

MC Simulations

1) Random number:

Computer doesn't know randomness. So the number generated by it randomly will be called pseudorandom numbers.

Due to finite arithmetics of computer for generating $U_k \sim U[1,1]$: we consider $i_k \in \{0, 1, \dots, m-1\}$ and $U_k = i_k/m$.

\Rightarrow Next, we only consider to generate random number $i_k \in \{0, 1, \dots, m-1\}$ uniformly.

Def: A Random number generator (RNG)

is $(X, x_0, T, h, \{0, 1, \dots, m-1\})$. X is finite set, $x_0 \in X$ is seed. $T: X \rightarrow X$ is transition func. $h: X \rightarrow \{0, 1, \dots, m-1\}$ is output func. We let $x_k = T(x_{k-1})$ $i_k = h(x_k)$ recursively.

Remark: i) i_k will be periodic. Note \exists

$x_k = x_k$. Then by recursion:

$$x_{k+1} = x_{k+1} \dots \Rightarrow i_{k+k} = i_{k+k} \quad \forall k,$$

ii) Criterion of goodness of RNg:

a) Stat. uniformity: No computationally feasible test can distinguish (i_k) and truly random number

b) Speed. c) Period length

e.g. (linear congruential generators)

$X = \{0, 1, \dots, m-1\}$. $i_k = h(x_k) = x_k$. And

$$x_{k+1} \equiv (ax_k + c) \pmod{m}.$$

Remark: i) For $c \neq 0$. we generally require

$$a) (a, c) = 1. \quad b) 4/m \Rightarrow 4/a-1$$

$$c) \forall p, \text{ prime}, p/m \Rightarrow p/a-1.$$

ii) Weakness: Note for truly

random (i_k, \dots, i_{k+n-1}) will be

uniformly distributed on $\{0, \dots, m-1\}^d$

But $KNH (i_1, \dots, i_{d-1})$ only falls
on a small hyperplane of $\{0, 1, \dots, m-1\}^d$

(2) Random Variables:

Assume we can produce (U_i) uniform
random number by some perfect RNH .

① Inversion method: F^{-1} is quantile of
r.v. X . Then: $F^{-1}(u) \sim X$.

Remark: i) Sometimes explicit inverse F^{-1}
doesn't exist. We can try the
numerical inverse.

ii) One goodness of it is that it
can transfer the structure

prop.: let $U^* = \max_{1 \leq i \leq d} \{U_i\} \Rightarrow$

$F^{-1}(U^*) \sim \max_{1 \leq i \leq d} \{X_i\}$.

② Acceptance-Reject method:

Let $f: \mathcal{X} \rightarrow \mathcal{X}^{\geq 0}$ is density we can

sample efficiently. And we want to sample from another density $f: \mathbb{R}^d \rightarrow \mathbb{R}^+$.

Assume: $\exists c \geq 1$. $\forall x$. $f(x) \leq c g(x)$. $\forall x \in \mathbb{R}^d$.

Algo.: Given RWH producing $X \stackrel{i.i.d.}{\sim} g$ and

$U \sim U[0,1]$. Let U is indep of X .

i) Generate one X and one U .

ii) If $U \leq f(X)/cg(X)$ return X .

else go back to i).

Remark: Note $P(g(X)=0) = \int_{\{g(x)=0\}} g(x) = 0$

Prop. Let Y is outcome of Algo. above. \Rightarrow

$Y \sim f$. And the loop in the algo.

has to be traversed c times average.

Pf: $P(U \leq f(X)/cg(X)) =$

$$\int \int I\{u \leq f(x)/cg(x)\} g(x) dx dX = 1/c$$

Remark: i) So we expect c is as small as possible.

ii) If $c < 1$. then $P(\Rightarrow) > 1$. So it isn't probn.!

If $c = 1$. then $X \text{ always} = Y$.

③ Zigzagging algo.:

It's inefficient if we feed with true r.v. to produce desired r.v. e.g., c is large
So next we construct an variant of acceptance-rejection method.

