Randomized Experiments and Covariates

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Motivating Example

Blocking: Randomization Improved

Applications: Seguro Popular, Perry Preschool, Ignatieff

Campaign

Two Issues: Outliers and Interference

Blocking for Sequential Experiments

Motivating Example

Random Assignment Mechanisms

- Simple/complete randomization (Bernoulli trial, prob π)
- ➤ Complete randomization / random allocation (fixed proportion to tr)
- ▶ Blocked randomizations
 (fixed proportion to tr, w/in group)
- Cluster randomizations
 (assignment at higher level)

"Would a canvassing policy increase enrollment in a health insurance program?"

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Precinct	Party	Canvass?	Enroll %
1	Dem		
2	Dem		
3	Rep		
4	Rep		

Suppose we observationally measure

Precinct	Party	Canvass?	Enroll $\%$
1	Dem	Yes	60
2	Dem	Yes	70
3	Rep	No	20
4	Rep	No	30
		Diff in Means:	40
		(Yes - No)	

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Causal claims? Concerns?

Suppose we randomly assign 2 Tr, 2 Co, and measure

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Seriously? Well, ...

```
x <- sample(0:1, size = 10, replace = TRUE)
x
## [1] 1 1 1 1 1 1 1 1 1</pre>
```

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 - ▶ 10^{th} → treated: $SE: \sqrt{\frac{6}{4} + \frac{4}{6}} = \sqrt{2.66}$

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 - ▶ 10^{th} → treated: $SE: \sqrt{\frac{6}{4} + \frac{4}{6}} = \sqrt{2.66}$
- ▶ If $Var(Y_i(0)) \neq Var(Y_i(1))$, allocate → higher-Variance condition

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Unit	Y_0	$Y_1 (+ cov)$	Y_1 (- cov)
1	0	0	10
2	5	5	5
3	10	10	0
Means	5	5	5

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1	0	0	10
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$$\widehat{ATE}_{+\text{cov}} = -7.5, 0, 7.5
\widehat{ATE}_{-\text{cov}} = 2.5, 0, -2.5$$

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$$\widehat{ATE}_{+\text{cov}} = -7.5, 0, 7.5$$

$$\widehat{ATE}_{-\text{cov}} = 2.5, 0, -2.5 \text{ (less variance!)}$$

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What would have happened to "No" precincts if "Yes"?

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What do we really want to know?

Does can vassing actually *change* enrollment in precinct? (Or, just Party \rightarrow Enrollment?)

What would have happened to "No" precincts if "Yes"?

What would have happened under *other* conditions?

Suppose we can know both potential outcomes . . .

			Enroll $\%$	Enroll $\%$
Precinct	Party	Canvass?	if No Canvass	if Canvass
1	Dem	_	20	60
2	Dem		30	70
3	Rep		20	30
4	Rep	_	30	40
		Means:	25	50

Suppose we can know both potential outcomes . . .

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Precinct	Party	Canvass?	if No Canvass	if Canvass
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2	Dem		30	70
3	Rep	_	20	30
4	Rep	_	30	40
		Means:	25	50

$$ATE = 50 - 25 = 25$$

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Precinct	Party	Canvass?	if No Canvass	if Canvass
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		Means:	25	50

$$ATE = 50 - 25 = 25$$

(True or an estimate?)

Another way to think about same information:

			Enroll $\%$	Enroll $\%$	True Precinct
Precinct	Party	Canvass?	if No Canvass	if Canvass	Effect
1	Dem		20	60	40
2	Dem	_	30	70	40
3	Rep	_	20	30	10
4	Rep	_	30	40	10
		Means:	25	50	25

$$ATE = (40 + 40 + 10 + 10)/4 = 25$$

The Fundamental Problem of Causal Inference

We can't observe both "Canvassed" and "Not Canvassed" for a precinct.

We can't observe both potential outcomes (counterfactuals).

So, how can we get a good causal estimate?

Suppose we observe ...

			Enroll %	Enroll $\%$
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-		Means:	25	65

Estimated ATE = 65 - 25 = 40

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Estimated ATE =
$$65 - 25 = 40$$
 © (too big)

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3	Rep	Yes		30
4	Rep	No	30	
-		Means:	30	45

Estimated ATE = 45 - 30 = 15

Or, we could have observed ...

