

MONASH BUSINESS SCHOOL

2017 Beijing Workshop on Forecasting

# Forecast Accuracy and Evaluation

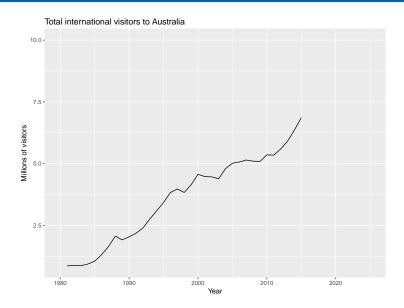
**Rob J Hyndman** 

robjhyndman.com/beijing2017

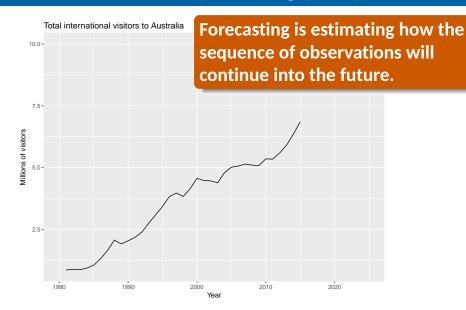
### **Outline**

- 1 The statistical forecasting perspective
- 2 Some simple forecasting methods
- 3 Forecasting residuals
- 4 Measuring forecast accuracy
- 5 Time series cross-validation
- 6 Probability scoring

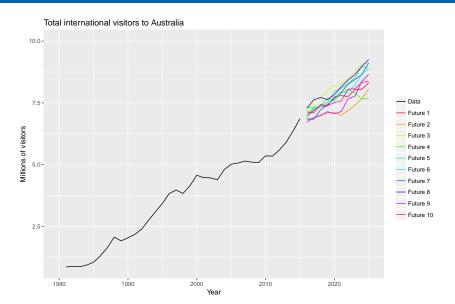
### The statistical forecasting perspective



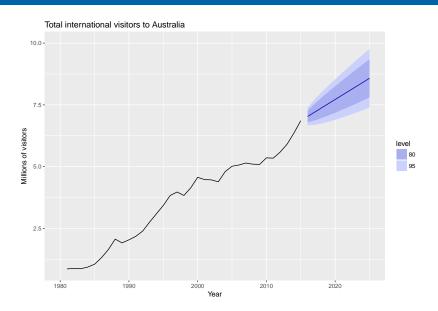
# The statistical forecasting perspective



# Sample futures



### **Forecast intervals**



### Statistical forecasting

■ Thing to be forecast: a random variable,  $y_t$ .

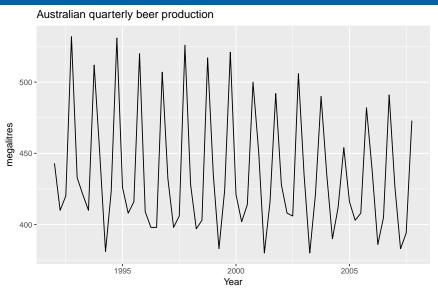
#### **Forecast distributions:**

$$y_{t|t-1} = y_t | \{y_1, y_2, \dots, y_{t-1}\}$$
  
 $y_{T+h|T} = y_{T+h} | \{y_1, y_2, \dots, y_T\}$ 

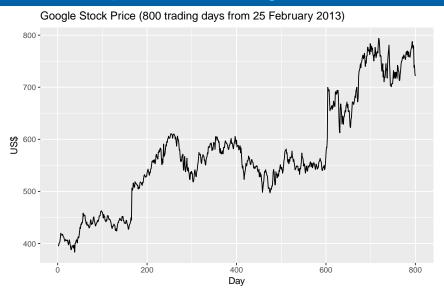
- The "point forecast" is the mean (or median) of  $y_{T+h|T} = y_{T+h} | \{y_1, y_2, \dots, y_T\}$
- The "forecast variance" is  $Var[y_{T+h}|y_1, y_2, ..., y_T]$
- A prediction interval or "interval forecast" is a range of values of  $y_t$  with high probability.

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#### Average method

- Forecast of all future values is equal to mean of historical data  $\{y_1, \dots, y_T\}$ .
- Forecasts:  $\hat{y}_{T+h|T} = \bar{y} = (y_1 + \cdots + y_T)/T$

#### Naïve method

- Forecasts equal to last observed value.
- Forecasts:  $\hat{y}_{T+h|T} = y_T$ .
- Consequence of efficient market hypothesis.

