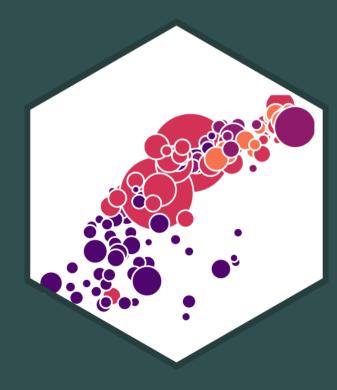
#### 2.7 — Inference for Regression

ECON 480 • Econometrics • Fall 2021

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#### **Outline**



**Hypothesis Testing** 

<u>Digression: p-Values and the Philosophy of Science</u>

<u>Hypothesis Testing by Simulation, with infer</u>

What R Calculates (Classical Statistical Inference)

The Use and Abuse of *p*-values



## **Hypothesis Testing**

#### **Estimation and Hypothesis Testing I**



- We want to **test** if our estimates are **statistically** significant and they describe the population
  - this is the "bread and butter" of using inferential statistics

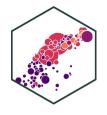
#### **Examples**:

- Does reducing class size actually improve test scores?
- Do more years of education increase your wages?
- Is the gender wage gap between men and women 23%?



 All modern science is built upon statistical hypothesis testing, so understand it well!

#### **Estimation and Hypothesis Testing II**



- Note, we can test a lot of hypotheses about a lot of population parameters, e.g.
  - $\circ$  A population mean  $\mu$ 
    - **Example**: average height of adults
  - A population proportion p
    - **Example:** percent of voters who voted for Trump
  - $\circ$  A difference in population means  $\mu_A \mu_B$ 
    - **Example**: difference in average wages of men vs. women
  - $\circ$  A difference in population proportions  $p_A p_B$ 
    - **Example**: difference in percent of patients reporting symptoms of drug A vs B
- We will focus on hypotheses about population regression slope  $(\hat{\beta}_1)$ , i.e. the causal effect of X on Y

<sup>&</sup>lt;sup>†</sup> With a model this simple, it's almost certainly **not** causal, but this is the ultimate direction we are heading...

#### **Null and Alternative Hypotheses I**



- All scientific inquiries begin with a **null hypothesis**  $(H_0)$  that proposes a specific value of a population parameter
  - $\circ$  Notation: add a subscript 0:  $\beta_{1,0}$  (or  $\mu_0$ ,  $p_0$ , etc)
- We suggest an alternative hypothesis  $(H_a)$ , often the one we hope to verify
  - $\circ$  Note, can be multiple alternative hypotheses:  $H_1, H_2, \ldots, H_n$
- Ask: "Does our data (sample) give us sufficient evidence to reject  $H_0$  in favor of  $H_a$ ?"
  - $\circ$  Note: the test is always about  $H_0!$
  - See if we have sufficient evidence to reject the status quo

#### **Null and Alternative Hypotheses II**



• Null hypothesis assigns a value (or a range) to a population parameter

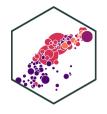
$$\circ$$
 e.g.  $\beta_1 = 2$  or  $\beta_1 \le 20$ 

- $\circ$  Most common is  $\beta_1 = 0 \implies X$  has no effect on Y (no slope for a line)
- Note: always an equality!
- Alternative hypothesis must mathematically *contradict* the null hypothesis

$$\circ$$
 e.g.  $\beta_1 \neq 2$  or  $\beta_1 > 20$  or  $\beta_1 \neq 0$ 

- Note: always an inequality!
- Alternative hypotheses come in two forms:
  - 1. One-sided alternative:  $\beta_1 > H_0$  or  $\beta_1 < H_0$
  - 2. Two-sided alternative:  $\beta_1 \neq H_0$ 
    - Note this means either  $\beta_1 < H_0$  or  $\beta_1 > H_0$

#### **Components of a Valid Hypothesis Test**



- All statistical hypothesis tests have the following components:
- 1. A null hypothesis,  $H_0$
- 2. An alternative hypothesis,  $H_a$
- 3. A **test statistic** to determine if we reject  $H_0$  when the statistic reaches a "critical value"
  - $\circ$  Beyond the critical value is the "rejection region", sufficient evidence to reject  $H_0$
- 4. A **conclusion** whether or not to reject  $H_0$  in favor of  $H_a$

#### Type I and Type II Errors I



- Sample statistic  $(\hat{\beta}_1)$  will rarely be exactly equal to the hypothesized parameter  $(\beta_1)$
- Difference between observed statistic and true parameter could be because:
- Parameter is *not* the hypothesized value
  - $\circ$   $H_0$  is false
- Parameter is truly hypothesized value but sampling variability gave us a different estimate
  - $\circ$   $H_0$  is *true*
- We cannot distinguish between these two possibilities with any certainty



#### Type I and Type II Errors II

- We can interpret our estimates probabilistically as commiting one of two types of error:
- 1. Type I error (false positive): rejecting  $H_0$  when it is in fact true
  - Believing we found an important result when there is truly no relationship
- 2. Type II error (false negative): failing to reject  $H_0$  when it is in fact false
  - Believing we found nothing when there was truly a relationship to find



