

## Module 9.3: Time Series Analysis

### Fall Term 2023

**Week 4:**

Unit Roots; Regressions between Time Series

# Outline in Weeks

- ➊ Introduction; Descriptive Modelling
- ➋ Returns; Autocorrelation; Stationarity
- ➌ ARMA Models
- ➍ Unit Roots; Regressions between Time Series
- ➎ Volatility Modelling
- ➏ Value at Risk
- ➐ Cointegration

# Outline

## 1 Unit Root Testing

## 2 ARIMA Models

## 3 Regressions with Time Series

## 4 Epilogue

# Unit Root Testing

- Recall that if a nonstationary series,  $y_t$  must be differenced  $d$  times before it becomes stationary, then it is said to be *integrated* of order  $d$ . Notation:  $I(d)$ . When we simply say 'integrated' we mean  $I(1)$ . Another synonym is 'the series has a *unit root*'.
- So far, our decision to take differences was based on the correlogram: if autocorrelations decay slowly and approximately linearly, then the series may be integrated and must be differenced.
- This procedure is subjective and unreliable: a *trend-stationary* series, such as  $Y_t = \beta_0 + \beta_1 t + U_t$ , will also have large and slowly decaying autocorrelations (see `simulation.xlsx`).
- Therefore, it is useful to have a formal testing procedure to distinguish integrated from (trend-) stationary time series. Such tests are called *unit root tests*. The (*Augmented*) *Dickey-Fuller* test is the most common.

## Example: the AR(1) Model

- The simplest example of stationary and integrated time series is the AR(1) model

$$Y_t = \phi_1 Y_{t-1} + U_t,$$

where  $Y_0$  is a constant.

- If  $-1 < \phi_1 < 1$ , then  $Y_t$  is stationary, with mean 0 and variance  $\sigma^2/(1 - \phi^2)$ .
- If  $\phi_1 = 1$ , then the model becomes a *random walk*,  $Y_t = Y_{t-1} + U_t$ , with mean  $Y_0$  and variance  $\sigma^2 t$ .

# Stochastic vs. Deterministic Trends

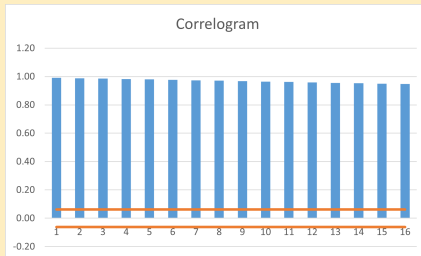
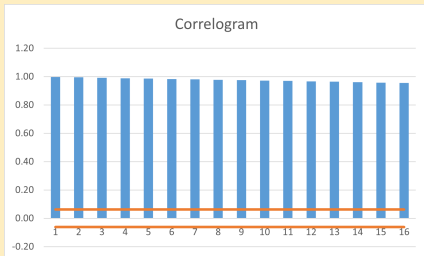
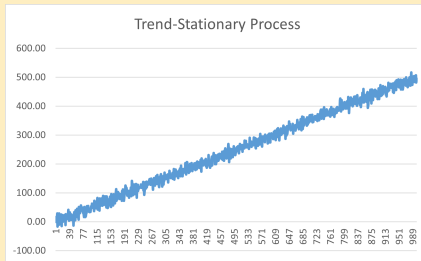
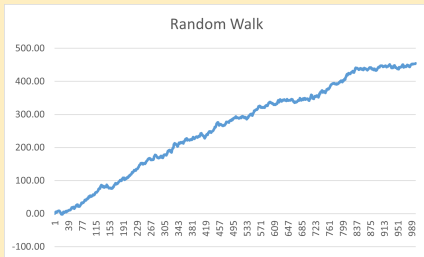
- Consider the two models

$$Y_{1,t} = \delta t + U_{1,t} \quad \text{and}$$

$$Y_{2,t} = \delta + Y_{2,t-1} + U_{2,t}.$$

- For both models,  $\mathbb{E}[\Delta Y_t] = \delta$ , i.e., both series trend upwards by  $\delta$  each period.
- $Y_{1,t}$  contains a **deterministic trend**:  $Y_{1,t} - \delta t = U_{1,t}$  is stationary. In practice, regress  $y_t$  on a linear trend like in week 1; the residuals will be stationary.
- $Y_{2,t}$  contains a **stochastic trend**:  $Y_{2,t} - \delta t$  is a random walk. It becomes stationary only by differencing, i.e., it is  $I(1)$ .

