

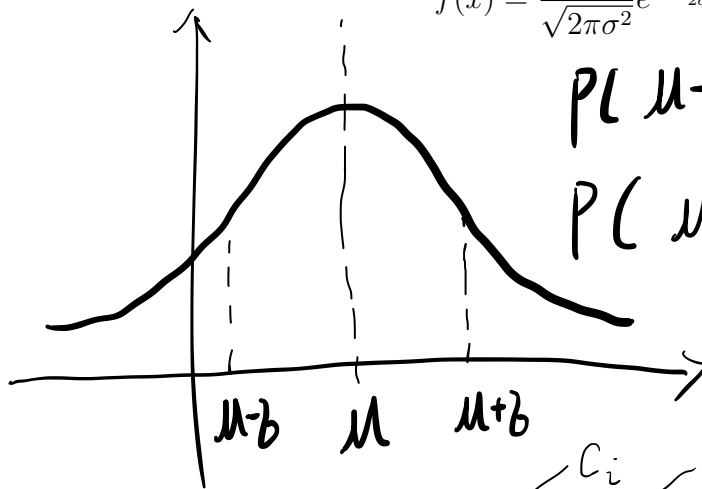
Packet 2: Functions of Random Variables

Chap 5.5 Random Variables related with Normal distributions

Normal distribution (Gaussian distribution) is originally found by observing that mean of sample often follows a special bell shaped distribution.

$X \sim N(\mu, \sigma^2)$, $E(X) = \mu$, $Var(X) = \sigma^2$, has p.d.f. and m.g.f.

$$f(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, M_X(t) = e^{\mu t + \frac{\sigma^2 t^2}{2}}$$



$$P(\mu - b < X < \mu + b) = 0.68$$

$$P(\mu - 2b < X < \mu + 2b) = 0.95$$

Theorem 5.5-1: If X_1, X_2, \dots, X_n are independent random variables with $X_i \sim N(\mu_i, \sigma_i^2)$, then $Y = \sum_{i=1}^n c_i X_i \sim N(\sum_{i=1}^n c_i \mu_i, \sum_{i=1}^n c_i^2 \sigma_i^2)$.

proof.

$$M_Y(t) = \prod_{i=1}^n M_{X_i}(c_i t) = \prod_{i=1}^n e^{\mu_i c_i t + \frac{\sigma_i^2 c_i^2 t^2}{2}}$$

$$= e^{\sum_{i=1}^n c_i \mu_i \cdot t + \sum_{i=1}^n c_i^2 \sigma_i^2 \times \frac{t^2}{2}}$$

mean variance

is the m.g.f. for a normal r.v.

$$\text{So } Y \sim N\left(\sum_{i=1}^n c_i \mu_i, \sum_{i=1}^n c_i^2 \sigma_i^2\right)$$

$$\text{if } c_i = \frac{1}{n} \quad \mu_i = \mu \quad \sigma_i^2 = \sigma^2$$

Corollary 5.5-1: If X_1, X_2, \dots, X_n are independent random variables with $X_i \sim N(\mu, \sigma^2)$, then $\bar{X} \sim N(\mu, \sigma^2/n)$.

sample mean $\bar{X} \sim N\left(\sum_{i=1}^n \frac{1}{n} \mu, \sum_{i=1}^n \left(\frac{1}{n}\right)^2 \sigma^2\right)$
 $\sim N\left(\mu, \frac{1}{n} \sigma^2\right)$ as $n \rightarrow \infty$
 $\sqrt{\text{Var}(\bar{X})} \rightarrow 0$

Theorem 5.5-2: If X_1, X_2, \dots, X_n are independent random variables with $X_i \sim N(\mu, \sigma^2)$, then the sample mean \bar{X} and the sample variance $S^2 = \frac{\sum_{i=1}^n (X_i - \bar{X})^2}{n-1}$ are independent,

$$\frac{S^2(n-1)}{\sigma^2} = \frac{\sum_{i=1}^n (X_i - \bar{X})^2}{\sigma^2} \sim \chi^2(n-1)$$

$$E\left(\frac{S^2(n-1)}{\sigma^2}\right) = n-1 \quad \frac{(n-1)}{\sigma^2} E(S^2) = n-1 \quad E(S^2) = \sigma^2$$

sample variance S^2 is an unbiased estimator for true variance σ^2

Theorem 5.5-3: Student's t distribution $T = \frac{Z}{\sqrt{U/r}} \sim t(r)$, where $Z \sim N(0, 1)$ and $U \sim \chi^2(r)$.

If X_1, X_2, \dots, X_n are independent random variables with $X_i \sim N(\mu, \sigma^2)$, then

$$T = \frac{\bar{X} - \mu}{S/\sqrt{n}} \sim t(n-1)$$

$$\bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right) \quad Z = \frac{\bar{X} - \mu}{\sqrt{\sigma^2/n}} \sim N(0, 1)$$

$$U = \frac{S^2(n-1)}{\sigma^2} \sim \chi^2(n-1)$$

$$T = \frac{Z}{\sqrt{U/(n-1)}} \sim t(n-1) \quad \frac{\frac{\bar{X} - \mu}{\sqrt{\sigma^2/n}}}{\sqrt{\frac{S^2(n-1)}{\sigma^2(n-1)}}} = \frac{\frac{\bar{X} - \mu}{\sqrt{\sigma^2/n}}}{\frac{S}{\sqrt{\sigma^2}}} = \frac{\bar{X} - \mu}{S/\sqrt{n}}$$

T statistic

Chap 5.6 The Central Limit Theorem (CLT)

i. i. d. independently and identically distributed

CLT tells us that, with sufficiently many i.i.d. samples collected, the sample mean \bar{X} follows $N(\mu, \sigma^2/n)$ approximately, regardless the true distribution of X_i .

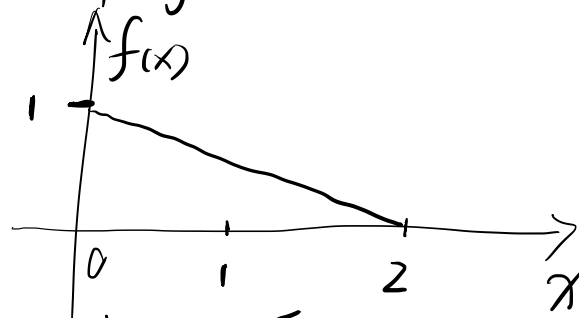
The distribution of X_i 's is complicated or unknown.

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i \sim N(\mu, \sigma^2/n) \text{ when } n \text{ is large}$$

where $E(X_i) = \mu$ $Var(X_i) = \sigma^2$

Example : $X_1 \dots X_{18}$ i.i.d. with p.d.f.

$$f(x) = 1 - \frac{1}{2}x \quad 0 < x < 2$$



Find the approximate distn of \bar{X} by CLT

$$\text{Sol: } \mu = E(X) = \int_0^2 x f(x) dx = 2/3$$

$$E(X^2) = \int_0^2 x^2 f(x) dx = 2/3$$

$$\sigma^2 = Var(X) = E(X^2) - (E(X))^2 = 2/3 - (2/3)^2 = 2/9$$

by CLT $\bar{X} \sim N(\mu, \frac{\sigma^2}{n}) = N(\frac{2}{3}, \frac{2/9}{18})$ $N(\frac{2}{3}, \frac{1}{81})$

* CLT is much easier than finding the exact distn of \bar{X}