Blocking

Ryan T. Moore

American University

The Lab @ DC

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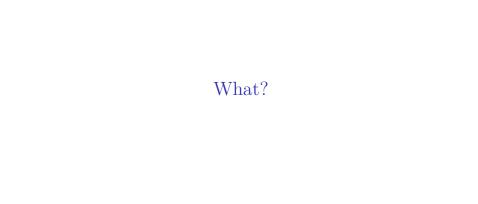
What?

Why?

How: blockTools

How: randomizr

Then what?



"Would a canvassing policy increase enrollment in a health insurance program?"

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| Precinct | Party | Canvass? | Enroll % |
|----------|-------|----------|----------|
| 1 | Dem | | |
| 2 | Dem | | |
| 3 | Rep | | |
| 4 | Rep | | |

Suppose we observationally measure

| Precinct | Party | Canvass? | Enroll $\%$ |
|----------|-------|----------------|-------------|
| 1 | Dem | Yes | 60 |
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| 4 | Rep | No | 30 |
| | | Diff in Means: | 40 |
| | | (Yes - No) | |

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Causal claims? Concerns?

Suppose we randomly assign 2 Tr, 2 Co, and measure

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Seriously? Well, ...

Figure 1: Unlucky

$$T_i \sim \mathrm{Bern}(\pi)$$

- ► → sample imbalance
- ▶ Might get no treateds!
- ► CLT SE for diff in means:

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- CLT SE for diff in means:

$$SE(\widehat{ATE}) = \sqrt{\frac{1}{N-1} \left[\frac{m \mathrm{Var}(Y_i(0))}{N-m} + \frac{(N-m) \mathrm{Var}(Y_i(1))}{m} + 2 \mathrm{Cov}(Y_i(0), Y_i(1)) \right]} + \frac{1}{N-m} \left[\frac{m \mathrm{Var}(Y_i(0))}{N-m} + \frac{(N-m) \mathrm{Var}(Y_i(1))}{m} + 2 \mathrm{Cov}(Y_i(0), Y_i(1)) \right]} \right]$$

If $Var(Y_i(0)) = Var(Y_i(1))$, allocate units equally.

$$T_i \sim \mathrm{Bern}(\pi)$$

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- CLT SE for diff in means:

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If $Var(Y_i(0)) = Var(Y_i(1))$, allocate units equally. Say 5 treated, 4 control. What to do with 10th village?

$$T_i \sim \mathrm{Bern}(\pi)$$

- >>> sample imbalance
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 - ▶ 10^{th} → treated: $SE: \sqrt{\frac{6}{4} + \frac{4}{6}} = \sqrt{2.66}$

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$$1 \quad \lceil m \operatorname{Var}(Y_{\cdot}(0)) \rceil$$

- $SE(\widehat{ATE}) = \sqrt{\frac{1}{N-1}} \left[\frac{m \mathrm{Var}(Y_i(0))}{N-m} + \frac{(N-m) \mathrm{Var}(Y_i(1))}{m} + 2 \mathrm{Cov}(Y_i(0), Y_i(1)) \right]$
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 - If $Var(Y_i(0)) \neq Var(Y_i(1))$, allocate \rightarrow higher-Variance condition

$$SE(\widehat{ATE}) = \sqrt{\frac{1}{N-1}} \left[\frac{m \mathrm{Var}(Y_i(0))}{N-m} + \frac{(N-m) \mathrm{Var}(Y_i(1))}{m} + 2 \mathrm{Cov}(Y_i(0), Y_i(1)) \right]$$

If $Cov(Y_i(0), Y_i(1)) > 0$, larger SE, less precision

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Demonstration:

| Unit | Y_0 | Y_1 (+ cov) | $Y_1 (- cov)$ |
|-------|-------|---------------|---------------|
| 1 | 0 | 0 | 10 |
| 2 | 5 | 5 | 5 |
| 3 | 10 | 10 | 0 |
| Means | 5 | 5 | 5 |

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 $\widehat{ATE}_{+\text{cov}} = -7.5, 0, 7.5$

