

# Strong Converses for High-dimensional Statistical Estimation

Ramji Venkataramanan  
University of Cambridge

Oliver Johnson  
University of Bristol

**Beyond IID 2018**

# Inference set-up

Want to estimate  $\theta \in \mathcal{F}$  from data  $\mathbf{Y} = (Y_1, \dots, Y_n)$

Data  $\mathbf{Y}$  generated according to  $P_\theta(\mathbf{Y})$

How well can we estimate  $\theta$  as the number of samples grows?

# Inference set-up

Want to estimate  $\theta \in \mathcal{F}$  from data  $\mathbf{Y} = (Y_1, \dots, Y_n)$

Data  $\mathbf{Y}$  generated according to  $P_\theta(\mathbf{Y})$

How well can we estimate  $\theta$  as the number of samples grows?

## Density Estimation [Yu '97]

$\mathcal{F}$ : smooth densities on  $[0, 1]$  with bounded second derivative

For  $\theta \in \mathcal{F}$ , samples  $Y_1, \dots, Y_n$  drawn i.i.d.  $\sim \theta$

Measure of performance:

$$d(\theta, \hat{\theta}) = \int_0^1 \left( \sqrt{\theta(x)} - \sqrt{\hat{\theta}(x)} \right)^2 dx$$

# Compressed sensing

$$\begin{bmatrix} \vdots \\ y \end{bmatrix} = \begin{bmatrix} \vdots \\ \theta \end{bmatrix} + \begin{bmatrix} \vdots \\ w \end{bmatrix}$$

Vector  $\theta \in \mathcal{F}$  observed through linear model:

$$\mathbf{y} = \mathbf{A} \theta + \text{noise}$$

$\mathcal{F}$ : unit norm vectors in  $\mathbb{R}^n$  with at most  $k$  non-zeros

How well can we estimate  $\theta$ ?

Measure of performance:

$$M^*(\mathbf{A}) := \inf_{\hat{\theta}} \sup_{\theta \in \mathcal{F}} \mathbb{E} \left[ \frac{1}{n} \|\hat{\theta}(\mathbf{y}) - \theta\|^2 \right],$$

# Loss function and Risk

Want to estimate  $\theta \in \mathcal{F}$  from data  $\mathbf{Y} = (Y_1, \dots, Y_n)$

Data  $\mathbf{Y}$  generated according to  $P_\theta(\mathbf{Y})$

Performance of an estimator  $\hat{\theta}$  measured via  $d(\theta, \hat{\theta}(\mathbf{Y}))$

Loss function  $d$  is a distance or semi-distance

Risk  $R(\theta, \hat{\theta}) = \mathbb{E} \left[ d(\theta, \hat{\theta}) \right]$

GOAL: Lower bounds on the *minimax risk*

$$\inf_{\hat{\theta}} \sup_{\theta \in \mathcal{F}} \mathbb{E} \left[ d(\theta, \hat{\theta}) \right]$$

## Standard approach (see Tsybakov 2009)

For any  $\psi_n > 0$ ,

$$\mathbb{P}\left(d(\theta, \hat{\theta}) \geq \psi_n\right) \leq \frac{1}{\psi_n} \mathbb{E}\left[d(\theta, \hat{\theta})\right]$$

## Standard approach (see Tsybakov 2009)

For any  $\psi_n > 0$ ,

$$\mathbb{P} \left( d(\theta, \hat{\theta}) \geq \psi_n \right) \leq \frac{1}{\psi_n} \mathbb{E} \left[ d(\theta, \hat{\theta}) \right]$$

Hence

$$\sup_{\theta \in \mathcal{F}} \mathbb{E} \left[ d(\theta, \hat{\theta}) \right] \geq \psi_n \sup_{\theta \in \mathcal{F}} \mathbb{P} \left( d(\theta, \hat{\theta}) \geq \psi_n \right)$$

