Introduction, Potential Outcomes, and Assignments

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Welcome

About Me

- Associate Prof of Government (American University)
- Associate Director, Center for Data Science (American University)
- Senior Social Scientist (The Lab @ DC)
- ➤ Fellow in Methodology (US Office of Evaluation Sciences: "OES")

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- Research agenda: political methodology, causal inference, experimental design, experiments in public policy

▶ Identify causal effects using the potential outcomes framework

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- ► (Walk through syllabus)

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- ▶ Often $\approx 1-2$ minutes
- ➤ We learn what we don't quite understand when we process/talk about it

"What caused the terror attacks of 9/11?"

VS.

"What is the effect of foreign policy X on domestic terror attacks?"

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"Causes of an effect"

VS.

"Effects of a cause"

We will focus on "effects of a cause"

We will focus on "effects of a cause", where the "cause" is well-defined.

"Canvassing": a systematic program where agents talk with residents (knock on doors) at the residents' home.



Figure 1: Credit: The Campaign Workshop

Citizen	Canvassed?	Enrolled?
1	Yes	Yes
2	Yes	Yes
3	No	No
4	No	No

Suppose we ask, "Would a canvassing policy increase enrollment in a health insurance program?"

Citizen	Canvassed?	Enrolled?
1	Yes	Yes
2	Yes	Yes
3	No	No
4	No	No

▶ What fraction enroll under canvassing vs. no canvassing?

Citizen	Canvassed?	Enrolled?
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- For each person canvassed, expect 1 more enrollment.)

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- ▶ What fraction enroll under canvassing vs. no canvassing?
- $\frac{2}{2} \frac{0}{2} = 1$
- For each person canvassed, expect 1 more enrollment.)
- ▶ Did the policy "work" (cause more enrollment)?

Motivating Example: Canvassing and Enrollment

But, is it causal?

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But, is it causal?

What do we really want to know?

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Would can vassing actually *change* anyone's enrollment?

Motivating Example: Canvassing and Enrollment

But, is it causal?

What do we really want to know?

Would can vassing actually *change* anyone's enrollment?

What would have happened under *other* conditions?

		Would Enroll if	Would Enroll if	
Citizen	Canvass?	Canvass?	No Canvass?	Enroll
1	Yes			Yes
2	Yes			Yes
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$$\frac{3}{4} - \frac{1}{4} = \frac{1}{2}$$

Empirical data consistent with $\it different\ unobserved$ outcomes:

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What is the true causal effect of canvass? What fraction enroll under canvass vs. no canvass? $\frac{4}{4} - \frac{1}{4} = \frac{3}{4}$

Well ... how do we know which?

Well ... how do we know which?

We can never know.

Can we know for one person?

Can we know for one person?

We can never know.

But I have some ideas.

But I have some ideas.

We could not canvass, then canvass later.

But I have some ideas.

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We can never know.

We can never observe both potential outcomes.

We can never observe both potential outcomes.

We can never observe both the factual and the counterfactual.

We can never observe both potential outcomes.

We can never observe both the factual and the counterfactual.

We can never know.

The Fundamental Problem of Causal Inference

We can never observe more than one potential outcome for a given unit.

So, how can we get a causal estimate?

So, how can we get a *causal* estimate?

We infer missing potential outcomes.

The problem with our naive estimate of effect:

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Knowing whether would "enroll under canvass" predicts whether canvassed!

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$$Pr(Canvass = Yes|Would if Canvassed = Yes) = \frac{2}{3}$$

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1	Yes	Yes	(Yes)	
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whether canvassed!

$$Pr(\text{Canvass} = \text{Yes}|\text{Would if Canvassed} = \text{Yes}) = \frac{2}{3}$$

But Pr(Canvass = Yes|Would if Canvassed = No) = 0

Yes Yes No

No

When comparing two groups does recover truth

Here, potential outcomes do **not** help predict Canvass:

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Knowing whether enroll if canvass not predictive.

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- What would we *observe* as the effect of canvass?

Good! ©

Neither *potential outcome* should help predict treatment/intervention.

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True underlying responses in "Canvass" group = True underlying responses in "No Canvass" group

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- ightharpoonup Aspirin \implies headache!
- ightharpoonup Canvass \implies turnout!
- ightharpoonup Insurance \implies health spending!
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But, full schedule of true underlying responses should **not** predict treatment.

How to ensure potential outcomes won't predict treatment?

