Car Jacking MPLS - Tract

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Spatial Data

```
#MN tracts
tracts <- get_acs(geography = "tract",</pre>
                  state = "MN",
                variables = "B01001_001E",
                output = "wide",
                survey = "acs5",
                year = 2020,
                geometry = T)
#Minneapolis Shapefile
mpls <- st_read("Data/mpls_city-shp/16cdbbfa-ad10-493c-afaf-52b61f2e76e42020329-1-180h9ap.whbo.shp") %>
   st_transform(st_crs(tracts))
## Reading layer `16cdbbfa-ad10-493c-afaf-52b61f2e76e42020329-1-180h9ap.whbo' from data source `C:\User
## using driver `ESRI Shapefile'
\#\# Simple feature collection with 1 feature and 4 fields
## Geometry type: POLYGON
## Dimension:
## Bounding box: xmin: -93.32911 ymin: 44.89059 xmax: -93.19433 ymax: 45.05125
## Geodetic CRS: WGS 84
mpls_tract <- tracts %>%
  st_filter(mpls, .predicate = st_intersects) %>%
  mutate(GEOID = as.numeric(GEOID),
         tract_area = as.numeric(st_area(.)),
         tract_area_sqkm = tract_area*.000001,
         tract_area_sqmi = tract_area_sqkm*.386102,
         intersection_area = as.numeric(st_area(st_intersection(., mpls))),
         perc_intersection = intersection_area/tract_area*100) %>%
  filter(perc_intersection >= 2) %>%
  select(-"B01001_001M")
```

ACS Covariates and Denominators

```
acs_17 <- get_acs(
  geography = "tract",
  variables = c("B01001_001E"),
  year = 2017,</pre>
```

```
state = "MN",
 county = "Hennepin",
  geometry = TRUE) %>%
 select(estimate)
## Getting data from the 2013-2017 5-year ACS
## Downloading feature geometry from the Census website. To cache shapefiles for use in future session
##
acs_18 <- get_acs(</pre>
 geography = "tract",
 variables = c("B01001_001E"),
 year = 2018,
 state = "MN",
 county = "Hennepin",
  geometry = TRUE
) %>%
select(estimate)
## Getting data from the 2014-2018 5-year ACS
## Downloading feature geometry from the Census website. To cache shapefiles for use in future session
##
acs_19 <- get_acs(</pre>
 geography = "tract",
 variables = c("B01001_001E"),
 year = 2019,
 state = "MN",
 county = "Hennepin",
  geometry = TRUE
) %>%
select(estimate)
## Getting data from the 2015-2019 5-year ACS
## Downloading feature geometry from the Census website. To cache shapefiles for use in future session
acs_20 <- get_acs(</pre>
 geography = "tract",
 variables = c("B01001 001E"),
 year = 2020,
 state = "MN",
 county = "Hennepin",
 geometry = TRUE
## Getting data from the 2016-2020 5-year ACS
## Downloading feature geometry from the Census website. To cache shapefiles for use in future session
library(tigris)
## To enable caching of data, set `options(tigris_use_cache = TRUE)`
## in your R script or .Rprofile.
```

```
hennepin_blocks <- blocks(
  "MN",
  "Hennepin",
  year = 2020
##
#2017 ACS interpolation
acs_1720 <- interpolate_pw(</pre>
 from = acs_17,
 to = acs_20,
 to_id = "GEOID",
  weights = hennepin_blocks,
  weight_column = "POP20",
  crs = 26993,
  extensive = TRUE) %>%
  mutate(year = 2017)
#2018 ACS interpolation
acs_1820 <- interpolate_pw(</pre>
 from = acs_18,
 to = acs_20,
 to_id = "GEOID",
  weights = hennepin_blocks,
  weight_column = "POP20",
  crs = 26993,
  extensive = TRUE) %>%
  mutate(year = 2018)
#2019 ACS interpolation
acs_1920 <- interpolate_pw(</pre>
 from = acs_19,
  to = acs_20,
  to_id = "GEOID",
  weights = hennepin_blocks,
  weight_column = "POP20",
  crs = 26993,
  extensive = TRUE) %>%
  mutate(year = 2019)
pop_denoms <- acs_20 %>%
  st_transform(crs = 26993) %>%
  mutate(year = 2020) \%
  select(-moe, -variable, -NAME) %>%
  rbind(acs_1720, acs_1820, acs_1920) %>%
  filter(GEOID %in% mpls_tract$GEOID)
#2021+2022: LOCF
pop_denom_21 <- pop_denoms \%>%
  filter(year==2020) %>%
  select(GEOID, year, estimate) %>%
```

```
mutate(year = 2021)
pop_denom_22 <- pop_denoms %>%
  filter(year==2020) %>%
  select(GEOID, year, estimate) %>%
  mutate(year = 2022)
pop denom locf <- pop denoms %>%
  rbind(pop_denom_21, pop_denom_22) %>%
  rename(total pop = estimate) %>%
  mutate(GEOID = as.numeric(GEOID)) %>%
  st drop geometry()
#ACS 2020 L-2 covariates
acs_2020 <- get_acs(geography = "tract",</pre>
            state = "MN",
               variables = c("B01001_001E", "B03002_003E", "B03002_004E", "B03002_005E",
                             "B03002_006E", "B03002_007E", "B03002_008E", "B03002_009E",
                             "B03002_010E", "B03002_011E", "B03002_012E", "B23025_002E",
                             "B23025_005E", "B17001_002E", "B19057_002E", "B11003_015E",
                             "B06009_002E", "B06009_005E", "C24010_001E", "C24010_003E",
                             "C24010_039E", "B11001_003E", "B01001_002E", "B05001_006E",
                             "B01001_003E", "B01001_004E", "B01001_005E", "B01001_006E",
                             "B01001_007E", "B01001_008E", "B01001_009E", "B01001_010E",
                             "B01001_011E", "B01001_012E", "B01001_013E", "B01001_014E",
                             "B01001_015E", "B01001_016E", "B01001_017E", "B01001_018E",
                             "B01001_019E", "B01001_020E", "B01001_021E", "B01001_022E",
                             "B01001_023E", "B01001_024E", "B01001_025E", "B01001_027E",
                             "B01001_028E", "B01001_029E", "B01001_030E", "B01001_031E",
                             "B01001_032E", "B01001_033E", "B01001_034E", "B01001_035E",
                             "B01001_036E", "B01001_037E", "B01001_038E", "B01001_039E",
                             "B01001_040E", "B01001_041E", "B01001_042E", "B01001_043E",
                             "B01001_044E", "B01001_045E", "B01001_046E", "B01001_047E",
                             "B01001_048E", "B01001_049E", "B07001_017E", "B25003_002E",
                             "B05002_013E", "B19013_001E"),
               output = "wide",
               survey = "acs5",
               vear = 2020) %>%
  select(-ends_with("M", ignore.case = F)) %>%
  rename(total_pop = B01001_001E, white_pop = B03002_003E, black_pop = B03002_004E,
         na_pop = B03002_005E, asian_pop = B03002_006E, hpi_pop = B03002_007E,
         other_pop = B03002_008E, biracial_pop = B03002_009E, biracial_other_pop = B03002_010E,
         biracial three pop = B03002 011E, hisp pop = B03002 012E, total ilf = B23025 002E,
         unemp = B23025_005E, povlevel = B17001_002E, pub_assist = B19057_002E,
         female_hh = B11003_015E, no_hs_dip = B06009_002E, bach_degree = B06009_005E,
         total employed = C24010 001E, employed mbsa male = C24010 003E,
         employed_mbsa_female = C24010_039E, mar_fam = B11001_003E, male = B01001_002E,
         noncitizen = B05001_006E,
        age_m_5_under = B01001_003E, age_m_5_9 = B01001_004E, age_m_10_14 = B01001_005E,
        age_m_15_17 = B01001_006E, age_m_18_19 = B01001_007E, age_m_20 = B01001_008E,
        age_m_21 = B01001_009E, age_m_22_24 = B01001_010E, age_m_25_29 = B01001_011E,
        age_m_30_34 = B01001_012E, age_m_35_39 = B01001_013E, age_m_40_44 = B01001_014E,
        age m 45 49 = B01001_015E, age m 50 54 = B01001_016E, age m 55 59 = B01001_017E,
```

```
age m 60 61 = B01001_018E, age m 62 64 = B01001_019E, age m 65 66 = B01001_020E,
      age_m_67_69 = B01001_021E, age_m_70_74 = B01001_022E, age_m_75_79 = B01001_023E,
      age m 80 84 = B01001 024E, age m 85 plus = B01001 025E, age f 5 under = B01001 027E,
      age f 5 9 = B01001 028E, age f 10 14 = B01001 029E, age f 15 17 = B01001 030E,
      age f 18 19 = B01001 031E, age f 20 = B01001 032E, age f 21 = B01001 033E,
      age f 22 24 = B01001 034E, age f 25 29 = B01001 035E, age f 30 34 = B01001 036E,
      age_f_35_39 = B01001_037E, age_f_40_44 = B01001_038E, age_f_45_49 = B01001_039E,
      age_f_50_54 = B01001_040E, age_f_55_59 = B01001_041E, age_f_60_61 = B01001_042E,
      age f 62 64 = B01001 043E, age f 65 66 = B01001 044E, age f 67 69 = B01001 045E,
      age_f_70_74 = B01001_046E, age_f_75_79 = B01001_047E, age_f_80_84 = B01001_048E,
      age_f_85_plus = B01001_049E, res_mob = B07001_017E,
      own_hh = B25003_002E, foreign = B05002_013E,
     med_hh_inc = B19013_001E) %>%
mutate(white_prop = white_pop/total_pop,
       black_prop = black_pop/total_pop,
       na_prop = na_pop/total_pop,
       asian_prop = asian_pop/total_pop,
       hpi_prop = hpi_pop/total_pop,
       other_prop = other_pop/total_pop,
       biracial_prop = (biracial_pop+biracial_other_pop+biracial_three_pop)/total_pop,
       hisp prop = hisp pop/total pop,
       white perc = 100*white pop/total pop,
       black_perc = 100*black_pop/total_pop,
       na_perc = 100*na_pop/total_pop,
       asian_perc = 100*asian_pop/total_pop,
       hpi perc = 100*hpi pop/total pop,
       other_perc = 100*other_pop/total_pop,
       biracial_perc = 100*(biracial_pop+biracial_other_pop+biracial_three_pop)/total_pop,
       hisp_perc = 100*hisp_pop/total_pop,
       unemp_rate = 100*unemp/total_ilf,
       pov_rate = 100*povlevel/total_pop,
       pub_assist_rate = 100*pub_assist/total_pop,
       female hh rate = 100*female hh/total pop,
       no_hs_dip_rate = 100*no_hs_dip/total_pop,
       bach degree rate = 100*bach degree/total pop,
       employed_mbsa = employed_mbsa_male+employed_mbsa_female,
       employed_mbsa_rate = 100*employed_mbsa/total_employed,
       mar fam rate = 100*mar fam/total pop,
       male rate = 100*male/total pop,
       noncitizen rate = 100*noncitizen/total pop,
       race_eth_hetero = 1-(white_prop^2+black_prop^2+na_prop^2+asian_prop^2+
                     hpi_prop^2+other_prop^2+other_prop^2+biracial_prop^2+hisp_prop^2),
       age_below_18_perc = 100*(age_m_5_under+age_f_5_under+age_m_5_9+
                                age_f_5_9+age_m_10_14+age_f_10_14+age_m_15_17+
                                age_f_15_17)/total_pop,
       age 19 29 perc = 100*(age_m_18_19+age_f_18_19+age_m_20+age_f_20+age_m_21+age_f_21+
                        age_m_22_24+age_f_22_24+age_m_25_29+age_f_25_29)/total_pop,
       age_30_49_perc = 100*(age_m_30_34+age_f_30_34+age_m_35_39+age_f_35_39+
                             age_m_40_44+age_f_40_44+age_m_45_49+age_f_45_49)/total_pop,
       age_50_69_perc = 100*(age_m_50_54+age_f_50_54+age_m_55_59+age_f_55_59+
                             age_m_60_61+age_f_60_61+age_m_62_64+age_f_62_64+
                             age_m_65_66+age_f_65_66+age_m_67_69+age_f_67_69)/total_pop,
       age_70_plus_perc = 100*(age_m_70_74+age_f_70_74+age_m_75_79+age_f_75_79+
```

```
age_m_80_84+age_f_80_84+age_m_85_plus+age_f_85_plus)/total_pop,
         res_mob_rate = 100-100*res_mob/total_pop,
         own_hh_rate = 100*own_hh/total_pop,
         foreign_rate = 100*foreign/total_pop)
## Getting data from the 2016-2020 5-year ACS
# 2020 Census 18+ Denominator
dc2020 <- get_decennial(</pre>
  geography = "tract",
 variables = c("P3_001N"),
 year = 2020,
  state = "MN".
  county = "Hennepin",
  geometry = F) %>%
  mutate(GEOID = is.numeric(GEOID)) %>%
  rename(total_pop = value) %>%
  select(-GEOID)
## Getting data from the 2020 decennial Census
## Using the PL 94-171 Redistricting Data summary file
## Note: 2020 decennial Census data use differential privacy, a technique that
## introduces errors into data to preserve respondent confidentiality.
## i Small counts should be interpreted with caution.
## i See https://www.census.gov/library/fact-sheets/2021/protecting-the-confidentiality-of-the-2020-cen
## This message is displayed once per session.
```