Want to choose  $\psi_n$  such that prob. is bounded below by a constant

# Packing set

$$\sup_{\theta \in \mathcal{F}} \mathbb{E} \left[ d(\theta, \hat{\theta}) \right] \geq \psi_n \sup_{\theta \in \mathcal{F}} \mathbb{P} \left( d(\theta, \hat{\theta}) \geq \psi_n \right)$$

Construct a **packing set**  $\{\theta_1, \dots, \theta_M\} \subseteq \mathcal{F}$  such that

$$d(\theta_i, \theta_j) \geq d_{\min} = 2\psi_n, \quad \text{for all } i \neq j$$

- Existence of packing set can be generally guaranteed via Gilbert-Varshamov bound or the probabilistic method



# Packing set

$$\sup_{\theta \in \mathcal{F}} \mathbb{E} \left[ d(\theta, \hat{\theta}) \right] \geq \psi_n \sup_{\theta \in \mathcal{F}} \mathbb{P} \left( d(\theta, \hat{\theta}) \geq \psi_n \right)$$

Construct a **packing set**  $\{\theta_1, \dots, \theta_M\} \subseteq \mathcal{F}$  such that

$$d(\theta_i, \theta_j) \geq d_{\min} = 2\psi_n, \quad \text{for all } i \neq j$$

- ▶ Existence of packing set can be generally guaranteed via Gilbert-Varshamov bound or the probabilistic method
- ▶ **Idea:** Any estimator  $\hat{\theta}$  defines an  $M$ -ary hypothesis test between  $\{\theta_1, \dots, \theta_M\}$

$$\hat{i} = \arg \min_{1 \leq j \leq M} d(\theta_j, \hat{\theta})$$

# Channel coding interpretation

- ▶ Channel  $P_{\mathbf{Y}|\theta} := P_{\theta}(\mathbf{Y})$
- ▶ Codebook  $\{\theta_1, \dots, \theta_M\}$
- ▶ 'Transmitted' codeword  $\theta = \theta_i$
- ▶ Channel output  $\mathbf{Y} = (Y_1, \dots, Y_n)$
- ▶ Minimum-distance decoder Distance measured via  $d(\cdot, \cdot)$   
Decode codeword that is closest to  $\hat{\theta}(\mathbf{Y})$ .

# Probability of decoding error

Minimum distance between codewords is  $d_{\min} = 2\psi_n \Rightarrow$

Decoder makes error only if  $d(\theta_i, \hat{\theta}) \geq \frac{d_{\min}}{2} = \psi_n \Rightarrow$

$$\mathbb{P}(\hat{i} \neq i \mid \theta_i \text{ true codeword}) \leq \mathbb{P}(d(\theta_i, \hat{\theta}) \geq \psi_n)$$

# Probability of decoding error

Minimum distance between codewords is  $d_{\min} = 2\psi_n \Rightarrow$

Decoder makes error only if  $d(\theta_i, \hat{\theta}) \geq \frac{d_{\min}}{2} = \psi_n \Rightarrow$

$$\mathbb{P}(\hat{i} \neq i \mid \theta_i \text{ true codeword}) \leq \mathbb{P}(d(\theta_i, \hat{\theta}) \geq \psi_n)$$

Therefore

$$\varepsilon_M := \frac{1}{M} \sum_{i=1}^M \mathbb{P}(\hat{i} \neq i \mid \theta_i \text{ true codeword}) \leq \sup_{\theta \in \mathcal{F}} \mathbb{P}(d(\theta, \hat{\theta}) \geq \psi_n)$$

## Probability of decoding error

Minimum distance between codewords is  $d_{\min} = 2\psi_n \Rightarrow$

Decoder makes error only if  $d(\theta_i, \hat{\theta}) \geq \frac{d_{\min}}{2} = \psi_n \Rightarrow$

$$\mathbb{P}(\hat{i} \neq i \mid \theta_i \text{ true codeword}) \leq \mathbb{P}(d(\theta_i, \hat{\theta}) \geq \psi_n)$$

