

## Module 9.3: Time Series Analysis with Python

### Fall Term 2023

**Week 7:**

Cointegration

# Outline in Weeks

- ➊ Introduction; Descriptive Modeling
- ➋ Returns; Autocorrelation; Stationarity
- ➌ ARMA Models
- ➍ Unit Roots; ARIMA Models
- ➎ Volatility Modeling
- ➏ Value at Risk
- ➐ Cointegration

# Outline

- 1 Cointegration and Common Trends
- 2 Error Correction Models and the Engle-Granger Procedure
- 3 Epilogue

# Cointegration and Common Trends

Suppose we have two time series  $Y_t$  and  $X_t$ , which are both  $I(1)$ , and we analyze a regression model of the form

$$Y_t = \beta_1 + \beta_2 X_t + U_t.$$

Here  $U_t$  has mean zero but may display autocorrelation. Two cases:

- $U_t \sim I(1)$ : if  $U_t$  displays no mean-reversion, then  $Y_t$  does not revert to the explained part  $\beta_1 + \beta_2 X_t$ . Even if  $\beta_2 = 0$ , its  $t$ -statistic and  $R^2$  will often seem significant (*spurious regressions*). To avoid this, one should estimate a model in differences, i.e.,  $\Delta Y_t = a_1 + a_2 \Delta X_t + \Delta U_t$ .
- $U_t \sim I(0)$ : now  $Y_t$  and  $X_t$  have a *common* stochastic trend, such that the linear combination  $Y_t - \beta_2 X_t$  does not have a trend. This is called *cointegration*.

# Example

- Consider the model

$$Y_t = \beta_1 + \beta_2 X_t + U_{1,t}$$

$$X_t = X_{t-1} + U_{2,t}$$

where  $\beta_2 \neq 0$ ,  $U_{1,t}, U_{2,t} \stackrel{\text{iid}}{\sim} (0, \sigma^2)$  independently of each other.

- $X_t$  is a random walk and thus nonstationary.  $Y_t$  contains  $X_t$  and is thus also nonstationary. But

$$Y_t - \beta_2 X_t = \beta_1 + U_{1,t}$$

is stationary: the RHS is white noise plus a constant.

- $(1, -\beta_2)$  is called the *cointegrating vector*.

# Cointegration and Common Trends, contd.

- The concept is easily extended to more than two series: if  $X_{2t}, \dots, X_{kt}$  are all  $I(1)$  variables, and

$$Y_t = \beta_1 + \beta_2 X_{2t} + \dots + \beta_k X_{kt} + U_t,$$

then this is a spurious regression if  $U_t \sim I(1)$  (and  $\beta_i = 0$ ), and a cointegrating relation if  $U_t$  is stationary.

- In other words, cointegration between  $k$  integrated series means that there exists a *linear combination*<sup>1</sup> of them which is stationary.
- Examples of possibly cointegrated time series:
  - exchange rates and relative prices (*purchasing power parity*);
  - spot and futures prices of assets or exchange rates;
  - short- and long-term interest rates (*term structure models*);
  - stock prices and dividends (*present value relations*).

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<sup>1</sup>i.e., a weighted sum

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# Testing for Cointegration

- For cointegrated series, one should exploit the long-run equilibrium relationship between variables for estimation rather than differencing. Differencing would *remove* that structure.
- Engle and Granger proposed the following procedure:
  - Conduct individual unit root tests to ensure all series are  $I(1)$ .
  - Estimate the regression model

$$Y_t = \beta_1 + \beta_2 X_{2t} + \dots + \beta_k X_{kt} + U_t$$

by ordinary least-squares. Estimates are (super)consistent, but standard errors are wrong because series are  $I(1)$ .

- Apply an ADF unit root test (with constant) to the residuals  $\hat{u}_t$  from this regression. This yields a test for  $H_0 : U_t \sim I(1)$  (spurious regression) against  $H_1 : U_t \sim I(0)$  (cointegration). The critical values depend on  $k$ . E.g., for  $k = 2$  ( $X_{2t}$  and an intercept), the 5% c.v. is -3.41.
- If  $H_0$  is rejected, estimate an *error correction model*.



