# Car Jacking MPLS - Tract

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## Spatial Data

```
#MN tracts
tracts <- get_acs(geography = "tract",</pre>
                  state = "MN",
                variables = "B01001_001E",
                output = "wide",
                survey = "acs5",
                year = 2020,
                geometry = T)
#Minneapolis Shapefile
mpls <- st_read("Data/mpls_city-shp/16cdbbfa-ad10-493c-afaf-52b61f2e76e42020329-1-180h9ap.whbo.shp") %>
   st_transform(st_crs(tracts))
## Reading layer `16cdbbfa-ad10-493c-afaf-52b61f2e76e42020329-1-180h9ap.whbo' from data source `C:\User
## using driver `ESRI Shapefile'
\#\# Simple feature collection with 1 feature and 4 fields
## Geometry type: POLYGON
## Dimension:
## Bounding box: xmin: -93.32911 ymin: 44.89059 xmax: -93.19433 ymax: 45.05125
## Geodetic CRS: WGS 84
mpls_tract <- tracts %>%
  st_filter(mpls, .predicate = st_intersects) %>%
  mutate(GEOID = as.numeric(GEOID),
         tract_area = as.numeric(st_area(.)),
         tract_area_sqkm = tract_area*.000001,
         tract_area_sqmi = tract_area_sqkm*.386102,
         intersection_area = as.numeric(st_area(st_intersection(., mpls))),
         perc_intersection = intersection_area/tract_area*100) %>%
  filter(perc_intersection >= 2) %>%
  select(-"B01001_001M")
```

### **ACS** Covariates

```
acs_17 <- get_acs(
  geography = "tract",
  variables = c("B01001_001E"),
  year = 2017,</pre>
```

```
state = "MN",
 county = "Hennepin",
  geometry = TRUE) %>%
 select(estimate)
## Getting data from the 2013-2017 5-year ACS
## Downloading feature geometry from the Census website. To cache shapefiles for use in future session
##
acs_18 <- get_acs(</pre>
 geography = "tract",
 variables = c("B01001_001E"),
 year = 2018,
 state = "MN",
 county = "Hennepin",
  geometry = TRUE
) %>%
select(estimate)
## Getting data from the 2014-2018 5-year ACS
## Downloading feature geometry from the Census website. To cache shapefiles for use in future session
##
acs_19 <- get_acs(</pre>
 geography = "tract",
 variables = c("B01001_001E"),
 year = 2019,
 state = "MN",
 county = "Hennepin",
  geometry = TRUE
) %>%
select(estimate)
## Getting data from the 2015-2019 5-year ACS
## Downloading feature geometry from the Census website. To cache shapefiles for use in future session
acs_20 <- get_acs(</pre>
 geography = "tract",
 variables = c("B01001 001E"),
 year = 2020,
 state = "MN",
 county = "Hennepin",
 geometry = TRUE
## Getting data from the 2016-2020 5-year ACS
## Downloading feature geometry from the Census website. To cache shapefiles for use in future session
library(tigris)
## To enable caching of data, set `options(tigris_use_cache = TRUE)`
## in your R script or .Rprofile.
```

```
hennepin_blocks <- blocks(
  "MN",
  "Hennepin",
  year = 2020
##
#2017 ACS interpolation
acs_1720 <- interpolate_pw(</pre>
 from = acs_17,
 to = acs_20,
 to_id = "GEOID",
  weights = hennepin_blocks,
  weight_column = "POP20",
  crs = 26993,
  extensive = TRUE) %>%
  mutate(year = 2017)
#2018 ACS interpolation
acs_1820 <- interpolate_pw(</pre>
 from = acs_18,
 to = acs_20,
 to_id = "GEOID",
  weights = hennepin_blocks,
  weight_column = "POP20",
  crs = 26993,
  extensive = TRUE) %>%
  mutate(year = 2018)
#2019 ACS interpolation
acs_1920 <- interpolate_pw(</pre>
 from = acs_19,
  to = acs_20,
  to_id = "GEOID",
  weights = hennepin_blocks,
  weight_column = "POP20",
  crs = 26993,
  extensive = TRUE) %>%
  mutate(year = 2019)
pop_denoms <- acs_20 %>%
  st_transform(crs = 26993) %>%
  mutate(year = 2020) \%
  select(-moe, -variable, -NAME) %>%
  rbind(acs_1720, acs_1820, acs_1920) %>%
  filter(GEOID %in% mpls_tract$GEOID)
#2021+2022: LOCF
pop_denom_21 <- pop_denoms \%>%
  filter(year==2020) %>%
  select(GEOID, year, estimate) %>%
```

```
mutate(year = 2021)
pop_denom_22 <- pop_denoms %>%
  filter(year==2020) %>%
  select(GEOID, year, estimate) %>%
  mutate(year = 2022)
pop denom locf <- pop denoms %>%
  rbind(pop_denom_21, pop_denom_22) %>%
  rename(total pop = estimate) %>%
  mutate(GEOID = as.numeric(GEOID)) %>%
  st drop geometry()
#ACS 2020 L-2 covariates
acs_2020 <- get_acs(geography = "tract",</pre>
            state = "MN",
               variables = c("B01001_001E", "B03002_003E", "B03002_004E", "B03002_005E",
                             "B03002_006E", "B03002_007E", "B03002_008E", "B03002_009E",
                             "B03002_010E", "B03002_011E", "B03002_012E", "B23025_002E",
                             "B23025_005E", "B17001_002E", "B19057_002E", "B11003_015E",
                             "B06009_002E", "B06009_005E", "C24010_001E", "C24010_003E",
                             "C24010_039E", "B11001_003E", "B01001_002E", "B05001_006E",
                             "B01001_003E", "B01001_004E", "B01001_005E", "B01001_006E",
                             "B01001_007E", "B01001_008E", "B01001_009E", "B01001_010E",
                             "B01001_011E", "B01001_012E", "B01001_013E", "B01001_014E",
                             "B01001_015E", "B01001_016E", "B01001_017E", "B01001_018E",
                             "B01001_019E", "B01001_020E", "B01001_021E", "B01001_022E",
                             "B01001_023E", "B01001_024E", "B01001_025E", "B01001_027E",
                             "B01001_028E", "B01001_029E", "B01001_030E", "B01001_031E",
                             "B01001_032E", "B01001_033E", "B01001_034E", "B01001_035E",
                             "B01001_036E", "B01001_037E", "B01001_038E", "B01001_039E",
                             "B01001_040E", "B01001_041E", "B01001_042E", "B01001_043E",
                             "B01001_044E", "B01001_045E", "B01001_046E", "B01001_047E",
                             "B01001_048E", "B01001_049E", "B07001_017E", "B25003_002E",
                             "B05002_013E", "B19013_001E"),
               output = "wide",
               survey = "acs5",
               vear = 2020) %>%
  select(-ends_with("M", ignore.case = F)) %>%
  rename(total_pop = B01001_001E, white_pop = B03002_003E, black_pop = B03002_004E,
         na_pop = B03002_005E, asian_pop = B03002_006E, hpi_pop = B03002_007E,
         other_pop = B03002_008E, biracial_pop = B03002_009E, biracial_other_pop = B03002_010E,
         biracial three pop = B03002 011E, hisp pop = B03002 012E, total ilf = B23025 002E,
         unemp = B23025_005E, povlevel = B17001_002E, pub_assist = B19057_002E,
         female_hh = B11003_015E, no_hs_dip = B06009_002E, bach_degree = B06009_005E,
         total employed = C24010 001E, employed mbsa male = C24010 003E,
         employed_mbsa_female = C24010_039E, mar_fam = B11001_003E, male = B01001_002E,
         noncitizen = B05001_006E,
        age_m_5_under = B01001_003E, age_m_5_9 = B01001_004E, age_m_10_14 = B01001_005E,
        age_m_15_17 = B01001_006E, age_m_18_19 = B01001_007E, age_m_20 = B01001_008E,
        age_m_21 = B01001_009E, age_m_22_24 = B01001_010E, age_m_25_29 = B01001_011E,
        age_m_30_34 = B01001_012E, age_m_35_39 = B01001_013E, age_m_40_44 = B01001_014E,
        age m 45 49 = B01001_015E, age m 50 54 = B01001_016E, age m 55 59 = B01001_017E,
```

```
age m 60 61 = B01001_018E, age m 62 64 = B01001_019E, age m 65 66 = B01001_020E,
      age_m_67_69 = B01001_021E, age_m_70_74 = B01001_022E, age_m_75_79 = B01001_023E,
      age m 80 84 = B01001 024E, age m 85 plus = B01001 025E, age f 5 under = B01001 027E,
      age f 5 9 = B01001 028E, age f 10 14 = B01001 029E, age f 15 17 = B01001 030E,
      age f 18 19 = B01001 031E, age f 20 = B01001 032E, age f 21 = B01001 033E,
      age f 22 24 = B01001 034E, age f 25 29 = B01001 035E, age f 30 34 = B01001 036E,
      age_f_35_39 = B01001_037E, age_f_40_44 = B01001_038E, age_f_45_49 = B01001_039E,
      age_f_50_54 = B01001_040E, age_f_55_59 = B01001_041E, age_f_60_61 = B01001_042E,
      age f 62 64 = B01001 043E, age f 65 66 = B01001 044E, age f 67 69 = B01001 045E,
      age_f_70_74 = B01001_046E, age_f_75_79 = B01001_047E, age_f_80_84 = B01001_048E,
      age_f_85_plus = B01001_049E, res_mob = B07001_017E,
      own_hh = B25003_002E, foreign = B05002_013E,
     med_hh_inc = B19013_001E) %>%
mutate(white_prop = white_pop/total_pop,
       black_prop = black_pop/total_pop,
       na_prop = na_pop/total_pop,
       asian_prop = asian_pop/total_pop,
       hpi_prop = hpi_pop/total_pop,
       other_prop = other_pop/total_pop,
       biracial_prop = (biracial_pop+biracial_other_pop+biracial_three_pop)/total_pop,
       hisp prop = hisp pop/total pop,
       white perc = 100*white pop/total pop,
       black_perc = 100*black_pop/total_pop,
       na_perc = 100*na_pop/total_pop,
       asian_perc = 100*asian_pop/total_pop,
       hpi perc = 100*hpi pop/total pop,
       other_perc = 100*other_pop/total_pop,
       biracial_perc = 100*(biracial_pop+biracial_other_pop+biracial_three_pop)/total_pop,
       hisp_perc = 100*hisp_pop/total_pop,
       unemp_rate = 100*unemp/total_ilf,
       pov_rate = 100*povlevel/total_pop,
       pub_assist_rate = 100*pub_assist/total_pop,
       female hh rate = 100*female hh/total pop,
       no_hs_dip_rate = 100*no_hs_dip/total_pop,
       bach degree rate = 100*bach degree/total pop,
       employed_mbsa = employed_mbsa_male+employed_mbsa_female,
       employed_mbsa_rate = 100*employed_mbsa/total_employed,
       mar fam rate = 100*mar fam/total pop,
       male rate = 100*male/total pop,
       noncitizen rate = 100*noncitizen/total pop,
       race_eth_hetero = 1-(white_prop^2+black_prop^2+na_prop^2+asian_prop^2+
                     hpi_prop^2+other_prop^2+other_prop^2+biracial_prop^2+hisp_prop^2),
       age_below_18_perc = 100*(age_m_5_under+age_f_5_under+age_m_5_9+
                                age_f_5_9+age_m_10_14+age_f_10_14+age_m_15_17+
                                age_f_15_17)/total_pop,
       age 19 29 perc = 100*(age_m_18_19+age_f_18_19+age_m_20+age_f_20+age_m_21+age_f_21+
                        age_m_22_24+age_f_22_24+age_m_25_29+age_f_25_29)/total_pop,
       age_30_49_perc = 100*(age_m_30_34+age_f_30_34+age_m_35_39+age_f_35_39+
                             age_m_40_44+age_f_40_44+age_m_45_49+age_f_45_49)/total_pop,
       age_50_69_perc = 100*(age_m_50_54+age_f_50_54+age_m_55_59+age_f_55_59+
                             age_m_60_61+age_f_60_61+age_m_62_64+age_f_62_64+
                             age_m_65_66+age_f_65_66+age_m_67_69+age_f_67_69)/total_pop,
       age_70_plus_perc = 100*(age_m_70_74+age_f_70_74+age_m_75_79+age_f_75_79+
```

```
age_m_80_84+age_f_80_84+age_m_85_plus+age_f_85_plus)/total_pop,
         res_mob_rate = 100-100*res_mob/total_pop,
         own_hh_rate = 100*own_hh/total_pop,
         foreign_rate = 100*foreign/total_pop)
## Getting data from the 2016-2020 5-year ACS
# 2020 Census 18+ Denominator
dc2020 <- get decennial(
  geography = "tract",
  variables = c("P3_001N"),
 year = 2020,
  state = "MN".
  county = "Hennepin",
  geometry = F) %>%
  mutate(GEOID = is.numeric(GEOID)) %>%
  rename(total_pop = value) %>%
  select(-GEOID)
## Getting data from the 2020 decennial Census
## Using the PL 94-171 Redistricting Data summary file
## Note: 2020 decennial Census data use differential privacy, a technique that
## introduces errors into data to preserve respondent confidentiality.
## i Small counts should be interpreted with caution.
## i See https://www.census.gov/library/fact-sheets/2021/protecting-the-confidentiality-of-the-2020-cen
## This message is displayed once per session.
```