Therefore

$$\varepsilon_M := \frac{1}{M} \sum_{i=1}^M \mathbb{P}(\hat{i} \neq i \mid \theta_i \text{ true codeword}) \leq \sup_{\theta \in \mathcal{F}} \mathbb{P}(d(\theta, \hat{\theta}) \geq \psi_n)$$

Plugging into our risk lower bound,

$$\sup_{\theta \in \mathcal{F}} \mathbb{E}[d(\theta, \hat{\theta})] \geq \psi_n \sup_{\theta \in \mathcal{F}} \mathbb{P}(d(\theta, \hat{\theta}) \geq \psi_n) \geq \psi_n \varepsilon_M$$

## Risk Lower Bound

$$\sup_{\theta \in \mathcal{F}} \mathbb{E} \left[ d(\theta, \hat{\theta}) \right] \geq \psi_n \varepsilon_M$$

$\varepsilon_M$  is average error probability of codebook with  $d_{\min} = 2\psi_n$

## Risk Lower Bound

$$\sup_{\theta \in \mathcal{F}} \mathbb{E} \left[ d(\theta, \hat{\theta}) \right] \geq \psi_n \varepsilon_M$$

$\varepsilon_M$  is average error probability of codebook with  $d_{\min} = 2\psi_n$

Fano's inequality is a standard way to lower bound  $\varepsilon_M$ :

$$\varepsilon_M \geq 1 - \frac{\log 2 + \frac{1}{M} \sum_{i=1}^M D(P_{\mathbf{Y}|\theta_i} \| \bar{P}_{\mathbf{Y}})}{\log M}, \quad \text{where } \bar{P}_{\mathbf{Y}} = \frac{1}{M} \sum_{i=1}^M P_{\mathbf{Y}|\theta_i}$$

If we show that  $\frac{1}{M} \sum_{i=1}^M D(P_{\mathbf{Y}|\theta_i} \| \bar{P}_{\mathbf{Y}}) \leq \alpha \log M$ , then

$$\varepsilon_M \geq 1 - \alpha - \frac{1}{\log M} > 0,$$

---

Ibragimov and Khasminskii, *Estimation of infinite dimensional parameter in Gaussian white noise*, 1977

# Improving on Fano

Generalized versions of Fano: [Birgé '05], [Sason-Verdú '18]

Other converse techniques:

Sphere-packing bound: [Shannon-Gallager-Berlekamp '67]

Based on information spectrum: [Wolfowitz '68], [Verdú-Han '94]

Based on general  $f$ -divergences: [Guntuboyina '11]

⋮

Based on binary hypothesis testing:

[Hayashi, Nagaoka '03]

[Polyanskiy, Poor, Verdú '10] (“Meta-converse”)

[Vazquez-Vilar, Tauste Campo, Guillén i Fàbregas, Martinez '16]



# Improving on Fano

Generalized versions of Fano: [Birgé '05], [Sason-Verdú '18]

Other converse techniques:

Sphere-packing bound: [Shannon-Gallager-Berlekamp '67]

Based on information spectrum: [Wolfowitz '68], [Verdú-Han '94]

Based on general  $f$ -divergences: [Guntuboyina '11]

⋮

Based on binary hypothesis testing:

[Hayashi, Nagaoka '03]

[Polyanskiy, Poor, Verdú '10] (“Meta-converse”)

[Vazquez-Vilar, Tauste Campo, Guillén i Fàbregas, Martinez '16]

$$\sup_{\theta \in \mathcal{F}} \mathbb{E} \left[ d(\theta, \hat{\theta}) \right] \geq \psi_n \varepsilon_M$$

$$\varepsilon_M = \frac{1}{M} \sum_{i=1}^M \mathbb{P} \left( \hat{i} \neq i \mid \theta_i \text{ true codeword} \right)$$

What we want is a lower bound for  $\varepsilon_M$  that:

- is computable for wide range of statistical problems
- with *existing* packing sets
- shows  $\varepsilon_M \rightarrow 1$  as  $M$  grows (strong converse)

