

# Piecewise-linear Modeling of Multivariate Geometric Extremes

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UNC–Chapel Hill



# Limit sets and gauge functions

- $\mathbf{X}_1, \dots, \mathbf{X}_n \stackrel{\text{iid}}{\sim} f_{\mathbf{X}}$ ,  $d$ -dimensional, Exponential(1) margins...



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- ▶ If  $f_{\mathbf{X}}$  satisfies

$$\frac{-\log f_{\mathbf{X}}(t\mathbf{x})}{t} \longrightarrow g(\mathbf{x}) \quad , \text{ as } t \rightarrow \infty,$$

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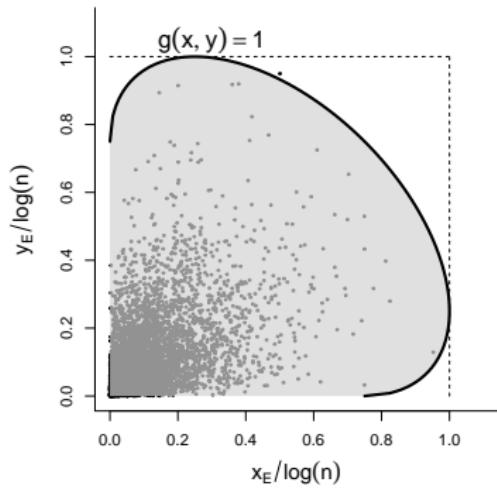
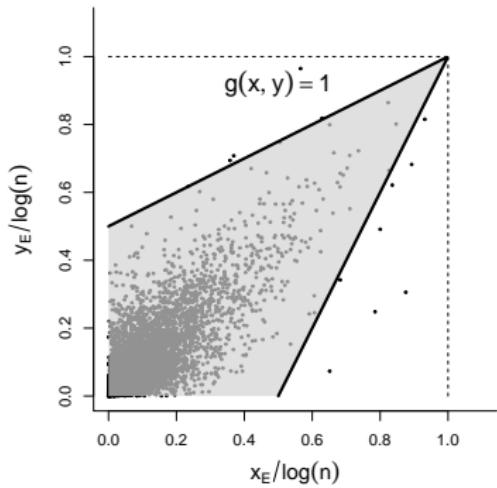
- ▶ ...then scaled sample clouds  $\left\{ \frac{\mathbf{X}_1}{\log n}, \dots, \frac{\mathbf{X}_n}{\log n} \right\}$  converge onto a **limit set**,

$$G := \left\{ \mathbf{x} \in \mathbb{R}^d \mid g(\mathbf{x}) \leq 1 \right\}$$

as  $n \rightarrow \infty$  (Balkema and Nolde, 2010; Nolde and Wadsworth, 2022).

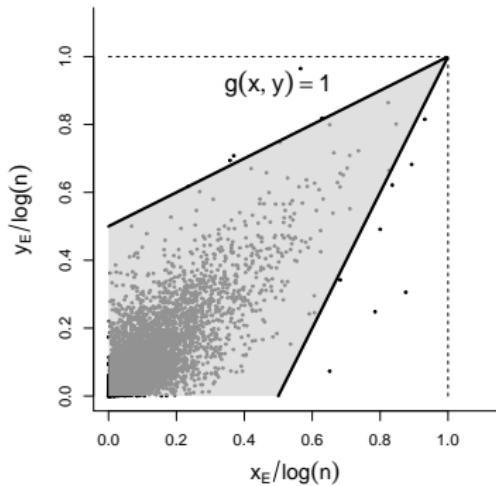


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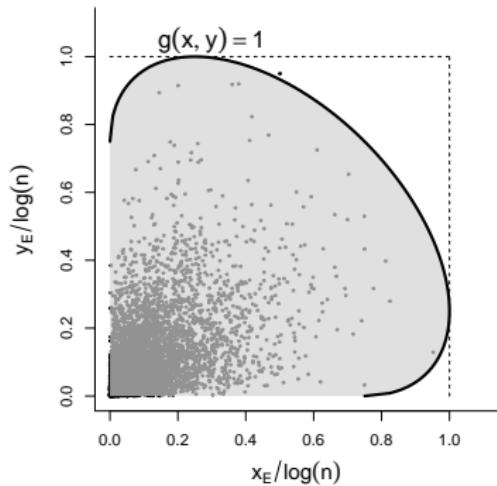




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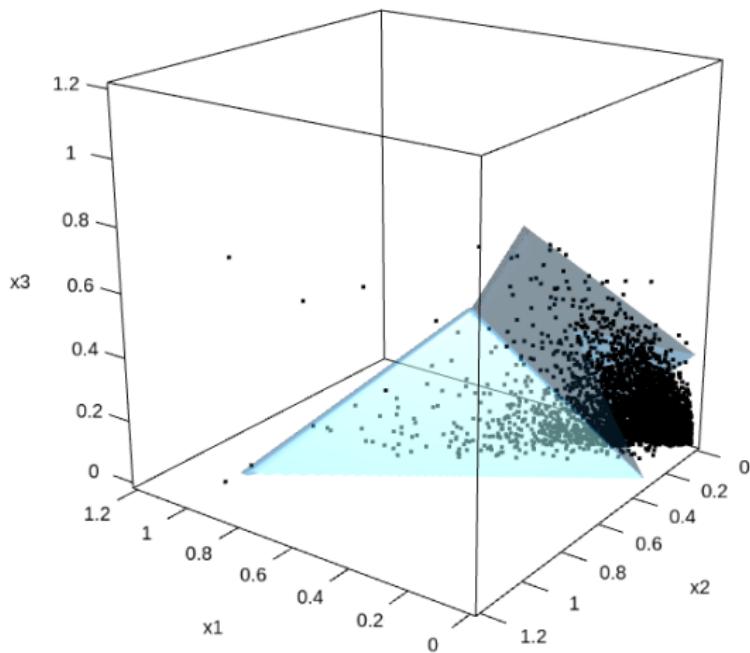
**“pointy”**  $\implies$  AD



**“not pointy”**  $\implies$  AI



# Limit sets and gauge functions

 Limit sets and gauge functions

**Is this approach useful for extremal statistical inference across the multivariate tail?**

**If so, how can we estimate  $g$  from data and use it for extremal statistical inference?**



## Extremal inference with gauge functions

- ▶ Define  $(R, \mathbf{W}) = (\|\mathbf{X}\|_1, \mathbf{X}/\|\mathbf{X}\|_1)$
- ▶  $f_{\mathbf{X}}$  satisfies  $-\log f_{\mathbf{X}}(r\mathbf{w}) \sim rg(\mathbf{w})$  as  $r \rightarrow \infty$ .



