Supporting Information for "Additive and Multiplicative Latent Factor Models for Network Inference"

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Abstract

There is growing interest in the study of social networks enabling scholars to move away from focusing on individual observations to examining more precisely the inter-relationships among observations. Many network approaches have been developed in a descriptive fashion, but attention to inferential approaches using statistical models of networks has been growing. We review a latent factor approach that models interdependencies among observations using additive and multiplicative effects (AME) which can be applied to binary, ordinal, and continuous network data. We compare this to alternative latent variable models as well as to exponential random graph models (ERGM). The AME approach can be easily implemented in the context of general linear models. It is both computationally straightforward and avoids degeneracy, which often plagues ERGM. In an out-of-sample context, it out-performs alternatives in terms of predicting links and capturing network dependencies.

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Limitations of Standard Framework

Generalized linear models are typically expressed via a stochastic and systematic component. The stochastic component reflects assumptions about the probability distribution from which the data are generated: $y_{ij} \sim P(Y|\theta_{ij})$, with a probability density or mass function such as the normal, binomial, or Poisson. Each dyad in the sample is independently drawn from a particular distribution, given θ_{ij}. The systematic component characterizes the model for the parameters of that distribution and describes how θ_{ij} varies as a function of a set of nodal and dyadic covariates, $\mathbf{X}_{ij}$: $\theta_{ij} = \boldsymbol{\beta}^T \mathbf{X}_{ij}$. A fundamental assumption we make when applying this modeling technique is that given $\mathbf{X}_{ij}$ and the parameters of the distribution, each of the dyadic observations is conditionally independent.

The importance of this assumption becomes clearer in the process of estimating a GLM via maximum likelihood. After having chosen a set of covariates and specifying a distribution, the joint density function over all dyads using the observations from Table 1 is given:

$$P(y_{ij}, y_{ik}, \ldots, y_{lk}|\theta_{ij}, \theta_{ik}, \ldots, \theta_{lk}) = P(y_{ij}|\theta_{ij}) \times P(y_{ik}|\theta_{ik}) \times \ldots \times P(y_{lk}|\theta_{lk})$$

$$P(\mathbf{Y} \mid \boldsymbol{\theta}) = \prod_{\alpha=1}^{n \times (n-1)} P(y_\alpha|\theta_\alpha) \tag{1}$$

The joint probability is converted into a likelihood: $\mathcal{L}(\boldsymbol{\theta}|\mathbf{Y}) = \prod_{\alpha=1}^{n \times (n-1)} P(y_\alpha|\theta_\alpha)$.

The parameters are estimated by maximizing the likelihood. The likelihood as defined above is only valid if y_{ij} is independent of y_{ji} and y_{ik} given the set of covariates we specified, or the values of θ_{ij}.

Ingold & Fischer Model Specification and Expected Effects

Variable	Description	Expected Effect
Conflicting policy preferences		
Business v. NGO	Binary, dyadic covariate that equals one when one actor is from the business sector and the other an NGO.	−
Opposition/alliance	Binary, dyadic covariate that equals one when i, sender, perceives j, receiver, as having similar policy objectives regarding climate change.	+
Preference dissimilarity	Transformation of four core beliefs into a Manhattan distance matrix, smaller the distance the closer the beliefs of i and j.	−
Transaction costs		
Joint forum participation	Binary, dyadic covariate that equals one when i and j belong to the same policy forum.	+
Influence		
Influence attribution	Binary, dyadic covariate that equals one when i considers j to be influential.	+
Alter's influence in-degree	Number of actors that mention i as being influential, this is a measure of reputational power.	+
Influence absolute diff.	Absolute difference in reputational power between i and j.	−
Alter = Government Actor	Binary, nodal covariate that equals one when j is a state actor.	+
Functional requirements		
Ego = Environment NGO	Binary, nodal covariate that equals one when i is an NGO.	+
Same actor type	Binary, dyadic covariate that equals when i and j are the same actor type.	+
Endogenous dependencies: ERGM Specific Parameters		
Mutuality	Captures concept of reciprocity, if i indicates they collaborated with j then j likely collaborates with i.	+
Outdegree popularity	Captures idea that actors sending more ties will be more popular targets themselves for collaboration.	+
Twopaths	Counts the number of two-paths in the network, two-path is an instance where i is connected to j, j to k, but i is not connected to k.	−
GWIdegree (2.0)	Takes into account how many ties a node sends in the network, used to capture network structures that result from some highly active nodes.	+
GWESP (1.0)	Counts the number of shared partners for each pair and sums across.	+
GWOdegree (0.5)	Takes into account how many ties a node receives in the network, used to capture networks structures that result from some highly popular nodes.	+

Table 1: Summary of variables to be included in model specification.