## Obtaining a tighter lower bound

- Channel  $P_{\mathbf{Y}|\theta}$
- Codebook  $\{\theta_1, \dots, \theta_M\}$  (equally likely codewords)
- Average error probability  $\varepsilon_M$

A channel decoder defines a hypothesis test to distinguish between:

$$H_0 : (\theta, \mathbf{Y}) \sim Q = P_\theta Q_{\mathbf{Y}}$$

$$H_1 : (\theta, \mathbf{Y}) \sim P = P_\theta P_{\mathbf{Y}|\theta}$$

Does the data look like it came from the true generating model ?

## Obtaining a tighter lower bound

- Channel  $P_{\mathbf{Y}|\theta}$
- Codebook  $\{\theta_1, \dots, \theta_M\}$  (equally likely codewords)
- Average error probability  $\varepsilon_M$

A channel decoder defines a hypothesis test to distinguish between:

$$H_0 : (\theta, \mathbf{Y}) \sim Q = P_\theta Q_{\mathbf{Y}}$$

$$H_1 : (\theta, \mathbf{Y}) \sim P = P_\theta P_{\mathbf{Y}|\theta}$$

Does the data look like it came from the true generating model ?

For the channel decoder based test [Polyanskiy, Poor, Verdú '10]:

$$Q[T = 1] = \frac{1}{M}, \quad P[T = 0] = \varepsilon_M$$

# Obtaining a tighter lower bound

- Channel  $P_{\mathbf{Y}|\theta}$
- Codebook  $\{\theta_1, \dots, \theta_M\}$  (equally likely codewords)
- Average error probability  $\varepsilon_M$

A channel decoder defines a hypothesis test to distinguish between:

$$H_0 : (\theta, \mathbf{Y}) \sim Q = P_\theta Q_{\mathbf{Y}}$$

$$H_1 : (\theta, \mathbf{Y}) \sim P = P_\theta P_{\mathbf{Y}|\theta}$$

Does the data look like it came from the true generating model ?

For the channel decoder based test [Polyanskiy, Poor, Verdú '10]:

$$Q[T = 1] = \frac{1}{M}, \quad P[T = 0] = \varepsilon_M$$

For any randomized hypothesis test  $T$  and  $\gamma > 0$ , we have

$$P[T = 1] - \gamma Q[T = 1] \leq P \left[ \frac{dP}{dQ} > \gamma \right].$$

Hence, in our case, for any  $\gamma > 0$

$$\frac{1}{M} \geq \frac{1}{\gamma} \left( 1 - \varepsilon_M - P_{\theta \mathbf{Y}} \left[ \frac{dP_{\mathbf{Y}|\theta}}{dQ_{\mathbf{Y}}} > \gamma \right] \right)$$

- ▶ Can bound  $P_{\theta \mathbf{Y}} \left[ \frac{dP_{\mathbf{Y}|\theta}}{dQ_{\mathbf{Y}}} > \gamma \right]$  in terms of Rényi divergences using Markov inequality type argument
- ▶ Can optimize over  $\gamma$  to deduce ...

## Theorem

For any  $\lambda > 0$ , and any distribution  $Q_{\mathbf{Y}}$  over  $\mathcal{Y}$  (satisfying mild absolute continuity condition),

$$\varepsilon_M \geq 1 - \frac{(1 + \lambda)}{(\lambda M)^{\frac{\lambda}{1+\lambda}}} \left[ \sum_{i=1}^M \frac{1}{M} \exp \left( \lambda D_{1+\lambda}(P_{\mathbf{Y}|\theta_i} \| Q_{\mathbf{Y}}) \right) \right]^{\frac{1}{1+\lambda}}.$$