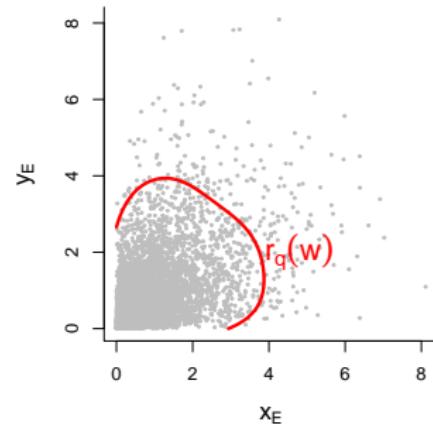
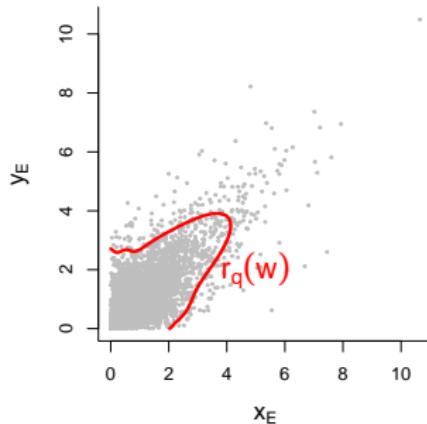
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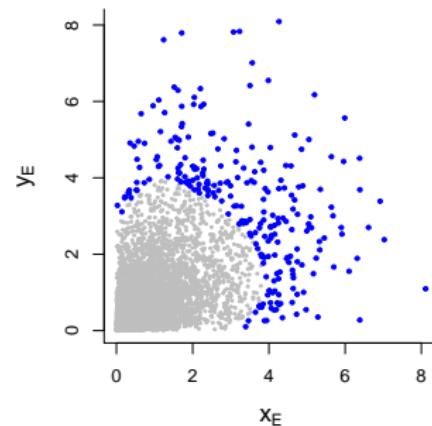
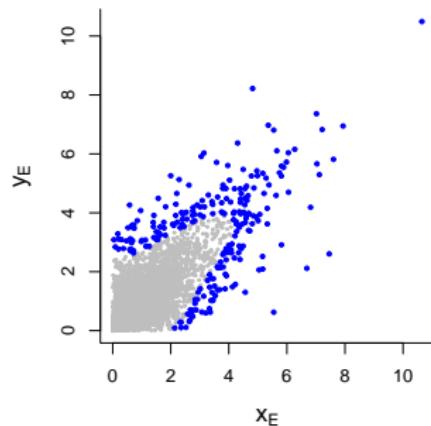
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## Geometric approach: model fitting

- Wadsworth and Campbell (2024): fit the model

$$R \mid \{W = w, R > r_q(w)\} \sim \text{truncGamma}(\alpha, g(w; \theta))$$

by maximizing

$$L(\theta; r_{1:n}, w_{1:n}) = \prod_{i:r_i > r_q(w_i)} \frac{f_{\text{Gamma}}(r_i; \alpha, g(w_i; \theta))}{\bar{F}_{\text{Gamma}}(r_q(w_i); \alpha, g(w_i; \theta))}$$



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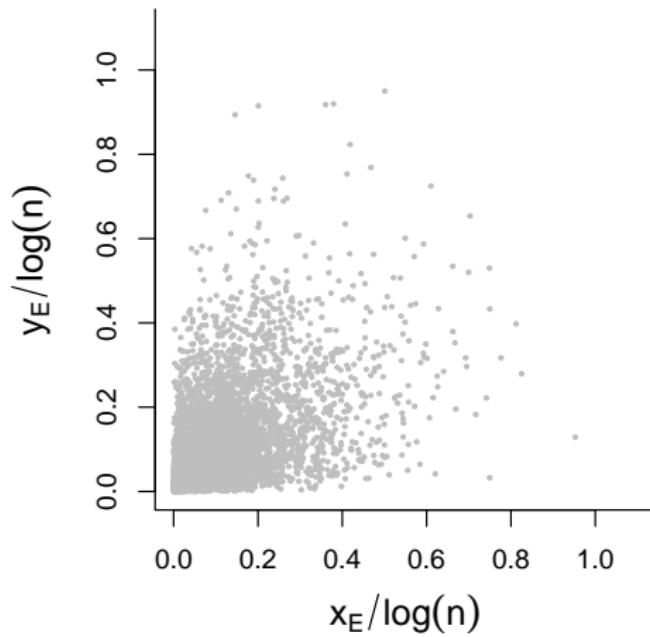
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- ▶ **Campbell and Wadsworth (2024): Define  $g(w; \theta)$  piecewise-linearly.**