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How to assign treatment so it won't predict potential outcomes?

Possible assignment mechanisms:

▶ Let Citizens decide whether to get Canvass

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- ➤ What if you randomly select whom gets Canvass? ("Citizen was Canvassed" won't help guess pot. out.) (We will estimate correct effect)

The Potential Outcomes Model

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The objective is to determine for some population of units ...the 'typical' causal effect of the [treatment vs. control conditions] on a dependent variable Y.—Rubin (1974)

▶ A "causal effect" is a comparative statement

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- Excludability (assignment effect only via Tr)

What is Causal Inference?

- Central definition for causal inference: "a well-defined treatment"
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- ➤ Timing of treatment: outcomes vs. covariates
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- ▶ One study, one causal effect

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- ► "No causation without manipulation"
- ➤ Timing of treatment: outcomes vs. covariates
- Exclusivity of treatment (to unit)
- Excludability (assignment effect only via Tr)
- ➤ One study, one causal effect (roughly)

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- ➤ Treatment: (putatively) causal variable of interest
- ▶ Potential outcome: outcome that would obtain if unit were to receive tr condition
- ► Assignment mechanism: means by which units come to be sorted into conditions

Stable Unit Treatment Value Assumption (SUTVA)

No versions of the treatment, varying in effectiveness

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- No versions of the treatment, varying in effectiveness
- No interference between units

▶ Units: index $i \in \{1, ..., 2n\}$

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- ▶ Pot outcome for *i* under $T_i = 1$: Y_{i1} or $Y_i(1)$

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- ▶ Pot outcome for *i* under $T_i = 0$: Y_{i0} or $Y_i(0)$

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- ▶ Binary treatment: $T_i \in \{0, 1\}$
- ▶ Pot outcome for *i* under $T_i = 1$: Y_{i1} or $Y_i(1)$
- ▶ Pot outcome for *i* under $T_i = 0$: Y_{i0} or $Y_i(0)$
- For vector of $Y_{i1} \quad \forall i$, write Y_1

Observed outcomes

▶ The observed outcome:

$$Y_i = Y_i(1) \cdot T_i + Y_i(0) \cdot (1-T_i)$$

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▶ If $T_i = 1$,

$$Y_i = Y_i(1) \cdot 1 + Y_i(0)(1-1) = Y_i(1)$$

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(sometimes written Y_i^{obs})

$$Y_i = Y_i(1) \cdot 1 + Y_i(0)(1-1) = Y_i(1)$$

▶ If $T_i = 0$,

$$Y_i = Y_i(1) \cdot 0 + Y_i(0)(1-0) = Y_i(0)$$

Observed outcomes

▶ The observed outcome:

$$Y_i = Y_i(1) \cdot T_i + Y_i(0) \cdot (1-T_i)$$

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$$Y_i = Y_i(1) \cdot 1 + Y_i(0)(1-1) = Y_i(1)$$

▶ If $T_i = 0$,

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Observed outcomes

▶ The observed outcome:

$$Y_i = Y_i(1) \cdot T_i + Y_i(0) \cdot (1-T_i)$$

(sometimes written Y_i^{obs})

 $\qquad \text{If } T_i = 1,$

$$Y_i = Y_i(1) \cdot 1 + Y_i(0)(1-1) = Y_i(1)$$

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The assignment mechanism *selects* which potential outcome we observe.

Statistical language

▶ parameter: unknown numeric value characterizing feature of prob model (Greek; θ , β)

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"The sample statistic \bar{x} is an estimator of true mean param μ ". μ is my estimand. 5.1 is my estimate.

For i, individual treatment effect

$$\tau_i = Y_i(1) - Y_i(0)$$

➤ True treatment effect?

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We can never know.

▶ Average treatment effect

$$ATE = \overline{Y_1 - Y_0} = \overline{Y_1} - \overline{Y_0}$$

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$$\begin{split} \overline{\tau} &= \overline{TE} = ATE &\equiv \overline{Y_1 - Y_0} \\ &= \frac{1}{2n} \sum_{i=1}^{2n} (Y_{i1} - Y_{i0}) \end{split}$$

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Statistical Independence

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Define

$$Y \perp \!\!\! \perp X$$

as

$$Pr(Y|X) = Pr(Y)$$

Statistical Independence

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$$Y \perp \!\!\! \perp X$$

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(Knowing X doesn't change probability of Y)

Something we can calculate:

If we know (Y_1, Y_0) indep of T

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If we know (Y_1,Y_0) indep of T (i.e., knowing true underlying response does ${f not}$ indicate T)

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Observed diff in Tr and Co group means gives \overline{TE} !

