```
nlogL = normlike(params, data)
```

returns the negative of the normal log-likelihood function.

R = normrnd(mu, sigma)

R = randn

generates random numbers from the normal distribution with mean parameter mu and standard deviation parameter sigma.

- Exponential Distribution:

```
Y = exppdf(X,mu)
```

returns the pdf of the exponential distribution with mean parameter mu, evaluated at the values in X.

```
P = expcdf(X, mu)
```

computes the exponential cdf at each of the values in X using the corresponding mean parameter mu.

R = exprnd(mu)

generates random numbers from the exponential distribution with mean parameter mu.

12.5 Exercises

1. Write MATLAB programs to generate Geometric(p) and Negative Binomial(i,p) random variables.

Solution: To generate a Geometric random variable, we run a loop of Bernoulli trials until the first success occurs. K counts the number of failures plus one success, which is equal to the total number of trials.

```
K=1;
p=0.2;
while(rand>p)
K=K+1;
end
K
```

Now, we can generate Geometric random variable i times to obtain a Negative Binomial(i, p) variable as a sum of i independent Geometric (p) random variables.

```
K = 1;
p = 0.2;
r = 2;
success = 0;
while(success < r)
if \quad rand > p
K = K + 1;
print = 0 \quad \%Failure
else \quad success = success + 1;
print = 1 \quad \%Success
end
end
K + r - 1 \quad \%Number \quad of \quad trials \quad needed \quad to \quad obtain \quad r \quad successes
```

2. (Poisson) Use the algorithm for generating discrete random variables to obtain a Poisson random variable with parameter $\lambda = 2$.

Solution: We know a Poisson random variable takes all nonnegative integer values with probabilities

$$p_i = P(X = x_i) = e^{-\lambda} \frac{\lambda^i}{i!}$$
 for $i = 0, 1, 2, \dots$

To generate a $Poisson(\lambda)$, first we generate a random number U. Next, we divide the interval [0,1] into subintervals such that the jth subinterval has length p_j (Figure 12.2). Assume

$$X = \begin{cases} x_0 & \text{if } (U < p_0) \\ x_1 & \text{if } (p_0 \le U < p_0 + p_1) \\ \vdots \\ x_j & \text{if } \left(\sum_{k=0}^{j-1} p_k \le U < \sum_{k=0}^{j} p_k\right) \\ \vdots \end{cases}$$

Here $x_i = i - 1$, so

$$X = i$$
 if $p_0 + \dots + p_{i-1} \le U < p_0 + \dots + p_{i-1} + p_i$
 $F(i-1) \le U < F(i)$ F is CDF

$$lambda = 2;$$
 $i = 0;$
 $U = rand;$
 $cdf = exp(-lambda);$
 $while(U >= cdf)$
 $i = i + 1;$
 $cdf = cdf + exp(-lambda) * lambda^i/gamma(i + 1);$
 $end;$
 $X = i;$

3. Explain how to generate a random variable with the density

$$f(x) = 2.5x\sqrt{x}$$
 for $0 < x < 1$

if your random number generator produces a Standard Uniform random variable U. Hint: use the inverse transformation method.

Solution:

$$F_X(X) = X^{\frac{5}{2}} = U \quad (0 < x < 1)$$

 $X = U^{\frac{2}{5}}$

$$U = rand;$$
$$X = U^{\frac{2}{5}};$$

We have the desired distribution.

4. Use the inverse transformation method to generate a random variable having distribution function

$$F(x) = \frac{x^2 + x}{2}, \quad 0 \le x \le 1$$

Solution:

$$\frac{X^2 + X}{2} = U$$

$$(X + \frac{1}{2})^2 - \frac{1}{4} = 2U$$

$$X + \frac{1}{2} = \sqrt{2U + \frac{1}{4}}$$

$$X = \sqrt{2U + \frac{1}{4}} - \frac{1}{2} \quad (X, U \in [0, 1])$$

By generating a random number, U, we have the desired distribution.

$$U = rand;$$

$$X = sqrt\left(2U + \frac{1}{4}\right) - \frac{1}{2};$$

5. Let X have a standard Cauchy distribution.

$$F_X(x) = \frac{1}{\pi}\arctan(x) + \frac{1}{2}$$

Assuming you have $U \sim Uniform(0,1)$, explain how to generate X. Then, use this result to produce 1000 samples of X and compute the sample mean. Repeat the experiment 100 times. What do you observe and why?

Solution: Using Inverse Transformation Method:

$$U - \frac{1}{2} = \frac{1}{\pi} \arctan(X)$$
$$\pi \left(U - \frac{1}{2}\right) = \arctan(X)$$
$$X = \tan\left(\pi(U - \frac{1}{2})\right)$$

Next, here is the MATLAB code:

$$U = zeros(1000, 1);$$

 $n = 100;$
 $average = zeros(n, 1);$
 $for \quad i = 1:n$
 $U = rand(1000, 1);$
 $X = tan(pi*(U - 0.5));$
 $average(i) = mean(X);$
 end
 $plot(average)$

Cauchy distribution has no mean (Figure 12.6), or higher moments defined.

6. (The Rejection Method) When we use the Inverse Transformation Method, we need a simple form of the cdf F(x) that allows direct computation of $X = F^{-1}(U)$. When F(x) doesn't have a simple form but the pdf f(x) is available, random variables with

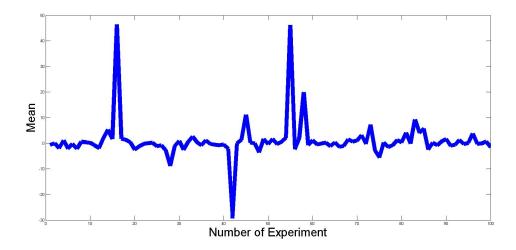


Figure 12.6: Cauchy Simulation

density f(x) can be generated by the **rejection method**. Suppose you have a method for generating a random variable having density function g(x). Now, assume you want to generate a random variable having density function f(x). Let c be a constant such that

$$\frac{f(y)}{g(y)} \le c$$
 (for all y)

Show that the following method generates a random variable with density function f(x).

- Generate Y having density g.
- Generate a random number U from Uniform (0,1).
- If $U \leq \frac{f(Y)}{cg(Y)}$, set X = Y. Otherwise, return to step 1.

Solution: The number of times N that the first two steps of the algorithm need to be called is itself a random variable and has a geometric distribution with "success" probability

$$p = P\left(U \le \frac{f(Y)}{cg(Y)}\right)$$

Thus, $E(N) = \frac{1}{p}$. Also, we can compute p:

$$\begin{split} P\left(U \leq \frac{f(Y)}{cg(Y)}|Y = y\right) &= \frac{f(y)}{cg(y)} \\ p &= \int_{-\infty}^{\infty} \frac{f(y)}{cg(y)} g(y) dy \\ &= \frac{1}{c} \int_{-\infty}^{\infty} f(y) dy \\ &= \frac{1}{c} \end{split}$$

Therefore, E(N) = c

Let F be the desired CDF (CDF of X). Now, we must show that the conditional distribution of Y given that $U \leq \frac{f(Y)}{cg(Y)}$ is indeed F, i.e. $P(Y \leq y | U \leq \frac{f(Y)}{cg(Y)}) = F(y)$. Assume $M = \{U \leq \frac{f(Y)}{cq(Y)}\}, K = \{Y \leq y\}$. We know $P(M) = p = \frac{1}{c}$. Also, we can compute

$$\begin{split} P(U \leq \frac{f(Y)}{cg(Y)} | Y \leq y) &= \frac{P(U \leq \frac{f(Y)}{cg(Y)}, Y \leq y)}{G(y)} \\ &= \int_{-\infty}^{y} \frac{P(U \leq \frac{f(y)}{cg(y)} | Y = v \leq y)}{G(y)} g(v) dv \\ &= \frac{1}{G(y)} \int_{-\infty}^{y} \frac{f(v)}{cg(v)} g(v) dv \\ &= \frac{1}{cG(y)} \int_{-\infty}^{y} f(v) dv \\ &= \frac{F(y)}{cG(y)} \end{split}$$

Thus,

$$\begin{split} P(K|M) &= P(M|K)P(K)/P(M) \\ &= P(U \le \frac{f(Y)}{cg(Y)}|Y \le y) \times \frac{G(y)}{\frac{1}{c}} \\ &= \frac{F(y)}{cG(y)} \times \frac{G(y)}{\frac{1}{c}} \\ &= F(y) \end{split}$$

7. Use the rejection method to generate a random variable having density function Beta(2, 4). Hint: Assume g(x) = 1 for 0 < x < 1. Solution:

$$f(x) = 20x(1-x)^3 \quad 0 < x < 1$$

$$g(x) = 1 \quad 0 < x < 1$$

$$\frac{f(x)}{g(x)} = 20x(1-x)^3$$

We need to find the smallest constant c such that $f(x)/g(x) \leq c$. Differentiation of this quantity yields

$$\frac{d\left(\frac{f(x)}{g(x)}\right)}{dx} = 0$$
Thus, $x = \frac{1}{4}$
Therefore, $\frac{f(x)}{g(x)} \le \frac{135}{64}$
Hence, $\frac{f(x)}{cg(x)} = \frac{256}{27}x(1-x)^3$

$$n = 1;$$
 $while(n == 1)$
 $U1 = rand;$
 $U2 = rand;$
 $if U2 <= 256/27 * U1 * (1 - U1)^3$
 $X = U1;$
 $n = 0;$
 end
 end