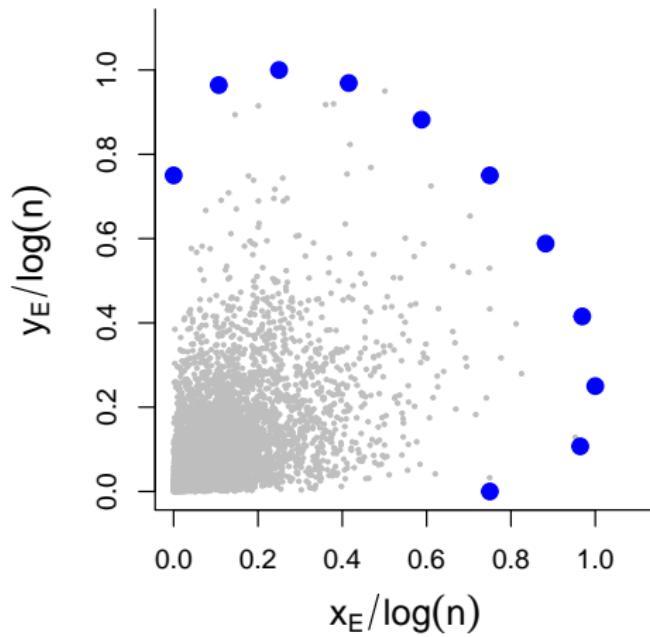


# Semiparametric piecewise-linear approach



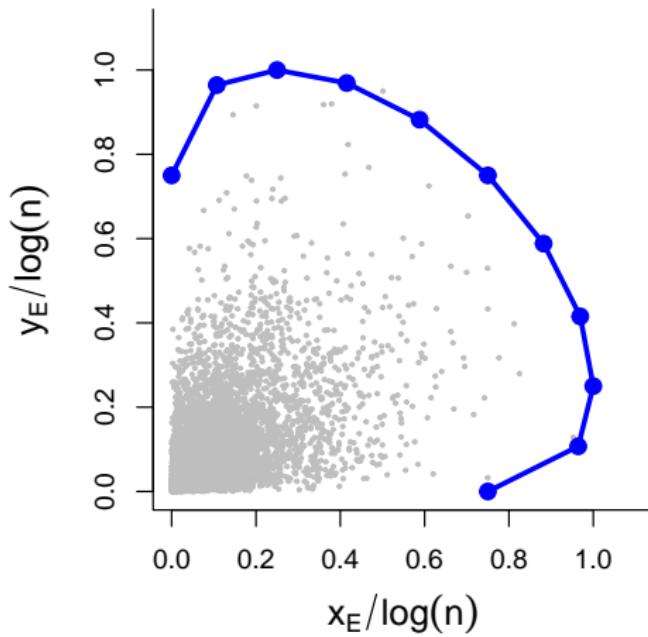


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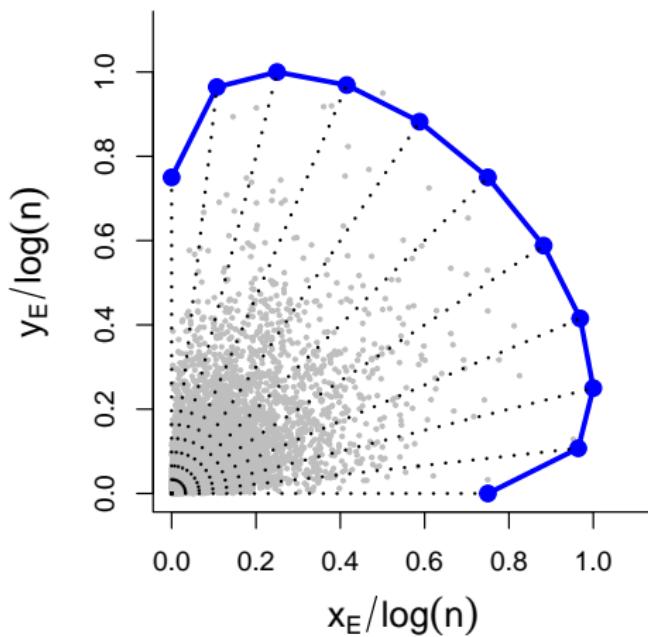


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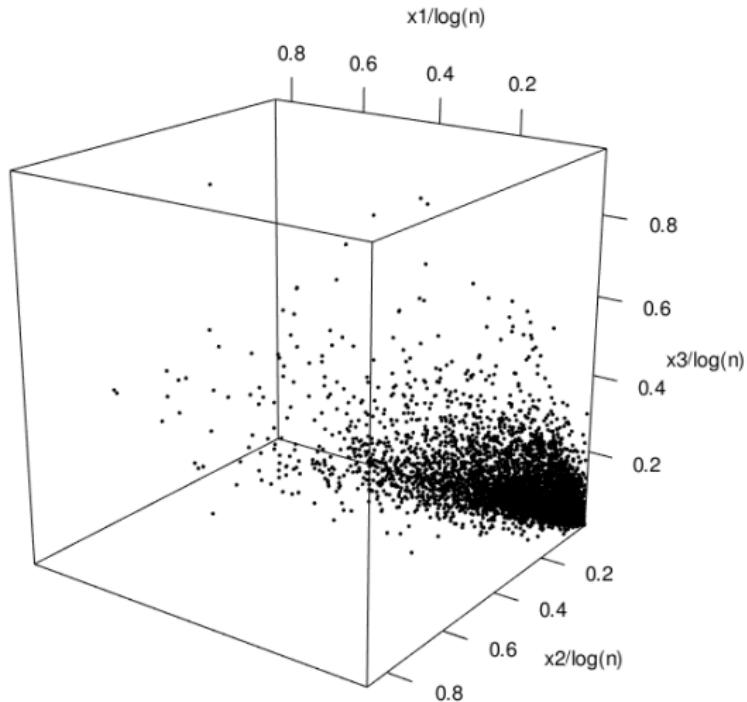


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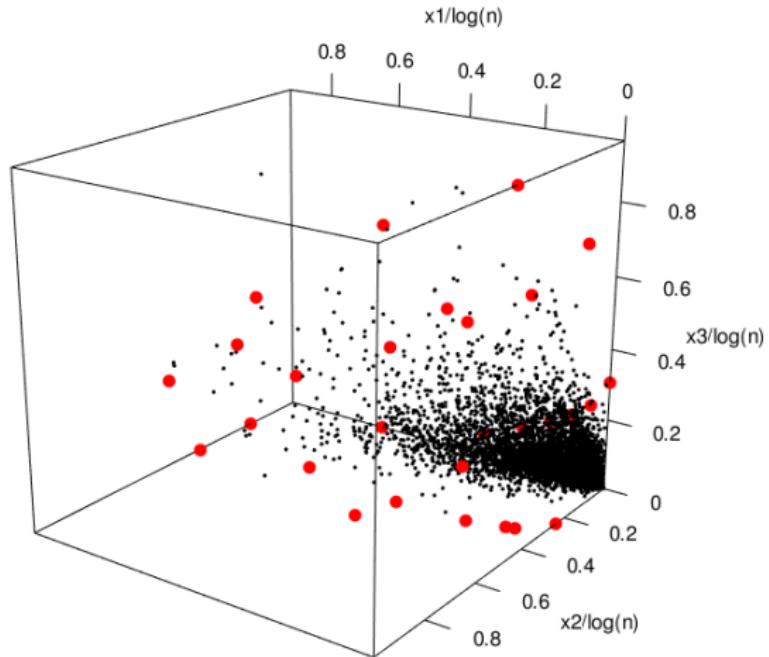


