biostatistical methods homework 4

library(tidyverse)
library(knitr)
library(patchwork)
library(readxl)

Problem1

(a)

$$b_{1} = \frac{n \sum X_{i}Y_{i} - \sum X_{i} \sum Y_{i}}{n \sum X_{i}^{2} - (\sum X_{i})^{2}}$$

$$= \frac{\sum X_{i}Y_{i} - n\bar{Y}\bar{X}}{\sum X_{i}^{2} - n\bar{X}^{2}}$$

$$b_{0} = \bar{Y} - b_{1}\bar{X}$$

$$\sum X_{i}Y_{i} - n\bar{Y}\bar{X} = \sum X_{i}Y_{i} - \bar{X} \sum Y_{i}$$

$$= \sum (X_{i} - \bar{X})Y_{i}$$

$$E\left\{\sum (X_{i} - \bar{X})Y_{i}\right\} = \sum (X_{i} - \bar{X})E(Y_{i})$$

$$= \sum (X_{i} - \bar{X})(\beta_{0} + \beta_{1}X_{i})$$

$$= \beta_{0} \sum X_{i} - n\bar{X}\beta_{0} + \beta_{1} \sum X_{i}^{2} - n\bar{X}^{2}\beta_{1}$$

$$= \beta_{1}(\sum X_{i}^{2} - n\bar{X}^{2})$$

$$E(b_{1}) = \frac{E\left\{\sum (X_{i} - \bar{X})Y_{i}\right\}}{\sum X_{i}^{2} - n\bar{X}^{2}}$$

$$= \frac{\beta_{1}(\sum X_{i}^{2} - n\bar{X}^{2})}{\sum X_{i}^{2} - n\bar{X}^{2}}$$

$$= \beta_{1}$$

$$E(b_{0}) = E(\bar{Y} - b_{1}\bar{X})$$

$$= \frac{1}{n} \sum E(Y_{i}) - E(b_{1})\bar{X}$$

$$= \frac{1}{n} \sum [\beta_{0} + \beta_{1}X_{i}] - \beta_{1}\bar{X}$$

$$= \frac{1}{n} [n\beta_{0} + n\beta_{1}\bar{X}] - \beta_{1}\bar{X}$$

(b)

$$Y_i = \hat{\beta}_1 X_i + \hat{\beta}_0$$
$$= \hat{\beta}_1 X_i + \bar{Y} - \hat{\beta}_1 \bar{X}$$
$$X_i = \bar{X}$$

$$X_i = X$$

$$Y_i = \hat{\beta}_1 \bar{X} + \bar{Y} - \hat{\beta}_1 \bar{X}$$

$$= \bar{Y}$$

So the Least Square line equation always goes through the point (\bar{X}, \bar{Y})

(c)

$$log_{e}L = -\frac{n}{2}log_{e}2\pi - \frac{n}{2}log_{e}\sigma^{2} - \frac{1}{2\sigma^{2}}\sum(Y_{i} - \beta_{0} - \beta_{1}X_{i})^{2}$$

$$\frac{\partial(log_{e}L)}{\partial\sigma^{2}} = -\frac{n}{2\sigma^{2}} + \frac{1}{2\sigma^{4}}\sum(Y_{i} - \beta_{0} - \beta_{1}X_{i})^{2}$$

$$\hat{\sigma}^{2} = \frac{\sum(Y_{i} - \hat{\beta}_{0} - \hat{\beta}_{1}X_{i})^{2}}{n}$$

$$= \frac{\sum(Y_{i} - \hat{Y}_{i})^{2}}{n}$$

Find its expected value

$$E(\hat{\sigma}^2) = E\left(\frac{SSE}{n}\right)$$

$$= E\left(\frac{SSE}{n-2} \times \frac{n-2}{n}\right)$$

$$= \frac{n-2}{n} \times E\left(\frac{SSE}{n-2}\right)$$

$$= \frac{n-2}{n}\sigma^2$$

Comment on the unbiasness property

As the result shown above, $\hat{\sigma}^2$ is a biasd estimator of σ^2 as the unbiased estimator of σ^2 is MSE:

$$s^{2} = MSE = \frac{SSE}{n-2} = \frac{\sum (Y_{i} - \hat{Y}_{i})^{2}}{n-2} = \frac{\sum e_{i}^{2}}{n-2}$$

$$E\{MSE\} = \sigma^{2}$$

Problem 2

For this problem, you will be using data 'HeartDisease.csv'.

```
heart_data = read_csv("./data/HeartDisease.csv")
```