Expanded MPLS Carjacking (Crime Incidents) Data

```
cj_exp <- read_csv("Car Jacking/MPDdata_2017to2022.csv") %>%
  mutate(date=mdy_hm(reporteddate),
         year=isoyear(date),
         week=isoweek(date)) %>%
  select(casenumber, year, week, latitude, longitude) %>%
  distinct(casenumber, .keep_all = TRUE) %>% #collapsing to incident-level
  drop_na(latitude, longitude) %>%
  st_as_sf(coords = c("longitude", "latitude"), crs = "NAD83", remove=F) %>%
  st_join(mpls_tract) %>% #spatial join neighborhoods
  st_drop_geometry() %>%
  drop_na(GEOID) %>%
  group_by(year, week, GEOID, .drop=F) %>%
  tally(name = "car_jack") %>%
  ungroup() %>%
  complete(year, week, GEOID=mpls_tract$GEOID, fill = list(car_jack = 0)) %>%
  filter(!(year==2021 & week==53)) %>%
  arrange(GEOID, year, week) %>%
  left_join(mpls_tract, by = "GEOID") %>%
  left_join(dc2020, by = c("NAME")) %>%
  mutate(car_jack_rate = car_jack/total_pop*1000) %>%
  st_as_sf()
## Rows: 3779 Columns: 28
```

-- Column specification -----

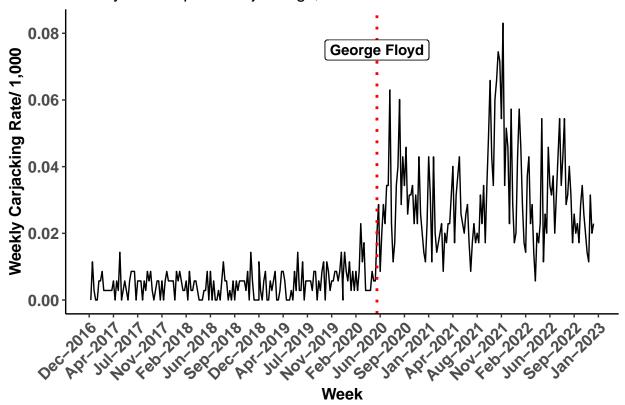
```
## Delimiter: ","
## chr (23): casenumber, dataset, closurecode, closurecode_mpd, reporteddate, c...
## dbl (4): precinct, latitude, longitude, age
## lgl (1): dateofbirth
##
## i Use `spec()` to retrieve the full column specification for this data.
## i Specify the column types or set `show_col_types = FALSE` to quiet this message.
```

MPLS Carjackings by Week - MPD Extended Data

```
#aggregate to week over tracts
cj_exp_week <- cj_exp %>%
  group by(year, week) %>%
  summarize(car_jack = sum(car_jack, na.rm = T),
            total_pop = sum(total_pop, na.rm = T)) %>%
  mutate(begin_date = ISOweek2date(paste(year, pasteO("W", sprintf("%02d", week)), 1,sep = "-")),
         end_date = begin_date+weeks(1)-days(1),
         car_jack_rate = car_jack/total_pop*1000,
         pre_post_floyd = ifelse(end_date <= as.Date("2020-05-25"), 0, 1)) %%</pre>
  filter(end_date <= as.Date("2022-12-31")) %>%
  ungroup() %>%
  mutate(csma = forecast::ma(car_jack_rate, order=5, centre=TRUE),
         tsma = TTR::SMA(car_jack_rate, n=5))
## `summarise()` has grouped output by 'year'. You can override using the
## `.groups` argument.
## Registered S3 method overwritten by 'quantmod': method from as.zoo.data.frame
pre_mean <- mean(cj_exp_week$car_jack_rate[cj_exp_week$pre_post_floyd==0], na.rm = T)</pre>
post_mean <- mean(cj_exp_week$car_jack_rate[cj_exp_week$pre_post_floyd==1], na.rm = T)</pre>
c(pre_mean, post_mean)
## [1] 0.004835884 0.030500004
post mean/pre mean
## [1] 6.307018
ggplot(cj exp week)+
  geom_line(aes(x=begin_date, y=car_jack_rate))+
  scale_x_date(date_labels = "%b-%Y", date_breaks = "15 weeks",
               limits = c(min(cj_exp_week$begin_date), max(cj_exp_week$begin_date)))+
  geom_vline(xintercept=cj_exp_week$begin_date[cj_exp_week$year==2020 &
                                                  cj_exp_week$week==isoweek(date("2020-05-25"))],
              linetype="dotted", color="red", size=1)+
   geom_label(aes(x=cj_exp_week$begin_date[cj_exp_week$year==2020 &
                                              cj_exp_week$week==isoweek(date("2020-05-25"))],
                 y=0.075),
             label = "George Floyd", show.legend = FALSE,
             fontface = "bold")+
  labs(title = "Weekly Minneapolis Carjackings, 2017-2022",
       x = "Week",
       y = "Weekly Carjacking Rate/ 1,000",
```

```
## Warning: Using `size` aesthetic for lines was deprecated in ggplot2 3.4.0.
## i Please use `linewidth` instead.
## This warning is displayed once every 8 hours.
## Call `lifecycle::last_lifecycle_warnings()` to see where this warning was ## generated.
## Warning: Use of `cj_exp_week$begin_date` is discouraged.
## i Use `begin_date` instead.
## Warning: Use of `cj_exp_week$year` is discouraged.
## i Use `year` instead.
## Warning: Use of `cj_exp_week$week` is discouraged.
## i Use `year` instead.
```

Weekly Minneapolis Carjackings, 2017–2022



ggsave(filename = "Car Jacking/Figures for publication/fig1.png", bg="white", width = 10, height = 8)

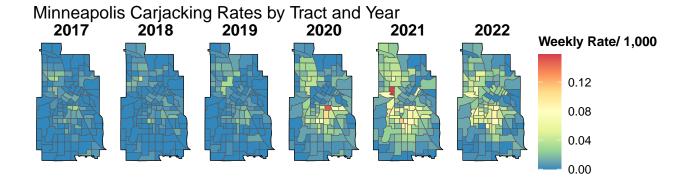
```
## Warning: Use of `cj_exp_week$begin_date` is discouraged.
## i Use `begin_date` instead.
## Warning: Use of `cj_exp_week$year` is discouraged.
## i Use `year` instead.
## Warning: Use of `cj_exp_week$week` is discouraged.
## i Use `week` instead.
```

Structural Change in Carjacking

```
library(strucchange)
## Loading required package: sandwich
## Attaching package: 'strucchange'
## The following object is masked from 'package:stringr':
##
##
       boundary
cj_exp_week <- cj_exp_week %>%
  ungroup() %>%
  mutate(t = row_number())
sctest(cj_exp_week$car_jack_rate~cj_exp_week$t,
       type = "Chow",
       point = 180)
##
## Chow test
## data: cj_exp_week$car_jack_rate ~ cj_exp_week$t
## F = 51.174, p-value < 2.2e-16
breakpoints(car_jack_rate~t,
            data = cj_exp_week,
            breaks = 1)
##
     Optimal 2-segment partition:
##
## Call:
## breakpoints.formula(formula = car_jack_rate ~ t, breaks = 1,
       data = cj_exp_week)
##
## Breakpoints at observation number:
## Corresponding to breakdates:
## 0.5714286
```

MPLS ZCTA Carjackings Map - MPD Extended Data

```
#aggregate to neighborhood-year level
cj_exp_tract_year <- cj_exp %>%
 group_by(GEOID, year) %>%
  summarize(car jack = sum(car jack, na.rm = T),
            total_pop = sum(B01001_001E, na.rm = T),
            car_jack_rate = car_jack/total_pop*1000) %>%
 mutate(GEOID = as.character(GEOID))
## `summarise()` has grouped output by 'GEOID'. You can override using the
## `.groups` argument.
ggplot() +
  geom_sf(data = cj_exp_tract_year, aes(geometry = geometry, fill = car_jack_rate)) +
  geom_sf(data = mpls, aes(geometry = geometry), color = "black", alpha = 0)+
 facet grid(~year)+
  scale_fill_distiller(palette = "Spectral")+
  labs(title = "Minneapolis Carjacking Rates by Tract and Year",
      fill = "Weekly Rate/ 1,000")+
  theme_void()+
  theme(strip.text.x = element_text(face = "bold", size = 12),
       legend.title = element_text(face = "bold", size = 10))
```



```
ggsave(filename = "Car Jacking/Figures for publication/fig2.png", bg="white", width = 10, height = 8)
```

MPLS Murder (Crime Incidents) Data

```
mpd_2016 <- read_csv("Data/Police_Incidents_2016.csv")</pre>
## Rows: 20155 Columns: 20
## -- Column specification -----
## Delimiter: ","
## chr (12): PublicAddress, CCN, Precinct, ReportedDate, BeginDate, Offense, D...
        (7): FID, ControlNbr, GBSID, Lat, Long, X, Y
## time (1): Time
## i Use `spec()` to retrieve the full column specification for this data.
## i Specify the column types or set `show_col_types = FALSE` to quiet this message.
mpd_2017 <- read_csv("Data/Police_Incidents_2017.csv")</pre>
## Rows: 22085 Columns: 20
## -- Column specification -----
## Delimiter: ","
## chr (12): PublicAddress, CCN, Precinct, ReportedDate, BeginDate, Offense, D...
        (7): FID, ControlNbr, GBSID, Lat, Long, X, Y
## time (1): Time
##
## i Use `spec()` to retrieve the full column specification for this data.
## i Specify the column types or set `show_col_types = FALSE` to quiet this message.
mpd_2018a <- read_csv("Data/Police_Incidents_2018.csv")</pre>
## Rows: 7350 Columns: 20
## Delimiter: ","
## chr (12): PublicAddress, CCN, Precinct, ReportedDate, BeginDate, Offense, D...
        (7): FID, ControlNbr, GBSID, Lat, Long, X, Y
## time (1): Time
## i Use `spec()` to retrieve the full column specification for this data.
## i Specify the column types or set `show_col_types = FALSE` to quiet this message.
mpd_2018b <- read_csv("Data/Police_Incidents_2018_PIMS.csv")</pre>
## Rows: 11603 Columns: 23
## -- Column specification -----
## Delimiter: ","
## chr (13): publicaddress, caseNumber, precinct, reportedDate, beginDate, repo...
## dbl (10): X, Y, reportedTime, beginTime, centergbsid, centerLong, centerLat,...
##
## i Use `spec()` to retrieve the full column specification for this data.
## i Specify the column types or set `show_col_types = FALSE` to quiet this message.
mpd_2019 <- read_csv("Data/Police_Incidents_2019.csv")</pre>
## Warning: One or more parsing issues, call `problems()` on your data frame for details,
## e.g.:
    dat <- vroom(...)</pre>
##
    problems(dat)
```