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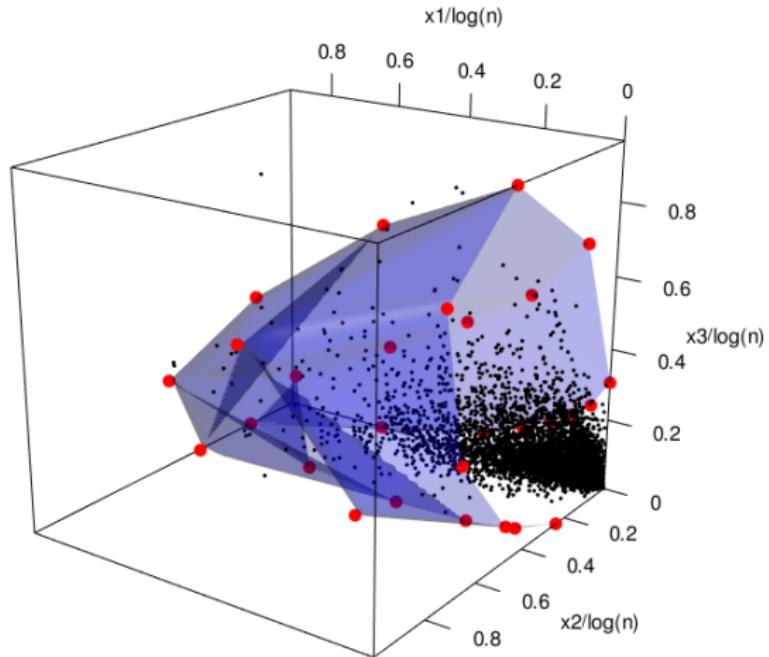


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In  $d$ -dimensions...

- ▶ Define a set of  $N$  reference angles  $\mathbf{w}^{*1}, \dots, \mathbf{w}^{*N} \in \mathcal{S}_{d-1}$ .
- ▶ Results in  $M$  regions:  $\triangle^{(1)}, \triangle^{(2)}, \dots, \triangle^{(M)}$ .



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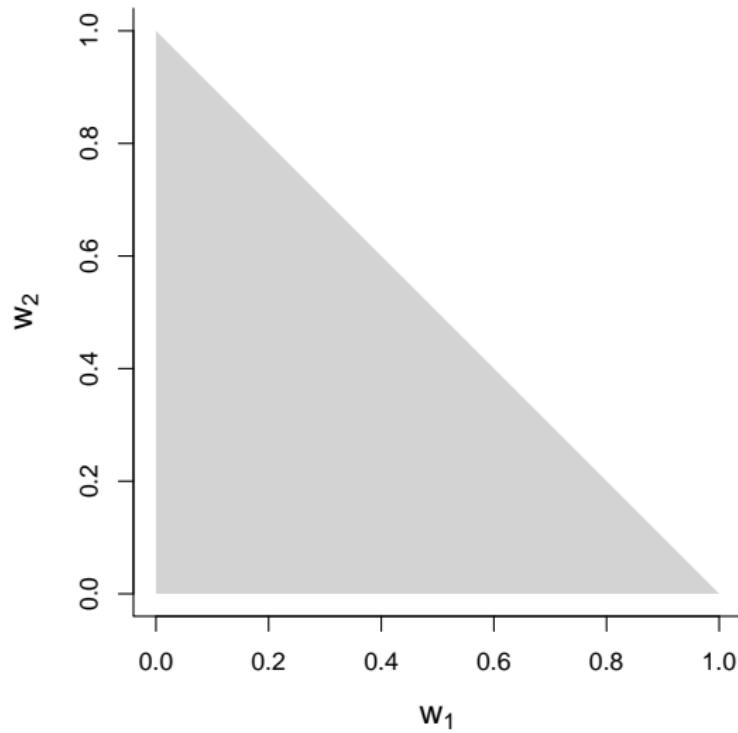
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- ▶ Results in  $M$  regions:  $\triangle^{(1)}, \triangle^{(2)}, \dots, \triangle^{(M)}$ .
- ▶ At a point  $\mathbf{x} \in \mathbb{R}^d$ , the gauge function value is given by

$$g_{\text{PWL}}(\mathbf{x}; \boldsymbol{\theta}) = \sum_{k=1}^M \mathbf{1}_{\triangle^{(k)}}(\mathbf{x}/\|\mathbf{x}\|) \frac{\mathbf{n}^{(k)\top} \mathbf{x}}{\mathbf{n}^{(k)\top} \boldsymbol{\theta}_1^{(k)} \mathbf{w}^{*(k),1}}$$



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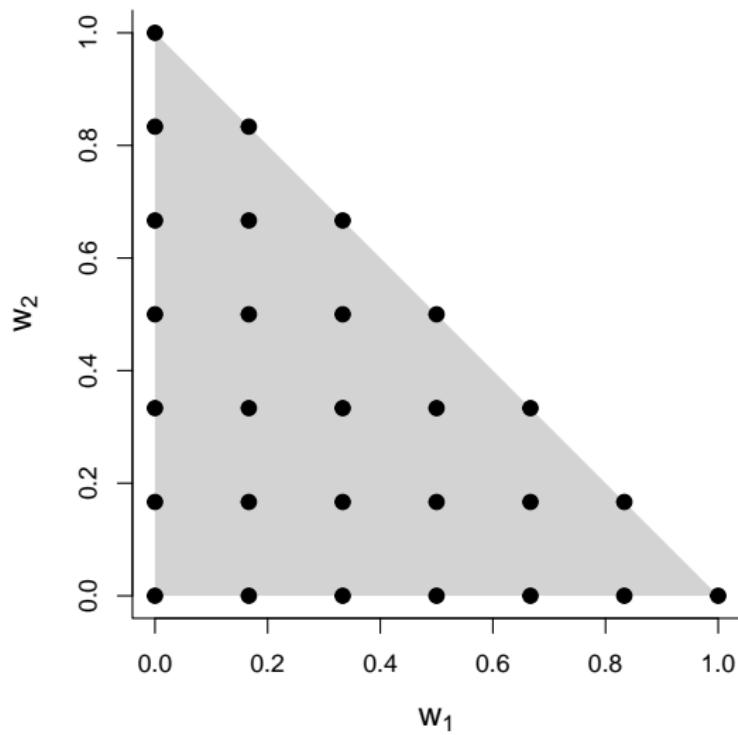
Partitioning the angular space,  $\mathcal{S}_2 = \{\mathbf{x} \in \mathbb{R}^3 : \|\mathbf{x}\|_1 = 1\}$





