

MC Simulations

1) Random number:

Computer doesn't know randomness. So the numbers generated by it randomly will be called pseudorandom numbers.

Due to finite arithmetics of computer for generating $u_k \sim U[0,1]$: we consider $i_k \in \{0, 1, \dots, m-1\}$ and $u_k = i_k/m$.

⇒ Next, we only consider to generate random number $i_k \in \{0, 1, \dots, m-1\}$ uniformly.

Def: A Random number generator (RNG) is $(X, x_0, T, h, \{0, 1, \dots, m-1\})$. X is finite set. $x_0 \in X$ is seed. $T: X \rightarrow X$ is transition func. $h: X \rightarrow \{0, 1, \dots, m-1\}$ is output func. We let $x_1 = T(x_0)$, $i_1 = h(x_1)$ recursively.

Rmk: i) i_L will be periodic. Note \exists

$x_L = x_1$. Then by recursion:

$$x_{L+1} = x_2, \dots \Rightarrow i_{L+1} = i_2. \forall k,$$

ii) Criterion of goodness of Rng:

a) Stat. uniformity: No computationally feasible test can distinguish (i_k) and truly random number

b) Speed. c) Period length

e.g. linear congruential generators)

$X = \{0, 1, \dots, m-1\}$. $i_L = h(x_L) = x_L$. And

$$x_{L+1} \equiv (ax_L + c) \bmod m.$$

Rmk: i) For $c \neq 0$. we generally require

a) $(a, c) = 1$. b) $4/m \Rightarrow 4/n-1$

c) H.p. prime. $p/m \Rightarrow p/n-1$.

ii) Weakness: note for truly

random (I_L, \dots, I_{L+n-1}) will be

uniformly distributed on $\{0, \dots, m-1\}^d$

But R_N (i.e., \dots, i_{d-1}) only falls

in a small hyperplane of $\{0, 1, \dots, m-1\}^d$

(2) Random Variables:

Assume we can produce ϵ_{RK} uniform random numbers by some perfect R_N .

① Inversion method: F^{-1} is quantile of r.v. X . Then: $F^{-1}(u) \sim X$.

Draw: i) sometimes explicit inverse F^{-1} doesn't exist. We can try the numerical inverse.

ii) One goodness of it is that it can transfer the structure

prop.: Let $u^* = \max_{1 \leq i \leq d} \{\epsilon_{Ri}\} \Rightarrow F^{-1}(u^*) \sim \max_{1 \leq i \leq d} \{X_i\}$.

② Acceptance - Rejection method:

Let $f: \mathbb{R}^d \rightarrow \mathbb{R}^{>0}$ is density we can

Sample efficiently. And we want to sample from another density $f: \mathbb{R}^d \rightarrow \mathbb{R}^d$.

Assume: $\exists c \geq 1$. s.t. $f(x) \leq c g(x)$. $\forall x \in \mathbb{R}^d$.

algo.: Given RWH producing $X \stackrel{\text{i.i.d.}}{\sim} f$ and $u \sim U[0,1]$. Let u is indep of X .

- i) Generate one X and one u .
- ii) If $u \leq f(X)/c g(X)$ return X .

else go back to i).

Rmk: Note $P(u \leq f(X)/c g(X)) = \int_{\{g(x)=0\}} f(x) = 0$

Prop. Let Y is outcome of Algo. above. \Rightarrow

$Y \sim f$. And the loop in the algo.
has to be traversed c times averagely.

Pf: $P(u \leq f(x)/c g(x)) =$

$$\iint I_{\{u \leq f(x)/c g(x)\}} f(x) du dx = 1/c$$

Rmk: i) So we expect c is as small as possible.

ii) If $c < 1$. then $P(c \geq 1) > 1$. So it
isn't prob.!

If $c = 1$. then X always = Y.

③ Ziggurate algo.:

It's inefficient if we deal with true r.v. to produce desired r.v. e.g. c is large
So next we construct an variant of accept-rejection method.

