

Rob J Hyndman

Functional time series

with applications in demography

5. Forecasting functional time series via PLS

Outline

- 1 Functional Partial Least Squares
- 2 Application: French mortality rates
- 3 Application: Australian fertility rates
- 4 Forecast accuracy comparisons
- **5** Bootstrap intervals
- **6** Comparisons
- 7 References

- PCA components are designed to explain historical variation. They do not necessarily provide the best predictors.
- Partial least squares extracts uncorrelated latent components cores by maximizing the covariance between predictors and response.
- Response is $s_t(x)$ and predictor is $s_{t-1}(x)$.
- We want to predict $s_t(x)$ using

- This is a functional ARH(1).
- How to choose b(x, u)?

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$$\hat{\boldsymbol{f}}^*(x) = \sum_{k=1}^{\infty} \beta_k \psi_k(x)$$
, and $\hat{\boldsymbol{g}}^*(x) = \sum_{k=1}^{\infty} \beta_k \phi_k(x)$

- $\mathbf{W} = \text{diagonal}(w_1, \dots, w_T)$, $w_t = \kappa (1 \kappa)^{T-t}$
- $\mathbf{f}^*(x) = \mathbf{W}[s_1^*(x), \dots, s_{T-1}^*(x)]'$ and $\mathbf{g}^*(x) = \mathbf{W}[s_2^*(x), \dots, s_T^*(x)]'$ are weighted decentralized functional predictors and responses
- lacksquare eta_k denotes common latent component scores
- $\psi_k(x)$ and $\phi_k(x)$ are latent components of predictors and responses respectively

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- **1** Let $\mathbf{f}_0^*(x) = \mathbf{f}^*(x)$ and $\mathbf{g}_0^*(x) = \mathbf{g}^*(x)$
- Obtain $w_k(x)$ iteratively, starting with $w_k^{(0)}(x) = 1$: $w_k^{(i)}(x) = \iint w_k^{(i-1)}(v) [\hat{\boldsymbol{f}}_{k-1}^*(v)]' \hat{\boldsymbol{g}}_{k-1}^*(u) [\hat{\boldsymbol{g}}_{k-1}^*(u)]' \hat{\boldsymbol{f}}_{k-1}^*(x) \, dv \, du$
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- Discretize $s_t^*(x)$ on a dense grid of q equally spaced points.
- Denote discretized $s_t^*(x)$ as $T \times q$ matrix G^* and let $G = WG^*$.
- Define G_{k-1} and F_{k-1} analogously
- $\mathbf{w}_1(x) = \text{largest eigenvector of } \mathbf{G}' \mathbf{F}$
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Functional autoregression coefficient:

$$b(x,u) = \sum_{k=1}^{\infty} w_k(u) \phi_k(x)$$

By orthogonality of β_k :

$$\phi_k(\mathbf{x}) = (\beta_k' \beta_k)^{-1} \beta_k' \hat{\mathbf{g}}^*(\mathbf{x}).$$

Therefore

$$\hat{b}(x,u) = \sum_{k=1}^K w_k(u) (\beta_k' \beta_k)^{-1} \beta_k' \hat{\boldsymbol{g}}^*(x),$$

for some finite *K*.

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Forecasted curves:

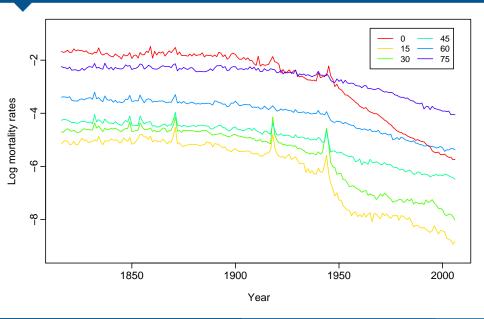
$$\hat{f}_{t+1|t}(x) = \hat{\mu}(x) + \int [\hat{f}_t(u) - \hat{\mu}(u)] \hat{b}(x,u) du.$$

For $\hat{f}_{t+h|t}(x)$ where h > 1, apply iteratively.