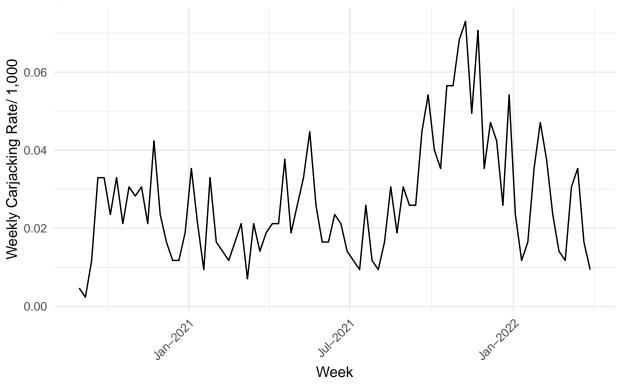
## Open Minneapolis Carjacking Data

```
#open minneapolis crime data 2019-4/7 (date of download)
cj_spatial <- read_csv("Car Jacking/crime_data.csv") %>%
  filter(Offense=="Carjacking - Subset of Robbery") %>% #filter carjackings
  mutate(date=ymd_hms(Occurred_Date),
        year=isoyear(date),
         week=isoweek(date)) %>%
  select(OBJECTID, year, week, Latitude, Longitude) %>%
  st_as_sf(coords = c("Longitude", "Latitude"), crs = "NAD83", remove=F) %>%
  st_join(mpls_tract) %>% #spatial join neighborhoods
  st_drop_geometry() %>%
  filter(!is.na(GEOID)) %>%
  group_by(year, week, GEOID, .drop=F) %>%
  tally(name = "car_jack") %>%
  ungroup() %>%
  complete(year, week, GEOID=mpls_tract$GEOID, fill = list(car_jack = 0)) %>%
  filter(!(year==2022 & week >= 14) & #removing unobserved/redundant completions
           !(year==2021 & week==53) & #removing unobserved/redundant completions
           !(year==2020 & week < 36)) %>% #removing weeks before Sept. 2020, isoweek 36
  arrange(GEOID, year, week) %>%
  left join(mpls tract, by = "GEOID") %>%
  left_join(pop_denom_locf, by = c("GEOID", "year")) %>%
  mutate(car_jack_rate = car_jack/total_pop*1000) %>%
  st_as_sf()
```

## MPLS Carjackings by Week

```
#aggregate to week over tracts
cj_week <- cj_spatial %>%
  group_by(year, week) %>%
  summarize(car_jack = sum(car_jack, na.rm = T),
            total_pop = sum(total_pop, na.rm = T)) %>%
  mutate(begin_date = ISOweek2date(paste(year, pasteO("W", sprintf("%02d", week)), 1,sep = "-")),
         end_date = begin_date+weeks(1)-days(1),
         car_jack_rate = car_jack/total_pop*1000)
## `summarise()` has grouped output by 'year'. You can override using the
## `.groups` argument.
ggplot(cj_week)+
  geom_line(aes(x=begin_date, y=car_jack_rate))+
  scale_x_date(date_labels = "%b-%Y",
              limits = c(min(cj_week$begin_date), max(cj_week$begin_date)))+
  labs(title = "Figure X: Weekly Minneapolis Carjackings, 8/31/2020-4/7/2022",
      subtitle = "openminneapolis Crime Data",
      x = "Week",
      y = "Weekly Carjacking Rate/ 1,000")+
  theme_minimal()+
    theme(axis.text.x=element_text(angle=45, hjust=1))
```

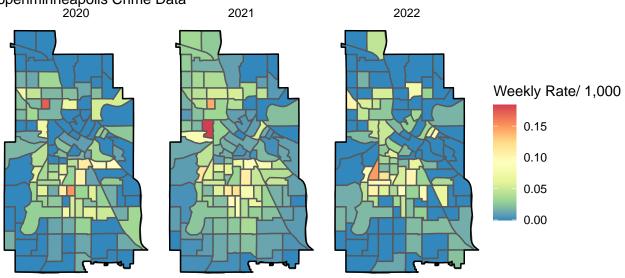
Figure X: Weekly Minneapolis Carjackings, 8/31/2020–4/7/2022 openminneapolis Crime Data



## MPLS ZCTA Carjackings Map

```
#aggregate to neighborhood-year level
cj_tract_year <- cj_spatial %>%
 group_by(GEOID, year) %>%
  summarize(car_jack = sum(car_jack, na.rm = T),
            total_pop = sum(B01001_001E, na.rm = T),
            car_jack_rate = car_jack/total_pop*1000)
## `summarise()` has grouped output by 'GEOID'. You can override using the
## `.groups` argument.
ggplot() +
 geom_sf(data = cj_tract_year, aes(geometry = geometry, fill = car_jack_rate)) +
  geom_sf(data = mpls, aes(geometry = geometry), color = "black", alpha = 0)+
 facet_wrap(~year)+
  scale_fill_distiller(palette = "Spectral")+
  labs(title = "Figure X: Minneapolis Carjacking Rates by Tract and Year",
      subtitle = "openminneapolis Crime Data",
      fill = "Weekly Rate/ 1,000")+
  theme_void()
```

Figure X: Minneapolis Carjacking Rates by Tract and Year openminneapolis Crime Data



Expanded MPLS Carjacking (Crime Incidents) Data

```
cj_exp <- read_csv("Data/MPDdata_082422.csv") %>%
   mutate(date=mdy_hm(reporteddate),
         year=isoyear(date),
         week=isoweek(date)) %>%
  select(CaseNumber, year, week, latitude, longitude) %>%
  distinct(CaseNumber, .keep_all = TRUE) %>%
  drop_na(latitude, longitude) %>%
  st as sf(coords = c("longitude", "latitude"), crs = "NAD83", remove=F) %%
  st_join(mpls_tract) %>% #spatial join neighborhoods
  st_drop_geometry() %>%
  drop_na(GEOID) %>%
  group_by(year, week, GEOID, .drop=F) %>%
  tally(name = "car_jack") %>%
  ungroup() %>%
  complete(year, week, GEOID=mpls_tract$GEOID, fill = list(car_jack = 0)) %>%
  filter(!(year==2021 & week==53)) %>%
  arrange(GEOID, year, week) %>%
  left_join(mpls_tract, by = "GEOID") %>%
  left_join(dc2020, by = c("NAME")) %>%
  mutate(car_jack_rate = car_jack/total_pop*1000) %>%
  st_as_sf()
## Rows: 3894 Columns: 28
## -- Column specification --
## Delimiter: ","
## chr (24): CaseNumber, dataset, closurecode, closurecode_MPD, reporteddate, c...
## dbl (4): precinct, latitude, longitude, age
## i Use `spec()` to retrieve the full column specification for this data.
## i Specify the column types or set `show_col_types = FALSE` to quiet this message.
```

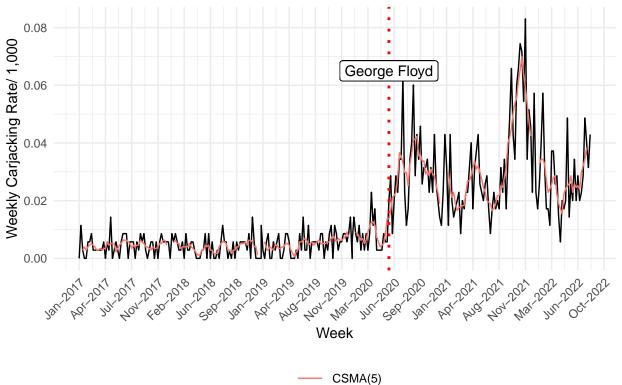
## MPLS Carjackings by Week - MPD Extended Data

```
#aggregate to week over tracts
cj_exp_week <- cj_exp %>%
 group_by(year, week) %>%
  summarize(car_jack = sum(car_jack, na.rm = T),
            total_pop = sum(total_pop, na.rm = T)) %>%
  mutate(begin_date = ISOweek2date(paste(year, pasteO("W", sprintf("%02d", week)), 1,sep = "-")),
         end_date = begin_date+weeks(1)-days(1),
         car_jack_rate = car_jack/total_pop*1000,
         pre_post_floyd = ifelse(end_date <= as.Date("2020-05-25"), 0, 1)) %%</pre>
  filter(end_date <= as.Date("2022-08-20")) %>%
  mutate(csma = forecast::ma(car_jack_rate, order=5, centre=TRUE),
         tsma = TTR::SMA(car_jack_rate, n=5))
## `summarise()` has grouped output by 'year'. You can override using the
## `.groups` argument.
## Registered S3 method overwritten by 'quantmod': method from as.zoo.data.frame
## zoo
pre_mean <- mean(cj_exp_week$car_jack_rate[cj_exp_week$pre_post_floyd==0], na.rm = T)</pre>
post_mean <- mean(cj_exp_week$car_jack_rate[cj_exp_week$pre_post_floyd==1], na.rm = T)</pre>
```

```
c(pre_mean, post_mean)
## [1] 0.004835884 0.030361407
post_mean/pre_mean
## [1] 6.278358
ggplot(cj_exp_week)+
  geom_line(aes(x=begin_date, y=car_jack_rate))+
  scale_x_date(date_labels = "%b-%Y", date_breaks = "15 weeks",
               limits = c(min(cj_exp_week$begin_date), max(cj_exp_week$begin_date)))+
  geom_vline(xintercept=cj_exp_week$begin_date[cj_exp_week$year==2020 &
                                                 cj_exp_week$week==isoweek(date("2020-05-25"))],
              linetype="dotted", color="red", size=1)+
  geom_label(aes(x=cj_exp_week$begin_date[cj_exp_week$year==2020 &
                                             cj_exp_week$week==isoweek(date("2020-05-25"))],
                 y=0.065),
            label = "George Floyd", show.legend = FALSE)+
  labs(title = "Figure 1: Weekly Minneapolis Carjackings, 1/1/2017-8/20/2022",
      x = "Week",
       y = "Weekly Carjacking Rate/ 1,000",
       color = NULL)+
  geom_line(aes(x=begin_date, y=csma, color = "CSMA(5)"))+
  theme_minimal()+
    theme(axis.text.x=element_text(angle=45, hjust=1)) +
  theme(legend.key.size = unit(0.8, "cm"),legend.position = "bottom")
```

