

Random Numbers

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CONTENTS

I	Uniform Random Numbers	1
II	Central Limit Theorem	2
III	From Uniform to Other	3
IV	Triangular Distribution	4
V	Maximul Likelihood	5
VI	Gaussian to Other	7

I. UNIFORM RANDOM NUMBERS

Let U be a uniform random variable between 0 and 1.

I.1 Generate 10^6 samples of U using a C program and save into a file called uni.dat .

Solution: Download the following files and execute the C program.

```
wget https://github.com/rudranshm/Random/
blob/main/Code/exrand.c
wget https://github.com/rudranshm/Random/
blob/main/Code/coeffs.h
```

Run the following commands:

```
cc exrand.c -lm
./a.out
```

I.2 Load the uni.dat file into python and plot the empirical CDF of U using the samples in uni.dat. The CDF is defined as

$$F_U(x) = \Pr(U \leq x) \quad (\text{I.1})$$

Solution: The following code plots Fig I.2.

```
https://github.com/rudranshm/Random/blob/
main/Code/cdf_plot.py
```

Run the following command in the terminal to run the code.

```
python uni_cdf.py
```

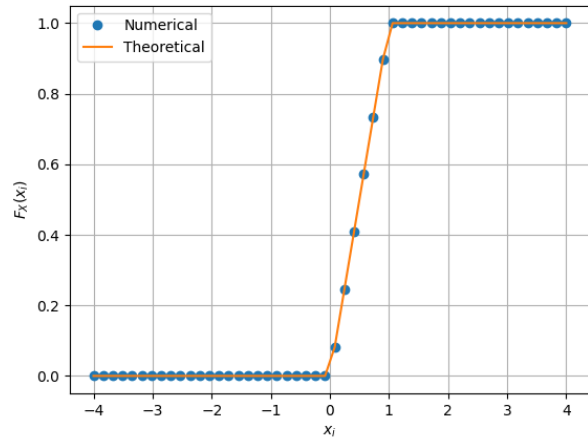


Fig. I.2: The CDF of U

I.3 Find a theoretical expression for $F_U(x)$.

Solution: Since U is an uniform random variable distribution, $P_U(x_i) = P_U(x_j) = k, \forall i, j$

$$F_U(x) = \int P_U(x) dx \quad (\text{I.2})$$

$$= \int k dx \quad (\text{I.3})$$

$$\text{we know that } \int_0^1 k dx = 1 \quad (\text{I.4})$$

$$\therefore k = 1 \quad (\text{I.5})$$

$$\therefore F_U(x) = x \quad (\text{I.6})$$

I.4 The mean of U is defined as

$$E[U] = \frac{1}{N} \sum_{i=1}^N U_i \quad (\text{I.7})$$

and its variance as

$$\text{var}[U] = E[U - E[U]]^2 \quad (\text{I.8})$$

Write a C program to find the mean and variance of U .

Solution:

```
wget https://github.com/rudranshm/Random/
blob/main/Code/meanvar1.4.c
```

Use below command to run file,

```
cc mean_var.c -lm
./a.out
```

running the code gives us the mean as 0.500031
an the variance as 0.083247

I.5 Verify your result theoretically given that

$$E[U^k] = \int_{-\infty}^{\infty} x^k dF_U(x) \quad (\text{I.9})$$

$$dF_U(x) = dx \quad (\text{I.10})$$

$$\therefore E[U^k] = \int_{-\infty}^{\infty} x^k dx \quad (\text{I.11})$$

$$E[U] = \int_0^1 x dx = \frac{1}{2} \quad (\text{I.12})$$

$$E[U^2] = \int_0^1 x^2 dx = \frac{1}{3} \quad (\text{I.13})$$

$$\therefore P_X(x) = 0, \forall x \in (1, \infty) \cap (-\infty, 0) \quad (\text{I.14})$$

$$\text{Var}(X) = E[U^2] - (E[U])^2 = \frac{1}{3} - \frac{1}{4} = \frac{1}{12} \quad (\text{I.15})$$

II. CENTRAL LIMIT THEOREM

II.1 Generate 10^6 samples of the random variable

$$X = \sum_{i=1}^{12} U_i - 6 \quad (\text{II.1})$$

using a C program, where $U_i, i = 1, 2, \dots, 12$
are a set of independent uniform random vari-
ables between 0 and 1 and save in a file called
gau.dat

Solution:

```
wget https://github.com/rudranshm/Random/
blob/main/Code/exrand.c
wget https://github.com/rudranshm/Random/
blob/main/Code/coeffs.h
```

Running the above codes generates uni.dat and
gau.dat file. Use the command

```
cc exrand.c -lm
./a.out
```

II.2 Load gau.dat in python and plot the empirical
CDF of X using the samples in gau.dat. What

properties does a CDF have?