Why Experiments?

Random assignment of treatment promotes

$$(Y_1,Y_0) \perp \!\!\! \perp T$$

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Why Experiments?

Random assignment of treatment promotes

$$(Y_1,Y_0) \perp \!\!\! \perp T$$

When (Y_1, Y_0) and T independent, we can substitute things we know

$$(\overline{Y_1}|T=1)$$
 and $(\overline{Y_0}|T=0)$

for things we can *never* know

$$\overline{Y_1}$$
 and $\overline{Y_0}$

Common Assumptions, Null Hyp's in Causal Inference

Constant effect:

$$\tau_i = Y_{i1} - Y_{i0} = \tau \quad \forall i$$

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▶ Sharp null hypothesis of no effect:

$$\tau_i = 0$$

Assigned Tr	Assigned Co	Type
Tr	Со	Complier

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Tr	Co	Complier
Tr	Tr	Always-taker
Co	Co	Never-taker
Co	Tr	Defier

Of course, we cannot observe i assigned both Tr and Co.

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		Treatment Taken	
		Treatment	Control
Tr Assigned	Tr	Complier/Always	Never/Defier
	Co	Always/Defier	Complier/Never

Common Estimates under Noncompliance

Let T_i be treatment assigned, D_i be treatment received.

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► Intent-to-treat effect (as-assigned)

$$\begin{split} ITT &=& (\overline{Y_1}|T=1) - (\overline{Y_0}|T=0) \\ &=& (\overline{Y_1}|T=1, D(T=1)) - (\overline{Y_0}|T=0, D(T=0)) \end{split}$$

Common Estimates under Noncompliance

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➤ As-treated effect

$$ASTRE \ = \ (\overline{Y_1}|D=1) - (\overline{Y_0}|D=0)$$

$$E(\hat{\theta}) = \theta$$

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$$\blacktriangleright \ E\left[\widehat{\overline{Y_1 - Y_0}}\right] = \overline{Y_1 - Y_0}$$

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Is it a "good" estimator?

The difference-in-means estimator is *unbiased* for the true average treatment effect.

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(See 01-unbiased.R)

▶ Individual TE

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► Average treatment effect

$$ATE = E(Y_1-Y_0) = E(Y_1) - E(Y_0) \label{eq:atention}$$

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$$\begin{split} E(Y_1 - Y_0) &= \frac{1}{2n} \sum_{i=1}^{2n} (Y_{i1} - Y_{i0}) \\ &= \frac{1}{2n} \sum_{i=1}^{2n} (Y_{i1}) - \frac{1}{2n} \sum_{i=1}^{2n} (Y_{i0}) \\ &= E(Y_1) - E(Y_0) \end{split}$$

- ▶ If we know (Y_1, Y_0) indep of T
- Then,

$$\begin{split} E\left(Y_{1}\right) &=& E(Y_{1}|T=1) \\ E\left(Y_{0}\right) &=& E(Y_{0}|T=0) \end{split}$$

▶ Then, can substitute

$$\begin{array}{lcl} ATE & = & E(Y_1) - E(Y_0) \\ & = & E(Y_1|T=1) - E(Y_0|T=0) \end{array}$$

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= $E(Y_1|T=1) - E(Y_0|T=0)$

Observed diff in Tr and Co group means gives ATE!

$$E(Y_t|S=t) - E(Y_c|S=c) \label{eq:energy}$$

Holland (1986): "prima facie effect":

$$E(Y_t|S=t) - E(Y_c|S=c) \label{eq:energy}$$

"It is important to recognize that $E(Y_t)$ and $E(Y_t|S=t)$ are not the same thing ..."

