

Estimations

i) m and $\gamma_{(k)}$:

Prop. If (X_t) is stationary with mean m

and autocov. $\gamma_{(k)}$. Then:

$$\gamma_{(n)} \xrightarrow{n \rightarrow \infty} 0 \Rightarrow \text{Var}(\bar{X}_n) \xrightarrow{n \rightarrow \infty} 0.$$

Additionally. $\sum_k |\gamma_{(k)}| < \infty \Rightarrow \text{Var}(\bar{X}_n) \sim \frac{\sum_k \gamma_{(k)}}{n}$

Pf. $\text{Var}(\bar{X}_n) \stackrel{i)}{=} \frac{1}{n} \sum_{h=1}^n (1 - \frac{|h|}{n}) \gamma_{(h)}$

$$\stackrel{ii)}{\leq} \frac{1}{n} \sum_{h=1}^n |\gamma_{(h)}|.$$

Apply Stolz on ii) to obtain the first conclusion. the latter is from i).

Cir. $(X_t)_2$ is stationary erg. st.

$$X_t = m + \sum_j \gamma_j Z_{t-j}, \quad Z_t \sim \text{IID}(0, \sigma^2)$$

If $\sum_k |\gamma_k| < \infty$. Then:

$$\bar{X}_n \sim \text{ANC}(m, \frac{\sum_k \gamma_{(k)}}{n})$$

Then. Under the conditions of cor. above.

If i) $\mathbb{E}(Z_t^4) < \infty$ or ii) $\sum_j |\gamma_j|^2 < \infty$.

Then: $\tilde{\epsilon}_{(k)} \sim \text{ANC}(0, \sigma^2), \quad N(n)$.

where $\hat{c}^{(h)} = (\hat{c}(0), \dots, \hat{c}(h))$. $\hat{c}(h) = \frac{\hat{y}^{(h)}}{\hat{y}(0)}$.

$$w_{ij} = \sum_{k=1}^{\infty} (c(k+i) + c(k-i) - 2c(i)c(k)) (c(k+j) + c(k-j) - 2c(j)c(k)).$$

(2) ARMA Models:

Next, we will model p, q, ϕ, θ and σ^2 .

① AR(p)-process:

i) Consider zero-mean causal AR(p) model:

$$\phi(\beta) X_t = z_t. \quad z_t \sim N(0, \sigma^2).$$

$$(B_p \text{ identity}): \quad X_t = z_t + \sum_{j=1}^p \gamma_j z_{t-j}$$

product with $X_{t-j}, j = 0, \dots, p$. We have:

$$\left\{ \begin{array}{l} \sigma^2 = y(0) - \phi^T y_p. \quad j=0 \\ I_p \phi = y_p. \quad j>0. \end{array} \right. \quad \text{by orthogonality.}$$

We have Yule-Walker estimator:

$$\left\{ \begin{array}{l} \hat{\sigma}^2 = \hat{y}(0) - \hat{\phi}^T \hat{y}_p \\ \hat{I}_p \hat{\phi} = \hat{y}_p. \end{array} \right. \quad \begin{array}{l} \hat{I}_p, \hat{\phi} \text{ no estimator} \\ \text{by Axta } (x_i). \end{array}$$

Then i) $\hat{\phi}$ and \hat{y} also satisfy:

$$\hat{y}^{(h)} - \hat{\phi}_1 \hat{y}^{(h-1)} - \dots - \hat{\phi}_p \hat{y}^{(p+h)} = \begin{cases} 0, & h=1, \dots, p \\ \hat{\sigma}^2, & h=0 \end{cases}$$

$$\text{ii) } n^{\frac{1}{2}}(\hat{\phi} - \phi) \xrightarrow{D} N(0, \sigma^2 \hat{I}_p^{-1}). \quad \hat{\sigma}^2 \xrightarrow{P} \sigma^2.$$

ii) ρ is commonly unknown.

