

$$e) p(\theta_{j,c} | x, \pi, c) \propto p(\theta_{j,c}) p(x, c | \theta_{j,c}, \pi) = p(\theta_{j,c}) \prod_{i=1}^n p(x^{(i)}, c^{(i)} | \theta_{j,c}, \pi) = p(\theta_{j,c}) \prod_{i=1}^n \left[ p(c^{(i)} | \pi) \prod_{j=1}^d p(x_j^{(i)} | c^{(i)}, \theta_{j,c}) \right]$$

$$\propto \theta_{j,c}^{\alpha-1} (1-\theta_{j,c})^{\beta-1} \prod_{i=1}^n \left[ p(c^{(i)} | \pi) \prod_{j=1}^d p(x_j^{(i)} | c^{(i)}, \theta_{j,c}) \right] \quad (\text{note } p(x, c | \theta_{j,c}, \pi) \text{ is from part a}).$$

$$\log(p(\theta_{j,c} | x, \pi, c)) \propto (\alpha-1) \log \theta_{j,c} + (\beta-1) \log(1-\theta_{j,c}) + \sum_{i=1}^n \left[ \log p(c^{(i)} | \pi) + \sum_{j=1}^d \log p(x_j^{(i)} | c^{(i)}, \theta_{j,c}) \right]$$

$$= (\alpha-1) \log \theta_{j,c} + (\beta-1) \log(1-\theta_{j,c}) + \sum_{i=1}^n \left[ \log \pi_{c^{(i)}} + \sum_{j=1}^d \log \left( \theta_{j,c}^{x_j^{(i)}} (1-\theta_{j,c})^{(1-x_j^{(i)})} \right) \right]$$

$$= (\alpha-1) \log \theta_{j,c} + (\beta-1) \log(1-\theta_{j,c}) + \sum_{i=1}^n \log \pi_{c^{(i)}} + \sum_{j=1}^d \sum_{i=1}^n \left[ x_j^{(i)} \log \theta_{j,c} + (1-x_j^{(i)}) \log(1-\theta_{j,c}) \right]$$

$$= (\alpha-1) \log \theta_{j,c} + (\beta-1) \log(1-\theta_{j,c}) + \sum_{i=1}^n \log \pi_{c^{(i)}} + \sum_{j=1}^d \left[ N_{\text{pixel } j=1} \log \theta_{j,c} + N_{\text{pixel } j=0} \log(1-\theta_{j,c}) \right]$$

where  $N_{\text{pixel } j=1}$  is the # of pixels where  $j=1$  for images in class  $c$ , and similarly for  $N_{\text{pixel } j=0}$ .

Set the derivative to 0, so

$$0 \stackrel{\text{set}}{=} \frac{d \log(p(\theta_{j,c} | x, \pi, c))}{d \theta_{j,c}} = \frac{\alpha-1}{\hat{\theta}_{j,c}} - \frac{\beta-1}{1-\hat{\theta}_{j,c}} + 0 + \frac{N_{\text{pixel } j=1}}{\hat{\theta}_{j,c}} - \frac{N_{\text{pixel } j=0}}{1-\hat{\theta}_{j,c}}$$

$$= \frac{N_{\text{pixel } j=1} + \alpha - 1}{\hat{\theta}_{j,c}} - \frac{N_{\text{pixel } j=0} + \beta - 1}{1-\hat{\theta}_{j,c}} = N_{\text{pixel } j=1} + \alpha - 1 - \hat{\theta}_{j,c} (N_{\text{pixel } j=1} + \alpha - 1) - \hat{\theta}_{j,c} (N_{\text{pixel } j=0} + \beta - 1)$$

$$\Rightarrow \hat{\theta}_{j,c} = \frac{N_{\text{pixel } j=1} + \alpha - 1}{N_{\text{pixel } j=1} + N_{\text{pixel } j=0} + \alpha + \beta - 2} \quad \text{which for } \alpha=3, \beta=3 \text{ equals } \frac{N_{\text{pixel } j=1} + 2}{N_{\text{pixel } j=1} + N_{\text{pixel } j=0} + 4}.$$

Note that the MLE estimator for  $\theta_{j,c}$  from part a can also be expressed as

$$\hat{\theta}_{j,c} = \frac{N_{\text{pixel } j=1}}{N_{\text{pixel } j=1} + N_{\text{pixel } j=0}} \quad \text{which is similar to the MAP estimator but without } \alpha-1 \text{ in the numerator}$$

and  $\alpha+\beta-2$  in the denominator, where  $\alpha$  and  $\beta$  are pseudo-counts.

Parts f) and g) on last page.

h) An advantage is that it is efficient; for instance, during training it only requires one pass through the data, and during testing applying Bayes's rule can be cheap due to the model structure.

A disadvantage is that it may be less accurate in practice since its independence assumption is strong or naïve; in this case, the probability of a pixel being 1 often affects the probability of neighboring pixels.