Here  $D_{1+\lambda}(P_{\mathbf{Y}|\theta_i} \| Q_{\mathbf{Y}})$  is the Rényi divergence of order  $(1 + \lambda)$ :

$$D_{1+\lambda}(P_{\mathbf{Y}|\theta_i} \| Q_{\mathbf{Y}}) := \frac{1}{\lambda} \log \left( \int_{\mathcal{Y}} \left( \frac{dP_{\mathbf{Y}|\theta_i}}{dQ_{\mathbf{Y}}} \right)^{1+\lambda} dQ_{\mathbf{Y}} \right).$$

## Theorem

For any  $\lambda > 0$ , and any distribution  $Q_Y$  over  $\mathcal{Y}$  (satisfying mild absolute continuity condition),

$$\varepsilon_M \geq 1 - \frac{(1 + \lambda)}{(\lambda M)^{\frac{\lambda}{1+\lambda}}} \left[ \sum_{i=1}^M \frac{1}{M} \exp \left( \lambda D_{1+\lambda}(P_{Y|\theta_i} \| Q_Y) \right) \right]^{\frac{1}{1+\lambda}}.$$

Here  $D_{1+\lambda}(P_{Y|\theta_i} \| Q_Y)$  is the Rényi divergence of order  $(1 + \lambda)$ :

$$D_{1+\lambda}(P_{Y|\theta_i} \| Q_Y) := \frac{1}{\lambda} \log \left( \int_{\mathcal{Y}} \left( \frac{dP_{Y|\theta_i}}{dQ_Y} \right)^{1+\lambda} dQ_Y \right).$$

- ▶ Pick a good  $Q_Y$ , compute lower bound for  $\varepsilon_M$  via upper bound for Rényi divergence, e.g., [Sason-Verdú '16]
- ▶ Have free choice of  $\lambda$ , often  $\lambda = 1$  works well enough



# Improved risk lower bounds

$$\sup_{\theta \in \mathcal{F}} \mathbb{E} \left[ d(\theta, \hat{\theta}) \right] \geq \psi_n \varepsilon_M$$

In paper we use the result to study three illustrative examples:

1. Compressed sensing
2. Density estimation problem
3. Active learning of a binary classifier. . . see paper.

In each case, get improved bounds with  $\varepsilon_M \rightarrow 1$  (strong converse), essentially for free.

# Application: Compressed Sensing

$$\begin{matrix} & \xrightarrow{n} \\ \begin{matrix} \uparrow m \\ \downarrow \end{matrix} & \left[ \begin{array}{c} \mathbf{A} \end{array} \right] & \left[ \begin{array}{c} \theta \end{array} \right] + \left[ \begin{array}{c} \mathbf{w} \end{array} \right] = \left[ \begin{array}{c} \mathbf{y} \end{array} \right] \end{matrix}$$

$$\mathbf{y} = \mathbf{A}\theta + \mathbf{w}, \quad \mathbf{w} \sim \mathcal{N}(0, \sigma^2 \mathbf{I})$$

$\mathcal{F}_k$ : unit norm vectors  $\theta$  in  $\mathbb{R}^n$  with at most  $k$  non-zeros

Want to lower bound

$$M^*(\mathbf{A}) := \inf_{\hat{\theta}} \sup_{\theta \in \mathcal{F}_k} \mathbb{E} \left[ \frac{1}{n} \|\hat{\theta}(\mathbf{y}) - \theta\|^2 \right],$$

# Packing set (see [Candès-Davenport '13])

Packing set of vectors  $\{\theta_1, \dots, \theta_M\} \in \mathbb{R}^n$  with:

- ▶  $\|\theta_i\|^2 = 1$  for all  $i$
- ▶  $\|\theta_i - \theta_j\|^2 \geq \frac{1}{2}$  for  $i \neq j$
- ▶  $\|\frac{1}{M} \sum_{i=1}^M \theta_i \theta_i^T - \frac{1}{n} \mathbf{I}\|_{\text{op}} \leq \frac{\beta}{n}$  for some small  $\beta > 0$

# Packing set (see [Candès-Davenport '13])

Packing set of vectors  $\{\theta_1, \dots, \theta_M\} \in \mathbb{R}^n$  with:

- ▶  $\|\theta_i\|^2 = 1$  for all  $i$
- ▶  $\|\theta_i - \theta_j\|^2 \geq \frac{1}{2}$  for  $i \neq j$
- ▶  $\|\frac{1}{M} \sum_{i=1}^M \theta_i \theta_i^T - \frac{1}{n} \mathbf{I}\|_{\text{op}} \leq \frac{\beta}{n}$  for some small  $\beta > 0$
- ▶ Size of packing set  $M = \left(\frac{n}{k}\right)^{k/4} = \exp\left(\frac{k}{4} \log\left(\frac{n}{k}\right)\right)$

# Packing set (see [Candès-Davenport '13])

Packing set of vectors  $\{\theta_1, \dots, \theta_M\} \in \mathbb{R}^n$  with:

- ▶  $\|\theta_i\|^2 = 1$  for all  $i$
- ▶  $\|\theta_i - \theta_j\|^2 \geq \frac{1}{2}$  for  $i \neq j$
- ▶  $\|\frac{1}{M} \sum_{i=1}^M \theta_i \theta_i^T - \frac{1}{n} \mathbf{I}\|_{\text{op}} \leq \frac{\beta}{n}$  for some small  $\beta > 0$
- ▶ Size of packing set  $M = \left(\frac{n}{k}\right)^{k/4} = \exp\left(\frac{k}{4} \log\left(\frac{n}{k}\right)\right)$

# Computing the Renyi Divergence

$$\varepsilon_M \geq 1 - \frac{(1 + \lambda)}{(\lambda M)^{\frac{\lambda}{1+\lambda}}} \left[ \sum_{i=1}^M \frac{1}{M} \exp \left( \lambda D_{1+\lambda}(P_{\mathbf{Y}|\theta_i} \| Q_{\mathbf{Y}}) \right) \right]^{\frac{1}{1+\lambda}}$$

Since  $\mathbf{y} = \mathbf{A}\theta + \mathbf{w}$ ,  $P_{\mathbf{Y}|\theta_i} \sim \mathcal{N}(\mathbf{A}\theta_i, \sigma^2\mathbf{I})$

Choose  $Q_{\mathbf{Y}} \sim \mathcal{N}(\mathbf{0}, \sigma^2\mathbf{I})$

# Computing the Renyi Divergence

$$\varepsilon_M \geq 1 - \frac{(1+\lambda)}{(\lambda M)^{\frac{\lambda}{1+\lambda}}} \left[ \sum_{i=1}^M \frac{1}{M} \exp \left( \lambda D_{1+\lambda}(P_{\mathbf{Y}|\theta_i} \| Q_{\mathbf{Y}}) \right) \right]^{\frac{1}{1+\lambda}}$$

Since  $\mathbf{y} = \mathbf{A}\theta + \mathbf{w}$ ,  $P_{\mathbf{Y}|\theta_i} \sim \mathcal{N}(\mathbf{A}\theta_i, \sigma^2\mathbf{I})$

Choose  $Q_{\mathbf{Y}} \sim \mathcal{N}(\mathbf{0}, \sigma^2\mathbf{I})$

Then

$$D_{1+\lambda}(P_{\mathbf{Y}|\theta_i} \| Q_{\mathbf{Y}}) = \frac{(1+\lambda)}{2\sigma^2} \|\mathbf{A}\theta_i\|^2$$

# Computing the Renyi Divergence

$$\varepsilon_M \geq 1 - \frac{(1 + \lambda)}{(\lambda M)^{\frac{\lambda}{1+\lambda}}} \left[ \sum_{i=1}^M \frac{1}{M} \exp \left( \lambda D_{1+\lambda}(P_{\mathbf{Y}|\theta_i} \| Q_{\mathbf{Y}}) \right) \right]^{\frac{1}{1+\lambda}}$$

