

exercise2_109301060

April 16, 2024

1 Exercises for Lecture 2

109301060

Q1.a

```
[ ]: import wooldridge as woo
import statsmodels.formula.api as smf
import numpy as np

ceosal2 = woo.dataWoo('ceosal2')
y = ceosal2['salary']
x = ceosal2['ceoten']
```

```
[ ]: # a. Average salary and tenure
avg_sal = np.mean(y)
avg_ten = np.mean(x)

print(round(avg_sal, 2))
print(round(avg_ten, 2))
```

865.86

7.95

Q1.b

```
[ ]: # b. How many CEO's are in their first year?
ten_year1 = np.sum(x==0)

print(ten_year1)

# Longest tenure
ten_max = np.max(x)

print(ten_max)
```

5

37

Q1.c

```
[ ]: # c. Estimate the regression
reg = smf.ols('np.log(salary) ~ ceoten', data = ceosal2)
results = reg.fit()
b = results.params

# Print parameter estimates
print(f'b: \n{b}\n')

# Print results using summary:
print(f'results.summary(): \n{results.summary()}\n')

# Print regression table
import pandas as pd

table = pd.DataFrame({'b': round(b, 4),
                      'se': round(results.bse, 4),
                      't': round(results.tvalues, 4),
                      'pval': round(results.pvalues, 4)})
print(f'table: \n{table}\n')
```

```
b:
Intercept      6.505498
ceoten          0.009724
dtype: float64
```

```
results.summary():
```

```

                                OLS Regression Results
=====
Dep. Variable:          np.log(salary)    R-squared:                0.013
Model:                  OLS               Adj. R-squared:         0.008
Method:                 Least Squares     F-statistic:              2.334
Date:                  Tue, 16 Apr 2024   Prob (F-statistic):       0.128
Time:                  09:10:56          Log-Likelihood:           -160.84
No. Observations:      177              AIC:                    325.7
Df Residuals:          175              BIC:                    332.0
Df Model:               1
Covariance Type:       nonrobust
=====

```

	coef	std err	t	P> t	[0.025	0.975]
Intercept	6.5055	0.068	95.682	0.000	6.371	6.640
ceoten	0.0097	0.006	1.528	0.128	-0.003	0.022

```
=====
Omnibus:                 3.858    Durbin-Watson:           2.084
Prob(Omnibus):           0.145    Jarque-Bera (JB):        3.907
Skew:                   -0.189    Prob(JB):                0.142
Kurtosis:                3.622    Cond. No.:               16.1
=====
```

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

table:

	b	se	t	pval
Intercept	6.5055	0.0680	95.6817	0.0000
ceoten	0.0097	0.0064	1.5278	0.1284

0.97% increase in salary given one more year as a CEO.

Q2.a

```
[ ]: import scipy.stats as stats
import numpy as np
import statsmodels.formula.api as smf
import pandas as pd

# a. Generate 500 uniform random variable x
# Lower = loc, Upper = loc + scale
x = stats.uniform.rvs(loc = 0, scale = 10, size = 500)

# Sample mean and standard deviation of x
mean_x = x.mean()
std_x = x.std()
print(mean_x)
print(std_x)
```

4.963334976122681
2.7863943061386656

Q2.b

```
[ ]: # b. Generate 500 N(0, 36) variable u
u = stats.norm.rvs(loc = 0, scale = np.sqrt(36), size = 500)
mean_u = u.mean()
std_u = u.std()
print(mean_u)
print(std_u)
```

0.02234090133816153
5.822551531428157

Although the u sample mean and standard deviation have some deviations from the theoretical values, they are within a reasonable range. Therefore, it can be considered that the generated u data satisfies the assumption of obeying the $N(0, 36)$.

Q2.c

```
[ ]: # c. Generate  $y = 1 + 2x + u$ 
y = 1 + 2*x + u

data1 = pd.DataFrame({'y':y, 'x': x})

# Run the regression
reg = smf.ols(formula = 'y~x', data = data1)
results = reg.fit()

# Parameter estimates
b = results.params
print(f'b:\n{b}\n')
```

```
b:
Intercept    0.264935
x            2.152600
dtype: float64
```

β_1 is close to the theoretical value.

β_0 is far different from the theoretical value, but the intercept β_1 is more easily affected by the randomness of u .

Overall, although the estimated value of β_0 deviates from the true value, the estimated value of β_1 is very accurate, and the overall regression model captures the linear relationship between y and x , so the regression result is satisfactory.