#### Type I and Type II Errors III



		Null is True	<b>Null is False</b>
Judgment	Reject Null	<b>Type I Error</b> (False Positive)	CORRECT (True Positive)
	Don't Reject Null	CORRECT (True Negative)	<b>Type II Error</b> (False Negative)

**Truth** 

• Depending on context, committing one type of error may be more serious than the other

#### Type I and Type II Errors IV



# Defendant is Innocent Defendant is Guilty Convict Type I Error (False Positive) Acquit CORRECT (True Positive) CORRECT (True Negative) Type II Error (False Negative)

- ullet Anglo-American common law *presumes* defendant is innocent:  $H_0$
- Jury judges whether the evidence presented against the defendant is plausible assuming the defendant were in fact innocent
- ullet If highly improbable (beyond a "reasonable doubt"): sufficient evidence to reject  $H_0$  and convict

#### Type I and Type II Errors V





"It is better that ten guilty persons escape than that one innocent suffer."

• Type I error is worse than a Type II error in law!

Blackstone, William, 1765-1770, Commentaries on the Laws of England

William Blackstone

(1723-1780)

#### **Type I and Type II Errors VI**



#### Type I and Type II Errors VI





Type I error: You think you're on mute but you're not

Type II error: You think they can hear you but you're on mute

Or did I get it backwards?

11:47 PM · 2020-09-11 · Twitter for iPhone

#### Significance Level, $\alpha$ , and Confidence Level $1-\alpha$



• The **significance level**,  $\alpha$ , is the probability of a **Type I error** 

$$\alpha = P(\text{Reject } H_0 | H_0 \text{ is true})$$

- The **confidence level** is defined as  $(1 \alpha)$ 
  - $\circ$  Specify *in advance* an  $\alpha$ -level (0.10, 0.05, 0.01) with associated confidence level (90%, 95%, 99%)
- The probability of a **Type II error** is defined as  $\beta$ :

$$\beta = P(\text{Don't reject } H_0 | H_0 \text{ is false})$$





#### **Truth**

**Null is True** 

Judgment

**Reject Null** 

**Don't Reject Null** 

Type I Error (α)	CORRECT $(1 - \beta)$

**Null is False** 

CORRECT  $(1 - \alpha)$  Type II Error  $(\beta)$ 

#### **Power and p-values**



• The statistical power of the test is  $(1 - \beta)$ : the probability of correctly rejecting  $H_0$  when  $H_0$  is in fact false (e.g. not convicting an innocent person)

Power = 
$$1 - \beta = P(\text{Reject } H_0 | H_0 \text{ is false})$$

• The *p*-value or significance probability is the probability that, if the null hypothesis were true, the test statistic from any sample will be *at least as extreme* as the test statistic from *our* sample

$$p(\delta \ge \delta_i | H_0 \text{ is true})$$

- ullet where  $\delta$  represents some test statistic
- $\delta_i$  is the test statistic we observe in our sample
- More on this in a bit

#### p-Values and Statistical Significance



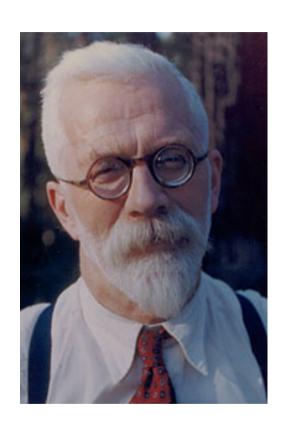
- After running our test, we need to make a *decision* between the competing hypotheses
- Compare p-value with pre-determined  $\alpha$  (commonly,  $\alpha=0.05$ , 95% confidence level)
- If  $p < \alpha$ : statistically significant evidence sufficient to reject  $H_0$  in favor of  $H_a$ 
  - $\circ$  Note this does **not** mean  $H_a$  is true! We merely have *rejected*  $H_0$ !
- If  $p \ge \alpha$ : insufficient evidence to reject  $H_0$ 
  - $\circ$  Note this does **not** mean  $H_0$  is true! We merely have failed to reject  $H_0$ !



# Digression: p-Values and the Philosophy of Science

#### Hypothesis Testing and the Philosophy of Science I





"The null hypothesis is never proved or established, but is possibly disproved, in the course of experimentation. Every experiment may be said to exist only in order to give the facts a chance of disproving the null hypothesis."

1931, The Design of Experiments

Sir Ronald A. Fisher

(1890 - 1962)

#### Hypothesis Testing and the Philosophy of Science I



- Modern philosophy of science is largely based off of hypothesis testing and falsifiability, which form the "Scientific Method"<sup>†</sup>
- For something to be "scientific", it must be falsifiable, or at least testable
- Hypotheses can be corroborated with evidence, but always tentative until falsified by data in suggesting an alternative hypothesis

"All swans are white" is a hypothesis rejected upon discovery of a single black swan



<sup>&</sup>lt;sup>†</sup> Note: economics is a very different kind of "science" with a different methodology!