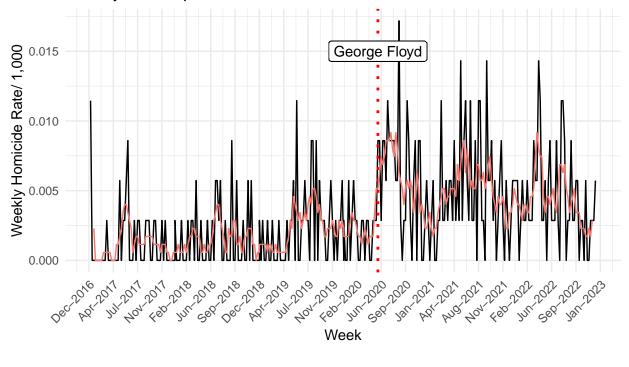
```
## Rows: 22934 Columns: 23
## -- Column specification -------
## Delimiter: ","
## chr (12): publicaddress, caseNumber, reportedDate, beginDate, reportedDateTi...
## dbl (11): X, Y, precinct, reportedTime, beginTime, centergbsid, centerLong, ...
## i Use `spec()` to retrieve the full column specification for this data.
## i Specify the column types or set `show_col_types = FALSE` to quiet this message.
mpd 2020 <- read csv("Data/Police Incidents 2020.csv")</pre>
## Warning: One or more parsing issues, call `problems()` on your data frame for details,
## e.g.:
    dat <- vroom(...)</pre>
##
    problems(dat)
## Rows: 24136 Columns: 23
## -- Column specification ------
## Delimiter: ","
## chr (12): publicaddress, caseNumber, reportedDate, beginDate, reportedDateTi...
## dbl (11): X, Y, precinct, reportedTime, beginTime, centergbsid, centerLong, ...
## i Use `spec()` to retrieve the full column specification for this data.
## i Specify the column types or set `show_col_types = FALSE` to quiet this message.
mpd_2021 <- read_csv("Data/Police_Incidents_2021.csv")</pre>
## Warning: One or more parsing issues, call `problems()` on your data frame for details,
## e.g.:
    dat <- vroom(...)</pre>
    problems(dat)
## Rows: 24755 Columns: 23
## -- Column specification -------
## Delimiter: ","
## chr (12): publicaddress, caseNumber, reportedDate, beginDate, reportedDateTi...
## dbl (11): X, Y, precinct, reportedTime, beginTime, centergbsid, centerLong, ...
## i Use `spec()` to retrieve the full column specification for this data.
## i Specify the column types or set `show_col_types = FALSE` to quiet this message.
mpd_2022 <- read_csv("Data/Police_Incidents_2022.csv")</pre>
## Rows: 26526 Columns: 23
## -- Column specification -----
## Delimiter: ","
## chr (13): publicaddress, caseNumber, precinct, reportedDate, beginDate, repo...
## dbl (9): X, Y, OBJECTID, reportedTime, beginTime, centerLong, centerLat, ce...
## lgl (1): centergbsid
## i Use `spec()` to retrieve the full column specification for this data.
## i Specify the column types or set `show_col_types = FALSE` to quiet this message.
pre_pims_base <- mpd_2016 %>%
 rbind(mpd 2017) %>%
 rbind(mpd_2018a) %>%
 rename(reportedDate = ReportedDate,
```

```
centerLong = Long,
         centerLat = Lat) %>%
  select(FID, centerLong, centerLat, Offense, reportedDate) %>%
  rename(OBJECTID = FID,
        X = centerLong,
         Y = centerLat,
         offense = Offense)
post_pims_base <- mpd_2018b %>%
  rbind(mpd_2019) %>%
  rbind(mpd_2020) %>%
  rbind(mpd_2021) %>%
  rbind(mpd_2022) %>%
  select(OBJECTID, X, Y, offense, reportedDate)
mpd <- pre_pims_base %>%
  rbind(post_pims_base)
#aggregate homicides to tract-week
homicide <- mpd %>%
  mutate(date=ymd hms(reportedDate),
         year=isoyear(date),
         week=isoweek(date)) %>%
  filter(offense=="MURDR" & year!=2016 & year!=2015) %>% #filter homicides
  select(OBJECTID, year, week, Y, X) %>%
  st_as_sf(coords = c("X", "Y"), crs = "NAD83", remove=F) %>%
  st_join(mpls_tract) %>% #spatial join neighborhoods
  st_drop_geometry() %>%
  filter(!is.na(GEOID)) %>%
  group_by(year, week, GEOID, .drop=F) %>%
  tally(name = "homicide") %>%
  ungroup() %>%
  complete(year, week, GEOID=mpls_tract$GEOID, fill = list(homicide = 0)) %>%
  filter(!(year==2021 & week==53)) %>%
  arrange(GEOID, year, week) %>%
  left_join(mpls_tract, by = "GEOID") %>%
  left_join(dc2020, by = c("NAME")) %>%
  mutate(homicide_rate = homicide/total_pop*1000) %>%
  st as sf()
```

MPLS Murder by Week

```
ungroup() %>%
  mutate(csma = forecast::ma(homicide_rate, order=5,centre=TRUE),
         tsma = TTR::SMA(homicide_rate, n=5))
## `summarise()` has grouped output by 'year'. You can override using the
## `.groups` argument.
pre_mean <- mean(homicide_week$homicide_rate[homicide_week$pre_post_floyd==0], na.rm = T)</pre>
post_mean <- mean(homicide_week$homicide_rate[homicide_week$pre_post_floyd==1], na.rm = T)</pre>
c(pre_mean, post_mean)
## [1] 0.001749826 0.004984354
post_mean/pre_mean
## [1] 2.848485
ggplot(homicide_week)+
 geom_line(aes(x=begin_date, y=homicide_rate))+
 scale_x_date(date_labels = "%b-%Y", date_breaks = "15 weeks",
               limits = c(min(homicide_week$begin_date), max(homicide_week$begin_date)))+
     labs(title = "Weekly Minneapolis Homicide, 2017-2022",
       x = "Week",
       y = "Weekly Homicide Rate/ 1,000",
       color = NULL)+
  geom_vline(xintercept=homicide_week$begin_date[homicide_week$year==2020 &
                                                 homicide_week$week==isoweek(date("2020-05-25"))],
              linetype="dotted", color="red", size=1)+
   geom_label(aes(x=homicide_week$begin_date[homicide_week$year==2020 &
                                             homicide week$week==isoweek(date("2020-05-25"))],
                 y=0.015),
             label = "George Floyd", show.legend = FALSE)+
  theme_minimal()+
  geom_line(aes(x=begin_date, y=csma, color = "CSMA(5)"))+
  theme(axis.text.x=element_text(angle=45, hjust=1)) +
  theme(legend.key.size = unit(0.8, "cm"),legend.position = "bottom")
## Warning: Use of `homicide_week$begin_date` is discouraged.
## i Use `begin_date` instead.
## Warning: Use of `homicide_week$year` is discouraged.
## i Use `year` instead.
## Warning: Use of `homicide_week$week` is discouraged.
## i Use `week` instead.
## Warning: Removed 4 rows containing missing values (`geom_line()`).
```

Weekly Minneapolis Homicide, 2017-2022



— CSMA(5)

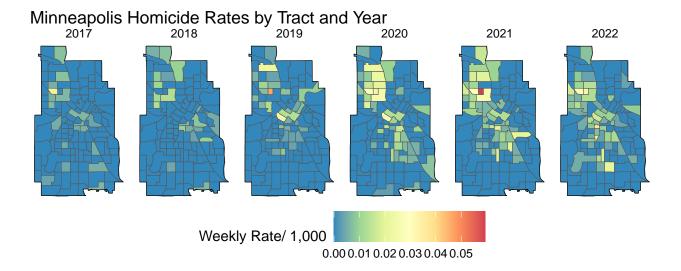
```
ggsave(filename = "Car Jacking/Figures for publication/fig3.png", bg="white", width = 10, height = 8)
## Warning: Use of `homicide_week$begin_date` is discouraged.
## i Use `begin_date` instead.
## Warning: Use of `homicide_week$year` is discouraged.
## i Use `year` instead.
## Warning: Use of `homicide_week$week` is discouraged.
## i Use `week` instead.
## warning: Removed 4 rows containing missing values (`geom_line()`).
```

Structural Change in Carjacking

```
## F = 16.772, p-value = 1.21e-07
breakpoints(homicide_rate~t,
            data = homicide_week,
            breaks = 1)
##
##
     Optimal 2-segment partition:
##
## Call:
## breakpoints.formula(formula = homicide_rate ~ t, breaks = 1,
       data = homicide week)
##
## Breakpoints at observation number:
## 180
## Corresponding to breakdates:
## 0.5714286
```

MPLS ZCTA Murder Map - MPD Extended Data

```
#aggregate to neighborhood-year level
homicide_tract_year <- homicide %>%
  group_by(GEOID, year) %>%
  summarize(homicide = sum(homicide, na.rm = T),
            total_pop = sum(B01001_001E, na.rm = T),
            homicide_rate = homicide/total_pop*1000) %>%
 mutate(GEOID = as.character(GEOID))
## `summarise()` has grouped output by 'GEOID'. You can override using the
## `.groups` argument.
ggplot() +
 geom_sf(data = homicide_tract_year, aes(geometry = geometry, fill = homicide_rate)) +
  geom_sf(data = mpls, aes(geometry = geometry), color = "black", alpha = 0)+
  facet_grid(~year)+
  scale_fill_distiller(palette = "Spectral")+
  labs(title = "Minneapolis Homicide Rates by Tract and Year",
      fill = "Weekly Rate/ 1,000")+
  theme void() +
  theme(legend.key.size = unit(0.8, "cm"),legend.position = "bottom")
```



```
ggsave(filename = "Car Jacking/Figures for publication/fig4.png", bg="white", width = 10, height = 8)
```

Dispersion of Change from 2017-2019 to 2020-2021

Car Jacking

```
crimedispersion <- function
(data1, unitID, time1, time2, method = "match") {

    # define variables to limit build warnings
    adjusted <- Ut1 <- Ut2 <- Rt1 <- Rt2 <- chg <- pct <- NULL

# ERROR CHECKING. Has user passed a data frame?
if (!is.data.frame(data1)) {
    stop("The input data specified is not a data.frame object. Please fix.")
}

# Build a local data.frame and populate with passed arguments
source_rows <- nrow(data1)
df1 <- data.frame(matrix(ncol = 3, nrow = source_rows))
colnames(df1) <- c("unit", "time1", "time2")
df1$unit <- data1[, unitID]
df1$time1 <- data1[, time1]
df1$time2 <- data1[, time2]
if (method == "remove") {</pre>
```

```
analysisMethod <- "remove"</pre>
} else {
 analysisMethod <- "match"</pre>
}
# ERROR CHECKING. Did user pass numeric columns where needed?
try (df1$time1 <- as.numeric(df1$time1), silent = TRUE)</pre>
try (df1$time2 <- as.numeric(df1$time2), silent = TRUE)</pre>
if (!class(df1$time1)[1] == "numeric") {
 stop("The time1 field is not a numeric object. Please fix.")
}
if (!class(df1$time2)[1] == "numeric") {
 stop("The time2 field is not a numeric object. Please fix.")
# MORE ERROR CHECKING:
# What if the user has NA or missing data?
# What if the crime problem is decreasing?
# Fun tasks for later...
# Set up parameters -----
# Set up initial parameters
count_Rt1 <- sum(df1$time1)</pre>
count_Rt2 <- sum(df1$time2)</pre>
chg_Rt1_Rt2 <- count_Rt2 - count_Rt1</pre>
pct_Rt1_Rt2 <- (chg_Rt1_Rt2 / count_Rt1) *100</pre>
# Add the field that has the volume of change, and order by it
df1 <- df1 %>%
 mutate (diff = time2 - time1) %>%
 mutate (diffPct = 100*(diff/time1)) %>%
 arrange(desc(diff))
# Grab some basic statistics here
numPositive <- length(which(df1$diff > 0))
numNeutral <- length(which(df1$diff == 0))</pre>
numNegative <- length(which(df1$diff < 0))</pre>
# Create the new data frame to hold the result
df2 <- data.frame(matrix(ncol =8, nrow = 0))</pre>
colnames(df2) <- c("unit", "adjusted", "Ut1", "Ut2", "Rt1", "Rt2", "chg", "pct")</pre>
df2 <- df2 %>%
 mutate(unit = as.character(unit)) %>%
 mutate(adjusted = as.numeric(adjusted)) %>%
 mutate(Ut1 = as.numeric(Ut1)) %>%
 mutate(Ut2 = as.numeric(Ut2)) %>%
 mutate(Rt1 = as.numeric(Rt1)) %>%
 mutate(Rt2 = as.numeric(Rt2)) %>%
```