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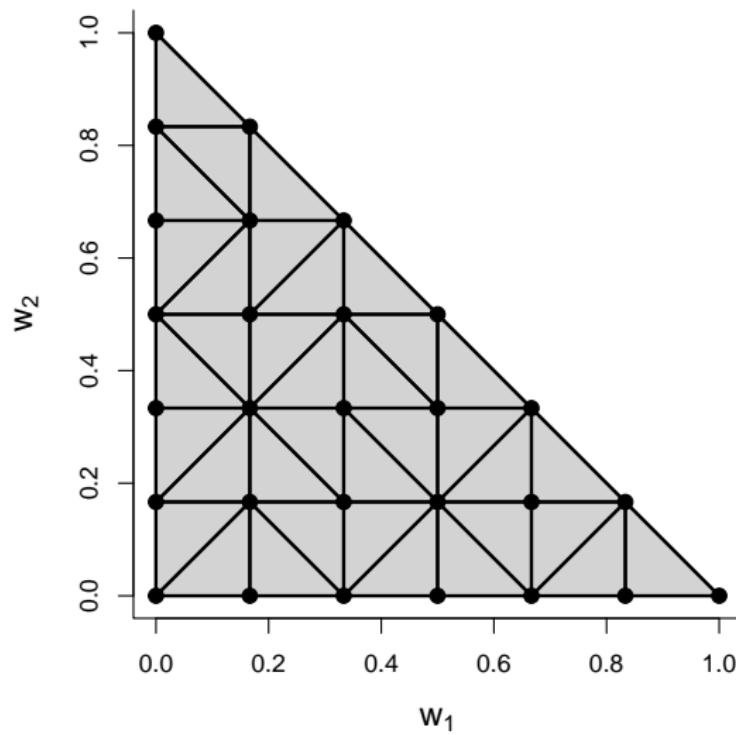
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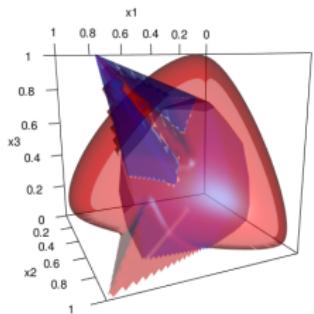
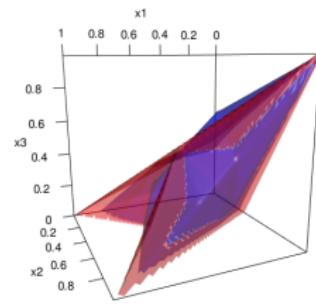
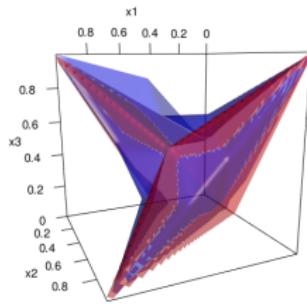
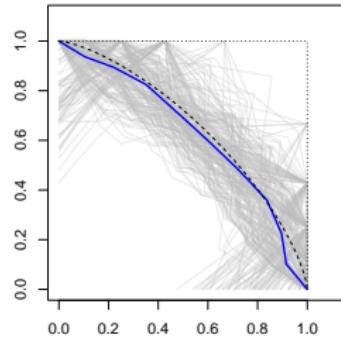
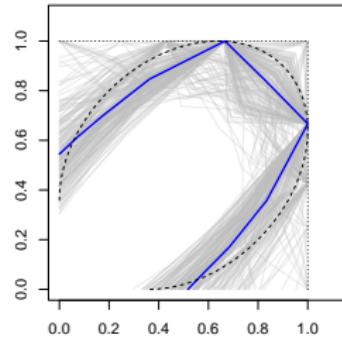
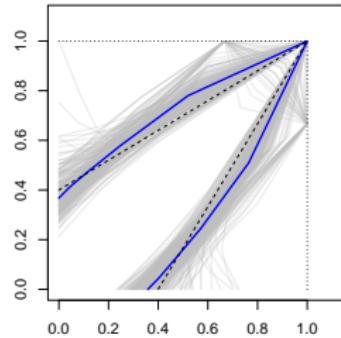
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## Geometric approach: extrapolation

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- ▶ Compute  $\Pr(\mathbf{X} \in B | R > r_q(\mathbf{W}))$  via sampling:

1. sample  $\mathbf{w}_1, \dots, \mathbf{w}_N$  from  $\mathbf{W} | \{R > r_q(\mathbf{W})\}$ .
2. sample  $r_i$  from  $\text{truncGamma}(\hat{\alpha}, g(\mathbf{w}_i; \hat{\theta}))$  for  $i = 1, \dots, N$
3. return  $\mathbf{x}_i = r_i \mathbf{w}_i$ ,  $i = 1, \dots, N$



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- ▶ Compute  $\Pr(R > r_q(\mathbf{W}))$  empirically.
- ▶ **Bonus!** Can extrapolate far into the tails using

$$\Pr(\mathbf{X} \in B) = \Pr(\mathbf{X} \in B | R > kr_q(\mathbf{W})) \Pr(R > kr_q(\mathbf{W}))$$

for  $k > 1$ .



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$$f_{\mathbf{W} \mid \{R > r_q(\mathbf{W})\}}(\mathbf{w}; \boldsymbol{\theta}) = \frac{g(\mathbf{w}; \boldsymbol{\theta})^{-d}}{d\text{vol}(\{\mathbf{x} : g(\mathbf{x}; \boldsymbol{\theta}) \leq 1\})}$$



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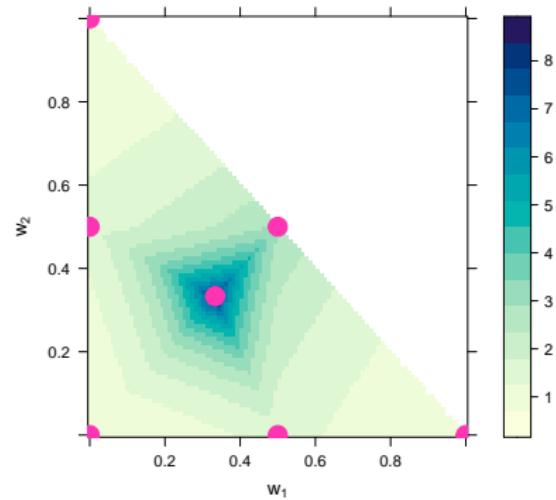
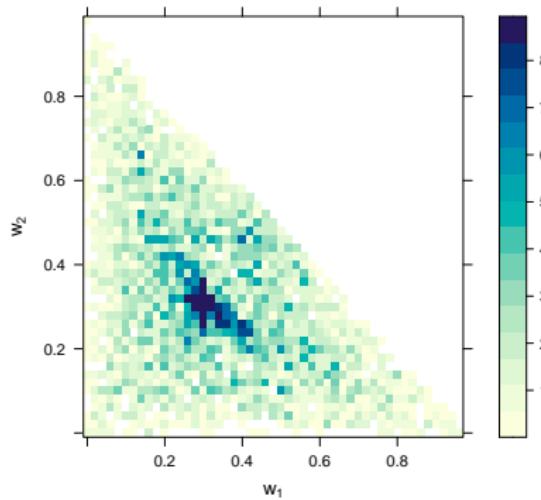
- ▶ Campbell and Wadsworth (2024):
  - ▶  $\text{vol}(\{\mathbf{x} : g_{\text{PWL}}(\mathbf{x}; \boldsymbol{\theta}) \leq 1\})$  has a closed-form expression.
  - ▶ samples are drawn from  $f_{\mathbf{W} \mid \{R > r_q(\mathbf{W})\}}$  using MCMC.