#### **Hypothesis Testing and p-Values**



 Hypothesis testing, confidence intervals, and p-values are probably the hardest thing to understand in statistics



Fivethirtyeight: Not Even Scientists Can Easily Explain P-values

#### **Hypothesis Testing: Which Test? I**



- Rigorous course on statistics (<u>ECMG 212</u> or **MATH 112**) will spend weeks going through different types of tests:
  - Sample mean; difference of means
  - Proportion; difference of proportions
  - Z-test vs t-test
  - ∘ 1 sample vs. 2 samples
  - $\circ \chi^2$  test

#### **Hypothesis Testing: Which Test? II**



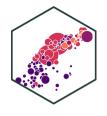
#### **There is Only One Test**

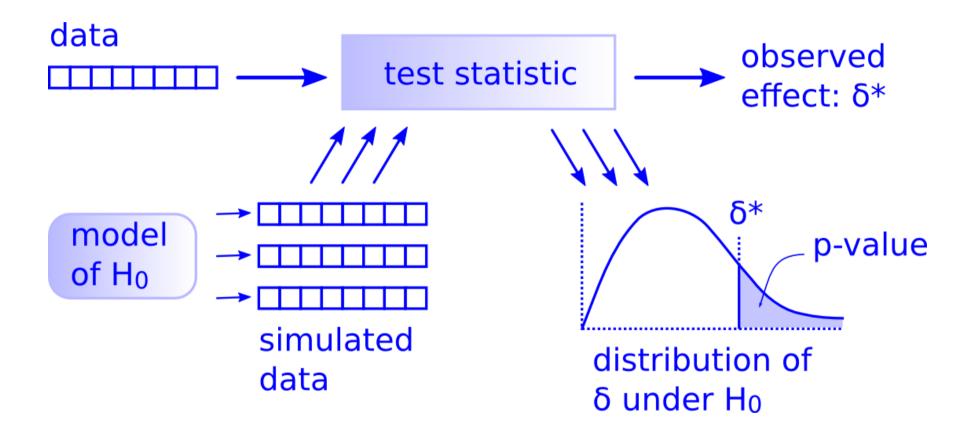


- Fortunately, some clever statisticians realized <u>"there is only one test"</u> and some built a nice R package called <u>infer</u>
- 1. **Calculate** a statistic,  $\delta_i^{\dagger}$ , from a sample of data
- 2. **Simulate** a world where  $\delta$  is null  $(H_0)$
- 3. **Examine** the distribution of  $\delta$  across the null world
- 4. **Calculate** the probability that  $\delta_i$  could exist in the null world
- 5. **Decide** if  $\delta_i$  is statistically significant

 $<sup>^{\</sup>dagger}\delta$  can stand in for any test-statistic in any hypothesis test! For our purposes,  $\delta$  is the slope of our regression sample,  $\hat{\beta}_1$ .

#### **Elements of a Hypothesis Test**





Alan Downey: "There is still only one test"



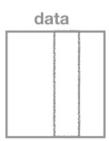
• R naturally runs the following hypothesis test on any regression as part of lm():

$$H_0: \beta_1 = 0$$

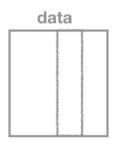
$$H_1: \beta_1 \neq 0$$

- infer allows you to run through these steps manually to understand the process:
- 1. specify() a model
- 2. hypothesize() the null
- 3. generate() simulations of the null world
- 4. calculate() the *p*-value
- 5. visualize() with a histogram (optional)



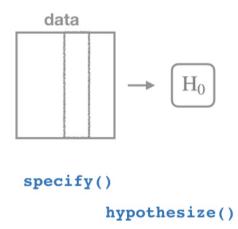




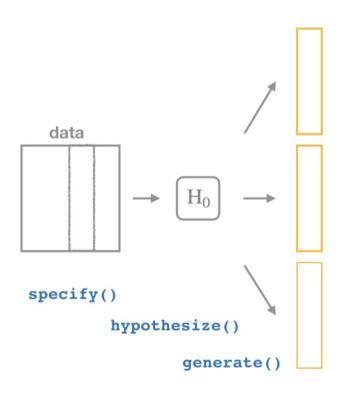


specify()

