ST540 HW2

Q2

(a) We first validate that $f(x)=rac{1}{b-a}\geq 0, orall x\in [a,b].$ It is obvious since b>a. Then we need to show that $\int_{[a,b]}f(x)dx=1$:

$$\int_{[a,b]} f(x) dx = \int_{[a,b]} rac{1}{b-a} dx = rac{1}{b-a} x|_a^b = 1$$

(b)

$$E(X) = \int_{[a,b]} x f(x) dx = \int_{[a,b]} rac{x}{b-a} dx = rac{1}{b-a} rac{1}{2} x^2 |_a^b = rac{a+b}{2}$$
 $E(X^2) = \int_{[a,b]} x^2 f(x) dx = \int_{[a,b]} rac{x^2}{b-a} dx = rac{1}{b-a} rac{1}{3} x^3 |_a^b = rac{b^2 + ab + a^2}{3}$
 $Var(X) = E(X^2) - [E(X)]^2 = rac{(b-a)^2}{12}$

Q3

 $Gamma(\alpha, \beta)$ is the easiest one that satisfies all of the constraints. It has positive support and the following properties:

Suppose $X \sim Gamma(\alpha, \beta)$, then

$$E(X) = \alpha \beta$$

$$Var(X) = lpha eta^2$$

Let E(X)=5 and Var(X)=3 and we have lpha=25/3 and eta=3/5.

Some other distributions with positive support also works like lognormal distribution but can be more complicated.

Q4

(a)

$$p(X_1=i) = \sum_{j=0}^1 p(X_1=i, X_2=j), orall i=0,1,2$$

Thus, the marginal distribution of X_1 is:

X_1	0	1	2	
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(b)

$$p(X_2=j) = \sum_{i=0}^2 p(X_1=i, X_2=j), orall j=0, 1$$

Thus, the marginal distribution of X_2 is:

X_2	0	1	
p	0.45	0.55	

(c)

$$p(X_1=i|X_2=j)=rac{p(X_1=i,X_2=j)}{p(X_2=j)}$$

Thus, the conditional distribution of $X_1 \mid X_2$ is:

$X_1 \setminus X_2$	0	1
0	1/3	3/11
1	1/3	4/11
2	1/3	4/11

(d)

$$p(X_2=j|X_2=i)=rac{p(X_1=i,X_2=j)}{p(X_1=i)}$$

Thus, the conditional distribution of $X_1 ig| X_2$ is:

$X_2 \setminus X_1$	0	1	2
0	1/2	3/7	3/7
1	1/2	4/7	4/7

(d) X_1 and X_2 are not independent.

$$p(X_1=0,X_2=0)=0.15
eq p(X_1=0)p(X_2=0)=0.3 imes 0.45$$

Q4

We first work out the kernel of the posterior of n.

$$egin{aligned} P(n=k|Y) & \propto P(Y|n=k)P(n=k) \ & = inom{k}{Y} heta^Y (1- heta)^{k-Y} \cdot rac{5^k e^{-5}}{k!} \ & \propto rac{(1- heta)^{k-Y} 5^k}{(k-Y)!} \ & \propto rac{(5-5 heta)^{k-Y}}{(k-Y)!} \end{aligned}$$

where k is an **integer** and $k \geq Y$.

Recall that the kernel of $X \sim Poisson(\lambda)$ is:

$$P(X=k) = rac{\lambda^k e^{-\lambda}}{k!} \propto rac{\lambda^k}{k!}$$

where $k \geq 0$.

Comparing the kernels of P(n=k|Y) and $Poisson(\lambda)$, we conclude that P(n=k|Y) follows a translated Poisson distribution with coefficient $\lambda=5-5\theta$. Thus, we have:

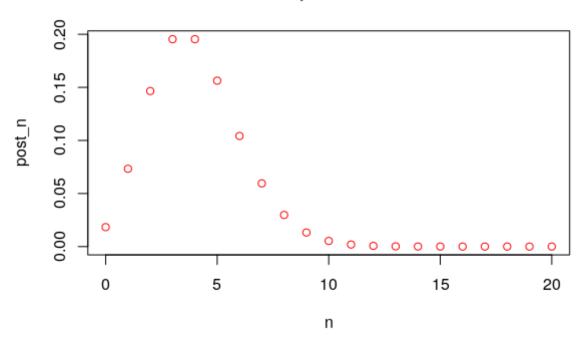
$$P(n = k|Y) = P(X = k - Y)$$

where $X \sim Poisson(5-5\theta)$.

