

1 General
 P-Norm: $\|x\|_p = (\sum_{i=1}^n |x_i|^p)^{\frac{1}{p}}$
 Frobenious Norm: $\|A\|_F = \sqrt{\sum_{i,j} a_{ij}^2}$
 Chain rule: $D(f(g(x))) = Df(g(x)) * Dg(x)$
 Positive (semi-)definiteness: $A \in \mathbb{R}^{n \times n}$ is p.s.d., then A is a real, symmetric matrix and $\forall x \in \mathbb{R}^n. x^T A x > (=) 0$.
 Joint distribution: X, Y are RVs $F_{X,Y}(x, y) = P(X \leq x, Y \leq y)$
 Joint density: $f_{X,Y}(x, y) = \frac{\delta^2 F}{\delta x \delta y}(x, y)$
 Conditional Probability: $P(A|B) = \frac{P(A \cap B)}{P(B)}$
 Total probability: $P(B) = \sum_{i=1}^n P(B|A_i)P(A_i)$
 Bayes rule: $P(A|B) = P(B|A) \frac{P(A)}{P(B)}$
 Variance: $\text{Var}(X) = \mathbb{E}[(X - \mathbb{E}[X])^2] = \mathbb{E}[X^2] - \mathbb{E}[X]^2 \geq 0$
 Gaussian and Multivariate (2D) Gaussian:
 $f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp(-\frac{(x-\mu)^2}{2\sigma^2})$,
 $f(x) = \frac{1}{2\pi\sqrt{|\Sigma|}} \exp(-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu))$.
 Convex function: $f: \mathbb{R}^d \rightarrow \mathbb{R}$ is convex $\Leftrightarrow x_1, x_2 \in \mathbb{R}^d, \lambda \in [0, 1]: f(\lambda x_1 + (1-\lambda)x_2) \leq \lambda f(x_1) + (1-\lambda)f(x_2)$.
 A twice differentiable function $f: \mathbb{R}^d \rightarrow \mathbb{R}$ is convex iff $\forall x \in \mathbb{R}^d$ its Hessian is p.s.d.
 If $w_i \geq 0$ and f_i all conv., then $\sum_i w_i f_i$ conv.

2 Regression
Linear Regression: Goal: Measure distance between predicted and target values
 $f(x) = w_1 x_1 + \dots + w_d x_d + w_0 = \tilde{w}^T \tilde{x}$ with $\tilde{w} = [w_1 \dots w_d, w_0]$ and $\tilde{x} = [x_1 \dots x_d, 1]$.
 Residual: $r_i = y_i - w^T x_i, x_i \in \mathbb{R}^d, y_i \in \mathbb{R}$.
 Objective function (convex): $\hat{R}(w) = \sum_{i=1}^n r_i^2$.
 $w^* = \arg \min_w \sum_{i=1}^n (y_i - w^T x_i)^2$.
 Closed form solution: $w^* = (X^T X)^{-1} X^T y$.
 Gradient: $\nabla_w \hat{R}(w) = -2 \sum_{i=1}^n r_i x_i^T$.
 Non-linear functions: $f(x) = \sum_{i=1}^D w_i \phi_i(x)$.
Gradient Descent: 1. Start $w_0 \in \mathbb{R}^D$.
 2. For $t = 1, 2, \dots$ do $w_{t+1} = w_t - \eta_t \nabla \hat{R}(w_t)$.
Empirical risk minimization: Assumption: Data set generated iid from unknown distribution $P: (x_i, y_i) \sim P(X, Y)$.
 True risk: $R(w) = \int P(x, y)(y - w^T x)^2 dx dy = \mathbb{E}_{x,y}[(y - w^T x)^2]$
 Emp. risk: $\hat{R}_D(w) = \frac{1}{|D|} \sum_{(x,y) \in D} (y - w^T x)^2$
 Generalization error: $|R(w) - \hat{R}_D(w)|$
 Uniform convergence: $\sup_w |R(w) - \hat{R}_D(w)| \rightarrow 0$ as $|D| \rightarrow 0$
 In general, it holds that: $\mathbb{E}_D[\hat{R}_D(\hat{w}_D)] \leq \mathbb{E}_D[\hat{R}(\hat{w}_D)]$, where $\hat{w}_D = \arg \min_w \hat{R}_D(w)$.
