = S_t \times T_t \times E_t$.

Time series decomposition

- Additive model appropriate if magnitude of seasonal fluctuations does not vary with level.
- If seasonal are proportional to level of series, then multiplicative model appropriate.
- Multiplicative decomposition more prevalent with economic series
- Logs turn multiplicative relationship into an additive relationship:

$$Y_t = S_t \times T_t \times E_t \quad \Rightarrow \quad \log Y_t = \log S_t + \log T_t + \log E_t.$$

Time series decomposition

- Additive model appropriate if magnitude of seasonal fluctuations does not vary with level.
- If seasonal are proportional to level of series, then multiplicative model appropriate.
- Multiplicative decomposition more prevalent with economic series
- Logs turn multiplicative relationship into an additive relationship:

$$Y_t = S_t \times T_t \times E_t \quad \Rightarrow \quad \log Y_t = \log S_t + \log T_t + \log E_t.$$

Time series decomposition

- Additive model appropriate if magnitude of seasonal fluctuations does not vary with level.
- If seasonal are proportional to level of series, then multiplicative model appropriate.
- Multiplicative decomposition more prevalent with economic series
- Logs turn multiplicative relationship into an additive relationship:

$$Y_t = S_t \times T_t \times E_t \quad \Rightarrow \quad \log Y_t = \log S_t + \log T_t + \log E_t.$$

Time series decomposition

- Additive model appropriate if magnitude of seasonal fluctuations does not vary with level.
- If seasonal are proportional to level of series, then multiplicative model appropriate.
- Multiplicative decomposition more prevalent with economic series
- Logs turn multiplicative relationship into an additive relationship:

$$Y_t = S_t \times T_t \times E_t \quad \Rightarrow \quad \log Y_t = \log S_t + \log T_t + \log E_t.$$

History of time series decomposition

- Classical method originated in 1920s.

In R: `decompose()`.

- Census II method introduced in 1957.

Basis for modern X-12-ARIMA method.

- STL method introduced in 1983. It only allows additive decomposition.

In R: `stl()`.

- TRAMO/SEATS introduced in 1990s.

History of time series decomposition

- Classical method originated in 1920s.
In R: `decompose()`.
- Census II method introduced in 1957.
Basis for modern X-12-ARIMA method.
- STL method introduced in 1983. It only allows additive decomposition.
In R: `stl()`.
- TRAMO/SEATS introduced in 1990s.

History of time series decomposition

- Classical method originated in 1920s.
In R: `decompose()`.
- Census II method introduced in 1957.
Basis for modern X-12-ARIMA method.
- STL method introduced in 1983. It only allows additive decomposition.
In R: `stl()`.
- TRAMO/SEATS introduced in 1990s.

History of time series decomposition

- Classical method originated in 1920s.
In R: `decompose()`.
- Census II method introduced in 1957.
Basis for modern X-12-ARIMA method.
- STL method introduced in 1983. It only allows additive decomposition.
In R: `stl()`.
- TRAMO/SEATS introduced in 1990s.

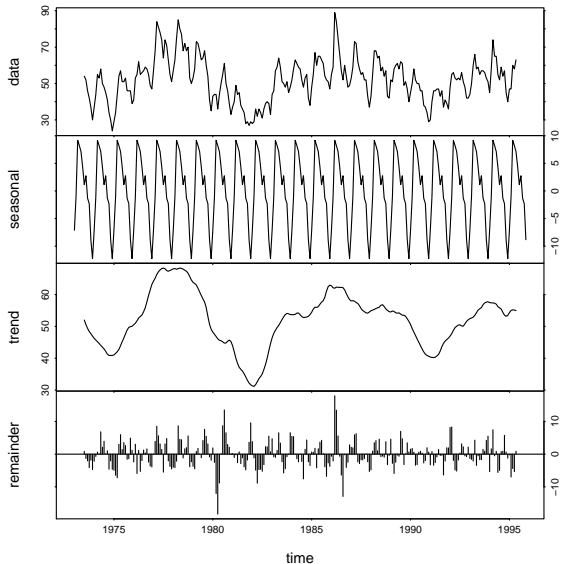
History of time series decomposition

- Classical method originated in 1920s.
In R: `decompose()`.
- Census II method introduced in 1957.
Basis for modern X-12-ARIMA method.
- STL method introduced in 1983. It only allows additive decomposition.
In R: `stl()`.
- TRAMO/SEATS introduced in 1990s.

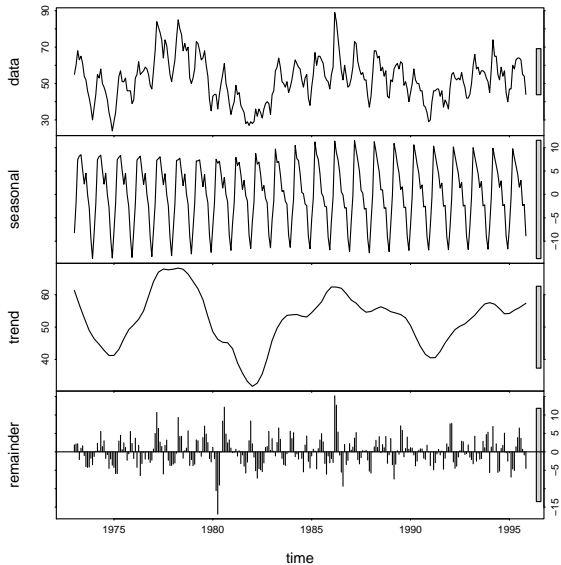
History of time series decomposition

- Classical method originated in 1920s.
In R: `decompose()`.
- Census II method introduced in 1957.
Basis for modern X-12-ARIMA method.
- STL method introduced in 1983. It only allows additive decomposition.
In R: `stl()`.
- TRAMO/SEATS introduced in 1990s.