#### Seasonal naïve method

- Forecasts equal to last value from same season
- Forecasts:  $\hat{y}_{T+h|T} = y_{T+h-km}$  where m = seasonal period and k = |(h-1)/m|+1.

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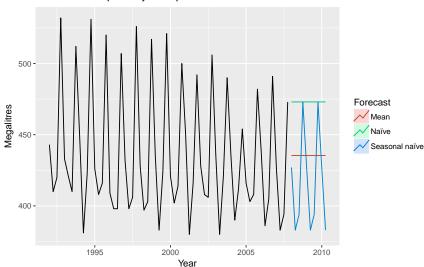
#### **Drift method**

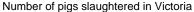
- Forecasts equal to last value plus average change.
- Forecasts:

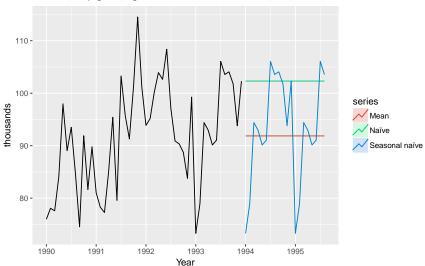
$$\hat{y}_{T+h|T} = y_T + \frac{h}{T-1} \sum_{t=2}^{T} (y_t - y_{t-1})$$
$$= y_T + \frac{h}{T-1} (y_T - y_1).$$

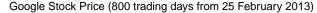
Equivalent to extrapolating a line drawn between first and last observations.













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### **Fitted values**

- $\hat{y}_{t|t-1}$  is the forecast of  $y_t$  based on observations  $y_1, \dots, y_{t-1}$ .
- We call these "fitted values".
- Sometimes drop the subscript:  $\hat{y}_t \equiv \hat{y}_{t|t-1}$ .
- Often not true forecasts since parameters are estimated on all data.

#### **Examples**:

- $\hat{y}_t = \bar{y}$  for average method.
- $\hat{y}_t = y_{t-1}$  for naive method
- $\hat{y}_t = y_{t-1} + (y_T y_1)/(T 1)$  for drift method.
- $\hat{y}_t = y_{t-m}$  for seasonal naive method

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**Residuals in forecasting:** difference between observed value and its fitted value:  $e_t = y_t - \hat{y}_{t|t-1}$ .

#### **Assumptions**

- $\{e_t\}$  uncorrelated. If they aren't, then information left in residuals that should be used in computing forecasts.
- $\{e_t\}$  have mean zero. If they don't, then forecasts are biased.

#### **Useful properties** (for prediction intervals)

- $\{e_t\}$  have constant variance.
- $\{e_t\}$  are normally distributed

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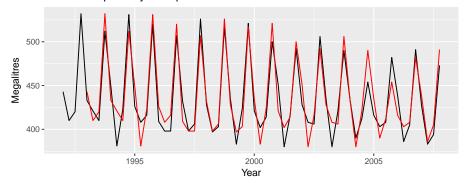
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#### Seasonal naïve forecast:

$$\hat{y}_{t|t-1} = y_{t-4}$$
  $e_t = y_t - y_{t-4}$ 

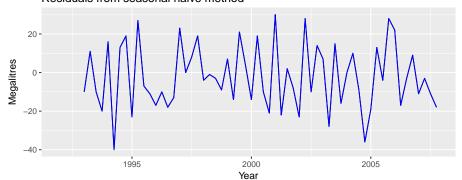
#### Australian quarterly beer production



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Residuals from seasonal naive method

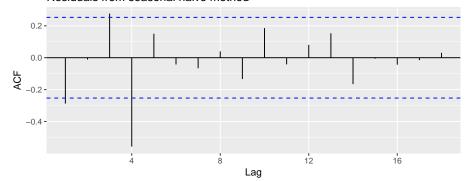


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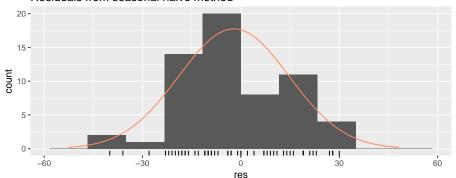
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$$\hat{y}_{t|t-1} = y_{t-4}$$
  $e_t = y_t - y_{t-4}$ 

#### Residuals from seasonal naive method



- Minimizing the size of forecasting residuals is used for estimating model parameters (e.g., minimizing MSE or maximizing likelihood.
- In general, forecasting residuals cannot be used (directly) for estimating forecast accuracy.
- Forecast accuracy can only be measured using genuine forecasts; i.e., on different data.
- Forecasting residuals can help suggest model improvements.