## Warning: Removed 4 row(s) containing missing values (geom\_path).



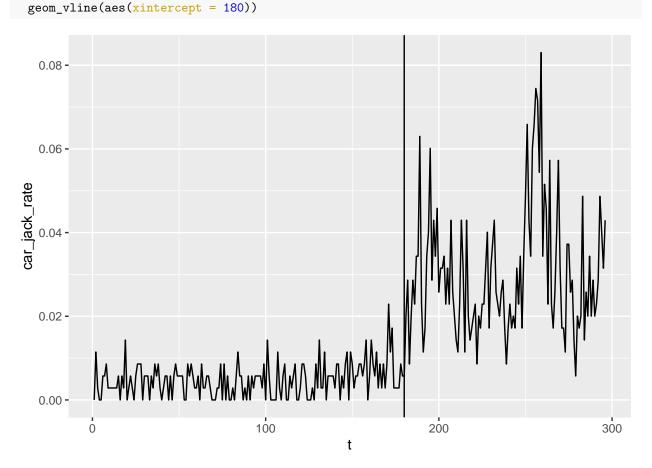


```
ggsave(filename = "Car Jacking/Figures for PAA/fig1.png", bg="white", width = 10, height = 8)
## Warning: Removed 4 row(s) containing missing values (geom_path).
```

# Structural Change in Carjacking

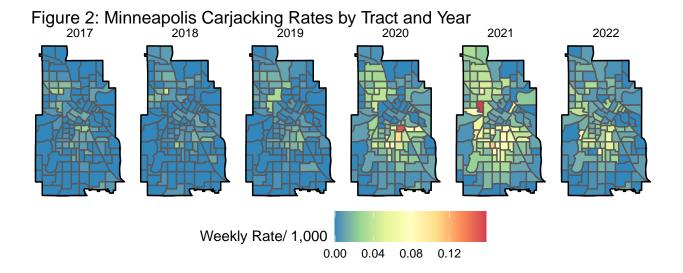
```
library(strucchange)
## Loading required package: sandwich
##
## Attaching package: 'strucchange'
## The following object is masked from 'package:stringr':
##
       boundary
##
cj_exp_week <- cj_exp_week %>%
 ungroup() %>%
 mutate(t = row_number())
sctest(cj_exp_week$car_jack_rate~cj_exp_week$t,
       type = "Chow",
       point = 180)
##
##
   Chow test
```

```
##
## data: cj_exp_week$car_jack_rate ~ cj_exp_week$t
## F = 49.567, p-value < 2.2e-16
breakpoints(car_jack_rate~t,
            data = cj_exp_week,
            breaks = 1)
##
##
     Optimal 2-segment partition:
##
## Call:
## breakpoints.formula(formula = car_jack_rate ~ t, breaks = 1,
       data = cj_exp_week)
##
## Breakpoints at observation number:
## 180
##
## Corresponding to breakdates:
## 0.6081081
ggplot(cj_exp_week)+
  geom_line(aes(x=t, y=car_jack_rate))+
```



## MPLS ZCTA Carjackings Map - MPD Extended Data

```
#aggregate to neighborhood-year level
cj_exp_tract_year <- cj_exp %>%
 group_by(GEOID, year) %>%
  summarize(car jack = sum(car jack, na.rm = T),
            total_pop = sum(B01001_001E, na.rm = T),
            car_jack_rate = car_jack/total_pop*1000) %>%
 mutate(GEOID = as.character(GEOID))
## `summarise()` has grouped output by 'GEOID'. You can override using the
## `.groups` argument.
ggplot() +
  geom_sf(data = cj_exp_tract_year, aes(geometry = geometry, fill = car_jack_rate)) +
  geom_sf(data = mpls, aes(geometry = geometry), color = "black", alpha = 0)+
 facet grid(~year)+
  scale_fill_distiller(palette = "Spectral")+
  labs(title = "Figure 2: Minneapolis Carjacking Rates by Tract and Year",
      fill = "Weekly Rate/ 1,000")+
  theme_void()+
  theme(legend.key.size = unit(0.8, "cm"),legend.position = "bottom")
```



```
ggsave(filename = "Car Jacking/Figures for PAA/fig2.png", bg="white", width = 10, height = 8)
```

## MPLS Murder (Crime Incidents) Data

```
mpd_2016 <- read_csv("Data/Police_Incidents_2016.csv")</pre>
## Rows: 20155 Columns: 20
## -- Column specification -----
## Delimiter: ","
## chr (12): PublicAddress, CCN, Precinct, ReportedDate, BeginDate, Offense, D...
        (7): FID, ControlNbr, GBSID, Lat, Long, X, Y
## time (1): Time
## i Use `spec()` to retrieve the full column specification for this data.
## i Specify the column types or set `show_col_types = FALSE` to quiet this message.
mpd_2017 <- read_csv("Data/Police_Incidents_2017.csv")</pre>
## Rows: 22085 Columns: 20
## -- Column specification -----
## Delimiter: ","
## chr (12): PublicAddress, CCN, Precinct, ReportedDate, BeginDate, Offense, D...
        (7): FID, ControlNbr, GBSID, Lat, Long, X, Y
## time (1): Time
##
## i Use `spec()` to retrieve the full column specification for this data.
## i Specify the column types or set `show_col_types = FALSE` to quiet this message.
mpd_2018a <- read_csv("Data/Police_Incidents_2018.csv")</pre>
## Rows: 7350 Columns: 20
## Delimiter: ","
## chr (12): PublicAddress, CCN, Precinct, ReportedDate, BeginDate, Offense, D...
        (7): FID, ControlNbr, GBSID, Lat, Long, X, Y
## time (1): Time
## i Use `spec()` to retrieve the full column specification for this data.
## i Specify the column types or set `show_col_types = FALSE` to quiet this message.
mpd_2018b <- read_csv("Data/Police_Incidents_2018_PIMS.csv")</pre>
## Rows: 11603 Columns: 23
## -- Column specification -----
## Delimiter: ","
## chr (13): publicaddress, caseNumber, precinct, reportedDate, beginDate, repo...
## dbl (10): X, Y, reportedTime, beginTime, centergbsid, centerLong, centerLat,...
##
## i Use `spec()` to retrieve the full column specification for this data.
## i Specify the column types or set `show_col_types = FALSE` to quiet this message.
mpd_2019 <- read_csv("Data/Police_Incidents_2019.csv")</pre>
## Warning: One or more parsing issues, call `problems()` on your data frame for details,
## e.g.:
    dat <- vroom(...)</pre>
##
    problems(dat)
```

```
## Rows: 22934 Columns: 23
## -- Column specification --------
## Delimiter: ","
## chr (12): publicaddress, caseNumber, reportedDate, beginDate, reportedDateTi...
## dbl (11): X, Y, precinct, reportedTime, beginTime, centergbsid, centerLong, ...
## i Use `spec()` to retrieve the full column specification for this data.
## i Specify the column types or set `show_col_types = FALSE` to quiet this message.
mpd 2020 <- read csv("Data/Police Incidents 2020.csv")</pre>
## Warning: One or more parsing issues, call `problems()` on your data frame for details,
## e.g.:
    dat <- vroom(...)</pre>
##
    problems(dat)
##
## Rows: 24136 Columns: 23
## -- Column specification ------
## Delimiter: ","
## chr (12): publicaddress, caseNumber, reportedDate, beginDate, reportedDateTi...
## dbl (11): X, Y, precinct, reportedTime, beginTime, centergbsid, centerLong, ...
## i Use `spec()` to retrieve the full column specification for this data.
## i Specify the column types or set `show_col_types = FALSE` to quiet this message.
mpd_2021 <- read_csv("Data/Police_Incidents_2021.csv")</pre>
## Warning: One or more parsing issues, call `problems()` on your data frame for details,
## e.g.:
##
    dat <- vroom(...)</pre>
##
    problems(dat)
## Rows: 24755 Columns: 23
## -- Column specification -------
## Delimiter: ","
## chr (12): publicaddress, caseNumber, reportedDate, beginDate, reportedDateTi...
## dbl (11): X, Y, precinct, reportedTime, beginTime, centergbsid, centerLong, ...
## i Use `spec()` to retrieve the full column specification for this data.
## i Specify the column types or set `show_col_types = FALSE` to quiet this message.
mpd_2022 <- read_csv("Data/Police_Incidents_2022.csv")</pre>
## Warning: One or more parsing issues, call `problems()` on your data frame for details,
## e.g.:
    dat <- vroom(...)</pre>
    problems(dat)
## Rows: 19555 Columns: 23
## -- Column specification -------
## Delimiter: ","
## chr (12): publicaddress, caseNumber, reportedDate, beginDate, reportedDateTi...
## dbl (11): X, Y, precinct, reportedTime, beginTime, centergbsid, centerLong, ...
## i Use `spec()` to retrieve the full column specification for this data.
## i Specify the column types or set `show_col_types = FALSE` to quiet this message.
```