Q2.d

```
[ ]: # d. OLS residuals
# Two ways to get the residuals
u_hat1 = results.resid

# Verify properties of residuals (use u_hat1)
sum_u_hat1 = u_hat1.sum()
print(sum_u_hat1)

x_u_hat1 = (x * u_hat1).sum()
print(x_u_hat1)
```

```
2.3661073100811336e-12
1.3812950783176348e-11
```

$\sum_{i=1}^n \hat{u}_i$ and $\sum_{i=1}^n x_i \hat{u}_i$ are all close to 0, and the characteristics are established.

Q2.e

```
[ ]: # e. Use u
sum_u = u.sum()
print(sum_u)
```

```
x_u = (x*u).sum()
print(x_u)
```

```
11.170450669080765
647.836058760705
```

Obviously $\sum_{i=1}^n u_i$ and $\sum_{i=1}^n x_i u_i$ are not equal to 0, and the characteristics do not hold.

This is reasonable because the least squares residual \hat{u}_i is specially constructed to satisfy these properties.

But the real error term u_i is randomly generated.

Q2.f

```
[ ]: # a. Generate 500 uniform random variable x
# Lower = loc, Upper = loc + scale
x1 = stats.uniform.rvs(loc = 0, scale = 10, size = 500)

# Sample mean and standard deviation of x
mean_x = x1.mean()
std_x = x1.std()
print('mean_x:', mean_x)
print('std_x:', std_x)

# b. Generate 500 N(0, 36) variable u
u1 = stats.norm.rvs(loc = 0, scale = np.sqrt(36), size = 500)
mean_u = u1.mean()
std_u = u1.std()
print('mean_u:', mean_u)
print('std_u:', std_u)

print('-----')
# c. Generate y = 1 + 2x + u
y1 = 1 + 2*x1 + u1

data2 = pd.DataFrame({'y': y1, 'x': x1})

# Run the regression
reg = smf.ols(formula = 'y~x', data = data2)
results = reg.fit()

# Parameter estimates
b = results.params
print(f'b:\n{b}\n')
print('-----')

# d. OLS residuals
```

```

# Two ways to get the residuals
u_hat1 = results.resid

# Verify properties of residuals (use u_hat1)
sum_u_hat1 = u_hat1.sum()
print(sum_u_hat1)

x_u_hat1 = (x1 * u_hat1).sum()
print(x_u_hat1)

# e. Use u
sum_u = u1.sum()
print(sum_u)

x_u = (x1*u1).sum()
print(x_u)

```

```

mean_x: 4.920882160167661
std_x: 2.792962130931689
mean_u: -0.38946532625981
std_u: 5.976240329242176

```

```

-----
b:
Intercept    0.538665
x            2.014605
dtype: float64

```

```

-----
-2.1316282072803006e-13
-3.609557097661309e-12
-194.732663129905
-901.2920781265744

```

The conclusion of Q2.E still holds.

Q3.a

```

[ ]: import numpy as np
import pandas as pd
import matplotlib.pyplot as plt
import statsmodels.formula.api as smf

# a. Import data
data = pd.read_csv('data_exercise_2c.csv')

```

Q3.b

```

[ ]: # b. Transfer DATE to date
data['DATE'] = pd.to_datetime(data['DATE'])

```

Q3.c

```
[ ]: # c. Calculate simple return
SP500_sr = np.diff(data['SP500'])/data['SP500'][:-1]
AAPL_sr = np.diff(data['AAPL_Close'])/data['AAPL_Close'][:-1]
AAPL_Adj_sr = np.diff(data['AAPL_Adj_Close'])/data['AAPL_Adj_Close'][:-1]

data['SP500_sr'] = np.append(np.nan, SP500_sr)
data['AAPL_sr'] = np.append(np.nan, AAPL_sr)
data['AAPL_Adj_sr'] = np.append(np.nan, AAPL_Adj_sr)
```

Q3.d

```
[ ]: # d. Calculate risk premia
data['MKT_rp'] = data['SP500_sr'] - data['RF_%']/100
data['AAPL_rp'] = data['AAPL_sr'] - data['RF_%']/100
data['AAPL_Adj_rp'] = data['AAPL_Adj_sr'] - data['RF_%']/100

# Calculate summary statistics of these risk premia
results_sst = data[['MKT_rp', 'AAPL_rp', 'AAPL_Adj_rp']].describe()
print(results_sst)
```

	MKT_rp	AAPL_rp	AAPL_Adj_rp
count	116.000000	116.000000	116.000000
mean	0.008940	0.023283	0.024555
std	0.033172	0.080546	0.080221
min	-0.191881	-0.185845	-0.185845
25%	-0.003751	-0.026152	-0.026152
50%	0.013634	0.023168	0.026221
75%	0.027767	0.078010	0.078010
max	0.063280	0.214280	0.214280

