

Lucerne University of Applied Sciences

Time Series Analysis

Final Exam

Last name: Answer Key

Given name:

mscbf_ rm01_ tsa, Fall term

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mscbf_rm01_tsa

Question 0**1.** Hi

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| |
| 0 pts |

Question 1 (27 points)

1. I have simulated 1000 observations $\{y_t\}$ from an ARMA(p, q) model. Figure 2 on Page 8 shows a plot of the correlogram. Based on it, what do you think p and q are, and why?

6 pts

Solution: The PACF becomes insignificant after 2 lags, whereas the ACF decays exponentially. This is indicative of an AR(2) process.

2. The 11th sample autocorrelation (not shown in the graph) is $\hat{\tau}_{11} = -0.045$. Test if $\hat{\tau}_{11}$ is significantly different from zero.

6 pts

Solution:

- The hypotheses are $H_0 : \tau_{11} = 0$ vs. $H_a : \tau_{11} \neq 0$.
- The test statistic is $\hat{\tau}_{11}$.
- Its distribution is $N(0, 1/T)$ under H_0 .
- The critical value is thus $1.96/\sqrt{1000} = 0.062$.
- The observed test statistic is $\hat{\tau}_{11} = -0.045$.
- Since $|-0.045| < 0.062$, we do not reject H_0 .
- Conclusion: τ_{11} is not significantly different from zero.

3. The output in Figure 3 on Page 8 shows the results of regressing Δy_t (y) on y_{t-1} (x1) and Δy_{t-1} (x2). Use it to test if the data are integrated.

6 pts

Solution:

- The hypotheses are $H_0 : y_t$ is integrated of order 1 vs. $H_a : y_t$ is stationary.
- The test statistic is the estimated coefficient on the lagged level, y_{t-1} (x1).
- Its follows a Dickey-Fuller distribution under H_0 .
- The critical value is -2.86, because a constant has been included.
- The observed test statistic is -6.52.
- Since $-6.52 < -2.86$, we reject H_0 .
- Conclusion: the data are stationary.

4. I have estimated a particular ARMA model. The estimation output is shown in Figure 4 on Page 9. Write down the estimated model in equation form.

3 pts

Solution: The estimated AR(2) model is

$$y_t = 0.2158 \cdot (1 - 0.5236 - 0.3595) + 0.5236 \cdot y_{t-1} + 0.3595 \cdot y_{t-2}.$$

Note `archmodel`'s peculiar interpretation of the intercept, `const`: in our notation, it corresponds to $\hat{\alpha}/(1 - \hat{\phi}_1 - \hat{\phi}_2)$, so that $\hat{\alpha} = \text{const} \cdot (1 - \hat{\phi}_1 - \hat{\phi}_2)$. See the exercises.

5. Use the estimated model from the previous question to forecast the value of the series at $t = 1001$. You may need some of the values below.

6 pts

| t | y_t | u_t |
|------|-------|-------|
| 999 | -0.77 | -4.18 |
| 1000 | -0.75 | -1.41 |

Solution: The forecast is

$$\begin{aligned}
 \hat{y}_{1001} &= \text{const} \cdot (1 - \hat{\phi}_1 - \hat{\phi}_2) + \hat{\phi}_1 \cdot y_{1000} + \hat{\phi}_2 \cdot y_{999} \\
 &= 0.2158 \cdot (1 - 0.5236 - 0.3595) + 0.5236 \cdot (-0.75) + 0.3595 \cdot (-0.77) \\
 &= -0.644.
 \end{aligned}$$

Question 2 (27 points) In this exercise, we analyze the daily returns on Tesla stock between 6/29/2010 and 12/14/2022. The returns are shown in Figure 1.

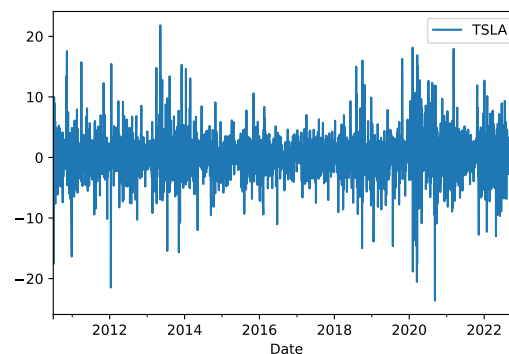


Figure 1: Returns on Tesla stock

The data clearly display volatility clustering, which we want to model using a GARCH model. The output is shown in Figure 5 on Page 9.