Classical decomposition

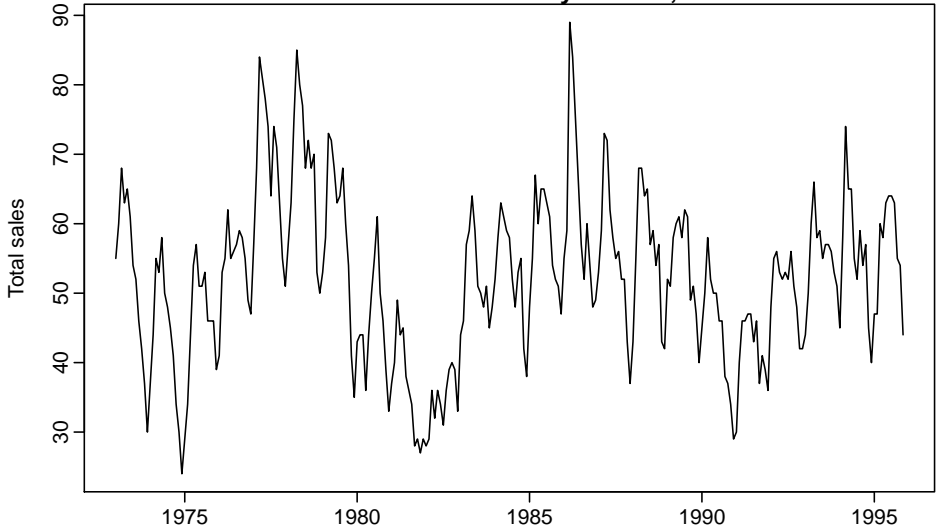


STL decomposition



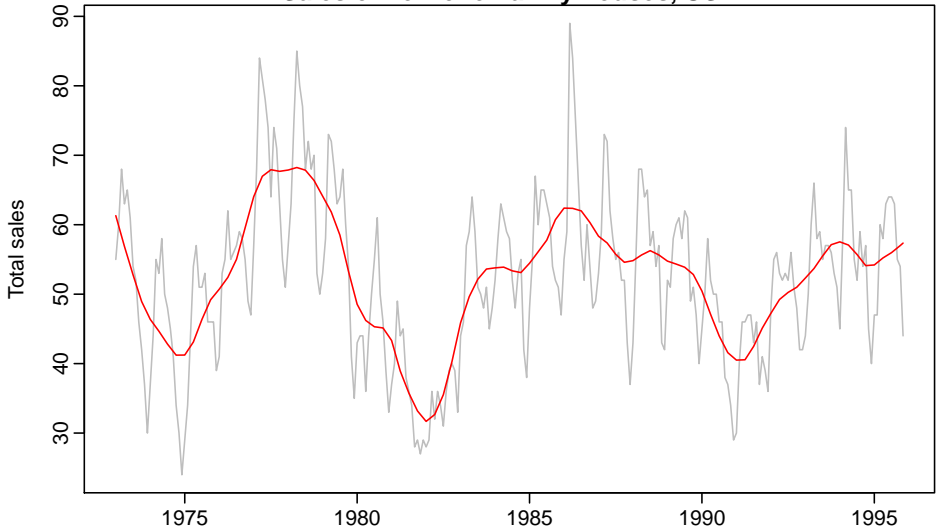
Examples: US house sales

Sales of new one-family houses, USA

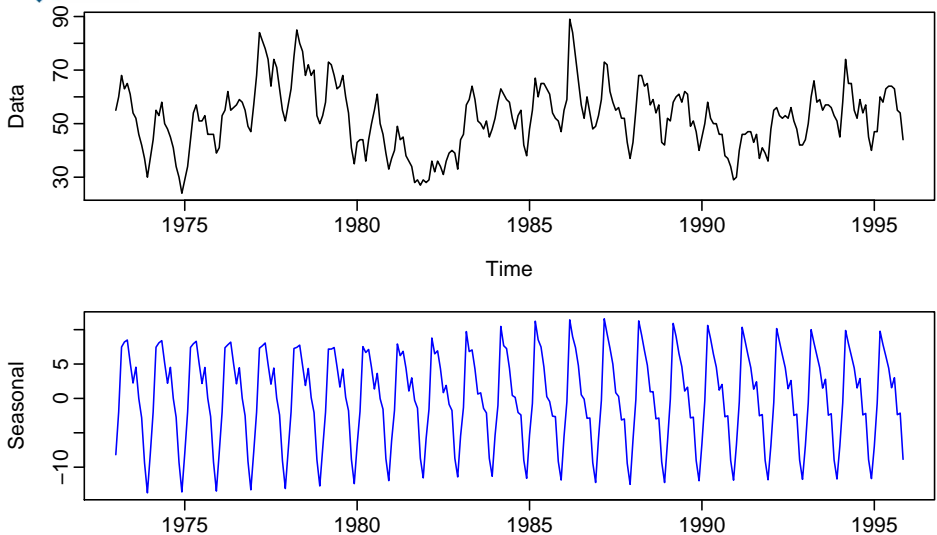


Examples: US house sales

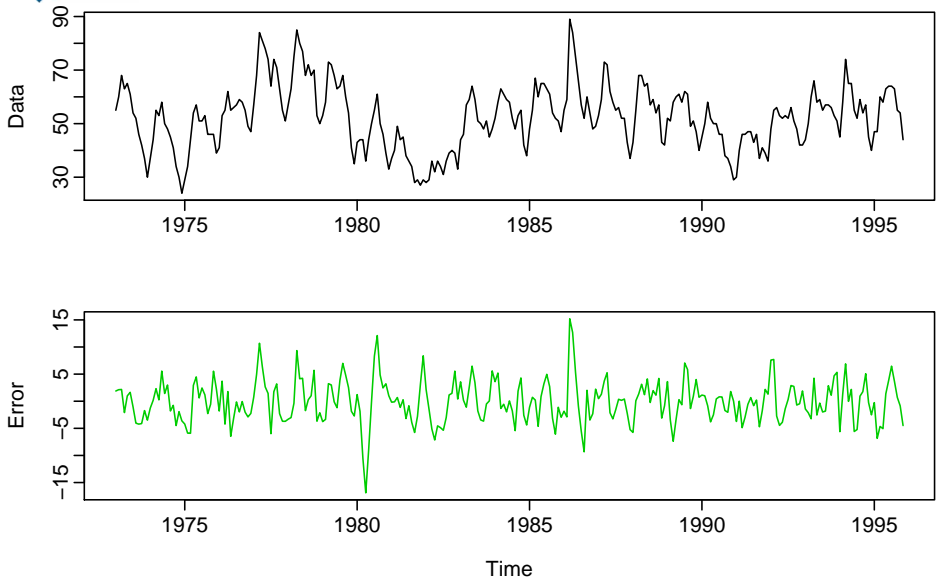