Since  $\mathbf{y} = \mathbf{A}\theta + \mathbf{w}$ ,  $P_{\mathbf{Y}|\theta_i} \sim \mathcal{N}(\mathbf{A}\theta_i, \sigma^2 \mathbf{I})$

Choose  $Q_{\mathbf{Y}} \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{I})$

Then

$$D_{1+\lambda}(P_{\mathbf{Y}|\theta_i} \| Q_{\mathbf{Y}}) = \frac{(1 + \lambda)}{2\sigma^2} \|\mathbf{A}\theta_i\|^2$$

We use a subset  $\mathcal{P}$  of the Candès-Davenport packing set with  $M' = \frac{M}{\log M}$  elements such that

$$\max_{\theta_i \in \mathcal{P}} \|\mathbf{A}\theta_i\|^2 \leq \frac{\|\mathbf{A}\|_F^2}{n} (1 + \delta) \quad \text{for some small } \delta > 0$$



### Proposition:

For any  $\lambda > 0$ ,  $\Delta \in (0, 1)$ , and  $M = (n/k)^{k/4}$ , we have

$$\varepsilon_M \geq 1 - (1 + \lambda) \left( \frac{(\log M) M^{-\Delta}}{\lambda} \right)^{\lambda/(1+\lambda)},$$

### Proposition:

For any  $\lambda > 0$ ,  $\Delta \in (0, 1)$ , and  $M = (n/k)^{k/4}$ , we have

$$\varepsilon_M \geq 1 - (1 + \lambda) \left( \frac{(\log M) M^{-\Delta}}{\lambda} \right)^{\lambda/(1+\lambda)},$$

and

$$\begin{aligned} M^*(\mathbf{A}) &= \inf_{\hat{\theta}} \sup_{\theta \in \mathcal{F}_k} \mathbb{E} \left[ \frac{1}{n} \|\hat{\theta}(\mathbf{y}) - \theta\|^2 \right] \\ &\geq \frac{\sigma^2}{4\|\mathbf{A}\|_F^2} \left( \frac{k}{4} \log \frac{n}{k} - 1 \right) \frac{(1 - \Delta)}{(1 + \lambda)} \varepsilon_M, \end{aligned}$$

For large  $n$  we have

$$M^*(\mathbf{A}) \geq \frac{\sigma^2}{4\|\mathbf{A}\|_F^2} \left( \frac{k}{4} \log \frac{n}{k} \right) (1 - o(1)).$$

- Improvement of factor close to 8 over Fano argument of Candès-Davenport, which gave

$$M^*(\mathbf{A}) \geq \frac{\sigma^2}{32\|\mathbf{A}\|_F^2(1 + \beta)} \left( \frac{k}{4} \log \frac{n}{k} - 2 \right).$$

- MSE of same order achievable with  $\mathbf{A}$  that satisfies RIP  $\Rightarrow$  improvement beyond constant factors not possible

# Density estimation

- ▶ Consider  $\mathcal{F}$ , set of probability densities  $\theta$  on  $[0, 1]$  such that

$$a_0 \leq \theta \leq a_1 \quad \text{and} \quad |\theta''(x)| \leq a_2$$

- ▶ We are given  $(Y_1, \dots, Y_n)$  generated IID from  $\theta$ .
- ▶ Wish to estimate density with  $\hat{\theta}_n = \hat{\theta}_n(Y_1, \dots, Y_n)$ .
- ▶ Measure performance by squared Hellinger distance

$$d(\theta, \hat{\theta}_n) = \int_0^1 \left( \sqrt{\theta(x)} - \sqrt{\hat{\theta}_n(x)} \right)^2 dx.$$

Wish to obtain lower bound on minimax risk  $\inf_{\hat{\theta}_n} \sup_{\theta \in \mathcal{F}} \mathbb{E} d(\theta, \hat{\theta}_n)$

## Packing set (see Yu '97)

Packing set consists of densities that are small perturbations of uniform density on  $[0, 1]$