# Angular model

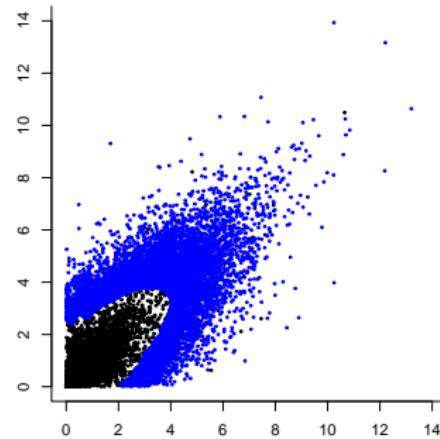
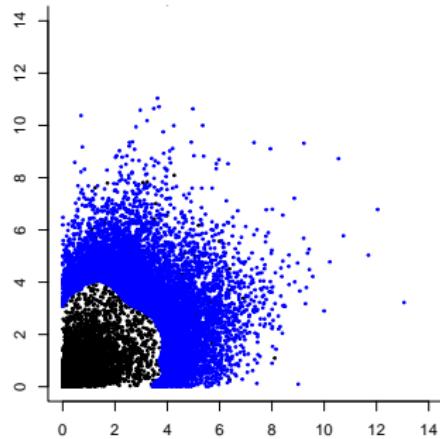
Fitting the angular density

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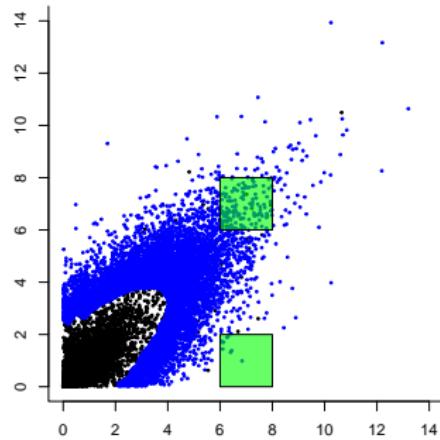
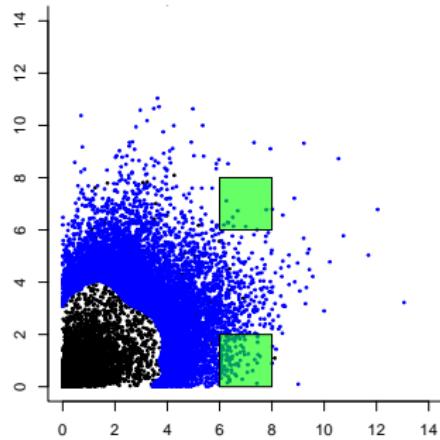


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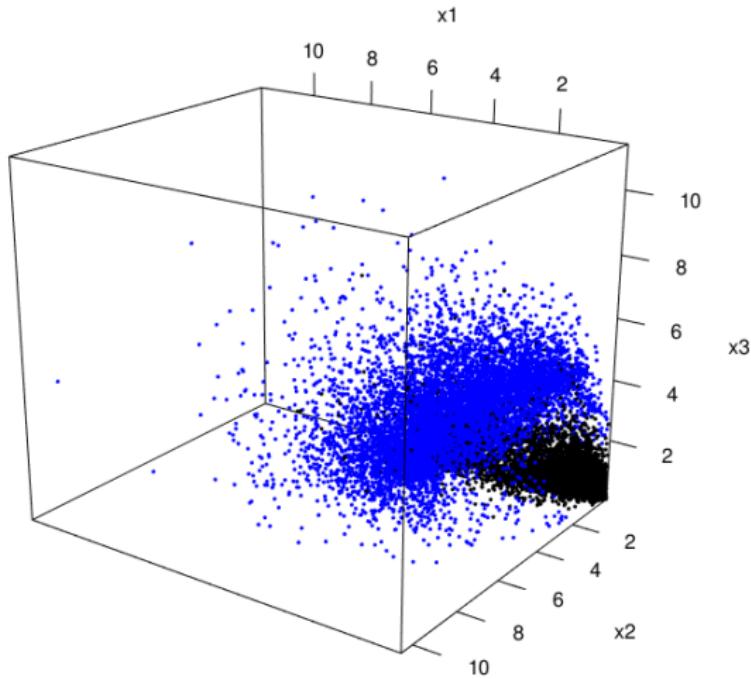


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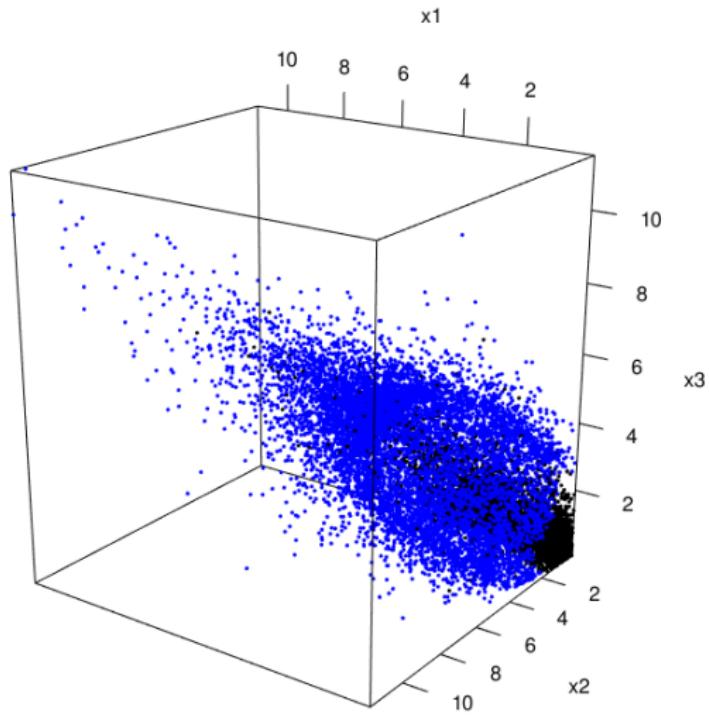