The investigator is mainly interested if there is an association between 'total cost' (in dollars) of patients diagnosed with heart disease and the 'number of emergency room (ER) visits'.

Further, the model will need to be adjusted for other factors, including 'age', 'gender', 'number of complications' that arose during treatment, and 'duration of treatment condition'.

a)

Provide a short description of the data set: what is the main outcome, main predictor and other important covariates.

This dataset include 10 variables and 788 observations. The main outcome is totalcost which represents the total cost (in dollars) of heart-diseased patients. The main predictor is the ERvisits which represents the number of emergency room (ER) visits. Other important covariates are age, gender, complications and duration.

Also, generate appropriate descriptive statistics for all variables of interest (continuous and categorical) – no test required.

```
mean_and_sd = function(x) {

if (!is.numeric(x)) {
    stop("Argument x should be numeric")
} else if (length(x) == 1) {
    stop("Cannot be computed for length 1 vectors")
}

mean_x = mean(x)
sd_x = sd(x)

tibble(
    mean = mean_x,
    sd = sd_x
)
}
```

totalcost

1 2800. 6690.

```
mean_and_sd(heart_data$totalcost)

## # A tibble: 1 x 2

## mean sd

## <dbl> <dbl>
```

The mean of the total cost is about 2800 with a standard deviation of 6690.26.

ERvisits

```
summary(heart_data$ERvisits)
```

```
## Min. 1st Qu. Median Mean 3rd Qu. Max.
## 0.000 2.000 3.000 3.425 5.000 20.000
```

The minimum number of emergency room (ER) visits is 0 and the maximum is 20. The median is 3 with 1st Qu. of 2 and 3rd Qu. of 5.

age

```
mean_and_sd(heart_data$age)
```

```
## # A tibble: 1 x 2
## mean sd
## <dbl> <dbl>
## 1 58.7 6.75
```

The distribution of age is centered at about 59 with a standard deviation of 6.75.

gender

```
(summary(as.factor(heart_data$gender)))
## 0 1
## 608 180
```

As 0 represents female and 1 represents male, there are 608 female and 180 male in the dataset.

complications

```
(summary(as.factor(heart_data$complications)))
## 0 1 3
## 745 42 1
```

As we observed from the dataset, there number of complications existing in this dataset is simply 0, 1 and 3. Using summary function, we can conclude that there are 745 patients have zero complications and 42 patients have one complications, and there is only 1 patient has 3 complication.

duration

1 164. 121.

```
mean_and_sd(heart_data$duration)

## # A tibble: 1 x 2

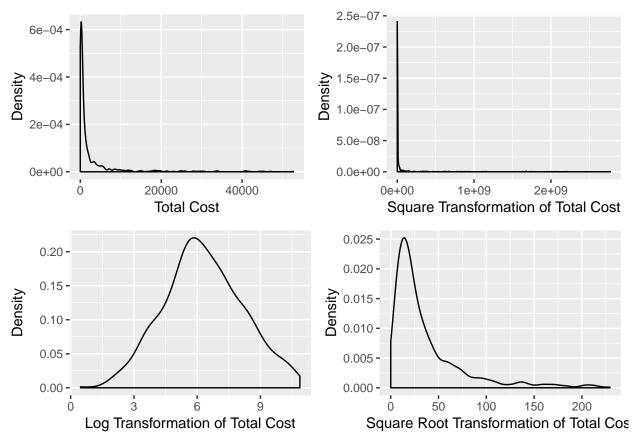
## mean sd
## <dbl> <dbl>
```

