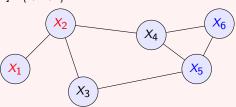
The recent development of dependency graphs

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Dependency Graphs

Graph G is a dependency graph for random variables $\mathbf{X} = (X_1, \dots, X_n)$ if

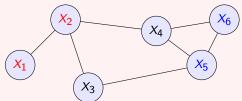
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- If disjoint subsets $I, J \subset [n]$ are non-adjacent in G, then $\{X_i\}_{i \in I}$ and $\{X_j\}_{j \in J}$ are independent.
 - ▶ In the above example, $\{X_1, X_2\}$ and $\{X_5, X_6\}$ are independent.
- There are ones with weaker assumptions, such as the one used in Lovász local lemma.
- The dependency graph for variables may not necessarily be unique, and the sparser ones are the more interesting ones.

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- Limiting distribution via dependency graph.
 - Janson's normality criterion via cumulants (1988)

A McDiarmid-type bounded difference inequality for graph-dependent variables.

Theorem (Z., Liu, Wang, Wang, 2019)

If we have

- f is c-Lipschitz,
- **2** F is a dependency graph for X and $F = \{T_i\}_{i \in [k]}$ is a forest,

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Theorem (Z., Liu, Wang, Wang, 2019)

If we have

- f is c-Lipschitz,
- **2** F is a dependency graph for X and $F = \{T_i\}_{i \in [k]}$ is a forest, then for every t > 0,

$$\mathbb{P}\left(f(\boldsymbol{X}) - \mathbb{E}\left[f(\boldsymbol{X})\right] \ge t\right) \le \exp\left(-\frac{2t^2}{\sum_{i=1}^k c_{\min,i}^2 + \sum_{\{i,j\} \in E(F)} (c_i + c_j)^2}\right),$$

where $c_{\min,i} = \min\{c_j : j \in V(T_i)\}$ is the minimum entry of \mathbf{c} in each tree T_i .

Janson, 2004

Let random variables $\{X_i\}_{i\in V(G)}$ be G-dependent such that every X_i takes values in an interval of length $c_i\geqslant 0$. Then, for every t>0,

$$\mathbb{P}\left(\sum_{i \in V(G)} X_i - \mathbb{E}\left[\sum_{i \in V(G)} X_i\right] \ge t\right) \le \exp\left(-\frac{2t^2}{\chi_f(G)\sum_{i \in V(G)} c_i^2}\right),$$

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Under the same setting,

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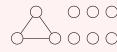
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- ullet It is no worse than Janson's, better when G is sparse.
- It generalises to certain decomposable Lipschitz functions, extending McDiarmid's.

Example

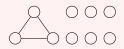
Let $\{X_i\}_{i\in[9]}$ be random indicators with the dependency graph G, and $X=\sum_{i\in[9]}X_i$.



$$\mathbb{P}(X - \mathbb{E}[X] \ge t) \le ?$$

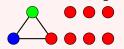
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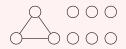
 Janson's idea: to cover vertices with weighted independent sets such that the sum of weights for each vertex equals 1



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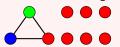
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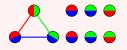
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 New idea: to cover vertices with weighted induced forests such that the sum of weights for each vertex equals 1



$$\mathbb{P}(X - \mathbb{E}[X] \ge t) \le \exp\left(-\frac{8t^2}{81}\right)$$

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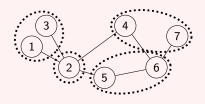
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Then we have a factorization property



$$\mu(1,3)\mu(2)\mu(4,7)\mu(5,6)$$

= $\mu(1)\mu(2)\mu(3)\mu(4)\mu(5,6)\mu(7)$.

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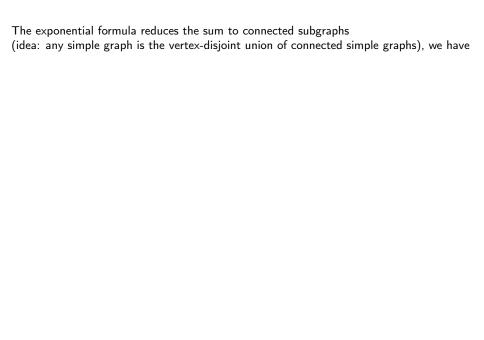
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where $C_1 \cup ... \cup C_k = S$ and $C_1,...,C_k$ induce maximal pairwise non-adjacent 'polymers' (that is, connected subgraphs; also called 'animals', suggested by Dobrushin, 1996).



(idea: any simple graph is the vertex-disjoint union of connected simple graphs), we have

Formal cluster expansion (Z. 2022)

$$\log (\mathbb{P}(X=0)) \stackrel{\text{formally}}{=} \sum_{k \geqslant 1} \sum_{(C_1, \dots, C_k)} \phi(C_1, \dots, C_k) \prod_{C \in (C_1, \dots, C_k)} (-1)^{|C|} \mu(C),$$

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• Polymers (connected subgraphs): $\{..., \triangleleft >, \triangleleft <, \triangleright >, ...\}$.