Sales of new one-family houses, USA



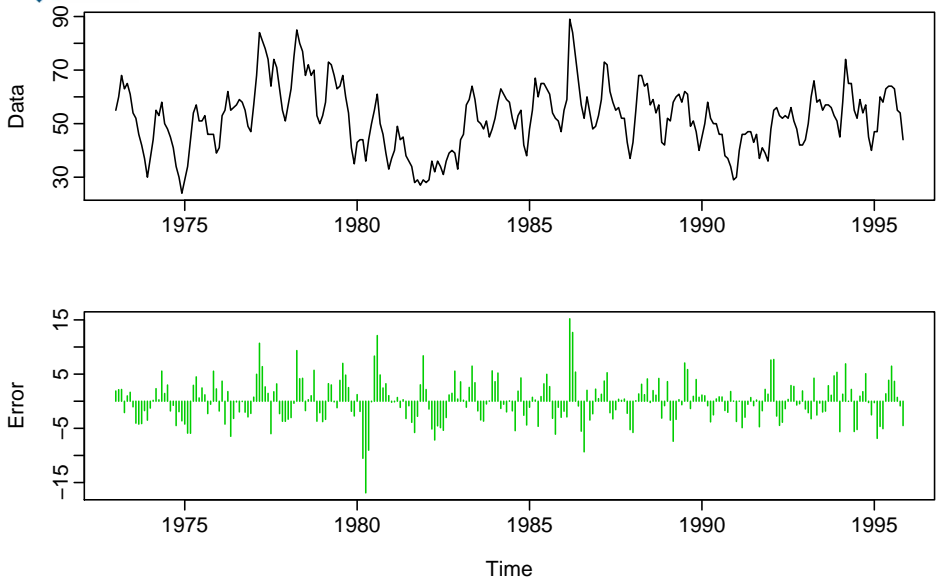
Examples: US house sales



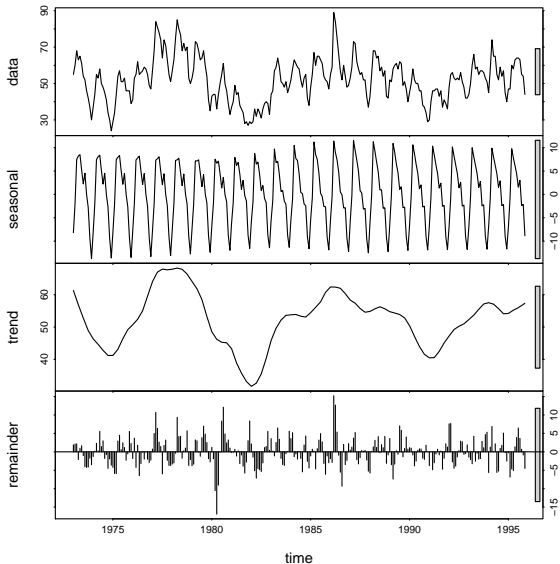
Examples: US house sales



Examples: US house sales



Examples: US house sales

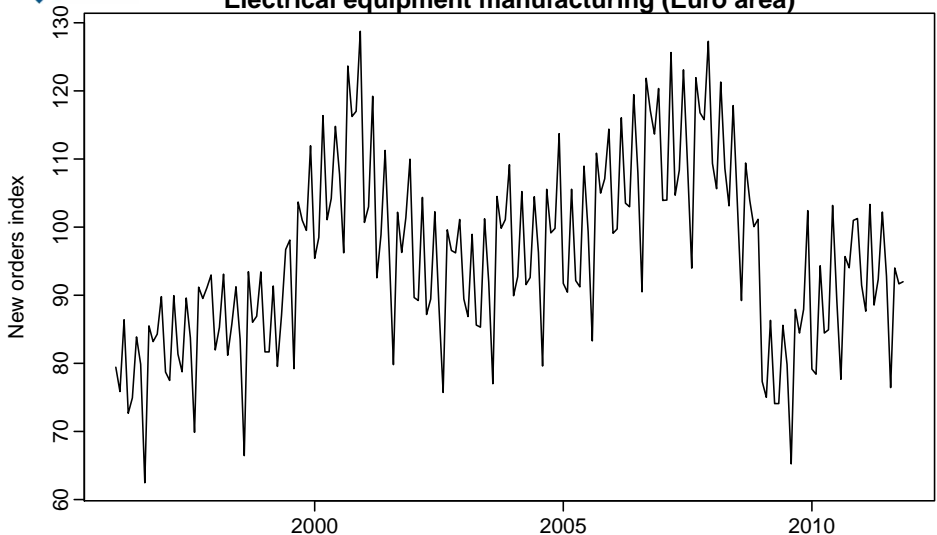


Time series decomposition in R

```
plot(decompose(hsales))  