The average duration of treatment condition is 164 with a standard deviation of 121.

b)

```
totalcost_non = heart_data %>%
  ggplot(aes(x = totalcost)) +
  geom_density() +
  labs(
       x = 'Total Cost',
       y = 'Density'
totalcost_sq = heart_data %>%
  ggplot(aes(x = (totalcost)^2)) +
  geom_density() +
  labs(
       x = 'Square Transformation of Total Cost',
       y = 'Density'
totalcost_log = heart_data %>%
  ggplot(aes(x = log(totalcost))) +
  geom_density() +
  labs(
       x = 'Log Transformation of Total Cost',
       y = 'Density'
       )
totalcost_sqrt = heart_data %>%
  ggplot(aes(x = sqrt(totalcost))) +
  geom_density() +
  labs(
      x = 'Square Root Transformation of Total Cost',
      y = 'Density'
(totalcost_non + totalcost_sq) / (totalcost_log + totalcost_sqrt)
```

Warning: Removed 3 rows containing non-finite values (stat_density).



The shape of the distribution for totalcost is right skewed. After trying different transformation we find that the log transformation makes the plot approximate to normal distribution.

c)

Create a new variable called 'comp_bin' by dichotomizing 'complications': 0 if no complications, and 1 otherwise.

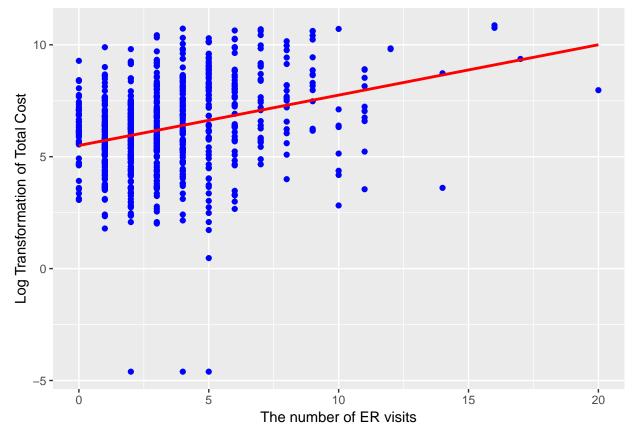
```
heart_data = heart_data %>%
  mutate(comp_bin = ifelse(complications == 0, 0, 1)) %>%
  mutate(comp_bin = as.character(comp_bin))
```

d)

Based on our decision in part b), fit a simple linear regression (SLR) between the original or transformed 'total cost' and predictor 'ERvisits'. This includes a scatterplot and results of the regression, with appropriate comments on significance and interpretation of the slope.

```
heart_data_trans = heart_data
heart_data_trans$totalcost[heart_data_trans$totalcost==0]=0.01
heart_data_trans = heart_data_trans %>%
    mutate(totalcost = log(totalcost))
```

```
heart_data_trans %>%
    ggplot(aes(x = ERvisits, y = totalcost)) +
    geom_point(color = 'blue') +
    geom_smooth(method = "lm", color = 'red', se = FALSE) +
    labs(
        x = 'The number of ER visits',
        y = 'Log Transformation of Total Cost'
        )
```



```
fit_SLR = lm(totalcost ~ ERvisits, data = heart_data_trans)
summary(fit_SLR)
```

```
##
## Call:
## lm(formula = totalcost ~ ERvisits, data = heart_data_trans)
## Residuals:
##
        Min
                  1Q
                       Median
                                    ЗQ
                                            Max
## -11.2321 -1.1013
                       0.0529
                                1.3055
                                          4.3224
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                 5.5016
                            0.1106 49.725
                                             <2e-16 ***
## ERvisits
                 0.2251
                            0.0256
                                     8.792
                                             <2e-16 ***
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
```

```
## Residual standard error: 1.894 on 786 degrees of freedom
## Multiple R-squared: 0.08954, Adjusted R-squared: 0.08838
## F-statistic: 77.3 on 1 and 786 DF, p-value: < 2.2e-16</pre>
```