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### **Training and test sets**



- The test set must not be used for *any* aspect of model development or calculation of forecasts.
- Forecast accuracy is based only on the test set.
- A model which fits the data well does not necessarily forecast well.
- A perfect fit can always be obtained by using a model with enough parameters. (Compare  $R^2$ )
- Problems can be overcome by measuring true out-of-sample forecast accuracy. Training set used to estimate parameters. Forecasts are made for test set.

Training set: T observations

Test set: H observations

$$\begin{aligned} \text{MAE} &= \frac{1}{H} \sum_{h=1}^{H} |y_{T+h} - \hat{y}_{T+h|T}| \\ \text{MSE} &= \frac{1}{H} \sum_{h=1}^{H} (y_{T+h} - \hat{y}_{T+h|T})^2 \quad \text{RMSE} \quad = \sqrt{\frac{1}{H} \sum_{h=1}^{H} (y_{T+h} - \hat{y}_{T+h|T})^2} \\ \text{MAPE} &= \frac{100}{H} \sum_{h=1}^{H} |y_{T+h} - \hat{y}_{T+h|T}| / |y_{T+h}| \end{aligned}$$

- MAE, MSE, RMSE are all scale dependent.
- MAPE is scale independent but is only sensible if  $y_{T+h} \gg 0$  for all h, and v has a natural zero.

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#### **Mean Absolute Scaled Error**

MASE = 
$$\frac{1}{H} \sum_{h=1}^{H} |y_{T+h} - \hat{y}_{T+h|T}|/Q$$

where Q is a stable measure of the scale of the time series  $\{y_t\}$ .

Proposed by Hyndman and Koehler (IJF, 2006).

For non-seasonal time series,

$$Q = \frac{1}{T-1} \sum_{t=2}^{I} |y_t - y_{t-1}|$$

works well. Then MASE is equivalent to MAE relative to a naïve method.

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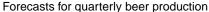
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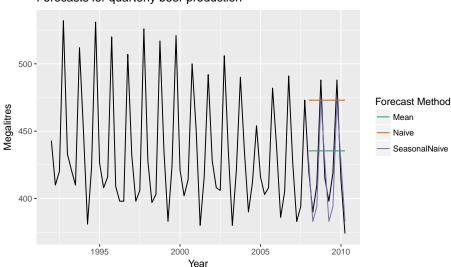
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For seasonal time series,

$$Q = \frac{1}{T - m} \sum_{t=m+1}^{T} |y_t - y_{t-m}|$$

works well. Then MASE is equivalent to MAE relative to a seasonal naïve method.





# Measures of forecast accuracy

	RMSE	MAE	MAPE	MASE
Mean method	38.5	34.8	8.28	2.44
Naïve method	62.7	57.4	14.18	4.01
Seasonal naïve method	14.3	13.4	3.17	0.94

# Measures of forecast accuracy

Scaling can be used with any measure, and with different scaling statistics.

### **Mean Squared Scaled Error**

MASE = 
$$\frac{1}{H} \sum_{h=1}^{H} (y_{T+h} - \hat{y}_{T+h|T})^2 / Q$$
  
ere  $Q = \frac{1}{T - m} \sum_{h=1}^{T} (y_t - y_{t-m})^2$ 

- Assumes  $\{y_t\}$  is difference stationary.
- Minimizing MSSE leads to conditional mean forecasts.
- MSSE < 1 : out-of-sample multi-step forecasts are more accurate than in-sample one-step forecasts.

# Measures of forecast accuracy

- Many suggested scale-free measures of forecast accuracy are degenerate due to infinite variance.
- The denominator must be positive with probability one.
- Distribution of most measures are highly skewed when applied to real data.