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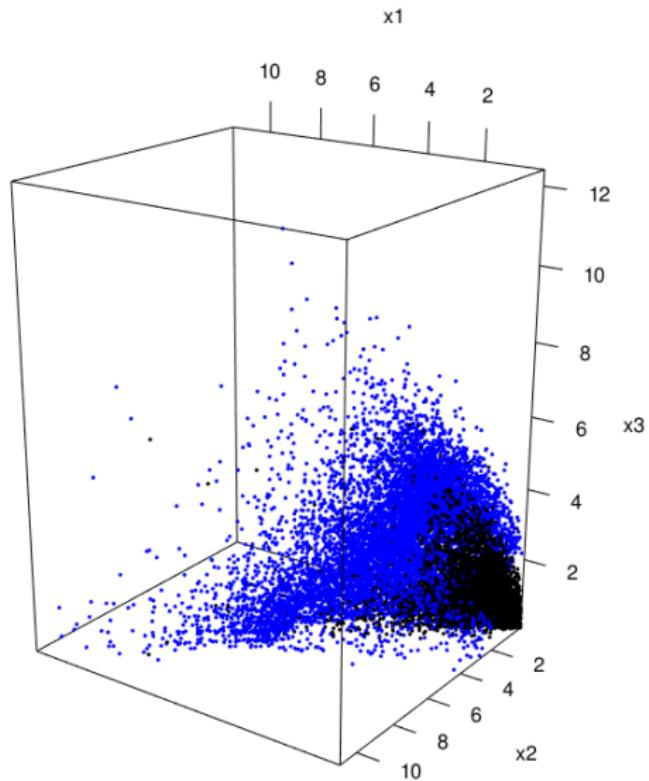


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## Application to air pollution measurements ( $d = 3$ )

- ▶ North Kensington site, London, UK
- ▶ carbon monoxide (CO, mg/m<sup>3</sup>), nitrogen dioxide (NO<sub>2</sub>, µg/m<sup>3</sup>), and particles with a diameter of 10 µm or less (PM10, mg/m<sup>3</sup>).
- ▶  $n = 5,584$  daily maximum measurements, October–April.  
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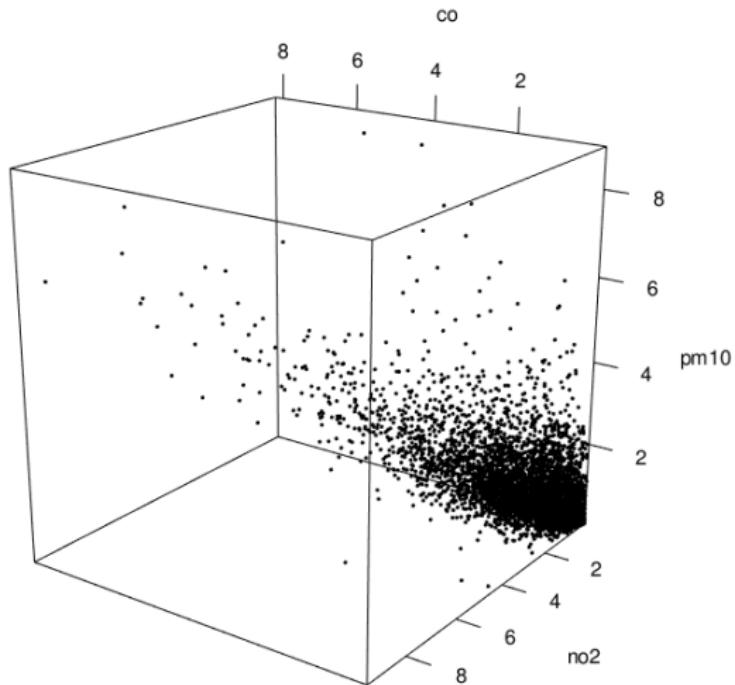


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- ▶  $n = 5,584$  daily maximum measurements, October–April. 1996–2024.
- ▶ Using methods from Simpson et al. (2020), evidence that
  - ▶  $\{\text{PM10}\}$  is large when  $\{\text{CO}, \text{NO}_2\}$  are small
  - ▶  $\{\text{CO}, \text{NO}_2, \text{PM10}\}$  grow large together.