$$\log\left(\frac{Y_2}{Y_1}\right) = \beta_1 = 0.22672$$

$$\frac{Y_2}{Y_1} = \exp^{0.2251} = 1.25$$

$$Y_2 = 1.25Y_1$$

The plot above shows the scatterplot and results of the regression. Using summary function, we can see that the estimate slope is 0.2251 with a p-value <2e-16, which strongly indicates that the slope is not equal to 0 and there is significant relationship with ERvisits and totalcost. The estimate of slope means that when the number of ER visits increase 1, total cost will increase 25%. However, the adjusted R-squated is 0.08838 which is really close to 0 indicating the SLR is not the best model.

e)

Fit a multiple linear regression (MLR) with 'comp_bin' and 'ERvisits' as predictors.

i)

Test if 'comp_bin' is an effect modifier of the relationship between 'total cost' and 'ERvisits'. Comment.

```
fit_MLR_interaction = lm(totalcost ~ comp_bin * ERvisits, data = heart_data_trans)
summary(fit_MLR_interaction)
```

```
##
## Call:
  lm(formula = totalcost ~ comp_bin * ERvisits, data = heart_data_trans)
##
## Residuals:
##
        Min
                  1Q
                       Median
                                    3Q
                                            Max
## -11.1126 -1.0605
                       0.0257
                                         4.4258
                                1.2181
##
## Coefficients:
##
                      Estimate Std. Error t value Pr(>|t|)
                                  0.11072
## (Intercept)
                       5.46319
                                           49.343 < 2e-16 ***
## comp bin1
                       2.21549
                                  0.58466
                                            3.789 0.000163 ***
## ERvisits
                       0.20884
                                  0.02626
                                            7.954 6.32e-15 ***
## comp_bin1:ERvisits -0.09686
                                  0.10154
                                           -0.954 0.340430
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.855 on 784 degrees of freedom
## Multiple R-squared: 0.1291, Adjusted R-squared: 0.1258
## F-statistic: 38.73 on 3 and 784 DF, p-value: < 2.2e-16
```

The definition of modifier is when the magnitude of association differs at different levels of another variable (in this case comp_bin), it suggests that effect modification is present. From the result shown above, comp_bin is not a modifier according to the p-value of comp_bin1:ERvisits is way larger than 0.05.

ii)

```
Test if 'comp_bin' is a confounder of the relationship between 'total cost' and 'ERvisits'. Comment.
```

```
lm(totalcost ~ ERvisits, data = heart_data_trans) %>%
  summary()
##
## Call:
## lm(formula = totalcost ~ ERvisits, data = heart_data_trans)
## Residuals:
       Min
                  1Q
                      Median
                                    3Q
                                            Max
## -11.2321 -1.1013
                       0.0529
                                1.3055
                                         4.3224
##
## Coefficients:
              Estimate Std. Error t value Pr(>|t|)
                            0.1106 49.725
## (Intercept)
                 5.5016
                                             <2e-16 ***
## ERvisits
                 0.2251
                            0.0256
                                     8.792
                                             <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.894 on 786 degrees of freedom
## Multiple R-squared: 0.08954,
                                    Adjusted R-squared: 0.08838
## F-statistic: 77.3 on 1 and 786 DF, p-value: < 2.2e-16
#coefficient estimate:
#ERvisits: 0.2251
lm(totalcost ~ comp_bin + ERvisits, data = heart_data_trans) %>%
  summary()
##
## lm(formula = totalcost ~ comp_bin + ERvisits, data = heart_data_trans)
## Residuals:
       Min
                  1Q
                       Median
                                    3Q
                                            Max
## -11.1017 -1.0561
                       0.0165
                                1.2104
                                         4.4301
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
                           0.10838 50.607 < 2e-16 ***
## (Intercept) 5.48475
                           0.29432
## comp_bin1
                1.73361
                                     5.890 5.72e-09 ***
## ERvisits
                0.20236
                           0.02536
                                     7.979 5.23e-15 ***
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.855 on 785 degrees of freedom
## Multiple R-squared: 0.1281, Adjusted R-squared: 0.1259
## F-statistic: 57.65 on 2 and 785 DF, p-value: < 2.2e-16
#coefficient estimate:
#ERvisits: 0.20236
```