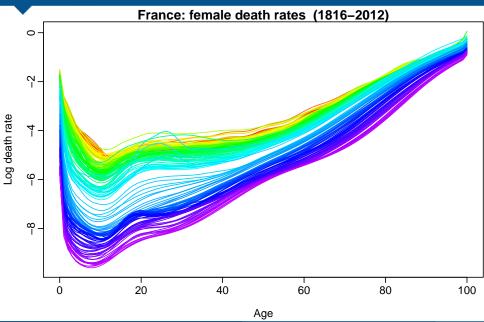
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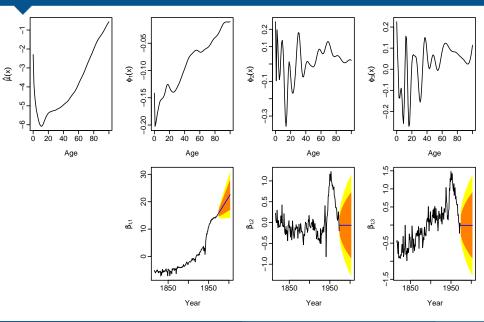
French female mortality rates



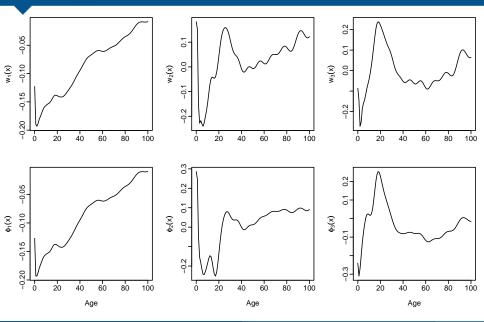
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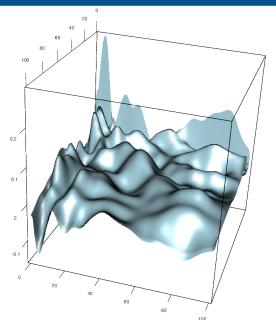
French mortality rate models



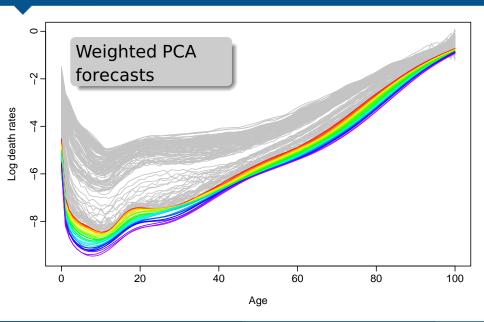
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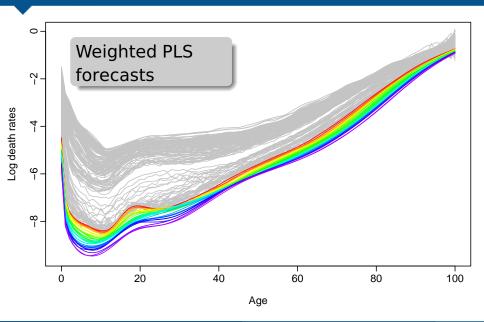
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French mortality rate forecasts