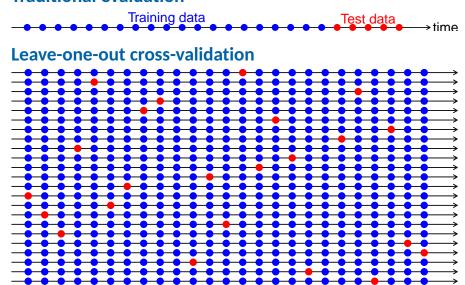
## Poll: true or false?

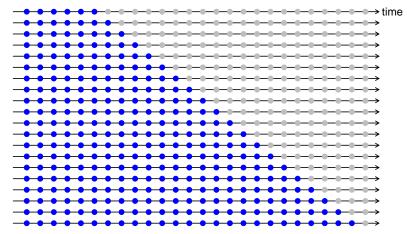
- Good forecast methods should have normally distributed residuals.
- A model with small residuals will give good forecasts.
- The best measure of forecast accuracy is MAPE.
- If your model doesn't forecast well, you should make it more complicated.
- Always choose the model with the best forecast accuracy as measured on the test set.

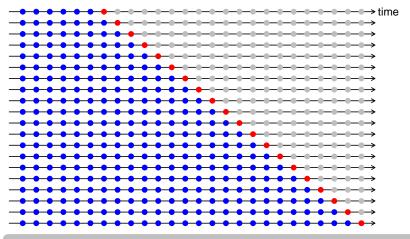
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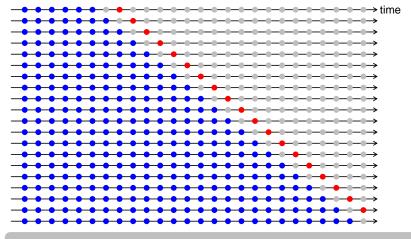
## **Traditional evaluation**



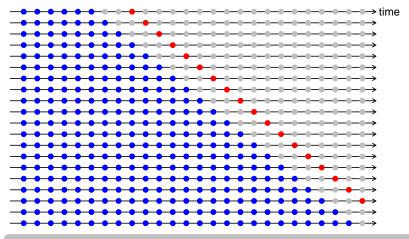




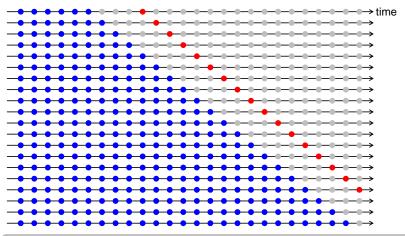
$$h = 1$$



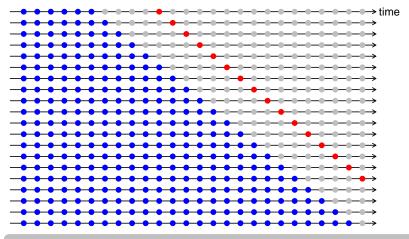
$$h = 2$$



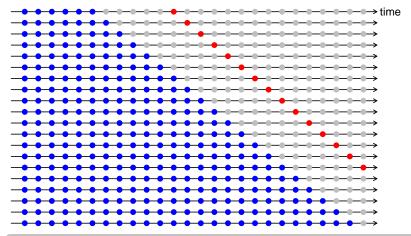
$$h = 3$$



$$h = 4$$

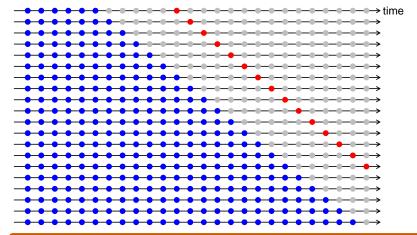


$$h = 5$$



$$h = 6$$

### Time series cross-validation

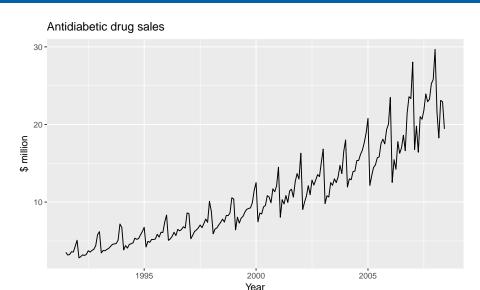


Also known as "Evaluation on a rolling forecast origin"

### Time series cross-validation

Assume *k* is the minimum number of observations for a training set.