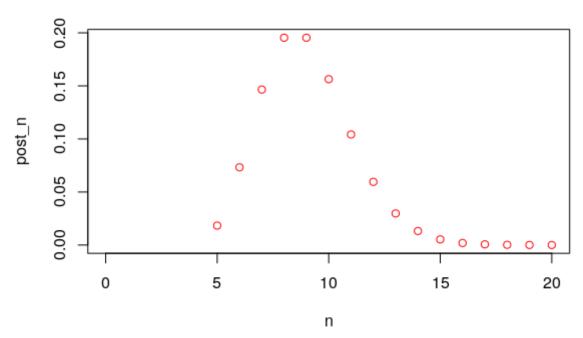
The following code is to plot the posterior of P(n=k|Y) when Y=0 and $\theta=0.2$. To plot with other values of n and theta, just change the first two lines of code.

```
1 Y = 0
2 theta = 0.2
3 # n is an integer and n >= Y
4 # we will calculate P(n|Y) at n = Y, Y+1, ..., 20
5 n = seq(Y, 20, 1)
6 post_n = dpois(n-Y, 5-5*theta)
7 # plot
8 plot(n, post_n, col = 'red', xlim = range(0:20), main = 'Y=0, theta=0.2')
```

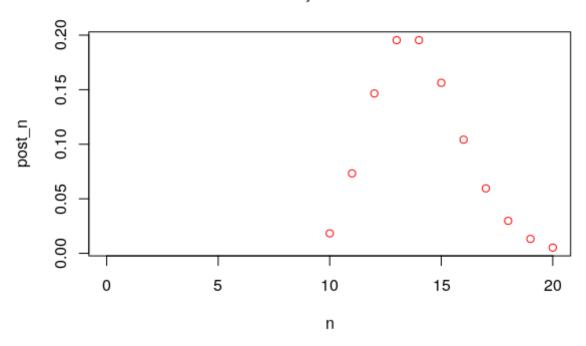
Y=0, theta=0.2



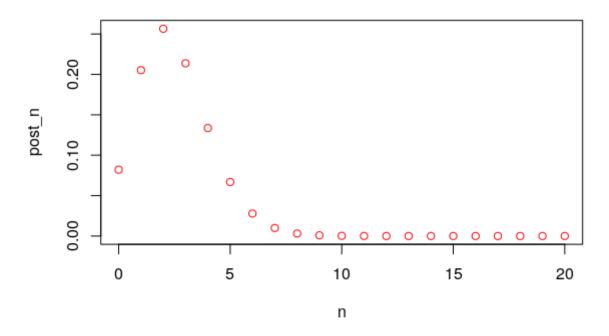




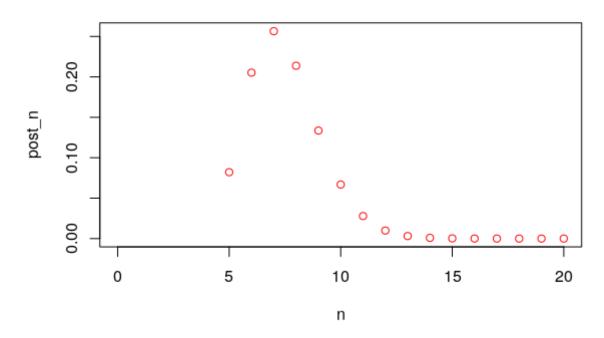
Y=10, theta=0.2



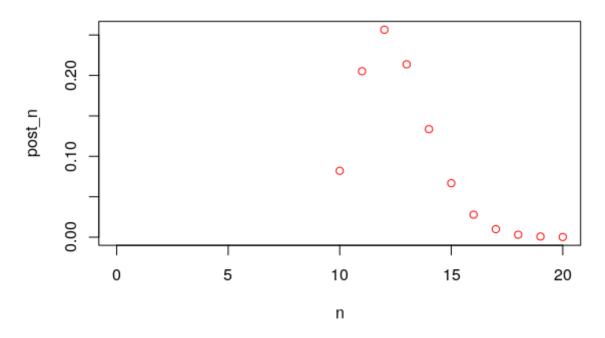
Y=0, theta=0.5



Y=5, theta=0.5



Y=10, theta=0.5



comment on $\boldsymbol{\theta}$ and \boldsymbol{Y}

- 1. Given θ , the posterior just moves to the right as Y increases with the shape stays the same the distribution to the right.
- 2. Given Y, the posterior gets more concentrated around the mode as θ increases. distribution.

Method 2: Approximate posterior

Instead of working on the kernel of the posterior, we can also approximate the posterior:

$$P(n=k|Y) = rac{P(Y|n=k)P(n=k)}{\sum_{k_0=Y}^{\infty} P(Y|n=k_0)P(n=k_0)}$$

Note that $n \sim Poisson(5)$ so that $P(n=k_0)$ becomes very small if k_0 is large. So we have:

$$\sum_{k_0=Y}^{\infty} P(Y|n=k_0) P(n=k_0) pprox \sum_{k_0=Y}^{K} P(Y|n=k_0) P(n=k_0)$$

where K is a large integer (e.g. 100).

```
1 Y = 0
2 theta = 0.2
3
4 # approximate the denominator. Take K = 100
5 n_grid = seq(Y, 100, 1)
6 like = dbinom(Y, n_grid, theta)
7 prior = dpois(n_grid, 5)
8 denominator = sum(like*prior)
9
10 # posterior of n at n=Y,Y+1,...,100
11 post = like*prior/denominator
12 plot(n_grid, post, col = 'red', xlim = range(0:20), main = 'Y=0, theta=0.2'
```