Q3.e

```
[ ]: # e. Fit CAPM
data = data.drop(data.index[0]) # remove the row containing NaN

# AAPL_rf
capm = smf.ols('AAPL_rp ~ MKT_rp', data=data)
result_capm = capm.fit()
print(f'result_capm.summary():\n{result_capm.summary()}\n')
```

result_capm.summary():

OLS Regression Results			
=====			
Dep. Variable:	AAPL_rp	R-squared:	0.210
Model:	OLS	Adj. R-squared:	0.203
Method:	Least Squares	F-statistic:	30.30
Date:	Tue, 16 Apr 2024	Prob (F-statistic):	2.31e-07
Time:	09:10:56	Log-Likelihood:	141.77
No. Observations:	116	AIC:	-279.5

Df Residuals: 114 BIC: -274.0
Df Model: 1
Covariance Type: nonrobust

	coef	std err	t	P> t	[0.025	0.975]
Intercept	0.0133	0.007	1.928	0.056	-0.000	0.027
MKT_rp	1.1127	0.202	5.505	0.000	0.712	1.513
Omnibus:	2.296	Durbin-Watson:	1.950			
Prob(Omnibus):	0.317	Jarque-Bera (JB):	2.128			
Skew:	-0.331	Prob(JB):	0.345			
Kurtosis:	2.951	Cond. No.	30.3			

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

$$\alpha = 0.0133$$

$$\beta = 1.1127$$

Q3.f

```
[ ]: # f. AAPL_adj_rf
capm_adj = smf.ols('AAPL_Adj_rp ~ MKT_rp', data=data)
result_capm_adj = capm_adj.fit()
print(f'result_capm_adj.summary():\n{result_capm_adj.summary()}\n')
```

result_capm_adj.summary():

OLS Regression Results						
Dep. Variable:	AAPL_Adj_rp	R-squared:	0.212			
Model:	OLS	Adj. R-squared:	0.205			
Method:	Least Squares	F-statistic:	30.71			
Date:	Tue, 16 Apr 2024	Prob (F-statistic):	1.96e-07			
Time:	09:10:56	Log-Likelihood:	142.41			
No. Observations:	116	AIC:	-280.8			
Df Residuals:	114	BIC:	-275.3			
Df Model:	1					
Covariance Type:	nonrobust					
	coef	std err	t	P> t	[0.025	0.975]
Intercept	0.0146	0.007	2.122	0.036	0.001	0.028
MKT_rp	1.1141	0.201	5.542	0.000	0.716	1.512
Omnibus:	2.567	Durbin-Watson:	1.945			

Prob(Omnibus):	0.277	Jarque-Bera (JB):	2.312
Skew:	-0.346	Prob(JB):	0.315
Kurtosis:	3.009	Cond. No.	30.3

=====

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

$\alpha = 0.0146$

$\beta = 1.1141$

Q3.g

```
[ ]: # g. Plot fitted results
fig = plt.figure()          # initialize figure window
fig.subplots_adjust(hspace=.5, wspace=0.4) # Use this to do some adjustments

# AAPL
x = data['MKT_rp']          # Market risk premium
y = data['AAPL_rp']        # Risk premium of individual stock
b = result_capm.params     # Estimated parameters

x_range = np.linspace(data['MKT_rp'].min(), data['MKT_rp'].max(), num = 200)
ax = fig.add_subplot(1, 2, 1)

plt.plot(x, y, color = 'blue', marker = 'o', linestyle = '')
plt.plot(x_range, b[0] + b[1]*x_range, color = 'red',
         linestyle = '--', linewidth = 2, label = 'Est. CAPM')
plt.xlabel('Mkt - Rf')
plt.ylabel('AAPL - Rf')
plt.legend()
ax.set_title('AAPL')        ## Use this to make a title for the plot

# AAPL_adj
x = data['MKT_rp']          # Market risk premium
y = data['AAPL_Adj_rp']    # Risk premium of individual stock
b = result_capm_adj.params # Estimated parameters

x_range = np.linspace(data['MKT_rp'].min(), data['MKT_rp'].max(), num = 200)
ax = fig.add_subplot(1, 2, 2)

plt.plot(x, y, color = 'blue', marker = 'o', linestyle = '')
plt.plot(x_range, b[0] + b[1]*x_range, color = 'red', linestyle = '--',
         linewidth = 2, label = 'Est. CAPM')
plt.xlabel('Mkt - Rf')
plt.ylabel('Risk Premium')
```

```
ax.set_title('AAPL Adjusted')
plt.legend()

# Plot and save the plot
plt.savefig('capm_plots.png')
```

