

1. (Normal- χ^2 relationship) Show that if $Z \sim N(0, 1)$ then $Z^2 \sim \chi_1^2$.

$Z \sim N(0, 1)$: PDF is $\phi(z) = (2\pi)^{-1/2} \exp\{-\frac{1}{2}z^2\}$; CDF is $\Phi(z)$

if $Y = Z^2$, the CDF is

$$\begin{aligned} P(Y \leq y) &= P(Z^2 \leq y) = P(-\sqrt{y} \leq Z \leq \sqrt{y}) = P(-\sqrt{y} < Z \leq \sqrt{y}) \\ &= P(Z \leq \sqrt{y}) - P(Z \leq -\sqrt{y}) \\ &= \Phi(\sqrt{y}) - \Phi(-\sqrt{y}) \end{aligned}$$

then the PDF is:

$$\begin{aligned} f(y) &= \frac{d}{dy} P(Y \leq y) = \frac{1}{2\sqrt{y}} \phi(\sqrt{y}) + \frac{1}{2\sqrt{y}} \phi(-\sqrt{y}) \\ &= \frac{1}{2\sqrt{y}} [\phi(\sqrt{y}) + \phi(\sqrt{y})] \quad \phi(z) = \phi(-z) \\ &= \frac{1}{\sqrt{y}} \phi(\sqrt{y}) \\ &= y^{-1/2} \frac{1}{\sqrt{2\pi}} \exp\{-\frac{1}{2}y\} \\ &= \underbrace{\frac{1}{\Gamma(1/2) 2^{1/2}}}_{\Gamma(1/2, 2) \text{ PDF}} y^{1/2-1} e^{-y/2}, \quad y > 0 \end{aligned}$$

score out of 4

(1pt) identify or use the distribution of Z correctly

(1pt) express CDF of Z^2 in terms of CDF of Z
(note: students may also use symmetry properties of the standard normal and may not obtain exactly the expression above)

(1pt) differentiate to obtain PDF

(1pt) express in the form of a gamma/chisq

So $Y \sim \Gamma(1/2, 2)$, i.e., $Y \sim \chi_1^2$

2. (Stochastic ordering) Two random variables X and Y are stochastically ordered if either $F_X(x) \leq F_Y(x)$ or the reverse inequality is true for every $x \in \mathbb{R}$. We say that:

$$X \geq_{st} Y \quad \text{if} \quad F_X(x) \leq F_Y(x) \quad \text{for every} \quad x \in \mathbb{R}$$

Show that the exponential distribution is stochastically ordered in its parameter: that is, if $X \sim \text{exponential}(\alpha)$ and $Y \sim \text{exponential}(\alpha + c)$ where $c > 0$, then $X \geq_{st} Y$. (Use the 'rate' parametrization: $f(x) = \alpha e^{-\alpha x}$, $x > 0$, $\alpha > 0$).

the CDFs are

$$\begin{cases} F_X(x) = 1 - e^{-\alpha x}, & x > 0 \\ F_Y(y) = 1 - e^{-(\alpha+c)y}, & y > 0 \end{cases}$$

note for $c > 0$, $\alpha > 0$, and any $x > 0$

$$e^{-\alpha x} > e^{-(\alpha+c)x}$$

$$1 - e^{-\alpha x} < 1 - e^{-(\alpha+c)x}$$

$$F_X(x) < F_Y(x)$$

so since $F_X(x) < F_Y(x)$ for every $x > 0$ and $F_X(x) = F_Y(y) = 0$ for $x \leq 0$,

$$F_X(x) \leq F_Y(x) \quad \forall x \in \mathbb{R}$$

$$\text{it-}, \quad X \geq_{st} Y$$

3. Let X be a random variable with moment generating function $m_X(t)$. Define $s_X(t) = \log(m_X(t))$. Show that $s'_X(0) = \mathbb{E}X$ and $s''_X(0) = \text{var}(X)$. Then use this approach to find the mean and variance of a random variable X when:

- $X \sim N(\mu, \sigma^2)$
- $X \sim \Gamma(\alpha, \beta)$

note that $m_X(0) = \mathbb{E}e^{0X} = \mathbb{E}(1) = 1$. then:

$$s'_X(0) = \left. \frac{d}{dt} m_X(t) \right|_{t=0} = \frac{m'_X(0)}{m_X(0)} = m'_X(0) = \mathbb{E}X$$

$$s''_X(0) = \left. \frac{d^2}{dt^2} m_X(t) \right|_{t=0} = \frac{m''_X(0)}{m_X(0)} - m'_X(0) \cdot \frac{m'_X(0)}{[m_X(0)]^2}$$

score out of 6

(2pt) express derivatives of s_X in terms of derivatives of m_X , or obtain equivalent solution

(1pt) obtain s_X for the gaussian
(1pt) differentiate to obtain mean and variance

(1pt) obtain s_X for the gamma
(1pt) differentiate to obtain mean and variance

$$= m''_X(0) - [m'_X(0)]^2$$

$$= \mathbb{E}X^2 - [\mathbb{E}X]^2$$

$$= \text{var } X$$

i. $X \sim N(\mu, \sigma^2)$: $m_X(t) = \exp\{\mu t + \frac{1}{2}t^2\sigma^2\}$ so $s_X(t) = \mu t + \frac{1}{2}t^2\sigma^2$

$$s'_X(0) = [\mu + t\sigma^2]_{t=0} = \mu$$

$$s''_X(0) = [\sigma^2]_{t=0} = \sigma^2$$

ii. $X \sim \Gamma(\alpha, \beta)$: $m_X(t) = (1 - \beta t)^{-\alpha}$ so $s_X(t) = -\alpha \log(1 - \beta t)$

$$s'_X(0) = \alpha \beta (1 - \beta t)^{-1} \big|_{t=0} = \alpha \beta$$

$$s''_X(0) = \alpha \beta^2 (1 - \beta t)^{-2} \big|_{t=0} = \alpha \beta^2$$