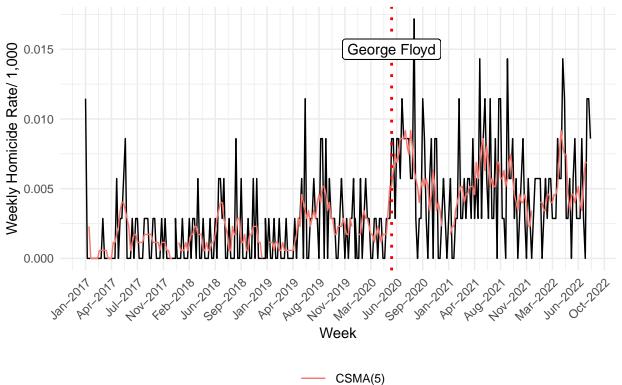
```
pre_pims_base <- mpd_2016 %>%
  rbind(mpd 2017) %>%
  rbind(mpd_2018a) %>%
  rename(reportedDate = ReportedDate,
         centerLong = Long,
         centerLat = Lat) %>%
  select(FID, centerLong, centerLat, Offense, reportedDate) %>%
  rename(OBJECTID = FID,
        X = centerLong,
         Y = centerLat.
         offense = Offense)
post_pims_base <- mpd_2018b %>%
  rbind(mpd_2019) %>%
  rbind(mpd_2020) %>%
  rbind(mpd_2021) %>%
  rbind(mpd_2022) %>%
  select(OBJECTID, X, Y, offense, reportedDate)
mpd <- pre_pims_base %>%
 rbind(post_pims_base)
#aggregate homicides to tract-week
homicide <- mpd %>%
 mutate(date=ymd hms(reportedDate),
         year=isoyear(date),
         week=isoweek(date)) %>%
  filter(offense=="MURDR" & year!=2016 & year!=2015) %>% #filter homicides
  select(OBJECTID, year, week, Y, X) %>%
  st_as_sf(coords = c("X", "Y"), crs = "NAD83", remove=F) %>%
  st_join(mpls_tract) %>% #spatial join neighborhoods
  st_drop_geometry() %>%
  filter(!is.na(GEOID)) %>%
  group_by(year, week, GEOID, .drop=F) %>%
  tally(name = "homicide") %>%
  ungroup() %>%
  complete(year, week, GEOID=mpls_tract$GEOID, fill = list(homicide = 0)) %>%
  filter(!(year==2021 & week==53)) %>%
  arrange(GEOID, year, week) %>%
  left_join(mpls_tract, by = "GEOID") %>%
  left_join(dc2020, by = c("NAME")) %>%
  mutate(homicide_rate = homicide/total_pop*1000) %>%
  st_as_sf()
```

## MPLS Murder by Week

```
end_date = begin_date+weeks(1)-days(1),
         homicide_rate = homicide/total_pop*1000,
         pre_post_floyd = ifelse(end_date <= as.Date("2020-05-25"), 0, 1)) %%</pre>
  filter(end_date <= as.Date("2022-08-20")) %>%
  mutate(csma = forecast::ma(homicide_rate, order=5,centre=TRUE),
         tsma = TTR::SMA(homicide_rate, n=5))
## `summarise()` has grouped output by 'year'. You can override using the
## `.groups` argument.
pre_mean <- mean(homicide_week$homicide_rate[homicide_week$pre_post_floyd==0], na.rm = T)</pre>
post_mean <- mean(homicide_week$homicide_rate[homicide_week$pre_post_floyd==1], na.rm = T)</pre>
c(pre_mean, post_mean)
## [1] 0.001749826 0.005331759
post_mean/pre_mean
## [1] 3.047022
ggplot(homicide_week)+
  geom_line(aes(x=begin_date, y=homicide_rate))+
 scale_x_date(date_labels = "%b-%Y", date_breaks = "15 weeks",
               limits = c(min(homicide week$begin date), max(homicide week$begin date)))+
     labs(title = "Figure 3: Weekly Minneapolis Homicide, 1/1/2017-8/20/2022",
       x = "Week",
       y = "Weekly Homicide Rate/ 1,000",
       color = NULL)+
  geom vline(xintercept=homicide week$begin date[homicide week$year==2020 &
                                                 homicide week$week==isoweek(date("2020-05-25"))],
              linetype="dotted", color="red", size=1)+
   geom_label(aes(x=homicide_week$begin_date[homicide_week$year==2020 &
                                             homicide_week$week==isoweek(date("2020-05-25"))],
                 y=0.015),
             label = "George Floyd", show.legend = FALSE)+
  theme_minimal()+
  geom_line(aes(x=begin_date, y=csma, color = "CSMA(5)"))+
  theme(axis.text.x=element_text(angle=45, hjust=1)) +
  theme(legend.key.size = unit(0.8, "cm"),legend.position = "bottom")
```

## Warning: Removed 4 row(s) containing missing values (geom\_path).





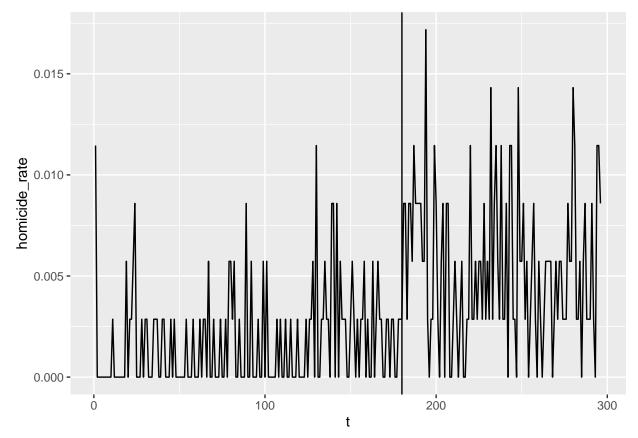
```
ggsave(filename = "Car Jacking/Figures for PAA/fig3.png", bg="white", width = 10, height = 8)
## Warning: Removed 4 row(s) containing missing values (geom_path).
```

# Structural Change in Carjacking

```
homicide_week <- homicide_week %>%
  ungroup() %>%
 mutate(t = row_number())
sctest(homicide_week$homicide_rate~homicide_week$t,
       type = "Chow",
       point = 180)
##
##
   Chow test
##
## data: homicide_week$homicide_rate ~ homicide_week$t
## F = 10.643, p-value = 3.454e-05
breakpoints(homicide_rate~t,
            data = homicide_week,
            breaks = 1)
##
##
     Optimal 2-segment partition:
```

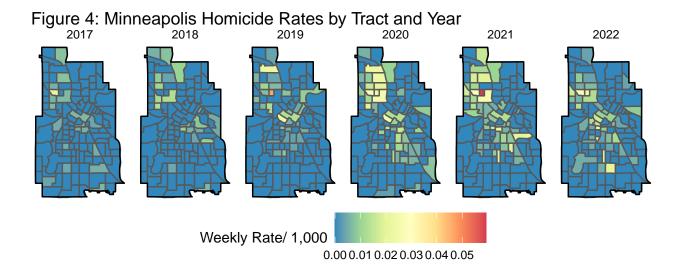
```
##
## Call:
## breakpoints.formula(formula = homicide_rate ~ t, breaks = 1,
## data = homicide_week)
##
## Breakpoints at observation number:
## 180
##
## Corresponding to breakdates:
## 0.6081081

ggplot(homicide_week)+
    geom_line(aes(x=t, y=homicide_rate))+
    geom_vline(aes(xintercept = 180))
```



## MPLS ZCTA Murder Map - MPD Extended Data

## `summarise()` has grouped output by 'GEOID'. You can override using the



```
ggsave(filename = "Car Jacking/Figures for PAA/fig4.png", bg="white", width = 10, height = 8)
```