Solution: The CDF of X is plotted in
II.2, Properties of the CDF:

- $\Phi(x) = P(Z \leq x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x \exp\left\{-\frac{u^2}{2}\right\} du$
- $\lim_{x \rightarrow \infty} \Phi(x) = 1, \lim_{x \rightarrow -\infty} \Phi(x) = 0$
- $\Phi(0) = \frac{1}{2}$
- $\Phi(-x) = 1 - \Phi(x)$

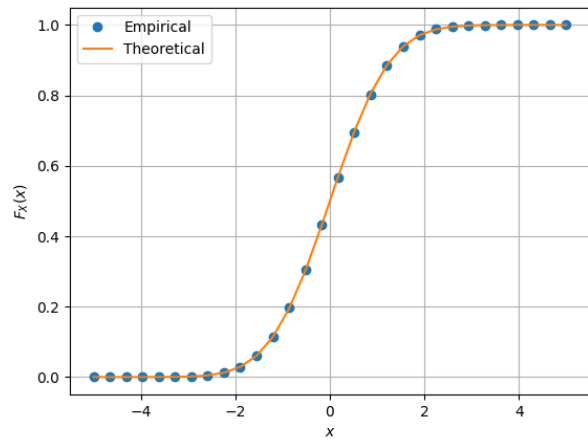


Fig. II.2: The CDF of X

II.3 Load gau.dat in python and plot the empirical
PDF of X using the samples in gau.dat. The
PDF of X is defined as

$$p_X(x) = \frac{d}{dx} F_X(x) \quad (\text{II.2})$$

What properties does the PDF have?

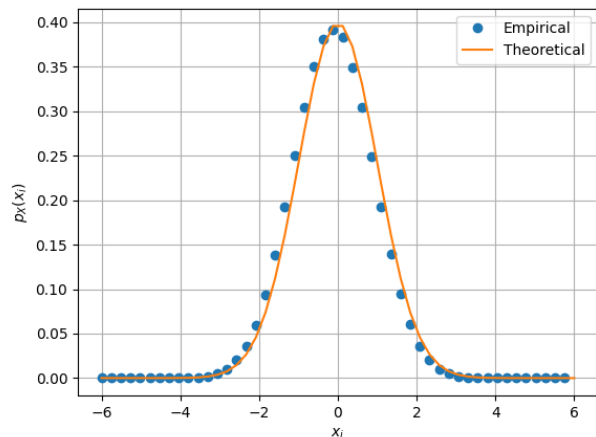


Fig. II.3: The PDF of X

Solution: The PDF of X is plotted in II.3 using

the code below

```
https://github.com/rudranshm/Random/blob/main/Code/pdf_plot.py
```

Use the below command to run the code:

```
python pdf_plot.py
```

Properties of PDF:

- PDF is symmetric about $x = 0$
- graph is bell shaped
- mean of graph is situated at the apex point of the bell

II.4 Find the mean and variance of X by writing a C program.

Solution: Running the below code gives the mean as 0.000685 and the variance as 1.000024

```
wget https://github.com/rudranshm/Random/blob/main/Code/meanvargaussian.c
```

Command used:

```
cc meanvargaussian.c -lm  
./a.out
```

II.5 Given that

$$p_X(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right), -\infty < x < \infty, \quad (\text{II.3})$$

repeat the above exercise theoretically.

Given, $p_X(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$

$$E[x] = \int_{-\infty}^{\infty} x p_X(x) dx \quad (\text{II.4})$$

$$= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} x e^{-\frac{x^2}{2}} dx \quad (\text{II.5})$$

$$\because x e^{-\frac{x^2}{2}} \text{ is a odd function,} \quad (\text{II.6})$$

$$E[x] = 0$$

$$E[x^2] = \int_{-\infty}^{\infty} x^2 p_X(x) dx \quad (\text{II.7})$$

$$= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} x^2 e^{-\frac{x^2}{2}} dx \quad (\text{II.8})$$

Using integration by parts:

$$= x \int x e^{-\frac{x^2}{2}} dx - \int \frac{d(x)}{dx} \int x e^{-\frac{x^2}{2}} dx \quad (\text{II.9})$$

$$I = \int x e^{-\frac{x^2}{2}} dx \quad (\text{II.10})$$

$$\text{Let } \frac{x^2}{2} = t \quad (\text{II.11})$$

$$\Rightarrow x dx = dt \quad (\text{II.12})$$

$$\Rightarrow \int e^{-t} dt = -e^{-t} + c \quad (\text{II.13})$$

$$\therefore \int x e^{-\frac{x^2}{2}} = -e^{-\frac{x^2}{2}} + c \quad (\text{II.14})$$