## Random walk (left) vs. trend stationary process (right).



# Why do we care?

- ❶ Inference: OLS estimators have non-standard distribution, so that standard inference is invalid.
- ❷ Forecasting:
  - stationary series display *mean-reversion*; deviations from the mean are corrected in the next period. Shocks  $U_t$  have a *transitory*, decreasing effect on future  $Y_{t+k}$  (in AR(1) model, the effect is  $\phi^k \rightarrow 0$ );
  - For integrated time series (of order 1): no mean-reversion, shocks  $U_t$  have a *persistent* effect on future  $Y_{t+k}$ .
- ❸ Spurious regression: regressing two drifting I(1) variables onto each other will spuriously result in significant estimates, because they happen to trend in the same direction.



# The Dickey-Fuller Test

- Consider again the AR(1) model

$$Y_t = \phi_1 Y_{t-1} + U_t.$$

- We wish to test the null hypothesis  $H_0 : \phi_1 = 1$ , against the *one-sided* alternative hypothesis  $H_1 : \phi_1 < 1$ .
- Under the null, the process is a random walk, and hence integrated. Under the alternative, it is stationary (if  $\phi_1 > -1$ , which we assume). Therefore, we will be testing  $H_0 : Y_t \sim I(1)$  against  $H_1 : Y_t \sim I(0)$ .
- The *Dickey-Fuller test* is based on the  $t$ -statistic for  $\phi_1 = 1$  in the AR(1) model, which may be reformulated as the  $t$ -statistic for  $\psi = 0$  in

$$\Delta y_t = \psi y_{t-1} + u_t,$$

where  $\psi := \phi_1 - 1$ ; note that  $\psi < 0$  under the alternative.

# The Dickey-Fuller Distribution

- We reject  $H_0$  if the  $t$ -statistic is less than the (negative) critical value.
- In classical regressions, the 5% critical value for a one-sided  $t$ -test would be  $-1.645$ . However, because the regressor  $Y_{t-1}$  is non-stationary under the null, a different distribution arises, and the appropriate 5% critical value is  $= -1.95$ .
- This critical value changes to  $= -2.86$  if we add a constant term to the regression:

$$\Delta y_t = \alpha + \psi y_{t-1} + u_t.$$

This is the relevant test if we want to allow for a non-zero mean  $\mathbb{E}[Y_t]$  under the alternative.

- If we want to test a random walk *with drift* against a *trend-stationary* alternative, then the relevant regression is

$$\Delta y_t = \alpha + \delta t + \psi y_{t-1} + u_t,$$

and the 5% critical value is  $= -3.41$ .

# Choice of Model

- The difference between the three tests is whether or not a constant and a time trend are included in the test regression.
- In practice, the test without constant and trend is almost never applicable.
- The test with an intercept in the regression test is relevant for series such as interest rates and real exchange rates, where we do not expect a linear trend under either the null or the alternative hypothesis.
- Many other economic and financial time series, such as GDP or (log-) asset prices, clearly display an upward trend, in which case a trend should be included.