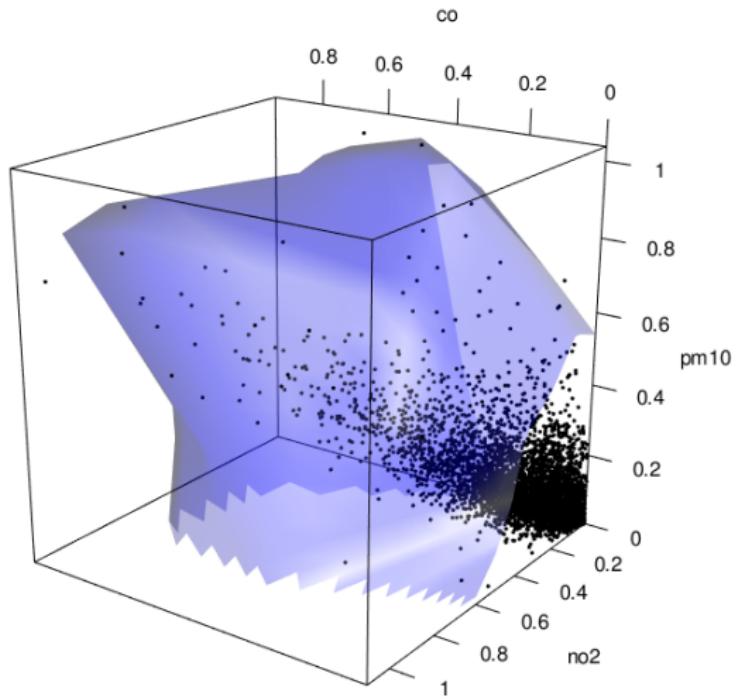


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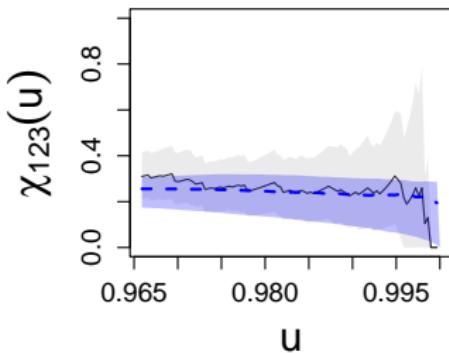
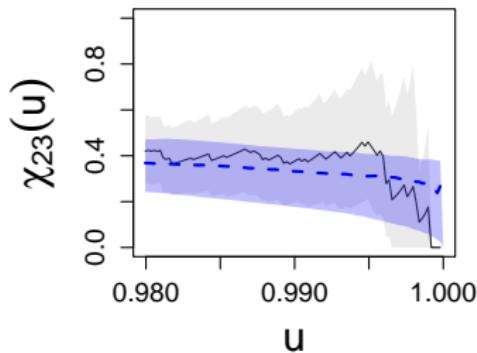
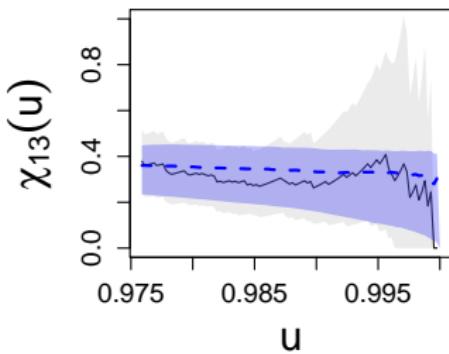
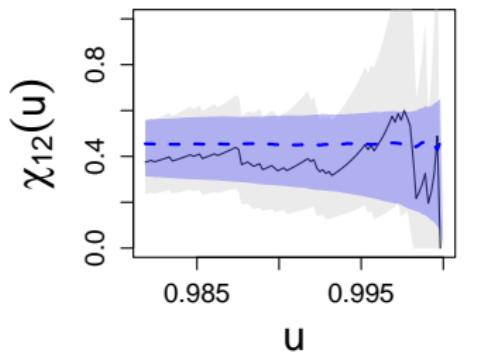


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## $\chi_C(u)$ estimates, $C \subseteq \{1, 2, 3\}$



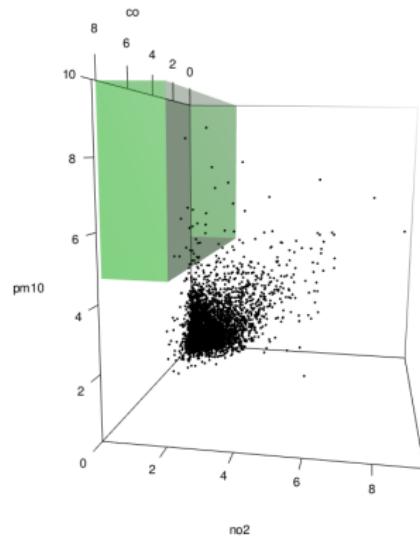
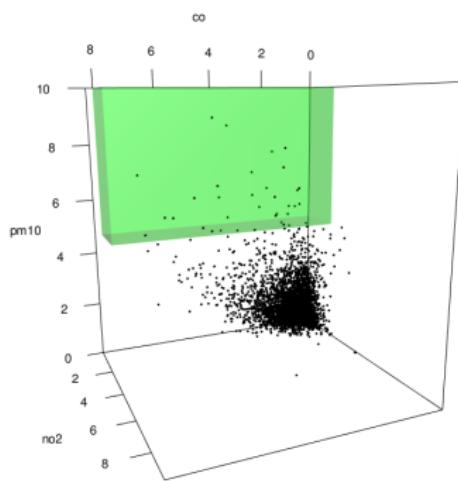


# An alternate tail probability

For  $d = 3$  pollution data, consider

$$\psi_{\{3\}}(u; \delta_1, \delta_2) = \Pr [F_E(X_1) < \delta_1, F_E(X_2) < \delta_2, F_E(X_3) > u]$$

for  $\delta_1, \delta_2 \in [0, 1]$ ,  $u > u_0$ ,  $u_0 \in [0, 1)$  close to 1.



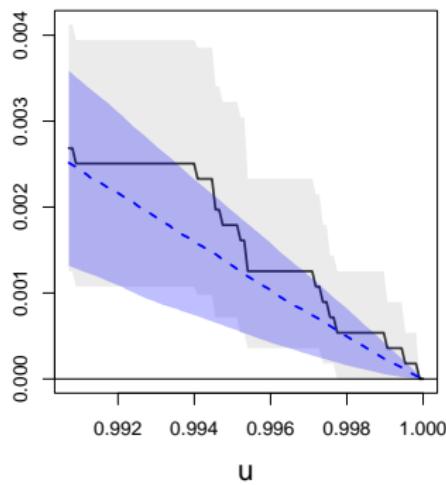


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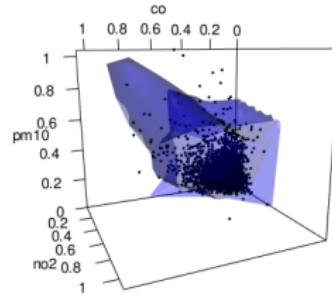
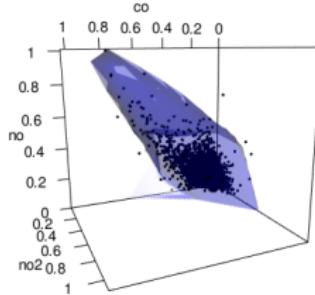
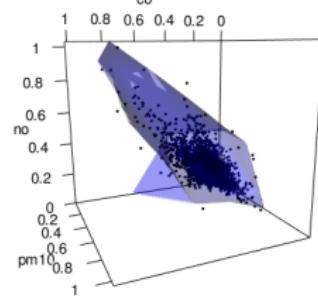
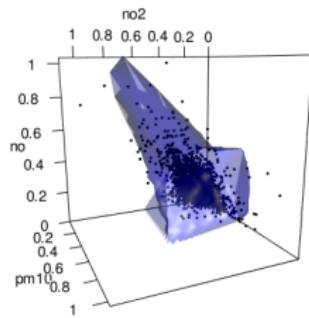
$$\psi_{\{3\}}(u; \delta_1, \delta_2) = \Pr [F_E(X_1) < \delta_1, F_E(X_2) < \delta_2, F_E(X_3) > u]$$

for  $\delta_1, \delta_2 \in [0, 1]$ ,  $u > u_0$ ,  $u_0 \in [0, 1)$  close to 1.





# Application to air pollution measurements ( $d = 4$ )



# Thank you!



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