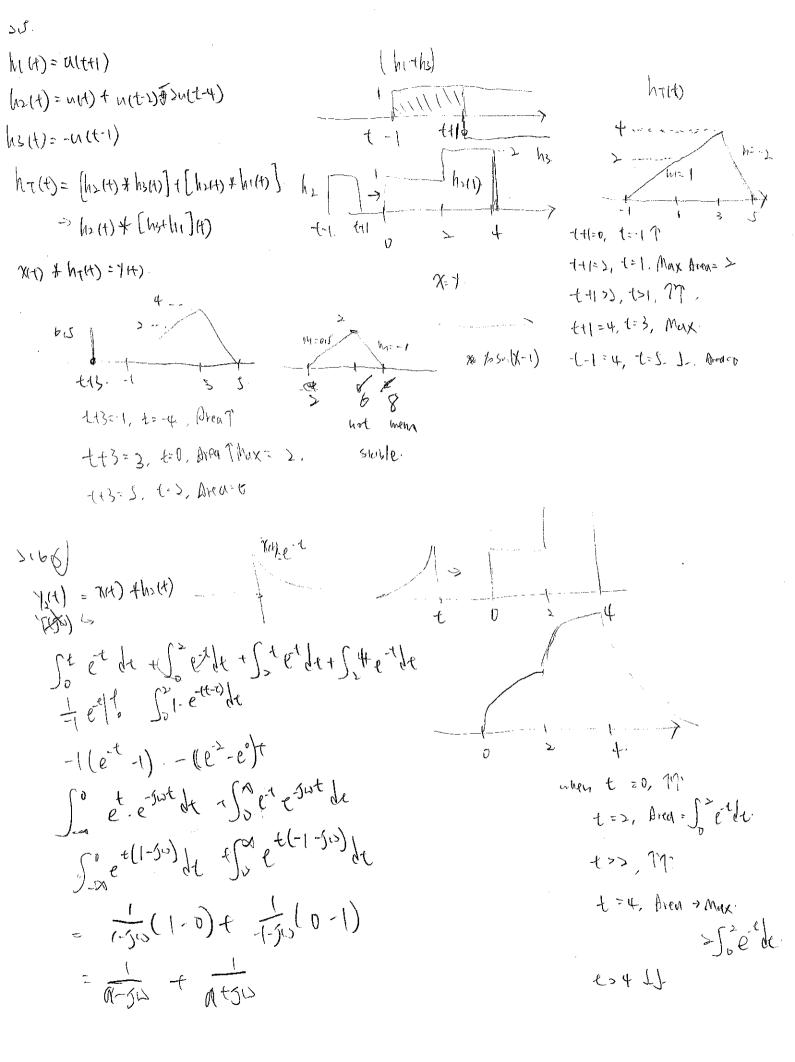
( 1 (1) It = 0 - ET. 1.4. pre-at-e-sent de unt) 1 5 t . (9(4) 2 = Pav Ja e-tlatsu) (4. Sp1: dt = 1, = ET 150 / elo > (0-1) = 250 45014 7(+)= 1 5t x 10/1/2 1 (t-to) = - State de St. 1. e swt h. 9 yet) = + [tx(t-to)dr > - 1/2 (e - JuT1 - e suT1). let i'= tto del = 1 9 2 e 51 - t 51 = > 5 5 5 h (0 Ti). I SINCY = SINTIY KX WII ( ) D = - 1 ) - x(z') dz' 、, 物質集 SINC TI = SMEATI > & SIN(UTI) J.F. T.