8. Use the rejection method to generate a random variable having the $Gamma(\frac{5}{2},1)$ density function. Hint: Assume g(x) is the pdf of the $Gamma\left(\alpha=\frac{5}{2},\lambda=1\right)$. Solution:

$$f(x) = \frac{4}{3\sqrt{\pi}} x^{\frac{3}{2}} e^{-x}, x > 0$$

$$g(x) = \frac{2}{5} e^{-\frac{2x}{5}} \quad x > 0$$

$$\frac{f(x)}{g(x)} = \frac{10}{3\sqrt{\pi}} x^{\frac{3}{2}} e^{-\frac{3x}{5}}$$

$$\frac{d\left(\frac{f(x)}{g(x)}\right)}{dx} = 0$$
Hence, $x = \frac{5}{2}$

$$c = \frac{10}{3\sqrt{\pi}} \left(\frac{5}{2}\right)^{\frac{3}{2}} e^{\frac{-3}{2}}$$

$$\frac{f(x)}{cg(x)} = \frac{x^{\frac{3}{2}} e^{\frac{-3x}{5}}}{\left(\frac{5}{2}\right)^{\frac{3}{2}} e^{\frac{-3}{2}}}$$

We know how to generate an Exponential random variable.

- Generate a random number U_1 and set $Y = -\frac{5}{2} \log U_1$.
- Generate a random number U_2 .
- If $U_2 < \frac{Y^{\frac{3}{2}}e^{\frac{-3Y}{5}}}{\left(\frac{5}{2}\right)^{\frac{3}{2}}e^{\frac{-3}{2}}}$, set X = Y. Otherwise, execute the step 1.
- 9. Use the rejection method to generate a standard normal random variable. Hint: Assume g(x) is the pdf of the exponential distribution with $\lambda = 1$.

Solution:

$$f(x) = \frac{2}{\sqrt{2\pi}}e^{-\frac{x^2}{2}} \quad 0 < x < \infty$$

$$g(x) = e^{-x} \quad 0 < x < \infty \quad \text{(Exponential density function with mean 1)}$$
 Thus,
$$\frac{f(x)}{g(x)} = \sqrt{\frac{2}{\pi}}e^{x-\frac{x^2}{2}}$$
 Thus,
$$x = 1 \quad \text{maximizes} \quad \frac{f(x)}{g(x)}$$
 Thus,
$$c = \sqrt{\frac{2e}{\pi}}$$

$$\frac{f(x)}{cg(x)} = e^{-\frac{(x-1)^2}{2}}$$

- Generate Y, an exponential random variable with mean 1.
- Generate a random number U.
- If $U \leq e^{\frac{-(Y-1)^2}{2}}$ set X = Y. Otherwise, return to step 1.
- 10. Use the rejection method to generate a Gamma(2,1) random variable conditional on its value being greater than 5, that is

$$f(x) = \frac{xe^{-x}}{\int_5^\infty xe^{-x}dx}$$
$$= \frac{xe^{-x}}{6e^{-5}} \quad (x \ge 5)$$

Hint: Assume g(x) be the density function of exponential distribution.

Solution: Since Gamma(2,1) random variable has expected value 2, we use an exponential distribution with mean 2 that is conditioned to be greater than 5.

$$f(x) = \frac{xe^{(-x)}}{\int_5^\infty xe^{(-x)}dx}$$
$$= \frac{xe^{(-x)}}{6e^{(-5)}} \quad x \ge 5$$
$$g(x) = \frac{\frac{1}{2}e^{(-\frac{x}{2})}}{e^{\frac{-5}{2}}} \quad x \ge 5$$
$$\frac{f(x)}{g(x)} = \frac{x}{3}e^{-(\frac{x-5}{2})}$$

We obtain the maximum in x = 5 since $\frac{f(x)}{g(x)}$ is decreasing. Therefore,

$$c = \frac{f(5)}{g(5)} = \frac{5}{3}$$

- Generate a random number V.

- $Y = 5 2\log(V)$.
- Generate a random number U. If $U < \frac{Y}{5}e^{-(\frac{Y-5}{2})}$, set X = Y; otherwise return to step 1.

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