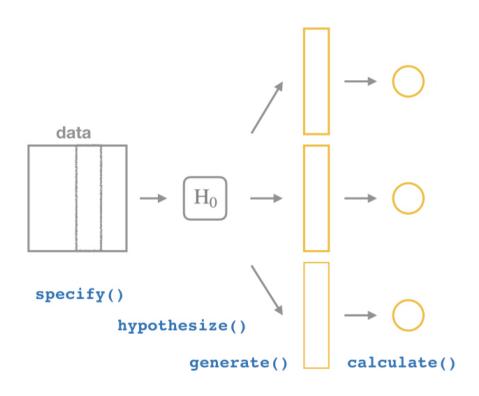




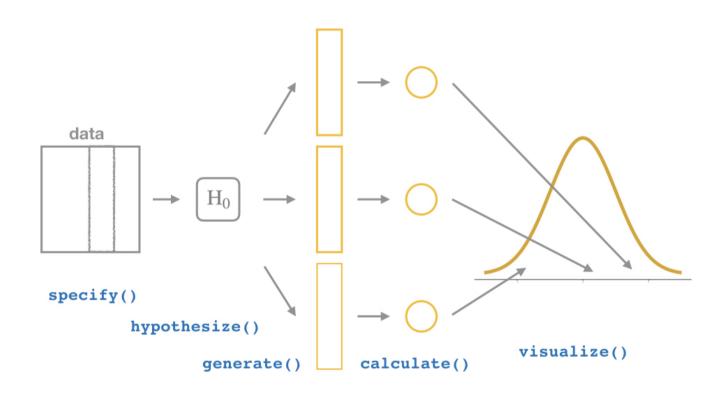












#### **Classical Inference: Critical Values of Test Statistic**



- Test statistic  $(\delta)$ : measures how far what we observed in our sample  $(\hat{\beta_1})$  is from what we would expect if the null hypothesis were true  $(\beta_1=0)$ 
  - $\circ$  Calculated from a sampling distribution of the estimator (i.e.  $\hat{\beta}_1$ )
  - $\circ$  In econometrics, we use t-distributions which have n-k-1 degrees of freedom<sup>†</sup>
- Rejection region: if the test statistic reaches a "critical value" of  $\delta$ , then we reject the null hypothesis

<sup>&</sup>lt;sup>†</sup> Again, see last class's <u>appendix</u> for more on the t-distribution. k is the number of independent variables our model has, in this case, with just one X, k=1. We use two degrees of freedom to calculate  $\hat{\beta}_0$  and  $\hat{\beta}_1$ , hence we have n-2 df.



### Hypothesis Testing by Simulation, with infer

# Imagine a Null World, where $H_0$ is True





Our world, and a world where  $\beta_1 = 0$  by assumption.

# **Comparing the Worlds I**



ullet From that null world where  $H_0: eta_1=0$  is true, we **simulate** another sample and calculate OLS estimators again

#### **Our Sample**

term	estimate	std.error	
<chr></chr>	<dpl></dpl>	<dbl></dbl>	
(Intercept)	698.932952	9.4674914	
str	-2.279808	0.4798256	
2 rows   1-3 of 5 columns			

#### **Another Sample**

term	estimate	std.error
<chr></chr>	<dbl></dbl>	<dpl></dpl>
(Intercept)	647.8027952	9.7147718
str	0.3235038	0.4923581
2 rows   1-3 of 5 columns		

# **Comparing the Worlds II**

• From that null world where  $H_0: \beta_1=0$  is true, let's **simulate 1,000** samples and calculate slope  $(\hat{\beta}_1)$  for each

sample	slope
<int></int>	<pre><dpl></dpl></pre>
1	-0.3027333296
2	-0.3624481355
3	0.6448518690
4	-0.0745971847
5	0.5969444290
6	0.5505335318
7	0.5927466147
8	0.0572148658
9	-0.0989989073
10	0.8043957511
1-10 of 1,000 rows	Previous <b>1</b> 2 3 4 5 6 100 Next

# Prepping the infer Pipeline



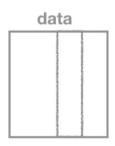
- Before I show you how to do this, let's first save our estimated slope from our *actual* sample
  - We'll want this later!

```
# save as obs_slope
sample_slope <- school_reg_tidy %>% # this is the regression tidied with broom's tidy()
filter(term=="str") %>%
pull(estimate)
# confirm what it is
sample_slope
```

```
## [1] -2.279808
```

# The *infer* Pipeline: Specify





specify()

# The infer Pipeline: Specify



#### Specify

```
data %>%
  specify(y ~ x)
```

 Take our data and pipe it into the specify() function, which is essentially a lm() function for regression (for our purposes)

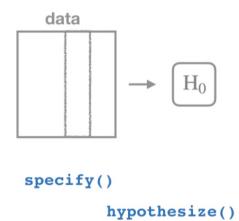
```
CASchool %>%
  specify(testscr ~ str)
```

	testscr	str
	<pre><dpl></dpl></pre>	<qpf></qpf>
	690.8	17.88991
	661.2	21.52466
	643.6	18.69723
3 rows		

Note nothing happens yet

# The infer Pipeline: Hypothesize





# The infer Pipeline: Hypothesize



#### Specify

#### Hypothesize

```
%>% hypothesize(null =
"independence")
```

- Describe what the null hypothesis is here
- In infer's language, we are hypothesizing that str and testscr are independent  $(\beta_1 = 0)^{\dagger}$

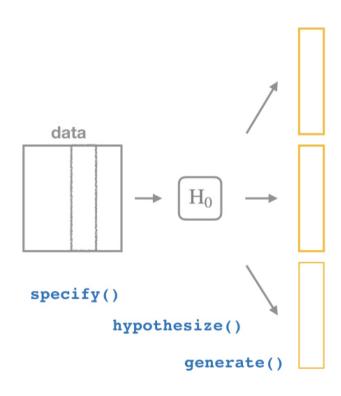
```
CASchool %>%
  specify(testscr ~ str) %>%
  hypothesize(null = "independence")
```

testse	r str
<db< th=""><th><dpl>&gt;</dpl></th></db<>	<dpl>&gt;</dpl>
690	8 17.88991
661	2 21.52466
643	6 18.69723
3 rows	

<sup>&</sup>lt;sup>†</sup> See more <u>here</u> about what other hypotheses you can test with <u>infer</u>.