Cross-validation: For each model m and for $i = 1 : k$:
 1. Split data: $D = D_{train}^{(i)} \sqcup D_{val}^{(i)}$.

2. Train model: $\hat{w}_{i,m} = \arg \min_w \hat{R}_{train}^{(i)}(w)$.
 3. Estimate error: $\hat{R}_m^{(i)} = \hat{R}_{val}^{(i)}(\hat{w}_{i,m})$.
 Select model: $\hat{m} = \arg \min_m \frac{1}{k} \sum_{i=1}^k \hat{R}_m^{(i)}$.
Ridge regression: Regularization (corresponds to MAP estimation):
 $\min_w \frac{1}{n} \sum_{i=1}^n (y_i - w^T x_i)^2 + \lambda \|w\|_2^2$.
 Closed form: $\hat{w} = (X^T X + \lambda I)^{-1} X^T y$
 Gradient: $-2 \sum_{i=1}^n r_i x_i^T + 2\lambda w$
Standardization: Goal: each feature: $\mu = 0, \sigma^2 = 1$: $\tilde{x}_{i,j} = \frac{(x_{i,j} - \mu_j)}{\sigma_j}$ with $\hat{\mu}_j = \frac{1}{n} \sum_{i=1}^n x_{i,j}, \hat{\sigma}_j^2 = \frac{1}{n} \sum_{i=1}^n (x_{i,j} - \hat{\mu}_j)^2$.
3 Classification
 0/1 loss: $\ell_{0/1}(w; x, y) = [y \neq \text{sign}(w^T x)]$
 Perc. loss: $\ell_p(w; x, y) = \max(0, -y w^T x)$
 Hinge loss: $\ell_H(w; x, y) = \max(0, 1 - y w^T x)$
 $\nabla_w \ell_p = \begin{cases} 0, & \text{if } w^T x_i y_i \geq 0 \text{ (1 if } \ell_H) \\ -y_i x_i & \text{else} \end{cases}$
Fisher consistency: A surrogate loss $\psi: Y \times S \rightarrow \mathbb{R}$, is said to be consistent to the loss $L: Y \times S \rightarrow \mathbb{R}$, if every minimizer f of surrogate risk function $R_\psi(f)$ is also a minimizer of risk function $R_L(f)$. E.g. hinge and logistic losses are consistent with 0/1 loss.
SGD: GD requires sum over all data \rightarrow slow.
 1. Choose random initial $w_0 \in \mathbb{R}^d$
 2. For $t = 1, 2, \dots$ do:
 (a) Choose $(x, y) \in D$ u.a.r (w/ replacement)
 (b) Set $w_{t+1} = w_t - \eta_t \nabla \ell(w_t; x, y)$
 SGD converges if $\sum_t \eta_t = \infty$ and $\sum_t \eta_t^2 < \infty$.
 Mini-batch: Choose multiple datapoints at random; **may converge faster**.
Perceptron: SGD with ℓ_p and $\eta = 1$.
 $\hat{w} = \arg \min_w \frac{1}{n} \sum_{i=1}^n \ell_p(w; x_i, y_i)$
 If data linearly separable finds separator.
SVM: SGD with ℓ_H and regularization.
 $\hat{w} = \arg \min_w \frac{1}{n} \sum_{i=1}^n \ell_H(w; x_i, y_i) + \lambda \|w\|_2^2$
 $w_{t+1} = w_t(1 - 2\eta_t \lambda) + \eta_t x_i y_i [y_i w^T x_i < 1]$
 Often $\eta_t = \frac{1}{\lambda t}$. **Works on non-linearly separable data**, finds best separator w.r.t. ℓ_H .
4 Feature selection
Greedy: Greedily add (or remove) features to maximize cross-validated prediction accuracy w.r.t. cost $c: V \Rightarrow \mathbb{R}$ of using features in subset of V . Forward (start with empty set) **is faster**, but backward is more **resilient to "dependent" features**. Applies to **any method**, but **slow** b/c trains many models and can be **suboptimal**.