Consider for m , use $\hat{\phi}_m = (\hat{\phi}_{m1} \dots \hat{\phi}_{mn})$

$= R_m^{-1} \hat{e}_m$. follows from the relation above.

$$\Rightarrow X_t - \hat{\phi}_{m1} X_{t-1} - \dots - \hat{\phi}_{mn} X_{t-n} = z_t \text{ where}$$

$$z_t \sim WN(0, \hat{\gamma}^{(0)}) - \hat{\phi}_m^T \hat{y}_p \stackrel{d}{=} WN(0, v_m)$$

Rmk: We can also use the Durbin-Ljung-Box algorithm to generate $(\hat{\phi}_m)$ if $\hat{\gamma}^{(0)} > 0$.
(It also satisfies the relation equations.)

$$\text{Thm. } n^{\frac{1}{2}}(\hat{\phi}_m - \phi_m) \xrightarrow{d} N(0, \sigma^2 I_m^{-1}).$$

OMA(γ)-process:

consider for m , $\hat{\theta}_m = (\hat{\theta}_{m1} \dots \hat{\theta}_{mn})$.

$$X_t = z_t + \hat{\theta}_{m1} z_{t-1} + \dots + \hat{\theta}_{mn} z_{t-m}, \quad z_t \sim WN(0, \hat{V}_m)$$

We can apply the innovation algorithm to obtain the estimators:

$$\hat{\theta}_{m,m-k} = \hat{V}_k^{-1} (\hat{y}^{(m-k)} - \sum_0^{k-1} \hat{\theta}_{m,m-j} \hat{\theta}_{t+k-j} \hat{v}_j), \quad k=1 \dots m-1$$

$$\hat{V}_0 = \hat{y}^{(0)} - \sum_0^{m-1} \hat{\theta}_{m,m-j} \hat{v}_j.$$

Thm. For causal invertible ARMA(p,q)-process

$$\phi(B)X_t = \theta(B)z_t, \quad z_t \sim IID(0, \sigma^2), \quad E(z_t^q) < \infty.$$

$$\gamma_{zz} = \theta(z)/\phi(z). \quad \text{Then: } A(\text{min})_{n=1..s+1} \text{ s.t.}$$

$\text{min}_n = o(n^{-\frac{1}{3}}), \quad (n \rightarrow \infty).$ we have:

$$n^{\frac{1}{2}}(\hat{\theta}_m - \gamma_1, \dots, \hat{\theta}_m - \gamma_p)^T \xrightarrow{d} N(0, A)$$

$$\hat{V}_m \xrightarrow{P} \sigma^2 I, \quad \text{where } \pi_{ij} = \sum_{k=1}^{i+j} \gamma_{i-k} \gamma_{j-k}.$$

Remark For MA(q)-process, $\hat{\theta}_m$ isn't consistent with θ (only for special (γ_{zz})).

③ ARMA(p,q)

For causal ARMA(p,q)-process:

$$\phi(B)X_t = \theta(B)z_t, \quad z_t \sim WN(0, \sigma^2).$$

$$\Rightarrow X_t = \gamma(B)z_t, \quad \gamma(z) = \theta(z)/\phi(z).$$

$$\text{where } \gamma_0 = 1, \quad \gamma_j = \theta_j + \sum_{i=1}^{j-p} \phi_i \gamma_{j-i}, \quad \theta_j = 0 \quad j > p.$$

By Thm. above. replace γ_i by $\hat{\theta}_{m,i}$. (approx.).

$$\Rightarrow \hat{\theta}_{m,j} = \theta_j + \sum_{i=1}^{j-p} \phi_i \hat{\theta}_{m,j-i}, \quad j = 1, \dots, p+2.$$

From the equation $j = 2+1, \dots, p+2$:

$$\text{we have: } \begin{pmatrix} \hat{\theta}_{m,2+1} \\ \vdots \\ \hat{\theta}_{m,p+2} \end{pmatrix} = (\hat{\theta}_{m,2+i-j})_{p,p} \begin{pmatrix} \phi_1 \\ \vdots \\ \phi_p \end{pmatrix}$$

∴ obtain estimate $\hat{\phi}$. replace into equation $j = 1 \sim q$.

to get estimate $\hat{\theta}$.