- Select observation t + h for test set, and use observations at times  $1, 2, \ldots, t$  to estimate model.
- **Compute error on forecast for time** t + h**.**
- Repeat for t = k, k + 1, ..., T h where T is total number of observations.
- Compute accuracy measure over all errors.

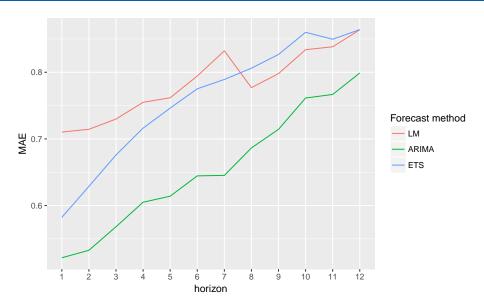


### Which of these models is best?

- Linear model with trend and seasonal dummies applied to log data.
- ARIMA model applied to log data
- ETS model applied to original data
- Forecast h = 12 steps ahead based on data to time t for t = 1, 2, ...
- Compare MAE values for each forecast horizon.

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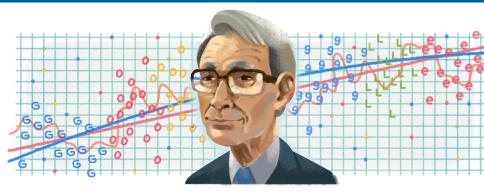
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# Example: R code

```
f1 <- function(y,h) {
  forecast(tslm(y ~ trend + season, lambda=0), h=h)
}
f2 <- function(y,h) {
  forecast(auto.arima(y, D=1, lambda=0), h=h)
f3 <- function(y,h) {
  forecast(ets(y), h=h)
}
e1 \leftarrow tsCV(a10, f1, h=12)
e2 \leftarrow tsCV(a10, f2, h=12)
e3 \leftarrow tsCV(a10, f3, h=12)
mae <- data.frame(</pre>
    LM = colMeans(abs(e1), na.rm=TRUE),
    ARIMA = colMeans(abs(e2), na.rm=TRUE),
    ETS = colMeans(abs(e3), na.rm=TRUE))
```

# Hirotugu Akaike (1927-2009)



Akaike, H. (1974), "A new look at the statistical model identification", *IEEE Transactions on Automatic Control*, **19**(6): 716–723.

$$AIC = -2\log(L) + 2k$$

where *L* is the model likelihood and *k* is the number of estimated parameters in the model.

■ If *L* is Gaussian, then AIC  $\approx c + T \log MSE + 2k$  where *c* is a constant, MSE is from one-step forecasts on **training set**, and *T* is the length of the series.

- AICc a bias-corrected small-sample version.
- AIC/AICc much faster than CV

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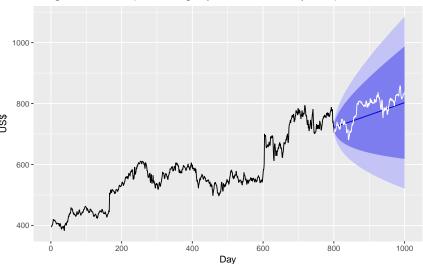
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# **Probablistic forecasting**

Google Stock Price (800 trading days from 25 February 2013)



# **Probabilistic forecasting**

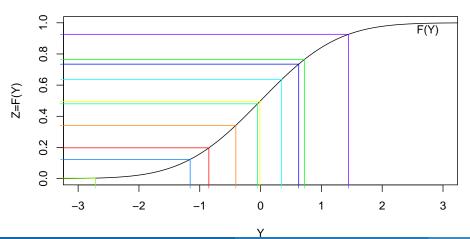
How to evaluate a forecast probability distribution?

- Forecast intervals: percentage of observations covered compared to nominal percentage.
- Density forecasting
- Quantile forecasting
- Distribution forecasting

# **Probability Integral Transform**

Let F = cdf of Y and Z = F(Y). If F is continuous, then Z is standard uniform.

#### **Probability Integral Transform**



### **Calibration**

 $Y_{T+h|T}$  has cdf  $F_{T+h|T}$ .  $\hat{F}_{T+h|T}$  is our forecast cdf.