L1-Regularization: regularize loss with $\|w\|_1$, automatic feature selection. Can be used for regression (Lasso), classification (L1-SVM) by replacing $\|w\|_2^2$ with $\|w\|_1$. Only works for **linear** models, but is **fast**.
5 Class imbalance
 Downsample **loses data** but **fast**, upsample **random perturbation maybe unsafe**. Use cost-sensitive metrics controlling tradeoff: $\ell_{CS} = c_y \ell(w; x, y)$.

	True label	Positive	Negative	Σ
Predict:	Positive	TP	FP	p_+
	Negative	FN	TN	p_-
	Σ	n_+	n_-	

Accuracy **bad** metric for imbalanced data.
 Accuracy: $\frac{TP+TN}{TP+FP+FN+TN}$
 Precision: $\frac{TP}{TP+FP}$, TPR, Recall: $\frac{TP}{TP+FN}$,
 FPR: $\frac{FP}{TN+FP}$, F1 score: $\frac{2TP}{2TP+FP+FN}$
Precision Recall Curve: Precision (y-axis) vs. Recall (x-axis). Precision = 1 and Recall = 1 is optimal. Area under curve (AUC) can be used for comparison of algos.
Receiver Operator Characteristic (ROC) Curve: TPR (y-axis) vs. FPR (x-axis). Random guessing achieves TPR = FPR line. TPR > FPR is better than random guessing. TPR = 1 and FPR = 0 is optimal. Area under curve (AUC) can be used for comparison of algos.
 Theorem: Alg 1 dominates Alg 2 in terms of ROC curve \Leftrightarrow Alg 1 dominates Alg 2 in terms of Precision Recall curve.
One-vs-all: **c classifiers**, one for each class, pick highest confidence. Class may **not be lin. sep.** from all others. Note **Scaling + Imbalance**.
One-vs-one **c(c-1)/2 classifiers**, voting scheme with highest number of positive prediction wins: **no confidence needed**. Ideally $\mathcal{O}(\log c)$ classifiers, theor. optimum.
Multi-class SVM: c weight vectors, $w^{(y)T} x \geq \max_{i \neq y} \{w^{(i)T} x\} + 1$ for correct label y .
 $\ell_{MC-H}(w^{(1)}, \dots, w^{(c)}; x, y) = \max(0, 1 + \max_{j \neq y} w^{(j)T} x - w^{(y)T} x)$
6 Kernels
Kernel trick: 1. Express problem s.t. it only depends on inner products $x_i^T x_j$.
 2. Replace inner products by kernels.
Reformulate problem: Fundamental insight: Optimal separating hyperplane lives in the span of data: $\hat{w} = \sum_{i=1}^n \alpha_i y_i x_i$, e.g. Perceptron: $\hat{w} = \arg \min_w \frac{1}{n} \sum_{i=1}^n \max(0, -y_i w^T x_i)$.
 $\hat{a} = \arg \min_{\alpha} \frac{1}{n} \sum_{i=1}^n \max(0, -\sum_{j=1}^n \alpha_j y_i y_j x_j^T x_i)$.
Replace inner products: For some feature transform $\phi: x \mapsto \phi(x)$, kernels solve $\phi(x)^T \phi(x')$ **efficiently** as $k(x, x')$.
 Perceptron example (Training):
 1. Initialize $\alpha_1 = \alpha_2 = \dots = \alpha_n = 0$.
 2. For $t = 1, 2, \dots$:
 (a) Pick $(x_i, y_i) \in D$ u.a.r.
 (b) Predict $\hat{y} = \text{sign}(\sum_{j=1}^n \alpha_j y_j k(x_j, x_i))$.
 (c) If $\hat{y} \neq y_i$ set $\alpha_i \leftarrow \alpha_i + \eta_t X \Rightarrow \mathbb{R}$ must be **symmetric**: $k(x, x') = k(x', x)$
 Gram matrix K must be p.s.d. ($\forall x. x^T K x \geq 0$)
 for any n , any set $\{x_1, \dots, x_n\} \subseteq X$. All p.s.d.

matrices are some kernel.