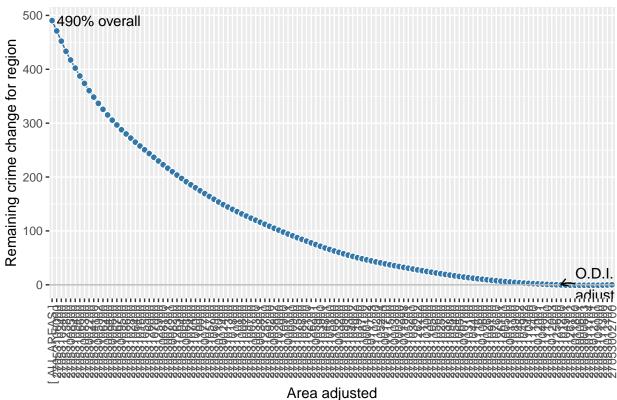
```
mutate(chg = as.numeric(chg)) %>%
 mutate(pct = as.numeric(pct))
# set up the initial row in the result data frame
df2 <- df2 %>% add_row(unit = "[ ALL AREAS ]", adjusted = 0,
                        Ut1 = 0, Ut2 = 0,
                        Rt1 = count_Rt1, Rt2 = count_Rt2,
                        chg = chg_Rt1_Rt2, pct = pct_Rt1_Rt2)
gain_from_row_removal <- row_to_remove <- NULL</pre>
# Loop through each row of the data
for (master_loop in 1:(source_rows)){
 df1 <- df1 %>% # order the data frame
    arrange(desc(diff))
  if (analysisMethod == "match"){
    #### 'Zero change the row' approach
    count_Rt1_temp <- count_Rt1</pre>
    count_Rt2_temp <- count_Rt2 - df1$diff[master_loop]</pre>
   pct_Rt1_Rt2 <- ((count_Rt1_temp - count_Rt2_temp) / count_Rt1) *100</pre>
 else { #analysisMethod == "remove"
    #### 'Remove entire row' approach, including remove t1 value
    count_Rt1_temp <- count_Rt1 - df1$time1[master_loop]</pre>
    count_Rt2_temp <- count_Rt2 - df1$time2[master_loop]</pre>
    pct_Rt1_Rt2 <- ((count_Rt1_temp - count_Rt2_temp) / count_Rt1) *100</pre>
 }
 row_to_remove <- 1 # Always row 1, but this is a legacy from
  # when I used a different approach...
  # Here, the row we are removing is
  # stored in row_to_remove
  if (analysisMethod == "remove"){
    #### Remove entire row approach
         This approach removes the impact of the area by subtracting
         both Rt1 and Rt2
    count_Rt1 <- count_Rt1 - df1$time1[row_to_remove]</pre>
    count_Rt2 <- count_Rt2 - df1$time2[row_to_remove]</pre>
    chg_Rt1_Rt2 <- count_Rt2 - count_Rt1</pre>
    pct_Rt1_Rt2 <- (chg_Rt1_Rt2 / count_Rt1) *100</pre>
    named_areas <- df1$unit[row_to_remove]</pre>
  if (analysisMethod == "match"){
    #### Zero change the row approach, as if Rt2 == Rt1 in the row
         The best row to remove is has been exhaustively calculated
         Here, the row we are removing is stored in row_to_remove
    count_Rt1 <- count_Rt1</pre>
    count_Rt2 <- count_Rt2 - df1$diff[row_to_remove]</pre>
```

```
chg_Rt1_Rt2 <- count_Rt2 - count_Rt1</pre>
   pct_Rt1_Rt2 <- (chg_Rt1_Rt2 / count_Rt1) *100</pre>
   named_areas <- df1$unit[row_to_remove]</pre>
  # Add result to the output data frame
 df2 <- df2 %>% add_row(unit = named_areas, adjusted = master_loop,
                        Ut1 = df1$time1[row_to_remove], Ut2 = df1$time2[row_to_remove],
                        Rt1 = count Rt1, Rt2 = count Rt2,
                        chg = chg_Rt1_Rt2, pct = pct_Rt1_Rt2)
  # Adjust the row we just used in one of two ways:
  # 1. remove the actual row entirely
  if (analysisMethod == "remove"){
   df1 <-df1[-c(row_to_remove), ]</pre>
  #2. adjust the Rt2 to match Rt1 resulting in a zero diff
  # but show that diff as < lowest diff in the data set so that
  # the program does not stall with too many zeros
 if (analysisMethod == "match"){
   df1$time2[row_to_remove] <- df1$time1[row_to_remove]</pre>
   df1$diff[row_to_remove] <- -999 # this should be changed to always less than
   # the lowest diff score in the data set
} # end master loop
# Calculate ODI and NCDI indices -----
NumContributed <- length(which(df2$chg > 0))
ODI <- NumContributed / source_rows</pre>
NCDI <- (numPositive - NumContributed) / source_rows</pre>
ODI.text <- paste("O.D.I. = ", format(ODI, digits = 3), "after \nadjusting",
                 NumContributed, "of the", source_rows, "units")
# Tidy up names for data frame -----
df2 <- df2 %>%
 rename(unit_t1 = Ut1, unit_t2 = Ut2, region_t1 = Rt1, region_t2 = Rt2)
# Plot -----
df3 <- df2
plot.adjustment <- ""</pre>
if (nrow(df3) > 151) {
 df3 <- df3[1:151, ]
 plot.adjustment <- "Plot only shows first\n100 areas adjusted"</pre>
p <- ggplot(df3, aes(x=reorder(unit, adjusted), y=pct, group = 1)) +</pre>
 geom_line(color="#3277a8") +
 geom_point(shape=21, color="white", fill="#3277a8", size=2) +
```

```
geom_hline(color="grey", yintercept=0) +
    labs(title="Dispersion of crime change",
         x ="Area adjusted", y = "Remaining crime change for region") +
      geom = "curve", x = NumContributed+4, y = 1.5,
     xend = NumContributed+1, yend = 0.2,
     curvature = .2, arrow = arrow(length = unit(2, "mm"))
    annotate(geom = "text", x = NumContributed+4.1, y = 1.5,
             label = ODI.text, hjust = "left") +
    theme(axis.text.x = element_text(angle = 90, vjust = 0.5, hjust=1)) +
    annotate(geom = "text", x = 2, y = df2\$pct[1],
             label = pasteO(format(df2$pct[1], digits = 3),"% overall"), hjust = "left") +
   theme(axis.text.x = element_text(angle = 90, vjust = 0.5, hjust=1))
  if (plot.adjustment != "") {
   p <- p +
      annotate(geom = "text", x = 100, y = df3$pct[1]-1, label = plot.adjustment, hjust = "right")
  }
 р
  # Create return list -----
 output <- list(df2, p, NumContributed, ODI, NCDI)</pre>
 return(output)
prepost_cj <- cj_exp %>%
  mutate(begin_date = ISOweek2date(paste(year, paste0("W", sprintf("%02d", week)), 1,sep = "-")),
         end_date = begin_date+weeks(1)-days(1),
         pre_post_floyd = ifelse(end_date <= as.Date("2020-05-25"), 0, 1)) %>%
  filter(end_date <= as.Date("2022-12-31")) %>%
  group_by(GEOID, pre_post_floyd) %>%
  summarize(car_jack = sum(car_jack, na.rm = T),
           total_pop = sum(total_pop, na.rm = T)) %>%
  mutate(car_jack_rate = car_jack/total_pop*1000) %>%
  select(GEOID, pre_post_floyd, car_jack, car_jack_rate) %>%
  st_drop_geometry() %>%
  pivot_wider(names_from = pre_post_floyd, values_from = c(car_jack, car_jack_rate)) %>%
 mutate(GEOID = as.character(GEOID))
## `summarise()` has grouped output by 'GEOID'. You can override using the
## `.groups` argument.
output <- crimedispersion(as.data.frame(prepost_cj), 'GEOID', 'car_jack_rate_0', 'car_jack_rate_1')
ouput_data <- output[[1]]</pre>
n_remove <- output[[3]]</pre>
odi <- output[[4]] #ratio of n removed to n overall</pre>
ncdi <- output[[5]] #ratio of areas not contributing to overall increase but still increase to overall
```

output[[2]]

Dispersion of crime change



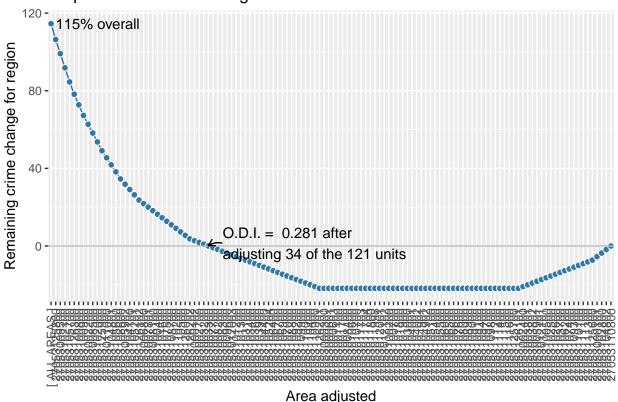
ggsave(filename = "Car Jacking/Figures for publication/fig5.png", bg="white", width = 10, height = 8)

Homicide

```
prepost_hom <- homicide %>%
 mutate(begin_date = ISOweek2date(paste(year, pasteO("W", sprintf("%02d", week)), 1,sep = "-")),
         end_date = begin_date+weeks(1)-days(1),
         pre_post_floyd = ifelse(end_date <= as.Date("2020-05-25"), 0, 1)) %%</pre>
  group_by(GEOID, pre_post_floyd) %>%
  summarize(homicide = sum(homicide, na.rm = T),
            total_pop = sum(total_pop, na.rm = T)) %>%
  mutate(homicide_rate = homicide/total_pop*1000) %>%
  st_drop_geometry() %>%
  select(GEOID, pre_post_floyd, homicide, homicide_rate) %>%
  pivot_wider(names_from = pre_post_floyd, values_from = c(homicide, homicide_rate)) %%
 mutate(GEOID = as.character(GEOID))
## `summarise()` has grouped output by 'GEOID'. You can override using the
## `.groups` argument.
output_homicide <- crimedispersion(as.data.frame(prepost_hom), 'GEOID', 'homicide_0', 'homicide_1')</pre>
ouput data <- output homicide[[1]]</pre>
n_remove <- output_homicide[[3]]</pre>
```

```
odi <- output_homicide[[4]] #ratio of n removed to n overall
ncdi <- output_homicide[[5]] #ratio of areas not contributing to overall increase but still increase to
output_homicide[[2]]</pre>
```

Dispersion of crime change

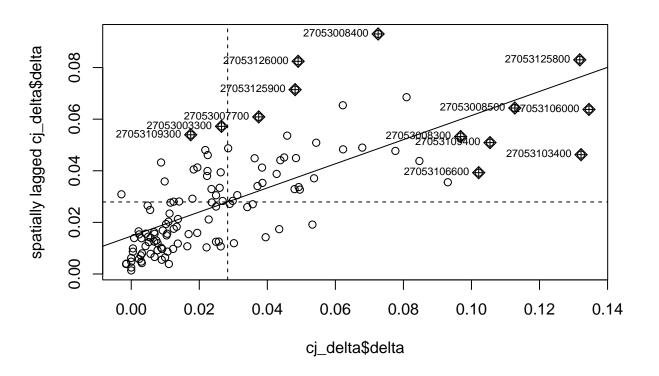


ggsave(filename = "Car Jacking/Figures for publication/fig6.png", bg="white", width = 10, height = 8)

Spatial Correlation Change in Carjackings and Homicide

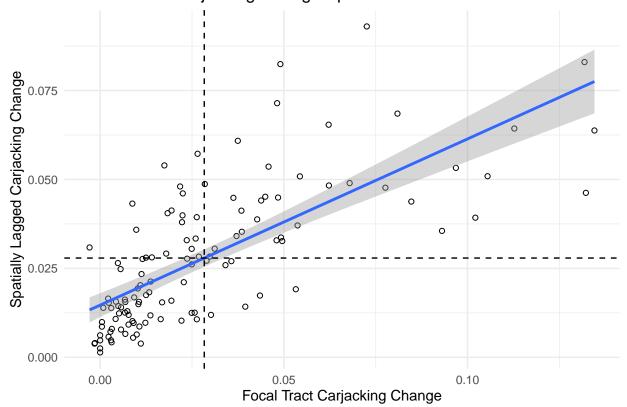
Carjacking

```
nb <- st_contiguity(cj_delta, queen=TRUE)</pre>
wt <- st_weights(nb, style = "W")
global_moran_test(
  cj_delta$delta,
  nb,
  wt,
  alternative = "greater",
  randomization = TRUE)
##
##
    Moran I test under randomisation
##
## data: x
## weights: listw
##
## Moran I statistic standard deviate = 8.9895, p-value < 2.2e-16
## alternative hypothesis: greater
   sample estimates:
## Moran I statistic
                            Expectation
                                                  Variance
##
         0.467232307
                           -0.008333333
                                               0.002798653
mp <- moran.plot(cj_delta$delta, nb2listw(nb),</pre>
                 labels = as.character(cj_delta$GEOID))
```



```
ggplot(mp, aes(x=x, y=wx)) +
geom_point(shape=1) +
```