Using (II.14) in (II.9)

$$= -x e^{-\frac{x^2}{2}} + \int e^{-\frac{x^2}{2}} dx \quad (\text{II.15})$$

$$\text{Also, } \int_{-\infty}^{\infty} e^{-\frac{x^2}{2}} dx = \sqrt{2\pi} \quad (\text{II.16})$$

$$\therefore \text{substituting limits we get, } E[x^2] = 1 \quad (\text{II.17})$$

$$\text{Var}(X) = E[x^2] - (E[x])^2 = 1 - 0 \quad (\text{II.18})$$

III. FROM UNIFORM TO OTHER

III.1 Generate samples of

$$V = -2 \ln(1 - U) \quad (\text{III.1})$$

and plot its CDF.

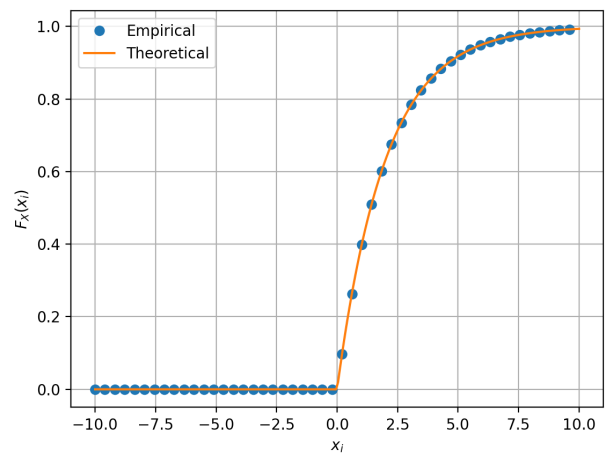


Fig. III.1: CDF for V

Solution: Running the below code generates samples of V from file uni.dat(U).

```
https://github.com/rudranshm/Random/blob/main/Code/var_v.py
```

Use the below command in the terminal to run the code:

```
python var_v.py
```

Now these samples are used to plot (III.1) by running the below code,

```
https://github.com/rudranshm/Random/blob/main/Code/cdfv.py
```

Use the below command to run the code:

```
python cdfv.py
```

$$\therefore P_T(t) = \int_0^1 P_{U2}(t-a) da$$

$$P_{U2}(x-a) = 1, \text{ for } 0 < x-a < 1 \therefore x-1 < a < x$$

If $x < 1, 0 < a < x$

$$P_T(t) = \int_0^x 1 da$$

$$\therefore P_T(t) = \begin{cases} 0 & t < 0 \\ t & 0 \leq t \leq 1 \\ 2-t & 1 < t \leq 2 \\ 0 & t > 2 \end{cases}$$

III.2 Find a theoretical expression for $F_V(x)$.

$$F_V(x) = P(V \leq x) \quad (\text{III.2})$$

$$= P(-2\ln(1-U) \leq x) \quad (\text{III.3})$$

$$= P(1 - e^{\frac{-x}{2}} \geq U) \quad (\text{III.4})$$

$$P(U < x) = \int_0^x dx = x \quad (\text{III.5})$$

$$\therefore P(1 - e^{\frac{-x}{2}} \geq U) = 1 - e^{\frac{-x}{2}}, \forall x \geq 0 \quad (\text{III.6})$$

$$\therefore F_T(t) = P(U_1 + U_2 < t) = \begin{cases} 0 & t < 0 \\ \frac{t^2}{2} & 0 \leq t \leq 1 \\ 1 - \frac{(2-t)^2}{2} & 1 < t \leq 2 \\ 1 & t > 2 \end{cases} \quad (\text{IV.2})$$

Find the PDF of T .

Solution:

IV. TRIANGULAR DISTRIBUTION

IV.1 Generate

$$T = U_1 + U_2 \quad (\text{IV.1})$$

Solution: Run the below code to generate T.dat

```
https://github.com/rudranshm/Random/blob/main/Code/Tvar.c
```

Run the command below in the terminal

```
cc Tvar.c -lm
./a.out
```

$$\therefore P_T(t) = \begin{cases} 0 & t < 0 \\ t & 0 \leq t \leq 1 \\ 2-t & 1 < t \leq 2 \\ 0 & t > 2 \end{cases} \quad (\text{IV.3})$$

Find the theoretical expressions for the PDF and CDF of T .