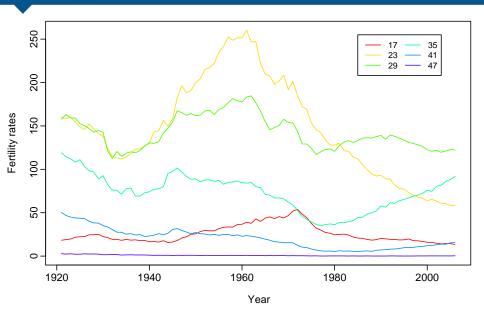
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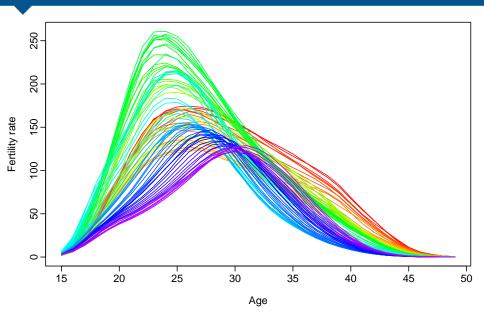
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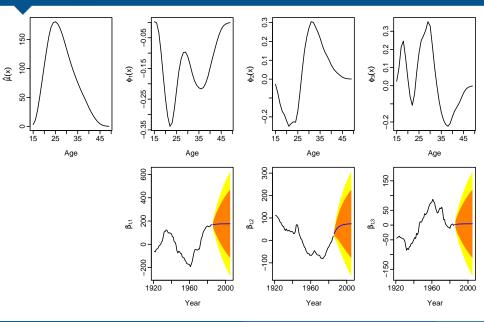
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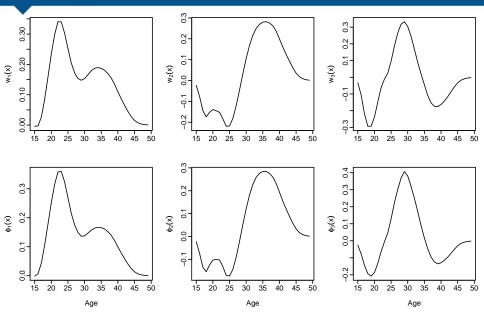
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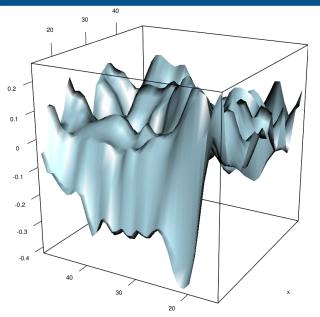
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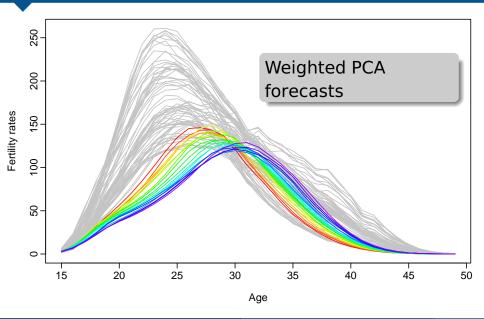
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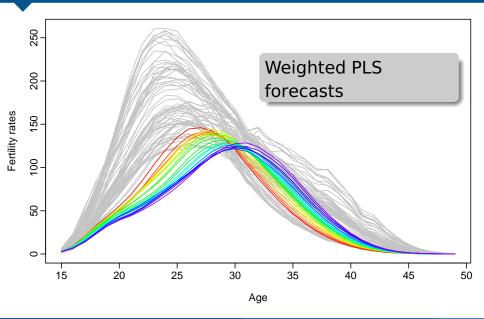
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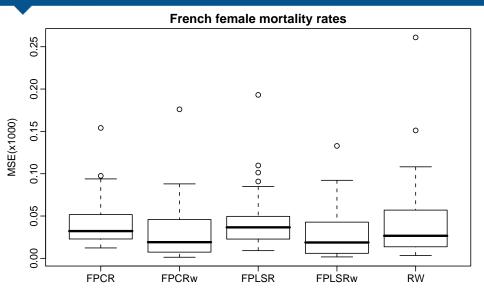
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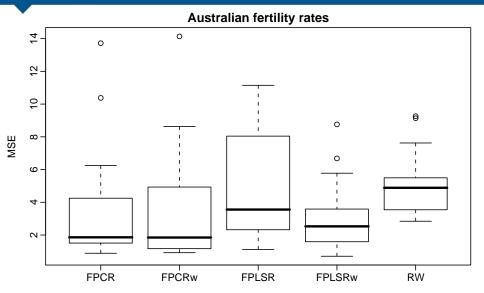
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MSE: French female mortality rates (x1000)

K	FPC	FPC _w	FPLSR	FPLSR _W	RW
1	0.5956	0.0293	0.5994	0.0607	
2	0.0537	0.0310	0.0738	0.0288	
3	0.0316	0.0310	0.0445	0.0288	
4	0.0296	0.0311	0.0428	0.0288	
5	0.0287	0.0311	0.0472	0.0297	
6	0.0425	0.0311	0.0474	0.0291	0.0437



MSE: Australian fertility rates

K	FPC	\mathbf{FPC}_{w}	FPLSR	$FPLSR_{\scriptscriptstyle{W}}$	RW
1	99.0611	16.7304	94.0311	53.8186	
2	56.3095	3.3019	54.3410	17.5883	
3	24.9330	3.2580	26.0428	10.2599	
4	15.6845	3.1995	19.7227	4.4818	
5	4.4495	3.2132	5.9299	4.0573	
6	3.4310	3.2123	4.9205	2.9046	4.9800