Moran's I Plot Carjacking Change Spatialautocorrelation

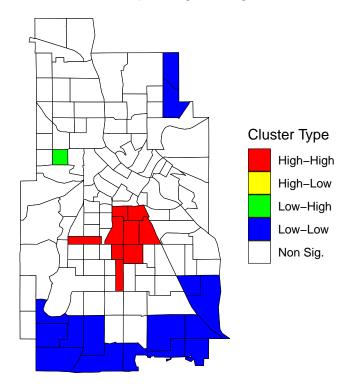


```
global_moran_test(
  cj_delta$car_jack_rate_1,
  nb,
  wt,
  alternative = "greater",
  randomization = TRUE)
```

```
##
## Moran I test under randomisation
##
## data: x
## weights: listw
##
## Moran I statistic standard deviate = 9.7351, p-value < 2.2e-16
## alternative hypothesis: greater
## sample estimates:</pre>
```

```
## Moran I statistic
                          Expectation
                                                  Variance
         0.506295621
                          -0.008333333
                                              0.002794527
##
#LISA
cj_lisa <- local_moran(cj_delta$delta,</pre>
                       nb = nb,
                       wt = wt,
                       nsim = 1000,
                       iseed = set.seed(7188)) %>%
 mutate(mean_p = ifelse(p_ii_sim <= 0.05, as.character(pysal), "Non Sig."),</pre>
         mean_p = factor(mean_p, levels = c("High-High", "High-Low", "Low-High",
                                             "Low-Low", "Non Sig.")))
cj_lisa_rate <- local_moran(cj_delta$car_jack_rate_1,</pre>
                       nb = nb,
                       wt = wt,
                       nsim = 1000,
                       iseed = set.seed(7188)) %>%
 mutate(mean_p = ifelse(p_ii_sim <= 0.05, as.character(pysal), "Non Sig."),</pre>
         mean_p = factor(mean_p, levels = c("High-High", "High-Low", "Low-High",
                                             "Low-Low", "Non Sig.")))
cj_delta %>%
  cbind(cj_lisa) %>%
  ggplot(aes(fill = mean_p)) +
  geom_sf() +
  geom_sf(lwd = 0.2, color = "black") +
  theme_void() +
  scale_fill_manual(values = c("red", "yellow", "green", "blue", "white"), drop = FALSE)+
  labs(title = "LISA Plot for Carjacking Change Pre/Post Police Murder",
       fill = "Cluster Type",
       caption = "Clusters significant at p < .05 with 1,000 simulations.")</pre>
```

LISA Plot for Carjacking Change Pre/Post Police Murder



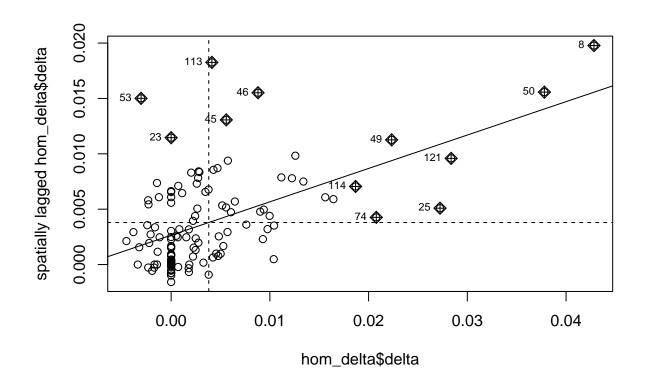
Clusters significant at p < .05 with 1,000 simulations.

```
ggsave(filename = "Car Jacking/Figures for publication/fig7.png", bg="white", width = 10, height = 8)
```

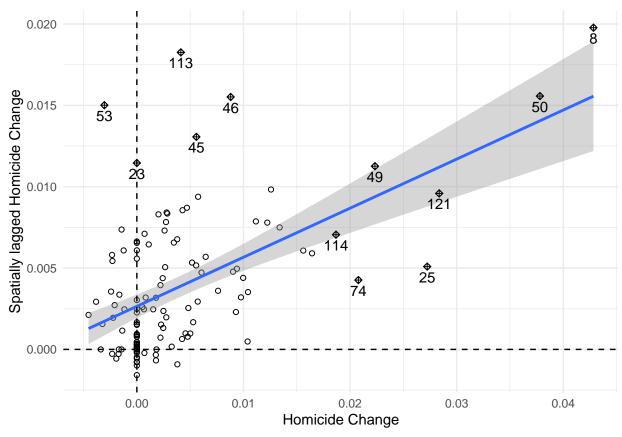
Homicide

weights: listw

```
hom_delta <- prepost_hom %>%
  mutate(delta = homicide_rate_1-homicide_rate_0,
           GEOID = as.numeric(GEOID)) %>%
  left_join(mpls_tract, by = "GEOID") %>%
  st_sf()
nb <- st_contiguity(hom_delta, queen=TRUE)</pre>
wt <- st_weights(nb, style = "W")</pre>
global_moran_test(
  hom_delta$delta,
  nb,
  wt,
 alternative = "greater",
 randomization = TRUE)
##
## Moran I test under randomisation
## data: x
```



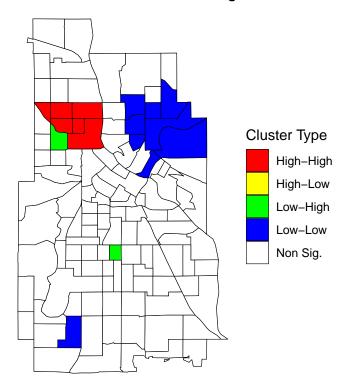
```
ggplot(mp, aes(x=x, y=wx)) +
  geom_point(shape=1) +
  geom_smooth(formula=y ~ x, method="lm") +
    #geom_hline(yintercept=mean(mp$wx), lty=2) +
  geom_hline(yintercept=0, lty=2) +
  #geom_vline(xintercept=mean(mp$x), lty=2) +
  geom_vline(xintercept=0, lty=2) +
  theme_minimal() +
  geom_point(data=mp[mp$is_inf,], aes(x=x, y=wx), shape=9) +
  geom_text(data=mp[mp$is_inf,], aes(x=x, y=wx, label=labels, vjust=1.5)) +
  xlab("Homicide Change") +
  ylab(paste0("Spatially lagged ", "Homicide Change"))
```



```
global_moran_test(
  hom_delta$homicide_1,
  nb,
  wt,
  alternative = "greater",
  randomization = TRUE)
##
##
    Moran I test under randomisation
##
## data: x
## weights: listw
##
## Moran I statistic standard deviate = 6.725, p-value = 8.777e-12
## alternative hypothesis: greater
## sample estimates:
## Moran I statistic
                            Expectation
                                                  Variance
##
         0.347800619
                           -0.008333333
                                               0.002804380
#LISA
hom_lisa <- local_moran(hom_delta$delta,</pre>
                        nb = nb,
                        wt = wt,
                        nsim = 1000,
                        iseed = set.seed(7188)) %>%
  mutate(mean_p = ifelse(p_ii_sim <= 0.05, as.character(pysal), "Non Sig."),</pre>
         mean_p = factor(mean_p, levels = c("High-High", "High-Low", "Low-High",
```

```
"Low-Low", "Non Sig.")))
hom_lisa_rate <- local_moran(hom_delta$homicide_1,</pre>
                        nb = nb,
                        wt = wt,
                        nsim = 1000,
                        iseed = set.seed(7188)) %>%
  mutate(mean_p = ifelse(p_ii_sim <= 0.05, as.character(pysal), "Non Sig."),</pre>
         mean_p = factor(mean_p, levels = c("High-High", "High-Low", "Low-High",
                                             "Low-Low", "Non Sig.")))
hom_delta %>%
  cbind(hom_lisa) %>%
  ggplot(aes(fill = mean_p)) +
  geom_sf() +
  geom_sf(lwd = 0.2, color = "black") +
  theme_void() +
  scale_fill_manual(values = c("red", "yellow", "green", "blue", "white"), drop = F)+
  labs(title = "LISA Plot for Homicide Change Pre/Post Police Murder",
       fill = "Cluster Type",
       caption = "Clusters significant at p < .05 with 1,000 simulations.")</pre>
```

LISA Plot for Homicide Change Pre/Post Police Murder



Clusters significant at p < .05 with 1,000 simulations.

```
ggsave(filename = "Car Jacking/Figures for publication/fig8.png", bg="white", width = 10, height = 8)
```

RE CJ Models

```
cj_exp_prepost <- cj_exp %>%
  group_by(GEOID) %>%
  mutate(begin_date = ISOweek2date(paste(year, paste0("W", sprintf("%02d", week)), 1,sep = "-")),
         end_date = begin_date+weeks(1)-days(1),
         post_floyd = as.numeric(begin_date >= as.Date("2020-05-25")),
         post_floyd_3 = as.numeric(begin_date >= as.Date("2020-05-25")+months(3)),
         weeks_post = as.numeric(begin_date-as.Date("2020-05-25"))/7,
         t_post_floyd = ifelse(weeks_post >=0,
                               weeks post,
                               0),
        # stay at home = as.numeric(begin date >= as.Date("2020-03-28") &
         #state_of_emerg = as.numeric(begin_date >= as.Date("2020-03-13")),
        period = factor(case when(
          post floyd==0 & post floyd 3==0 ~ "Pre-Killing",
          post floyd>=1 & post floyd 3==0 ~ "0-3 Months Post-Killing",
          post_floyd>=1 & post_floyd_3>=1 ~ "3+ Months Post-Killing"),
           levels = c("Pre-Killing", "0-3 Months Post-Killing", "3+ Months Post-Killing")),
        GEOID = as.character(GEOID),
        anyjack = ifelse(car_jack==0, 0, 1),
        t = 1:length(car_jack_rate)) %>%
  left_join(acs_2020, by = c("GEOID")) \%>%
  select(-med_hh_inc) %>%
  drop_na()
library(lme4)
## Warning: package 'lme4' was built under R version 4.2.3
library(lmerTest)
library(lavaan)
## Warning: package 'lavaan' was built under R version 4.2.3
cd_model_1 <- ' cd =~ unemp_rate + pov_rate + female_hh_rate + no_hs_dip_rate + black_perc</pre>
                 black_perc ~~ unemp_rate'
cfa_cd <- cfa(cd_model_1, data = cj_exp_prepost, std.lv = T)</pre>
modificationindices(cfa cd)
##
                                                  epc sepc.lv sepc.all sepc.nox
                lhs op
                                   rhs
                                            mi
                             pov_rate 943.629 7.286
## 13
         unemp rate ~~
                                                      7.286
                                                                0.198
                                                                         0.198
## 14
         unemp_rate ~~ female_hh_rate 704.567 -1.273 -1.273
                                                                -0.200
                                                                        -0.200
         unemp_rate ~~ no_hs_dip_rate
## 15
                                       6.434 0.255
                                                      0.255
                                                                 0.021
                                                                         0.021
                                                               -0.122
                                                                        -0.122
## 16
           pov_rate ~~ female_hh_rate 443.948 -2.705 -2.705
## 17
           pov_rate ~~ no_hs_dip_rate 58.088 1.936
                                                      1.936
                                                                0.046
                                                                         0.046
## 18
                           black_perc 110.356 -8.406 -8.406
                                                               -0.144
                                                                         -0.144
           pov_rate ~~
## 19 female_hh_rate ~~ no_hs_dip_rate 105.301 0.598
                                                      0.598
                                                                0.082
                                                                          0.082
## 20 female_hh_rate ~~
                          black_perc 719.519 5.183 5.183
                                                                0.514
                                                                         0.514
## 21 no_hs_dip_rate ~~
                          black_perc 296.794 -7.281 -7.281 -0.382 -0.382
```

summary(cfa_cd, fit.measures=TRUE, standardized = T)

