# Additive and Multiplicative Latent Factor Models for Network Inference

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#### Motivation

#### Relational data consists of

- a set of units or nodes
- a set of measurements,  $y_{ij}$ , specific to pairs of nodes (i, j)

Sender	Receiver	Event			$i$	j	k	l
i	j	$y_{ij}$		$\overline{i}$	NA		21:1.	21:1
	k	$y_{ik}$	$\longrightarrow$	-	14/	$y_{ij}$	$y_{ik}$	$y_{il}$
:	l	$y_{il}$	,	j	$y_{ji}$	NA	$y_{jk}$	$y_{jl}$
j	i	$y_{ji}$		k	$y_{ki}$	$y_{kj}$	NA	$y_{kl}$
	k	$y_{jk}$		,				
:	l	$y_{jl}$		l	$y_{li}$	$y_{lj}$	$y_{lk}$	NA
k	i	$y_{ki}$						
	j	$y_{kj}$				*		
:	l	$y_{kl}$						
l	i	$y_{li}$			//			
	j	$y_{lj}$						
:	k	$y_{lk}$				_ \		

#### Relational data assumptions

GLM: 
$$y_{ij} \sim \beta^T X_{ij} + e_{ij}$$

Networks typically show evidence against independence of  $e_{ii}$ :  $i \neq j$ 

Not accounting for dependence can lead to:

- biased effects estimation
- uncalibrated confidence intervals
- poor predictive performance
- inaccurate description of network phenomena

#### We've been hearing this concern for decades now:

Thompson & Walker (1982)	Beck et al. (1998)	Snijders (2011)
Frank & Strauss (1986)	Signorino (1999)	Erikson et al. (2014)
Kenny (1996)	Li & Loken (2002)	Aronow et al. (2015)
Krackhardt (1998)	Hoff & Ward (2004)	Athey et al. (2016)

#### Outline

- Nodal and dyadic dependencies in networks
  - Can model using the "A" in AME
- Third order dependencies
  - Can model using the "M" in AME
- Application

# What network phenomena? Sender heterogeneity

Values across a row, say  $\{y_{ij}, y_{ik}, y_{il}\}$ , may be more similar to each other than other values in the adjacency matrix because each of these values has a common sender i

	i	j	k	1
i	NA	Уij	Yik	Yil
j	Ујі	NA	Уjk	YjI
k	Уki	Уkj	NA	YkI
1	Уli	Уij	Yık	NA

# What network phenomena? Receiver heterogeneity

Values across a column, say  $\{y_{ji}, y_{ki}, y_{li}\}$ , may be more similar to each other than other values in the adjacency matrix because each of these values has a common receiver i

	i	j	k	1
i	NA	Уij	Yik	Yil
j	Ујі	NA	Уjk	YjI
k	Уki	Уkj	NA	YkI
1	Уli	Уij	Уlk	NA

#### What network phenomena? Sender-Receiver Covariance

Actors who are more likely to send ties in a network may also be more likely to receive them

	i	j	k	1
i	NA	Уij	Уik	Yil
j	Ујі	NA	Уjk	УјІ
k	Уki	Уkj	NA	YkI
1	Уli	Уij	Уlk	NA

# What network phenomena? Reciprocity

Values of  $y_{ii}$  and  $y_{ii}$  may be statistically dependent

	i	j	k	1
i	NA	Уij	Yik	Yil
j	Ујі	NA	Уjk	YjI
k	Уki	Уkj	NA	YkI
1	Ун	Уij	Уlk	NA

We use this model to form the additive effects portion of AME

$$y_{ij} = \mu + e_{ij}$$
 $e_{ij} = a_i + b_j + \epsilon_{ij}$ 
 $\{(a_1, b_1), \dots, (a_n, b_n)\} \sim N(0, \Sigma_{ab})$ 
 $\{(\epsilon_{ij}, \epsilon_{ji}) : i \neq j\} \sim N(0, \Sigma_{\epsilon}), \text{ where}$ 
 $\Sigma_{ab} = \begin{pmatrix} \sigma_a^2 & \sigma_{ab} \\ \sigma_{ab} & \sigma_b^2 \end{pmatrix} \quad \Sigma_{\epsilon} = \sigma_{\epsilon}^2 \begin{pmatrix} 1 & \rho \\ \rho & 1 \end{pmatrix}$ 

- $\mu$  baseline measure of network activity
- $e_{ii}$  residual variation that we will use the SRM to decompose

$$egin{aligned} y_{ij} &= \mu + e_{ij} \ e_{ij} &= a_i + b_j + \epsilon_{ij} \ \{ ig( a_1, b_1 ig), \ldots, ig( a_n, b_n ig) \} &\sim \mathit{N}(0, \Sigma_{ab}) \ \{ ig( \epsilon_{ij}, \epsilon_{ji} ig) : i 
eq j \} &\sim \mathit{N}(0, \Sigma_{\epsilon}), \text{ where} \ \Sigma_{ab} &= egin{pmatrix} \sigma_a^2 & \sigma_{ab} \ \sigma_{ab} & \sigma_b^2 \end{pmatrix} & \Sigma_{\epsilon} &= \sigma_{\epsilon}^2 egin{pmatrix} 1 & \rho \ \rho & 1 \end{pmatrix} \end{aligned}$$

- row/sender effect  $(a_i)$  & column/receiver effect  $(b_i)$
- Modeled jointly to account for correlation in how active an actor is in sending and receiving ties

$$y_{ij} = \mu + e_{ij}$$
  $e_{ij} = a_i + b_j + \epsilon_{ij}$   $\{(a_1, b_1), \dots, (a_n, b_n)\} \sim N(0, \Sigma_{ab})$   $\{(\epsilon_{ij}, \epsilon_{ji}) : i \neq j\} \sim N(0, \Sigma_{\epsilon}), \text{ where}$   $\Sigma_{ab} = \begin{pmatrix} \sigma_a^2 & \sigma_{ab} \\ \sigma_{ab} & \sigma_b^2 \end{pmatrix}$   $\Sigma_{\epsilon} = \sigma_{\epsilon}^2 \begin{pmatrix} 1 & \rho \\ \rho & 1 \end{pmatrix}$ 

- $\sigma_a^2$  and  $\sigma_b^2$  capture heterogeneity in the row and column means
- $\sigma_{ab}$  describes the linear relationship between these two effects (i.e., whether actors who send [receive] a lot of ties also receive [send] a lot of ties)

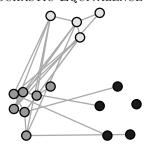
$$egin{aligned} y_{ij} &= \mu + e_{ij} \ e_{ij} &= a_i + b_j + \epsilon_{ij} \ \{(a_1,b_1),\ldots,(a_n,b_n)\} &\sim \textit{N}(0,\Sigma_{ab}) \ \{(\epsilon_{ij},\epsilon_{ji}): i 
eq j\} &\sim \textit{N}(0,\Sigma_{\epsilon}), ext{ where} \ \Sigma_{ab} &= egin{pmatrix} \sigma_{a}^2 & \sigma_{ab} \ \sigma_{ab} & \sigma_{b}^2 \end{pmatrix} & \Sigma_{\epsilon} &= \sigma_{\epsilon}^2 egin{pmatrix} 1 & \rho \ \rho & 1 \end{pmatrix} \end{aligned}$$

- $\epsilon_{ij}$  captures the within dyad effect
- Second-order dependencies are described by  $\sigma_\epsilon^2$
- Reciprocity, aka within dyad correlation, represented by  $\boldsymbol{\rho}$

#### Third Order Dependencies

# HOMOPHILY

#### STOCHASTIC EQUIVALENCE



To account for these patterns we can build on what we have so far and find an expression for  $\gamma$ :

$$y_{ij} \approx \beta^T X_{ij} + a_i + b_j + \gamma(u_i, v_j)$$

#### Latent Factor Model: The "M" in AME

Each node i has an unknown latent factor

$$\mathbf{u}_i, \mathbf{v}_j \in \mathbb{R}^k \ i, j \in \{1, \dots, n\}$$

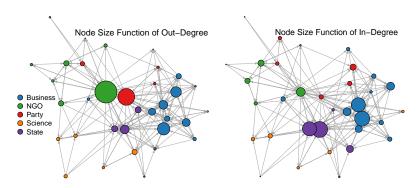
The probability of a tie from i to j depends on their latent factors

$$\begin{split} \gamma(\mathbf{u}_i, \mathbf{v}_j) &= \mathbf{u}_i^\mathsf{T} D \mathbf{v}_j \\ &= \sum_{k \in K} d_k u_{ik} v_{jk} \\ D \text{ is a } K \times K \text{ diagonal matrix} \end{split}$$

Can account for both stochastic equivalence and homophily

#### Swiss Climate Change Application

Cross-sectional network measuring whether an actor indicated that they collaborated with another during the policy design of the Swiss  $CO_2$  act (Ingold 2008)



#### Parameter Estimates

	Expected Effect	Logit	MRQAP	LSM	ERGM	AME
Conflicting policy preferences						
Business vs. NGO	_	-0.86	-0.87*	-1.37*	-1.11*	-1.37*
Opposition/alliance	+	1.21*	1.14*	0.00	1.22*	1.08*
Preference dissimilarity	_	-0.07	-0.60	-1.76*	-0.44	-0.79*
Transaction costs						
Joint forum participation	+	0.88*	0.75*	1.51*	0.90*	0.92*
Influence						
Influence attribution	+	1.20*	1.29*	0.08	1.00*	1.09*
Alter's influence indegree	+	0.10*	0.11*	0.01	0.21*	0.11*
Influence absolute diff.	_	-0.03*	-0.06*	0.04	-0.05*	-0.07*
$Alter = Government \ actor$	+	0.63*	0.68	-0.46	1.04*	0.55
Functional requirements						
$Ego = Environmental \; NGO$	+	0.88*	0.99	-0.60	0.79*	0.67
Same actor type	+	0.74*	1.12*	1.17*	0.99*	1.04*