## Engle-Granger critical values

Number of series (w/o constant)	2	3	4	5	6
Critical value	-3.41	-3.80	-4.16	-4.49	-4.74

# Error Correction Models

- Cointegration between  $Y_t$  and  $X_t$  implies that deviations ( $Y_{t-1} - \beta_1 - \beta_2 X_{t-1}$ ) from the equilibrium level should be (partially) corrected in the next period, by  $Y_t$ ,  $X_t$ , or both.
- This leads to a *vector error correction model* (VECM), which in the simplest form is

$$\begin{aligned}\Delta Y_t &= c_1 + \alpha_1(Y_{t-1} - \beta_1 - \beta_2 X_{t-1}) + e_{1t}, \\ \Delta X_t &= c_2 + \alpha_2(Y_{t-1} - \beta_1 - \beta_2 X_{t-1}) + e_{2t},\end{aligned}$$

where  $e_{1t}$  and  $e_{2t}$  are two white noise errors (possibly correlated), and where we expect  $\alpha_1 < 0$  and/or  $\alpha_2 \beta_2 > 0$ .

- We might need to add lags of  $\Delta Y_t$  and/or  $\Delta X_t$  on RHS to combat autocorrelation.
- The *Granger representation theorem* states that cointegration implies an error correction model (possibly with more lags), and vice versa; see exercises.

# Engle-Granger Procedure

- The VECM

$$\begin{aligned}\Delta Y_t &= c_1 + \alpha_1(Y_{t-1} - \beta_1 - \beta_2 X_{t-1}) + e_{1t}, \\ \Delta X_t &= c_2 + \alpha_2(Y_{t-1} - \beta_1 - \beta_2 X_{t-1}) + e_{2t}.\end{aligned}$$

is estimated by replacing  $Y_{t-1} - \beta_1 - \beta_2 X_{t-1}$  by OLS residual  $\hat{u}_{t-1} = Y_{t-1} - \hat{\beta}_1 - \hat{\beta}_2 X_{t-1}$ , and estimating  $\alpha_1$  and  $\alpha_2$  by OLS.

- Note that  $\hat{u}$  is stationary, so this is a valid regression!
- If  $\alpha_2 = 0$ , then all correction is done by  $Y_t$ , and not by  $X_t$ . In that case it makes sense to treat  $X_t$  as exogenous and  $Y_t$  as endogenous, and consider the “single-equation” error correction model

$$\Delta Y_t = c + \alpha_1(Y_{t-1} - \beta_1 - \beta_2 X_{t-1}) + e_t.$$

- In general, both  $Y_t$  and  $X_t$  are endogenous.

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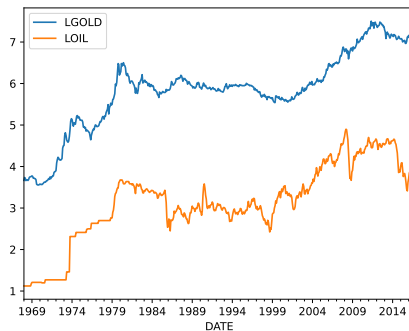
# Example

- Until 1971, as part of the Bretton-Woods system of fixed exchange rates, the US dollar was convertible to gold, i.e., it was possible for foreign central banks to redeem US dollars for gold at a fixed rate of 35\$ per troy ounce, so that the price of gold was fixed.
- In 1971, US president Nixon unilaterally cancelled the direct convertibility, ultimately ending the Bretton-Woods agreement.
- Gold became a floating asset, and its price increased sharply; in other words, the US\$ was massively devalued.