# Dispersion of Change from 2017-2019 to 2020-2021

### Car Jacking

```
crimedispersion <- function
(data1, unitID, time1, time2, method = "match") {

# define variables to limit build warnings
adjusted <- Ut1 <- Ut2 <- Rt1 <- Rt2 <- chg <- pct <- NULL

# ERROR CHECKING. Has user passed a data frame?
if (!is.data.frame(data1)) {</pre>
```

```
stop("The input data specified is not a data.frame object. Please fix.")
}
# Build a local data.frame and populate with passed arguments
source_rows <- nrow(data1)</pre>
df1 <- data.frame(matrix(ncol = 3, nrow = source_rows))</pre>
colnames(df1) <- c("unit", "time1", "time2")</pre>
df1$unit <- data1[, unitID]</pre>
df1$time1 <- data1[, time1]</pre>
df1$time2 <- data1[, time2]</pre>
if (method == "remove") {
  analysisMethod <- "remove"</pre>
} else {
  analysisMethod <- "match"</pre>
# ERROR CHECKING. Did user pass numeric columns where needed?
try (df1$time1 <- as.numeric(df1$time1), silent = TRUE)</pre>
try (df1$time2 <- as.numeric(df1$time2), silent = TRUE)</pre>
if (!class(df1$time1)[1] == "numeric") {
  stop("The time1 field is not a numeric object. Please fix.")
if (!class(df1$time2)[1] == "numeric") {
  stop("The time2 field is not a numeric object. Please fix.")
# MORE ERROR CHECKING:
# What if the user has NA or missing data?
# What if the crime problem is decreasing?
# Fun tasks for later...
# Set up parameters -----
# Set up initial parameters
count Rt1 <- sum(df1$time1)</pre>
count_Rt2 <- sum(df1$time2)</pre>
chg_Rt1_Rt2 <- count_Rt2 - count_Rt1</pre>
pct_Rt1_Rt2 <- (chg_Rt1_Rt2 / count_Rt1) *100</pre>
# Add the field that has the volume of change, and order by it
df1 <- df1 %>%
  mutate (diff = time2 - time1) %>%
  mutate (diffPct = 100*(diff/time1)) %>%
  arrange(desc(diff))
# Grab some basic statistics here
numPositive <- length(which(df1$diff > 0))
numNeutral <- length(which(df1$diff == 0))</pre>
numNegative <- length(which(df1$diff < 0))</pre>
```

```
# Create the new data frame to hold the result
df2 <- data.frame(matrix(ncol =8, nrow = 0))</pre>
colnames(df2) <- c("unit", "adjusted", "Ut1", "Ut2", "Rt1", "Rt2", "chg", "pct")</pre>
df2 <- df2 %>%
 mutate(unit = as.character(unit)) %>%
 mutate(adjusted = as.numeric(adjusted)) %>%
 mutate(Ut1 = as.numeric(Ut1)) %>%
 mutate(Ut2 = as.numeric(Ut2)) %>%
 mutate(Rt1 = as.numeric(Rt1)) %>%
 mutate(Rt2 = as.numeric(Rt2)) %>%
 mutate(chg = as.numeric(chg)) %>%
 mutate(pct = as.numeric(pct))
# set up the initial row in the result data frame
df2 <- df2 %>% add_row(unit = "[ ALL AREAS ]", adjusted = 0,
                        Ut1 = 0, Ut2 = 0,
                        Rt1 = count_Rt1, Rt2 = count_Rt2,
                        chg = chg_Rt1_Rt2, pct = pct_Rt1_Rt2)
gain_from_row_removal <- row_to_remove <- NULL</pre>
# Loop through each row of the data
for (master_loop in 1:(source_rows)){
 df1 <- df1 %>% # order the data frame
    arrange(desc(diff))
 if (analysisMethod == "match"){
    #### 'Zero change the row' approach
    count_Rt1_temp <- count_Rt1</pre>
    count_Rt2_temp <- count_Rt2 - df1$diff[master_loop]</pre>
    pct_Rt1_Rt2 <- ((count_Rt1_temp - count_Rt2_temp) / count_Rt1) *100</pre>
  else { #analysisMethod == "remove"
    #### 'Remove entire row' approach, including remove t1 value
    count_Rt1_temp <- count_Rt1 - df1$time1[master_loop]</pre>
    count Rt2 temp <- count Rt2 - df1$time2[master loop]</pre>
    pct_Rt1_Rt2 <- ((count_Rt1_temp - count_Rt2_temp) / count_Rt1) *100</pre>
 }
 row_to_remove <- 1 # Always row 1, but this is a legacy from</pre>
  # when I used a different approach...
  # Here, the row we are removing is
  # stored in row_to_remove
  if (analysisMethod == "remove"){
    #### Remove entire row approach
         This approach removes the impact of the area by subtracting
         both Rt1 and Rt2
    count_Rt1 <- count_Rt1 - df1$time1[row_to_remove]</pre>
    count_Rt2 <- count_Rt2 - df1$time2[row_to_remove]</pre>
    chg_Rt1_Rt2 <- count_Rt2 - count_Rt1</pre>
```

```
pct_Rt1_Rt2 <- (chg_Rt1_Rt2 / count_Rt1) *100</pre>
   named_areas <- df1$unit[row_to_remove]</pre>
 }
  if (analysisMethod == "match"){
    #### Zero change the row approach, as if Rt2 == Rt1 in the row
    # The best row to remove is has been exhaustively calculated
    # Here, the row we are removing is stored in row_to_remove
    count Rt1 <- count Rt1</pre>
    count_Rt2 <- count_Rt2 - df1$diff[row_to_remove]</pre>
    chg_Rt1_Rt2 <- count_Rt2 - count_Rt1</pre>
   pct_Rt1_Rt2 <- (chg_Rt1_Rt2 / count_Rt1) *100</pre>
   named_areas <- df1$unit[row_to_remove]</pre>
  # Add result to the output data frame
  df2 <- df2 %>% add_row(unit = named_areas, adjusted = master_loop,
                         Ut1 = df1$time1[row_to_remove], Ut2 = df1$time2[row_to_remove],
                         Rt1 = count_Rt1, Rt2 = count_Rt2,
                         chg = chg_Rt1_Rt2, pct = pct_Rt1_Rt2)
  # Adjust the row we just used in one of two ways:
  # 1. remove the actual row entirely
 if (analysisMethod == "remove"){
   df1 <-df1[-c(row to remove), ]</pre>
 }
  #2. adjust the Rt2 to match Rt1 resulting in a zero diff
  # but show that diff as < lowest diff in the data set so that
    the program does not stall with too many zeros
  if (analysisMethod == "match"){
   df1$time2[row_to_remove] <- df1$time1[row_to_remove]</pre>
   df1$diff[row_to_remove] <- -999 # this should be changed to always less than
    # the lowest diff score in the data set
 }
} # end master_loop
# Calculate ODI and NCDI indices -----
NumContributed <- length(which(df2$chg > 0))
ODI <- NumContributed / source_rows</pre>
NCDI <- (numPositive - NumContributed) / source_rows</pre>
ODI.text <- paste("0.D.I. = ", format(ODI, digits = 3), "after \nadjusting",
                  NumContributed, "of the", source_rows, "units")
# Tidy up names for data frame -----
df2 <- df2 %>%
 rename(unit_t1 = Ut1, unit_t2 = Ut2, region_t1 = Rt1, region_t2 = Rt2)
df3 <- df2
```

```
plot.adjustment <- ""</pre>
  if (nrow(df3) > 151) {
   df3 <- df3[1:151, ]
   plot.adjustment <- "Plot only shows first\n100 areas adjusted"
  }
  p <- ggplot(df3, aes(x=reorder(unit, adjusted), y=pct, group = 1)) +
   geom line(color="#3277a8") +
    geom_point(shape=21, color="white", fill="#3277a8", size=2) +
   geom_hline(color="grey", yintercept=0) +
   labs(title="Dispersion of crime change",
         x ="Area adjusted", y = "Remaining crime change for region") +
    annotate(
      geom = "curve", x = NumContributed+4, y = 1.5,
     xend = NumContributed+1, yend = 0.2,
     curvature = .2, arrow = arrow(length = unit(2, "mm"))
    annotate(geom = "text", x = NumContributed+4.1, y = 1.5,
             label = ODI.text, hjust = "left") +
   theme(axis.text.x = element_text(angle = 90, vjust = 0.5, hjust=1)) +
    annotate(geom = "text", x = 2, y = df2$pct[1],
             label = pasteO(format(df2$pct[1], digits = 3),"% overall"), hjust = "left") +
    theme(axis.text.x = element_text(angle = 90, vjust = 0.5, hjust=1))
  if (plot.adjustment != "") {
   p <- p +
      annotate(geom = "text", x = 100, y = df3$pct[1]-1, label = plot.adjustment, hjust = "right")
  р
  # Create return list -----
  output <- list(df2, p, NumContributed, ODI, NCDI)</pre>
 return(output)
prepost_cj <- cj_exp %>%
  mutate(begin_date = ISOweek2date(paste(year, paste0("W", sprintf("%02d", week)), 1,sep = "-")),
         end_date = begin_date+weeks(1)-days(1),
         pre_post_floyd = ifelse(end_date <= as.Date("2020-05-25"), 0, 1)) %>%
  filter(end_date <= as.Date("2022-08-20")) %>%
  group_by(GEOID, pre_post_floyd) %>%
  summarize(car_jack = sum(car_jack, na.rm = T),
           total_pop = sum(total_pop, na.rm = T)) %>%
  mutate(car_jack_rate = car_jack/total_pop*1000) %>%
  select(GEOID, pre_post_floyd, car_jack, car_jack_rate) %>%
  st_drop_geometry() %>%
  pivot_wider(names_from = pre_post_floyd, values_from = c(car_jack, car_jack_rate)) %>%
```

```
mutate(GEOID = as.character(GEOID))

## `summarise()` has grouped output by 'GEOID'. You can override using the

## `.groups` argument.

output <- crimedispersion(as.data.frame(prepost_cj), 'GEOID', 'car_jack_rate_0', 'car_jack_rate_1')

ouput_data <- output[[1]]

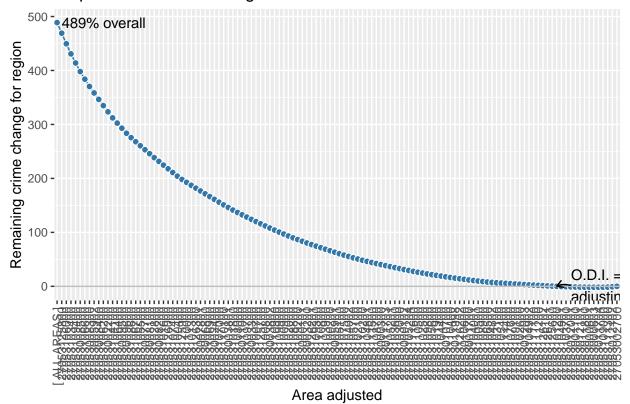
n_remove <- output[[3]]

odi <- output[[4]] #ratio of n removed to n overall

ncdi <- output[[5]] #ratio of areas not contributing to overall increase but still increase to overall

output[[2]]</pre>
```

### Dispersion of crime change



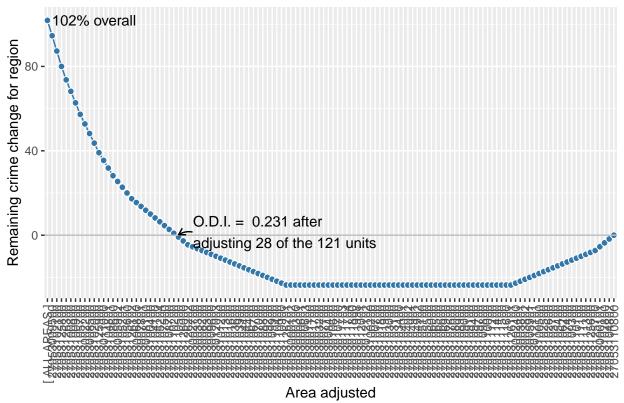
ggsave(filename = "Car Jacking/Figures for PAA/fig5.png", bg="white", width = 10, height = 8)

#### Homicide

```
st_drop_geometry() %>%
select(GEOID, pre_post_floyd, homicide, homicide_rate) %>%
pivot_wider(names_from = pre_post_floyd, values_from = c(homicide, homicide_rate)) %>%
mutate(GEOID = as.character(GEOID))

## `summarise()` has grouped output by 'GEOID'. You can override using the
## `.groups` argument.
output_homicide <- crimedispersion(as.data.frame(prepost_hom), 'GEOID', 'homicide_0', 'homicide_1')
ouput_data <- output_homicide[[1]]
n_remove <- output_homicide[[3]]
odi <- output_homicide[[4]] #ratio of n removed to n overall
ncdi <- output_homicide[[5]] #ratio of areas not contributing to overall increase but still increase to
output_homicide[[2]]</pre>
```

## Dispersion of crime change



ggsave(filename = "Car Jacking/Figures for PAA/fig6.png", bg="white", width = 10, height = 8)

## Spatial Correlation Change in Carjackings and Homicide

## Carjacking

```
library(sfdep)
cj_delta <- prepost_cj %>%
```

```
mutate(delta = car_jack_rate_1-car_jack_rate_0,
         GEOID = as.numeric(GEOID)) %>%
  left_join(mpls_tract, by = "GEOID") %>%
  st_as_sf()
nb <- st_contiguity(cj_delta, queen=TRUE)</pre>
wt <- st_weights(nb, style = "W")
global_moran_test(
 cj_delta$delta,
 nb,
 wt,
 alternative = "greater",
randomization = TRUE)
##
## Moran I test under randomisation
##
## data: x
## weights: listw
## Moran I statistic standard deviate = 8.736, p-value < 2.2e-16
## alternative hypothesis: greater
## sample estimates:
## Moran I statistic
                          Expectation
                                                 Variance
         0.453555803
                          -0.008333333
##
                                              0.002795420
#I.TSA
cj_lisa <- local_moran(cj_delta$delta,</pre>
                       nb = nb,
                       wt = wt,
                       nsim = 1000,
                       iseed = set.seed(7188)) %>%
 mutate(mean_p = ifelse(p_ii_sim <= 0.05, as.character(pysal), "Non Sig."),</pre>
         mean_p = factor(mean_p, levels = c("High-High", "High-Low", "Low-High",
                                             "Low-Low", "Non Sig.")))
cj_delta %>%
  cbind(cj_lisa) %>%
  ggplot(aes(fill = mean_p)) +
  geom_sf() +
  geom_sf(lwd = 0.2, color = "black") +
 theme_void() +
  scale_fill_manual(values = c("red", "yellow", "green", "blue", "white"), drop = FALSE)+
  labs(title = "Figure 7: LISA Plot for Carjacking Change Pre/Post Police Murder",
       fill = "Cluster Type",
       caption = "Clusters significant at p < .05 with 1,000 simulations.")</pre>
```

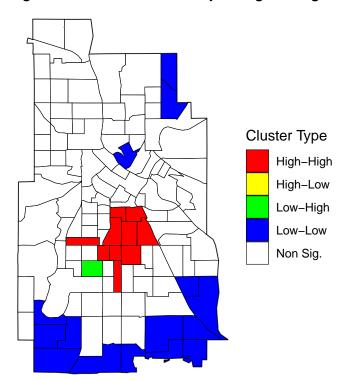


Figure 7: LISA Plot for Carjacking Change Pre/Post Police Mul

Clusters significant at p < .05 with 1,000 simulations.