$$K = \begin{bmatrix} k(x_1, x_1) & \dots & k(x_1, x_n) \\ \vdots & \ddots & \vdots \\ k(x_n, x_1) & \dots & k(x_n, x_n) \end{bmatrix}$$

For k_1, k_2 kernel, $c > 0$ and f polyn. with pos. coef. or exp. $k_1 + k_2, k_1 \cdot k_2, c \cdot k_1$ and $f(k_1(x, x'))$ are kernels.
 Poly. degree = d $(x^T x')^d$
 Poly. degree $\leq d$ $(x^T x' + 1)^d$
 Gaussian(RBF) $\exp(-\|x - x'\|_2^2 / (2h^2))$
 Laplacian $\exp(-\|x - x'\|_1 / h)$
 Note: $h > 0$ is bandwidth, $h \rightarrow 0$ overfits.
 k-NN: $y = \text{sign}(\sum_{i=1}^k y_i x_i)$ is k-NN of x]
No training, but **depends on all data**.
 Can use kernel as **similarity function**: $y = \text{sign}(\sum_{i=1}^n y_i \alpha_i k(x_i, x))$, **Improved performance**, depends only on **wrongly classified data**, can capture **global trends**, but **requires training**.
Parametric vs. nonparametric learning:
 Parametric have finite set of parameters (regression, perceptron), while *nonparametric* increase complexity with size of data (kernelized perceptron, k-NN).
 Can kernelize other tasks, such as SVM: $\arg \min_{\alpha} \frac{1}{n} \sum_{i=1}^n \max\{0, 1 - y_i \alpha^T k_i\} + \lambda \alpha^T D_y K D_y \alpha$ with $k_i = [y_1 k(x_1, x_i), \dots, y_n k(x_n, x_i)]$.
Linear regression: Training: $\hat{\alpha} = \arg \min_{\alpha} \frac{1}{n} \|\alpha^T K - y\|_2^2 + \lambda \alpha^T K \alpha$, Closed form: $\hat{\alpha} = (K + n\lambda I)^{-1} y$, Prediction: $\hat{y} = \sum_{i=1}^n \hat{\alpha}_i k(x_i, x)$
 Est. kernel parameters via CV. Choosing kernels requires **domain knowledge**. Deal w/ overfit by regularization.
7 Neural Networks
 Parametrized feature maps. Can learn nonlinear features. **Forward propagation** in layer ℓ : $(v^{(0)} = x): v^{(\ell)} = \varphi(z^{(\ell)}) = \varphi(W^{(\ell)} v^{(\ell-1)})$.
 Output layer: $y = f = W^{(L-1)} v^{(L-1)}$.
 Weight optimization: Apply loss $\ell(y - f(x, W))$, optimize weights to minimize loss.
 For multi-outputs sum losses.
 Sigmoid: $\frac{1}{1+\exp(-z)}$, Tanh: $\frac{\exp(z) - \exp(-z)}{\exp(z) + \exp(-z)}$,
 ReLu: $\max(0, z)$. Sigmoid and Tanh: **diff. everywhere**, but gradient ≈ 0 if $x \neq 0$. ReLu: **not diff.** at 0, $\phi' = 1$ if $z > 0$.
 Back prop.: $\delta^{(L)} = l'(f) = [l'(f_1), \dots, l'(f_p)]$, and $\nabla_{W^{(L)}} \ell(W; y, x) = \delta^{(L)} v^{(L-1)T}$. For $\ell < L$: $\delta^{(\ell)} = \varphi'(z^{(\ell)}) \odot (W^{(\ell+1)T} \delta^{(\ell+1)})$ and $\nabla_{w^{(\ell)}} \ell(W; y, x) = \delta^{(\ell)} v^{(\ell-1)T}$
Nonconvex optimization, initialization matters.

Glorot (tanh): $w_{i,j} \sim \mathcal{N}(0, 1/n_{in})$
 $w_{i,j} \sim \mathcal{N}(0, 2/(n_{in} + n_{out}))$
 He (ReLU): $w_{i,j} \sim \mathcal{N}(0, 2/(n_{in}))$
 Learning rate η_t decreasing (e.g. $\min(0.1, 100/t)$) to prevent oscillation.