# The *infer* Pipeline: Generate I





## The *infer* Pipeline: Generate I



Specify

Hypothesize

#### Generate

```
%>% generate(reps = n, type =
"permute")
```

- Now the magic starts, as we run a number of simulated samples
- Set the number of reps and set the type equal to "permute"
  - we want permutation (not bootstrap!) because we are simulating a world where  $\beta_1 = 0$  by construction!

# The infer Pipeline: Generate II



Specify

Hypothesize

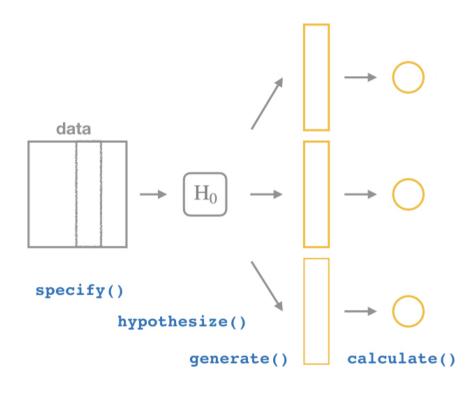
#### Generate

```
%>% generate(reps = n, type =
"permute")
```

testscr	str	str				replicate
<dpl></dpl>	<dbl></dbl>					<int></int>
693.95	17.88991					1
642.40	21.52466					1
680.45	18.69723					1
672.70	17.35714					1
666.45	18.67133		1			
654.20	21.40625					1
671.95	19.50000					1
671.75	20.89412					1
624.55	19.94737	19.94737				1
699.10	20.80556					1
1-10 of 10,000 rows	Previous <b>1</b>	2	3	4	5	6 1000Next

# The *infer* Pipeline: Calculate I





## The *infer* Pipeline: Calculate I



Specify

Hypothesize

Generate

Calculate

```
%>% calculate(stat = "")
```

- We calculate sample statistics for each of the 1,000 replicate samples
- In our case, calculate the slope,  $(\hat{\beta}_1)$  for each replicate

• Other stats for calculation: "mean", "median", "prop", "diff in means", "diff in props", etc. (see <a href="mailto:package information">package information</a>)

# The *infer* Pipeline: Calculate II



Specify

Hypothesize

Generate

Calculate

%>% calculate(stat = "")

replicate	stat
<int></int>	<dpl></dpl>
1	0.384783281
2	0.241700895
3	0.268799843
4	-0.189039951
5	1.215030315
6	0.511783627
7	-0.457378304
8	1.008206723
9	0.092043084
10	0.233837801
1-10 of 1,000 rows Previou	s <b>1</b> 2 3 4 5 6 100Next

## The infer Pipeline: Get p Value



Specify

Hypothesize

Generate

Calculate

#### Get p Value

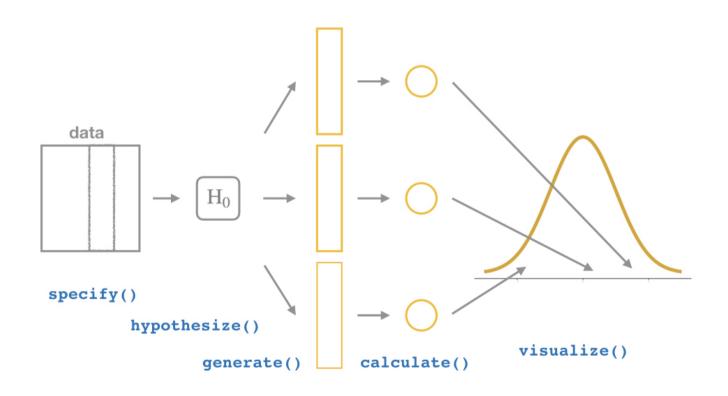
```
%>% get_p_value(obs stat = "",
direction = "both")
```

- We can calculate the *p*-value
  - the probability of seeing a value at least as large as our
     sample\_slope (-2.28) in our simulated null distribution
- Two-sided alternative  $H_a: \beta_1 \neq 0$ , we double the raw p-value

```
p_value <dbl>
0
1 row
```

# The infer Pipeline: Visualize I





## The infer Pipeline: Visualize I



Specify

Hypothesize

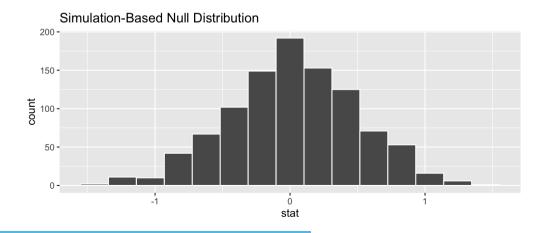
Generate

Calculate

Visualize

%>% visualize()

- Make a histogram of our null distribution of  $eta_1$ 
  - $\circ$  Note it is centered at  $\beta_1=0$  because that's  $H_0!$



## The infer Pipeline: Visualize II



Specify

Hypothesize

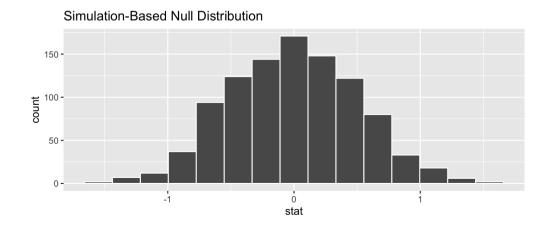
Generate

Calculate

Visualize

%>% visualize()

• Add our sample\_slope to show our finding on the null
distr.