1. Write the model down as an equation (only the volatility equation).

Solution:

$$\sigma_{t+1}^2 = 0.14 + 0.0323 \cdot u_t^2 + 0.9564 \cdot \sigma_t^2.$$

3 pts

2. Does your model incorporate a leverage effect? Justify your answer.

Solution: No. Since u_t (the demeaned return) only enters as its square, the news impact curve is symmetric: volatility reacts equally strongly to a positive and negative shock of the same magnitude.

6 pts

3. Explain what the standardized residuals from a GARCH model are, ideally with an equation. Also, explain what their properties should be if the volatility model is correct.

Solution: The standardized residuals are defined as $\hat{z}_t = \frac{r_t - \hat{\mu}_t}{\hat{\sigma}_t}$. If the model for σ_t is correct, then they should no longer display volatility clustering. This means that there should be no autocorrelation in their squares.

6 pts

4. Regressing the squared standardized residuals on an intercept and 5 of their own lags results in an R^2 of 0.00364. Use this to test if the GARCH model has successfully removed the volatility clustering.

Solution: This is an ARCH-LM test.

- The hypotheses are H_0 : no remaining volatility clustering vs. H_a : there is remaining volatility clustering.
- The test statistic is $T \cdot R^2$ in the auxiliary regression.
- Its distribution is $\chi^2(5)$ under H_0 .
- The critical value is thus 11.07.
- The observed test statistic is $3138 \cdot 0.00364 = 11.42$.

6 pts

- Since $11.42 > 11.07$, we reject H_0 (although barely).
- Conclusion: Some volatility clustering remains.

5. Use the model to predict the variance σ_{t+1}^2 for 12/15/2022, using the following values on 12/14/2022: $\hat{u}_t = -2.72$ and $\sigma_t^2 = 15.198$.

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| 6 pts |
|-------|

Solution:

$$\begin{aligned}\hat{\sigma}_{t+1}^2 &= 0.14 + 0.0323 \cdot u_t^2 + 0.9564 \cdot \sigma_t^2 \\ &= 0.14 + 0.0323 \cdot (-2.72)^2 + 0.9564 \cdot 15.198 \\ &= 14.91.\end{aligned}$$

Question 3 (18 points) We now turn our attention to Value at Risk forecasting.

1. Use the model from the previous question to predict the 1% Value at Risk for 12/15/2022. Note: If you weren't successful in predicting the variance in the previous question, you can use the value $\hat{\sigma}_{t+1}^2 = 15.00$.

| |
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| |
| 6 pts |

Solution: The VaR forecast is

$$\begin{aligned} VaR_{t+1}^{0.01} &= -\mu_{t+1} - \sigma_{t+1} \cdot \Phi_{0.01}^{-1} \\ &= -\mu_{t+1} - \sigma_{t+1} \cdot \Phi_{0.01}^{-1} \\ &= -0.1118 - \sqrt{14.91} \cdot (-2.326) \\ &= 8.87. \end{aligned}$$

2. In a backtesting exercise, we have created 1% VaR forecasts for the entire sample. From this and the returns, we have created the hit series $\{I_t\}$.

| |
|--------|
| |
| 12 pts |

- (a) (6 pts) Explain how the hit series is defined, and how many times you expect it to equal 1 if the VaR model is correct.

Solution: The hit series equals 1 whenever a VaR violation occurs, i.e.,

$$I_{t+1} = \begin{cases} 1, & \text{if } R_{t+1} < -VaR_{t+1}^p, \\ 0, & \text{if } R_{t+1} > -VaR_{t+1}^p. \end{cases}$$

If the model is correct, we would expect a violation 1% of the time. Since $T = 3138$, we would therefore expect around 31 VaR violations.

- (b) (6 pts) Figure 6 on page 10 shows the result of regressing $(I_t - 0.01)$ on an intercept and I_{t-1} . Use it to test the independence of the VaR violations.