 Momentum ($m < 1, a \leftarrow m \cdot a + \eta_t \nabla_w \ell(W; y, x), W \leftarrow W - a$), avoid local minima.
 Many parameters \rightarrow overfitting! Early stop or regularization ($\lambda \|W\|_F^2$) to prevent. Also dropout, randomly set weights to 0 with prob. p , set $W = W \odot p$ after training.
 Batch normalization, standardize some batch $\{x_{1..m}\}$ and set $y_i = \gamma \hat{x}_i + \beta = BN_{\gamma, \beta}(x_i)$. Then $\varphi(Wx) = \varphi(W(BN_{\gamma, \beta}(x)))$.
 Convolutional NN: Apply $m_f \times f$ filters to $n \times n$ image, padding p and stride s : Leaves with $\alpha \times \alpha \times m$ output, where $\alpha = \frac{n+2p-f}{s}$.
8 Unsupervised Learning
k-Means clust.: Assign point $x_i \in \mathbb{R}^d$ to nearest center $\mu_j \in \mathbb{R}^d$. $\hat{R}(\mu) = \sum_{i=1}^n \min_{j \in [k]} \|x_i - \mu_j\|_2^2$, nonconvex.
 Lloyd: Initialize $\mu^{(0)}$. Then assign x_i to closest center $z_i^{(t)} = \arg \min_j \|x_i - \mu_j^{(t-1)}\|_2^2$. Then update mean: $\mu_j^{(t)} = \frac{1}{|i: z_i^{(t)} = j|} \sum_{i: z_i^{(t)} = j} x_j$.
 $\mathcal{O}(nkd)$ per it., converg. **pot. slowly** but **monotonically** to **local optimum** \rightarrow **Mult. iter.**
k-Means++: Let $\mu^{(0)} = x$ u.a.r. from X . Assign centers $2 \dots k$ randomly, prop. to sq. dist. to closest sel. cent. $\mathbb{E}[\text{cost}]$ is $\mathcal{O}(\log k)$ of opt. $\Pr[\mu_j = x_i] = \frac{1}{2} \min_{k \in [j-1]} \|\mu_k - x_i\|_2^2$. Choosing k **difficult**. Heuristic: When $k+1$ yields **diminishing returns**, or **regularization** with λk . Only models circ. clust.: use kernels.
Dimension Reduction: Embed $\{x_1, \dots, x_n\}$, $x_i \in \mathbb{R}^d$ in \mathbb{R}^k where $k \leq \frac{d}{n}$.
PCA: Center data $\mu = \frac{1}{n} \sum_{i=1}^n x_i = 0$ and construct emp. cov. matrix: $\Sigma = \frac{1}{n} \sum_{i=1}^n x_i x_i^T$.
 PCA problem: $(W, z_1, \dots, z_n) = \arg \min_{W, z} \sum_{i=1}^n \|W z_i - x_i\|_2^2$ where W orthogonal, $z_i \in \mathbb{R}^k$ has sol. $z_i = W^T x_i$. Sol. for W is k principal eigenvectors of Σ .
 $W = (v_1 | \dots | v_k)$ where $\Sigma = \sum_{i=1}^d \lambda_i v_i v_i^T$ with $\lambda_1 \geq \dots \geq \lambda_d \geq 0$. Can apply to any matrix $X = USV^T$, first k columns of V are first k principal components. Choose k by CV for feature ind. or s.t. var is explained (see k-Means).
Kernel PCA: $w = \sum_{j=1}^n \alpha_j \phi(x_j)$, $K = \sum_{i=1}^n \lambda_i v_i v_i^T$, $\arg \max_{\|w\|_2=1} \sum_{i=1}^n (w^T \phi(x_i))^2 = \arg \max_{\alpha^T K \alpha = 1} \alpha^T K^T K \alpha$, $\alpha^{(i)} = v_i / \sqrt{\lambda_i}$.
 New point: $z_i = w^{(i)T} x = \sum_{j=1}^n \alpha_j^{(i)} k(x, x_j)$.
 Centering a kernel: $K' = K - KE - EK +$

EKE where $E = \frac{1}{n} [1, \dots, 1] [1, \dots, 1]^T$. Complexity grows with the # of data points.