## The *infer* Pipeline: Visualize p-value



Specify

Hypothesize

Generate

Calculate

Visualize

%>%
visualize()+shade\_p\_value()

Add shade\_p\_value() to see what p is

## The infer Pipeline: Visualize Confidence Intervals



Specify

Hypothesize

Generate

Calculate

Visualize

```
%>%
visualize()+shade_ci()
```

- To shade confidence interval, we first need a vector of what they are
  - I've saved the outputted tibble of them from 4 slides ago as ci\_values

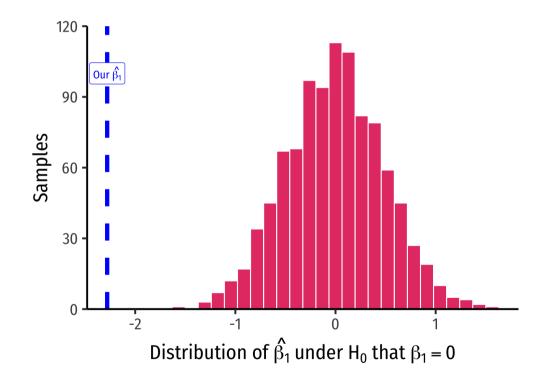
```
simulations %>%
visualize(obs_stat = sample_slope)+
shade_confidence_interval(ci_values)
```

# The infer Pipeline: Visualize is a Wrapper of ggplot



- infer's visualize() function is just a wrapper function for ggplot()
  - you can take your simulations tibble and just ggplot a normal histogram

```
simulations %>%
  ggplot(data = .)+
  aes(x = stat)+
  geom histogram(color="white", fill="#e64173")+
  geom vline(xintercept = sample slope,
             color = "blue",
             size = 2.
            linetype = "dashed")+
  annotate(geom = "label",
           x = -2.28
           y = 100,
           label = expression(paste("Our ", hat(beta[1]))),
           color = "blue")+
  scale y continuous(lim=c(0,120),
                     expand = c(0,0)+
  labs(x = expression(paste("Distribution of ", hat(beta[1]),
       v = "Samples")+
    theme_classic(base_family = "Fira Sans Condensed",
           base size=20)
```





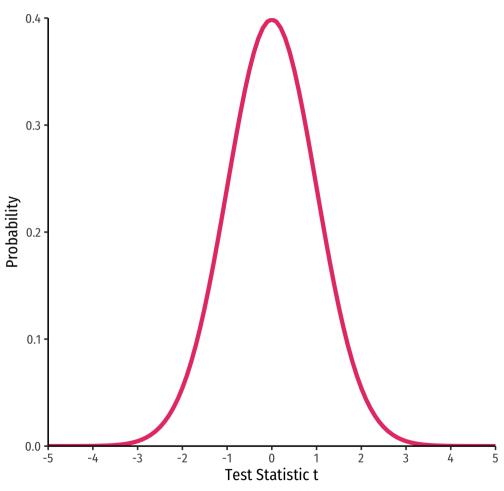
# What R Calculates (Classical Statistical Inference)

#### **What R Does: Classical Statistical Inference I**



- R does things the old-fashioned way, using a theoretical null distribution instead of simulation
- A *t*-distribution with n-k-1 df<sup>†</sup>
- Calculate a t-statistic for  $\hat{\beta}_1$ :

test statistic = 
$$\frac{\text{estimate} - \text{null hypothesis}}{\text{standard error of estimate}}$$



t measures our sample statistic ( $t_i$ )'s deviation from  $H_0$  if  $H_0$  were True

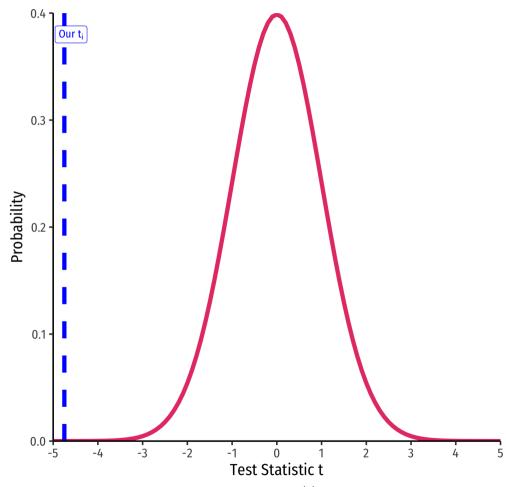
 $<sup>^{\</sup>dagger}k$  is the number of X variables.