### **Calibration**

- (a)  $\hat{F}$  is marginally calibrated if  $E[\hat{F}(y)] = P(Y \leq y) \ \forall y \in \mathbb{R}$ .
- **(b)**  $\hat{F}$  is probabilistically calibrated if  $Z = \hat{F}(Y)$  has a standard uniform distribution.
- We could plot a histogram of  $Z = \hat{F}(Y)$  and check that it looks uniform.
- → This is a more sophisticated version of testing if prediction intervals have the correct coverage.

# **Sharpness**

 $Y_{T+h|T}$  has cdf  $F_{T+h|T}$ .  $\hat{F}_{T+h|T}$  is our forecast cdf.

### **Sharpness**

- → A "sharp" forecast distribution has narrow prediction intervals.
  - A good probabilistic forecast is both calibrated and sharp.
  - Scoring rules combine calibration and sharpness in a single measure.

 $Y_{T+h|T}$  has cdf  $F_{T+h|T}$ .  $\hat{F}_{T+h|T}$  is our forecast cdf. A scoring rule assigns numerical score  $S(\hat{F}_{T+h|T}, y_{T+h})$ .

Dawid-Sebastiani score:

$$DSS(\hat{F}, y) = \frac{(y - \mu_{\hat{F}})^2}{\sigma_{\hat{F}}^2} + 2\log\sigma_{\hat{F}}$$

Generalization of MSE assuming normality.

A "proper" scoring rule has the property:

$$\mathsf{E}_{\mathsf{F}}[S(\mathsf{F},\mathsf{Y})] < \mathsf{E}_{\mathsf{F}}[S(\hat{\mathsf{F}},\mathsf{Y})]$$

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, y) =  $\frac{(y - \mu_{\hat{F}})^2}{\sigma_{\hat{F}}^2}$  +  $2 \log \sigma_{\hat{F}}$ 

Generalization of MSE assuming normality.

A "proper" scoring rule has the property:

$$E_F[S(F, Y)] < E_F[S(\hat{F}, Y)]$$

 $Y_{T+h|T}$  has cdf  $F_{T+h|T}$ .  $\hat{F}_{T+h|T}$  is our forecast cdf. A scoring rule assigns numerical score  $S(\hat{F}_{T+h|T}, y_{T+h})$ .

### **Continuous Ranked Probability Score:**

CRPS(
$$\hat{F}, y$$
) =  $\int [\hat{F}(x) - 1_{\{y \le x\}}]^2 dx = E_{\hat{F}}|Y - y| - \frac{1}{2}E_{\hat{F}}|Y - Y'|$  where Y and Y' have cdf  $\hat{F}$ . Generalization of MAE.

### **Continuous Ranked Probability Score:**

Let  $\hat{Q}_{T+h|T} = \hat{F}_{T+h|T}^{-1}$  be the forecast quantile function

CRPS(
$$\hat{Q}, y$$
) = 2  $\int_{0}^{1} [\hat{Q}(p) - y] [1_{\{y < \hat{Q}(p)\}} - p] dp$ 

 $Y_{T+h|T}$  has cdf  $F_{T+h|T}$ .  $\hat{F}_{T+h|T}$  is our forecast cdf. A scoring rule assigns numerical score  $S(\hat{F}_{T+h|T}, y_{T+h})$ .

### **Energy Score**

$$ES(\hat{F}, y) = E_{\hat{F}}|Y - y|^{\alpha} - \frac{1}{2}E_{\hat{F}}|Y - Y'|^{\alpha}$$

where Y and Y' have cdf  $\hat{F}$  and  $\alpha \in (0, 2]$ .

### **Log Score**

$$\log S(\hat{F}, y) = -\log \hat{f}(y)$$

where  $\hat{f} = d\hat{F}/dy$  is density corresponding to  $\hat{F}$ .

# R packages



https://github.com/earowang/tsibble

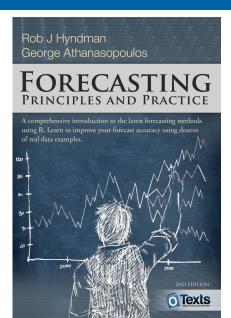
http://pkg.earo.me/sugrrants

https://github.com/mitchelloharawild/fasster

http://pkg.robjhyndman.com/forecast

http://pkg.earo.me/hts

## **Textbook**



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