Solution:

- The hypotheses are H_0 : the VaR violations are independent vs. H_a : they are not.
- The test statistic is the t statistic for the lagged hit series.
- Its distribution is (asymptotically) standard normal under H_0 .
- The critical value is thus 1.96.
- The observed test statistic is 1.187.
- Since $|1.187| < 1.96$, we do not reject H_0 .
- Conclusion: there is no evidence of dependence between the VaR violations.

Note: in this case, since I didn't delete the p -value, you could alternatively have answered as follows.

- The hypotheses are H_0 : the VaR violations are independent vs. H_a : they are not.
- The test statistic is the t statistic for the lagged hit series.
- The observed p -value is $0.235 > 0.05$, so we do not reject H_0 .
- Conclusion: there is no evidence of dependence between the VaR violations.

Question 4 (18 points) Answer the questions below.

1.

- (a) (3 pts) Spurious regressions can occur between cointegrated variables.

☐ True☒ False

- (b) (3 pts) In a stationary time series, shocks
- U_t
- have a transitory effect on the future of the series.

☒ True☐ False

- (c) (3 pts) An ARCH(
- q
-) model for the returns corresponds to an AR(
- q
-) for the squared returns.

☒ True☐ False

- (d) (3 pts) The order
- q
- of an MA(
- q
-) model can be determined from the correlogram.

☒ True☐ False

- (e) (3 pts) In the presence of the leverage effect, the news impact curve is steeper to the right of the origin than to the left.

☐ True☒ False

- (f) (3 pts) A VaR model is correctly specified if no VaR violations occur.

☐ True☒ False

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| 18 pts |

End of exam.

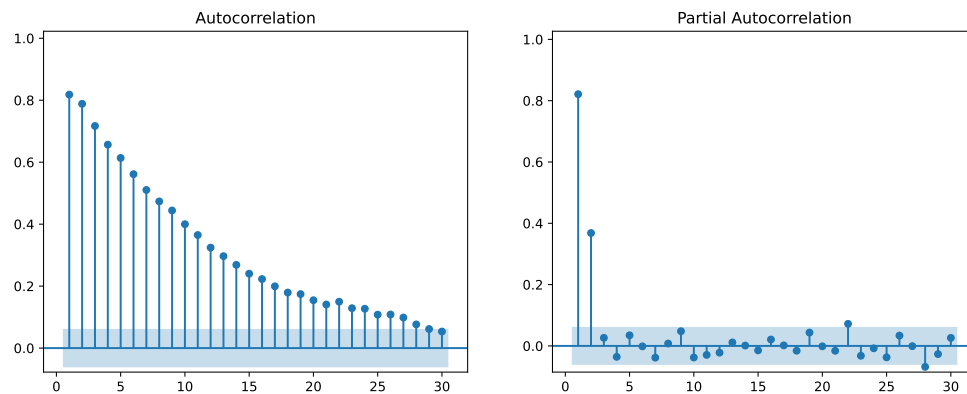


Figure 2: Correlogram of simulated data.

| OLS Regression Results | | | | | | |
|------------------------|------------------|---------------------|----------|-------|--------|--------|
| ===== | | | | | | |
| Dep. Variable: | y | R-squared: | 0.208 | | | |
| Model: | OLS | Adj. R-squared: | 0.207 | | | |
| Method: | Least Squares | F-statistic: | 130.9 | | | |
| Date: | Mon, 16 Oct 2023 | Prob (F-statistic): | 3.47e-51 | | | |
| Time: | 15:04:00 | Log-Likelihood: | -2086.9 | | | |
| No. Observations: | 998 | AIC: | 4180. | | | |
| Df Residuals: | 995 | BIC: | 4195. | | | |
| Df Model: | 2 | | | | | |
| Covariance Type: | nonrobust | | | | | |
| ===== | | | | | | |
| | coef | std err | t | P> t | [0.025 | 0.975] |
| ----- | | | | | | |
| x1 | -0.1161 | 0.018 | -6.520 | 0.000 | -0.151 | -0.081 |
| x2 | -0.3594 | 0.030 | -12.164 | 0.000 | -0.417 | -0.301 |
| const | 0.0309 | 0.062 | 0.497 | 0.619 | -0.091 | 0.153 |
| ===== | | | | | | |
| Omnibus: | 0.529 | Durbin-Watson: | 2.023 | | | |
| Prob(Omnibus): | 0.768 | Jarque-Bera (JB): | 0.404 | | | |
| Skew: | 0.003 | Prob(JB): | 0.817 | | | |
| Kurtosis: | 3.098 | Cond. No. | 3.76 | | | |
| ===== | | | | | | |