## Example continued

- We want to analyze the hypothesis that the increasing price (in US\$) of oil is not a consequence of an increased demand for (or a reduced supply of) oil, but rather of a continued devaluation of the US\$.
- We have at our disposal monthly data from April 1968 to January 2017 (586 observations) on the following variables:
  - $\text{gold}_t$ , the spot price of one troy ounce of gold in US\$;
  - $\text{oil}_t$ , the spot price of one barrel of WTI crude oil in US\$.
- Idea: if the relative price of oil expressed in units of gold  $\text{oil}_t/\text{gold}_t$  is stationary, then this implies that  $\log(\text{oil}_t) - \log(\text{gold}_t)$  is stationary, so that  $\log(\text{oil}_t)$  and  $\log(\text{gold}_t)$  must be cointegrated if the individual series are integrated.

## $\log(\text{gold}_t)$ and $\log(\text{oil}_t)$



# Example continued

## Step 1 Test that the variables are integrated (ADF test with constant and trend)

```
ADF, p, crits, res = tsa.stattools.adfuller(df["LOIL"], regression='ct', autolag='AIC', store=True)
print("ADF = ", ADF, "\np = ", p)
```

```
ADF = -2.5524224546607397
p = 0.3022670842559252
```

```
ADF, p, crits, res = tsa.stattools.adfuller(df["LGOLD"], regression='ct', autolag='AIC', store=True)
print("ADF = ", ADF, "\np = ", p)
```

```
ADF = -2.559158224027972
p = 0.29904292389348935
```

Neither test rejects, so the series are  $I(1)$ .



# Example continued

Step 2 Estimate long-run relationship  $loil_t = \beta_1 + \beta_2 lgold_t + U_t$

## Estimated long run relationship

```

=====
                        OLS Regression Results
=====
Dep. Variable:          LOIL      R-squared:                0.873
Model:                  OLS      Adj. R-squared:           0.873
Method:                 Least Squares  F-statistic:         4015.
Date:                   Wed, 11 Oct 2023  Prob (F-statistic):    6.99e-264
Time:                   15:03:24   Log-Likelihood:       -183.59
No. Observations:      586        AIC:                  371.2
Df Residuals:          584        BIC:                  379.9
Df Model:               1
Covariance Type:       nonrobust
=====

```

	coef	std err	t	P> t	[0.025	0.975]
Intercept	-2.2590	0.086	-26.167	0.000	-2.429	-2.089
LGOLD	0.9262	0.015	63.364	0.000	0.898	0.955

```

=====
Omnibus:                4.994   Durbin-Watson:           0.074
Prob(Omnibus):           0.082   Jarque-Bera (JB):       4.313
Skew:                    -0.133  Prob(JB):               0.116
Kurtosis:                 2.675   Cond. No.:              38.3
=====

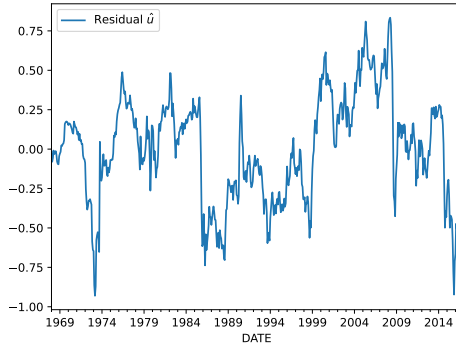
```

The cointegrating vector is  $(1, -\beta_2) = (1, -0.926)$  (if the Engle-Granger test rejects).

*Careful:* standard errors are wrong, because variables are  $I(1)$ .

# Example continued

Residuals (=equilibrium error)  $\hat{u}_t$



## Example continued

Step 3 Apply ADF test (with intercept) to  $u$ .

```
ADF, p, crits, res = tsa.stattools.adfuller(u, regression='c', autolag='AIC', store=True)
print("ADF = ", ADF)
```

```
ADF = -3.7247713389879844
```

*Careful*: we need to use the Engle-Granger 5% critical value of  $-3.41$ , **not** the one given by `adfuller` (nor the  $p$ -value). Conclusion: test rejects the null of “no cointegration”.