			Enroll $\%$	Enroll $\%$
Precinct	Party	Canvass?	if No Canvass	if Canvass
1	Dem	Yes		60
2	Dem	No	30	
3	Rep	Yes		30
4	Rep	No	30	
		Means:	30	45

Estimated ATE =
$$45 - 30 = 15$$
 \odot

⊚ (too small; closer)

Assignments	Est ATE
YYNN	40

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YYNN	40
NYNY	35
YNNY	25
NYYN	25
YNYN	15
NNYY	10

In our random allocation, possible data were

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YYNN	40
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- $E(\hat{\bar{\tau}}) = \frac{1}{6} \cdot 40 + \frac{1}{6} \cdot 35 + \frac{1}{3} \cdot 25 + \frac{1}{6} \cdot 15 + \frac{1}{6} \cdot 10 = 25 = \bar{\tau}$

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Two assignments (YYNN and NNYY) leave treatment perfectly confounded with party.

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Two assignments (YYNN and NNYY) leave treatment perfectly confounded with party.

We can never see all of Y_1 , Y_0 . But we can see all of X! Let's ensure X does not predict T.

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Creating pre-treatmnt groups that look same on *predictors*.

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(Then, randomize within groups.)

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(Then, randomize within groups.)

Blocking restricts possible data to

Assignments	Est TE
YYNN	40
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$\mathbf{Y}\mathbf{N}\mathbf{N}\mathbf{Y}$	25
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Under random allocation, 5 possible estimates: 40, 15, 25, 35, 10.

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- ▶ Blocked pairs randomization: $2^{\frac{n}{2}} = 2^{\frac{4}{2}} = 4$

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- ► Guidelines for limited/uncertain resources

Imai, King, and Stuart (2008):

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and observed difference in sample means, D:

$$D \equiv \frac{1}{n/2} \sum_{i \in \{I_i = 1, T_i = 1\}} Y_i - \frac{1}{n/2} \sum_{i \in \{I_i = 1, T_i = 0\}} Y_i$$

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then, the total estimation error is

$$\Delta \equiv \text{PATE} - D$$

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Further,

$$\Delta = \underbrace{\Delta_{S_X}}_{\text{sampling error from obs}} + \underbrace{\Delta_{S_U}}_{\text{sampling error from unobs}} + \underbrace{\Delta_{T_X}}_{\text{treatment imbalance from obs}} + \underbrace{\Delta_{T_U}}_{\text{treatment imbalance from unobs}}$$

Blocking controls Δ_{T_X}

The *estimation error* can be decomposed:

$$\Delta = \underbrace{\Delta_S}_{\text{sampling error}} + \underbrace{\Delta_T}_{\text{treatment imbalance}}$$

Further,

$$\Delta$$
 = Δ_{S_X} + Δ_{S_U} +
sampling error from obs sampling error from unobs
$$\Delta_{T_X}$$
 + Δ_{T_U}
treatment imbalance from obs treatment imbalance from unobs

Blocking controls Δ_{T_X} , and, to the degree correlated, Δ_{T_U} .

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- Exact on one or two discrete covariates (best predictor)
- More covariates: no exact comparable units (5 covars, 3 levels: $3^5 = 243$)
- ► More covariates often done informally

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- Like matching, select other attributes (caliper, etc.)

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- ► Create blocks of similar units
- ► Randomize within blocks

A Matrix of Distances

```
(First 5 rows of x100, variables b1 and b2)

## 1001 1002 1003 1004 1005

## 1001 0.00 2.68 2.51 2.46 1.36

## 1002 2.68 0.00 2.80 1.77 1.71

## 1003 2.51 2.80 0.00 1.09 1.53

## 1004 2.46 1.77 1.09 0.00 1.11

## 1005 1.36 1.71 1.53 1.11 0.00
```

Classes of Blocking Algorithms

➤ Optimal: consider all blockings; pick best. (High-order problem)

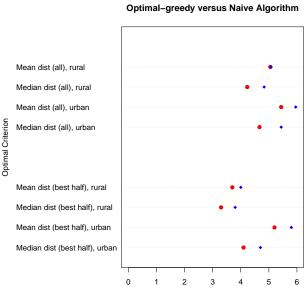
Classes of Blocking Algorithms

- ➤ Optimal: consider all blockings; pick best. (High-order problem)
- ▶ Optimal-greedy: consider all distances, pick best.