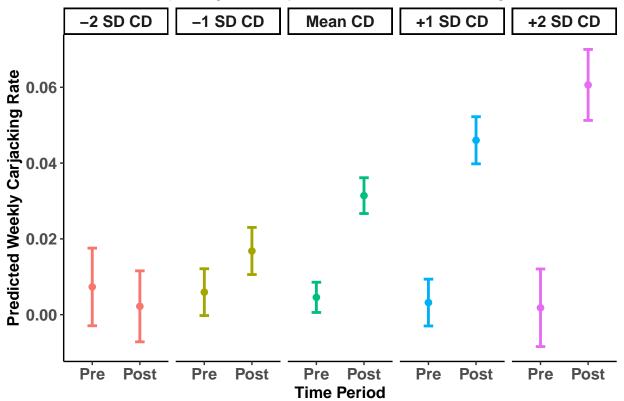
```
## lavaan 0.6.15 ended normally after 32 iterations
##
##
     Estimator
                                                         ML
     Optimization method
##
                                                     NLMINB
##
     Number of model parameters
                                                         11
##
##
     Number of observations
                                                      38357
##
## Model Test User Model:
##
     Test statistic
                                                   1610.419
##
##
     Degrees of freedom
                                                      0.000
     P-value (Chi-square)
##
##
## Model Test Baseline Model:
                                                  96458.724
##
     Test statistic
     Degrees of freedom
##
                                                         10
     P-value
                                                      0.000
##
##
## User Model versus Baseline Model:
##
     Comparative Fit Index (CFI)
                                                      0.983
##
     Tucker-Lewis Index (TLI)
##
                                                      0.958
##
## Loglikelihood and Information Criteria:
##
     Loglikelihood user model (HO)
##
                                               -606558.430
     Loglikelihood unrestricted model (H1)
##
                                               -605753.221
##
##
     Akaike (AIC)
                                               1213138.861
                                               1213232.962
##
     Bayesian (BIC)
##
     Sample-size adjusted Bayesian (SABIC)
                                               1213198.004
##
## Root Mean Square Error of Approximation:
##
##
    RMSEA
                                                      0.102
                                                      0.098
##
     90 Percent confidence interval - lower
     90 Percent confidence interval - upper
##
                                                      0.107
     P-value H_0: RMSEA <= 0.050
                                                      0.000
##
     P-value H_0: RMSEA >= 0.080
##
                                                      1.000
##
## Standardized Root Mean Square Residual:
##
##
     SRMR
                                                      0.029
##
## Parameter Estimates:
##
##
     Standard errors
                                                   Standard
##
     Information
                                                   Expected
##
     Information saturated (h1) model
                                                Structured
##
```

```
## Latent Variables:
##
                      Estimate Std.Err z-value P(>|z|)
                                                            Std.lv Std.all
##
     cd =~
                                  0.023 148.576
                                                     0.000
                                                              3.428
                                                                       0.726
##
                         3.428
       unemp_rate
##
       pov rate
                         8.295
                                  0.067 123.883
                                                     0.000
                                                              8.295
                                                                       0.590
##
                         2.155
                                  0.013 164.248
                                                    0.000
                                                                       0.740
       female hh rate
                                                              2.155
##
                         4.667
                                  0.026 177.181
                                                     0.000
                                                              4.667
                                                                       0.783
       no hs dip rate
                                  0.074 236.523
                                                     0.000
                                                             17.544
                                                                       0.960
##
       black_perc
                        17.544
##
## Covariances:
##
                      Estimate Std.Err z-value P(>|z|)
                                                             Std.lv Std.all
##
    .unemp_rate ~~
                        -9.264
                                  0.237 - 39.009
                                                     0.000
                                                             -9.264
##
      .black_perc
                                                                      -0.555
##
## Variances:
##
                      Estimate Std.Err z-value P(>|z|)
                                                             Std.lv Std.all
##
                        10.512
                                  0.103 101.640
                                                     0.000
                                                            10.512
                                                                       0.472
      .unemp_rate
##
                       128.929
                                  0.964 133.780
                                                     0.000 128.929
                                                                       0.652
      .pov rate
                         3.844
                                                             3.844
##
                                  0.031 122.692
                                                     0.000
                                                                       0.453
      .female_hh_rate
                                  0.120 114.796
##
      .no hs dip rate
                        13.746
                                                     0.000
                                                             13.746
                                                                       0.387
##
      .black_perc
                        26.463
                                  1.011 26.187
                                                     0.000
                                                             26.463
                                                                       0.079
##
                         1.000
                                                              1.000
                                                                       1.000
cd_predict <- as.vector(lavPredict(cfa_cd, newdata = as.data.frame(cj_exp_prepost)))</pre>
cj_exp_prepost$conc_dis <- cd_predict</pre>
#predicted probability/rate plots
  #over time
  #stratified by CD (mean, +1SD, -1SD, +2SD, -2SD)
re <- lmer(car_jack_rate~t+post_floyd+t_post_floyd+
             conc dis+
             age_19_29_perc+age_30_49_perc+age_50_69_perc+
             age_70_plus_perc+ post_floyd:conc_dis+t_post_floyd:conc_dis+
             (1|GEOID),
          data = cj_exp_prepost)
summary(re)
## Linear mixed model fit by REML. t-tests use Satterthwaite's method [
## lmerModLmerTest]
## Formula:
## car_jack_rate ~ t + post_floyd + t_post_floyd + conc_dis + age_19_29_perc +
##
       age_30_49_perc + age_50_69_perc + age_70_plus_perc + post_floyd:conc_dis +
##
       t post floyd:conc dis + (1 | GEOID)
##
      Data: cj_exp_prepost
##
## REML criterion at convergence: -71411.7
## Scaled residuals:
              10 Median
                            3Q
## -1.255 -0.254 -0.120 0.018 31.639
## Random effects:
## Groups
            Name
                         Variance Std.Dev.
```

```
## GEOID
            (Intercept) 0.0001701 0.01304
                        0.0090060 0.09490
## Residual
## Number of obs: 38357, groups: GEOID, 121
##
## Fixed effects:
##
                         Estimate Std. Error
                                                     df t value Pr(>|t|)
## (Intercept)
                         5.038e-02 1.649e-02 1.162e+02 3.054 0.00280 **
                                                        1.204 0.22858
                         1.490e-05 1.238e-05 3.823e+04
## t
                         2.684e-02 1.954e-03 3.823e+04 13.736
## post_floyd
                                                                 < 2e-16 ***
## t_post_floyd
                        -1.775e-05 2.237e-05 3.823e+04 -0.794 0.42745
## conc_dis
                        4.779e-04 2.031e-03 1.261e+02
                                                        0.235
                                                                0.81436
## age_19_29_perc
                        -4.629e-04 1.690e-04 1.150e+02 -2.738
                                                                 0.00716 **
## age_30_49_perc
                        -3.520e-04 2.709e-04 1.150e+02 -1.299
                                                                 0.19652
## age_50_69_perc
                        -1.033e-03 3.113e-04 1.150e+02 -3.318 0.00121 **
                        -6.118e-04 3.773e-04 1.150e+02 -1.622
                                                                 0.10760
## age_70_plus_perc
## post_floyd:conc_dis
                         1.598e-02 1.630e-03 3.823e+04
                                                         9.805
                                                                < 2e-16 ***
## t_post_floyd:conc_dis -2.725e-05 1.897e-05 3.823e+04 -1.436 0.15097
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Correlation of Fixed Effects:
              (Intr) t
##
                            pst_fl t_pst_ cnc_ds a_19_2 a_30_4 a_50_6 a_70__
              -0.068
## t
             0.026 - 0.573
## post floyd
## t_post_flyd 0.038 -0.553 -0.223
## conc dis
              -0.724 0.000 0.000 0.000
## ag_19_29_pr -0.936 0.000 0.000 0.000 0.695
## ag_30_49_pr -0.884 0.000 0.000 0.000 0.597
                                                 0.767
## ag_50_69_pr -0.749 0.000 0.000 0.000 0.618 0.724
                                                        0.435
## ag_70_pls_p -0.300 0.000
                            0.000 0.000 0.190 0.279
                                                        0.276 - 0.127
## pst_flyd:c_ 0.000 0.000
                            0.000 0.000 -0.129
                                                 0.000
                                                        0.000
                                                              0.000 0.000
## t_pst_fly:_ 0.000 0.000 0.000 0.000 0.000 0.000 0.000 0.000 0.000
##
              pst_:_
## t
## post floyd
## t_post_flyd
## conc dis
## ag_19_29_pr
## ag_30_49_pr
## ag_50_69_pr
## ag 70 pls p
## pst flyd:c
## t_pst_fly:_ -0.792
library(emmeans)
pred_raw <- emmeans(re, c("conc_dis", "post_floyd"),</pre>
                   at = list(conc_dis = c(-2, -1, 0, 1, 2),
                             post_floyd = c(0,1),
                             t_post_floyd = mean(cj_exp_prepost$t_post_floyd[cj_exp_prepost$post_floyd
                             t = mean(cj_exp_prepost$t[cj_exp_prepost$post_floyd==0]))) %>%
 as.data.frame() %>%
 mutate(conc_dis = factor(case_when(
    conc_dis==-2~"-2 SD CD",
```

```
conc_dis==-1~"-1 SD CD",
   conc_dis==0~"Mean CD",
   conc_dis==1~"+1 SD CD",
   conc_dis==2~"+2 SD CD"),
   levels = c("-2 SD CD",
               "-1 SD CD",
               "Mean CD",
               "+1 SD CD",
               "+2 SD CD")),
   post_floyd = factor(case_when(
   post_floyd==1~"Post",
   post_floyd==0~"Pre"),
   levels = c("Pre", "Post")))
ggplot(pred_raw, aes(x = post_floyd, y = emmean, color = conc_dis))+
 geom_point(size = 2)+
  geom_errorbar(aes(ymin=asymp.LCL, ymax=asymp.UCL), width = .2, size = 1)+
  labs(title = "Predicted Carjacking Rates by Concentrated Disadvantage, Pre- and Post-Killing",
      y = "Predicted Weekly Carjacking Rate",
      x = "Time Period") +
 facet_grid(~conc_dis)+
  guides(color = "none")+
 theme_classic()+
 theme(axis.text = element_text(face = "bold", size = 12),
        axis.title.x = element_text(face = "bold", size = 12),
       axis.title.y = element_text(face = "bold", size = 12),
        strip.text.x = element_text(face = "bold", size = 12))
```





```
ggsave(filename = "Car Jacking/Figures for publication/fig9.png", bg="white", width = 8, height = 8)
re_logit_cd <- glmer(anyjack ~ t+post_floyd+t_post_floyd+conc_dis+</pre>
             age_19_29_perc+age_30_49_perc+age_50_69_perc+
             age_70_plus_perc+ post_floyd:conc_dis+t_post_floyd:conc_dis+
             (1 | GEOID),
                  data = cj_exp_prepost, family = binomial)
## Warning in checkConv(attr(opt, "derivs"), opt$par, ctrl = control$checkConv, :
## Model failed to converge with max|grad| = 0.00861892 (tol = 0.002, component 1)
## Warning in checkConv(attr(opt, "derivs"), opt$par, ctrl = control$checkConv, : Model is nearly unide:
## - Rescale variables?; Model is nearly unidentifiable: large eigenvalue ratio
## - Rescale variables?
summary(re_logit_cd)
## Generalized linear mixed model fit by maximum likelihood (Laplace
     Approximation) [glmerMod]
##
   Family: binomial (logit)
##
  Formula: anyjack ~ t + post_floyd + t_post_floyd + conc_dis + age_19_29_perc +
##
       age_30_49_perc + age_50_69_perc + age_70_plus_perc + post_floyd:conc_dis +
       t_post_floyd:conc_dis + (1 | GEOID)
##
##
      Data: cj_exp_prepost
##
```

