

# Packet 3: Point Estimation

## Maximum Likelihood Estimator

**Likelihood function** (R. A. Fisher, 1922) of a model  $f(x | \theta)$  is the joint probability density or mass function of the observed data  $x = \{x_1, x_2, \dots, x_n\}$ , viewed as a function of  $\theta$ . For example, if  $X = \{X_1, X_2, \dots, X_n\}$  are continuous r.v.s,

$$L(\theta) = f(x | \theta) = f(x_1, x_2, \dots, x_n | \theta) = \prod_{i=1}^n f(x_i | \theta), \text{ if independent.}$$

If the data are discrete r.v.s,

$$L(\theta) = P(X = x | \theta) = P(X_1 = x_1, \dots, X_n = x_n | \theta) = \prod_{i=1}^n P(X_i = x_i | \theta), \text{ if independent.}$$

In this discrete case, the likelihood function is the “probability” that we observe the data  $\{X = x\}$  under  $\theta$ . For example, let's say  $L(0.8) \gg L(0.2)$ . It means that the probability of observing the current data  $P(X = x | \theta)$  is much higher when  $\theta = 0.8$ . So, the data seem to support  $\theta = 0.8$  much more than  $\theta = 0.2$ ; the data themselves speak about  $\theta$ ! In general,  $L(\theta)$  indicates how likely the observed data are as a function of  $\theta$ , and maximizing the likelihood function determines the parameters that are most likely to produce the observed data.

*Example:* We want to know the number of ducks living at Penn State Duck Pond (Hintz Alumni Garden) in this summer, and we count the number of ducks in 3 consecutive days  $x = (12, 13, 17)$ . Assume the number of observed ducks follows a uniform distribution,  $\text{Uniform}[0, \theta]$ , where  $\theta$  is the total number of ducks. The p.d.f. of  $\text{Uniform}[0, \theta]$  is given by

$$f(x | \theta) = \frac{1}{\theta} I_{\{0 \leq x \leq \theta\}}. \quad \text{means } f(x | \theta) = 0 \text{ if } x < 0 \text{ or } x > \theta$$

Which  $\theta$  most likely generate those three observations?

A:  $\theta = 30$ , B:  $\theta = 20$ , C:  $\theta = 10$ .

*Likelihood*

$$A: L(\theta=30) = P(X_1=12 | \theta=30) P(X_2=13 | \theta=30) P(X_3=17 | \theta=30) \\ = \frac{1}{30} \times \frac{1}{30} \times \frac{1}{30}$$

$$B: L(\theta=20) = \frac{1}{20} \times \frac{1}{20} \times \frac{1}{20}$$

$$C: L(\theta=10) = 0 \times 0 \times 0 = 0$$

$$* L(\theta=16) = \frac{1}{16} \times \frac{1}{16} \times 0 = 0 \quad \Leftarrow \begin{array}{l} X \sim U[0, 16] \text{ p.m.f.} = 0 \\ \text{if } X \text{ is out of support} \end{array} \\ L(\theta=17) = \frac{1}{17} \times \frac{1}{17} \times \frac{1}{17} \quad \hat{\theta} = 17 \text{ MLE}$$

**Maximum likelihood estimator:** A widely used method of obtaining a point estimate for a parameter  $\theta$  is to find the **maximum likelihood estimate (MLE)**. As the name suggests, the MLE is defined as **some value maximizing  $L(\theta)$**  in the parameter space  $\Omega$ .

In practice, we obtain the MLE by **maximizing  $\ell(\theta) = \log(L(\theta))$**  instead of maximizing  $L(\theta)$  for a few reasons.

1. Since  **$L(\theta)$  involves a product** when the data are independent, it is mathematically more convenient to work with the (natural) **logarithm of the likelihood function**. sum of log p.d.f or p.m.f.
2. The logarithmic function is strictly increasing, preserving the maximizing value, i.e., **the value of  $\theta$  that maximizes  $\ell(\theta)$  also maximizes  $L(\theta)$** .
3. When an analytic solution is not available, we need to find a **numerical solution** and it is computationally more stable to find the value of  $\theta$  that maximizes  $\ell(\theta)$ .

data

*Example:* If **we knew there were 10 ducks** and observed 8 of them on a random day. We assume that  $X_1, X_2, \dots, X_n \stackrel{i.i.d.}{\sim} \text{Bernoulli}(\theta)$  for some  $\theta \in [0, 1]$ , where  $X_i$  is 1 if we observe duck  $i$  and 0 otherwise. We want to find the most likely value of  $\theta$  that maximizes the probability of observing these data.  $\theta$  is the prob of observing a duck

Write down the likelihood function and log-likelihood function.