Classes of Blocking Algorithms

- ➤ Optimal: consider all blockings; pick best. (High-order problem)
- ▶ Optimal-greedy: consider all distances, pick best.
- ▶ Naive greedy: Get best block for this unit now.

Which Multivariate Blocking Algorithm? (optimal greedy)



Average Mahalanobis Distance between Pairs

Deploying Limited Resources: opt-greedy algorithm

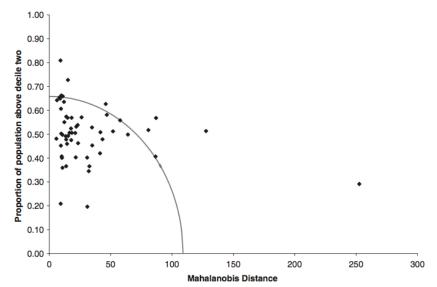
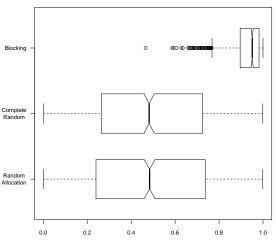


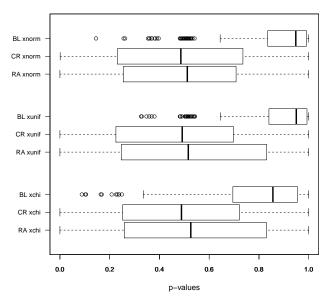
Figure 4. Choosing a subset of pairs for survey. We conducted our survey in 100 of $\frac{103}{203}$

Simulation study: 100 units, $X_1 \sim N(0,1)$, $X_2 \sim \text{Unif}(0,1)$, $X_3 \sim \chi_2^2$; 1000 such experiments. Assg treatmnt in 3 ways. p-value from xBalance.

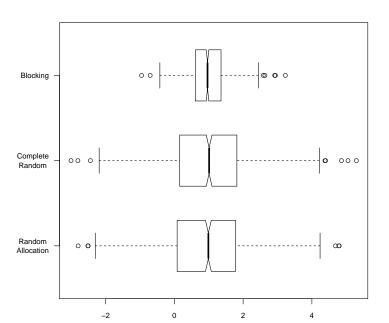
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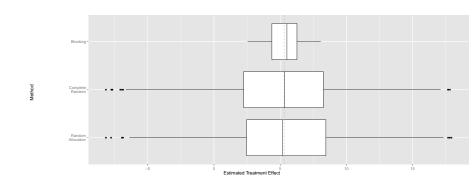
Kolmogorov-Smirnov Tests



Why Block: Efficiency



Why Block: Efficiency under TE Heterogeneity



Applications: Seguro Popular, Perry Preschool, Ignatieff Campaign

Experiment: Randomized Health Infrastructure & Insurance

King et al. (2009)

- ▶ **Intervention**: resources for medical services, preventive care, pharmaceuticals, access, and financial health protection
- ▶ Beneficiaries: 50M Mexicans (half the pop) w/o access to health care
- ▶ Cost 2005 full +1% GDP new money annually
- ▶ One of largest health reforms of any country in 2 decades
- Most visible accomplishment of Fox administration
- ► Major **issue** 2006 pres. campaign (vote choice effect: 7-11%)
- ▶ Randomized 74/148 health clusters to first roll-out, infrastructure spending, encouragement to enroll
- ▶ But do **better** than pure coin flip

How can we learn **most** from 148 cluster experiment?

- ► Ensure causes of HH expenditures evenly balanced betwn Tr and Co (otherwise, Tr might be poorer, or Co more insured)
- ▶ Make sure effect estimates as *precise* as possible (prefer estimate of "9 to 11%" to "−10 to 30%")

▶ Data on clusters:				
cluster	population	educ years	doctors	nurses
MCSSA000364	3130	4.00	1	1
MCSSA000504	6492	5.11	1	1
MCSSA008221	5096	4.26	2	2

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▶ Calculate how different clusters are from each other

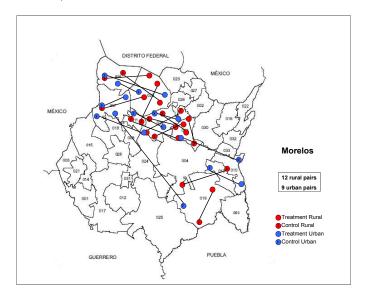
- ▶ Data on clusters:
 cluster
 population
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 MCSSA000364
 3130
 4.00
 1
 1