#### **What R Does: Classical Statistical Inference II**



test statistic = 
$$\frac{\text{estimate} - \text{null hypothesis}}{\text{standard error of estimate}}$$

- t has the same interpretation as Z, number of std. dev. away from the distribution's center<sup>†</sup>
- Compares to a critical value of  $t^*$  (determined by  $\alpha \& n-k-1$ )
  - $\circ$  For 95% confidence,  $\alpha=0.05$ ,  $t^*pprox 2^{\ddagger}$



t measures our sample statistic (t<sub>i</sub>)'s deviation from H<sub>0</sub> if H<sub>0</sub> were True

<sup>&</sup>lt;sup>†</sup> Think of our simulated distribution, the center was 0.

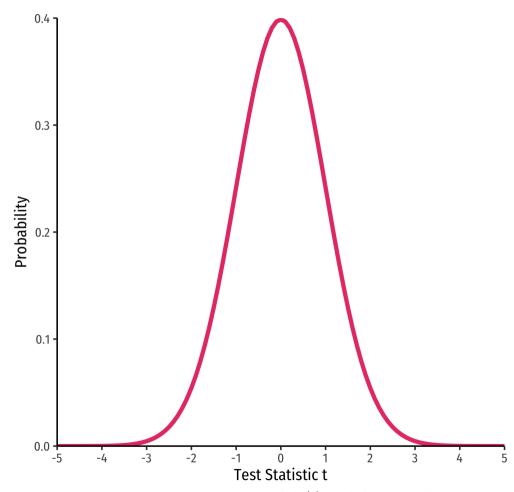
<sup>&</sup>lt;sup>‡</sup> The 68-95-99.7% empirical rule!

#### **What R Does: Classical Statistical Inference III**



$$t = \frac{\hat{\beta}_1 - \beta_{1,0}}{se(\hat{\beta}_1)}$$
$$t = \frac{-2.28 - 0}{0.48}$$
$$t = -4.75$$

- ullet Our sample slope is 4.75 standard deviations below the mean under  $H_0$
- p-value: prob. of a test statistic at least as large (in magnitude) as ours if the null hypothesis were true<sup>†</sup>
  - $\circ p$ -value is **2-sided** for  $H_a: \beta_1 \neq 0$



t measures our sample statistic (t<sub>i</sub>)'s deviation from H<sub>0</sub> if H<sub>0</sub> were True

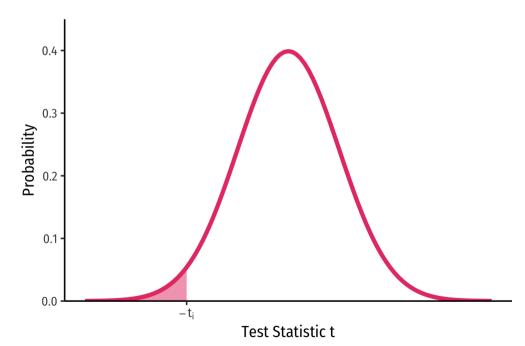
<sup>&</sup>lt;sup>†</sup> Think of our simulated distribution, the center was 0.

## 1-Sided vs. 2-Sided p-values I



 $H_a: \beta_1 < 0$ 

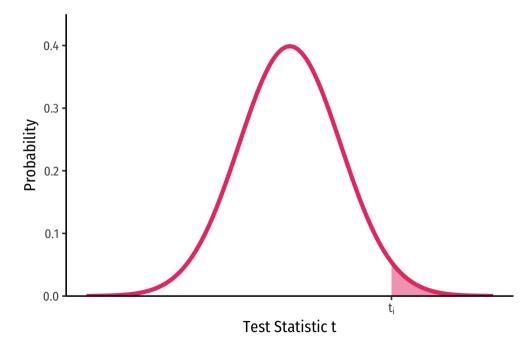
p-value:  $Prob(t < t_i)$ 



t measures our sample statistic (t<sub>i</sub>)'s deviation from H<sub>0</sub> if H<sub>0</sub> were True

 $H_a: \beta_1 > 0$ 

p-value:  $Prob(t > t_i)$ 



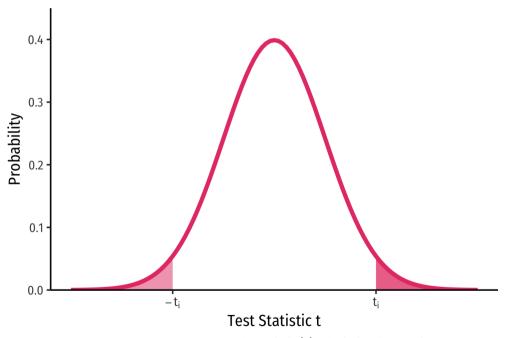
t measures our sample statistic (t<sub>i</sub>)'s deviation from H<sub>0</sub> if H<sub>0</sub> were True

# 1-Sided vs. 2-Sided p-values I



$$H_a:\beta_1\neq 0$$

p-value:  $2 \times Prob(t > |t_i|)$ 



t measures our sample statistic ( $t_i$ )'s deviation from  $H_0$  if  $H_0$  were True

## **Hypothesis Tests in Regression Output I**



```
summary(school_reg)
```

```
##
## Call:
## lm(formula = testscr ~ str. data = CASchool)
##
## Residuals:
      Min
##
          1Q Median 3Q Max
## -47.727 -14.251 0.483 12.822 48.540
##
## Coefficients:
             Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 698.9330 9.4675 73.825 < 2e-16 ***
       -2.2798 0.4798 -4.751 2.78e-06 ***
## str
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 18.58 on 418 degrees of freedom
## Multiple R-squared: 0.05124, Adjusted R-squared: 0.04897
## F-statistic: 22.58 on 1 and 418 DF, p-value: 2.783e-06
```

# **Hypothesis Tests in Regression Output II**



• In broom's tidy() (with confidence intervals)

```
tidy(school_reg, conf.int=TRUE)
```

term	estimate	std.error	statistic	p.value	conf.low	conf.high
<chr></chr>	<dpl></dpl>	<dpl></dpl>	<dbl></dbl>	<dpl>&lt;</dpl>	<dbl></dbl>	<dpl>&lt;</dpl>
(Intercept)	698.932952	9.4674914	73.824514	6.569925e-242	680.32313	717.542779
str	-2.279808	0.4798256	-4.751327	2.783307e-06	-3.22298	-1.336637

2 rows

### **Conclusions**

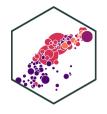


$$H_0: \beta_1 = 0$$

$$H_a:\beta_a\neq 0$$

- Because the hypothesis test's p-value  $< \alpha$  (0.05)...
- We have sufficient evidence to reject  $H_0$  in favor of our alternative hypothesis. Our sample suggests that there is a relationship between class size and test scores.
- Using the confidence intervals:
- We are 95% confident that the true marginal effect of class size on test scores is between -3.22 and -1.34.

## Hypothesis Testing vs. Confidence Intervals



Confidence intervals are all two-sided by nature

$$CI_{0.95} = \left( \left[ \hat{\beta}_1 - 2 \times se(\hat{\beta}_1) \right], \left[ \hat{\beta}_1 + 2 \times se(\hat{\beta}_1) \right] \right)$$

• Hypothesis test (t-test) of  $H_0: \beta_1 = 0$  computes a t-value of

$$t = \frac{\hat{\beta_1}}{se(\hat{\beta_1})}$$

and p<0.05 when  $t \ge 2$ 

• If a confidence interval contains the  $H_0$  value (i.e. 0, for our test), then we fail to reject  $H_0$ .

<sup>&</sup>lt;sup>1</sup> Since our null hypothesis is that  $\beta_{1,0}=0$ , the test statistic simplifies to this neat fraction.



# The Use and Abuse of p-values

# **Common Misconceptions about p-values**



So how do we interpret p again?

#### $\times p$ is the probability that the alternative hypothesis is false

• We can never *prove* an alternative hypothesis, only tentatively reject a null hypothesis

#### imes p is the probability that the null hypothesis is true

• We're not proving the  $H_0$  is false, only saying that it's very unlikely that if  $H_0$  were true, we'd obtain a slope as rare as our sample's slope

#### $\times p$ is the probability that our observed effects were produced purely by random chance

• p is computed under a specific model (think about our null world) that assumes  $H_0$  is true

#### $\times p$ tells us how significant our finding is

- p tells us nothing about the *size* or the *real world significance* of any effect deemed "statistically significant"
- it only tells us that the slope is statistically significantly different from 0 (if  $H_0$  is  $\beta_1 = 0$ )

# **Abusing p-Values I**



HOW SCIENCE REPORTING WORKS:





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# **Abusing p-Values II**





"The widespread use of 'statistical significance' (generally interpreted as  $(p \le 0.05)$  as a license for making a claim of a scientific finding (or implied truth) leads to considerable distortion of the scientific process."

Wasserstein, Ronald L. and Nicole A. Lazar, (2016), "The ASA's Statement on p-Values: Context, Process, and Purpose," The American Statistician 30(2): 129-133

# **Abusing p-Values II**



"No economist has achieved scientific success as a result of a statistically significant coefficient. Massed observations, clever common sense, elegant theorems, new policies, sagacious economic reasoning, historical perspective, relevant accounting, these have all led to scientific success. Statistical significance has not."

McCloskey, Dierdre N and Stephen Ziliak, 1996, The Cult of Statistical Significance, p. 112)

## p-value Clarification

- Again, p-value is the probability that, if the null hypothesis were true, we obtain (by pure random chance) a test statistic at least as extreme as the one we estimated for our sample
- A low p-value means either (and we can't distinguish which):
  - 1.  $H_0$  is true and a highly improbable event has occurred OR
  - 2.  $H_0$  is false

# **Significance In Regression Tables**



	Test Score	
Intercept	698.93 ***	
	(9.47)	
STR	-2.28 ***	
	(0.48)	
N	420	
R-Squared	0.05	
SER	18.58	
*** n < 0.001. ** n < 0.01. * n < 0.05		

<sup>\*\*\*</sup> p < 0.001; \*\* p < 0.01; \* p < 0.05.

 Statistical significance is shown by asterisks, common (but not always!) standard:

- $\circ$  1 asterisk: significant at  $\alpha=0.10$
- $\circ$  2 asterisks: significant at  $\alpha=0.05$
- $\circ$  3 asterisks: significant at  $\alpha=0.01$
- ullet Rare, but sometimes regression tables include p-values for estimates