# The Augmented Dickey-Fuller Test

- The *augmented* Dickey-Fuller (ADF) test is an extension of the procedure to the  $AR(p)$  model

$$Y_t = \phi_1 Y_{t-1} + \dots + \phi_p Y_{t-p} + U_t.$$

- This process is integrated under  $H_0 : \phi_1 + \dots + \phi_p = 1$ , and stationary under the alternative hypothesis  $H_1 : \phi_1 + \dots + \phi_p < 1$ .
- It can be shown (see exercises) that this is equivalent to testing  $H_0 : \psi = 0$  against  $H_1 : \psi < 0$  in the regression

$$\Delta y_t = \psi y_{t-1} + \alpha_1 \Delta y_{t-1} + \dots + \alpha_{p^*} \Delta y_{t-p^*} + u_t,$$

where  $\psi = \sum_{i=1}^p \phi_i - 1$ , and  $p^* = p - 1$ .

- The interpretation of the null and alternative hypothesis, the role of the constant and trend, and the critical values are the same as in the first-order model.

# Choosing $p$

- In practice,  $p$  is, of course, *unknown*.
- The number of lags in the test regression must be chosen large enough, such that the residuals have no autocorrelation, but not too large, because this would decrease test power.
- Often the choice is made based on *model selection criteria* (AIC, BIC). EViews has a built-in option to do this automatically.
- The AIC is preferred in this case, because it tends to include more lags than the BIC. This is preferred because our goal in this case is not to find the “correct” model, but to combat autocorrelation in the test regression.
- So effectively, the ADF test is just the DF test (regress  $\Delta y_t$  on  $y_{t-1}$ ), but with enough lags of  $\Delta Y_t$  thrown in to remove any autocorrelation.

# Outline

- 1 Unit Root Testing
- 2 ARIMA Models**
- 3 Regressions with Time Series
- 4 Epilogue

# ARIMA Models

- As discussed last week, the first step in the Box-Jenkins procedure is to make sure that the data are stationary.
- This is usually decided by the ADF test.
- If the test doesn't reject, then one models the first difference  $\Delta Y_t$  as an ARMA process.
- The model for the levels  $Y_t$  is then called an *ARIMA(p,d,q)* model: the data are differenced  $d$  times, and the result modeled as an ARMA(p,q) process. Usually,  $d = 1$ .
- To forecast an ARIMA process (with  $d=1$ ), first predict  $\Delta Y_{t+1}$ , and then let
$$\hat{Y}_{t+1} = Y_t + \widehat{\Delta Y_{t+1}}.$$
- Longer horizon forecast are obtained recursively as usual.

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# Regression with Time Series

Some considerations when analyzing the relation between the time series  $\{y_t\}$  and  $\{x_t\}$  in the regression model

$$y_t = \alpha + \beta' x_t + u_t :$$

- Model is only reasonable if  $Y_t$  and  $X_t$  are integrated of same order. If both are  $I(1)$ , then in general the error will be  $I(1)$ , except in case of *cointegration* (stationary linear combination of  $I(1)$  series).
- In case of non-cointegrated  $I(1)$  series, consider regression in first differences:  
 $\Delta y_t = \alpha + \beta \Delta x_t + e_t$ .
- Even when non-stationary is not a concern, autocorrelation usually is (most economic/financial series are autocorrelated).
- We will discuss two tests for autocorrelation in a regression: the *Breusch-Godfrey* test and the *Durbin-Watson* test.

# Breusch-Godfrey Test

- The *Breusch-Godfrey LM test* is done in four steps:

- 1 Estimate the regression, e.g.,

$$y_t = \beta_1 + \beta_2 x_{2t} + \beta_3 x_{3t} + u_t,$$

by OLS, and get the residuals  $\hat{u}_t$ .

- 2 Estimate the auxiliary regression

$$\hat{u}_t = \gamma_1 + \gamma_2 x_{2t} + \gamma_3 x_{3t} + \tau_1 \hat{u}_{t-1} + \dots + \tau_r \hat{u}_{t-r} + v_t.$$

Denote by  $R_{aux}^2$  the  $R^2$  of this regression.

- 3 Test the null hypothesis of *no residual autocorrelation*

$$H_0 : \tau_1 = \dots = \tau_r = 0,$$

by the usual F-test, or by the LM test:

$$T \cdot R_{aux}^2 \sim \chi^2(r)$$

- 4 The test rejects for large values of  $T \cdot R_{aux}^2$ .