To calculate the percentage change in the parameter estimate, we use the following formula:

$$\frac{|\beta_{crude} - \beta_{adjusted}|}{|\beta_{crude}|} = \frac{|0.2251 - 0.20236|}{|0.2251|} = 0.1010218$$

Here we use 10% rule of thumb. 0.1010218 is greater than 10%, so we consider comp_bin as a confounder.

iii)

Decide if 'comp_bin' should be included along with 'ERvisits'. Why or why not? Hypothese:

$$\begin{aligned} Model \ 1: \ Y_i &= \beta_0 \ + \ \beta_1 X_{i \ ERvisits} \ + \ \varepsilon_i \\ Model \ 2: \ Y_i &= \beta_0 \ + \ \beta_1 X_{i \ ERvisits} \ + \ \beta_2 X_{i \ comp_bin} \ + \ \varepsilon_i \\ H_0: \beta_2 &= 0 \\ H_1: \beta_2 \neq 0 \end{aligned}$$

Decision rule:

$$F^* = \frac{(SSE_S - SSE_L)/(df_L - df_S)}{\frac{SSE_L}{df_L}} \sim F_{df_L - df_S, df_L}$$

$$df_S = n - p_S - 1, \ df_L = n - p_L - 1$$

$$F^* = 34.694$$

$$Pr(F > F^*) < 0.05, \ reject \ H_0$$

$$Pr(F > F^*) \ge 0.05, \ fail \ to \ reject \ H_0$$

$$Pr(F > F^*) = 5.721e - 09 < 0.05, \ reject \ H_0$$

```
fit_without_comp = lm(totalcost ~ ERvisits, data = heart_data_trans)
fit_with_comp = lm(totalcost ~ ERvisits + comp_bin, data = heart_data_trans)
anova(fit_without_comp, fit_with_comp)
```

```
## Analysis of Variance Table
##
## Model 1: totalcost ~ ERvisits
## Model 2: totalcost ~ ERvisits + comp_bin
              RSS Df Sum of Sq
                                    F
##
    Res.Df
                                         Pr(>F)
       786 2820.0
## 1
## 2
       785 2700.6
                        119.36 34.694 5.721e-09 ***
                  1
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

According to the result of partial F-test, the larger model including comp_bin is prefered. Besides, we already proof that comp_bin is a confounder so it should be included in the model along with ERvisits.

f)

Use your choice of model in part e) and add additional covariates (age, gender, and duration of treatment).

i)

Fit a MLR, show the regression results and comment.

Regression model in e):

```
lm(totalcost ~ comp_bin + ERvisits, data = heart_data_trans) %>%
  summary()
##
## Call:
## lm(formula = totalcost ~ comp_bin + ERvisits, data = heart_data_trans)
## Residuals:
##
        Min
                                    3Q
                                            Max
                  1Q
                       Median
## -11.1017 -1.0561
                       0.0165
                                1.2104
                                         4.4301
##
## Coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 5.48475
                           0.10838 50.607 < 2e-16 ***
                1.73361
                           0.29432
                                     5.890 5.72e-09 ***
## comp_bin1
## ERvisits
                0.20236
                           0.02536
                                     7.979 5.23e-15 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.855 on 785 degrees of freedom
## Multiple R-squared: 0.1281, Adjusted R-squared: 0.1259
## F-statistic: 57.65 on 2 and 785 DF, p-value: < 2.2e-16
lm(totalcost ~ comp_bin + ERvisits + age + gender + duration, data = heart_data_trans) %>%
  summary()
##
## lm(formula = totalcost ~ comp_bin + ERvisits + age + gender +
##
       duration, data = heart_data_trans)
##
## Residuals:
##
       Min
                1Q Median
                                3Q
                                       Max
## -9.9436 -1.0080 -0.0886 0.9771
                                   4.3492
##
## Coefficients:
##
                 Estimate Std. Error t value Pr(>|t|)
## (Intercept) 5.8469672 0.5387065 10.854 < 2e-16 ***
                                       5.593 3.09e-08 ***
## comp_bin1
                1.5258252
                           0.2728204
## ERvisits
                0.1736943
                           0.0238252
                                       7.290 7.58e-13 ***
## age
               -0.0198581
                           0.0091556
                                     -2.169
                                               0.0304 *
## gender
               -0.2848042
                           0.1463906
                                     -1.946
                                               0.0521 .
## duration
                0.0059649
                          0.0005159
                                     11.561 < 2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

```
##
## Residual standard error: 1.714 on 782 degrees of freedom
## Multiple R-squared: 0.2583, Adjusted R-squared: 0.2536
## F-statistic: 54.47 on 5 and 782 DF, p-value: < 2.2e-16
```