Gerber and Green (2012):

When Y(1) and Y(0) indep of T,

$$ATE = E(Y_i(1)|T_i=1) - E(Y_i(0)|T_i=0)$$

When
$$Y(1)$$
 and $Y(0)$ indep of T ,

$$ATE = E(Y_i(1)|T_i = 1) - E(Y_i(0)|T_i = 0)$$

$$= E(Y_i(1)|T_i = 1) - E(Y_i(0)|T_i = 1) + E(Y_i(0)|T_i = 1) - E(Y_i(0)|T_i = 0)$$

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$$\begin{split} ATE &= E(Y_i(1)|T_i=1) - E(Y_i(0)|T_i=0) \\ &= E(Y_i(1)|T_i=1) - \underbrace{E(Y_i(0)|T_i=1)}_{E(Y_i(0)|T_i=1) - E(Y_i(0)|T_i=0)}_{E(Y_i(1)|T_i=1) - E(Y_i(0)|T_i=1)] + \\ &= [E(Y_i(0)|T_i=1) - E(Y_i(0)|T_i=0)] \end{split}$$

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$$\underbrace{E(Y_i(1) - Y_i(0) | T_i = 1)}_{\text{ATT}} + \underbrace{E(Y_i(0) | T_i = 1) - E(Y_i(0) | T_i = 0)}_{\text{Selection Bias}}$$

Assignment of Treatment

Timing clearly defined

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- Covariates: causally prior to treatment

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Sen and Wasow (2016): "Race as a Bundle of Sticks: Designs that Estimate Effects of Seemingly Immutable Characteristics" (elements of attributes varyingly manipulable)

Attributes

```
"Causal effect of race"?
```

```
data(resume, package = "qss")
dim(resume)
```

[1] 4870 4

Attributes

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```
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	0	1
black	2278	157
white	2200	235

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Bertrand and Mullainathan (2004): "Are Emily and Greg More Employable Than Lakisha and Jamal? A Field Experiment on Labor Market Discrimination"

Simple/complete randomization (Bernoulli trial, prob π)

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- Cluster randomizations
 (assignment at higher level)

The Potential Outcomes Model: Assignment

Observed outcome:
$$Y_i = Y_i(1) \cdot T_i + Y_i(0) \cdot (1 - T_i)$$

The assignment mechanism *selects* which potential outcome we observe.

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The assignment mechanism *selects* which potential outcome we observe.

(Gerber and Green (2012) use $Y_i = Y_i(1) \cdot d_i + Y_i(0) \cdot (1-d_i)$ to highlight that we observe pot outcome from treatment actually taken, not hypothetical or assigned treatment.)

The Potential Outcomes Model: Assignment

"Assignment mechanisms" are really missing-data-generating procedures.

Ignorability

Assignment mechanism is ignorable if Y_{mis} conditnly indep of T

$$P(T|X,Y_{obs},Y_{mis}) = P(T|X,Y_{obs}) \label{eq:problem}$$

Ignorability

Assignment mechanism is ignorable if Y_{mis} conditnly indep of T

$$P(T|X,Y_{obs},Y_{mis}) = P(T|X,Y_{obs}) \label{eq:problem}$$

Nothing in unobserved Y_{mis} informs relationship between Y_{obs} , T.

Unconfoundedness

Some ignorable mechanisms are *unconfounded*, too.

$$P(T|X,Y_{obs},Y_{mis}) = P(T|X) \label{eq:problem}$$

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Unconfoundedness

Some ignorable mechanisms are *unconfounded*, too.

$$P(T|X,Y_{obs},Y_{mis}) = P(T|X)$$

Nothing in Y informs T.

These are special cases of conditional independence.

An Assignment Mechanism

Little & Rubin (2000):

Patient	Y	Т
1	6	1
2	12	1
3	9	0
4	11	0

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Clearly, treatment is harmful. $\overline{Y(1)} - \overline{Y(0)} = 9 - 10 = -1$.

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Mean	10	9		

Clearly, treatment is harmful.

$$\overline{Y(1)|T=1} - \overline{Y(0)|T=0} = 9 - 10 = -1$$

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Patient	Y(0)	Y(1)	au	Т
1	(1)	6		1
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4	11	(10)		0
Mean	10	9		

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Patient	Y(0)	Y(1)	au	Т
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Clearly, treatment is beneficial:

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			(~)	
Mean	10	9	(3)	

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This assg mechanism is non-ignorable, confounded.

So, the assignment mechanism defines

So, the assignment mechanism defines how causal our empirical estimate may be.

Next: Inference for Experiments

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Inference for Experiments (or, "There is only one test")

- Bertrand, Marianne, and Sendhil Mullainathan. 2004. "Are Emily and Greg More Employable Than Lakisha and Jamal? A Field Experiment on Labor Market Discrimination." *American Economic Review* 94 (4): 991–1013.
- Gerber, Alan S., and Donald P. Green. 2012. Field Experiments: Design, Analysis, and Interpretation. New York, NY: WW Norton.
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