- where the sum is over ordered multisets of polymers, whose union is connected
- Clusters: ..., 炎 ,....

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 - ightharpoonup the tight instance for lower bound is when every polymer is \mathcal{K}_1 (also by Shearer, 1985).

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• Wormald, 1996: if $p = o(n^{-2/3})$, then

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$$-\widetilde{\mathbb{E}}\Big[\bigcirc \widehat{\mathbb{E}} \Big] - \widetilde{\mathbb{E}}\Big[\bigcirc \widehat{\mathbb{E}} \Big] + \widetilde{\mathbb{E}}\Big[\bigcirc \widehat{\mathbb{E}} \Big] + \dots$$

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$$\begin{split} \log \mathbb{P}(X = 0) & \stackrel{\text{formally}}{=} -\mathbb{E}\left[\circlearrowleft \right] + \mathbb{E}\left[\circlearrowleft \right] - \mathbb{E}\left[\circlearrowleft \right] - \mathbb{E}\left[\circlearrowleft \right] - \mathbb{E}\left[\circlearrowleft \right] \\ + \mathbb{E}\left[\circlearrowleft \right] - \mathbb{E}\left[\circlearrowleft \right] + \mathbb{E}\left[\circlearrowleft \right] + \mathbb{E}\left[\circlearrowleft \right] + \mathbb{E}\left[\circlearrowleft \right] + \mathbb{E}\left[\circlearrowleft \right] \\ - \mathbb{E}\left[\circlearrowleft \right] - \mathbb{E}\left[\circlearrowleft \right] + \mathbb{E}\left[\circlearrowleft \right] + \mathbb{E}\left[\circlearrowleft \right] + \dots \end{split}$$

Stark and Wormald, 2018; Mousset, Noever, Panagiotou, Samotij, 2020

If $p = o(n^{-7/11})$, then

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Joint cumulant (1929)

For any multiset of random variables $\{X_i\}_{i \in S}$,

$$\kappa(S) = \sum_{\pi \in \Pi(S)} (-1)^{|\pi|-1} (|\pi|-1)! \prod_{P \in \pi} \mu(P).$$

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if S does not induce a connected subgraph in G, then $\kappa(S) = 0$.

- Isaev, McKay, Z. (2022+): accurate asymptotic enumeration of Eulerian orientations, digraphs, oriented graphs.

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- Isaev, McKay, Z. (2022+):
- accurate asymptotic enumeration of Eulerian orientations, digraphs, oriented graphs.
- Stark and Wormald (2018) also have results for $\mathcal{G}_{n,m}$: if $d = m\binom{n}{2}^{-1} = o\left(n^{-7/11}\right)$, then

$$\mathbb{P}(X=0) = \exp\left(-\frac{1}{6}n^3d^3 - \frac{1}{9}n^4d^6 - \frac{1}{2}n^2d^3 + o(1)\right).$$

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If $p = o(n^{-3/2})$, then

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Z. 2022: Wormald and Z., 2022+

If $p = o(n^{-7/5})$, then

$$\mathbb{P}(H_3(n,p) \text{ is linear}) = \exp\left(-\frac{1}{4}n^4p^2 + \frac{2}{3}n^5p^3 - \frac{55}{24}n^6p^4 + \frac{3}{2}n^3p^2 + o(1)\right).$$

- In Z. 2022, the truncation error is by Mousset, Noever, Panagiotou, Samotij's results.
- Coefficients match McKay's conjecture based upon numerical simulation.

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Wormald and Z. 2022+

If $d = {n \choose 3}^{-1} m = o(n^{-7/5})$, then via perturbation method,

$$\mathbb{P}(H_3(n,m) \text{ is linear}) = \exp\left(-\frac{1}{4}n^4d^2 - \frac{1}{12}n^5d^3 - \frac{1}{24}n^6d^4 + \frac{3}{2}n^3d^2 + o(1)\right).$$

Dependency digraphs

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A digraph $\mathbf{D} = ([d], E)$ is a dependency digraph for the events $\{A_i\}_{i \in [d]}$ if A_i is mutually independent of all the non-adjacent $\{A_i : (i,j) \not\in E\}$ for every $i \in [d]$.

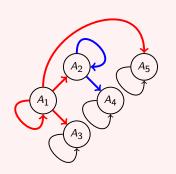
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Example

Let $\{A_i\}_{i\in[5]}$ be events with the following dependency digraph.

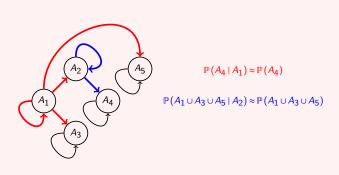


 A_1 is independent of A_4 .

 A_2 is independent of $\{A_1, A_3, A_5\}$.

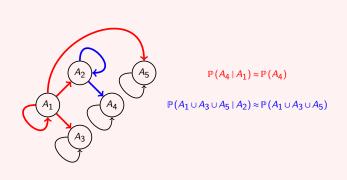
A natural refinement is

- adjacent vertices \rightarrow the pairs of 'strongly dependent' events, $D_i \subseteq [d]$ denotes the 'strongly dependent neighbours' of event A_i .
- ullet non-adjacent vertices ullet the pairs of 'weakly dependent' events.