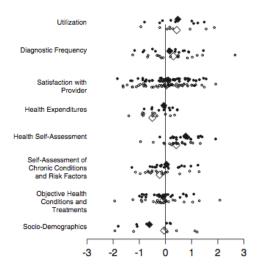
 MCSSA000504
 6492
 5.11
 1
 1

 MCSSA008221
 5096
 4.26
 2
 2
- ▶ Calculate how different clusters are from each other
- Mahalanobis distance between every pair of units: $MD_{ij} = \sqrt{(\mathbf{x}_i - \mathbf{x}_j)'\hat{\Sigma}^{-1}(\mathbf{x}_i - \mathbf{x}_j)}$

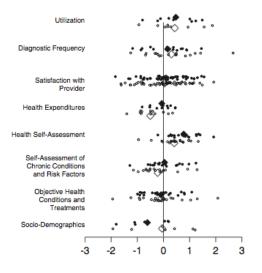
Blocked Pairs, Morelos



Balance in Baseline Outcomes, Seguro Popular King et al. (2007)



Balance in Baseline Outcomes, Seguro Popular King et al. (2007)



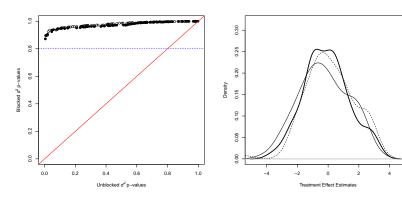
(Plus, SE's would have been $2 \times$ to $6 \times$ larger!)

Blocking in Applications: Balance and Efficiency

Moore (2012): Perry Preschool Experiment

Left: QQ plot of balance (100 blocked vs. unblocked)

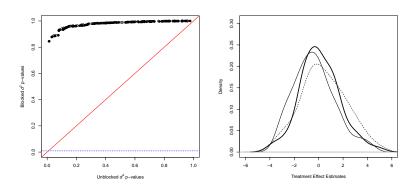
Right: Est TE under sharp null (100 blocked vs. unblocked)



(SES, sex, IQ)

Balance in Applications: Balance and Efficiency

Considering more variables ...



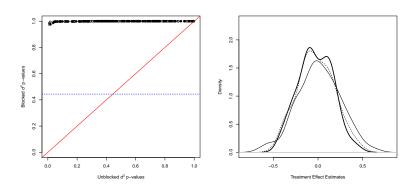
(+ siblings, AFDC, mom empl, educ, father, ...)

Balance in Applications

Loewen and Rubenson (2011): persuade party delegates to support Ignatieff

Left: QQ plot of balance (100 blocked vs. unblocked)

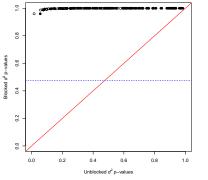
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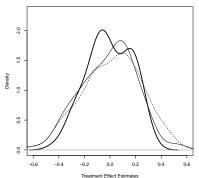


(province, pledged, special constituency)

Balance in Applications

Considering more variables ...





(+ attention, interest)

Moore and Schnakenberg (2023)

In R,

```
library(blockTools)
data(x100)
head(x100)
```

```
## id id2 b1 b2 g ig
## 1 1001 101 156 795 b 729
## 2 1002 102 813 469 a 627
## 3 1003 103 950 978 a 959
## 4 1004 104 991 781 a 661
## 5 1005 105 613 759 a 819
## 6 1006 106 654 838 b 643
```

Block:

```
block.out <- block(data = x100, id.vars = "id",
block.vars = c("b1", "b2"))
```

Block:

```
##
    Unit 1 Unit 2 Distance
      1043
             1040 0.01240000
## 1
## 2
    1100 1020 0.02259275
## 3
    1065
             1027 0.02912651
             1081 0.03498815
## 4
    1085
## 5
      1088
             1061 0.04789253
```

Assign:

```
assg.out <- assignment(block.out, seed = 157)</pre>
```

Assign:

```
assg.out <- assignment(block.out, seed = 157)</pre>
```

##		Treatment 1	Treatment 2	Distance
##	1	1043	1040	0.01240000
##	2	1020	1100	0.02259275
##	3	1027	1065	0.02912651
##	4	1085	1081	0.03498815
##	5	1061	1088	0.04789253

