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5.1-5. The pdf of X is $f(x) = \theta x^{\theta-1}$, $0 < x < 1$, $0 < \theta < \infty$. Let $Y = -2\theta \ln X$. How is Y distributed?

$$f(x) = \theta x^{\theta-1}$$

$$Y = -2\theta \ln x$$

$$x = e^{-\frac{Y}{2\theta}}$$

$$Y \in (0, \infty)$$

$$\frac{dy}{dx} = \frac{-2\theta}{x}$$

$$f_Y(y) = \theta x^{\theta-1} \cdot \frac{x}{2\theta}$$

$$Y \in (0, \infty)$$

$$= \frac{1}{2} (x)^\theta$$

$$= \frac{1}{2} e^{-\frac{Y}{2\theta} \theta}$$

$$= \frac{1}{2} e^{-\frac{Y}{2}}$$

$$Y \sim \exp\left(\frac{1}{2}\right)$$

~~5.1-15.~~ Let $Y = X^2$.

(a) Find the pdf of Y when the distribution of X is $N(0, 1)$.

(b) Find the pdf of Y when the pdf of X is $f(x) = (3/2)x^2, -1 < x < 1$.

a) $X \sim N(0, 1)$

$$Y = X^2$$

$$X = \pm \sqrt{Y}$$

for $X = 0$

$$Y = 0$$

$$f(y) = \frac{1}{2} \frac{1}{\sqrt{\frac{1}{2}}} y^{\frac{1}{2}-1} e^{-\frac{1}{2}y}$$

$$= \frac{1}{\sqrt{2\pi}y} e^{-\frac{y}{2}} \quad 0 < y < \infty$$

b $f(x) = \frac{3}{2} x^2 \quad -1 < x < 1$

$$F(x) = \int_{-1}^x \frac{3}{2} t^2 dt = \frac{1}{2} (x^3 + 1)$$

$$y = x^2$$

$$x = \pm \sqrt{y}$$

$$F(y) = \frac{y^{\frac{3}{2}+1}}{2} - \frac{-y^{\frac{3}{2}-1}}{2} = y^{\frac{3}{2}}$$

$$f(y) = \frac{d}{dy} F(y) = \frac{d}{dy} y^{\frac{3}{2}} = \frac{3}{2} \sqrt{y} \quad 0 < y < 1$$

5.2-1. Let X_1, X_2 denote two independent random variables, each with a $\chi^2(2)$ distribution. Find the joint pdf of $Y_1 = X_1$ and $Y_2 = X_2 + X_1$. Note that the support of Y_1, Y_2 is $0 < y_1 < y_2 < \infty$. Also, find the marginal pdf of each of Y_1 and Y_2 . Are Y_1 and Y_2 independent?

$$Y_1 = X_1$$

$$X_1 = Y_1$$

$$Y_2 = X_2 + X_1$$

$$X_2 = Y_2 - X_1 = Y_2 - Y_1$$

$$\begin{aligned} f(x_1, x_2) &= \frac{1}{2} e^{-\frac{x_1}{2}} \frac{1}{2} e^{-\frac{x_2}{2}} \\ &= \frac{1}{4} e^{-\frac{x_1 + x_2}{2}} \end{aligned}$$

$$|J| = \begin{vmatrix} \frac{\partial x_1}{\partial y_1} & \frac{\partial x_1}{\partial y_2} \\ \frac{\partial x_2}{\partial y_1} & \frac{\partial x_2}{\partial y_2} \end{vmatrix}$$

$$= \begin{vmatrix} 1 & 0 \\ -1 & 1 \end{vmatrix} = 1$$

$$f(y_1, y_2) = f(x_1, x_2) |J|$$

$$= \frac{1}{4} e^{-\frac{y_2}{2}}$$

$$f(y_1) = \int_{-\infty}^{\infty} \frac{1}{4} e^{\frac{y_1}{2}} dy_1$$

$$= \frac{1}{4} e^{\frac{y_1}{2}}$$

$$f(y_2) = \frac{1}{4} e^{\frac{y_2}{2}}$$

$$f(y_1, y_2) \neq f(y_1) \cdot f(y_2)$$

So they are not independent.

5.2-5. Let the distribution of W be $F(8, 4)$. Find the following:

(a) $F_{0.01}(8, 4)$.

(b) $F_{0.99}(8, 4)$.