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$$\widehat{ATE}_{+\text{cov}} = -7.5, 0, 7.5$$

 $\widehat{ATE}_{cov} = 2.5, 0, -2.5$

▶ If
$$Cov(Y_i(0), Y_i(1)) > 0$$
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Unit V V (+ corr) V (corr)

Demonstration:

| Unit | r_0 | $Y_1 (+ \text{cov})$ | $Y_1 (- \text{cov})$ |
|-------|-------------|----------------------|---|
| 1 | 0 | 0 | 10 |
| 2 | 5 | 5 | 5 |
| 3 | 10 | 10 | 0 |
| Means | 5 | 5 | 5 |
| | 1 2 3 | 1 0 2 5 3 10 | $ \begin{array}{cccccccccccccccccccccccccccccccccccc$ |

Suppose assign 1 to Tr, 2 to Co.

$$\widehat{ATE}_{+\text{cov}} = -7.5, 0, 7.5$$

 $\widehat{ATE}_{-\text{cov}} = 2.5, 0, -2.5 \text{ (less variance!)}$

But, is obs difference causal?

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What do we really want to know?

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Does can vassing actually <u>change</u> enrollment in precinct? (Or, just Party \rightarrow Enrollment?)

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Does can vassing actually <u>change</u> enrollment in precinct? (Or, just Party \rightarrow Enrollment?)

What <u>would have</u> happened to "No" precincts if "Yes"?

But, is obs difference causal?

What do we really want to know?

Does can vassing actually <u>change</u> enrollment in precinct? (Or, just Party \rightarrow Enrollment?)

What would have happened to "No" precincts if "Yes"?

What would have happened under <u>other</u> conditions?

Suppose we can know both potential outcomes \dots

| | | | Enroll % | Enroll $\%$ |
|----------|-------|----------|---------------|-------------|
| Precinct | Party | Canvass? | if No Canvass | if Canvass |
| 1 | Dem | _ | 20 | 60 |
| 2 | Dem | _ | 30 | 70 |
| 3 | Rep | | 20 | 30 |
| 4 | Rep | _ | 30 | 40 |
| | | Means: | 25 | 50 |

Suppose we can know both <u>potential outcomes</u> ...

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$$ATE = 50 - 25 = 25$$

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| 4 | Rep | _ | 30 | 40 |
| | | Means: | 25 | 50 |

$$ATE = 50 - 25 = 25$$

(True or an estimate?)

Another way to think about same information:

| | | | Enroll % | Enroll % | True Precinct |
|----------|----------------------|----------|---------------|------------|---------------|
| Precinct | Party | Canvass? | if No Canvass | if Canvass | Effect |
| 1 | Dem | | 20 | 60 | 40 |
| 2 | Dem | _ | 30 | 70 | 40 |
| 3 | Rep | _ | 20 | 30 | 10 |
| 4 | Rep | _ | 30 | 40 | 10 |
| | | Means: | 25 | 50 | 25 |

$$ATE = (40 + 40 + 10 + 10)/4 = 25$$

The "Fundamental Problem of Causal Inference",

We can't observe both "Canvassed" and "Not Canvassed" for a precinct.

We can't observe both <u>potential outcomes</u> (counterfactuals).

So, how can we get a good causal estimate?

Suppose we observe \dots

| | | | Enroll $\%$ | Enroll $\%$ |
|----------|-------|----------|---------------|-------------|
| Precinct | Party | Canvass? | if No Canvass | if Canvass |
| 1 | Dem | Yes | | 60 |
| 2 | Dem | Yes | | 70 |
| 3 | Rep | No | 20 | |
| 4 | Rep | No | 30 | |
| | | Means: | 25 | 65 |

Estimated ATE = 65 - 25 = 40

Suppose we observe \dots

| | | | Enroll $\%$ | Enroll $\%$ |
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| 3 | Rep | No | 20 | |
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| | | Means: | 25 | 65 |

Estimated ATE =
$$65 - 25 = 40$$
 \odot (too big)