Computation time

- Weighted FPLSR more efficient than weighted FPC as FPC requires many univariate time series models.
- Time to fit 100 replications:

Method	Fertility data	Mortality data
FPC	34.1072	62.2797
FPC_w	33.1424	60.8426
FPLSR	0.4287	2.9184
FPLSR _w	0.4537	3.1602
RW	0.0000	0.0002

(Intel Xeon 2.33GHz processor)

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Functional PLS

- **I** smoothing error in estimating $s_t(x)$
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- Let $\hat{\xi}_{t,k} = \hat{\beta}_{t,k} \hat{\beta}_{t|t-1,k}$ be 1-step errors of PC scores.
- $lacksquare \{\xi_k^{(\ell)}\}$ sampled with replacement from $\{\hat{\xi}_{t,k}\}$.
- Simulate future sample paths of $\beta_{T+h|T,k}$ using these bootstrapped residuals: $\{\beta_{T+h|T,k}^{(\ell)}\}$.
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$$\frac{1}{mph} \sum_{t=T-m+1}^{T} \sum_{j=1}^{h} \sum_{i=1}^{p} 1\left(\hat{y}_{t+j|t}^{(0.025)}(x_i) < y_{t+j}(x_i) < \hat{y}_{t+j|t}^{(0.975)}(x_i)\right)$$

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- Let $[\ell_h(x), u_h(x)]$ be $100(1-\alpha)\%$ *h*-step-ahead prediction intervals obtained from bootstrap methods.
- Ideally, $u_1(x) \ell_1(x) = d(x)$.

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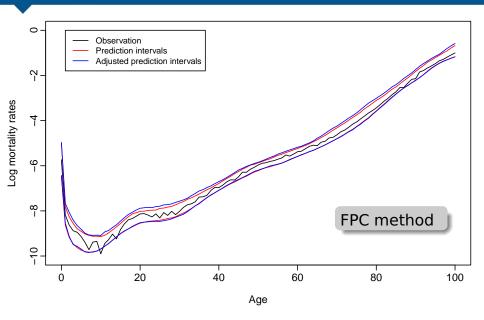
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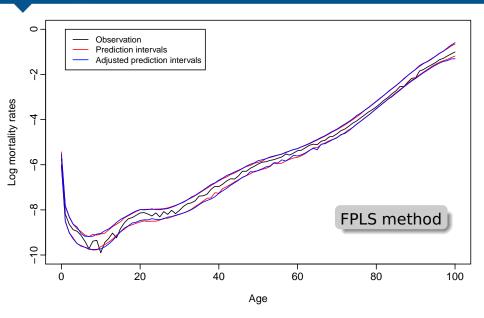
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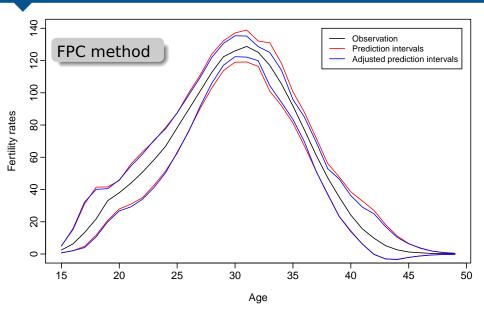
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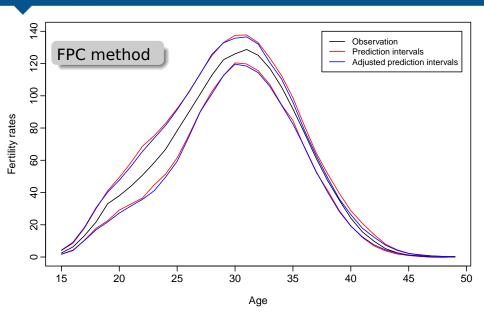
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- 2 Application: French mortality rates
- 3 Application: Australian fertility rates
- 4 Forecast accuracy comparisons
- **5** Bootstrap intervals
- **6** Comparisons
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Selected references



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