Autoencoder: NN where hidden layers usually smaller (k) than in- and output (d). Try to learn identity function. Compr. from input to smallest HL, Decompr. from smallest HL to output. $f(x; \theta) = f_{dec}(f_{enc}(x; \theta_1); \theta_2)$. If $\varphi(z) = z$ then autoencoder is equivalent to PCA.
9 Probabilistic modeling
Bayes optimal predictor: $h^*(x) = \mathbb{E}[Y | X = x]$, **unattainable in pr.** Can try to estimate conditional distr. $\hat{P}(Y | X)$.
Parametric estimation: Have $\hat{P}(Y | X, \theta)$, MLE: $\theta^* = \arg \max_{\theta} \hat{P}(Y | X, \theta) = \arg \min_{\theta} - \sum_{i=1}^n \log \hat{P}(y_i | x_i, \theta)$.
 E.g. Gauss. noise, lin. reg.: $y_i \sim \mathcal{N}(w^T x_i, \sigma^2)$
 $\arg \min_{\theta} = \frac{n}{2} \log(2\pi\sigma^2) \sum_{i=1}^n \frac{(y_i - w^T x_i)^2}{2\sigma^2}$, MLE equiv. to LSQ est.. In general, MLE with gauss. noise with const. var. **equiv.** to LSQ sol.
Bias Variance Tradeoff: Prediction error = Bias² + Variance + Noise. Bias: Risk of best model compared to minimal risk given $P(X, Y)$. Variance: Risk due to estimating model from limited data. Noise: Risk by optimal model. **Trade bias and variance via model selection / regularization**
MAP est.: $\arg \max_w P(w | x_{1:n}, y_{1:n}) = \arg \max_w \frac{P(w)P(y_{1:n} | x_{1:n}, w)}{P(y_{1:n} | x_{1:n})} = \arg \min_w - \log P(w) - \log P(y_{1:n} | x_{1:n}, w)$.
 For Gaussian noise and Gaussian prior: MAP = Ridge regression. Regularized estimation can often be understood as MAP inference: $\arg \min_w \sum_{i=1}^n \ell(w^T x_i; x_i, y_i) + C(w) = \arg \max_w \prod_i P(y_i | x_i, w) P(w) = \arg \max_w P(w | D)$ where $C(w) = -\log P(w)$ and $\ell(w^T x_i; x_i, y_i) = -\log P(y_i | x_i, w)$.
 1. Choose likelihood function \rightarrow loss function
 2. Choose prior \rightarrow regularizer
 3. Optimize for MAP parameters, choose hyperparameters through cross-validation
 4. Make predictions via B. Decision Theory
Bayes optimal classifier: $h^*(x) = \arg \max_y P(Y = y | X = x)$
Logistic regression: $P(Y = y | x) = \frac{1}{1 + \exp(-y w^T x)}$. Replaces Gauss. noise as-sympt. with Ber. noise: $P(y | x, w) = \text{Ber}(y; \sigma(w^T x))$. MLE: $\hat{R}(w) = \sum_{i=1}^n \log(1 + \exp(-y_i w^T x_i))$, is **convex** and $\nabla_w \ell_w = \frac{1}{1 + \exp(-y w^T x)} \exp(-y w^T x) (-y x)$, and $\exp(-y w^T x) = 1$ if misclassified. Thus $\nabla_w \ell_w = \frac{-y x}{1 + \exp(y w^T x)}$. With L2 regularizer, take step in direction $w(1 - 2\lambda \eta_t) - \eta_t \nabla_w \ell_w(y, x)$. Nonlinear classification via

kernels: $\hat{\alpha} = \arg \min_{\alpha} \sum_{i=1}^n \log(1 + \exp(-y_i \alpha^T K_i)) + \lambda \alpha^T K \alpha$ and $\hat{P}(y | x, \hat{\alpha}) = \frac{1}{1 + \exp(-y \sum_{j=1}^n \alpha_j k(x_j, x))}$ with $w = \sum_i \alpha_i x_i$.
 M-class: $P(Y = i | x, w_{[c]}) = \frac{\exp(w_i^T x)}{\sum_{j=1}^c \exp(w_j^T x)}$.