plot(stl(hsales,s.window="periodic"))  
plot(stl(hsales,s.window=15))
```

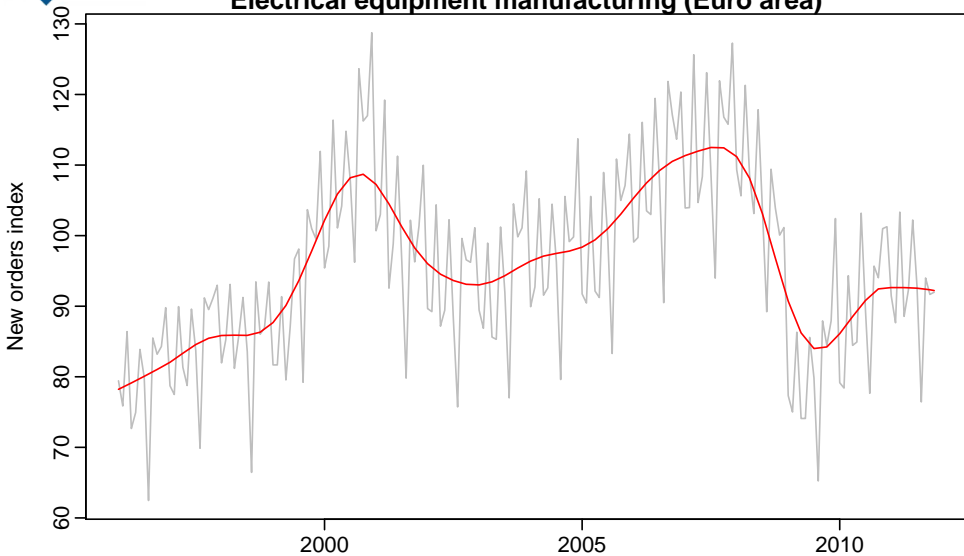

Euro electrical equipment

Electrical equipment manufacturing (Euro area)

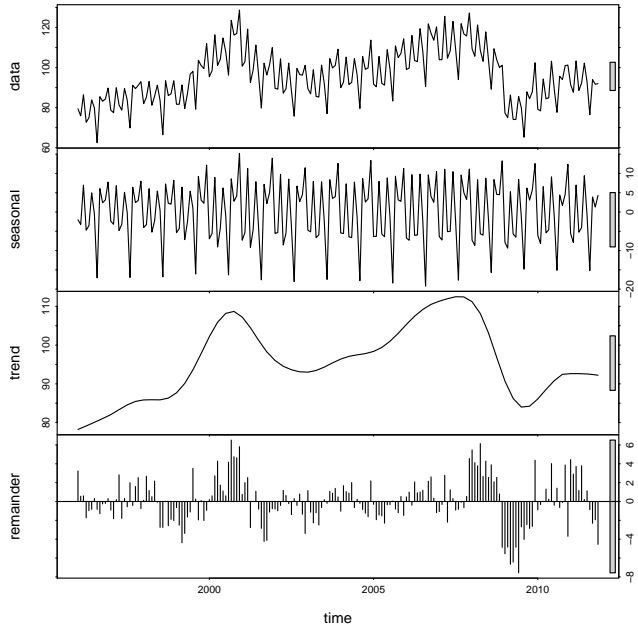


Euro electrical equipment

Electrical equipment manufacturing (Euro area)