38345

logLik deviance df.resid

##

BIC

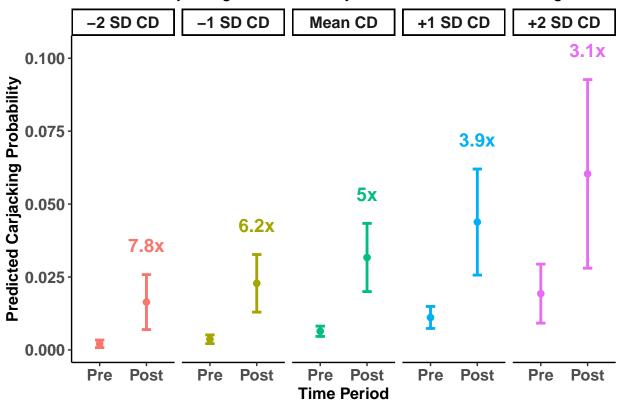
11182.8 11285.4 -5579.4 11158.8

```
## Scaled residuals:
##
      Min
           1Q Median
                               30
                                      Max
## -0.7196 -0.2197 -0.1384 -0.0726 21.5221
## Random effects:
                      Variance Std.Dev.
## Groups Name
## GEOID (Intercept) 0.4674 0.6837
## Number of obs: 38357, groups: GEOID, 121
##
## Fixed effects:
                          Estimate Std. Error z value Pr(>|z|)
## (Intercept)
                        -4.4947951 0.9185632 -4.893 9.92e-07 ***
## t
                         0.0038399 0.0011363
                                               3.379 0.000727 ***
                         1.6256029   0.1310456   12.405   < 2e-16 ***
## post_floyd
                        -0.0031597 0.0013849 -2.281 0.022519 *
## t_post_floyd
## conc_dis
                         0.6378147
                                   0.1190326
                                               5.358 8.40e-08 ***
## age_19_29_perc
                        -0.0028888 0.0093469
                                              -0.309 0.757271
## age_30_49_perc
                         0.0149162 0.0149808
                                               0.996 0.319402
## age_50_69_perc
                        -0.0491073 0.0173871
                                              -2.824 0.004738 **
## age_70_plus_perc
                        -0.0147769 0.0208016
                                              -0.710 0.477472
## post_floyd:conc_dis
                        -0.2215576  0.0776478  -2.853  0.004326 **
## t_post_floyd:conc_dis -0.0011664 0.0007388 -1.579 0.114378
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##
              (Intr) t
                            pst_fl t_pst_ cnc_ds a_19_2 a_30_4 a_50_6 a_70__
## t
              -0.124
              0.048 -0.699
## post_floyd
## t_post_flyd 0.101 -0.820 0.335
## conc_dis
              -0.714 0.002 0.122 0.000
## ag_19_29_pr -0.928  0.000 -0.002  0.001  0.675
## ag_30_49_pr -0.878  0.000 -0.001  0.001  0.584
                                                 0.766
## ag_50_69_pr -0.739 0.000 -0.002 0.000 0.581 0.720 0.431
## ag_70_pls_p -0.278  0.000 -0.005 -0.001  0.167  0.262  0.264 -0.147
## pst_flyd:c_ 0.026 -0.003 -0.303 0.119 -0.298 0.001 0.001 0.000 0.011
## t_pst_fly:_ 0.001 0.000 0.127 -0.172 -0.006 -0.001 -0.001 -0.001 0.004
##
              pst_:_
## t
## post_floyd
## t post flyd
## conc dis
## ag_19_29_pr
## ag_30_49_pr
## ag_50_69_pr
## ag_70_pls_p
## pst_flyd:c_
## t_pst_fly:_ -0.631
## optimizer (Nelder_Mead) convergence code: 0 (OK)
## Model failed to converge with max|grad| = 0.00861892 (tol = 0.002, component 1)
## Model is nearly unidentifiable: very large eigenvalue
## - Rescale variables?
## Model is nearly unidentifiable: large eigenvalue ratio
## - Rescale variables?
```

```
pred_prob_raw <- emmeans(re_logit_cd, c("conc_dis", "post_floyd"),</pre>
                    at = list(conc_dis = c(-2, -1, 0, 1, 2),
                              post_floyd = c(0,1),
                              t_post_floyd = mean(cj_exp_prepost$t_post_floyd[cj_exp_prepost$post_floyd
                              t = mean(cj_exp_prepost$t[cj_exp_prepost$post_floyd==0])),
                    trans = "response") %>%
  as.data.frame() %>%
  mutate(conc dis = factor(case when(
    conc dis==-2~"-2 SD CD",
    conc_dis==-1^{-"}-1 SD CD",
    conc_dis==0~"Mean CD",
    conc_dis==1~"+1 SD CD",
    conc_dis==2~"+2 SD CD"),
    levels = c("-2 SD CD",
               "-1 SD CD",
               "Mean CD",
               "+1 SD CD",
               "+2 SD CD")),
    post_floyd = factor(case_when(
    post_floyd==1~"Post",
    post_floyd==0~"Pre"),
    levels = c("Pre", "Post")))
ratios <- pred_prob_raw %>%
  select(conc dis, post floyd, prob) %>%
  pivot_wider(names_from = post_floyd, values_from = prob) %>%
  mutate(ratio = Post/Pre)
pred_prob_raw <- pred_prob_raw %>%
  left_join(ratios, by = "conc_dis") %>%
  mutate(ratio = ifelse(post_floyd=="Pre", NA_integer_, round(ratio,1)),
         ratio = ifelse(is.na(ratio), ratio, paste0(ratio, "x")))
ggplot(pred_prob_raw, aes(x = post_floyd, y = prob, color = conc_dis))+
  geom_point(size = 2)+
  geom_errorbar(aes(ymin=asymp.LCL, ymax=asymp.UCL), width = .2, size = 1)+
  labs(title = "Predicted Carjacking Probabilities by Concentrated Disadvantage, Pre- and Post-Killing"
       y = "Predicted Carjacking Probability",
       x = "Time Period") +
  theme_classic()+
  facet_grid(~conc_dis)+
  guides(color = "none")+
  geom_text(aes(x = post_floyd, y = asymp.UCL+.01, label = ratio),
            fontface = "bold", size = 5)+
  theme(axis.text = element_text(face = "bold", size = 12),
        axis.title.x = element_text(face = "bold", size = 12),
        axis.title.y = element_text(face = "bold", size = 12),
        strip.text.x = element_text(face = "bold", size = 12))
```

Warning: Removed 5 rows containing missing values (`geom_text()`).



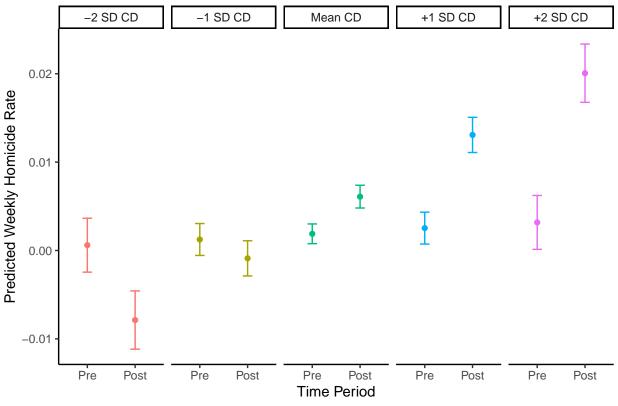


```
ggsave(filename = "Car Jacking/Figures for publication/fig10.png", bg="white", width = 8, height = 8)
## Warning: Removed 5 rows containing missing values (`geom_text()`).
homicide <- homicide %>% st_drop_geometry()
cj_exp_prepost <- cj_exp_prepost %>%
 mutate(GEOID = as.numeric(GEOID)) %>%
  left_join(homicide, by = c("GEOID", "year", "week")) %>%
  mutate(anyhom = ifelse(homicide==0, 0, 1))
re_homicide <- lmer(homicide_rate~t+post_floyd+t_post_floyd+conc_dis+
             age_19_29_perc+age_30_49_perc+age_50_69_perc+
             age_70_plus_perc+ post_floyd:conc_dis+t_post_floyd:conc_dis+
             (1|GEOID),
          data = cj_exp_prepost)
summary(re_homicide)
## Linear mixed model fit by REML. t-tests use Satterthwaite's method [
## lmerModLmerTest]
## Formula:
## homicide_rate ~ t + post_floyd + t_post_floyd + conc_dis + age_19_29_perc +
       age_30_49_perc + age_50_69_perc + age_70_plus_perc + post_floyd:conc_dis +
##
##
       t_post_floyd:conc_dis + (1 | GEOID)
##
      Data: cj_exp_prepost
##
## REML criterion at convergence: -131581.4
```

```
##
## Scaled residuals:
     Min
             1Q Median
## -1.095 -0.108 -0.033 -0.007 44.509
## Random effects:
  Groups
            Name
                       Variance Std.Dev.
            (Intercept) 1.527e-05 0.003907
## GEOID
## Residual
                       1.879e-03 0.043348
## Number of obs: 38357, groups: GEOID, 121
## Fixed effects:
                         Estimate Std. Error
                                                   df t value Pr(>|t|)
## (Intercept)
                        7.689e-03 5.403e-03 1.174e+02
                                                      1.423 0.15735
## t
                        1.165e-05 5.653e-06 3.823e+04
                                                        2.061 0.03929 *
## post_floyd
                        4.206e-03 8.925e-04
                                            3.823e+04
                                                       4.713 2.45e-06 ***
## t_post_floyd
                       -3.297e-05 1.022e-05 3.823e+04 -3.226 0.00126 **
## conc dis
                        1.681e-03 6.778e-04 1.371e+02
                                                        2.480
                                                               0.01434 *
## age_19_29_perc
                       -5.925e-05 5.524e-05
                                            1.150e+02 -1.073
                                                              0.28567
## age_30_49_perc
                       -1.355e-04 8.853e-05
                                            1.150e+02 -1.530
## age_50_69_perc
                        2.834e-05 1.017e-04 1.150e+02
                                                       0.279 0.78100
                       -2.832e-04 1.233e-04 1.150e+02 -2.297 0.02343 *
## age_70_plus_perc
                                                       8.514 < 2e-16 ***
## post_floyd:conc_dis
                        6.338e-03 7.444e-04 3.823e+04
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##
              (Intr) t
                           pst_fl t_pst_ cnc_ds a_19_2 a_30_4 a_50_6 a_70__
## t
              -0.095
## post_floyd
             0.036 - 0.573
## t_post_flyd 0.052 -0.553 -0.223
## conc_dis
             -0.707 0.000 0.000
                                  0.000
## ag_19_29_pr -0.934 0.000 0.000
                                  0.000
                                        0.681
## ag_30_49_pr -0.881
                    0.000
                           0.000
                                  0.000 0.585
                                               0.767
## ag_50_69_pr -0.747 0.000 0.000 0.000 0.605 0.724
                                                      0.435
## ag_70_pls_p -0.299 0.000
                           0.000
                                  0.000 0.186 0.279
                                                      0.276 - 0.127
## pst_flyd:c_ 0.000 0.000 0.000 0.000 -0.177
                                               0.000
                                                      0.000 0.000 0.000
## t_pst_fly:_ 0.000 0.000 0.000 0.000 0.000 0.000 0.000 0.000 0.000
##
             pst_:_
## t
## post_floyd
## t_post_flyd
## conc_dis
## ag_19_29_pr
## ag_30_49_pr
## ag_50_69_pr
## ag_70_pls_p
## pst_flyd:c_
## t_pst_fly:_ -0.792
pred_raw_homicide <- emmeans(re_homicide, c("conc_dis", "post_floyd"),</pre>
                  at = list(conc_dis = c(-2, -1, 0, 1, 2),
                            post_floyd = c(0,1)) %>%
```

```
as.data.frame() %>%
  mutate(conc_dis = factor(case_when(
    conc_dis==-2~"-2 SD CD",
    conc_dis==-1~"-1 SD CD",
   conc_dis==0~"Mean CD",
    conc_dis==1~"+1 SD CD",
   conc_dis==2~"+2 SD CD"),
   levels = c("-2 SD CD",
               "-1 SD CD".
               "Mean CD".
               "+1 SD CD",
               "+2 SD CD")),
   post_floyd = factor(case_when(
   post_floyd==1~"Post",
   post_floyd==0~"Pre"),
   levels = c("Pre", "Post")))
## Note: D.f. calculations have been disabled because the number of observations exceeds 3000.
## To enable adjustments, add the argument 'pbkrtest.limit = 38357' (or larger)
## [or, globally, 'set emm_options(pbkrtest.limit = 38357)' or larger];
## but be warned that this may result in large computation time and memory use.
## Note: D.f. calculations have been disabled because the number of observations exceeds 3000.
## To enable adjustments, add the argument 'lmerTest.limit = 38357' (or larger)
## [or, globally, 'set emm_options(lmerTest.limit = 38357)' or larger];
## but be warned that this may result in large computation time and memory use.
ggplot(pred_raw_homicide, aes(x = post_floyd, y = emmean, color = conc_dis))+
 geom_point()+
  geom errorbar(aes(ymin=asymp.LCL, ymax=asymp.UCL), width=0.2)+
  labs(title = "Predicted Homicide Rates by Concentrated Disadvantage, Pre- and Post-Killing",
      y = "Predicted Weekly Homicide Rate",
      x = "Time Period")+
  theme_classic()+
  facet_grid(~conc_dis)+
  guides(color = "none")
```

Predicted Homicide Rates by Concentrated Disadvantage, Pre- and Post-



```
re_logit_cd_homicide <- glmer(anyhom ~ t+post_floyd+t_post_floyd+conc_dis+
             age_19_29_perc+age_30_49_perc+age_50_69_perc+
             age_70_plus_perc+ post_floyd:conc_dis+t_post_floyd:conc_dis+
             (1|GEOID),
                  data = cj_exp_prepost, family = binomial)
## Warning in checkConv(attr(opt, "derivs"), opt$par, ctrl = control$checkConv, :
## Model failed to converge with max|grad| = 0.267118 (tol = 0.002, component 1)
## Warning in checkConv(attr(opt, "derivs"), opt$par, ctrl = control$checkConv, : Model is nearly unide:
## - Rescale variables?; Model is nearly unidentifiable: large eigenvalue ratio
## - Rescale variables?
summary(re_logit_cd_homicide)
## Generalized linear mixed model fit by maximum likelihood (Laplace
##
     Approximation) [glmerMod]
   Family: binomial (logit)
##
## Formula: anyhom ~ t + post_floyd + t_post_floyd + conc_dis + age_19_29_perc +
##
       age_30_49_perc + age_50_69_perc + age_70_plus_perc + post_floyd:conc_dis +
##
       t_post_floyd:conc_dis + (1 | GEOID)
##
     Data: cj_exp_prepost
```