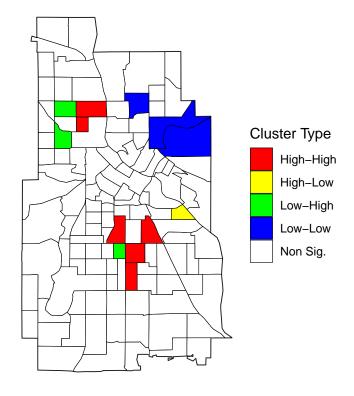
```
ggsave(filename = "Car Jacking/Figures for PAA/fig7.png", bg="white", width = 10, height = 8)
```

### Homicide

```
hom_delta <- prepost_hom %>%
  mutate(delta = homicide_1-homicide_0,
           GEOID = as.numeric(GEOID)) %>%
  left_join(mpls_tract, by = "GEOID") %>%
  st_sf()
nb <- st_contiguity(hom_delta, queen=TRUE)</pre>
wt <- st_weights(nb, style = "W")</pre>
global_moran_test(
  hom_delta$delta,
  nb,
  wt,
  alternative = "greater",
 randomization = TRUE)
##
## Moran I test under randomisation
## data: x
## weights: listw
```

```
##
## Moran I statistic standard deviate = 3.6635, p-value = 0.0001244
## alternative hypothesis: greater
## sample estimates:
## Moran I statistic
                           Expectation
                                                 Variance
##
         0.185646900
                          -0.008333333
                                              0.002803644
#LISA
hom_lisa <- local_moran(hom_delta$delta,</pre>
                       nb = nb,
                       wt = wt,
                       nsim = 1000,
                       iseed = set.seed(7188)) %>%
  mutate(mean_p = ifelse(p_ii_sim <= 0.05, as.character(pysal), "Non Sig."))</pre>
hom_delta %>%
  cbind(hom_lisa) %>%
  ggplot(aes(fill = mean_p)) +
  geom_sf() +
  geom_sf(lwd = 0.2, color = "black") +
  theme_void() +
  scale_fill_manual(values = c("red", "yellow", "green", "blue", "white"))+
  labs(title = "Figure 8: LISA Plot for Homicide Change Pre/Post Police Murder",
       fill = "Cluster Type",
       caption = "Clusters significant at p < .05 with 1,000 simulations.")</pre>
```

Figure 8: LISA Plot for Homicide Change Pre/Post Police Murc



Clusters significant at p < .05 with 1,000 simulations.

### **RE CJ Models**

```
cj_exp_year_acs <- cj_exp_tract_year %>%
  left_join(acs_2020, by = c("GEOID")) %>%
  ungroup() %>%
  mutate(yearzero = year-2017,
         pov_rate_center = pov_rate-mean(pov_rate, na.rm = T),
         white_perc_center = scale(white_perc, center = T, scale = F),
         black_perc_center = scale(black_perc, center = T, scale = F),
         anyjack = ifelse(car_jack==0, 0, 1)) %>%
  drop_na()
cj_exp_prepost <- cj_exp %>%
  mutate(begin_date = ISOweek2date(paste(year, paste0("W", sprintf("%02d", week)), 1,sep = "-")),
         end date = begin date+weeks(1)-days(1),
        post_floyd = as.numeric(begin_date >= as.Date("2020-05-25")),
        post_floyd_3 = as.numeric(begin_date >= as.Date("2020-05-25")+months(3)),
        # stay at home = as.numeric(begin date >= as.Date("2020-03-28") &
         \#state of emerg = as.numeric(begin date >= as.Date("2020-03-13")),
         period = factor(case_when(
          post_floyd==0 & post_floyd_3==0 ~ "Pre-Killing",
          post_floyd>=1 & post_floyd_3==0 ~ "0-3 Months Post-Killing",
          post_floyd>=1 & post_floyd_3>=1 ~ "3+ Months Post-Killing"),
          levels = c("Pre-Killing", "0-3 Months Post-Killing", "3+ Months Post-Killing")),
       GEOID = as.character(GEOID),
        anyjack = ifelse(car_jack==0, 0, 1),
        t = 1:length(car_jack_rate)) %>%
  left_join(acs_2020, by = c("GEOID")) %>%
  drop_na()
library(lme4)
library(lmerTest)
library(lavaan)
cd_model_1 <- ' cd =~ unemp_rate + pov_rate + female_hh_rate + no_hs_dip_rate + black_perc</pre>
                 black_perc ~~ unemp_rate'
cfa_cd <- cfa(cd_model_1, data = cj_exp_prepost, std.lv = T)</pre>
modificationindices(cfa_cd)
##
                 lhs op
                                            mi
                                                  epc sepc.lv sepc.all sepc.nox
                                   rhs
## 13
          unemp_rate ~~
                             pov_rate 776.997 6.514
                                                        6.514
                                                                 0.184
                                                                          0.184
## 14
         unemp_rate ~~ female_hh_rate 640.565 -1.195 -1.195
                                                                -0.195
                                                                         -0.195
## 15
         unemp_rate ~~ no_hs_dip_rate 10.933 0.327
                                                       0.327
                                                                 0.028
                                                                          0.028
## 16
            pov_rate ~~ female_hh_rate 414.268 -2.638 -2.638
                                                                -0.118
                                                                        -0.118
## 17
           pov rate ~~ no hs dip rate 62.778 2.032
                                                       2.032
                                                                 0.048
                                                                         0.048
           pov_rate ~~
## 18
                           black_perc 107.800 -8.340 -8.340
                                                                -0.141
                                                                         -0.141
## 19 female_hh_rate ~~ no_hs_dip_rate 88.342 0.548
                                                       0.548
                                                                 0.075
                                                                          0.075
```

```
black_perc 703.770 5.041
## 21 no_hs_dip_rate ~~
                            black_perc 291.676 -7.088 -7.088
summary(cfa_cd, fit.measures=TRUE, standardized = T)
## lavaan 0.6-12 ended normally after 32 iterations
##
##
     Estimator
                                                         ML
                                                     NLMINB
##
     Optimization method
##
     Number of model parameters
                                                         11
##
##
    Number of observations
                                                     38040
##
## Model Test User Model:
##
     Test statistic
                                                   1433.656
##
##
     Degrees of freedom
##
     P-value (Chi-square)
                                                     0.000
##
## Model Test Baseline Model:
##
##
     Test statistic
                                                 96858.000
##
     Degrees of freedom
                                                         10
##
     P-value
                                                     0.000
##
## User Model versus Baseline Model:
##
##
     Comparative Fit Index (CFI)
                                                     0.985
##
     Tucker-Lewis Index (TLI)
                                                     0.963
##
## Loglikelihood and Information Criteria:
##
##
     Loglikelihood user model (HO)
                                               -600873.408
##
     Loglikelihood unrestricted model (H1)
                                               -600156.580
##
##
     Akaike (AIC)
                                               1201768.815
##
     Bayesian (BIC)
                                               1201862.826
     Sample-size adjusted Bayesian (BIC)
##
                                               1201827.868
##
## Root Mean Square Error of Approximation:
##
     RMSEA
##
                                                     0.097
     90 Percent confidence interval - lower
                                                     0.093
##
##
     90 Percent confidence interval - upper
                                                     0.101
     P-value RMSEA <= 0.05
##
                                                     0.000
##
## Standardized Root Mean Square Residual:
##
##
     SRMR
                                                     0.026
##
## Parameter Estimates:
##
##
     Standard errors
                                                   Standard
##
     Information
                                                   Expected
     Information saturated (h1) model
                                                Structured
```