# Remarks

- How do we choose  $r$ ? Consult the correlogram, and take account of the frequency of the data, e.g.,  $r = 4$  or 5 for quarterly data.
- This is a general autocorrelation test. It has power against any degree of possible autocorrelation.
- EViews: after estimation, click View → Residual Diagnostics → Serial Correlation LM test

# The Durbin-Watson Test

- The *Durbin Watson* test is a test of first-order autocorrelation only.
- This test is performed as follows:
  - 1 Get the OLS residuals  $\hat{u}_t$  from the regression as before.
  - 2 Compute the statistic

$$DW = \frac{\sum_{t=2}^n (\hat{u}_t - \hat{u}_{t-1})^2}{\sum_{t=2}^n \hat{u}_t^2} \approx 2(1 - \hat{\tau}).$$

where  $\hat{\tau}$  is OLS estimator of regression of  $\hat{u}_t$  on  $\hat{u}_{t-1}$ . Since  $-1 \leq \hat{\tau} \leq 1$ , it follows that  $0 \leq DW \leq 4$ . If there is no autocorrelation, we expect  $DW \approx 2$ .

- 3 Look up the critical values for the test in a table. The test has a non-standard distribution. There are five different regions that lead to different conclusions (see next slide).
- The DW test is included in EViews estimation output by default.

# Test Decision

- The critical values depend on the regressors. However, one can obtain bounds on the critical values.
- The table gives the DW lower ( $d_l$ ) and upper ( $d_u$ ) bounds, at the 5% level, for different values of  $n$  and  $k$ . Test decision:

$0$	$<$	$DW$	$<$	$d_l$	$\Rightarrow$	reject (positive AC)
$d_l$	$<$	$DW$	$<$	$d_u$	$\Rightarrow$	inconclusive
$d_u$	$<$	$DW$	$<$	$4 - d_u$	$\Rightarrow$	do not reject
$4 - d_u$	$<$	$DW$	$<$	$4 - d_l$	$\Rightarrow$	inconclusive
$4 - d_l$	$<$	$DW$	$<$	$4$	$\Rightarrow$	reject (negative AC)

## Critical Values of the DW Test

	$k = 2$		$k = 3$		$k = 4$		$k = 6$		$k = 10$	
$n$	$d_L$	$d_U$	$d_L$	$d_U$	$d_L$	$d_U$	$d_L$	$d_U$	$d_L$	$d_U$
15	1.08	1.36	0.95	1.54	0.81	1.75	0.56	2.22	0.17	3.22
16	1.11	1.37	0.98	1.54	0.86	1.73	0.61	2.16	0.22	3.09
17	1.13	1.38	1.01	1.54	0.90	1.71	0.66	2.10	0.27	2.97
18	1.16	1.39	1.05	1.53	0.93	1.70	0.71	2.06	0.32	2.87
19	1.18	1.40	1.07	1.54	0.97	1.68	0.75	2.02	0.37	2.78
20	1.20	1.41	1.10	1.54	1.00	1.68	0.79	1.99	0.42	2.70
22	1.24	1.43	1.15	1.54	1.05	1.66	0.86	1.94	0.50	2.57
24	1.27	1.45	1.19	1.55	1.10	1.66	0.92	1.90	0.58	2.46
26	1.30	1.46	1.22	1.55	1.14	1.65	0.98	1.87	0.66	2.38
28	1.33	1.48	1.25	1.56	1.18	1.65	1.03	1.85	0.72	2.31
30	1.35	1.49	1.28	1.57	1.21	1.65	1.07	1.83	0.78	2.25
35	1.40	1.52	1.34	1.58	1.28	1.65	1.16	1.80	0.91	2.14
40	1.44	1.54	1.39	1.60	1.34	1.66	1.23	1.79	1.01	2.07
50	1.50	1.58	1.46	1.63	1.42	1.67	1.33	1.77	1.16	1.99
75	1.58	1.65	1.57	1.68	1.54	1.71	1.45	1.77	1.37	1.90
100	1.65	1.69	1.63	1.71	1.60	1.74	1.57	1.78	1.48	1.87
200	1.76	1.78	1.75	1.79	1.74	1.80	1.72	1.82	1.67	1.86

$\alpha = 5\%$ ,  $k =$  number of regressors (including intercept).