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| 1 | Dem | Yes | | 60 |
| 2 | Dem | No | 30 | |
| 3 | Rep | Yes | | 30 |
| 4 | Rep | No | 30 | |
| | | Means: | 30 | 45 |

Estimated ATE = 45 - 30 = 15

Or, we could have observed ...

| | | | Enroll % | Enroll $\%$ |
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| Precinct | Party | Canvass? | if No Canvass | if Canvass |
| 1 | Dem | Yes | | 60 |
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Estimated ATE = 45 - 30 = 15 \odot

(too small; closer)

| Assignments | Est ATE |
|-------------|---------|
| YYNN | 40 |

| Assignments | Est ATE |
|-------------|---------|
| YYNN | 40 |
| NYNY | 35 |
| YNNY | 25 |
| NYYN | 25 |
| YNYN | 15 |
| NNYY | 10 |
| | |

In our random allocation, possible data were

| Assignments | Est ATE |
|-------------|---------|
| YYNN | 40 |
| NYNY | 35 |
| YNNY | 25 |
| NYYN | 25 |
| YNYN | 15 |
| NNYY | 10 |
| | |

Some closer to truth

| Assignments | Est ATE |
|-------------|---------|
| YYNN | 40 |
| NYNY | 35 |
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| NYYN | 25 |
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| NNYY | 10 |

- Some closer to truth
- $E(\hat{\tau}) = \frac{1}{6} \cdot 40 + \frac{1}{6} \cdot 35 + \frac{1}{3} \cdot 25 + \frac{1}{6} \cdot 15 + \frac{1}{6} \cdot 10 = 25 = \bar{\tau}$

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- Some closer to truth
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In practice, we don't know all potential outcomes.

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Two assignments (YYNN and NNYY) leave treatment perfectly confounded with party.

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We can never see all of Y_1 , Y_0 . But we can see all of X!

In practice, we don't know all potential outcomes.

Two assignments (YYNN and NNYY) leave treatment perfectly confounded with party.

We can never see all of Y_1 , Y_0 . But we can see all of X! Let's ensure X does not predict T.

 $\underline{Blocking}:$

Creating pre-treatmnt groups that look same on $\underline{\text{predictors}}$.

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Creating pre-treatmnt groups that look same on predictors.

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| Precinct | Party | Canvass? | if No Canvass | if Canvass |
| 1 | Dem | | | |
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Blocking:

Creating pre-treatmnt groups that look same on <u>predictors</u>.

| | | | Enroll % | Enroll $\%$ |
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| 4 | Rep | | | |

(Then, randomize within groups.)

Blocking:

Creating pre-treatmnt groups that look same on predictors.

| | | | Enroll $\%$ | Enroll $\%$ |
|----------|-------|----------|---------------|-------------|
| Precinct | Party | Canvass? | if No Canvass | if Canvass |
| 1 | Dem | Y | | 60 |
| 2 | Dem | N | 30 | |
| 3 | Rep | N | 20 | |
| 4 | Rep | Y | | 40 |

(Then, randomize within groups.)

Blocking restricts possible data to

| Assignments | Est TE |
|-------------|--------|
| YYNN | 40 |
| NYNY | 35 |
| YNNY | 25 |
| NYYN | 25 |
| YNYN | 15 |
| NNYY | 10 |

Blocking restricts possible data to

| Assignments | Est TE | |
|--|--------|--|
| YYNN | 40 | |
| NYNY | 35 | |
| $\mathbf{Y}\mathbf{N}\mathbf{N}\mathbf{Y}$ | 25 | |
| NYYN | 25 | |
| $\mathbf{Y}\mathbf{N}\mathbf{Y}\mathbf{N}$ | 15 | |
| NNYY | 10 | |

Estimates have less variance, are closer to true ATE.

Blocking restricts possible data to

| Assignments | Est TE |
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Estimates have less variance, are closer to true ATE.

Under random allocation, 5 possible estimates: 40, 15, 25, 35, 10. Blocking restricts to 3 best: 15, 25, 35.

Blocking restricts possible data to

| Assignments | Est TE |
|-------------|------------|
| YYNN | 40 |
| NYNY | 35 |
| YNNY | ${\bf 25}$ |
| NYYN | ${\bf 25}$ |
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Estimates have less variance, are closer to true ATE.

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Covariate balance

- Covariate balance
- Estimate closer to truth

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- ▶ Block-level effects
 - \rightsquigarrow different actors interested in different effects

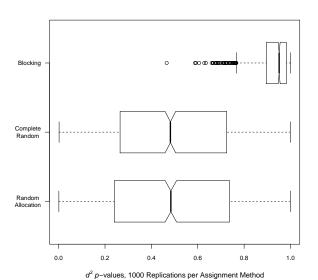
- Covariate balance
- Estimate closer to truth
- ► Increased efficiency
- ▶ Triply-robust estimates: block, randomize, adjust
- ▶ Block-level effects
 - \rightarrow different actors interested in different effects
- ▶ Guidelines for limited/uncertain resources

Why Block: Balance

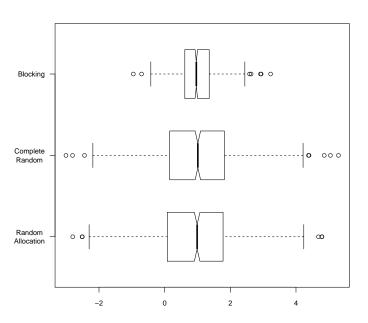
Simulation study: 100 units, $X_1 \sim N(0,1)$, $X_2 \sim \text{Unif}(0,1)$, $X_3 \sim \chi_2^2$; 1000 such experiments. Assg treatmnt in 3 ways.

Why Block: Balance

Simulation study: 100 units, $X_1 \sim N(0,1)$, $X_2 \sim \text{Unif}(0,1)$, $X_3 \sim \chi_2^2$; 1000 such experiments. Assg treatmnt in 3 ways.



Why Block: Efficiency

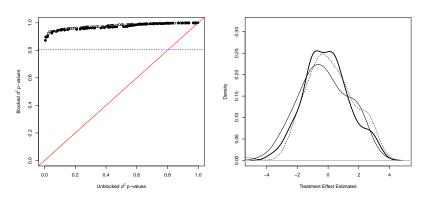


Blocking in Applications: Balance and Efficiency

Moore (2012): Perry Preschool Experiment

Left: QQ plot of balance (100 blocked vs. unblocked)

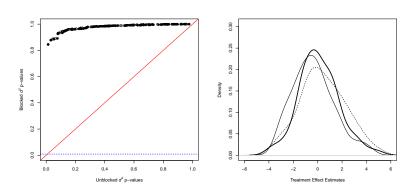
Right: Est TE under sharp null (100 blocked vs. unblocked)



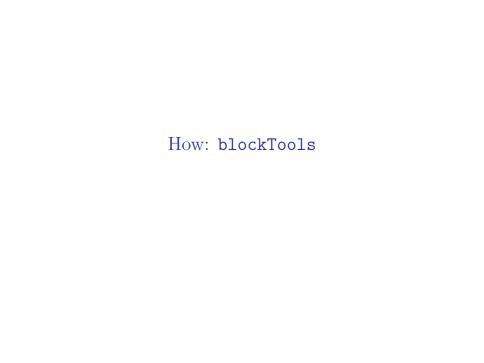
(SES, sex, IQ)

Balance in Applications: Balance and Efficiency

Considering more variables ...



(+ siblings, AFDC, mom empl, educ, father, ...)



Start with some sample data:

```
library(blockTools)
data(x100)

x100 |> head()
```

```
id id2 b1 b2 g ig
1 1001 101 156 795 b 729
2 1002 102 813 469 a 627
3 1003 103 950 978 a 959
4 1004 104 991 781 a 661
5 1005 105 613 759 a 819
6 1006 106 654 838 b 643
```

(Moore 2012; Moore and Schnakenberg 2023)

```
bl <- b$blocks$`1`
bl |> head()
```

```
Unit 1 Unit 2 Distance

1 1043 1040 0.01240000

2 1100 1020 0.02259275

3 1065 1027 0.02912651

4 1085 1081 0.03498815

5 1088 1061 0.04789253

6 1064 1014 0.07985116
```

Why all this?

```
bl <- b$blocks$`1`</pre>
```

We are extracting just the blocked pairs themselves.

▶ Why b\$blocks? Since b has 3 components:

```
names(b)
```

```
[1] "blocks" "level.two" "call"
```

▶ Why blocks\$1'? Since this is (default-named) first (and only) "group":

```
names(b$blocks)
```

```
[1] "1"
```

What else could we do?

```
b_3groups_3conditions <- block(</pre>
 x100,
 groups = "g",
                                 # (Factor variable in data)
 n.tr = 3,
 id.vars = "id",
 block.vars = c("b1", "b2"),
 distance = "mve"
```

```
b_3groups_3conditions$blocks
$a
   Unit 1 Unit 2 Unit 3 Max Distance
```

1 1076 1039 1056 0.2509889

1065 1061 1002 0.4335040 2 3 1084 1058 1017 0.4373315 1 O E O 1010 1 0 0 1 0 4E04767

Some rows from each "group":

```
rows_a <- b_3groups_3conditions$blocks$a |> slice(1:2) |> mutate
rows_b <- b_3groups_3conditions$blocks$b |> slice(1:2) |> mutate
rows_c <- b_3groups_3conditions$blocks$c |> slice(1:2) |> mutate
bind_rows(rows_a, rows_b, rows_c)
```

```
Unit 1 Unit 2 Unit 3 Max Distance group
   1076
         1039
              1056
                     0.2509889
                                а
2
   1065 1061 1002 0.4335040 a
   1043 1040 1019 0.1664608
3
                                b
4
   1048 1031 1062 0.2288718
                                h
5
   1095 1092 1049 0.3778655
                                С
6
   1045
         1015 1082
                     0.3808569
                                C.
```

Other arguments to block()

- vcov.data
- **proups:** for exact-blocks
- n.tr
- id.vars
- block.vars
- algorithm: optGreedy, optimal, naiveGreedy, randGreedy, sortGreedy
- \blacktriangleright distance: mahalanobis, mcd, mve, euclidean, $k \times k$ matx
- weight
- ▶ level.two: block states by most similar cities
- valid.var, valid.range: Goldilocks
- seed.dist: (for mcd and mve)

Assign

```
a <- assignment(b, seed = 71573706)
a</pre>
```

Assignments:

| | Treatment 1 | Treatment 2 | Distance |
|----|-------------|-------------|------------|
| 1 | 1040 | 1043 | 0.01240000 |
| 2 | 1100 | 1020 | 0.02259275 |
| 3 | 1065 | 1027 | 0.02912651 |
| 4 | 1081 | 1085 | 0.03498815 |
| 5 | 1088 | 1061 | 0.04789253 |
| 6 | 1014 | 1064 | 0.07985116 |
| 7 | 1032 | 1070 | 0.08279625 |
| 8 | 1097 | 1098 | 0.08882421 |
| 9 | 1038 | 1018 | 0.09316331 |
| 10 | 1031 | 1048 | 0.10391953 |
| 11 | 1084 | 1058 | 0.10835825 |

Get Assignments

```
a |> extract conditions(x100, id.var = "id")
 [1] 2 1 2 2 2 2 1 2 2 1 1 1 1 1 1 1 1 2 2 1 2 2 2 1 2 1 1 1
 [38] 1 1 1 2 2 2 2 2 2 2 1 2 1 2 2 2 2 1 1 2 1 2 1 2 2 1 3
 x100 |> mutate(
 condition = extract conditions(a, x100, id.var = "id"))
     id id2 b1 b2 g ig condition
   1001 101 156 795 b 729
   1002 102 813 469 a 627
   1003 103 950 978 a 959
3
4
   1004 104 991 781 a 661
5
   1005 105 613 759 a 819
   1006 106 654 838 b 643
6
   1007 107 640 645 c 12
8
   1008 108 681 404 a 221
```

Assign 3 Conditions, within Groups

```
a3 <- assignment(b_3groups_3conditions, seed = 979677744)
```

Assignments:

```
Group: a
    Treatment 1
                   Treatment 2
                                 Treatment 3
                                                Max Distance
    1056
                   1076
                                 1039
                                                0.2509889
2
    1002
                   1061
                                 1065
                                                0.4335040
3
    1058
                   1084
                                 1017
                                                0.4373315
4
    1046
                   1081
                                 1059
                                                0.4524767
5
    1073
                   1098
                                 1029
                                                0.4599752
6
    1060
                   1067
                                 1004
                                                0.5921446
    1089
                   1024
                                 1032
                                                0.7554402
8
    1052
                   1054
                                 1030
                                                0.8848444
9
    1093
                   1026
                                 1068
                                                0.8975495
10
    1036
                   1008
                                 1091
                                                1.3586913
```



Blocking with randomizr::block ra()

1007 107 640 645 c 12 1 1008 108 681 404 a 221

823 b 321

1009 109 530

8 9

```
library(randomizr)
tr <- block_ra(x100$g)
# Better:
x100 \mid > mutate(tr = block ra(x100$g))
      id id2 b1 b2 g ig tr
1
    1001 101 156 795 b 729 1
    1002 102 813 469 a 627 1
3
    1003 103 950 978 a 959 1
    1004 104 991 781 a 661 1
4
5
    1005 105 613 759 a 819 0
6
    1006 106 654 838 b 643 1
```



blockTools: diagnose, get block IDs, check balance

Diagnose:

blockTools: diagnose, get block IDs, check balance

Diagnose:

Get block IDs

```
createBlockIDs(a, data = x100, id.var = "id")
```

blockTools: diagnose, get block IDs, check balance

Diagnose:

Get block IDs

```
createBlockIDs(a, data = x100, id.var = "id")
```

Get balance:

```
assg2xBalance(a, x100, id.var = "id",
bal.vars = c("b1", "b2"))
```

▶ Generally, use Lin or Blocked Diff-in-Means

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- ▶ LSDV (block indicators) weights w/in block TE's by

$$p_j(1-p_j)n_j \\$$

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- ▶ LSDV (block indicators) weights w/in block TE's by

$$p_j(1-p_j)n_j$$

where

- $\triangleright p_i = \text{share of block } j \text{ treated}$
- $n_i = \text{size of block } j$
- $(I.e., p_j(1 p_j) = var(TE) \text{ in block } j)$

- ► Generally, use Lin or Blocked Diff-in-Means
- LSDV (block indicators) weights w/in block TE's by

$$p_j(1-p_j)n_j$$

where

- $p_i = \text{share of block } j \text{ treated}$
- $n_i = \text{size of block } j$
- (i.e., $p_i(1-p_i) = var(TE)$ in block j)

Safer when

- $\triangleright p_i$ constant across blocks j
- n_i constant across blocks j

```
library(estimatr)
df \leftarrow x100 \mid x1
lm lin(b1 ~ tr, covariates = ~ g, data = df)
                                                                                                      Estimate Std. Error t value Pr(>|t|)
 (Intercept) 492.800000 39.08723 12.60769882 6.405020e-29
                                      -50.760000 51.70890 -0.98164922 3.287925e-0
tr
gb_c -4.673611 100.16089 -0.04666104 9.628824e-0
gc_c -4.111111 92.84835 -0.04427770 9.647770e-0
tr:gb_c -190.763889 129.80730 -1.46959290 1.450111e-0
tr:gc_c -112.826389 128.65776 -0.87694974 3.827497e-0
```

DF

94

94

94

(Intercept) 94

tr

gb_c

gc c

Can I just ignore blocks and pool?

Can I just ignore blocks and pool?

ightharpoonup If p_j varies, no

Thanks!

rtm@american.edu
www.ryantmoore.org

References I

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 Assignment, and Diagnosing Interference in Randomized Experiments.
 http://www.ryantmoore.org/html/software.blockTools.html.