## 20 female\_hh\_rate ~~

5.041

0.494

-0.367

0.494

-0.367

```
##
## Latent Variables:
##
                      Estimate Std.Err z-value P(>|z|)
                                                            Std.lv Std.all
##
     cd =~
##
       unemp_rate
                         3.520
                                  0.023 154.568
                                                    0.000
                                                             3.520
                                                                       0.749
##
                         8.350
                                  0.067 123.785
                                                    0.000
                                                             8.350
                                                                       0.591
       pov rate
##
                         2.148
                                  0.013 162.986
                                                    0.000
                                                             2.148
                                                                       0.738
       female hh rate
                                  0.026 176.182
                                                    0.000
                                                             4.664
                                                                       0.782
##
       no_hs_dip_rate
                         4.664
##
       black_perc
                        17.533
                                  0.075 234.988
                                                    0.000
                                                             17.533
                                                                       0.959
##
## Covariances:
##
                      Estimate Std.Err z-value P(>|z|)
                                                             Std.lv Std.all
##
    .unemp_rate ~~
      .black_perc
                        -9.425
                                  0.237 -39.771
                                                    0.000
                                                                      -0.583
##
                                                            -9.425
##
## Variances:
##
                      Estimate Std.Err z-value P(>|z|)
                                                             Std.lv Std.all
                                  0.100 97.059
##
                         9.680
                                                    0.000
                                                             9.680
                                                                       0.439
      .unemp_rate
##
                       129.663
                                  0.973 133.264
                                                    0.000 129.663
                                                                       0.650
      .pov_rate
                                  0.032 122.544
##
      .female hh rate
                         3.867
                                                    0.000
                                                             3.867
                                                                       0.456
                                  0.121 114.574
##
      .no_hs_dip_rate
                        13.811
                                                    0.000
                                                           13.811
                                                                       0.388
##
      .black_perc
                        26.971
                                  1.022 26.403
                                                    0.000
                                                             26.971
                                                                       0.081
##
                         1.000
                                                              1.000
                                                                       1.000
       cd
cd_predict <- as.vector(lavPredict(cfa_cd, newdata = cj_exp_prepost))</pre>
cj exp prepost$conc dis <- cd predict</pre>
re <- lmer(car_jack_rate~t+post_floyd+post_floyd_3+conc_dis+
             age_19_29_perc+age_30_49_perc+age_50_69_perc+
             age_70_plus_perc+ post_floyd:conc_dis+post_floyd_3:conc_dis+
             (1 | GEOID),
          data = cj_exp_prepost)
## Warning: Some predictor variables are on very different scales: consider
## rescaling
## Warning: Some predictor variables are on very different scales: consider
## rescaling
summary(re)
## Linear mixed model fit by REML. t-tests use Satterthwaite's method [
## lmerModLmerTest]
## Formula:
## car_jack_rate ~ t + post_floyd + post_floyd_3 + conc_dis + age_19_29_perc +
##
       age_30_49_perc + age_50_69_perc + age_70_plus_perc + post_floyd:conc_dis +
       post floyd 3:conc dis + (1 | GEOID)
##
##
     Data: cj_exp_prepost
## REML criterion at convergence: -74814
##
## Scaled residuals:
     Min
              10 Median
                            3Q
                                  Max
## -1.291 -0.242 -0.121 0.005 33.376
```

```
##
## Random effects:
## Groups
                        Variance Std.Dev.
## GEOID
             (Intercept) 0.000132 0.01149
## Residual
                        0.008113 0.09007
## Number of obs: 38040, groups: GEOID, 120
## Fixed effects:
##
                          Estimate Std. Error
                                                      df t value Pr(>|t|)
## (Intercept)
                         5.052e-02 1.473e-02 1.132e+02
                                                           3.430 0.000842 ***
## t
                         1.379e-07 1.079e-07 1.130e+02
                                                          1.278 0.203926
                         2.486e-02 2.281e-03 3.792e+04 10.895 < 2e-16 ***
## post_floyd
## post_floyd_3
                        -1.500e-03 2.319e-03 3.792e+04 -0.647 0.517690
                                                         0.383 0.702547
## conc_dis
                         6.992e-04 1.827e-03 1.253e+02
## age_19_29_perc
                        -4.491e-04 1.509e-04 1.130e+02 -2.976 0.003577 **
## age_30_49_perc
                        -4.090e-04 2.465e-04
                                               1.130e+02
                                                         -1.659 0.099904
                        -1.056e-03 2.793e-04 1.130e+02 -3.779 0.000253 ***
## age_50_69_perc
## age_70_plus_perc
                        -5.403e-04 3.368e-04 1.130e+02 -1.604 0.111424
                         1.850e-02 2.320e-03 3.792e+04
## post_floyd:conc_dis
                                                         7.975 1.56e-15 ***
## post_floyd_3:conc_dis -8.651e-03 2.359e-03 3.792e+04 -3.668 0.000245 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##
              (Intr) t
                            pst_fl pst__3 cnc_ds a_19_2 a_30_4 a_50_6 a_70__
              -0.021
## post_floyd -0.011 -0.005
## post_flyd_3 0.000 -0.003 -0.913
## conc_dis
              -0.725 0.050 0.000
                                   0.000
## ag_19_29_pr -0.924 -0.052 0.000 0.000 0.680
## ag_30_49_pr -0.878 -0.108
                             0.000
                                   0.000 0.591
                                                  0.747
## ag_50_69_pr -0.741 -0.125
                             0.001
                                   0.000 0.604 0.719
                                                        0.444
## ag_70_pls_p -0.303 -0.036
                             0.000
                                   0.000 0.191
                                                 0.278
                                                        0.282 - 0.119
## pst_flyd:c_ 0.000 0.000
                             0.000 0.000 -0.092 0.000
                                                        0.000 0.000 0.000
## pst_fly_3:_ 0.000 0.000 0.000 0.000 0.000 0.000 0.000 0.000 0.000
##
              pst_:_
## t
## post_floyd
## post_flyd_3
## conc_dis
## ag 19 29 pr
## ag_30_49_pr
## ag_50_69_pr
## ag_70_pls_p
## pst_flyd:c_
## pst_fly_3:_ -0.913
## fit warnings:
## Some predictor variables are on very different scales: consider rescaling
re_logit_cd <- glmer(anyjack ~ t+post_floyd+post_floyd_3+conc_dis+</pre>
            age_19_29_perc+age_30_49_perc+age_50_69_perc+
            age_70_plus_perc+ post_floyd:conc_dis+post_floyd_3:conc_dis+
             (1|GEOID),
                 data = cj_exp_prepost, family = binomial)
```

```
## Warning: Some predictor variables are on very different scales: consider
## rescaling
## Warning in checkConv(attr(opt, "derivs"), opt$par, ctrl = control$checkConv, :
## Model failed to converge with max|grad| = 3.43449 (tol = 0.002, component 1)
## Warning in checkConv(attr(opt, "derivs"), opt$par, ctrl = control$checkConv, : Model is nearly unide:
## - Rescale variables?; Model is nearly unidentifiable: large eigenvalue ratio
## - Rescale variables?
summary(re_logit_cd)
## Generalized linear mixed model fit by maximum likelihood (Laplace
    Approximation) [glmerMod]
## Family: binomial (logit)
## Formula: anyjack ~ t + post_floyd + post_floyd_3 + conc_dis + age_19_29_perc +
##
      age_30_49_perc + age_50_69_perc + age_70_plus_perc + post_floyd:conc_dis +
##
      post_floyd_3:conc_dis + (1 | GEOID)
##
     Data: cj_exp_prepost
##
##
                BIC
                     logLik deviance df.resid
       AIC
   10380.8 10483.3 -5178.4 10356.8
##
## Scaled residuals:
      Min
               1Q Median
                               3Q
                                      Max
## -0.7191 -0.2141 -0.1329 -0.0772 21.2478
## Random effects:
## Groups Name
                      Variance Std.Dev.
## GEOID (Intercept) 0.4228
                               0.6502
## Number of obs: 38040, groups: GEOID, 120
##
## Fixed effects:
                          Estimate Std. Error z value Pr(>|z|)
##
## (Intercept)
                        -3.500e+00 9.026e-01 -3.877 0.000106 ***
## t
                         1.337e-05 8.136e-06
                                              1.644 0.100256
## post_floyd
                         1.786e+00 1.288e-01 13.866 < 2e-16 ***
## post_floyd_3
                         5.340e-02 1.144e-01
                                               0.467 0.640649
## conc_dis
                         5.760e-01 1.181e-01
                                               4.878 1.07e-06 ***
## age_19_29_perc
                        -8.335e-03 9.180e-03 -0.908 0.363887
## age_30_49_perc
                        1.447e-03 1.493e-02
                                               0.097 0.922832
## age_50_69_perc
                        -6.457e-02 1.718e-02 -3.759 0.000171 ***
## age_70_plus_perc
                        -1.858e-02 2.026e-02 -0.917 0.359044
## post_floyd:conc_dis
                       -8.203e-02 1.031e-01 -0.795 0.426430
## post_floyd_3:conc_dis -2.557e-01 9.589e-02 -2.667 0.007655 **
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Correlation of Fixed Effects:
              (Intr) t
                            pst_fl pst__3 cnc_ds a_19_2 a_30_4 a_50_6 a_70__
## t
              -0.096
## post_floyd -0.040 0.010
## post_flyd_3 -0.001 0.001 -0.803
## conc_dis
              -0.727 0.137 0.125 0.004
## ag_19_29_pr -0.924 0.004 -0.002 0.000 0.667
## ag_30_49_pr -0.867 -0.103 -0.002 0.000 0.570 0.749
```

```
## ag_50_69_pr -0.733 -0.096 -0.004 0.000 0.563 0.718 0.446
## ag_70_pls_p -0.276 -0.021 -0.004 -0.003 0.164 0.258 0.263 -0.141
## pst flyd:c 0.020 -0.008 -0.450 0.347 -0.223 0.001 0.002 0.001 0.005
## pst_fly_3:_ 0.002 -0.017 0.329 -0.408 -0.013 0.001 0.002 0.001 0.009
              pst_:_
## t
## post floyd
## post flyd 3
## conc dis
## ag_19_29_pr
## ag_30_49_pr
## ag_50_69_pr
## ag_70_pls_p
## pst_flyd:c_
## pst_fly_3:_ -0.802
## fit warnings:
## Some predictor variables are on very different scales: consider rescaling
## optimizer (Nelder Mead) convergence code: 0 (OK)
## Model failed to converge with max|grad| = 3.43449 (tol = 0.002, component 1)
## Model is nearly unidentifiable: very large eigenvalue
## - Rescale variables?
## Model is nearly unidentifiable: large eigenvalue ratio
## - Rescale variables?
#build in police covariates
#what other covariates do we need here?
  #crude model - no post-treatment control
#businesses - crime generators
#percent single males
#percent "isolated" youth
#similar story with homicide?
#FE models
cj_exp_prepost <- cj_exp_prepost %>%
  mutate(GEOID = as.numeric(GEOID)) %>%
  left_join(homicide, by = c("GEOID", "year", "week")) %>%
  mutate(anyhom = ifelse(homicide==0, 0, 1))
re_homicide <- lmer(homicide_rate~t+post_floyd+post_floyd_3+conc_dis+
             age_19_29_perc+age_30_49_perc+age_50_69_perc+
             age_70_plus_perc+ post_floyd:conc_dis+post_floyd_3:conc_dis+
             (1|GEOID),
          data = cj_exp_prepost)
## Warning: Some predictor variables are on very different scales: consider
## rescaling
## Warning: Some predictor variables are on very different scales: consider
## rescaling
summary(re_homicide)
## Linear mixed model fit by REML. t-tests use Satterthwaite's method [
## lmerModLmerTest]
## Formula:
```

```
## homicide_rate ~ t + post_floyd + post_floyd_3 + conc_dis + age_19_29_perc +
##
      age_30_49_perc + age_50_69_perc + age_70_plus_perc + post_floyd:conc_dis +
##
      post floyd 3:conc dis + (1 | GEOID)
##
     Data: cj_exp_prepost
##
## REML criterion at convergence: -131908.3
## Scaled residuals:
     Min
             10 Median
                           30
                                 Max
## -1.398 -0.100 -0.030 -0.007 45.410
## Random effects:
## Groups
                        Variance Std.Dev.
            Name
             (Intercept) 1.617e-05 0.004021
## GEOID
                        1.810e-03 0.042550
## Residual
## Number of obs: 38040, groups: GEOID, 120
##
## Fixed effects:
                          Estimate Std. Error
                                                      df t value Pr(>|t|)
##
## (Intercept)
                         9.667e-03 5.489e-03 1.133e+02
                                                           1.761
                                                                  0.0809
## t
                         1.817e-08 4.021e-08 1.130e+02
                                                           0.452
                                                                   0.6522
## post floyd
                         7.919e-03 1.078e-03 3.792e+04
                                                           7.348 2.06e-13 ***
## post_floyd_3
                        -4.939e-03 1.096e-03 3.792e+04 -4.508 6.56e-06 ***
## conc dis
                                                           2.340
                         1.617e-03 6.910e-04
                                               1.331e+02
                                                                   0.0208 *
## age_19_29_perc
                        -6.446e-05 5.623e-05 1.130e+02 -1.146
                                                                  0.2541
## age_30_49_perc
                        -1.635e-04 9.186e-05 1.130e+02 -1.780
                                                                  0.0777
                         1.222e-05 1.041e-04 1.130e+02
                                                          0.117
                                                                  0.9068
## age_50_69_perc
## age_70_plus_perc
                        -2.764e-04 1.255e-04 1.130e+02 -2.203
                                                                  0.0296 *
                         9.761e-03 1.096e-03 3.792e+04
                                                         8.906 < 2e-16 ***
## post_floyd:conc_dis
## post_floyd_3:conc_dis -6.798e-03 1.114e-03 3.792e+04 -6.102 1.06e-09 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##
              (Intr) t
                            pst_fl pst__3 cnc_ds a_19_2 a_30_4 a_50_6 a_70__
## t
              -0.021
## post floyd -0.014 -0.004
## post_flyd_3 0.000 -0.003 -0.913
              -0.714 0.050 0.000 0.000
## conc dis
## ag_19_29_pr -0.923 -0.052 0.000 0.000 0.669
## ag 30 49 pr -0.878 -0.108
                             0.000
                                   0.000 0.582
                                                 0.747
## ag_50_69_pr -0.741 -0.125
                             0.000
                                   0.000 0.595
                                                 0.719
                                                         0.444
## ag_70_pls_p -0.303 -0.036
                             0.000
                                   0.000 0.188
                                                 0.278
                                                        0.282 - 0.119
## pst_flyd:c_ 0.000 0.000 0.000 0.000 -0.114 0.000
                                                        0.000 0.000 0.000
## pst_fly_3:_ 0.000 0.000 0.000 0.000 0.000 0.000 0.000 0.000 0.000
##
              pst_:_
## t
## post_floyd
## post_flyd_3
## conc_dis
## ag_19_29_pr
## ag 30 49 pr
## ag_50_69_pr
## ag_70_pls_p
```

```
## pst flvd:c
## pst_fly_3:_ -0.913
## fit warnings:
## Some predictor variables are on very different scales: consider rescaling
re_logit_cd_homicide <- glmer(anyhom ~ t+post_floyd+post_floyd_3+conc_dis+
            age_19_29_perc+age_30_49_perc+age_50_69_perc+
            age_70_plus_perc+ post_floyd:conc_dis+post_floyd_3:conc_dis+
             (1 | GEOID),
                 data = cj_exp_prepost, family = binomial)
## Warning: Some predictor variables are on very different scales: consider
## rescaling
## Warning in checkConv(attr(opt, "derivs"), opt$par, ctrl = control$checkConv, :
## unable to evaluate scaled gradient
## Warning in checkConv(attr(opt, "derivs"), opt$par, ctrl = control$checkConv, :
## Model failed to converge: degenerate Hessian with 1 negative eigenvalues
summary(re_logit_cd)
## Generalized linear mixed model fit by maximum likelihood (Laplace
     Approximation) [glmerMod]
## Family: binomial (logit)
## Formula: anyjack ~ t + post_floyd + post_floyd_3 + conc_dis + age_19_29_perc +
##
       age_30_49_perc + age_50_69_perc + age_70_plus_perc + post_floyd:conc_dis +
##
       post_floyd_3:conc_dis + (1 | GEOID)
##
     Data: cj_exp_prepost
##
##
        ATC
                BIC
                      logLik deviance df.resid
   10380.8 10483.3 -5178.4 10356.8
##
## Scaled residuals:
             1Q Median
##
      Min
                               3Q
## -0.7191 -0.2141 -0.1329 -0.0772 21.2478
##
## Random effects:
## Groups Name
                       Variance Std.Dev.
## GEOID (Intercept) 0.4228 0.6502
## Number of obs: 38040, groups: GEOID, 120
##
## Fixed effects:
##
                          Estimate Std. Error z value Pr(>|z|)
## (Intercept)
                        -3.500e+00 9.026e-01 -3.877 0.000106 ***
## t
                         1.337e-05 8.136e-06
                                               1.644 0.100256
## post floyd
                         1.786e+00 1.288e-01 13.866 < 2e-16 ***
## post_floyd_3
                         5.340e-02 1.144e-01
                                              0.467 0.640649
## conc dis
                         5.760e-01 1.181e-01
                                                4.878 1.07e-06 ***
## age_19_29_perc
                        -8.335e-03 9.180e-03 -0.908 0.363887
## age_30_49_perc
                         1.447e-03 1.493e-02 0.097 0.922832
                        -6.457e-02 1.718e-02 -3.759 0.000171 ***
## age_50_69_perc
## age_70_plus_perc
                        -1.858e-02 2.026e-02 -0.917 0.359044
## post_floyd:conc_dis -8.203e-02 1.031e-01 -0.795 0.426430
## post_floyd_3:conc_dis -2.557e-01 9.589e-02 -2.667 0.007655 **
## ---
```

```
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##
               (Intr) t
                             pst_fl pst__3 cnc_ds a_19_2 a_30_4 a_50_6 a_70__
## t
               -0.096
## post_floyd -0.040 0.010
## post flyd 3 -0.001 0.001 -0.803
             -0.727 0.137 0.125 0.004
## conc dis
## ag_19_29_pr -0.924 0.004 -0.002 0.000 0.667
## ag_30_49_pr -0.867 -0.103 -0.002 0.000 0.570 0.749
## ag_50_69_pr -0.733 -0.096 -0.004 0.000 0.563 0.718 0.446
## ag_70_pls_p -0.276 -0.021 -0.004 -0.003 0.164 0.258 0.263 -0.141
## pst_flyd:c_ 0.020 -0.008 -0.450 0.347 -0.223 0.001 0.002 0.001 0.005
## pst_fly_3:_ 0.002 -0.017 0.329 -0.408 -0.013 0.001 0.002 0.001 0.009
##
              pst_:_
## t
## post_floyd
## post flyd 3
## conc dis
## ag_19_29_pr
## ag_30_49_pr
## ag_50_69_pr
## ag_70_pls_p
## pst_flyd:c_
## pst_fly_3:_ -0.802
## fit warnings:
## Some predictor variables are on very different scales: consider rescaling
## optimizer (Nelder_Mead) convergence code: 0 (OK)
## Model failed to converge with max|grad| = 3.43449 (tol = 0.002, component 1)
## Model is nearly unidentifiable: very large eigenvalue
## - Rescale variables?
## Model is nearly unidentifiable: large eigenvalue ratio
## - Rescale variables?
class(re) <- "lmerMod"</pre>
class(re_logit_cd) <- "lmerMod"</pre>
stargazer(re, re_logit_cd,
         title = "Interrupted Time Series Models of Carjackings, MPLS 2017-2022",
          covariate.labels = c("T", "Post-Killing", "Post-Killing 3 Months",
                              "Conc. Dis.", "Age 19-29", "Age 30-49",
                              "Age 50-69", "Age 70+",
                              "Post-Killing X Conc. Dis.",
                              "Post-Killing 3 Months X Conc. Dis."),
         header = F,
         dep.var.caption = "Carjacking",
         dep.var.labels = c("Rate per 1,000", "Any Carjacking"),
         model.names = FALSE,
         column.labels = c("RE HLM", "RE Logit"),
         report = "vcs",
         ci=TRUE,
         ci.level=0.95,
         ci.separator = "|",
         notes = "95\\% Confidence Intervals in parentheses",
```