# Consequences of Autocorrelation

- Only in a truly static model, the OLS estimator remains unbiased.
- In general,  $\hat{\beta}$  will be biased.
- OLS is no longer BLUE even when it is unbiased. More efficient estimators are available (GLS).
- The OLS standard errors are wrong.
- Inference based on usual  $t$ - and  $F$ -tests is unreliable.

# Solutions

- One should at least use Newey-West *heteroskedasticity and autocorrelation consistent* (HAC) standard errors (in Eviews: Click on the Options tab when entering the regression equation)<sup>1</sup>.
- The more modern approach is to model the autocorrelation explicitly, by including lags of  $y_t$  ( $\rightarrow$  *ARX* model) and possibly  $x_t$  ( $\rightarrow$  *ADL*<sup>2</sup> model). MA terms can also be added ( $\rightarrow$  *ARMAX* model).

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<sup>1</sup>EViews also provides *White standard errors*. These are consistent under heteroskedasticity only, not under autocorrelation.

<sup>2</sup>autoregressive distributed lag



# Example

- Consider the ADL model

$$Y_t = \alpha_0 + \alpha_1 Y_{t-1} + \beta_0 X_t + \beta_1 X_{t-1} + u_t. \quad (1)$$

- The *long-run equilibrium* is the *steady state* of the system  $(Y_t, X_t)$  when all future shocks  $U_{t+1}, U_{t+2}, \dots$  are set to zero. If the system is stationary, then

$$Y_t \rightarrow Y^* \text{ and } X_t \rightarrow X^*.$$

So the long-run equilibrium can be found from the equation

$$\begin{aligned} Y^* &= \alpha_0 + \alpha_1 Y^* + \beta_0 X^* + \beta_1 X^* \Rightarrow \\ Y^* &= \frac{\alpha_0}{1 - \alpha_1} + \frac{\beta_0 + \beta_1}{1 - \alpha_1} X^*. \end{aligned} \quad (2)$$

# Why is a dynamic model needed?

A dynamic model can be necessary for the following reasons:

- Inertia of the dependent variable. E.g., adjustment costs, habits in individual decisions, or market frictions.
- Trending behaviour. E.g., if returns are unpredictable, prices must be trending. Hence, prices are autocorrelated.
- Seasonality; see Lecture 1.
- Over-reactions. Typically the result of some market failure, but think of exchange rate overshooting.

# Residual Autocorrelation

- Now assume that instead of the dynamic model (1), the steady state equation (2) is estimated:

$$y_t = a_0 + b_0 x_t + u_t$$

- This is known as *dynamic misspecification*. The effect of the lagged variables is then in the error term. Result: autocorrelated residuals.
- In this case, not only the standard errors will be wrong, but the estimates will be biased:  $\hat{b}_0$  will be an estimate of  $\frac{\beta_0 + \beta_1}{1 - \alpha_1}$  in (2), not of  $\beta_0$  in (1).
- The modern view is that autocorrelation is not a problem, but an *opportunity to improve the model*.
- Thus, instead of correcting the standard errors, one should estimate a dynamic model directly.
- Start with a dynamic model (include lags of  $Y_t$  and  $X_t$  in the model). Test for autocorrelation, and for the significance of dynamics.

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# Learning Goals

## Students

- understand the ADF test and can use it to test for unit roots,
- are able to forecast ARIMA models,
- understand the difference between static and dynamic models, and
- know how to test for autocorrelation in a regression

# Homework

- Exercise 4
- Questions 2 and 3 from Chapter 8 of Brooks (2019)