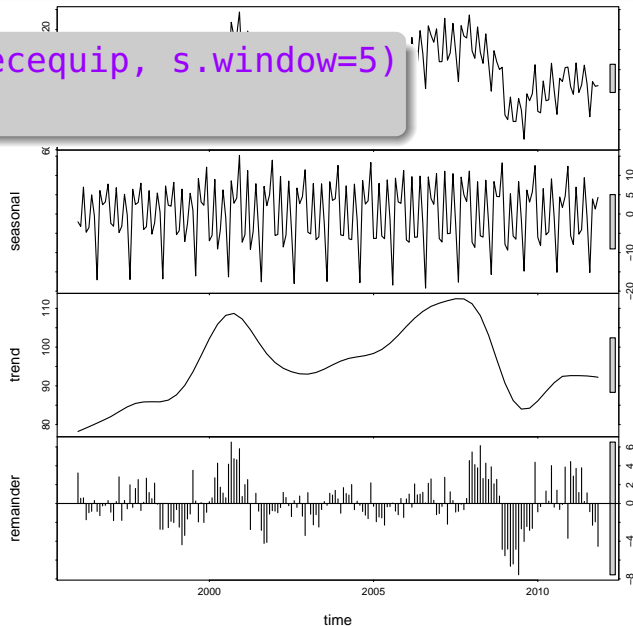


Euro electrical equipment

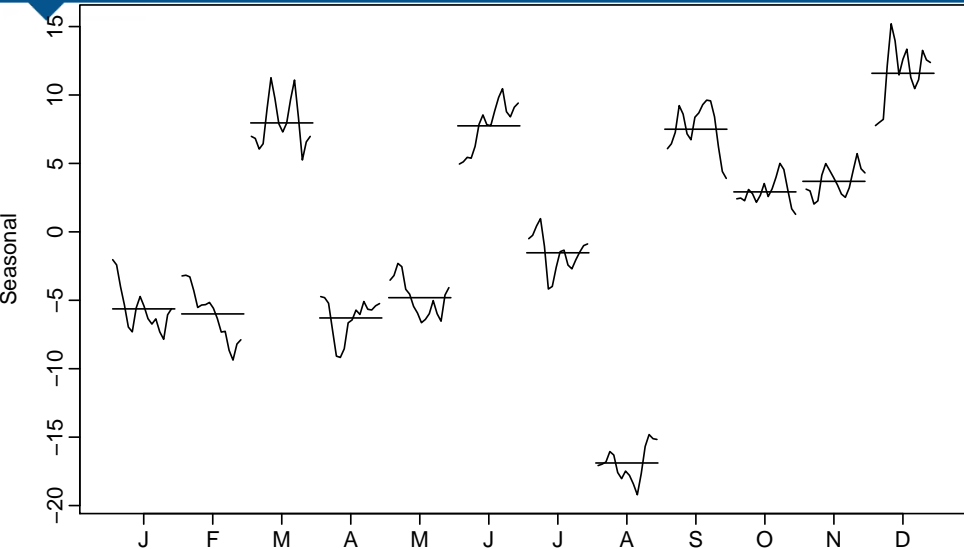


Euro electrical equipment

```
fit <- stl(elecequip, s.window=5)  
plot(fit)
```



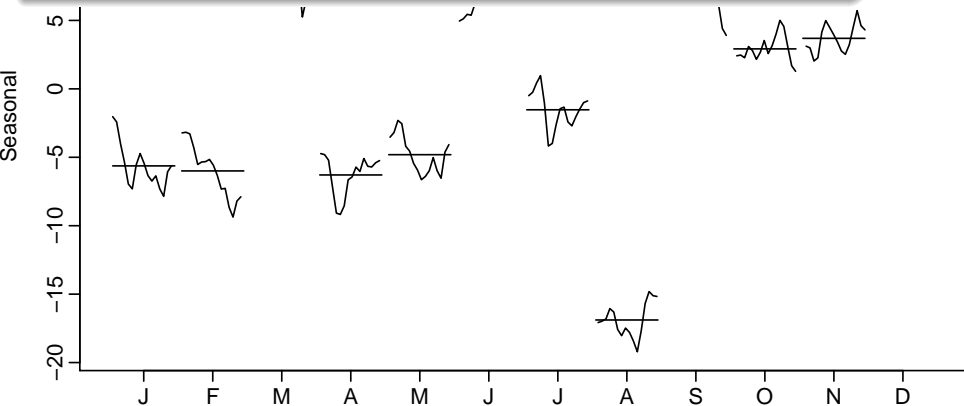
Euro electrical equipment



Seasonal sub-series plot of the seasonal component

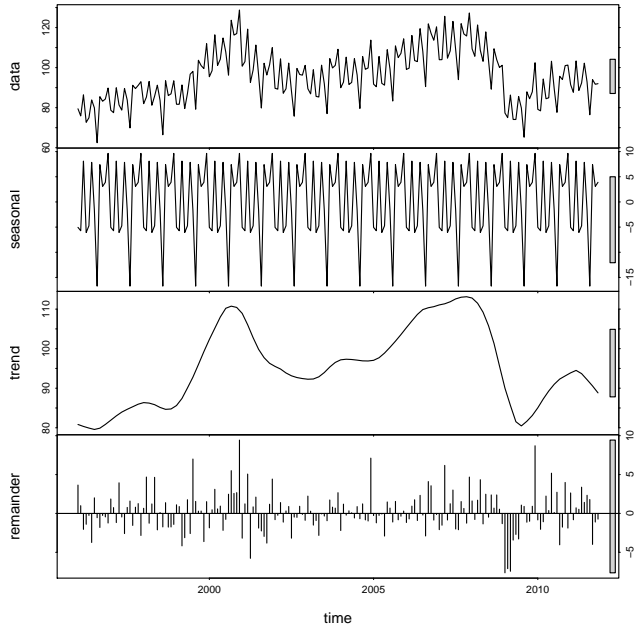
Euro electrical equipment

```
monthplot(fit$time.series[, "seasonal"],  
          ylab="Seasonal")
```



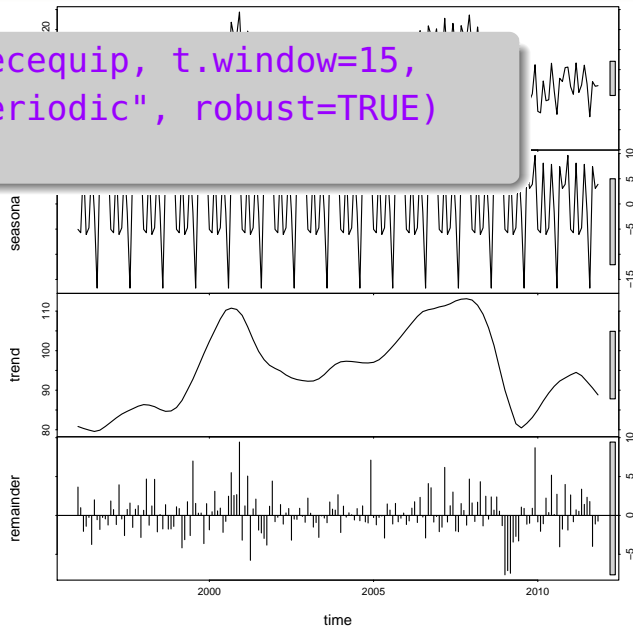
Seasonal sub-series plot of the seasonal component

Euro electrical equipment



Euro electrical equipment

```
fit <- stl(elecequip, t.window=15,  
  s.window="periodic", robust=TRUE)  
plot(fit)
```

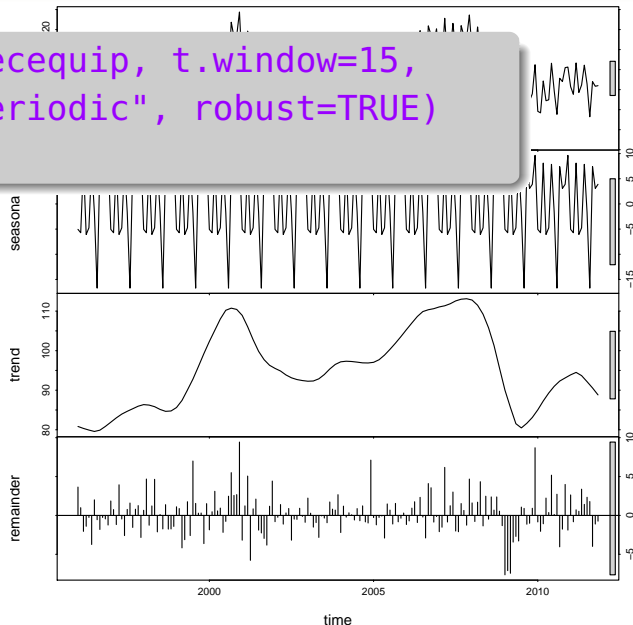


Euro electrical equipment

```
fit <- stl(elecequip, t.window=15,  
  s.window="periodic", robust=TRUE)  
plot(fit)
```

- `t.window` controls wiggleness of trend component.

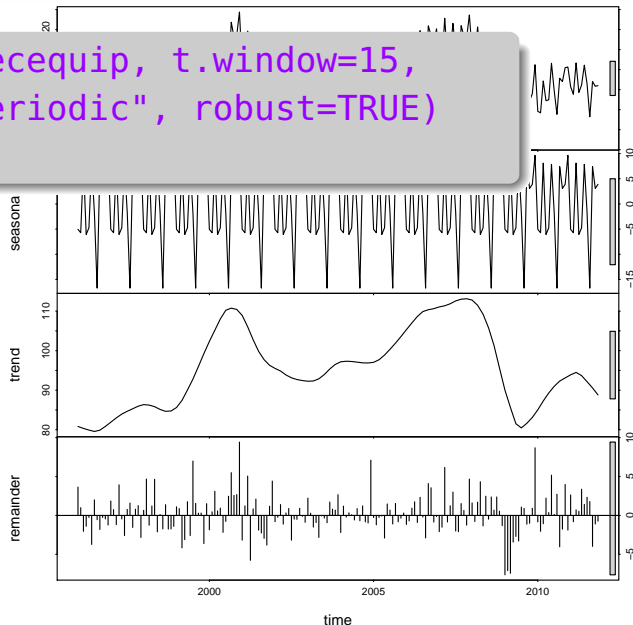
- `s.window` controls variation in seasonal component.



Euro electrical equipment

```
fit <- stl(elecequip, t.window=15,  
  s.window="periodic", robust=TRUE)  
plot(fit)
```

- **t.window**
controls
wiggleness of
trend
component.
- **s.window**
controls
variation in
seasonal
component.



Outline

- 1 White noise
- 2 Time series decomposition
- 3 Seasonal adjustment**
- 4 Forecasting and decomposition

Seasonal adjustment

- Useful by-product of decomposition: an easy way to calculate seasonally adjusted data.
- Additive decomposition: seasonally adjusted data given by

$$Y_t - S_t = T_t + E_t$$

- Multiplicative decomposition: seasonally adjusted data given by

$$Y_t/S_t = T_t \times E_t$$

Seasonal adjustment

- Useful by-product of decomposition: an easy way to calculate seasonally adjusted data.
- Additive decomposition: seasonally adjusted data given by