# Example continued

**Step 4a** Estimate the VECM equation for `loil`. Note: I threw in a lag of the dependent variable, as there was autocorrelation without it (cf. selected lag length in ADF test).

## VECM equation for `loil`

```

                        OLS Regression Results
=====
Dep. Variable:          df.LOIL.diff()          R-squared:          0.078
Model:                  OLS                    Adj. R-squared:     0.074
Method:                 Least Squares          F-statistic:        24.46
Date:                   Wed, 11 Oct 2023        Prob (F-statistic):  6.31e-11
Time:                   16:25:01               Log-Likelihood:     654.56
No. Observations:       584                   AIC:                -1303.
Df Residuals:           581                   BIC:                -1290.
Df Model:                2
Covariance Type:        nonrobust
=====
                        coef    std err          t      P>|t|      [0.025    0.975]
-----
Intercept                0.0036     0.003      1.111    0.267    -0.003     0.010
u.shift()               -0.0350     0.010    -3.518    0.000    -0.054    -0.015
df.LOIL.diff().shift()  0.2565     0.040     6.399    0.000     0.178     0.335
=====
Omnibus:                 325.732                Durbin-Watson:        2.007
Prob(Omnibus):            0.000                Jarque-Bera (JB):     11453.205
Skew:                     1.828                Prob(JB):              0.00
Kurtosis:                 24.385                Cond. No.              12.3
=====
```

# Example continued

Step 4b Estimate the VECM equation for `lgold`.

## VECM equation for `lgold`

```

=====
                        OLS Regression Results
=====
Dep. Variable:          df.LGOLD.diff()      R-squared:                0.003
Model:                  OLS                  Adj. R-squared:           0.001
Method:                 Least Squares        F-statistic:              1.616
Date:                   Wed, 11 Oct 2023      Prob (F-statistic):       0.204
Time:                   16:29:45              Log-Likelihood:           850.41
No. Observations:       585                  AIC:                     -1697.
Df Residuals:           583                  BIC:                     -1688.
Df Model:                1
Covariance Type:        nonrobust
=====

```

	coef	std err	t	P> t	[0.025	0.975]
Intercept	0.0059	0.002	2.504	0.013	0.001	0.010
u.shift()	0.0090	0.007	1.271	0.204	-0.005	0.023

```

=====
Omnibus:                69.639      Durbin-Watson:            1.933
Prob(Omnibus):           0.000      Jarque-Bera (JB):         356.493
Skew:                    0.368      Prob(JB):                 3.88e-78
Kurtosis:                6.753      Cond. No.                  3.02
=====

```

The adjustment coefficient  $\alpha_2$  in the equation for `d(lgold)` is insignificant ( $p = 0.204$ ). So all the adjustment is done by `loil` → single-equation ECM.

## Example continued

- The final model is the single-equation ECM

$$\Delta \text{loil}_t = 0.0036 - 0.035(\text{loil}_{t-1} - 0.926 \text{lgold}_{t-1} + 2.26) + 0.25 \Delta \text{loil}_{t-1} + e_{1t}.$$

In our earlier notation,

$$\Delta Y_t = c + \alpha_1(Y_{t-1} - \beta_1 - \beta_2 X_{t-1}) + \gamma \Delta Y_{t-1} + e_{1t},$$

with  $c = 0.0036$ ,  $\alpha_1 = -0.035 < 0$  as desired,  $\beta_1 = -2.26$ ,  $\beta_2 = 0.926$ , and  $\gamma = 0.25$ .

- Interpretation: there is an equilibrium relationship between  $\text{loil}$  and  $\text{lgold}$ . In case of a disequilibrium,  $\text{loil}$  adjusts towards the equilibrium. The adjustment amounts to 3.5% per period.

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# Learning Goals

## Students

- Understand the concept of cointegration,
- are able to test for cointegration using the Engle-Granger procedure,
- and are able to estimate an error correction model.



# Homework

- Exercise 7