(c) $P(0.198 \leq W \leq 8.98)$.

a $F_{0.01}(8, 4) = 14.8$

b
$$\begin{aligned} F_{0.99}(8, 4) &= \frac{1}{F_{0.01}(8, 4)} \\ &= \frac{1}{14.8} \\ &= 0.06756756756756757 \end{aligned}$$

c. $P(0.198 \leq W \leq 8.98)$

$$P(W \leq 8.98) = 0.975$$

$$P(W \leq 0.198) = 0.025$$

$$P(0.198 \leq W \leq 8.98) = 0.975 - 0.025 = 0.95$$

5.2-6. Let X_1 and X_2 have independent gamma distributions with parameters α, θ and β, θ , respectively. Let $W = X_1/(X_1 + X_2)$. Use a method similar to that given in the derivation of the F distribution (Example 5.2-4) to show that the pdf of W is

$$g(w) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} w^{\alpha-1} (1-w)^{\beta-1}, \quad 0 < w < 1.$$

We say that W has a beta distribution with parameters α and β . (See Example 5.2-3.)

$$W = \frac{X_1}{X_1 + X_2}$$

$$g(w) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} w^{\alpha-1} (1-w)^{\beta-1} \quad 0 < w < 1$$

X_1 & X_2 are independent gamma dist
with α, θ & β, θ

~~5.3-4.~~ Let X_1 and X_2 be a random sample of size $n = 2$ from the exponential distribution with pdf $f(x) = 2e^{-2x}$, $0 < x < \infty$. Find

(a) $P(0.5 < X_1 < 1.0, 0.7 < X_2 < 1.2)$.

(b) $E[X_1(X_2 - 0.5)^2]$.

$$a) f(x_1, x_2) = 2e^{-2x_1} \cdot 2e^{-2x_2} = 4e^{-2x_1 - 2x_2}$$

$$P(0.5 < X_1 < 1.0, 0.7 < X_2 < 1.2)$$

$$\int_{0.5}^{1.0} \int_{0.7}^{1.2} 4e^{-2x_1 - 2x_2} dx_2 dx_1$$

$$= 4 \int_{0.5}^{1.0} e^{-2x_1} \left. \frac{e^{-2x_2}}{-2} \right|_{0.7}^{1.2} dx_1$$

$$= 0.31176 \left. \frac{e^{-2x_1}}{-2} \right|_{0.5}^{1.0}$$

$$= 0.03625$$

$$\begin{aligned}
E(X_1(X_2 - 0.5)) &= E(X_1(X_2^2 + 0.25 - X_2)) \\
&= E(X_1X_2^2 + 0.25X_1 - X_1X_2) \\
&= E(X_1X_2^2) + E(0.25X_1) - E(X_1X_2) \\
&= E(X_1)E(X_2^2) + 0.25E(X_1) - E(X_1) \cdot E(X_2) \\
&= \frac{1}{2} \cdot \frac{2}{4} + \frac{1}{4} \cdot \frac{1}{2} - \frac{1}{2} \cdot \frac{1}{2} \\
&= \frac{1}{4} + \frac{1}{8} - \frac{1}{4} \\
&= \frac{1}{8}
\end{aligned}$$

5.3-11. Let X_1, X_2, X_3 be three independent random variables with binomial distributions $b(4, 1/2)$, $b(6, 1/3)$, and $b(12, 1/6)$, respectively. Find

(a) $P(X_1 = 2, X_2 = 2, X_3 = 5)$.

(b) $E(X_1 X_2 X_3)$.

(c) The mean and the variance of $Y = X_1 + X_2 + X_3$.

$$X_1 \sim b(4, \frac{1}{2})$$

$$X_2 \sim b(6, \frac{1}{3})$$

$$X_3 \sim b(12, \frac{1}{6})$$

$$\begin{aligned} a) \quad P(X_1=2, X_2=2, X_3=5) &= P(X_1=2)P(X_2=2)P(X_3=5) \\ &= \binom{4}{2} \left(\frac{1}{2}\right)^2 \left(1-\frac{1}{2}\right)^{4-2} \cdot \binom{6}{2} \left(\frac{1}{3}\right)^2 \left(1-\frac{1}{3}\right)^{6-2} \\ &\quad \cdot \binom{12}{5} \left(\frac{1}{6}\right)^5 \left(1-\frac{1}{6}\right)^{12-5} \\ &= 0.375 \cdot 0.329 \cdot 0.0284 \\ &= 0.00351 \end{aligned}$$

$$\begin{aligned} b) \quad E(X_1 X_2 X_3) &= E(X_1) \cdot E(X_2) \cdot E(X_3) \\ &= 4 \cdot \frac{1}{2} \cdot 6 \cdot \frac{1}{3} \cdot 12 \cdot \frac{1}{6} \\ &= 2^3 \\ &= 8 \end{aligned}$$

c. mean & var d) $Y = X_1 + X_2 + X_3$

$$E(Y) = E(X_1 + X_2 + X_3)$$

$$= E(X_1) + E(X_2) + E(X_3)$$

$$= 2 + 2 + 2$$

$$= 6$$

$$\text{Var}(Y) = \text{Var}(X_1) + \text{Var}(X_2) + \text{Var}(X_3)$$

$$= 4 \cdot \frac{1}{2} \cdot \frac{1}{2} + 6 \cdot \frac{1}{3} \cdot \frac{2}{3} + 12 \cdot \frac{1}{6} \cdot \frac{5}{6}$$

$$= 1 + \frac{12}{9} + \frac{60}{36}$$

$$= 4$$