$$Y_t - S_t = T_t + E_t$$

- Multiplicative decomposition: seasonally adjusted data given by

$$Y_t / S_t = T_t \times E_t$$

Seasonal adjustment

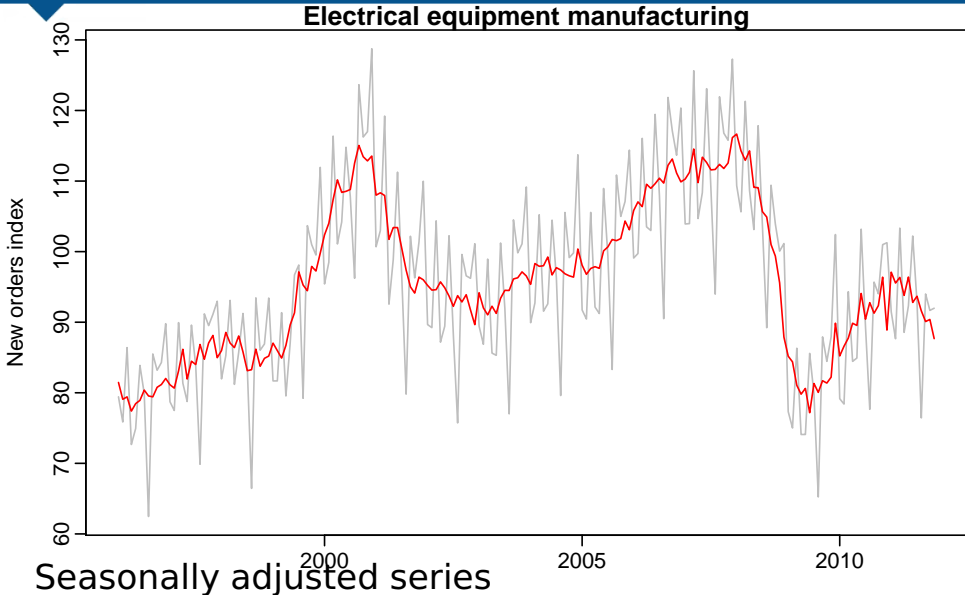
- Useful by-product of decomposition: an easy way to calculate seasonally adjusted data.
- Additive decomposition: seasonally adjusted data given by

$$Y_t - S_t = T_t + E_t$$

- Multiplicative decomposition: seasonally adjusted data given by

$$Y_t / S_t = T_t \times E_t$$

Euro electrical equipment



Seasonal adjustment in R

```
seasadj(obj)
```

where `obj` is the output from `stl()` or `decompose()`.

Example

```
plot(hsales,col="gray")  
fit <- stl(hsales,s.window=15)  
hsales.sa <- seasadj(fit)  
lines(hsales.sa, col="red")
```


Seasonal adjustment in R

```
seasadj(obj)
```

where `obj` is the output from `stl()` or `decompose()`.

Example

```
plot(hsales,col="gray")  
fit <- stl(hsales,s.window=15)  
hsales.sa <- seasadj(fit)  
lines(hsales.sa, col="red")
```

Outline

- 1 White noise
- 2 Time series decomposition
- 3 Seasonal adjustment
- 4 Forecasting and decomposition**

Forecasting and decomposition

- Forecast seasonal component by repeating the last year
- Forecast seasonally adjusted data using non-seasonal time series method. E.g.,
 - Holt's method — next topic
 - Random walk with drift model
- Combine forecasts of seasonal component with forecasts of seasonally adjusted data to get forecasts of original data.
- Sometimes a decomposition is useful just for understanding the data before building a separate forecasting model.

Forecasting and decomposition

- Forecast seasonal component by repeating the last year
- Forecast seasonally adjusted data using non-seasonal time series method. E.g.,
 - Holt's method — next topic
 - Random walk with drift model
- Combine forecasts of seasonal component with forecasts of seasonally adjusted data to get forecasts of original data.
- Sometimes a decomposition is useful just for understanding the data before building a separate forecasting model.

Forecasting and decomposition

- Forecast seasonal component by repeating the last year
- Forecast seasonally adjusted data using non-seasonal time series method. E.g.,
 - Holt's method — next topic
 - Random walk with drift model
- Combine forecasts of seasonal component with forecasts of seasonally adjusted data to get forecasts of original data.
- Sometimes a decomposition is useful just for understanding the data before building a separate forecasting model.

Forecasting and decomposition

- Forecast seasonal component by repeating the last year
- Forecast seasonally adjusted data using non-seasonal time series method. E.g.,
 - Holt's method — next topic
 - Random walk with drift model
- Combine forecasts of seasonal component with forecasts of seasonally adjusted data to get forecasts of original data.
- Sometimes a decomposition is useful just for understanding the data before building a separate forecasting model.

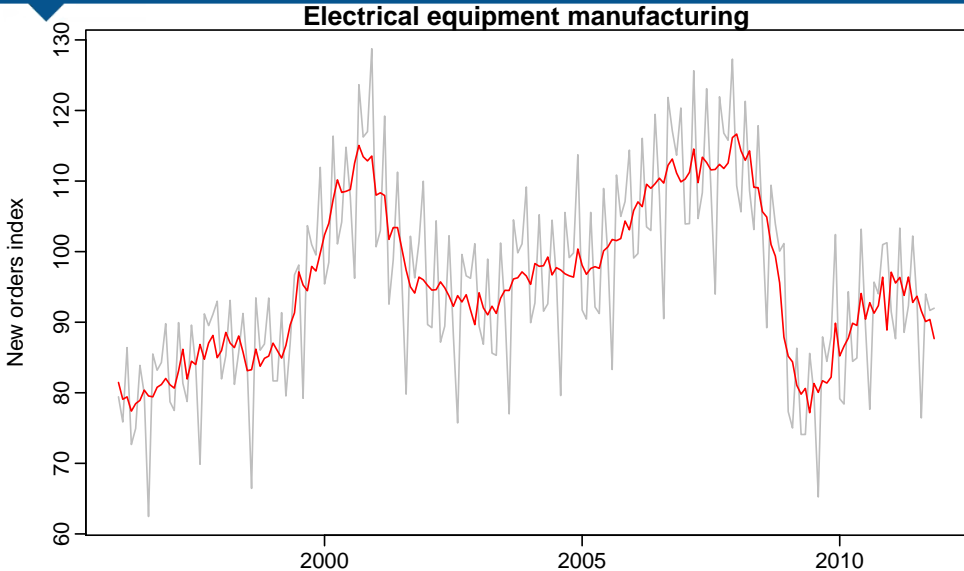
Forecasting and decomposition

- Forecast seasonal component by repeating the last year
- Forecast seasonally adjusted data using non-seasonal time series method. E.g.,
 - Holt's method — next topic
 - Random walk with drift model
- Combine forecasts of seasonal component with forecasts of seasonally adjusted data to get forecasts of original data.
- Sometimes a decomposition is useful just for understanding the data before building a separate forecasting model.