The idea is: sample from the density f
 \Rightarrow Sample a pt uniformly from the area
between 0 and graph of f. Besides we
can approx. the area by rectangles.

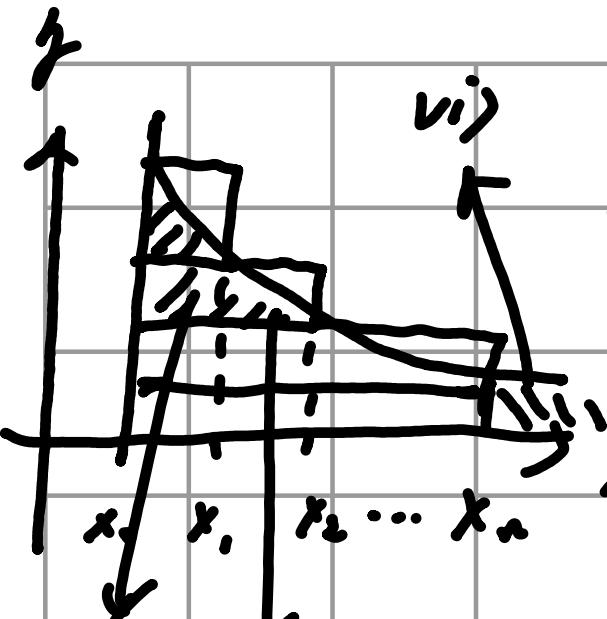
Assume: i) density $f: \mathbb{R}^{>0} \rightarrow \mathbb{R}^{>0}$ ↓ like exp.

ii) $0 = x_0 < x_1 < \dots < x_n$. sc. $y_i = f(x_i)$ and

$$x_i(y_{i+1} - y_i) = x_n y_n + \int_{x_n}^{\infty} f(x) dx =: V.$$

iii) We know list. $x_1 x_2 \dots x_n$.

Algo.: (goal is Sampling $X \sim f$).

- 
 i) generate $i \in \{1, \dots, n\}$ uniformly
 < choose a rectangle randomly >
 ii) $i = n \Rightarrow$ go to iv)
 iii) generate $u_i \sim U[0,1]$. set $x = u_i x_i$
 iv) if $x < x_{i-1}$, return x .
 < fall inside one of graph f >
 v) otherwise generate $u_e \sim U[0,1]$. set
 $y = y_i + u_e(y_{i+1} - y_i)$. if $y < f(x)$, \Rightarrow
 return x . else go back to i)
 vi) generate $u_r \sim U[0,1]$. set $x = V u_r / y_n$
 vii) if $x < x_n$, return x . otherwise return
 a sample from list. $x/x > x_n$.
Remark: Must stop at iv). It's fast.
- Next, we introduce 2 methods to generate
 $(X_1, X_2) \sim N(0, I_2)$:
- ② Box-Muller method:
Algo.: i) generate $u_i \stackrel{i.i.d.}{\sim} U[0,1]$. $i=1, 2$.

i) set $\theta = 22\pi$. $\ell = (-2/\ln u_1)^{\frac{1}{2}}$.

ii) return $(X_1, X_2) = (\ell \cos \theta, \ell \sin \theta)$

$$\text{prob: } \begin{pmatrix} X_1 \\ X_2 \end{pmatrix} = h(u_1, u_2) = \begin{pmatrix} (-2/\ln u_1)^{\frac{1}{2}} \cos 22\pi \\ (-2/\ln u_1)^{\frac{1}{2}} \sin 22\pi \end{pmatrix}. \text{ So}$$

density is $|h'|/h_1 = -e^{-\frac{1}{2}(\ln u_1)^2}/2\pi$.

(5) Polar method:

Alg. i) generate $u_i \stackrel{i.i.d.}{\sim} U(-1, 1)$. $i=1, 2$.

ii) $s = u_1^2 + u_2^2$.

iii) if $s < 1$. return $(Y_1, Y_2) = \begin{pmatrix} u_1 \cdot (-2 \ln s / s)^{\frac{1}{2}} \\ u_2 \cdot (-2 \ln s / s)^{\frac{1}{2}} \end{pmatrix}$
else go back to i).

Rmk: It's more efficient than Box-Muller method since it doesn't need to compute trigonometric functions.

Rmk: To guarantee $N \in \mathcal{N}(\mu, \Sigma)$. If $\Sigma = AA^\top$, then
 $\mu + Ax \sim N(\mu, \Sigma)$.

(2) Monte Carlo:

P.a.s.

β_1 SLLN: $\lim_{n \rightarrow \infty} \mathbb{E}[f(x)] = \bar{E}[f(x)] =: I[f]$

where $I_m(f, x) = \frac{1}{m} \sum_i^m f(x_i)$. $x_i \stackrel{i.i.d.}{\sim} x$.

$$\text{S.E.} = \text{E}_m = I(f, x) - I_m(f, x)$$

Rmk: $\bar{E}(E_m) = 0 \Rightarrow I_m$ is unbiased.

Prop. Let $\sigma = \sigma(f, x) < \infty$. std. var. of $f(x)$.

$$\Rightarrow \bar{E}(E_m(f, x))^2)^{\frac{1}{2}} = \sigma / \sqrt{m}$$

And $I_m | E_m \sim AN(0, \sigma^2)$.

Rmk: \Rightarrow Error E_m is probabilistic.

\Rightarrow MSE has order $\frac{1}{2}$.

\Rightarrow The analysis should assume we know σ^2 .

① Curse of dim:

Compare to the traditional method for

k -dim r.v. $X \sim U[\varepsilon_1, 1]^k$. $I(f, x) = \int_{\varepsilon_1, 1]^k} f(x) dx$

It base on grid $\{\mathbf{x}_1, \dots, \mathbf{x}_N\}^k$. size is N^k .

Given method of order k . error $\sim N^{-k}$.

Bkt we have to evaluate N^k pts. So:

its accuracy is $N^{-k/\lambda}$.

prob. So it's unfeasible if n large.

But MC method has no such problem.

② Variance reduction:

The idea is to find r.v. Y and func. g

$$\text{st. } E[g(Y)] = E[f(x)], \text{Var}(g(Y)) < \text{Var}(f(x)).$$

i) Antithetic variates:

Assume we know a simple transf. $\tilde{x} \sim x$.

$$\text{e.g. } B \sim N(0, 1^2), -B \sim B$$

$$\text{Def: } I_m^A(f, x) = \bar{m} \cdot \tilde{I}_m^{\sim}(f(x_i) + f(\tilde{x}_i))/2.$$

Note computational cost of I_m^A won't exceed compute of I_{2m} . So hopefully we require:

$$\text{Var}(\frac{1}{2}(f(x_i) + f(\tilde{x}_i))) / m < \text{Var}(f(x_i)) / 2m.$$

$$\Leftrightarrow \text{Cov}(f(x), f(\tilde{x})) < 0.$$

ii) Control variates:

Assume we have r.v. Y and func. f . s.t.

$\mathbb{E}(g(Y)) = I(g \cdot Y)$ is known.

$$\Rightarrow \bar{\mathbb{E}}(f(x) - \lambda(g(Y) - I(g \cdot Y))) = I(f \cdot x).$$

So: let $I_m^{\text{est}}(f \cdot x) = m^{-1} \sum_i^m (f(x_i) - \lambda g(Y_i)) + \lambda I(g \cdot Y)$
where $X_i, Y_i \stackrel{\text{i.i.d.}}{\sim} X, Y$.

And I_m^{est} lost at most twice lost of
 I_m . (When $X = Y$ it takes less)

We choose optimal $\lambda^* = \text{Cov}(f(X), g(Y)) / \text{Var}(g(Y))$
to reduce $\text{Var}(I_m^{\text{est}})$

Rmk: i) $\text{Cov}(g(Y), f(X)) \neq 0$. Var - reduced \uparrow .

ii) $\text{Cov}(g(Y), f(X))$ need to be known.

iii) Stratified sampling:

The idea is to partition $X = \bigcup A_k$. his-joint strata. And estimate:

$$\bar{\mathbb{E}}(f(x)) = \sum_1^L \bar{\mathbb{E}}(f(x) | X \in A_k, p_k) \text{ where } p_k = P(X \in A_k).$$

Set $m_L = \#\{x_k \in A_L \mid k \leq m\}$. $z_L = M_L/m$.

$$\Rightarrow I_m^{\text{str}}(f, x) = \sum_1^L p_L \cdot \frac{1}{m_L} \sum_{x_k \in A_L} f(x_k)$$
$$= \frac{1}{m} \sum_1^L \frac{p_L}{z_L} \sum_{x_k \in A_L} f(x_k)$$

Rank: Survey can depend on another r.v.

\geq called stratifying r.v. And let

$$p_L = P(z \in A_L).$$

Next, we want to optimize the following

variables: $z, (A_L), (m_L)$, and also know how to efficiently sample from $(x, z) \mid z \in A_L$.

Protocol: $M_L = \bar{E}(f(x) \mid z \in A_L)$.

$$\hat{\sigma}_L^2 = \text{Var}(f(x) \mid z \in A_L).$$

For m large enough. $z_L \sim p_L$. Assume:

proportional allocation $z_L = p_L$ holds.

$$\text{So: } \text{MSE}(I_m) = \text{Var}(I_m) = m^{-1} \sum_1^L p_L \hat{\sigma}_L^2.$$

$$\text{MSE}(I_m) = \text{Var}(I_m) = m^{-1} \text{Var}(f(x))$$

$$= m^{-1} \left(\sum_1^L p_L \bar{E}(f(x) \mid z \in A_L) - \left(\sum_1^L p_L \mu_L \right)^2 \right)$$

$$= m^{-1} \left(\sum_1^L p_L (\hat{\sigma}_L^2 + \mu_L^2) - \left(\sum_1^L p_L \mu_L \right)^2 \right)$$

$$\Rightarrow \text{MSE}(\bar{I}_m) = \text{MSE}(\bar{I}_m') + \mu^{-1} (\sum p_k \bar{\mu}_k - (\sum p_k \mu_k))^2$$

With Cauchy inequality: $\sum p_k \bar{\mu}_k \geq (\sum p_k \mu_k)^2$.

So stratified sampling can really reduce the variance / MSE.

Point: In general, $\text{Var}(\bar{I}_m) = \mu^{-1} \sum_i \frac{p_i^2}{z_i} \sigma_i^2$.

and the optimizer $z_i^* = p_i \sigma_i / \sum p_k \sigma_k$.

$$\Rightarrow \text{Var}(\bar{I}_m^{st.*}) = \mu^{-1} (\sum_i p_i \sigma_i)^2$$

Point: We should know σ_i here as well. \Rightarrow simulate z_i^* first.

(4) Importance sampling:

Given another r.v. $Y \sim \text{density } q$. Note:

$$I(f, X) = \int f p = \int f \frac{1}{Z} z dx = I(f \frac{1}{Z}, Y)$$

$$\Rightarrow \text{Est. } \tilde{I}_m(f, X) = \mu^{-1} \sum_i f(Y_i) p(Y_i) / Z^c Y_i$$

To choose Z , i.e. reduce $\text{Var}(f(Y_i) \frac{p(Y_i)}{Z^c Y_i})$, take

Z is proportional to $f p$. $\Rightarrow f p / Z$ is flat

so var. ≈ 0 . In fact, we need to know

$I(f) = \int f p$ to normalize Z (most accurate)