Diagnose:

Diagnose:

Get block IDs

```
createBlockIDs(assg.out, data = x100, id.var = "id")
```

Diagnose:

Get block IDs

```
createBlockIDs(assg.out, data = x100, id.var = "id")
```

Get balance:

Extract conditions:

```
extract_conditions(assg.out, x100, id.var = "id")
## [1] 2 1 1 2 2 2 2 2 2 1 2 2 1 1 2 1 2 2 2 2 2 1 2 1 1 1 2
## [38] 1 1 2 2 1 1 1 2 2 2 1 2 2 2 1 2 1 1 2 1 2 1 2 1 1
## [75] 1 2 1 2 2 2 2 2 1 1 1 1 2 1 2 2 2 1 1 1 2 1
```

Extract conditions:

##

8

x100 |> mutate(condition = extract_conditions(assg.out, x10)

1 1 2 1 2 2 2 1

##		id	id2	b1	b2	g	ig	${\tt condition}$
##	1	1001	101	156	795	b	729	2
##	2	1002	102	813	469	a	627	1
##	3	1003	103	950	978	a	959	1

404 a 221

2 1002 102 813 469 a 627 1
3 1003 103 950 978 a 959 1
4 1004 104 991 781 a 661 2
5 1005 105 613 759 a 819 2
6 1006 106 654 838 b 643 2
7 1007 107 640 645 c 12 2

1008 108 681

1 2 1 2 2 2 2 1 1

Two Issues: Outliers and Interference

$$MD_{ij} = \sqrt{(\mathbf{x}_i - \mathbf{x}_j)'\Sigma^{-1}(\mathbf{x}_i - \mathbf{x}_j)}$$

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- \blacktriangleright How to estimate Σ ?

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 - resistant covariance
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 - manual weighting

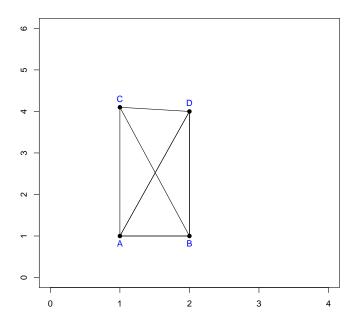
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 - ► MCD (Minimum Covariance Determinant) 1 constraint: include h points in interior

$$MD_{ij} = \sqrt{(\mathbf{x}_i - \mathbf{x}_j)'\Sigma^{-1}(\mathbf{x}_i - \mathbf{x}_j)}$$

- \blacktriangleright How to estimate Σ ?
 - resistant covariance
 - regular covariance
 - manual weighting
 - ▶ identity matrix (Euclidean dist)
- ► Resistant covariance estimators
 - ► MCD (Minimum Covariance Determinant) 1 constraint: include h points in interior
 - MVE (Minimum Volume Ellipsoid) 2 constraints: include h points, p+1 points on boundary



Given outliers, use resistant estimates of Σ and blocks will stay same

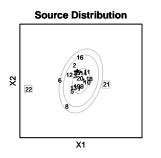
Covariate Weightings: Resistant Scaling

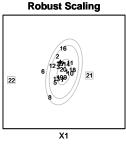
Given outliers, use resistant estimates of Σ and blocks will stay same:

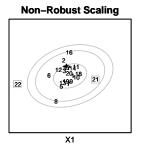
		Distance from A		
Scale Matx	Points	B	C	D
Cov	A-D	1.73	1.76	2.45
Cov	A-D + (1,100)	2.00	0.08	2.03
Cov	A-D + (20, 100)	1.72	1.01	0.74
MCD	A- D	1.73	1.76	2.45
MCD	A-D + (1,100)	1.73	1.76	2.45
MCD	A-D+(20,100)	1.73	1.76	2.45
MVE	A- D	1.73	1.76	2.45
MVE	A-D + (1,100)	1.73	1.76	2.45
MVE	A- $D + (20, 100)$	1.73	1.76	2.45