Solution:

$$P_T(t) = \begin{cases} 0 & t < 0 \\ t & 0 \leq t \leq 1 \\ 2-t & 1 < t \leq 2 \\ 0 & t > 2 \end{cases} \quad (\text{IV.4})$$

IV.2 Find the CDF of T .

$$P_T(t) = (U_1 + U_2 = x) = \int_{-\infty}^{\infty} P_{U1}(a)P_{U2}(t-a) da$$

$$F_T(t) = \begin{cases} 0 & t < 0 \\ \frac{t^2}{2} & 0 \leq t \leq 1 \\ 1 - \frac{(2-t)^2}{2} & 1 < t \leq 2 \\ 1 & t > 2 \end{cases} \quad (\text{IV.5})$$

As $P_{U1}(a) = 1$ for $0 < a < 1$ and 0 otherwise

Verify your results through a plot.

Solution: Run the below code to get the cdf

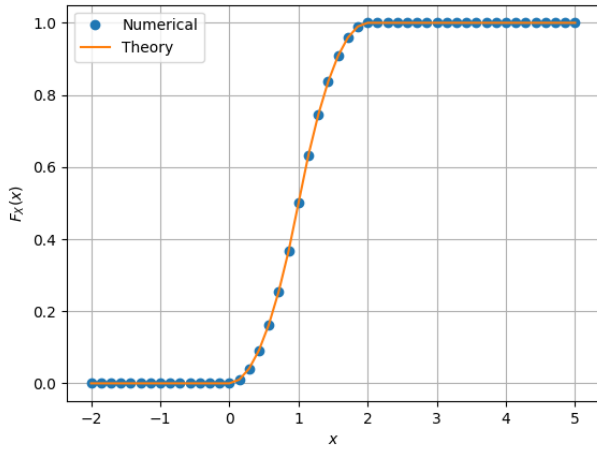


Fig. 2: CDF for (4)

<https://github.com/rudranshm/Random/blob/main/Code/cdf.py>

Use the following command in the terminal to run the code

```
python cdf.py
```

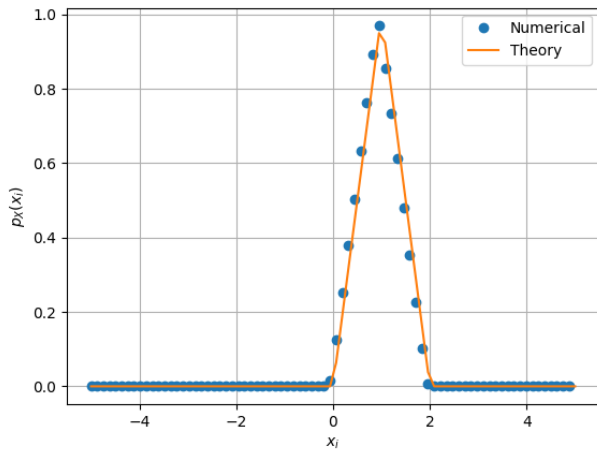


Fig. 2: PDF for (4)

Run the below code to get the pdf

<https://github.com/rudranshm/Random/blob/main/Code/pdf.py>

Use the following command in the terminal to run the code

```
python pdf.py
```

V. MAXIMUM LIKELIHOOD

V.1 Generate equiprobable $X \in \{1, -1\}$.

Solution: Run the below code,

<https://github.com/rudranshm/Random/blob/main/Code/bernoullivar.c>

Use the below command in the terminal to run the code

```
cc bernoullivar.c -lm
./a.out
```

V.2 Generate

$$Y = AX + N, \quad (V.1)$$

where $A = 5$ dB, and $N \sim \mathcal{N}(0, 1)$.

Solution: Run the below code for generating samples of Y ,

<https://github.com/rudranshm/Random/blob/main/Code/Yvar.c>

Use the below command in the terminal to run the code

```
cc Yvar.c -lm
./a.out
```

V.3 Plot Y using a scatter plot.

Solution:

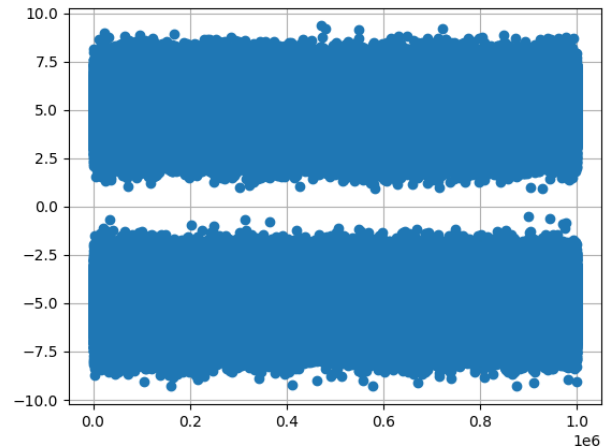


Fig. V.3: plot for (5.3)

Run the following code to generate the scatter plot

https://github.com/rudranshm/Random/blob/main/Code/y_plot.py

Use the below command to run the code,

```
python y_plot.py
```

V.4 Guess how to estimate X from Y .