Figure 3: Output for ADF test.

```

SARIMAX Results
=====
Dep. Variable:          y      No. Observations:          1000
Model:                ARIMA(2, 0, 0)  Log Likelihood      -2091.958
Date:                Mon, 16 Oct 2023  AIC                  4191.917
Time:                15:10:43         BIC                  4211.548
Sample:              0             HQIC                   4199.378
                  - 1000
Covariance Type:      opg
=====
              coef      std err          z      P>|z|      [0.025      0.975]
-----
const          0.2158        0.528        0.409      0.683      -0.818        1.250
ar.L1          0.5236        0.032       16.396      0.000        0.461        0.586
ar.L2          0.3595        0.032       11.326      0.000        0.297        0.422
sigma2         3.8369        0.169       22.770      0.000        3.507        4.167
=====
Ljung-Box (L1) (Q):          0.13      Jarque-Bera (JB):          0.36
Prob(Q):                    0.71      Prob(JB):          0.83
Heteroskedasticity (H):      0.86      Skew:          0.00
Prob(H) (two-sided):        0.18      Kurtosis:         3.09
=====

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Figure 4: Estimated ARMA model.

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Constant Mean - GARCH Model Results
=====
Dep. Variable:          TSLA      R-squared:          0.000
Mean Model:          Constant Mean  Adj. R-squared:      0.000
Vol Model:          GARCH      Log-Likelihood:     -8267.70
Distribution:        Normal      AIC:               16543.4
Method:          Maximum Likelihood  BIC:               16567.6
                  No. Observations:      3138
Date:          Mon, Oct 16 2023  Df Residuals:      3137
Time:          15:33:03      Df Model:          1
                  Mean Model
=====
              coef      std err          t      P>|t|      95.0% Conf. Int.
-----
mu          0.1118  5.774e-02        1.936  5.281e-02  [-1.359e-03,  0.225]
Volatility Model
=====
              coef      std err          t      P>|t|      95.0% Conf. Int.
-----
omega       0.1400        0.117        1.196      0.232  [-8.938e-02,  0.369]
alpha[1]    0.0323     1.341e-02        2.409  1.599e-02  [6.024e-03, 5.859e-02]
beta[1]     0.9564     2.205e-02       43.369      0.000      [ 0.913,  1.000]
=====

```

Figure 5: Estimated GARCH model.

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=====
                        OLS Regression Results
=====
Dep. Variable:      np.subtract(I, 0.01)    R-squared:                 0.000
Model:              OLS                    Adj. R-squared:            0.000
Method:             Least Squares          F-statistic:              1.409
Date:               Mon, 16 Oct 2023        Prob (F-statistic):       0.235
Time:               18:54:56               Log-Likelihood:           1976.8
No. Observations:   3137                  AIC:                     -3950.
Df Residuals:       3135                  BIC:                     -3938.
Df Model:           1
Covariance Type:    nonrobust
=====

```

| | coef | std err | t | P> t | [0.025 | 0.975] |
|------------|--------|---------|-------|-------|--------|--------|
| b0 | 0.0065 | 0.002 | 2.817 | 0.005 | 0.002 | 0.011 |
| I.shift(1) | 0.0212 | 0.018 | 1.187 | 0.235 | -0.014 | 0.056 |

```

=====
Omnibus:                 4069.691    Durbin-Watson:              2.001
Prob(Omnibus):            0.000      Jarque-Bera (JB):           412759.104
Skew:                     7.492      Prob(JB):                   0.00
Kurtosis:                 57.160      Cond. No.                   7.76
=====

```

Figure 6: Test regression for the Value at Risk backtest.