 $(\mathbf{Bold} = \mathbf{closest} \ \mathbf{to} \ A)$

Covariate Weightings: Resistant Scaling







Covariate Weightings

Baseline: Exclude outliers, regular MD

	Unit 1	Unit 2	Distance
1	15	4	0.18
2	9	7	0.37
3	13	5	0.37
4	18	10	0.43
5	17	3	0.44
6	14	11	0.50
7	20	1	0.65
8	16	2	0.95
9	12	6	1.32
10	19	8	2.12

Covariate Weightings

Include outliers, resistant MD: Same blocks! \odot

	Unit 1	Unit 2	Distance
1	15	4	0.18
2	13	5	0.37
3	9	7	0.42
4	18	10	0.46
5	17	3	0.50
6	14	11	0.59
7	20	1	0.76
8	16	2	1.03
9	12	6	1.56
10	19	8	2.18
11	22	21	13.43

Covariate Weightings

Include outliers, non-resistant MD: Blocks shuffled by outliers ©

	Unit 1	Unit 2	Distance
1	17	15	0.12
2	9	7	0.24
3	19	13	0.25
4	14	11	0.28
5	10	1	0.33
6	12	3	0.45
7	4	2	0.58
8	20	18	0.69
9	8	5	1.45
10	22	6	2.24
11	21	16	3.64

Interference

- ► We worry that effects may spillover: interference
- ▶ But nearby units are similar!
- ▶ Restrict blocks to units in a range

Multivariate Continuous Blocking with blockTools

Other arguments to block()

- vcov.data
- n.tr
- algorithm
- distance: mahalanobis, mcd, mve
- weight
- ▶ level.two: block states by most similar cities
- ▶ valid.var, valid.range: Goldilocks
- ▶ seed: (for mcd and mve)

Blocking for Sequential Experiments

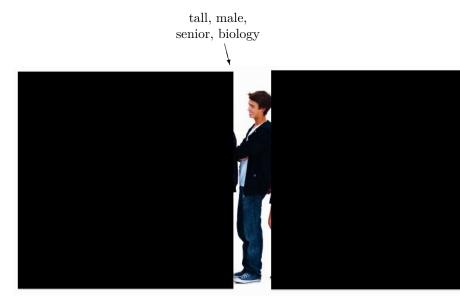
What about Sequential Experiments?

Moore and Moore (2013)



What about Sequential Experiments?

Moore and Moore (2013)



What about Sequential Experiments?

Moore and Moore (2013)

Have info!



Biased coins

Biased coins

Define unique covariate profile ${\bf p}$

Biased coins

Define unique covariate profile \mathbf{p}

use

$$Pr(t = T) = \begin{cases} \frac{2}{3} & \text{if } n_{\mathbf{p}T} < n_{\mathbf{p}C} \\ \frac{1}{3} & \text{if } n_{\mathbf{p}T} > n_{\mathbf{p}C} \end{cases}$$

Minimization

Minimization

Given **p**, score each variable s_j , $\left(\frac{n_{\mathbf{p}C} - n_{\mathbf{p}T}}{n_{\mathbf{p}C} + n_{\mathbf{p}T} + 1}\right)$ combine (sum), rank conditions, assign π 's

Minimization

Given **p**, score each variable s_j , $\left(\frac{n_{\mathbf{p}C} - n_{\mathbf{p}T}}{n_{\mathbf{p}C} + n_{\mathbf{p}T} + 1}\right)$ combine (sum), rank conditions, assign π 's

	Sex		Age	
	\mathbf{M}	\mathbf{F}	Young	Old
Control	1	3	3	3
Treatment	2	4	3	1
$s_j \text{ (Old M)}$	$-\frac{1}{4}$			$\frac{2}{5}$

Minimization

Given **p**, score each variable s_j , $\left(\frac{n_{\mathbf{p}C} - n_{\mathbf{p}T}}{n_{\mathbf{p}C} + n_{\mathbf{p}T} + 1}\right)$ combine (sum), rank conditions, assign π 's

	Sex		Age	
	M	\mathbf{F}	Young	Old
Control	1	3	3	3
Treatment	2	4	3	1
s_j (Old M)	$-\frac{1}{4}$			$\frac{2}{5}$

(bias toward Treatment)

1. Biased coins

→ require unique, replicated covariate profiles

- 1. Biased coins
 - → require unique, replicated covariate profiles
- 2. Minimization
 - → ignores joint distribution

		Sex		Age	Age	
Perfect balance?		M	F	Young	Old	
	Control	3	3	3	3	
	Treatment	3	3	3	3	

1. Biased coins

→ require unique, replicated covariate profiles

2. Minimization

→ ignores joint distribution

		$ $ S ϵ	ex	Age	Age	
Perfect balance?		M	\mathbf{F}	Young	Old	
	Control	3	3	3	3	
	Treatment	3	3	3	3	

Perfect imbalance?

	Young M	Old M	Young F	Old F
Control	3	0	0	3
Treatment	0	3	3	0

▶ Allow discrete covariates, restriction to exact profile **p**

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 - ▶ Define dissimilarity between current unit and conditions (Average pairwise Mahalanobis distance

$$\overline{MD}_{qt} = \frac{1}{R} \sum_{r=1}^{R} MD_{qr})$$

- ► Allow discrete covariates, restriction to exact profile **p**
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$$\overline{MD}_{qt} = \frac{1}{R} \sum_{r=1}^{R} MD_{qr})$$

▶ Assign higher prob to more dissimilar conditions

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$$\overline{MD}_{qt} = \frac{1}{R} \sum_{r=1}^{R} MD_{qr})$$

Assign higher prob to more dissimilar conditions (k-times, proportional to MD)

- Allow discrete covariates, restriction to exact profile p
- ▶ Move beyond counts in conditions
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$$\overline{MD}_{qt} = \frac{1}{R} \sum_{r=1}^{R} MD_{qr})$$

- Assign higher prob to more dissimilar conditions (k-times, proportional to MD)
- ► Test on variety of data

- ► Allow discrete covariates, restriction to exact profile **p**
- ► Move beyond counts in conditions
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$$\overline{MD}_{qt} = \frac{1}{R} \sum_{r=1}^{R} MD_{qr})$$

- Assign higher prob to more dissimilar conditions (k-times, proportional to MD)
- ► Test on variety of data
 - ► Uncorrelated MVN

- Allow discrete covariates, restriction to exact profile p
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- Assign higher prob to more dissimilar conditions (k-times, proportional to MD)
- ► Test on variety of data
 - Uncorrelated MVN
 - Correlated MVN
 - ► Correlated MVN with outliers

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$$\overline{MD}_{qt} = \frac{1}{R} \sum_{r=1}^{R} MD_{qr})$$

- Assign higher prob to more dissimilar conditions (k-times, proportional to MD)
- ► Test on variety of data
 - Uncorrelated MVN
 - Correlated MVN
 - ► Correlated MVN with outliers
 - ▶ Bimodal (extreme correlated MVN)

- ► Allow discrete covariates, restriction to exact profile **p**
- ► Move beyond counts in conditions
- ► Incorporate ordered, continuous covariates
 - ▶ Define dissimilarity between current unit and conditions (Average pairwise Mahalanobis distance

$$\overline{MD}_{qt} = \frac{1}{R} \sum_{r=1}^{R} MD_{qr})$$

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- ► Test on variety of data
 - Uncorrelated MVN
 - Correlated MVN
 - ► Correlated MVN with outliers
 - ▶ Bimodal (extreme correlated MVN)
 - ► Realistic data

My Approach

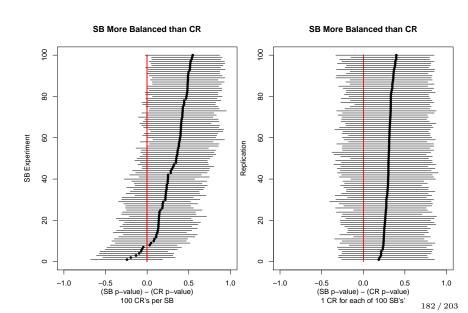
- ► Allow discrete covariates, restriction to exact profile **p**
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- ► Test on variety of data
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 - ► Realistic data
- ▶ Implement in actual sequential trial

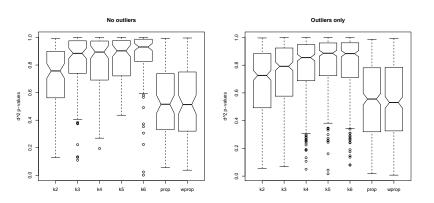
Seq. Blocking More Balanced than CR

Correlated MVN



Selecting a Method for Experiments

Simulate 100 datasets, each has (48 units, realistic means, SDs, bivariate corrs for 10 covariates)



- ▶ Background: PTSD impairs retrieval of specific autobiographical memories
- ▶ Question: Does practice lead to short-term changes in PTSD symptoms?