- ▶ Fix a smooth, bounded  $g(x)$  with

$$\int_0^1 g(x) dx = 0 \quad \text{and} \quad \int_0^1 (g(x))^2 dx = a.$$

- ▶ Partition  $[0, 1]$  into  $m$  subintervals of length  $1/m$
- ▶ Perturb uniform density in each subinterval by small amount proportional to rescaled version of  $g$
- ▶ That is, for some  $c$  define

$$g_j(x) = \frac{c}{m^2} g(mx - j) \mathbb{I} \left( \frac{j}{m} \leq x < \frac{j+1}{m} \right), \quad \text{for } j = 0, \dots, m-1.$$

## Packing set (contd.)

- ▶ Hypercube class of  $2^m$  densities

$$\left\{ f_{\tau}(y) = 1 + \sum_{j=0}^{m-1} \tau_j g_j(y) : \tau = (\tau_1, \dots, \tau_m) \in \{\pm 1\}^m \right\}$$

(In subinterval  $j$ , perturb uniform by either  $g_j$  or  $-g_j$ )

- ▶ Bandwidth parameter  $m$  chosen later to optimize lower bound

## Packing set (contd.)

- ▶ Hypercube class of  $2^m$  densities

$$\left\{ f_{\tau}(y) = 1 + \sum_{j=0}^{m-1} \tau_j g_j(y) : \tau = (\tau_1, \dots, \tau_m) \in \{\pm 1\}^m \right\}$$

(In subinterval  $j$ , perturb uniform by either  $g_j$  or  $-g_j$ )

- ▶ Bandwidth parameter  $m$  chosen later to optimize lower bound

Pick packing set corresponding to well-separated sequences in  $\{-1, 1\}^m$  (guaranteed by Gilbert-Varshamov)

- ▶  $\mathcal{A} \subseteq \{-1, 1\}^m$  whose elements have pairwise Hamming distance  $\geq m/3$
- ▶ Size of  $\mathcal{A} \geq \exp(c_0 m)$ , where  $c_0 \simeq 0.082$
- ▶ Resulting packing set  $\{f_{\tau} : \tau \in \mathcal{A}\}$  has minimum squared Hellinger distance  $d_{\min} = ac^2/(3m^4)$  (see Bin Yu)

## Using main theorem

For  $Q_Y$  uniform and  $\lambda = 1$  in main theorem, Rènyi term is

$$\left[ \sum_{\tau \in \mathcal{A}} \frac{1}{M} \int_{[0,1]^n} f_{\tau}^n(\mathbf{y})^2 d\mathbf{y} \right]^{\frac{1}{2}} \leq \exp \left( \frac{c^2 a n}{2m^4} \right).$$



### Proposition:

With  $m = n^{1/5}/\nu$  for any positive constant  $\nu < (c_0/(c^2a))^{1/5}$ , the minimax risk satisfies

$$\inf_{\hat{\theta}_n} \sup_{\theta \in \mathcal{F}} \mathbb{E} d(\hat{\theta}_n, \theta) \geq \frac{c^2 a \nu^4}{6} n^{-4/5} \varepsilon_M,$$

where

$$\varepsilon_M \geq 1 - 2 \exp \left( \frac{-n^{1/5}}{2\nu} (c_0 - \nu^5 c^2 a) \right).$$

- Bin Yu method uses same packing set + Fano, but gives  $\varepsilon_M$  bounded away from zero, not converging to 1

# Summary

Lower bounds on minimax risk: packing set + lower bound on  $\varepsilon_M$

- ▶ Computable via bounding Rényi divergence, gives strong converse
- ▶ Other example in paper: active learning of binary classifier
- ▶ Improvements over main theorem possible (Baraud arxiv:1807.05410)

## Further work:

- ▶ Can this method give improved minimax rates, rather than just improved constants?
- ▶ Extend results to work with global metric entropy features [Yang-Barron '99], [Guntuboyina '11]

Paper in *Electronic Journal of Statistics* (OA), 2018

doi: 10.1214/18-EJS1419

<https://arxiv.org/abs/1706.04410>