Comment:

The result shows that there is a relationship between Y and the set of covariates.

ii)

Compare the SLR and MLR models. Which model would you use to address the investigator's objective and why?

Hypothese:

$$\begin{split} Model \ 1: \ Y_i &= \beta_0 \ + \ \beta_1 X_{i \ ERvisits} \ + \ \varepsilon_i \\ Model \ 2: \ Y_i &= \beta_0 \ + \ \beta_1 X_{i \ ERvisits} \ + \ \beta_2 X_{i \ comp_bin} \ + \ \beta_3 X_{i \ gender} \ + \ \beta_4 X_{i \ age} \ + \ \varepsilon_i \\ H_0: \beta_2 &= \beta_3 = \beta_4 = 0 \\ H_1: at \ least \ one \ \beta \ not \ equal \ to \ zero \end{split}$$

 $F^* = \frac{(SSE_S - SSE_L)/(df_L - df_S)}{\frac{SSE}{df_L}} \sim F_{df_L - df_S, df_L}$

Decision rule:

##

1

2

Conclusion:

$$F^* = 44.49$$

$$Pr(F > F^*) < 0.05, \ reject \ H_0$$

$$Pr(F > F^*) \ge 0.05, \ fail \ to \ reject \ H_0$$

$$Pr(F > F^*) \ge 0.05, \ fail \ to \ reject \ H_0$$
 fit_SLR = lm(totalcost ~ ERvisits, data = heart_data_trans) fit_MLR = lm(totalcost ~ ERvisits + comp_bin + age + gender + duration, data = heart_data_trans) anova(fit_SLR, fit_MLR)
Analysis of Variance Table
Model 1: totalcost ~ ERvisits + comp_bin + age + gender + duration
Res.Df RSS Df Sum of Sq F Pr(>F)
1 786 2820.0
2 782 2297.2 4 522.78 44.49 < 2.2e-16 ***
"---
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1

Given the result of anova test, it's obviously that the MLR is preferred. So, I would choose MLR to address the investigator's objective.

Problem 3

A hospital administrator wishes to test the relationship between 'patient's satisfaction' (Y) and 'age', 'severity of illness', and 'anxiety level' (data 'PatSatisfaction.xlsx'). The administrator randomly selected 46 patients, collected the data, and asked for your help with the analysis.

```
sat_data = read_excel("./data/PatSatisfaction.xlsx")

colnames(sat_data)[1] <- "satisfaction"

sat_data = sat_data %>%
    janitor::clean_names()
```

a)

Create a correlation matrix and interpret your initial findings.

```
cor(sat_data, method = "pearson")
```

```
##
                satisfaction
                                     age
                                           severity
                                                       anxiety
## satisfaction
                   1.0000000 -0.7867555 -0.6029417 -0.6445910
## age
                  -0.7867555
                              1.0000000 0.5679505
                                                     0.5696775
## severity
                  -0.6029417
                              0.5679505
                                          1.0000000
                                                     0.6705287
                  -0.6445910
                              0.5696775
                                         0.6705287
## anxiety
                                                     1.0000000
```

The result is a table containing the correlation coefficients between each variable and the others. We can observe that age, severity and axiety have negative relationship with satisfaction. Among those three variables, age has the strongest negative relationship with satisfaction.

b)

Fit a multiple regression model and test whether there is a regression relation. State the hypotheses, decision rule and conclusion.