Table 1: Interrupted Time Series Models of Carjackings, MPLS 2017-2022

	Carjacking	
	Rate per 1,000 RE HLM (1)	Any Carjacking RE Logit (2)
Т	$ 0.00000 \\ (-0.00000 0.00000) $	$0.00001 \\ (-0.00000 0.00003)$
Post-Killing	$0.025 \\ (0.020 0.029)$	$ \begin{array}{c} 1.786 \\ (1.533 2.038) \end{array} $
Post-Killing 3 Months	$   \begin{array}{c}     -0.002 \\     (-0.006 0.003)   \end{array} $	$0.053 \\ (-0.171 0.278)$
Conc. Dis.	$0.001 \\ (-0.003 0.004)$	$0.576 \\ (0.345 0.807)$
Age 19-29	$ \begin{array}{c} -0.0004 \\ (-0.001 -0.0002) \end{array} $	$   \begin{array}{c}     -0.008 \\     (-0.026 0.010)   \end{array} $
Age 30-49	$ \begin{array}{c} -0.0004 \\ (-0.001 0.0001) \end{array} $	$0.001 \\ (-0.028 0.031)$
Age 50-69	$ \begin{array}{c} -0.001 \\ (-0.002 -0.001) \end{array} $	$ \begin{array}{c} -0.065 \\ (-0.098 -0.031) \end{array} $
Age 70+	$\begin{array}{c} -0.001 \\ (-0.001 0.0001) \end{array}$	$   \begin{array}{c}     -0.019 \\     (-0.058 0.021)   \end{array} $
Post-Killing X Conc. Dis.	$0.019 \\ (0.014 0.023)$	$   \begin{array}{c}     -0.082 \\     (-0.284 0.120)   \end{array} $
Post-Killing 3 Months X Conc. Dis.	$ \begin{array}{c} -0.009 \\ (-0.013 -0.004) \end{array} $	$ \begin{array}{c} -0.256 \\ (-0.444 -0.068) \end{array} $
Constant	$0.051 \\ (0.022 0.079)$	$ \begin{array}{c} -3.500 \\ (-5.269 -1.731) \end{array} $
SD(Tract) SD(Residual) Observations	0.008 0.067 38,040	0.65 38,040
Log Likelihood	37,407.000	-5,178.390

Note:

95% Confidence Intervals in parentheses

### **RE CJ Models**

```
class(re_homicide) <- "lmerMod"</pre>
class(re_logit_cd_homicide) <- "lmerMod"</pre>
stargazer(re_homicide, re_logit_cd_homicide,
          title = "Interrupted Time Series Models of Homicide, MPLS 2017-2022",
          covariate.labels = c("T", "Post-Killing", "Post-Killing 3 Months",
                              "Conc. Dis.", "Age 19-29", "Age 30-49",
                              "Age 50-69", "Age 70+",
                              "Post-Killing X Conc. Dis.",
                              "Post-Killing 3 Months X Conc. Dis."),
          header = F,
          dep.var.caption = "Homicide",
          dep.var.labels = c("Rate per 1,000", "Any Homicide"),
          model.names = FALSE,
          column.labels = c("RE HLM", "RE Logit"),
          report = "vcs",
          ci=TRUE,
          ci.level=0.95,
          ci.separator = "|",
          notes = "95\\% Confidence Intervals in parentheses",
          single.row = F,
          omit.stat = c("adj.rsq", "aic", "bic"),
          \#star.cutoffs = c(.05, .01, .001), star.char = c("*", "**", "***"),
          add.lines = list(c("SD(Tract)", .003, .031),
                           c("SD(Residual)", .065, "-")),
          notes.append = F)
```

Table 2: Interrupted Time Series Models of Homicide, MPLS 2017-2022

	Homicide	
	Rate per 1,000 RE HLM (1)	Any Homicide RE Logit (2)
T	$ 0.00000 \\ (-0.00000 0.00000) $	$ 0.00002 \\ (-0.00000 0.00004) $
Post-Killing	$0.008 \\ (0.006 0.010)$	$1.449 \\ (0.966 1.932)$
Post-Killing 3 Months	$ \begin{array}{c} -0.005 \\ (-0.007   -0.003) \end{array} $	$ \begin{array}{c} -0.573 \\ (-1.033 -0.112) \end{array} $
Conc. Dis.	$0.002 \\ (0.0003 0.003)$	$0.915 \\ (0.566 1.264)$
Age 19-29	$ \begin{array}{c} -0.0001 \\ (-0.0002 0.00005) \end{array} $	$0.016 \\ (-0.011 0.043)$
Age 30-49	$   \begin{array}{c}     -0.0002 \\     (-0.0003 0.00002)   \end{array} $	$0.012 \\ (-0.032 0.056)$
Age 50-69	$0.00001 \\ (-0.0002 0.0002)$	$0.017 \\ (-0.035 0.070)$
Age 70+	$ \begin{array}{c} -0.0003 \\ (-0.001 -0.00003) \end{array} $	$ \begin{array}{c} -0.043 \\ (-0.107 0.020) \end{array} $
Post-Killing X Conc. Dis.	$0.010 \\ (0.008 0.012)$	$0.145 \\ (-0.175 0.465)$
Post-Killing 3 Months X Conc. Dis.	$ \begin{array}{c} -0.007 \\ (-0.009 -0.005) \end{array} $	$   \begin{array}{c}     -0.182 \\     (-0.487 0.123)   \end{array} $
Constant	$0.010 \\ (-0.001 0.020)$	$ \begin{array}{c} -7.125 \\ (-9.722 -4.528) \end{array} $
SD(Tract) SD(Residual) Observations Log Likelihood	0.003 0.065 38,040 65,954.170	0.031 $ 38,040$ $-1,661.554$

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Note:

95% Confidence Intervals in parentheses