3. a) Note that  $p(y=0|x, \theta) = 1 - \frac{1}{1 + \exp(-x^T \theta)}$ , so

$$L(\theta|x, y) = \prod_{i=1}^N \left[ \left( \frac{1}{1 + \exp(-x^{(i)T} \theta)} \right)^{y_i} \left( 1 - \frac{1}{1 + \exp(-x^{(i)T} \theta)} \right)^{1-y_i} \right]$$

$$\begin{aligned} \log L(\theta|x, y) &= \sum_{i=1}^N \log \left[ \left( \frac{1}{1 + \exp(-x^{(i)T} \theta)} \right)^{y_i} \left( 1 - \frac{1}{1 + \exp(-x^{(i)T} \theta)} \right)^{1-y_i} \right] \\ &= \sum_{i=1}^N \left[ y_i \log \left( \frac{1}{1 + \exp(-x^{(i)T} \theta)} \right) + (1-y_i) \log \left( \frac{\exp(-x^{(i)T} \theta)}{1 + \exp(-x^{(i)T} \theta)} \right) \right] \\ &= \sum_{i=1}^N \left[ y_i \left( \log \left( \frac{1}{1 + \exp(-x^{(i)T} \theta)} \right) - \log \left( \frac{\exp(-x^{(i)T} \theta)}{1 + \exp(-x^{(i)T} \theta)} \right) \right) + \log \left( \frac{\exp(-x^{(i)T} \theta)}{1 + \exp(-x^{(i)T} \theta)} \right) \right] \\ &= \sum_{i=1}^N \left[ y_i x_i^{(i)T} \theta + \log \left( \frac{\exp(-x^{(i)T} \theta)}{1 + \exp(-x^{(i)T} \theta)} \right) \right] \end{aligned}$$

I would optimize this using gradient descent or another iterative method like the Newton-Raphson method since there is no closed form solution to this maximization problem.

b) Note  $p(\theta) = (2\pi)^{-p/2} |\sigma^2 I|^{-1/2} \exp \left( -\frac{1}{2} x^T (\sigma^2 I)^{-1} x \right) = (2\pi)^{-p/2} |\sigma^2 I|^{-1/2} \exp \left( -\frac{\|x\|^2}{2\sigma^2} \right)$

$$p(\mathcal{D}|\theta) = L(\theta|x, y) = \prod_{i=1}^N \left[ \left( \frac{1}{1 + \exp(-x^{(i)T} \theta)} \right)^{y_i} \left( \frac{\exp(-x^{(i)T} \theta)}{1 + \exp(-x^{(i)T} \theta)} \right)^{1-y_i} \right]$$

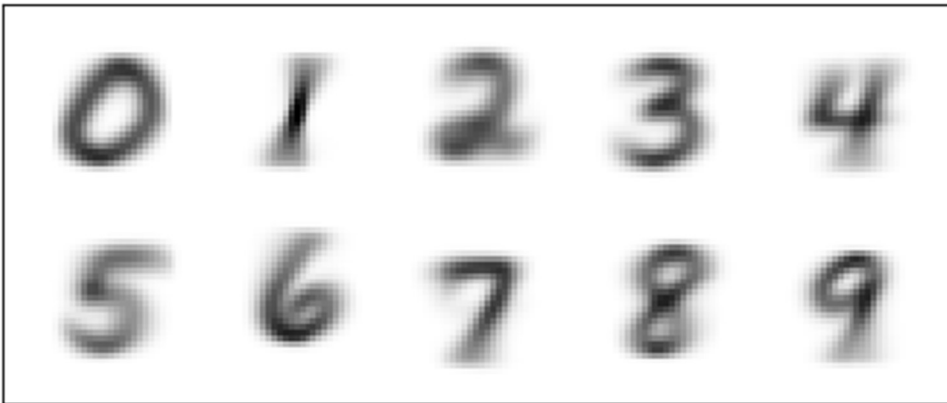
$$p(\theta|\mathcal{D}) \propto p(\theta) p(\mathcal{D}|\theta) \quad \text{so}$$

$$\log p(\theta|x, y) \propto \log p(\theta) + \log p(\mathcal{D}|\theta)$$

$$= -\frac{p}{2} \log 2\pi - \frac{1}{2} \log |\sigma^2 I| - \frac{1}{2\sigma^2} \|x\|^2 + \sum_{i=1}^N \left[ y_i x_i^{(i)T} \theta + \log \left( \frac{\exp(-x^{(i)T} \theta)}{1 + \exp(-x^{(i)T} \theta)} \right) \right] \quad (\text{from part a})$$

Question 2.

(d) The plot of the MLE estimator is as follows:



(f)

Average log-likelihood for MAP is -3.3570631378602918  
Training accuracy for MAP is 0.8352166666666667  
Test accuracy for MAP is 0.816

(g) The plot of the MAP estimator is as follows:

