STAT 4224 HW #4

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```
set.seed(1)
knitr::opts_chunk$set(echo = TRUE)
```

1.

a.

```
y \leftarrow c(95,70,80,62,108,62,61,74) / 10
J <- length(y)</pre>
n \leftarrow c(25,23,20,24,24,22,22,20)
sigma <- sqrt(14.3/n); rm(n);</pre>
Sim.size <- 10000
log.post.tau.fun <- function(tau, y, sigma)</pre>
 V.mu <- 1 / sum( 1/(sigma^2 + tau^2) )</pre>
 mu.hat <- V.mu * sum( y / (sigma^2 + tau^2) )</pre>
 log.post \leftarrow (log(V.mu) - sum(log(sigma^2 + tau^2))) / 2
 log.post \leftarrow log.post - 0.5 * sum((y-mu.hat)^2 / (sigma^2 + tau^2))
 log.post
}
log.tau \leftarrow seq(-2.5, 2.5, .01);
T <- length(log.tau);</pre>
log.post.tau <- rep(NA, T)</pre>
for(t in 1:T)
 log.post.tau[t] <- log.post.tau.fun(tau=exp(log.tau[t]), y, sigma)</pre>
maxie <- max(log.post.tau)</pre>
log.post.tau <- log.post.tau - maxie; rm(maxie);</pre>
```

```
post.tau <- exp(log.post.tau)</pre>
S <- Sim.size
theta.sim <- matrix(NA, J, S)
delta <- (log.tau[2] - log.tau[1]) / 2</pre>
post.log.tau <- exp(log.tau) * post.tau</pre>
for(s in 1:S)
{
t <- sample(T, 1, prob=post.log.tau)
tau <- exp(log.tau[t] + runif(1, -delta, delta))</pre>
V.mu <- 1 / sum( 1 / (sigma^2 + tau^2) )
mu.hat <- V.mu * sum( y / (sigma^2 + tau^2) )</pre>
 mu <- rnorm(1, mean=mu.hat, sd=sqrt(V.mu))</pre>
 V <- 1 / (1/sigma^2 + 1/tau^2)
theta.hat <- V * (y/sigma^2 + mu/tau^2)</pre>
 theta.sim[,s] <- rnorm(J, mean=theta.hat, sd=sqrt(V))</pre>
}
CIs<-data.frame(seq(1,8,1), apply(theta.sim,1,function(x) quantile(x,0.025)),apply(theta.sim,1,
function(x) quantile(x,0.975)))
colnames(CIs) <- c("School", "95% CI Lower Bound", "95% CI Upper Bound")
     School 95% CI Lower Bound 95% CI Upper Bound
## 1
          1
                        7.774868
                                           10.625420
## 2
          2
                        5.663680
                                            8.594743
## 3
          3
                      6.417326
                                           9.418949
## 4
          4
                       5.012549
                                           7.845976
## 5
          5
                       8.705884
                                          11.806266
## 6
          6
                       4.985416
                                            7.938029
## 7
          7
                                           7.889974
                       4.909415
## 8
                       5.963686
                                           8.949002
b.
mean(theta.sim[7,] < theta.sim[6,])</pre>
## [1] 0.5312
mean(apply(theta.sim, 2, which.min)==7)
## [1] 0.3229
P(\theta_7 < \theta_6 | y) = 0.53
P(min(\theta_1, ..., \theta_8) = \theta_7 | y) = 0.3229
\mathbf{c}.
```

```
ytilde.sim <- matrix(rnorm(J*S, theta.sim, sd=sqrt(14.3)), J, S)
mean(ytilde.sim[7,] < ytilde.sim[6,])</pre>
## [1] 0.5024
mean(apply(ytilde.sim, 2, which.min)==7)
## [1] 0.1805
P(\tilde{y}_7 < \tilde{y}_6 | y) = 0.5024
P(min(\tilde{y}_1,...,\tilde{y}_8) = \tilde{y}_7|y) = 0.1805
2.
a.
x \leftarrow c(-3.3, +0.1, -1.1, +2.7, +2.0, -0.4)
y \leftarrow c(-2.6, -0.2, -1.5, +1.5, +1.9, -0.3)
n <- length(y)</pre>
Sxx \leftarrow sum(x^2); Syy \leftarrow sum(y^2); Sxy \leftarrow sum(x*y);
neg.log.q <- function(theta, Sxx, Syy, Sxy, n)</pre>
n/2 * log(1 - theta^2) + 1/(2 * (1 - theta^2)) *
   ( Sxx - 2 * theta * Sxy + Syy )
foo <- optim(0.5, fn=neg.log.q, method="L-BFGS-B",
  lower=0.000001, upper=0.999999,
  Sxx=Sxx, Syy=Syy, Sxy=Sxy, n=n)
theta.hat.MAP <- foo$par;</pre>
theta.hat.MAP
## [1] 0.8468805
The MAP estimator approximates theta to be 0.8468.
b.
theta.vals <- seq(.005, .995, .01); length(theta.vals);</pre>
## [1] 100
log.q <- -1 * neg.log.q(theta.vals, Sxx, Syy, Sxy, n)</pre>
maxie \leftarrow max(log.q); log.q \leftarrow log.q - maxie; rm(maxie);
q <- exp(log.q); rm(log.q);</pre>
```

```
theta <- theta.vals;</pre>
sum(theta * q) / sum(q);
                              sum(theta^2 * q) / sum(q);
## [1] 0.7754085
## [1] 0.6135621
Using quadrature, we approximate E[\theta|x,y] to be ~0.7754 and E[\theta^2|x,y] to be 0.6134.
c.
theta.sim <- sample(theta.vals, 1000, replace=T, prob=q)</pre>
theta.sim <- theta.sim + runif(1000, -.005, +.005) # 'jitter'
mean(theta.sim)
## [1] 0.7802546
mean(theta.sim^2)
## [1] 0.6208263
Using MC Integration, we approximate E[\theta|x,y] to be ~0.7721 and E[\theta^2|x,y] to be ~0.60876.
d.
S<-1000
neg.log.M <- neg.log.q(theta.hat.MAP, Sxx, Syy, Sxy, n)</pre>
theta.prop <- rep(NA, S); accept <- rep(NA, S);
for(s in 1:S)
{
theta.star <- runif(1)</pre>
log.prob <- neg.log.M - neg.log.q(theta.star,Sxx,Syy,Sxy,n)</pre>
accept[s] <- ( log(runif(1)) < log.prob )</pre>
theta.prop[s] <- theta.star; rm(theta.star); rm(log.prob);</pre>
}
theta.RS <- theta.prop[accept]</pre>
length(theta.RS)
## [1] 196
```

mean(theta.RS)

[1] 0.7582437

mean(theta.RS^2)

[1] 0.5908891

Using Rejection Sampling with a uniform proposal distribution, we approximate $E[\theta|x,y]$ to be ~0.772 and $E[\theta^2|x,y]$ to be ~0.611. The effective sample size is 227.

e.

```
theta.IS <- runif(S)
w <- exp(-neg.log.q(theta.IS, Sxx, Syy, Sxy, n))
sum(w * theta.IS) / sum(w)

## [1] 0.7708949
sum(w * theta.IS^2) / sum(w)

## [1] 0.6065504
w.tilde <- w / sum(w)
1 / sum(w.tilde^2)</pre>
```

[1] 318.5525

Using Importance Sampling with a uniform proposal distribution, we approximate $E[\theta|x,y]$ to be ~0.771 and $E[\theta^2|x,y]$ to be ~0.607. The effective sample size is ~320.