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- ▶ 52 assigned total; 46 assigned Nov 09 Jan 11; 39 follow-up.

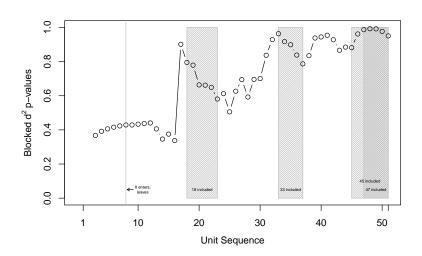
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 - Exclusion: psychotic disorder, bipolar disorder, substance abuse in last month, substance dependence in last year, recent psychiatric hospitalization, current suicidal intent
- ▶ 52 assigned total; 46 assigned Nov 09 Jan 11; 39 follow-up.
- ► Two Conditions (daily practice)
 - memory: retrieve specific memories of life events in response to cue words
 - ▶ anagrams: rearrange letters of cue words to create new words (control)

Inputting Data in Real Time: Generalized Query

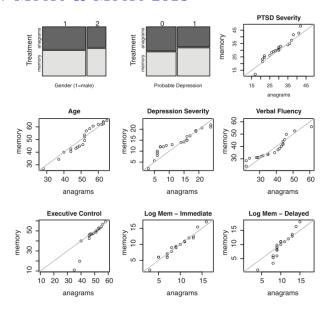
Unit 1

```
> seqblock1(query=T)
How many identification variables are there?
> 1
Enter the name of ID variable 1 without quotation marks.
> id
Enter the value of 'id'.
> 10624
How many exact blocking variables are there?
> 0
How many blocking variables are there?
> 2
Enter the name of blocking variable 1.
> x1
Enter the value of 'x1'.
> 100
Should 'x1' be restricted to certain values? [n/y]
> no
Enter the name of blocking variable 2.
> x2
Enter the value of 'x2'.
> 80
How many experimental/treatment conditions are there?
> 2
```

Balance: Moore & Moore 2013



Balance: Moore & Moore 2013



Get RI confidence interval from sequentially blocked data:

Get RI confidence interval from sequentially blocked data:

```
set.seed(407357912)
df \leftarrow tibble(id = 1:50,
             prov = sample(1:2, 50, replace = TRUE),
             age = sample(seq(18, 55, 1), 50, replace)
             treat = sample(0:1, 50, replace = TRUE),
             yobs = treat + sample(15:20, 50, replace
```

Check summary statistics:

```
df |> group_by(treat) |> summarise(m yobs = mean(yobs)
```

treat m yobs ## <int> <dbl> ## 1 0 17.7

A tibble: 2 x 2

2 1 18.3 193 / 203

Get RI confidence interval from sequentially blocked data:

Get RI confidence interval from sequentially blocked data:

```
$ci95

[1] -0.16  0.11  0.37  0.63  0.89  1.16  1.42

$ci90

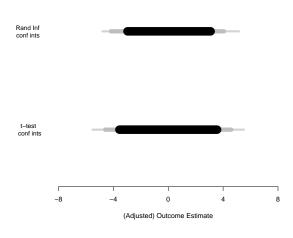
[1] -0.16  0.11  0.37  0.63  0.89  1.16

$ci80

[1]  0.11  0.37  0.63  0.89  1.16
```

RI Confidence Intervals: Moore & Moore 2013

Blocked estimates 9-15% more efficient



Analysis for Blocked Designs

If assignment probability for i varies by block j, let $p_{ij} = m_{ij}/N_j$. Weight each obs with

$$w_{ij} = \frac{d_i}{p_{ij}} + \frac{1 - d_i}{1 - p_{ij}}$$

Testing for Blocked Designs

'As ye randomise so shall ye analyse'

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If blocks used to randomize, incorporate into RI.

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'As ye randomise so shall ye analyse'

If blocks used to randomize, incorporate into RI.

- 1. Use actual blocks created, when possible
- 2. Reassign hypothetical treatment 1000 times
- 3. Obtain distribution of ATEs under (e.g.) sharp null
- 4. Compare observed ATE to randomisation distribution

Next: Clustered Designs Regression and Experiments

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