Solution: if the received signal is greater than 0, we can assume that signal s_1 , corresponding to $X = 1$ was transmitted.

if the received signal is less than or equal to 0, we can assume that signal s_2 , corresponding to $X = -1$ was transmitted. Here threshold 0 is taken to be the decision boundary.

$$y > 0 \implies s_1 \quad (\text{V.2})$$

$$y \leq 0 \implies s_0 \quad (\text{V.3})$$

V.5 Find

$$P_{e|0} = \Pr(\hat{X} = -1 | X = 1) \quad (\text{V.4})$$

and

$$P_{e|1} = \Pr(\hat{X} = 1 | X = -1) \quad (\text{V.5})$$

Solution: Here s_1 and s_2 are equally probable ie, $p(s_1) = p(s_0) = \frac{1}{2}$

$$Q(x) = \frac{1}{\sqrt{2\pi}} \int_x^\infty e^{-\frac{x^2}{2}} dx \quad (\text{V.6})$$

$$\begin{aligned} p(e|s_1) &= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^0 e^{-\frac{(y-A)^2}{2}} dy \\ &= Q(A) \end{aligned} \quad (\text{V.7})$$

$$\begin{aligned} p(e|s_0) &= \frac{1}{\sqrt{2\pi}} \int_0^\infty e^{-\frac{(y+A)^2}{2}} dy \\ &= Q(A) \end{aligned} \quad (\text{V.8})$$

V.6 Find P_e assuming that X has equiprobable symbols.

Solution: Total probability of bit error:

$$P_e = p(s_1)p(e|s_1) + p(s_0)p(e|s_0) \quad (\text{V.9})$$

$$= \frac{1}{2}[Q(A) + Q(A)] \quad (\text{V.10})$$

$$\because p(s_1) = p(s_0) = \frac{1}{2}, X \text{ has equiprobable symbols} \quad (\text{V.11})$$

$$= Q(A)$$

$$= Q(5) \quad (\text{V.12})$$

Since $A=5$

V.7 Verify by plotting the theoretical P_e with respect to A from 0 to 10 dB.

Solution:

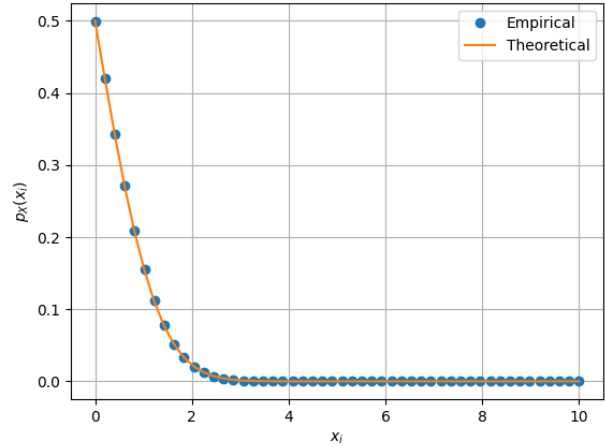


Fig. V.7: plot for (5.7)

Run the following code to generate the scatter plot

```
https://github.com/rudranshm/
Random/blob/main/Code/pe_plot
.py
```

Use the below command to run the code,

```
python pe_plot.py
```

V.8 Now, consider a threshold δ while estimating X from Y . Find the value of δ that minimizes the theoretical P_e .

Solution: Threshold= δ ,

$$y > \delta \implies s_1 \quad (\text{V.13})$$

$$y \leq \delta \implies s_0 \quad (\text{V.14})$$

$$p(e|s_1) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\delta} e^{-\frac{(y-A)^2}{2}} dy \quad (\text{V.15})$$

$$p(e|s_0) = \frac{1}{\sqrt{2\pi}} \int_{\delta}^{\infty} e^{-\frac{(y+A)^2}{2}} dy$$

$$P_e = \frac{1}{2\sqrt{2\pi}} \left(\int_{-\infty}^{\delta} e^{-\frac{(y-A)^2}{2}} dy + \int_{\delta}^{\infty} e^{-\frac{(y+A)^2}{2}} dy \right)$$

$$P_e = \frac{Q(\delta + A) + Q(A - \delta)}{2} \quad (\text{V.16})$$

$$P_e = f(\delta) \quad (\text{V.17})$$

$$\text{to minimize } P_e, \frac{d(f(\delta))}{d\delta} = 0 \text{ and } f''(\delta) > 0 \quad (\text{V.18})$$

$$e^{-\frac{(A-\delta)^2}{2}} - e^{-\frac{(A+\delta)^2}{2}} = 0 \quad (\text{V.19})$$

$$\therefore A - \delta = A + \delta, \implies \delta = 0 \quad (\text{V.20})$$

$$f''(\delta) = k((A - \delta)e^{-\frac{(A-\delta)^2}{2}} + (A + \delta)e^{-\frac{(A+\delta)^2}{2}}) > 0 \quad (\text{V.21})$$