#### Latent Factor Visualization

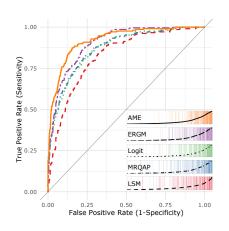


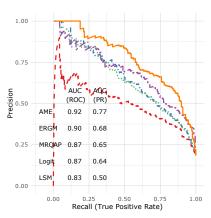
#### Out of Sample Performance Assessment

- Randomly divide the  $n \times (n-1)$  data points into S sets of roughly equal size, letting  $s_{ii}$  be the set to which pair  $\{ij\}$  is assigned.
- For each  $s \in \{1, \dots, S\}$ :
  - Obtain estimates of the model parameters conditional on  $\{y_{ij}: s_{ij} \neq s\}$ , the data on pairs not in set s.
  - For pairs  $\{kl\}$  in set s, let  $\hat{y}_{kl} = E[y_{kl}|\{y_{ij}: s_{ij} \neq s\}]$ , the predicted value of  $y_{kl}$  obtained using data not in set s.

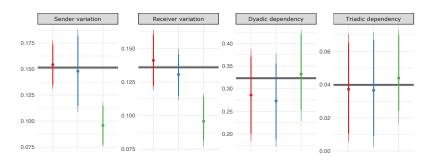
This procedure generates a sociomatrix of out-of-sample predictions of the observed data

# Performance Comparison





# Network Dependencies



♦ AME ♦ ERGM ♦ LDM

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#### Hope of Exponential Random Graph Models

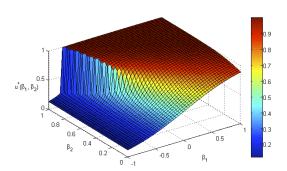
- Early 1970s development of pseudolikelihood estimation: Ove Frank (1971); Julian Besag (1972) proposed using a logistic regression with network characteristics as covariates. Birth of ERGM.
- $P(x) = exp(\Theta^T s(x) \psi(\theta))$ , where x is an adjacency matrix that is a graph,  $s(\theta)$  are some set of sufficient statistics for the graph, and  $\psi(\theta)$  is a normalizing constant, often set to be  $\log \sum_x e^{\theta^T zx}$ . This is often estimated via pseudolikelihood, simply by regressing  $x \sim \log i(z(x))$ .
- Maximum Likelihood is a better approach with Robbins and Monro, or the importance sampling approach of Geyer & Thompson. More advances with Bayesian approaches are available now with MCMC (Koskinen, Robins & Pattison).
- In the 1990s, networks became more widely recognized as and the ERGM approach was employed to estimate models in a variety of network domains. Needle sharing communities, HIV infections, for example.

#### There is a problem with ERGM.

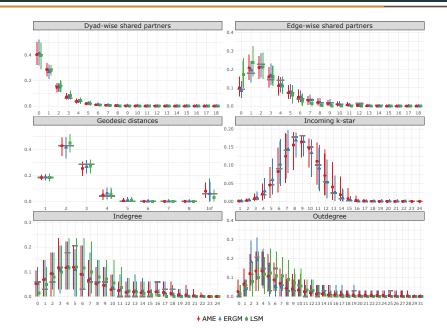
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#### It is a bug, not a feature.

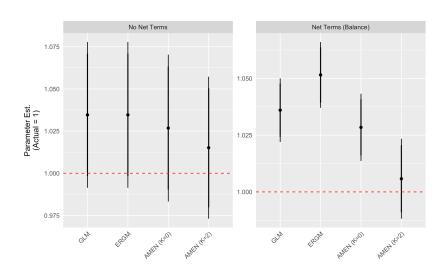
- •Probabilistic ERGM models place almost all of the probability on networks that are either nearly empty (degenerate) with no linkages or nearly saturated with all nodes being interconnected.
- •The likelihood surface contains steep or discontinuous gradients that render it impossible to solve numerically (or analytically). Even (especially) for very small networks this is problematic.



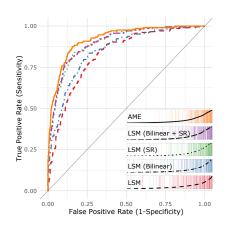
# Standard Network Dependence Measures

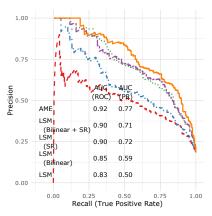


#### Simulation Comparison

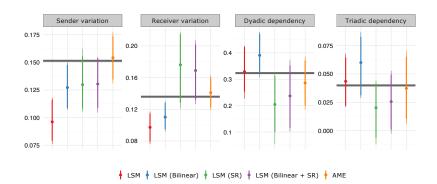


#### AMEN v LSM Performance

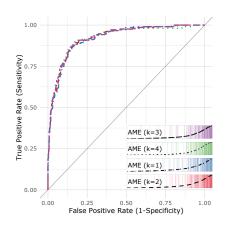


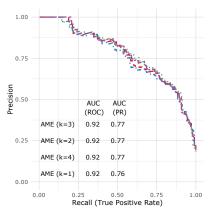


#### AMEN versus LSM Net Dependence



#### AMEN varying K Performance





# AMEN varying K Net Dependence