AME Model Convergence

Trace plot for AME model presented in paper.
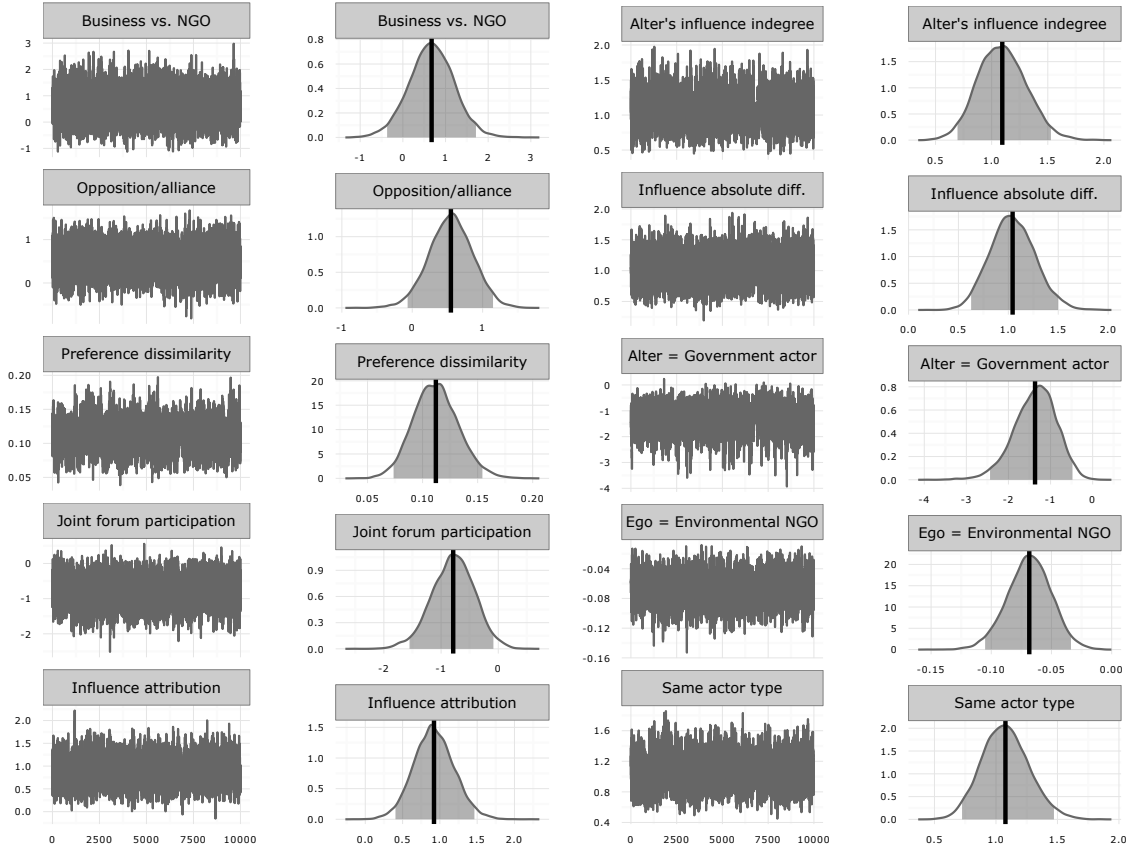
Figure 1: Trace plot for AME model presented in paper. In this model, we utilize the SRM to account for first and second-order dependence. To account for third order dependencies we use the latent factor approach with $K = 2$.

Other Network Goodness of Fit Tests

Below we show a standard set of statistics upon which comparisons are usually conducted:[1]

Variable	Description
Dyad-wise shared partners	Number of dyads in the network with exactly i shared partners.
Edge-wise shared partners	Similar to above except this counts the number of dyads with the same number of edges.
Geodesic distances	The proportion of pairs of nodes whose shortest connecting path is of length k, for $k = 1, 2, \ldots$. Also, pairs of nodes that are not connected are classified as $k = \infty$.
Incoming k-star	Propensities for individuals to have connections with multiple network partners.
Indegree	Proportion of nodes with the same value of the attribute as the receiving node.
Outdegree	Proportion of nodes with the same value of the attribute as the sending node.

Table 2: Description of a set of standard statistics used to assess whether a model captures network dependencies.