38345

logLik deviance df.resid

3433.5

Max

3Q

##

##

##

##

AIC

Scaled residuals:

3457.5

Min

BIC

1Q Median

3560.2 -1716.8

```
## -0.338 -0.092 -0.061 -0.042 35.080
##
## Random effects:
## Groups Name
                      Variance Std.Dev.
## GEOID (Intercept) 0.6719 0.8197
## Number of obs: 38357, groups: GEOID, 121
## Fixed effects:
##
                         Estimate Std. Error z value Pr(>|z|)
## (Intercept)
                        -7.959083
                                  1.318054 -6.039 1.56e-09 ***
## t
                         0.004882
                                  0.001890 2.583 0.00981 **
## post_floyd
                                             3.172 0.00152 **
                         0.777797
                                   0.245234
                        -0.007360
                                  0.002811 -2.618 0.00884 **
## t_post_floyd
                         0.957280
## conc_dis
                                  0.177431 5.395 6.84e-08 ***
## age_19_29_perc
                         0.019009
                                  0.013565
                                             1.401 0.16112
## age_30_49_perc
                         0.027486
                                   0.021594
                                              1.273 0.20308
## age_50_69_perc
                        0.026644
                                  0.025918
                                             1.028 0.30394
## age_70_plus_perc
                       -0.047473
                                   0.032278 -1.471 0.14136
## post_floyd:conc_dis
                        0.162234
                                   0.137858
                                             1.177 0.23927
## t_post_floyd:conc_dis -0.002923
                                   0.001564 -1.869 0.06164 .
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##
              (Intr) t
                            pst_fl t_pst_ cnc_ds a_19_2 a_30_4 a_50_6 a_70__
              -0.148
## post_floyd 0.044 -0.599
## t_post_flyd 0.099 -0.672 -0.001
## conc_dis
              -0.738 0.001 0.129 0.007
## ag_19_29_pr -0.913 -0.001 0.002 -0.001 0.689
## ag_30_49_pr -0.851 -0.001 0.002 -0.001 0.579 0.724
## ag_50_69_pr -0.693 -0.001 0.001 0.000 0.563 0.666 0.335
## ag_70_pls_p -0.256  0.000  0.004 -0.006  0.134  0.241  0.252 -0.202
## pst_flyd:c_ 0.033 -0.001 -0.468 0.295 -0.289 -0.003 -0.003 -0.001 -0.009
## t_pst_fly:_ 0.000 0.000 0.297 -0.404 -0.018 0.003 0.002 0.000 0.014
##
              pst_:_
## t
## post_floyd
## t_post_flyd
## conc_dis
## ag 19 29 pr
## ag_30_49_pr
## ag_50_69_pr
## ag_70_pls_p
## pst_flyd:c_
## t_pst_fly:_ -0.652
## optimizer (Nelder_Mead) convergence code: 0 (OK)
## Model failed to converge with max|grad| = 0.267118 (tol = 0.002, component 1)
## Model is nearly unidentifiable: very large eigenvalue
## - Rescale variables?
## Model is nearly unidentifiable: large eigenvalue ratio
## - Rescale variables?
```

```
class(re) <- "lmerMod"</pre>
class(re_logit_cd) <- "lmerMod"</pre>
stargazer(re, re_logit_cd,
          title = "Interrupted Time Series Models of Carjackings, MPLS 2017-2022",
          covariate.labels = c("T", "Post-Killing", "T Post-Killing",
                               "Conc. Dis.", "Age 19-29", "Age 30-49",
                               "Age 50-69", "Age 70+",
                               "Post-Killing X Conc. Dis.",
                               "T Post-Killing X Conc. Dis."),
          header = F,
          dep.var.caption = "Carjacking",
          dep.var.labels = c("Rate per 1,000", "Any Carjacking"),
          model.names = FALSE,
          column.labels = c("RE HLM", "RE Logit"),
          report = "vcs",
          ci=TRUE,
          ci.level=0.95,
          ci.separator = "|",
          notes = "95\\% Confidence Intervals in parentheses",
          single.row = F,
          omit.stat = c("adj.rsq", "aic", "bic"),
          \#star.cutoffs = c(.05, .01, .001), star.char = c("*", "**", "***"),
          add.lines = list(c("SD(Tract)", .013, .674),
                           c("SD(Residual)", .094, "")),
          notes.append = F)
```

Spatial Panel Models

```
library(splm)
library(plm)
## Attaching package: 'plm'
## The following object is masked from 'package:lavaan':
##
##
       nobs
## The following objects are masked from 'package:dplyr':
##
##
       between, lag, lead
nb <- st_contiguity(mpls_tract, queen=TRUE)</pre>
wt <- st_weights(nb, style = "W", allow_zero = T)</pre>
w <- recreate_listw(nb, wt)</pre>
cj_exp_prepost <- cj_exp_prepost %>%
 group_by(GEOID) %>%
 arrange(year, week) %>%
 mutate(WEEKID = row_number())
#write_csv(cj_exp_prepost, "Car Jacking/cj_exp_prepost.csv")
```

Table 1: Interrupted Time Series Models of Carjackings, MPLS 2017-2022

	Carjacking	
	Rate per 1,000 RE HLM	Any Carjacking RE Logit
	(1)	(2)
T	$0.00001 \\ (-0.00001 0.00004)$	$0.004 \\ (0.002 0.006)$
Post-Killing	$0.027 \\ (0.023 0.031)$	$1.626 \\ (1.369 1.882)$
T Post-Killing	$ \begin{array}{c} -0.00002 \\ (-0.0001 0.00003) \end{array} $	$ \begin{array}{c} -0.003 \\ (-0.006 -0.0004) \end{array} $
Conc. Dis.	$0.0005 \\ (-0.004 0.004)$	$0.638 \\ (0.405 0.871)$
Age 19-29	$ \begin{array}{c} -0.0005 \\ (-0.001 -0.0001) \end{array} $	$ \begin{array}{c} -0.003 \\ (-0.021 0.015) \end{array} $
Age 30-49	$ \begin{array}{c} -0.0004 \\ (-0.001 0.0002) \end{array} $	$0.015 \\ (-0.014 0.044)$
Age 50-69	$ \begin{array}{c} -0.001 \\ (-0.002 -0.0004) \end{array} $	$ \begin{array}{c} -0.049 \\ (-0.083 -0.015) \end{array} $
Age 70+	$ \begin{array}{c} -0.001 \\ (-0.001 0.0001) \end{array} $	$-0.015 \\ (-0.056 0.026)$
Post-Killing X Conc. Dis.	$0.016 \\ (0.013 0.019)$	$ \begin{array}{c} -0.222 \\ (-0.374 -0.069) \end{array} $
T Post-Killing X Conc. Dis.	$ \begin{array}{c} -0.00003 \\ (-0.0001 0.00001) \end{array} $	$ \begin{array}{c} -0.001 \\ (-0.003 0.0003) \end{array} $
Constant	$0.050 \\ (0.018 0.083)$	$ \begin{array}{c} -4.495 \\ (-6.295 -2.694) \end{array} $
SD(Tract) SD(Residual)	0.013 0.094	0.674
Observations Log Likelihood	38,357 35,705.860	38,357 $-5,579.381$

Note:

95% Confidence Intervals in parentheses

```
cj_exp_prepost_panel <- pdata.frame(cj_exp_prepost, index = c("GEOID", "WEEKID"), drop.index = F)</pre>
slmtest(car_jack_rate~t+post_floyd+t_post_floyd+
             conc dis+
             age_19_29_perc+age_30_49_perc+age_50_69_perc+
             age_70_plus_perc+ post_floyd:conc_dis+t_post_floyd:conc_dis,
       data=cj_exp_prepost_panel,
        listw =w,
        model = "random",
        test = "lml")
##
## LM test for spatial lag dependence
##
## data: formula (random transformation)
## LM = 1.0311, df = 1, p-value = 0.3099
## alternative hypothesis: spatial lag dependence
#SAR Linear CJ
sar <- spml(car_jack_rate~t+post_floyd+t_post_floyd+</pre>
             conc_dis+
             age_19_29_perc+age_30_49_perc+age_50_69_perc+
             age_70_plus_perc+ post_floyd:conc_dis+t_post_floyd:conc_dis,
                     data = cj_exp_prepost,
             index = c("GEOID"),
            effect="individual",
            model="random",
            listw = w,
            lag=T,
            spatial.error="none")
summary(sar)
## Warning in sqrt(diag(object$vcov.arcoef)): NaNs produced
## ML panel with spatial lag, random effects
##
## Call:
## spreml(formula = formula, data = data, index = index, w = listw2mat(listw),
       w2 = listw2mat(listw2), lag = lag, errors = errors, cl = cl)
##
## Residuals:
##
       Min. 1st Qu.
                     Median
                                  Mean 3rd Qu.
                                                     Max.
## -0.12788 -0.02415 -0.01091 0.00015 -0.00335 3.04165
##
## Error variance parameters:
##
       Estimate Std. Error t-value Pr(>|t|)
## phi 0.0178066 0.0028774 6.1884 6.079e-10 ***
##
## Spatial autoregressive coefficient:
           Estimate Std. Error t-value Pr(>|t|)
##
## lambda 0.0084483
                           NaN
                                   {\tt NaN}
##
## Coefficients:
                            Estimate Std. Error t-value Pr(>|t|)
##
                          5.0394e-02 1.6084e-02 3.1332 0.0017290 **
## (Intercept)
```

```
1.4758e-05 1.2374e-05 1.1926 0.2330308
## t
## post floyd
                         2.6618e-02 1.9536e-03 13.6247 < 2.2e-16 ***
                        -1.7571e-05 2.2370e-05 -0.7855 0.4321814
## t post floyd
## conc_dis
                         4.7086e-04 1.9828e-03 0.2375 0.8122918
## age_19_29_perc
                        -4.6335e-04 1.6484e-04 -2.8110 0.0049392 **
                        -3.5226e-04 2.6419e-04 -1.3334 0.1824035
## age 30 49 perc
                        -1.0343e-03 3.0353e-04 -3.4076 0.0006554 ***
## age 50 69 perc
                        -6.1286e-04 3.6786e-04 -1.6660 0.0957136 .
## age_70_plus_perc
## post_floyd:conc_dis
                         1.5984e-02 1.6296e-03 9.8089 < 2.2e-16 ***
## t_post_floyd:conc_dis -2.7238e-05 1.8971e-05 -1.4358 0.1510713
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
#SAR LPM CJ
sar_lpm <- spml(anyjack~t+post_floyd+t_post_floyd+</pre>
             conc_dis+
             age_19_29_perc+age_30_49_perc+age_50_69_perc+
            age_70_plus_perc+ post_floyd:conc_dis+t_post_floyd:conc_dis,
                     data = cj_exp_prepost,
            index = c("GEOID"),
            effect="individual",
            model="random",
            listw = w,
            lag=T,
            spatial.error="none")
summary(sar_lpm)
## Warning in sqrt(diag(object$vcov.arcoef)): NaNs produced
## ML panel with spatial lag, random effects
##
## Call:
## spreml(formula = formula, data = data, index = index, w = listw2mat(listw),
       w2 = listw2mat(listw2), lag = lag, errors = errors, cl = cl)
##
##
## Residuals:
      Min. 1st Qu. Median
                                 Mean 3rd Qu.
                                                   Max.
## -0.22521 -0.06171 -0.02737 -0.00018 -0.00872 1.01940
## Error variance parameters:
       Estimate Std. Error t-value Pr(>|t|)
## phi 0.0209482 0.0033372 6.2773 3.446e-10 ***
##
## Spatial autoregressive coefficient:
           Estimate Std. Error t-value Pr(>|t|)
## lambda -0.0045863
                           {\tt NaN}
                                   NaN
                                            NaN
## Coefficients:
##
                           Estimate Std. Error t-value Pr(>|t|)
## (Intercept)
                         7.7400e-02 3.4805e-02 2.2238 0.026160 *
## t
                         5.0628e-05 2.4981e-05 2.0267 0.042692 *
## post floyd
                         5.6788e-02 3.9439e-03 14.3989 < 2.2e-16 ***
                        -3.0913e-05 4.5160e-05 -0.6845 0.493650
## t_post_floyd
## conc dis
                         2.5904e-03 4.2791e-03 0.6054 0.544943
## age_19_29_perc
                        -5.9952e-04 3.5683e-04 -1.6801 0.092929 .
```