V.9 Repeat the above exercise when

$$p_X(0) = p \quad (\text{V.22})$$

Solution: $p_X(0) = p$

$$\implies p_X(1) = 1 - p$$

$$P_e = pP(e|s_0) + (1 - p)P(e|s_1) \quad (\text{V.23})$$

$$= pQ(A + \delta) + (1 - p)Q(A - \delta) \quad (\text{V.24})$$

$$\frac{d(P_e)}{d(\delta)} = 0 \quad (\text{V.25})$$

$$\implies e^{\frac{(A+\delta)^2 - (A-\delta)^2}{2}} = \frac{p}{1-p} \quad (\text{V.26})$$

$$\therefore \delta = \frac{1}{2A} \log\left(\frac{p}{1-p}\right) \quad (\text{V.27})$$

$$\frac{d(P_e)}{d(\delta)} \text{ at } \delta + \epsilon > 0 \quad (\text{V.28})$$

$$\frac{d(P_e)}{d(\delta)} \text{ at } \delta - \epsilon < 0$$

$$\therefore \delta = \frac{1}{2A} \log\left(\frac{p}{1-p}\right) \rightarrow \text{minima}$$

$$A = 5 \implies \delta = \frac{1}{10} \log\left(\frac{p}{1-p}\right) \quad (\text{V.29})$$

V.10 Repeat the above exercise using the MAP

criterion.

Solution:

$$P_{X|Y}(x|y) \Big|_{X=1} = \frac{P(Y = y|X = 1)P(X = 1)}{P(Y = y)} \quad (\text{V.30})$$

$$P(Y = y) = P(Y = y|X = 1)P(X = 1) + P(Y = y|X = -1)P(X = -1) \quad (\text{V.31})$$

$$P(Y = y|X = 1)P(X = 1) = pP(Y = A + N) \quad (\text{V.32})$$

$$= p \left(\frac{1}{\sqrt{2\pi}} e^{-\frac{(y-A)^2}{2}} \right) \quad (\text{V.33})$$

$$\therefore P_{X|Y}(x|y) \Big|_{X=1} = \frac{p \left(\frac{1}{\sqrt{2\pi}} e^{-\frac{(y-A)^2}{2}} \right)}{P(Y = y)} \quad (\text{V.34})$$

$$P_{X|Y}(x|y) \Big|_{X=-1} = \frac{P(Y = y|X = -1)P(X = -1)}{P(Y = y)} \quad (\text{V.35})$$

$$P(Y = y|X = -1)P(X = -1) = (1 - p)P(Y = -A + N) = (1 - p) \left(\frac{1}{\sqrt{2\pi}} e^{-\frac{(y+A)^2}{2}} \right) \quad (\text{V.36})$$

$$\therefore P_{X|Y}(x|y) \Big|_{X=-1} = \frac{(1 - p) \left(\frac{1}{\sqrt{2\pi}} e^{-\frac{(y+A)^2}{2}} \right)}{P(Y = y)} \quad (\text{V.37})$$

Now comparing $a = P_{X|Y}(x|y) \Big|_{X=-1}$ and $b = P_{X|Y}(x|y) \Big|_{X=1}$, if $a > b$, $X = -1$ is more likely, $a < b$, $X = 1$ is more likely.

$$pe^{-\frac{(y-A)^2}{2}} \geq (1 - p)e^{-\frac{(y+A)^2}{2}}$$

$$\implies e^{2Ay} \geq \frac{1-p}{p}$$

$$\implies y \geq \frac{1}{2A} \log\left(\frac{1-p}{p}\right)$$

$$\delta = \frac{1}{2A} \log\left(\frac{1-p}{p}\right)$$

$$y > \delta \implies X=1 \text{ is more likely}$$

$$y < \delta \implies X=-1 \text{ is more likely}$$

VI. GAUSSIAN TO OTHER

VI.1 Let $X_1 \sim \mathcal{N}(0, 1)$ and $X_2 \sim \mathcal{N}(0, 1)$. Plot the CDF and PDF of

$$V = X_1^2 + X_2^2 \quad (\text{VI.1})$$

Solution:

Method1: The sum of squares of k independent standard random normal variables is nothing but a χ^2 distribution with k degrees of freedom.

$$\chi^2(k) = \frac{x^{\frac{k}{2}-1}}{2^{\frac{k}{2}}\Gamma(\frac{k}{2})} e^{-\frac{x}{2}}, \forall x \geq 0$$

Here $k=2$,

$$\therefore \chi^2(2) = P_V(v) = \frac{e^{-\frac{v}{2}}}{2} \quad (\text{VI.2})$$

$$\Rightarrow F_V(v) = \int_0^v \frac{e^{-\frac{x}{2}}}{2} dx \quad (\text{VI.3})$$

$$= 1 - e^{-\frac{v}{2}} \quad (\text{VI.4})$$

Method2:

$$X_1 = R \cos \theta \quad (\text{VI.5})$$

$$X_2 = R \sin \Theta \quad (\text{VI.6})$$

$R \in [0, \infty), \Theta \in [0, 2\pi)$. Jacobian Matrix is given by

$$J = \begin{pmatrix} \frac{\partial X_1}{\partial R} & \frac{\partial X_2}{\partial R} \\ \frac{\partial X_1}{\partial \Theta} & \frac{\partial X_2}{\partial \Theta} \end{pmatrix} \quad (\text{VI.7})$$

$$= \begin{pmatrix} \cos \Theta & \sin \Theta \\ -R \sin \Theta & R \cos \Theta \end{pmatrix} \quad (\text{VI.8})$$

$$\Rightarrow |J| = R \quad (\text{VI.9})$$

We also know that

$$|J|p_{X_1, X_2}(x_1, x_2) = p_{R, \Theta}(r, \theta) \quad (\text{VI.10})$$

$$\Rightarrow p_{R, \Theta}(r, \theta) = R p_{X_1}(x_1) p_{X_2}(x_2) \quad (\text{VI.11})$$

$$= \frac{R}{2\pi} e^{-\left(\frac{x_1^2 + x_2^2}{2}\right)} \quad (\text{VI.12})$$

$$= \frac{R}{2\pi} e^{-\left(\frac{R^2}{2}\right)} \quad (\text{VI.13})$$

X_1, X_2 are iid random variables. Thus,

$$p_R(r) = \int_0^{2\pi} p_{R, \Theta}(r, \theta) d\theta \quad (\text{VI.14})$$

$$= R e^{-\left(\frac{R^2}{2}\right)} \quad (\text{VI.15})$$

However, $V = X_1^2 + X_2^2 = R^2 \geq 0$, thus $F_V(x) = 0$ for $x \geq 0$.

$$F_V(x) = F_R(\sqrt{x}) \quad (\text{VI.16})$$

$$= \int_0^{\sqrt{x}} r e^{-\left(\frac{r^2}{2}\right)} dr \quad (\text{VI.17})$$

$$= \int_0^{\frac{x}{2}} e^{-t} dt = 1 - e^{-\frac{x}{2}} \quad (\text{VI.18})$$

To generate data for V , run the following code,

```
https://github.com/rudranshm/Random/blob/main/Code/var_v.py
```

Run the below command in terminal,

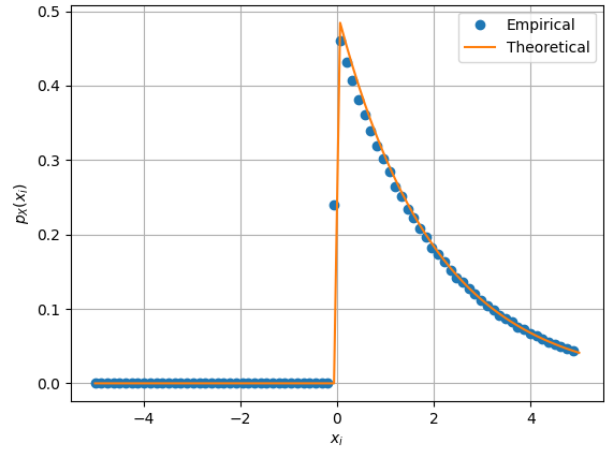


Fig. VI.1: PDF for (6.1)

```
cc var_v.c -lm
./a.out
```

The PDF plot of the $\chi^2(2)$ can be obtained by running the code below,

```
https://github.com/rudranshm/Random/blob/main/Code/chi_pdf.py
```

Use the following command in the terminal to run the code

```
python chi_pdf.py
```

The CDF plot of the $\chi^2(2)$ can be obtained by running the code below,

```
https://github.com/rudranshm/Random/blob/main/Code/chi_cdf.py
```

Use the following command in the terminal to run the code

```
python chi_cdf.py
```

VI.2 If

$$F_V(x) = \begin{cases} 1 - e^{-\alpha x} & x \geq 0 \\ 0 & x < 0, \end{cases} \quad (\text{VI.19})$$

find α .