What is the MLE of  $\theta$ ?

Sol ① p.m.f.  $P(X_i=1) = \theta$   $P(X_i=0) = (1-\theta)$   $X_i = 0, 1$

likelihood  $L(\theta) = \prod_{i=1}^{10} P(X_i=x_i) = \theta^8 (1-\theta)^2$

log likelihood  $\log L(\theta) = \ell(\theta) = \log[\theta^8] + \log[(1-\theta)^2]$   
 $= 8 \log \theta + 2 \log(1-\theta)$

②  $\ell(\theta)$  based on a standard prob. distr is log-concave

it is maximized at  $\frac{\partial \ell(\theta)}{\partial \theta} = 0$  set  $\frac{\partial \ell(\theta)}{\partial \theta} = 0$

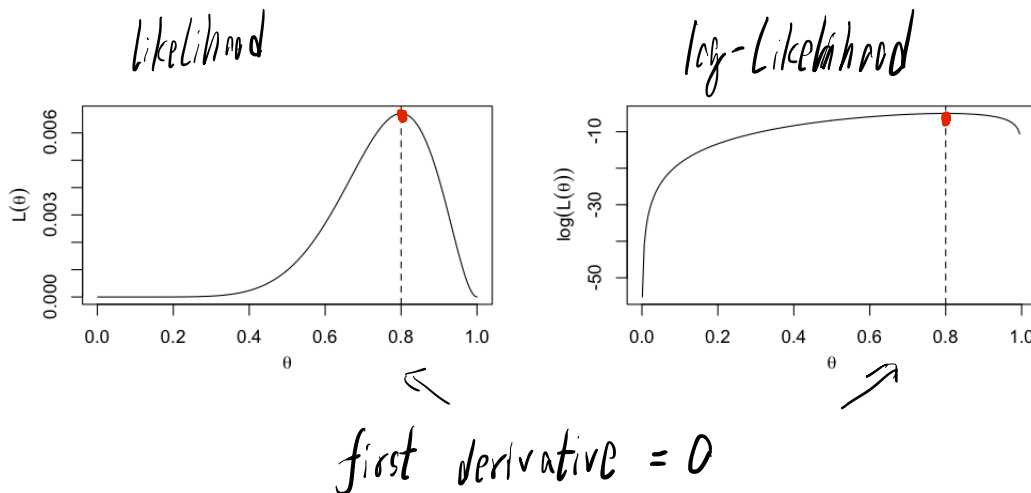
$\frac{\partial \ell(\theta)}{\partial \theta} = \frac{8}{\theta} + \frac{-2}{1-\theta} = 0$

multiply  $\theta(1-\theta)$  on both sides

③ solve equation  $8(1-\theta) - 2\theta = 0$

$8 - 8\theta - 2\theta = 0$

sample proportion  
 $\hat{\theta} = 0.8$   
**MLE**



*Example:* The lifetime of a particular type of light bulb can be modeled by an exponential distribution, and its p.d.f. is

$$f(x | \theta) = \frac{1}{\theta} \exp\left(-\frac{x}{\theta}\right) \text{ for } x > 0.$$

Suppose the average lifetime  $\theta$  is unknown, and we want to estimate it. We independently observe the lifetime of  $n$  such light bulbs,  $x_1, x_2, \dots, x_n$ . What is the MLE of the expected lifetime  $\theta$ ?

Sol

$$\begin{aligned} \textcircled{1} \quad L(\theta) &= \prod_{i=1}^n f(x_i | \theta) = \prod_{i=1}^n \frac{1}{\theta} \exp\left(-\frac{x_i}{\theta}\right) \\ &= \theta^{-n} \exp\left(-\frac{\sum x_i}{\theta}\right) \end{aligned}$$

$$l(\theta) = \log L(\theta) = -n \log \theta - \frac{\sum x_i}{\theta}$$

$$\textcircled{2} \quad l'(\theta) = -\frac{n}{\theta} - \left(-\frac{\sum_{i=1}^n x_i}{\theta^2}\right) = -\frac{n}{\theta} + \frac{\sum x_i}{\theta^2}$$

$$\textcircled{3} \quad l'(\theta) = -\frac{n}{\theta} + \frac{\sum x_i}{\theta^2} = 0$$

$$\sum x_i = n\theta$$

$$\hat{\theta} = \frac{\sum_{i=1}^n x_i}{n} = \bar{x} \quad \text{is the MLE of } \theta$$