Can obtain class probabilities, but dense sols.
10 Decision Theory
 Have $P(y | x)$, actions \mathcal{A} and cost C : $\mathcal{Y} \times \mathcal{A} \rightarrow R$. Min. Exp. cost: $a^* = \arg \min_a \mathbb{E}_y[C(y, a) | x]$. Asymmetric costs: Assume $A = \{-1, +1\}$, $c_{FP}, c_{FN} > 0$. Let $\hat{P}(y | x) = p$, then $\mathbb{E}_y[C(y, +1)] = (1-p)c_{FP}$ and $\mathbb{E}_y[C(y, -1)] = pc_{FN}$. Predict +1 when $p > \frac{c_{FP}}{c_{FP} + c_{FN}}$. Uncertainty sampling: Ask user to label example we most uncertain about: $i_t \in \arg \min_i |0.5 - \hat{P}(Y_i | x_i)|$. Active learning **violates i.i.d. assumption.**
11 Generative Modeling
 Est. joint distribution $P(X, Y)$ instead of $P(Y | X)$. Pred. $y = \text{sign}(f(x))$. Cond. distr. derived from joint: 1. Est. $P(Y)$, 2. Est. $P(x | y) \forall y$, 3. Use Bayes' rule: $P(y | x) = P(y)P(x | y) / P(x)$ where $P(x) = \sum_{y'} P(x, y')$. For $c = 2$ discriminant function $f(x) = \log \frac{P(Y=1|x)}{P(Y=-1|x)}$, which is +1 if $P(Y = 1 | x) > 0.5$.
Naive Bayes: Model y as categorical, and features **conditionally independent** given y , e.g. Gaussian NB, assumes $P(x_i | y) = \mathcal{N}(x_i | \mu_{y,i}, \sigma_{y,i}^2)$. Is linear class., equiv. to log. reg. if assumpt. are met. $f(x) = w^T x + w_0$ where $w_0 = \log \frac{p_+}{1-p_+} + \sum_{i=1}^d \frac{\mu_{-,i}^2 - \mu_{+,i}^2}{2\sigma_i^2}$ and $w_i = \frac{\mu_{+,i} - \mu_{-,i}}{\sigma_i^2}$. Cond. indep. assumpt. can make pred. **overconfident**. # parameters = $\mathcal{O}(cd)$, Compl. (mem. + inference) **lin. in d**.
Categorical Naive Bayes: $P(X_i = c | Y = y) = \theta_{c|y}^{(i)}$. MLE prior: $\hat{p}_y = \frac{\#y}{n}$, MLE feat. distr.: $\theta_{c|y}^{(i)} = \frac{\text{Count}(X_i=c, Y=y)}{\#y}$. **Exponentially (in d) parameters, Way to overfit.**
Gaussian Bayes: $P(x | y) = \mathcal{N}(x; \mu_y, \Sigma_y)$. MLE prior: $\hat{p}_y = \frac{\#y}{n}$, MLE feat. distr.: $\hat{\mu}_k = \frac{1}{\#y} \sum_{i: y_i=y} x_i$ and $\hat{\Sigma}_y = \frac{1}{\#y} \sum_{i: y_i=y} (x_i - \hat{\mu}_y)(x_i - \hat{\mu}_y)^T$. $f(x) = \log \frac{p_-}{1-p_-} + \frac{1}{2} [\log \frac{|\hat{\Sigma}_-|}{|\hat{\Sigma}_+|} + ((x - \hat{\mu}_-)^T \hat{\Sigma}_-^{-1} (x - \hat{\mu}_-) - ((x - \hat{\mu}_+)^T \hat{\Sigma}_+^{-1} (x - \hat{\mu}_+))]$. **Capt. corr. of feat., no overconf., # param. = $\mathcal{O}(cd^2)$, compl. quadr. in d.**

Fischer's LDA: Assume $p = 0.5$ and $\hat{\Sigma}_- = \hat{\Sigma}_+ = \hat{\Sigma}$. Then $f(x) = x^T \hat{\Sigma}^{-1} (\hat{\mu}_+ - \hat{\mu}_-) + \frac{1}{2} (\hat{\mu}_-^T \hat{\Sigma}^{-1} \hat{\mu}_- - \hat{\mu}_+^T \hat{\Sigma}^{-1} \hat{\mu}_+)$.