Forecasting and decomposition

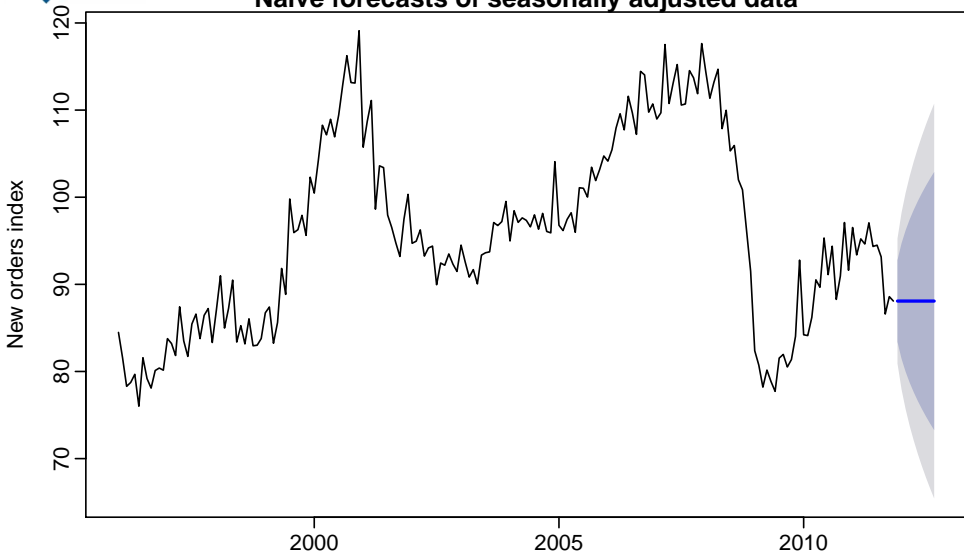
- Forecast seasonal component by repeating the last year
- Forecast seasonally adjusted data using non-seasonal time series method. E.g.,
 - Holt's method — next topic
 - Random walk with drift model
- Combine forecasts of seasonal component with forecasts of seasonally adjusted data to get forecasts of original data.
- Sometimes a decomposition is useful just for understanding the data before building a separate forecasting model.

Seas adj elec equipment



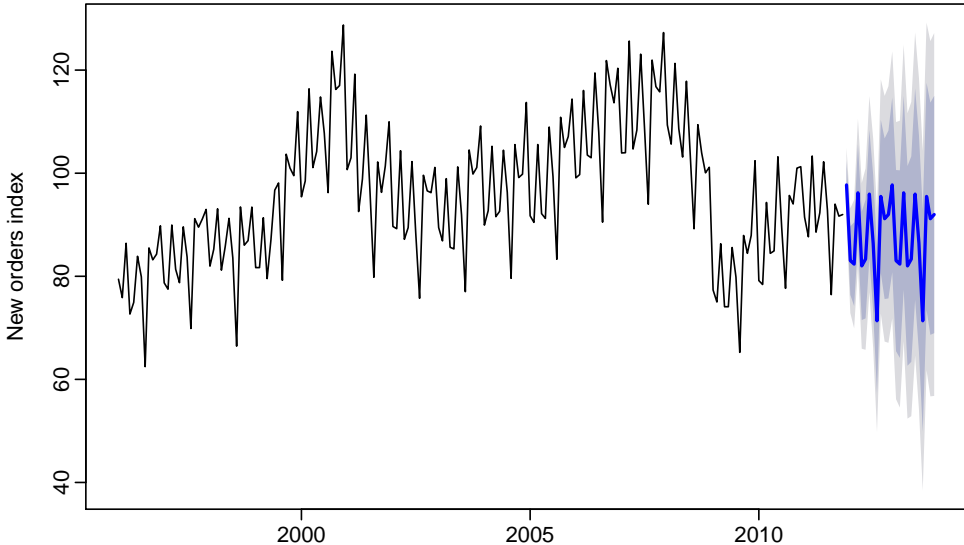
Seas adj elec equipment

Naive forecasts of seasonally adjusted data



Seas adj elec equipment

Forecasts from STL + Random walk



How to do this in R

```
fit <- stl(elecequip, t.window=15,  
  s.window="periodic", robust=TRUE)  
  
eeadj <- seasadj(fit)  
plot(naive(eeadj), xlab="New orders index")  
  
fcast <- forecast(fit, method="naive")  
plot(fcast, ylab="New orders index")
```

- It is common to take the prediction intervals from the seasonally adjusted forecasts and modify them with the seasonal component.
- This ignores the uncertainty in the seasonal component estimate.
- It also ignores the uncertainty in the future seasonal pattern.

Decomposition and prediction intervals

- It is common to take the prediction intervals from the seasonally adjusted forecasts and modify them with the seasonal component.
- This ignores the uncertainty in the seasonal component estimate.
- It also ignores the uncertainty in the future seasonal pattern.

- It is common to take the prediction intervals from the seasonally adjusted forecasts and modify them with the seasonal component.
- This ignores the uncertainty in the seasonal component estimate.
- It also ignores the uncertainty in the future seasonal pattern.