The idea is: sample from the density f
(\Rightarrow) sample a pt uniformly from the area between 0 and graph of f . Besides we can approxi. the area by rectangles.

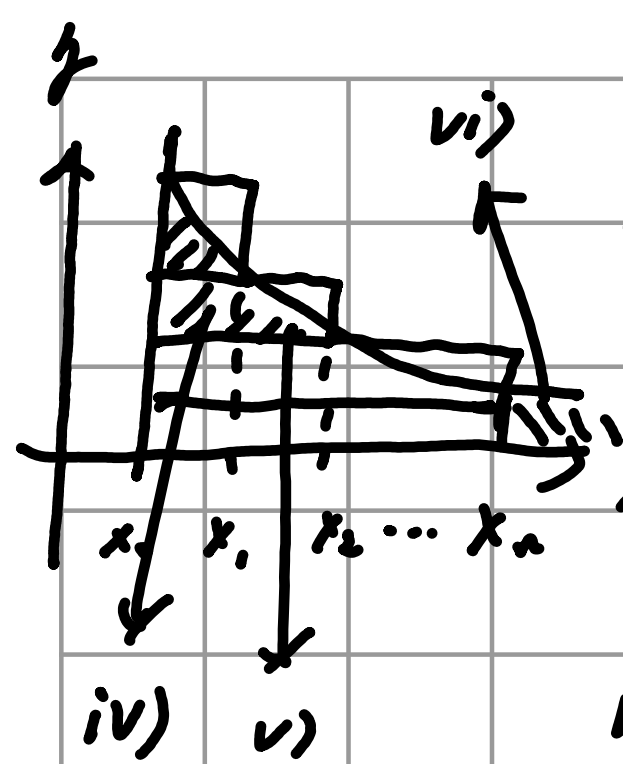
Assume: i) density $f: \mathbb{R}^{\geq 0} \rightarrow \mathbb{R}^{\geq 0} \searrow$ like exp.

ii) $0 = x_0 < x_1 < \dots < x_n$. s.t. $g_i = f(x_i)$ and

$$x_i < g_{i+1} - g_i = x_n g_n + \int_{x_n}^{\infty} f(x) dx =: V.$$

iii) We know dist. $X | X > x_n$.

Algo. (goal is sampling $X \sim f$).



i) generate $i \in \{1, \dots, n\}$ uniformly

ii) choose a rectangle randomly

iii) $i = n \Rightarrow$ go to iv)

iv) generate $u_1 \sim U[0, 1]$. Set $X = u_1 x_i$

v) if $X < x_{i-1}$, return X .

vi) fall inside area of graph f

vii) otherwise generate $u_2 \sim U[0, 1]$. Set

$y = y_i + u_2 (y_{i+1} - y_i)$. if $y < f(x)$, \Rightarrow

return X . else go back to i)

viii) generate $u_1 \sim U[0, 1]$. Set $X = V u_1 / y_n$

ix) if $X < x_n$, return X . otherwise return

a sample from dist. $X | X > x_n$.

Remark: Must stop at iv). It's fast.

Next, we introduce 2 methods to generate

$(X_1, X_2) \sim N(0, I_2)$:

④ Box-muller method:

Algo.: i) generate $u_i \sim U[0, 1]$. $i = 1, 2$.

ii) Set $\theta = 2\pi u_2$. $r = (-2 \log u_1)^{\frac{1}{2}}$.

iii) Return $(X_1, X_2) = (r \cos \theta, r \sin \theta)$

Proof: $\begin{pmatrix} X_1 \\ X_2 \end{pmatrix} = h(u_1, u_2) = \begin{pmatrix} (-2 \log u_1)^{\frac{1}{2}} \cos 2\pi u_2 \\ (-2 \log u_1)^{\frac{1}{2}} \sin 2\pi u_2 \end{pmatrix} \cdot J_0$

Density is $|Jh'|/|Jh| = -e^{-\frac{1}{2}(x_1^2 + x_2^2)} / 2\pi$.

