

$$\begin{aligned}
P_{X|A}(-2) &= \frac{13}{8}P(X = -2, A) \\
&= \frac{13}{8}P_{XY}(-2, 0) = \frac{1}{8}, \\
P_{X|A}(-1) &= \frac{13}{8}P(X = -1, A) \\
&= \frac{13}{8}[P_{XY}(-1, 0) + P_{XY}(-1, 1)] = \frac{2}{8} = \frac{1}{4}, \\
P_{X|A}(0) &= \frac{13}{8}P(X = 0, A) \\
&= \frac{13}{8}[P_{XY}(0, 0) + P_{XY}(0, 1)] = \frac{2}{8} = \frac{1}{4}, \\
P_{X|A}(1) &= \frac{13}{8}P(X = 1, A) \\
&= \frac{13}{8}[P_{XY}(1, 0) + P_{XY}(1, 1)] = \frac{2}{8} = \frac{1}{4}, \\
P_{X|A}(2) &= \frac{13}{8}P(X = 2, A) \\
&= \frac{13}{8}P_{XY}(2, 0) = \frac{1}{8}.
\end{aligned}$$

Thus, we have

$$\begin{aligned}
E[X|A] &= \sum_{x_i \in R_X} x_i P_{X|A}(x_i) \\
&= (-2)\frac{1}{8} + (-1)\frac{1}{4} + (0)\frac{1}{4} + (1)\frac{1}{4} + (2)\frac{1}{8} = 0.
\end{aligned}$$

c. To find $E[|X| \mid -1 < Y < 2]$, we use the conditional PMF and LOTUS. We have

$$\begin{aligned}
E[|X| \mid A] &= \sum_{x_i \in R_X} |x_i| P_{X|A}(x_i) \\
&= |-2| \cdot \frac{1}{8} + |-1| \cdot \frac{1}{4} + 0 \cdot \frac{1}{4} + 1 \cdot \frac{1}{4} + 2 \cdot \frac{1}{8} = 1.
\end{aligned}$$

Conditional expectation has some interesting properties that are used commonly in practice. Thus, we will revisit conditional expectation in [Section 5.1.5](#), where we discuss properties of conditional expectation, conditional variance, and their applications.

Law of Total Probability:

Remember the law of total probability: If B_1, B_2, B_3, \dots is a partition of the sample space S , then for any event A we have

$$P(A) = \sum_i P(A \cap B_i) = \sum_i P(A|B_i)P(B_i).$$

If Y is a discrete random variable with range $R_Y = \{y_1, y_2, \dots\}$, then the events $\{Y = y_1\}, \{Y = y_2\}, \{Y = y_3\}, \dots$ form a partition of the sample space. Thus, we can use the law of total probability. In fact we have already used the law of total probability to find the marginal PMFs:

$$P_X(x) = \sum_{y_j \in R_Y} P_{XY}(x, y_j) = \sum_{y_j \in R_Y} P_{X|Y}(x|y_j)P_Y(y_j).$$

We can write this more generally as

$$P(X \in A) = \sum_{y_j \in R_Y} P(X \in A|Y = y_j)P_Y(y_j), \quad \text{for any set } A.$$

We can write a similar formula for expectation as well. Indeed, if B_1, B_2, B_3, \dots is a partition of the sample space S , then

$$EX = \sum_i E[X|B_i]P(B_i).$$

To see this, just write the definition of $E[X|B_i]$ and apply the law of total probability. The above equation is sometimes called the law of total expectation [\[2\]](#).

Law of Total Probability:

$$P(X \in A) = \sum_{y_j \in R_Y} P(X \in A|Y = y_j)P_Y(y_j), \quad \text{for any set } A.$$

Law of Total Expectation:

1. If B_1, B_2, B_3, \dots is a partition of the sample space S ,

$$EX = \sum_i E[X|B_i]P(B_i) \quad (5.3)$$

2. For a random variable X and a discrete random variable Y ,

$$EX = \sum_{y_j \in R_Y} E[X|Y = y_j]P_Y(y_j) \quad (5.4)$$

Example 5.6

Let $X \sim \text{Geometric}(p)$. Find EX by conditioning on the result of the first "coin toss."

Solution

Remember that the random experiment behind $\text{Geometric}(p)$ is that we have a coin with $P(H) = p$. We toss the coin repeatedly until we observe the first heads. X is the total number of coin tosses. Now, there are two possible outcomes for the first coin toss: H or T . Thus, we can use the law of total expectation (Equation 5.3):

$$\begin{aligned} EX &= E[X|H]P(H) + E[X|T]P(T) \\ &= pE[X|H] + (1-p)E[X|T] \\ &= p \cdot 1 + (1-p)(EX + 1). \end{aligned}$$

In this equation, $E[X|T] = 1 + EX$, because the tosses are independent, so if the first toss is tails, it is like starting over on the second toss. Solving for EX , we obtain

$$EX = \frac{1}{p}.$$

Example 5.7

Suppose that the number of customers visiting a fast food restaurant in a given day is $N \sim \text{Poisson}(\lambda)$. Assume that each customer purchases a drink with probability p , independently from other customers and independently from the value of N . Let X be the number of customers who purchase drinks. Find EX .

Solution

By the above information, we conclude that given $N = n$, then X is a sum of n independent $\text{Bernoulli}(p)$ random variables. Thus, given $N = n$, X has a binomial distribution with parameters n and p . We write

$$X|N = n \sim \text{Binomial}(n, p).$$

That is,

$$P_{X|N}(k|n) = \binom{n}{k} p^k (1-p)^{n-k}.$$

Thus, we conclude

$$E[X|N = n] = np.$$

Thus, using the law of total probability, we have

$$\begin{aligned} E[X] &= \sum_{n=0}^{\infty} E[X|N = n]P_N(n) \\ &= \sum_{n=0}^{\infty} npP_N(n) \\ &= p \sum_{n=0}^{\infty} nP_N(n) = pE[N] = p\lambda. \end{aligned}$$