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If we change ' \approx ' to ' \geqslant ', then it is the 'negative dependency graph' by Erdős and Spencer (1991) to obtain lopsided Lovász local lemma.

For any events $\{A_i\}_{i\in[d]}$ and digraph **D** on [d]:

•
$$\varphi$$
-mixing coefficient is used to measure 'weak dependencies':
$$\alpha = \max_{\alpha \in \mathbb{R}^n} \left| \mathbb{P} \left(-1, 1, \ldots, \frac{1}{2} \right) - \mathbb{P} \left(-1, 1, \ldots, \frac{1}{2} \right) \right|$$

$$\varphi = \max_{i \in [d]} \left| \mathbb{P}\left(\bigcup_{j \in [i-1] \setminus D_i} A_j \mid A_i\right) - \mathbb{P}\left(\bigcup_{j \in [i-1] \setminus D_i} A_j\right) \right|.$$

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• The influence of 'strongly dependent' events is
$$\Delta_1 = \sum_{i \in [I,I]} \mathbb{P} \left(A_i \cap \bigcup_{i \in [I,I]} A_j \right) \prod_{k \in [I,I]} \mathbb{P} \left(\overline{A_k} \right)$$

$$\Delta_{1} = \sum_{i \in [d]} \mathbb{P} \left(A_{i} \cap \bigcup_{j \in [i-1] \cap D_{i}} A_{j} \right) \prod_{k \in [d] \setminus [i]} \mathbb{P} \left(\overline{A_{k}} \right)$$

$$\Delta_{2} = \sum_{i \in [d]} \mathbb{P} \left(A_{i} \right) \mathbb{P} \left(\bigcup_{i \in [i-1] \cap D_{i}} A_{j} \right) \prod_{k \in [d] \setminus [i]} \mathbb{P} \left(\overline{A_{k}} \right).$$

For any events $\{A_i\}_{i\in[d]}$ and digraph **D** on [d]:

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• The influence of 'strongly dependent' events is
$$\Delta_1 \leqslant \Delta_1' = \sum_{i \in [d]} \sum_{j \in [i-1] \cap D_i} \mathbb{P}\left(A_i \cap A_j\right),$$

$$\Delta_2 \leqslant \Delta_2' = \sum_{i \in [d]} \sum_{j \in [i-1] \cap D_i} \mathbb{P}\left(A_i\right) \mathbb{P}\left(A_j\right).$$

For any events $\{A_i\}_{i\in[d]}$ and digraph **D** on [d]:

 \bullet φ -mixing coefficient is used to measure 'weak dependencies'

$$\varphi = \max_{i \in [d]} \left| \mathbb{P}\left(\bigcup_{j \in [i-1] \setminus D_i} A_j \mid A_i\right) - \mathbb{P}\left(\bigcup_{j \in [i-1] \setminus D_i} A_j\right) \right|.$$

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$$\begin{split} &\Delta_1 \leq \Delta_1' = \sum_{i \in [d]} \sum_{j \in [i-1] \cap D_i} \mathbb{P}\left(A_i \cap A_j\right), \\ &\Delta_2 \leq \Delta_2' = \sum_{i \in [d]} \sum_{j \in [i-1] \cap D_i} \mathbb{P}\left(A_i\right) \mathbb{P}\left(A_j\right). \end{split}$$

- Chen–Stein Poisson approximation involves similar terms as φ , Δ'_1 , and Δ'_2 . Chen (1975) also considered the φ -mixing condition.
- Suen's inequality (1990) also involves Δ_1' and Δ_2' .

For any events $\{A_i\}_{i\in[d]}$ and digraph **D** with vertex set [d], we have

$$\left| \mathbb{P}\left(\bigcap_{i \in [d]} \overline{A_i}\right) - \prod_{i \in [d]} \mathbb{P}\left(\overline{A_i}\right) \right| \leq \left(1 - \prod_{i \in [d]} \mathbb{P}\left(\overline{A_i}\right)\right) \varphi + \max(\Delta_1, \Delta_2).$$

 The lower bound generalises Dubickas' bound (2008), and the upper bound is via a similar inductive proof.

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- If $\varphi = o(1)$, $\Delta_1 = o(1)$, and $\Delta_2 = o(1)$, then we have the asymptotic independence.

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 - 'Two Moments Suffice for Poisson Approximations: The Chen-Stein Method' by Arratia, Goldstein and Gordon (1989); also Chen (1975).

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- Max codegrees, max clique-extension counts of random (hyper)graphs, etc. are asymptotically Gumbel.
- Lu and Székely (2009) also introduced a notion of 'positive' dependency graph to obtain upper bounds for Poisson approximation for random matchings.

Open problem

 \bullet To establish the absolute convergence of cluster expansion.

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- To design a suitable model for $\mathcal{G}(n,m)$; e.g., Féray's weighted dependency graphs for normality of subgraph counts.

Thanks!