Approximating MLEs with Variational Inference and Monte Carlo EM

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Setting

Data: (Yo, Ym)

Latent Parameters: θ

Example: Poisson Generalized Linear Mixed Model

Latent Variables: β , σ^2 , τ^2

$$Y_{it} \sim \operatorname{Pois}(\lambda_{it})$$

$$\log \lambda_{it} = X_{it}\boldsymbol{\beta} + \gamma_i + \epsilon_{it}$$

$$\epsilon_{it} \stackrel{\text{iid}}{\sim} N(0, \sigma^2) \quad \gamma_i \stackrel{\text{iid}}{\sim} N(0, \tau^2) \quad \boldsymbol{\epsilon} \perp \boldsymbol{\gamma}$$

Observed Data	Missing Data
X11, Y11	γ1
X ₁₂ , Y ₁₂	γ1
X ₂₁ , Y ₂₁	γ2
X ₂₂ , Y ₂₂	γ2

Problem

Goal: Maximize (intractable) observed likelihood

$$\mathscr{L}(\boldsymbol{\theta}|\mathbf{Y}_o) = p(\mathbf{Y}_o|\boldsymbol{\theta}) = \int p(\mathbf{Y}_o, \mathbf{Y}_m|\boldsymbol{\theta}) d\mathbf{Y}_m$$

We explore two distinct ways to do this:

Monte Carlo Expectation Maximization: Iteratively converge to "viable" estimate of latent variables by approximating a surrogate function using MCMC, then maximizing it

Variational Inference (Frequentist): Simultaneously estimate the latent variables and approximate the missing data by picking a distribution from a family of "friendly" distributions that, together, maximize a lower bound on the observed likelihood

EM - A Review

Goal: Maximize

$$\mathscr{L}(\boldsymbol{\theta}|\mathbf{Y}_o) = p(\mathbf{Y}_o|\boldsymbol{\theta}) = \int p(\mathbf{Y}_o, \mathbf{Y}_m|\boldsymbol{\theta}) d\mathbf{Y}_m$$

Method: Iteratively

(E) Compute:

$$Q\left(\boldsymbol{\theta} \mid \boldsymbol{\theta}^{(t)}\right) = \mathrm{E}_{\boldsymbol{\theta}^{(t)}}[\log p(\mathbf{Y}_o, \mathbf{Y}_m; \boldsymbol{\theta} \mid \mathbf{Y}_o, \boldsymbol{\theta}^{(t)})]$$

(M) Maximize with respect to θ

Notes:

- $\theta^{(t)}$ is only used to integrate out the missing data \mathbf{Y} This method has ascent property

Issue: what if this is intractable too?

MCEM: Basic Version

Idea: Approximate the expectation using MCMC

(E): Markov Chain generates samples

$$\mathbf{Y}_m^{(t,j)} \sim p(\mathbf{Y}_m | \mathbf{Y}_o, \boldsymbol{\theta}^{(t)})$$

Then, compute

$$Q\left(\boldsymbol{\theta} \mid \boldsymbol{\theta}^{(t)}\right) = \mathrm{E}_{\boldsymbol{\theta}^{(t)}}[\log p(\mathbf{Y}_o, \mathbf{Y}_m; \boldsymbol{\theta} \mid \mathbf{Y}_o, \boldsymbol{\theta}^{(t)})] \approx \frac{\sum_{j=1}^{m_t} \log p(\mathbf{Y}_o, \mathbf{Y}_m^{(t,j)}; \boldsymbol{\theta})}{m_t}$$

- Does not inherently preserve ascent property, more detailed methods have been developed (Ascent Based MCEM, Caffo et al, 2005)
 Achieved by computing and controlling asymptotic lower bound for

$$Q\left(\boldsymbol{\theta}^{(t+1)}|\boldsymbol{\theta}^{(t)}\right) - Q\left(\boldsymbol{\theta}^{(t)}|\boldsymbol{\theta}^{(t)}\right)$$

Variational Inference

- Uses Importance sampling and Jensen's inequality to maximize the bound on how badly the log likelihood can be approximated (Evidence Lower BOund)
- $q(\cdot; \boldsymbol{\omega})$ is from a family of distributions \boldsymbol{D} meant to approximate the missing data

$$p(\mathbf{Y}_{o}; \boldsymbol{\theta}) = \int p(\mathbf{Y}_{o}, \mathbf{Y}_{m}; \boldsymbol{\theta}) d\mathbf{Y}_{m}$$

$$= \int \frac{p(\mathbf{Y}_{o}, \mathbf{Y}_{m}; \boldsymbol{\theta})}{q(\mathbf{Y}_{m}; \boldsymbol{\omega})} q(\mathbf{Y}_{m}; \boldsymbol{\omega}) d\mathbf{Y}_{m}$$

$$= \mathbb{E}_{\mathbf{Y}_{m} \sim q(\cdot; \boldsymbol{\omega})} \left(\frac{p(\mathbf{Y}_{o}, \mathbf{Y}_{m}; \boldsymbol{\theta})}{q(\mathbf{Y}_{m}; \boldsymbol{\omega})} \right)$$

Variational Inference

- Uses Importance sampling and Jensen's inequality to maximize the bound on how badly the log likelihood can be approximated (Evidence Lower BOund)
- q is a family of distributions meant to approximate missing data and to have computation-friendly densities

$$\ell(\boldsymbol{\theta} \mid \mathbf{Y}_{o}) = \log p(\mathbf{Y}_{o}; \boldsymbol{\theta})$$

$$= \log \mathbb{E}_{\mathbf{Y}_{m} \sim q(\cdot; \boldsymbol{\omega})} \left(\frac{p(\mathbf{Y}_{o}, \mathbf{Y}_{m}; \boldsymbol{\theta})}{q(\mathbf{Y}_{m}; \boldsymbol{\omega})} \right)$$

$$\geq \mathbb{E}_{\mathbf{Y}_{m} \sim q(\cdot; \boldsymbol{\omega})} \left(\log \frac{p(\mathbf{Y}_{o}, \mathbf{Y}_{m}; \boldsymbol{\theta})}{q(\mathbf{Y}_{m}; \boldsymbol{\omega})} \right)$$

$$= \mathbb{E}_{\mathbf{Y}_{m} \sim q(\cdot; \boldsymbol{\omega})} \log p(\mathbf{Y}_{o}, \mathbf{Y}_{m}; \boldsymbol{\theta}) - \mathbb{E}_{\mathbf{Y}_{m} \sim q(\cdot; \boldsymbol{\omega})} (\log q(\mathbf{Y}_{m}; \boldsymbol{\omega}))$$

$$= \text{ELBO}(\boldsymbol{\omega}, \boldsymbol{\theta} \mid \mathbf{Y}_{o}).$$

Future Plans: Simulation Study

Apply Variational Inference and basic MCEM to Poisson GLMM

$$Y_{it} \sim \operatorname{Pois}(\lambda_{it})$$

$$\log \lambda_{it} = X_{it}\boldsymbol{\beta} + \gamma_i + \epsilon_{it}$$

$$\epsilon_{it} \stackrel{\text{iid}}{\sim} N(0, \sigma^2) \quad \gamma_i \stackrel{\text{iid}}{\sim} N(0, \tau^2) \quad \boldsymbol{\epsilon} \perp \boldsymbol{\gamma}$$

- X being Gaussian noise
- Balanced design
- Testing under 2 random intercept variances * 2 numbers of random intercept groups
- Metrics: MSE of parameter estimates, runtime