⑤ Polar method:

Alg. i) generate $u_i \sim U(-1, 1)$. $i = 1, 2$.

ii) $S = u_1^2 + u_2^2$.

iii) if $S < 1$. return $\begin{pmatrix} Y_1 \\ Y_2 \end{pmatrix} = \begin{pmatrix} u_1 (-2 \ln S / S)^{\frac{1}{2}} \\ u_2 (-2 \ln S / S)^{\frac{1}{2}} \end{pmatrix}$
else go back to i).

Proof: It's more efficient than Box-Muller method since it doesn't need to compute trigo. func.

Proof: To generate $N(\mu, \Sigma)$. let $\Sigma = AA^T$, then
 $\mu + AX \sim N(\mu, \Sigma)$.

(3) Monte Carlo:

By SLLN: $\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n f(x_i) = E(f(x)) =: I(f, x)$

Where $I_m(f, x) = \frac{1}{m} \sum_{i=1}^m f(x_i)$. $x_i \stackrel{i.i.d}{\sim} x$.

$$f_{\text{est}} : \varepsilon_m = I(f, x) - I_m(f, x)$$

Prop.: $\mathbb{E}(\varepsilon_m) = 0 \Rightarrow I_m$ is unbiased.

Prop.: Let $\sigma = \sigma(f, x) < \infty$. Std. var. of $f(x)$.

$$\Rightarrow \mathbb{E}(\varepsilon_m(f, x)^2)^{\frac{1}{2}} = \sigma / \sqrt{m}.$$

And $I_m \varepsilon_m \sim AN(0, \sigma^2)$.

Prop.: \Rightarrow Error ε_m is probabilistic.

i) MSE per row $1/2$.

ii) The analysis should assume we know σ^2 .

① Curse of Dim:

Compare to the traditional method for

k-dim r.v. $X \sim \mathcal{U}([0, 1]^k)$. $I(f, x) = \int_{[0, 1]^k} f(x) dx$

is based on grid $\{x_1, \dots, x_n\}^k$. Size is n^k .

Given mesh of order k , error $\sim n^{-k}$.

But we have to evaluate n^k pts. So:

its accuracy is $n^{-k/k}$.

proof So it's unfeasible if n large.

But MC method has no such problem.

② Variance Reduction:

The idea is to find r.v. Y and func. g

$$\text{s.t. } \mathbb{E}(g(Y)) = \mathbb{E}(f(x)), \quad \text{Var}(g(Y)) < \text{Var}(f(x)).$$

i) Antithetic Variates:

Assume we know a simple transf. $\tilde{X} \sim X$.

e.g. $B \sim N(0, \sigma^2), -B \sim B$

Def: $I_m^A(f, X) = m^{-1} \sum_{i=1}^m (f(X_i) + f(\tilde{X}_i)) / 2$.

Note computational cost of I_m^A won't exceed compute of I_{2m} . So hopefully

we require:

$$\text{Var}(\frac{1}{2}(f(X_i) + f(\tilde{X}_i))) / m < \text{Var}(f(X_i)) / 2m.$$

$$\Leftrightarrow \text{Cov}(f(X), f(\tilde{X})) < 0.$$

ii) Control Variates:

Assume we have r.v. Y and func. g . s.t.

$E(g(Y)) = I(g, Y)$ is known.

$$\Rightarrow E(f(x) - \lambda(g(Y) - I(g, Y))) = I(f, x).$$

So: let $I_n^{c, \lambda}(f, x) = n^{-1} \sum_{i=1}^n (f(x_i) - \lambda(g(Y_i) - I(g, Y))) + \lambda I(g, Y)$
where $X_i, Y_i \stackrel{i.i.d.}{\sim} X, Y$.