To build a multiple regression model, we add age, severity and anxiety as predictors:

$$Y_i = \beta_0 + \beta_1 X_{i \ age} + \beta_2 X_{i \ anxiety} + \beta_3 X_{i \ severity} + \varepsilon_i$$

Hypotheses:

$$H_0: \ \beta_1 = \beta_2 = \beta_3 = 0$$

 $H_1: \ at \ least \ one \ \beta \ is \ not \ zero$

Decision rule:

 $Test\ statistic:$

$$F^* = \frac{MSR}{MSE} > F(1-\alpha; p, n-p-1), \ reject \ H_0.$$

 $The \ null \ model \ contains \ only \ the \ intercept:$

$$F^* = \frac{MSR}{MSE} \leqslant F(1-\alpha; p, n-p-1), \ fail \ to \ reject \ H_0$$

$$F^* = 30.05$$

$$Pr(F > F^*) < 0.05$$
, reject H_0
 $Pr(F > F^*) \ge 0.05$, fail to reject H_0

$$Pr(F > F^*) = 1.542e - 10 < 0.05, reject H_0$$

```
sat_fit = lm(satisfaction ~ age + severity + anxiety, data = sat_data)
summary(sat_fit)
```

```
##
## Call:
## lm(formula = satisfaction ~ age + severity + anxiety, data = sat_data)
##
## Residuals:
##
       Min
                 1Q
                      Median
                                   3Q
                                           Max
## -18.3524 -6.4230
                      0.5196
                               8.3715 17.1601
##
## Coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 158.4913
                          18.1259
                                   8.744 5.26e-11 ***
               -1.1416
                           0.2148 -5.315 3.81e-06 ***
## severity
               -0.4420
                           0.4920 -0.898
                                            0.3741
## anxiety
              -13.4702
                           7.0997 -1.897
                                            0.0647
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 10.06 on 42 degrees of freedom
## Multiple R-squared: 0.6822, Adjusted R-squared: 0.6595
## F-statistic: 30.05 on 3 and 42 DF, p-value: 1.542e-10
```

Judging from the p-value, we reject the null, which means there is a relationship between predictors and outcome.

 $\mathbf{c})$

Show the regression results for all estimated coefficients with 95% CIs. Interpret the coefficient and 95% CI associated with 'severity of illness'.

CI:

$$E(Y_h) = \beta_0 + \beta_1 X_h$$

$$\hat{Y}_h = \hat{\beta}_0 + \hat{\beta}_1 X_h$$

$$\hat{\beta}_0 + \hat{\beta}_1 X_h \pm t_{n-2,1-\alpha/2} \times se(\hat{\beta}_0 + \hat{\beta}_1 X_h)$$

$$se(\hat{\beta}_0 + \hat{\beta}_1 X_h) = \sqrt{MSE\left\{\frac{1}{n} + [(X_h - \bar{X})^2 / \sum_{i=1}^n (X_i - \bar{X})^2]\right\}}$$

```
sat_fit = lm(satisfaction ~ age + severity + anxiety, data = sat_data)
summary(sat_fit)
```

```
##
## Call:
  lm(formula = satisfaction ~ age + severity + anxiety, data = sat_data)
##
##
  Residuals:
##
        Min
                  1Q
                       Median
                                     3Q
                                             Max
                        0.5196
##
   -18.3524 -6.4230
                                 8.3715
                                         17.1601
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
##
  (Intercept) 158.4913
                            18.1259
                                      8.744 5.26e-11 ***
                -1.1416
                                     -5.315 3.81e-06 ***
                             0.2148
                -0.4420
                             0.4920
                                     -0.898
                                              0.3741
## severity
                             7.0997
                                              0.0647 .
## anxiety
               -13.4702
                                     -1.897
## ---
                   0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 10.06 on 42 degrees of freedom
## Multiple R-squared: 0.6822, Adjusted R-squared: 0.6595
## F-statistic: 30.05 on 3 and 42 DF, p-value: 1.542e-10
confint(sat_fit)
##
                    2.5 %
                                97.5 %
## (Intercept) 121.911727 195.0707761
                -1.575093
                           -0.7081303
## severity
                -1.434831
                             0.5508228
## anxiety
               -27.797859
                             0.8575324
```