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> Sin(kusti) 25/n(kusti)
T kW. 7/1 11+) = \$ 8(t-T) 11+) = \$ 8(t-T) 11-> T 3kwt dt 11-> T 3kwt dt 11-> T 3kwt dt Sh(khf) 1+ k=0 (H) q(t)= X(toti) - X(t+ti) Q(k)= X(k)etskitres - X(r)etskist, 3 ( essent - essent) - 3/5 in( ku/li) Sort-e-sortle to the s = So etlissu)de x(r) eskwt 1(t) = X(t) \* ha). 8: (1+KW) XIA -> X (56). H(50) 7(1) - 1+1-11 x x 14 - (4521) 4

J= 1.2 d+ = >(4.2)=4

t-65 HOU W 3 T=1, \$1 TK). YOU = 2 TON. e That MO = + I Ay e shot de 156-1-eskut de t. I 5/1-eskut de 1/25kw (e° - e-5kw) + 31kw (e°). - e - e 1 (564(02(KD)) 1 (kg (kg))

hip > HOW= t In the ort H(J(w-5ty) = h(t)-esmt 1 tole = 0, t= 30, T h.(4). t. M=100, t=150. Lo - (litt-0). MI [x |hm| 4= 1 [x (+bm) dt. John to 7 2 SP PORT of

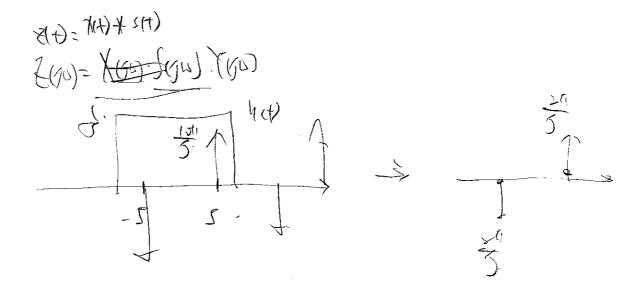
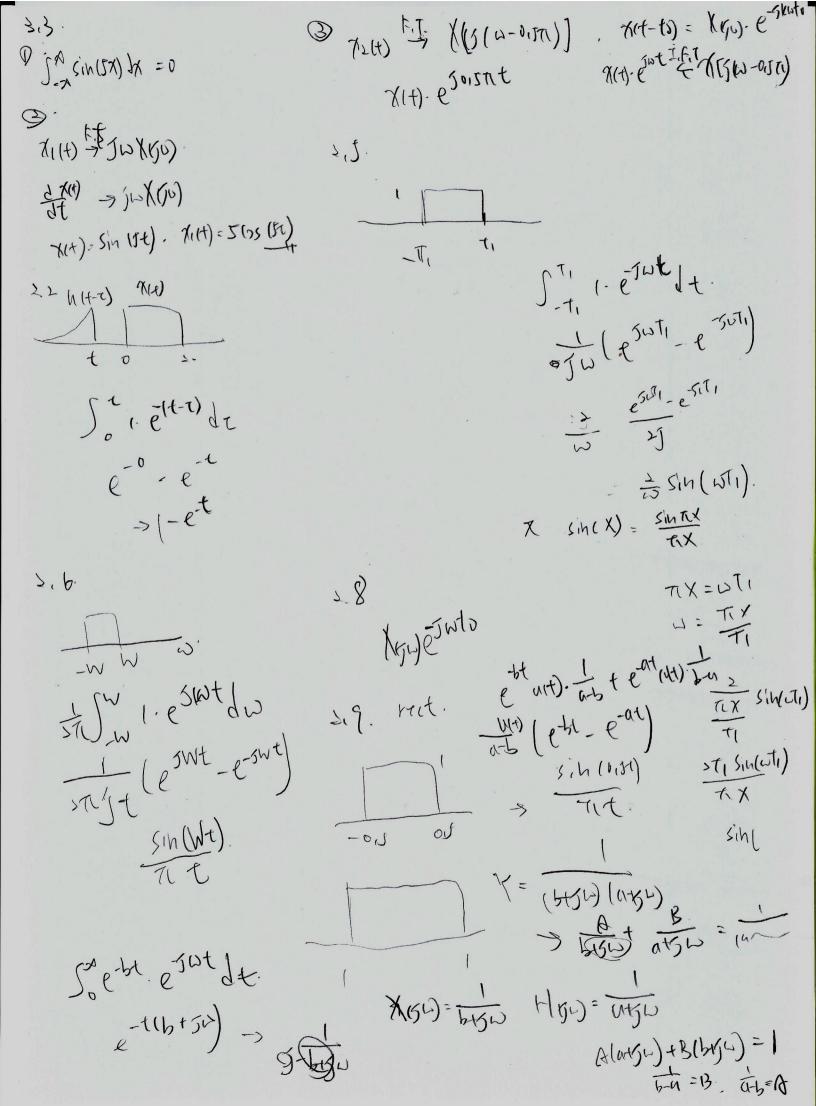
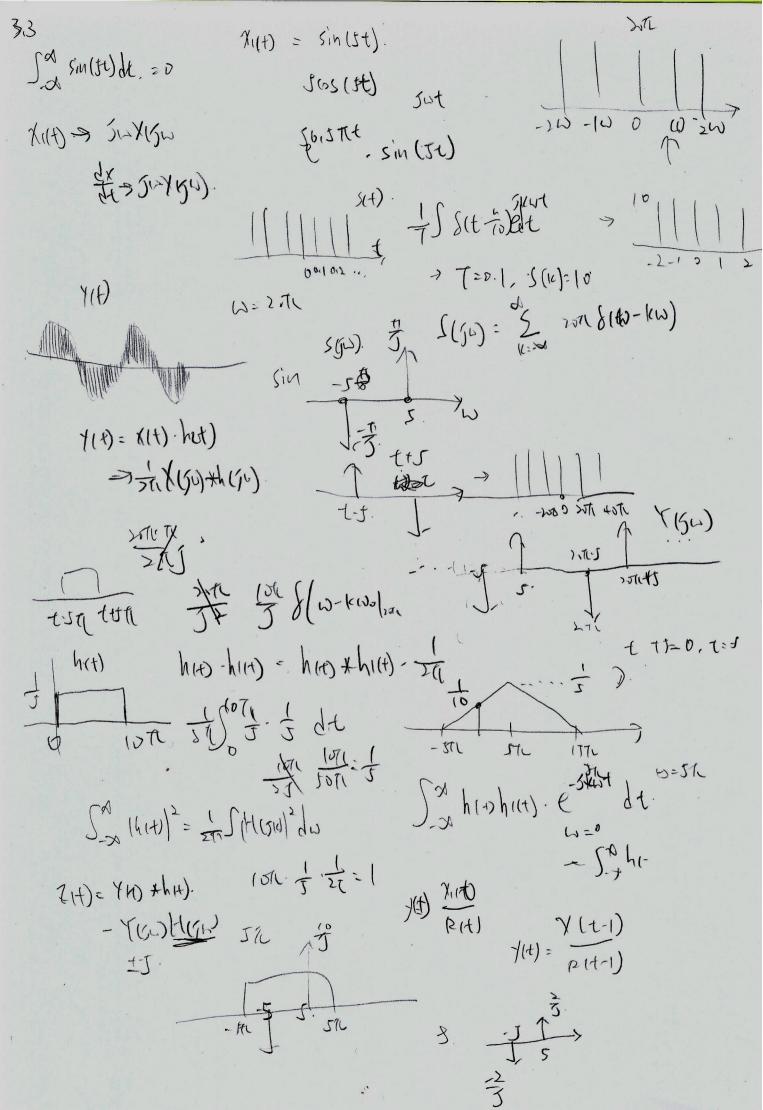


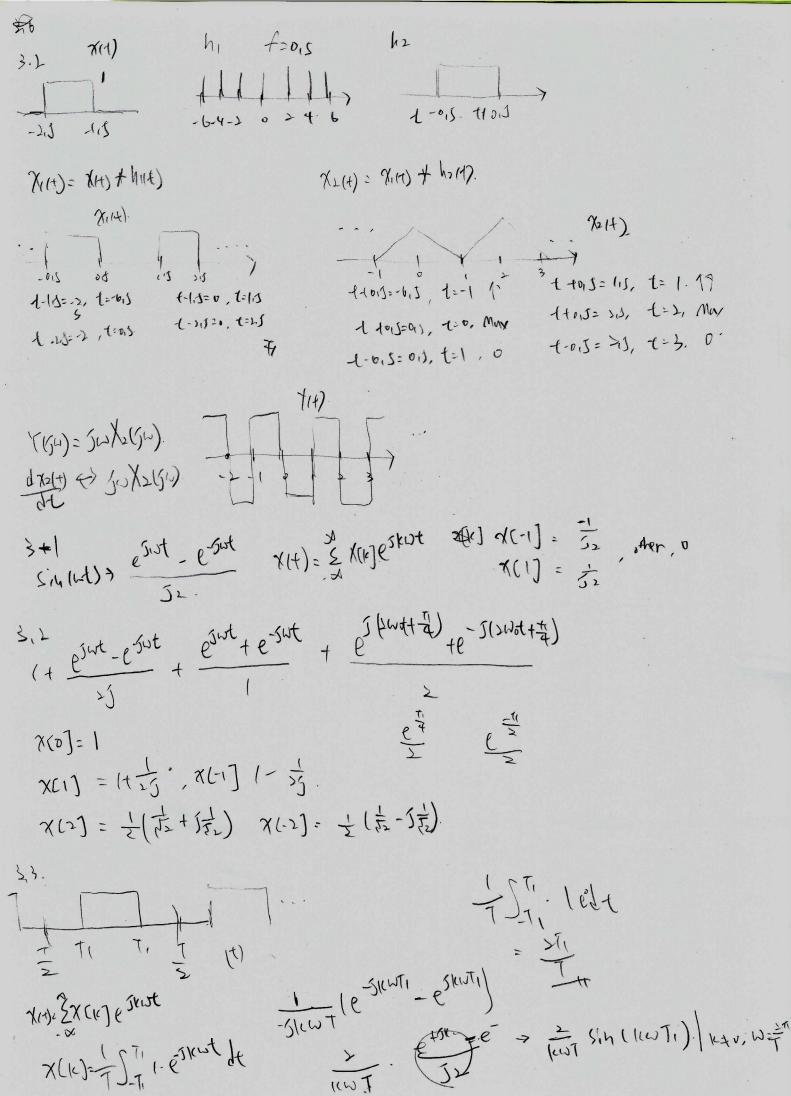
Table 1

行きます 3.2 KA=Sin (5t) D Sx similarly = 0 X(1) = X(1) - S(1) > Y(1) = X(1) X S(1) X 11 7-10-3 Twxgu). 果 Alt) = simist) NA) ( = 1 = 1 X(K) e - SKWt be TESSW (VU) 水丸社 X(Ju). X= \$x = 50 SEX)# ちかてち Lhtihth 3 ANOT 1 e 5001 - > /1510-044) 自転車 x e soont sin(si).e some (> Flx S(t) e skut dt t==> , (c=t) say = + 12 stoot 7 = 21 ともたち友をのいかっ かれ级 720,1 (C+3=1. かのじょかり kt= sa, k= 3% かぞく K.J. DoTL 100 1/4 -82-8.1 0 Bil 01 25/L-1 Hilos 151(-1 WHI. -100-60 0 sto 2000

2. £x(+) (+2) 50 X(50). 5 7/4) estable (+1+-5-05 Ja. A. du + Lod Jo- Edu + 5 1 205 du = 3602/01 + 550/00 + 550/01 7th ( 1,25 - 1 + 1-6,25) + # ( 0,52-6,52) = 0H (16) Signi (-37m) dw + Signi Jav (57m) dw So with dot - I gwardin 压了! - 至心! 歩(0-1) - で(1-0) = -た







# 離散

分佈名稱	Pmf	var	Mean
白努力 Bernoulli(p)	$P_{x}(x) = p, x = 1$	Var = p(1-p)	E(x) = p
	$P_{x}(x)=1-p, x=0$		
二項式 Binomial(n,p)	$C_k^n p^k (1-p)^{n-k}$	Var = np(1-p)	E(x) = np
	0, other		
幾何 Geometric(p)	$P_{\chi}(k) = (1-p)^{k-1}p$	$Var = \frac{1-p}{p^2}$	$F(r) = \frac{1}{r}$
	0, other	$var = \frac{1}{p^2}$	$E(x) = \frac{1}{p}$
Poisson( $\lambda$ )	$P_{x}(k) = e^{-\lambda} \frac{\lambda^{k}}{k!}$	$Var = \lambda$	$E(x) = \lambda$
	0, other		
Hypergeometric(R,W,n)	$P_{x}(k) = \frac{C_{k}^{R} C_{n-k}^{W}}{C_{n}^{R+W}}$	$n\frac{R}{R+W}\frac{W}{R+W}\frac{R+W-n}{R+W-1}$	$E(x) = \frac{R}{R + W}n$
負二項(p,r)	$C_{r-1}^{k-1}p^{r-1}(1 - p)^{k-r}p$	$Var = r \frac{1 - p}{p^2}$	$E(x) = r\frac{1}{p}$

連續

$$E(x) = \int_{-\infty}^{\infty} x f(x) dx, E(h(x)) = \int_{-\infty}^{\infty} h(x) f(x) dx$$

**Moments** 

$$E(x^n) = \int_{-\infty}^{\infty} x^n f(x) dx$$
,  $E(h(x))$ 

標準差σ

$$\sqrt{Var(x)} = \sigma_X = \sqrt{\sigma_X^2}$$

分佈名稱	PDF	var	Mean
Uniform(a,b)	$f(x) = \frac{1}{b-a}$	$\frac{1}{12}(b-a)^2$	$E(x) = \frac{b-a}{1}$
	$F(x) = \begin{cases} 0, x < a \\ \frac{x-a}{x-b}, a < x < b \\ 1, x > b \end{cases}$		
Exponentia( $\lambda$ )	$f(x) = u(x)\lambda e^{-\lambda x}$ $F(x) = u(x)(1 - e^{-\lambda x})$	$Var(x) = \frac{1}{\lambda^2}$	$E(x) = \frac{1}{\lambda}, E(x^2) = \frac{2}{\lambda^2}$
	$P(x > t) = 1 - F(x)$ $= e^{-\lambda t}$		
$Gamma(\pmb{lpha},\pmb{\lambda})$	$u(x)\lambda^{\alpha}\frac{x^{\alpha-1}}{T(\alpha)}e^{-\lambda x}$	$Var(x) = \frac{\alpha}{\lambda^2}$	$E(x)=\frac{\alpha}{\lambda}$
Weibull $(\alpha, \lambda)$	$u(x)\alpha\lambda(\lambda x)^{\alpha-1}e^{-(\lambda x)^{\alpha}}$ $F(x) = (1 - e^{-(\lambda x)^{\alpha}})$	$\frac{1}{\lambda^2} \left( T \left( 1 + \frac{2}{\alpha} \right) - T \left( 1 + \frac{1}{\alpha} \right)^2 \right)$	$E=\frac{1}{\lambda}\mathrm{T}(1+\frac{1}{\alpha})$
Gaussian( $\mu,\sigma^2$ )	$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}}e^{\frac{-(x-\mu)^2}{2\sigma^2}}$	$Var(x) = \sigma^2$	$E = \mu$
	$F(x) = \phi(\frac{x-\mu}{\sigma})$		

定理:

$$T(\alpha + 1) = \alpha T(\alpha)$$

範例:

$$\frac{5}{2}! = T\left(\frac{5}{2} + 1\right), = \frac{5}{2} \times T\left(\frac{5}{2}\right), = \frac{5}{2} \times T\left(\frac{3}{2} + 1\right) = \frac{5}{2} \times \frac{3}{2}T\left(\frac{3}{2}\right) = \frac{5}{2} \times \frac{3}{2}T\left(\frac{1}{2} + 1\right) = \frac{5}{2} \times \frac{3}{2} \times \frac{1}{2}T(\frac{1}{2})$$

又
$$T\left(\frac{1}{2}\right) = \sqrt{\pi}$$
,所以 $\frac{5}{2}! = \frac{5}{2} \times \frac{3}{2} \times \frac{1}{2} \times \sqrt{\pi} = \frac{15}{8}\sqrt{\pi}$ 

定理:

$$f_x|dx| = f_y|dy|, f_y = f_x|\frac{dy}{dx}|$$

範例:

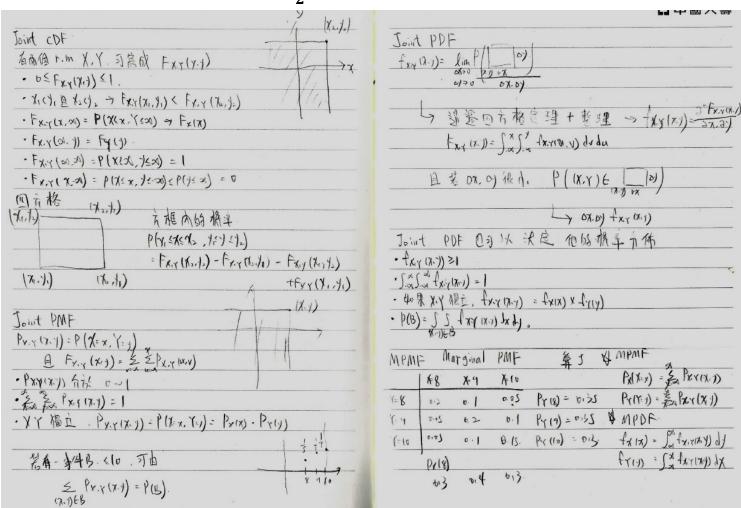
假設有一分佈

$$f_X(x) = \begin{cases} \frac{3}{2}(2x - x^2), & \text{if } 0 < x < 1 \\ 0, & \text{other} \end{cases}$$

有另外一個隨機變數Y = ln(X) ,求 $f_Y(y)$ 

可知 $x = e^y \circ \frac{dx}{dy} = (e^y)' = e^y$  所以根據 $f_y = f_x | \frac{dy}{dx} |$ 可得:

$$\frac{3}{2}(2e^{y}-e^{2y})e^{y}$$



### Joint 系列特性

- $\bullet \quad 0 \leq F_{XY}(x,y) \leq 1$
- $f_X(x) = \int_0^x f_{XY}(x, y) dy$ ,  $0 \le y \le x \le 1$
- $f_{X|Y}(X|Y) = \frac{f_{XY}}{f_Y}, f_{XY} = f_X \times f_Y$ 當 XY獨立。
- $E(f_X(x)) = \int_0^1 x \int_0^x f_{xy}(x,y) dy dx$ ,  $0 \le y \le x \le 1$

· Var (Y+Y) = E[(A+Y - E(X,Y))] = E[(X+Y - MX-MY)] (ov = E[(X-MX)+V-M)]] = E[(X+Y - MX-MY)] (ov = Var(X) + Var(Y) + 2 (ov (X,Y)) 又獨立 (ov = 0. 女 X, Y, 獨立 愛 Var(X+Y) = Var(X)+ Var(Y)	するよけ
(B) - *4· PXY1B(XY) = PXY(XY) (X1/16B	E(h(x))   Y=y] = Son har)farr (x)y) dx E[h(x))   X=x]= Son h(x)) frankly) dy  图直 X Y .
(は日本子 (文) (ス) - P(B) , (ハ) (D) (D) (D) (D) (D) (D) (D) (D) (D) (D	$P_{X,Y}(X,y) = P_{X}(X) \cdot P_{Y}(y) \qquad E_{X}(\mathcal{G}(X) \cdot h(Y))$ $P_{X,Y}(X,y) = f_{X}(X) \cdot P_{Y}(y) \qquad = E_{X}(\mathcal{G}(X)) \cdot E_{Y}(\mathcal{F}(X))$ $P_{X,Y}(X,y) = F_{X}(X) \cdot P_{Y}(y) \qquad \forall \chi : Port 4.C \( \frac{1}{2} \) \( \frac{1}{2} \) \( \frac{1}{2} \)$
E(hixiy) 1/=y) = \$\frac{\x}{\x} \frac{\x}{\x} \frac{\x}{\x	

#### 相關 Cov

- Correlation coefficient  $\rho = \frac{Cov(X,Y)}{\sigma_x \sigma_y}$
- Covariance  $Cov(X,Y) = E[E(X \mu_x)(Y \mu_y)] = E(XY) E(X)E(Y)$
- 標準差 :  $\sigma_x = \sqrt{Var(x)}$
- $\bullet \quad Var(x) = E(x^2) E(x)^2$
- $E(x) = \sum_{-\infty}^{\infty} xP(x) \text{ or } \int_{-\infty}^{\infty} xf(x)dx$
- $E(x^n) = \sum_{-\infty}^{\infty} x^n P(x) \text{ or } \int_{-\infty}^{\infty} x^n f(x) dx$

#### **Bivariate Normal Distribution**

$$f(x) = \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}}e^{\frac{-1}{2}\left[\frac{1}{1-\rho^2}(\frac{(x-\mu_x)^2}{\sigma_x^2}-2\rho\frac{x-\mu_xy-\mu_y}{\sigma_x}+\frac{(y-\mu_y)^2}{\sigma_y^2})\right]}$$

## MGF:

MGF:

離散 :  $M_X(t) = E(e^{tX}) = \sum_x e^{tX} P_X(x)$ 

連續:  $M_X(t) = E(e^{tX}) = \int_{Y} e^{tX} f_X(x) dx$ 

分布	mgf
Poisson( $\lambda$ )	$M_X(t)e^{-\lambda(1-e^{-t})}$
白努力(p)	$M_X(t) = 1 - p + pe^t$
二項式(n,p)	$M_X(t) = (1 - p + pe^t)^n$
Exponential( $\lambda$ )	$M_X(t) = \frac{\lambda}{\lambda - t}$
高斯 G(0,1)	$M_X(t) = e^{rac{t^2}{2}}$
高斯 $G(\mu, \sigma^2)$	$M_X(t) = e^{\mu t + \frac{\sigma^2 t^2}{2}}$

定理:

$$E(x^n) = M_X(0)^n$$

$$Z = X_1 + \dots + X_n, M_Z(t) = M_{X_1}(t) \times \dots \times M_{X_n}(t)$$

若高斯分布的隨機變數Y=3X-7,可得 $\sigma=3$ ,  $\mu=7(X=\frac{Y+7}{3})$ 

# 中央極限定理:

 $X_1, X_2, ..., X_n$ ,若獨立,且有相同機率分布,當 n 接近 $\infty$ ,

$$X = X_1 + X_2 + \dots + X_n \sim N(\mu_{X_1 + X_2 + \dots + X_n}, \sigma_{X_1 + X_2 + \dots + X_n}^2) = G(n\mu_{X_1}, n\sigma_{X_1}^2)$$

節例:

5.(10%) A fair die is rolled 200 times. What is the approximate probability that the sum of the outcomes is between (00) and (723) (7+b) < Hint> Find the mean and variance of rolling a die. Then, apply the central limit theorem.

<P.S.> You can write down your answer as a numerical value. Or, if you do not have the relevant table for carrying out the numerical computation, you can express your answer in terms of  $\Phi(\cdot)$ : the cumulative distribution function of the standard Gaussian random variable.

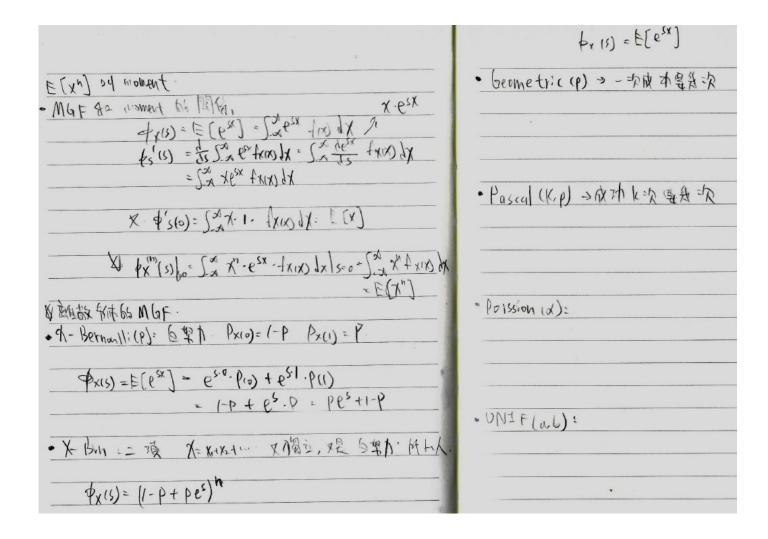
投散子
$$E(X) = \frac{1+\cdots+6}{6} = 3.5, E(X^2) = \frac{1+2^2+\cdots+6^2}{6} = 15.17, Var(x) = E(X^2) - E(X)^2 = 2.92$$

投很多次變常態分布,又 $Var(x) = \sigma^2 = 2.92, \mu = 3.5$ ,大數法則:

$$G(n\mu_{X_1}, n\sigma_{X_1}^2) = G(200 \times 3.5, 200 \times 2.92) = G(700, 584)$$

且找P(700 < X < 725) = P(X = 725) - P(X = 700)又高斯分布 cdf  $F_X(X) = \phi(\frac{x-\mu}{\sigma})$ 所以可得:

$$P(700 < X < 725) = \phi\left(\frac{725 - 700}{\sqrt{584}}\right) - \phi\left(\frac{700 - 700}{\sqrt{584}}\right) = \phi(1.03) - \phi(0)$$



■ In the special case:  $(X, Y) \sim \text{bi-Gauss}(\mu_1, \mu_2, \sigma_1^2, \sigma_2^2, \rho)$ , the conditional mean turns out to be in a linear form:

$$\mu_2 + \rho \cdot \frac{\sigma_2}{\sigma_1} \cdot (x - \mu_1) \quad (\#17)$$

**Bivariate Normal Distribution** 

$$f(x) = \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}}e^{\frac{-1}{2}\left[\frac{1}{1-\rho^2}\left(\frac{(x-\mu_x)^2}{\sigma_x^2}-2\rho\frac{x-\mu_xy-\mu_y}{\sigma_x}+\frac{(y-\mu_y)^2}{\sigma_y^2}\right)\right]}$$

Based on an observation on X, we try to estimate Y (i.e. produce a value that Y is likely to assume). The estimator is expressed as Y = g(X), where  $g(\cdot)$  is some function.

According to the criterion of minimum mean squared error (MMSE), the goal is to minimize

$$E((Y - g(X))^{2} | X = x)$$
 (#15)

The solution to the minimization of Eq.(#15) is:

$$g(X) = E(Y|X = x)$$
 (#16)  $(x = \frac{1}{2})^{\frac{1}{2}} = \frac{1}{2}$   
# Prf: LHS of (#15) =  $\int_{-\infty}^{\infty} (y - g(x))^2 \cdot f_{Y|X}(y|x) dy$