And $I_n^{c, \lambda}$ cost at most twice cost of I_n . (When $X=Y$ it takes less)

We choose optimal $\lambda^* = \text{Cov}(f(x), g(Y)) / \text{Var}(g(Y))$
to reduce $\text{Var}(I_n^{c, \lambda})$

Remark: i) $\text{Cov}(g(Y), f(x)) \uparrow$. Var - reduced \uparrow .

ii) $\text{Cov}(g(Y), f(x))$ need to be known.

iii) Stratified sampling:

The idea is to partition $X = \bigcup_{l=1}^L A_l$. Dis-joint strata. And estimate:

$$E(f(x)) = \sum_{l=1}^L E(f(x) | X \in A_l) p_l. \text{ where } p_l = P(X \in A_l).$$

Let $m_c = \#\{x_k \in A_c \mid k \leq m\}$. $z_c = m_c/m$.

$$\begin{aligned} \Rightarrow I_m^{ST}(f, x) &= \sum_{c=1}^L p_c \cdot \frac{1}{m_c} \sum_{x_k \in A_c} f(x_k) \\ &= \frac{1}{m} \sum_{c=1}^L \frac{p_c}{z_c} \sum_{x_k \in A_c} f(x_k) \end{aligned}$$

Remark: Strata can depend on another r.v.

z called stratifying r.v. And let

$$p_c = \mathbb{P}(z \in A_c).$$

Next, we want to optimize the following variables: $z, (A_c)_c, (m_c)_c$ and also know how to efficiently sample from $(x, z) \mid z \in A_c$.

Define: $\mu_c = \mathbb{E}(f(x) \mid z \in A_c)$.

$$\sigma_c^2 = \text{Var}(f(x) \mid z \in A_c).$$

For m large enough, $z_c \sim p_c$. Assume:

proportional allocation $z_c = p_c$ holds.

$$\text{So: } m \text{SE}(I_m^{ST}) = \text{Var}(I_m^{ST}) = m^{-1} \sum_{c=1}^L p_c \sigma_c^2.$$

$$m \text{SE}(I_m) = \text{Var}(I_m) = m^{-1} \text{Var}(f(x))$$

$$= m^{-1} \left(\sum_{c=1}^L p_c \mathbb{E}(f(x)^2 \mid z \in A_c) - \left(\sum_{c=1}^L p_c \mu_c \right)^2 \right)$$

$$= m^{-1} \left(\sum_{c=1}^L p_c (\sigma_c^2 + \mu_c^2) - \left(\sum_{c=1}^L p_c \mu_c \right)^2 \right)$$

$$\Rightarrow \text{MSE}(I_m) = \text{MSE}(I_m^{st}) + \mu' (\sum p_k \mu_k^2 - (\sum p_k \mu_k)^2)$$

With Cauchy inequality: $\sum p_k \mu_k^2 \geq (\sum p_k \mu_k)^2$.

So stratified sampling can really reduce the variance / MSE.

Proof: In general, $\text{Var}(I_m^{st}) = \mu^{-1} \sum_k \frac{p_k^2}{z_k} \sigma_k^2$.

and the optimizer $z_k^* = p_k \sigma_k / \sum p_k \sigma_k$.

$$\Rightarrow \text{Var}(I_m^{st,*}) = \mu^{-1} (\sum_k p_k \sigma_k)^2$$

Proof: We should know σ_k here as

well. \Rightarrow minimize z_k^* first.

④ Importance sampling:

Given another r.v. $Y \sim$ density q . Note:

$$I(f, X) = \int f p = \int f \frac{p}{z} z dx = I(f \frac{p}{z}, Y)$$

$$\Rightarrow \text{Let } \tilde{I}_m(f, X) = \mu^{-1} \sum_i \tilde{f}(Y_i) p(Y_i) / z(Y_i)$$

To choose z , i.e. reduce $\text{Var}(f(Y) \frac{p(Y)}{z(Y)})$, take

z is proportional to $f p$. $\Rightarrow f p / z$ is flat

So $\text{Var.} \approx 0$. (In fact, we need to know

$I(f) = \int f p$ to normalize q (most common))