Interpretation:

The coefficient associated with severity means that while other predictors hold constant the satisfaction will decrease 0.44 for each additional unit of the severity of the illness.

The function of **confint** shows the CIs for each estimated coefficients. From the result we can conclude that we are 95% confidence that the mean satisfaction increases by somewhere between -1.434831 and 0.5508228 for each additional unit of the severity of the illness as other predictors hold constant.

 \mathbf{d}

Obtain an interval estimate for a new patient's satisfaction when Age=35, Severity=42, Anxiety=2.1. Interpret the interval.

For a given value of x, the interval estimate of the dependent variable y is called the prediction interval.

PI:

$$\hat{\beta}_0 + \hat{\beta}_1 X_h \pm t_{n-2,1-\alpha/2} \times se(\hat{\beta}_0 + \hat{\beta}_1 X_h)$$

$$se(\hat{\beta}_0 + \hat{\beta}_1 X_h) = \sqrt{MSE\left\{\frac{1}{n} + [(X_h - \bar{X})^2 / \sum_{i=1}^n (X_i - \bar{X})^2] + 1\right\}}$$

```
pi_data = data.frame(age = 35, severity = 42, anxiety = 2.1)
predict(sat_fit, pi_data, interval="predict")
```

```
## fit lwr upr
## 1 71.68332 50.06237 93.30426
```

The 95% prediction interval of the satisfaction for the age is 35, severity is 42 and anxiety is 2.1 is between 50.06237 and 93.30426. The result means that the probability is 0.95 that this prediction interval will give a correct prediction for the satisfaction when age is 35, severity is 42 and anxiety is 2.1.

e)

Test whether 'anxiety level' can be dropped from the regression model, given the other two covariates are retained. State the hypotheses, decision rule and conclusion.

Hypothese:

$$\begin{split} Model \ 1: \ Y_i &= \beta_0 \ + \ \beta_1 X_i \ _{age} \ + \ \beta_2 X_i \ _{severity} \ + \ \varepsilon_i \\ Model \ 2: \ Y_i &= \beta_0 \ + \ \beta_1 X_i \ _{age} \ + \ \beta_2 X_i \ _{severity} \ + \ \beta_3 X_i \ _{anxiety} \ + \ \varepsilon_i \\ H_0: \beta_3 &= 0 \\ H_1: \beta_3 \neq 0 \end{split}$$

Decision rule:

$$F^* = \frac{(SSE_S - SSE_L)/(df_L - df_S)}{\frac{SSE_L}{df_L}} \sim F_{df_L - df_S, df_L}$$

$$df_S = n - p_S - 1, df_L = n - p_L - 1$$

$$F^* = 3.5997$$

$$Pr(F > F^*) < 0.05, reject H_0$$

$$Pr(F > F^*) \ge 0.05, fail to reject H_0$$

$$Pr(F > F^*) = 0.06468 > 0.05, fail to reject H_0$$

```
anova(lm(satisfaction ~ age + severity, data = sat_data),
    lm(satisfaction ~ age + severity + anxiety, data = sat_data))
```

```
## Analysis of Variance Table
##
## Model 1: satisfaction ~ age + severity
## Model 2: satisfaction ~ age + severity + anxiety
## Res.Df RSS Df Sum of Sq F Pr(>F)
## 1 43 4613.0
## 2 42 4248.8 1 364.16 3.5997 0.06468 .
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

Conclusion:

The result shows that the anxiety should NOT be included in the model due to the large p-value at 0.05 significance level and the desicion rule. So we tend to use the smaller model.