## Packet 7: More Testing

### Chap 9.1 Chi-Square Goodness-of-Fit Test

#### The frequency table and the multinomial distribution:

Sample units are classified into K mutually exclusive categories, the number of units falling into each category is recorded.

The frequency table is a "One way" table in which units are classified according to a single categorical variable.

E.g., Eye color: Brown, Blue, Black, Green, and others. (unordered categorical variable or nominal variable).

E.g., Attitude toward war: strongly agree, agree, disagree, strongly disagree. (ordered but no numerical scores).

E.g., number of children in a family: 0,1,2, ..., 22. (ordered with numerical values).

#### General setting:

Suppose a random experiment has K possible outcomes, say  $A_1, A_2, \dots, A_K$ .

Let  $p_i = P(A_i)$ , and thus  $\sum_i p_i = 1$ .

Repeat experiments n times independently. Let  $X_i$  be the number of  $A_i$  in n trials.

#### Assumptions:

 $X = (X_1, X_2, \dots, X_K) \sim \text{multinomial distribution}(n, p)$ , where n is often known and  $p = (p_1, p_2, \dots, p_k)$ . The probability mass function is

$$f(x_1, x_2, \dots, x_k) = \frac{n!}{x_1! x_2! \dots x_k!} p_1^{x_1} p_2^{x_2} \dots p_k^{x_k}.$$

The critical assumptions:

- N trials are independent !!!
- The parameter p remains constant from trial to trial.

The most common violation occurs when clustering is presented in the data.

E.g., Suppose eye color samples were not collected from unrelated individuals , but from multiple families. Persons within a family are more likely to have the same color than persons from different families.

Pearson (Chi-square) goodness of fit

$$H_0$$
:  $p_1 = p_{10}, p_2 = p_{20}, \cdots, p_K = p_{K0},$ 

 $H_1$ : at least one equation does not hold.

Special case: Let  $K=2, X_1 \sim Bin(n,p_1), X_2 \sim Bin(n,p_2).$   $\bigvee_{1} + \bigvee_{2} = \bigvee_{1} + \bigvee_{2} = \bigvee_{2} + \bigvee_{2} + \bigvee_{2} = \bigvee_{2} + \bigvee_{2} + \bigvee_{2} = \bigvee_{2} + \bigvee_{2} + \bigvee_{2} + \bigvee_{2} = \bigvee_{2} + \bigvee_{2} +$ 

$$\chi_1 + \chi_2 = \mathcal{H}$$
  $P_1 + P_2 =$ 

$$X_1$$
 and  $X_2$  are not independent,  $X_2 = n - X_1$  is fully determined after knowing  $X_1$ 

$$Z = \frac{\hat{P}_{10} - P_{10}}{\sqrt{P_{10}(1-P_{10})/n}} \sim N(0,1)$$

$$\frac{1}{P_{10}(1-P_{10})} = \frac{1-P_{10}}{P_{10}(1-P_{10})} + \frac{P_{10}}{P_{10}(1-P_{10})} = \frac{(\chi_1 - n P_{10})^2}{n P_{10}} + \frac{(\chi_1 - n P_{10})^2}{n (1-P_{10})} = \frac{1}{P_{10}} + \frac{1}{P_{10}(1-P_{10})} = \frac{(\chi_1 - n P_{10})^2}{n (1-P_{10})} + \frac{(\chi_1 - n P_{10})^2}{n (1-P_{10})}$$

$$\frac{1 - l_{10} = l_{20}}{1 + 1 + 1} = 11 = 11 + 11 + 11 + 11 = 11$$

$$Z = \frac{\int_{l_0}^{l_0} - P_{l_0}}{\sqrt{P_{l_0}(l - P_{l_0})/\eta}} \sim N(0, 1)$$

$$Z^2 = \frac{\left(\frac{\hat{P}_l - P_{l_0}}{P_{l_0}(l - P_{l_0})/\eta} \times M^2\right)}{\sqrt{P_{l_0}(l - P_{l_0})/\eta} \times M^2} \sim \sqrt{2}$$

$$= \frac{\left( \chi_{1} - n \rho_{10} \right)^{2}}{n \rho_{10} \left( 1 - \rho_{10} \right)}$$

pre defined values

$$= \frac{\left(\chi_{1} - n \rho_{10}\right)^{2}}{n \rho_{10}} + \frac{\left(\chi_{1} - n \rho_{10}\right)^{2}}{n \left(1 - \rho_{10}\right)}$$

$$= \frac{(\chi_{1} - n \gamma_{10})^{2}}{n \gamma_{10}} + \frac{(\chi_{2} - n \gamma_{20})^{2}}{n \gamma_{20}}$$

$$= \frac{\left(O_1 - E_1\right)^2}{E_1} + \frac{\left(O_2 - E_2\right)^2}{E_2}$$

 $(\chi_1 - nP_{10})^2 = (nP_{20} - \chi_2)^2$  Chi-square test statistic If Ho is wrong, then  $Z^2$  will be large.

For i=1,2

Ti is the observed count in category i, denoted by Oi

In Pio is the expected count under Ho2, denoted by Ei

Extend to k categories: 
$$Q = \sum_{i=1}^{K} \frac{(O_i - E_i)^2}{E_i} \sim \chi^2_{(K-1)}.$$
 Chisy (1-d, K-1)

If  $O_i$  is far away from  $E_i$  for some categories then test statistic Q is large, reject  $H_0$ .

# C. L. T. $Q \sim \chi^2$ if n is large enough to have $E_j = np_j < 5$ for no more than 20% of the cells in the table. None of $E_j$ should fall below 1.

If the above condition is not satisfied, we often combine cells until all  $E_i$  are large enough.

Example 1: A bag of candies have 4 colors. Test if the 4 colors are in equal proportions at  $\alpha = 0.05$ .

 $H_0: p1 = p2 = p3 = p4$  v.s.  $H_1:$  not all equal.

$$n=224, \ \text{observe} \ X_1=42, \ X_2=64, \ X_3=53, \ X_4=65.$$
 O<sub>2</sub> O<sub>3</sub> O<sub>4</sub>

$$E_1 = E_2 = E_3 = E_4 = M \times 0.25 = 224 \times 0.25 = 56$$

$$Q = \sum_{i=1}^{\frac{4}{5}} \frac{(0_i - E_i)^2}{E_i} = \frac{(42-56)^2}{56} + \frac{(64-56)^2}{56} + \frac{(53-56)^2}{56} + \frac{(65-56)^2}{56}$$

Critical Region 
$$Q > 1^2 \circ 0.05 (4-1)$$
 gchisg (0.95, 3) = 7.815  
Because observed  $Q = 6.25 < 7.815$  We do not reject Ho.

Example 2: Flip 3 coins at same time and record the number of heads X = 0, 1, 2, 3 If P(head) = 0.3 and 3 coins are independently flipped, we should have  $X \sim Binomial(3, 0.3)$ .

Suppose flip n=200 times, and observe  $Y_0 = 57$ ,  $Y_1 = 95$ ,  $Y_2 = 38$ ,  $Y_3 = 10$ .  $O_1$   $O_2$   $O_3$   $O_4$   $H_0: X \sim Binomial(3, 0.3)$  v.s.  $H_1: X$  follows other distributions.  $\alpha = 0.05$ .

$$\frac{H_0: X \sim Binomial(3,0.3)}{P_1 = P_{10} = P(X=0) = \binom{3}{0} 0.3^0 0.7^3 = 0.343} \qquad E_1 = 200 \times P_{10} = 68.6$$

$$P_2 = P_{20} = P(X=1) = \binom{3}{1} 0.3^1 0.7^2 = 0.441$$

$$P_3 = P_{30} = P(X=2) = \binom{3}{2} 0.3^2 0.7^1 = 0.189$$

$$P_4 = P_{40} = P(X=3) = \binom{3}{3} 0.3^3 0.7^0 = 0.027$$

$$E_1 = 200 \times P_{10} = 68.6$$

$$E_2 = 200 \times P_{20} = 88.2$$

$$E_3 = 200 \times P_{30} = 37.8$$

$$E_4 = 200 \times 0.027 = 5.4$$

$$E_1 = 200 \times P_{10} = 68.6$$

double check the calculation & Ei = M

$$Q = \frac{4}{5} \frac{(0i - Ei)^2}{Ei} = \frac{(57 - 68.6)^2}{68.6} + \frac{(95 - 88.2)^2}{88.2} + \frac{(38 - 37.8)^2}{37.8} + \frac{(10 - 5.4)^2}{5.4}$$

$$= 6.4$$

Critical Region 
$$Q = \chi_{0.05}^2 (4-1) = 7.815$$

Sometimes you are interested in testing whether a data set fits a probability model with dparameters left unspecified.

For instance what if the probability of head was unspecificed in the previous example and you simply want to know whether the distribution has a form of binomial?

- 1. Estimate the d parameters e.g. using the maximum likelihood method.
- 2. Calculate the chi-square statistic Q using the obtained estimates.
- 3. Compare the chi-square statistic to a chi-square distribution with k-1-d degrees of d is the number of unknown parameters that we need to estimate in order to get

Example 3: Flip 3 coins at same time and record the number of heads X = 0, 1, 2, 3. Suppose

flip n=200 times, and observe  $Y_0 = 57$ ,  $Y_1 = 95$ ,  $Y_2 = 38$ ,  $Y_3 = 10$ .  $0_1 \quad 0_2 \quad 0_3 \quad 0_4$   $H_0: X \sim Binomial(3, \underline{p}) \text{ where } \underline{p} \text{ is unknown} \text{ v.s. } H_1: X \text{ follows other distributions.}$  $\alpha = 0.05$ .

To estimate 
$$P$$
, we use sample proportion  $P$ 
 $200 \times 3 = 600$  flips  $P = \frac{0 \times 57 + 1 \times 95 + 2 \times 38 + 3 \times 10}{600} = \frac{201}{600} = 0.335$ 
 $P_{10} = P(X=0) = {3 \choose 0} 0.335^{\circ} (1-0.335)^{3}$ 

$$E_{i} = 200 \times P_{i}, \qquad Q = \frac{4}{i} \frac{\left(0_{i} - E_{i}\right)^{2}}{E_{i}}$$

we estimated P = 0.335

critical region 
$$Q > \chi^2$$

$$\chi^2$$

$$\chi^$$

Sometimes, we also collapse categories with small probabilities. The chi-square distribution relies on C.L.T. which needs relatively large sample size for each category, e.g.  $E_i \geq 5$ .

Example 4: We observe the number of small particles and conducted 100 experiments. Let X be the number of particles in each experiment and Y be the frequency of getting each set of X values.