Solution: From (VI.4) $\alpha = 0.5$

VI.3 Plot the CDF and Pdf of

$$A = \sqrt{V} \quad (\text{VI.20})$$

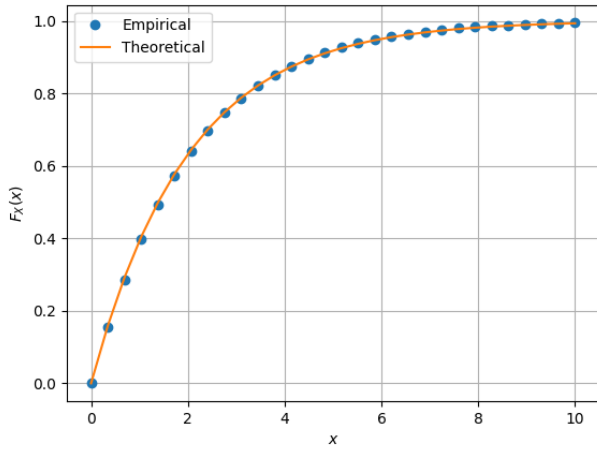


Fig. VI.1: CDF for (6.1)

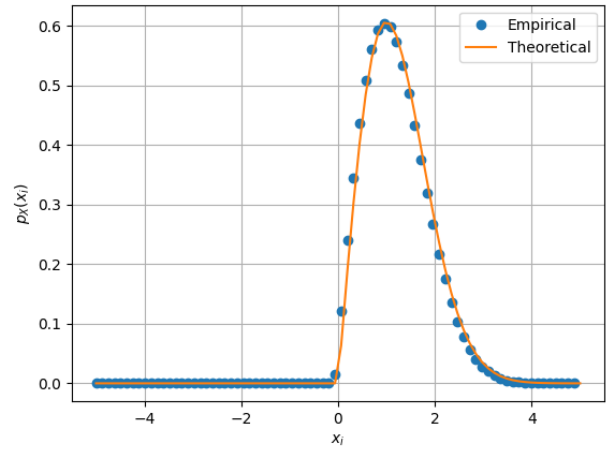


Fig. VI.3: PDF for (6.3)

Solution:

$$F_A(a) = P(A < a) = P(V < a^2) \quad (\text{VI.21})$$

$$\text{from (VI.4), } = \begin{cases} 1 - e^{-\frac{a^2}{2}} & a > 0 \\ 0 & a \leq 0 \end{cases} \quad (\text{VI.22})$$

$$\Rightarrow P_A(a) = \frac{d(F_A(a))}{da} \quad (\text{VI.23})$$

$$= \begin{cases} ae^{-\frac{a^2}{2}} & a > 0 \\ 0 & a \leq 0 \end{cases} \quad (\text{VI.24})$$

To generate data for A, run the following code,

```
https://github.com/rudranshm/Random/blob/main/Code/varA.c
```

Run the below command in terminal,

```
cc varA.c -lm
./a.out
```

The PDF plot of A can be obtained by running the code below,

```
https://github.com/rudranshm/Random/blob/main/Code/A_pdf.py
```

Use the following command in the terminal to run the code

```
python A_pdf.py
```

The CDF plot of the A can be obtained by running the code below,

```
https://github.com/rudranshm/Random/blob/main/Code/A_cdf.py
```

Use the following command in the terminal to run the code

```
python A_cdf.py
```

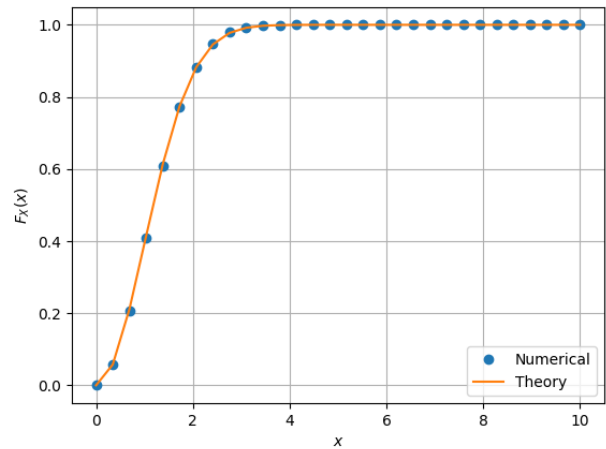


Fig. VI.3: CDF for (6.3)