 Equiv. to log. reg. if assumpt. are met. LDA can be viewed as proj. to a 1-dim. subsp. that maxes ratio of between-class and within-class var.. PCA ($k=1$) maxes the var. of the res. 1-dim. proj.. Gen. model, used to **detect outliers**, not rob. against viol. of **normality of X assumpt.**
Conjugate distr.: Posterior is same family as prior. Ex.: Beta($\theta, \alpha_+, \alpha_-$) and Beta($\theta, n_+ + \alpha_+, n_- + \alpha_-$).
12 Generalized Mixture Models
 Labels maybe unknown, want to cluster data. Model $P(x, \theta)$ as **conv. comb.** of Gauss. distr.: $\sum_{i=1}^c w_i \mathcal{N}(x; \mu_i, \Sigma_i)$ with $w_i > 0$ and $\sum_i w_i = 1$. $(\mu^*, \Sigma^*, w^*) = \arg \min - \sum_{i=1}^n \log \sum_{j=1}^k w_j \mathcal{N}(x_i; \mu_j, \Sigma_j)$.
 Constr. (Σ p.s.d.) **hard to maint.** in SGD. Fitting GMM = Train a GBC without labels.
Hard-EM: E-step: Pred. most likely class for all x_i : $z_i^{(t)} = \arg \max_z P(z | x_i, \theta^{(t-1)}) = \arg \max_z P(z | \theta^{(t-1)}) P(x_i | z, \theta^{(t-1)})$ with $\theta^{(t)} = [w_{1:c}^{(t)}, \mu_{1:c}^{(t)}, \Sigma_{1:c}^{(t)}]$. M-step: Comp. $\theta^{(t)}$ as MLE as for GBC: $\theta^{(t)} = \arg \max_{\theta} P(D^{(t)} | \theta)$. **Too much info** extr. from each label, **overlap-ping** clusters not detected, **fixed label** if model uncertain. k-Means Algo is equiv. to Hard-EM if $w_j = 1/k$ and $\Sigma_j = I \cdot \sigma^2$.
Soft-EM: $\gamma_j(x)$ is prob. x in clust. j : $\gamma_j(x) = \frac{w_j P(x | \Sigma_j, \mu_j)}{\sum_{\ell} w_{\ell} P(x | \Sigma_{\ell}, \mu_{\ell})}$.
 E-Step: Calculate prob $\gamma_j(x_i)$ for all i, j based on $\mu^{(t-1)}, \Sigma^{(t-1)}$ and $w^{(t-1)}$.
 M-Step: Adjust parameters: $w_j^{(t)} = \frac{1}{n} \sum_{i=1}^n \gamma_j^{(t)}(x_i)$, $\mu_j^{(t)} = \frac{\sum_{i=1}^n \gamma_j^{(t)}(x_i) x_i}{\sum_{i=1}^n \gamma_j^{(t)}(x_i)}$ and $\Sigma_j^{(t)} = \frac{\sum_{i=1}^n \gamma_j^{(t)}(x_i) (x_i - \mu_j^{(t)})(x_i - \mu_j^{(t)})^T}{\sum_{i=1}^n \gamma_j^{(t)}(x_i)}$.
Init. sensitive, nonconvex objective. Init. w with unif. distr., μ by k-means++ and Σ as spherical (acc. to emp. var. in data). Choose k **via CV**. Avoid **degeneracy** by **add. term $\nu^2 I$** to $\Sigma_j^{(t)} \rightarrow$ Wishart-prior.
Semi-supervised learning: For points with label y_i : $\gamma_j^{(t)}(x_i) = [j = y_i]$
GANs: $\min_{w_G} \max_{w_D} \mathbb{E}_{x \sim \text{Data}} \log D(x; w_D) + \mathbb{E}_{x \sim \mathcal{N}} \log(1 - D(G(z; w_G); w_D))$, **Mode Collapse:** gen. prod. limit. variety of samples, **Data memorization**, Simult. training \rightarrow **oscillations, Can't comp. likelih. on holdout set.**