We simulate 1,000 networks from the LDM, ERGM, and AME model and compare how well they align with the observed network in terms of the statistics described in Table 2. The results are shown in Figure 2. Values for the observed network are indicated by a gray bar and average values from the simulated networks for the AME, ERGM, and LDM are represented by a diamond, triangle, and square, respectively. The densely shaded interval around each point represents the 95% interval from the simulations and the taller, less dense the 90% interval.[2] Looking across the panels in Figure 2 it is clear that there is little difference between the ERGM and AME models in terms of how well they capture network dependencies. The LDM model, however, does perform somewhat worse in comparison here as well. Particularly, when it comes to assessing the number of edge-wise shared partners and in terms of capturing the indegree and outdegree distributions of the collaboration network.

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[1]See Morris et al. (2008) for details on each of these parameters. If one was to examine goodness of fit in the **ergm** package these parameters would be calculated by default.

[2]Calculation for the incoming k-star statistic is not currently supported by the **latentnet** package.
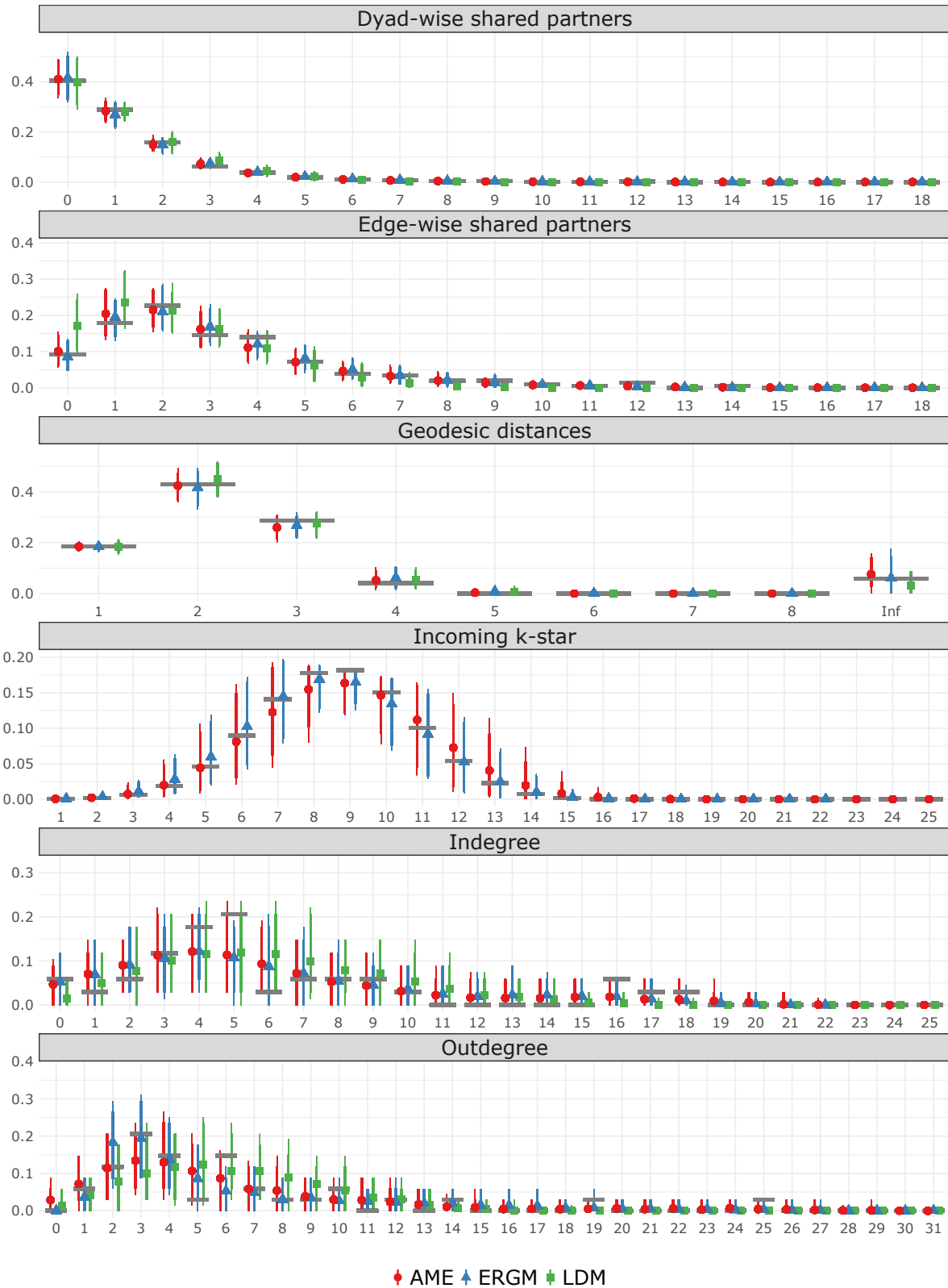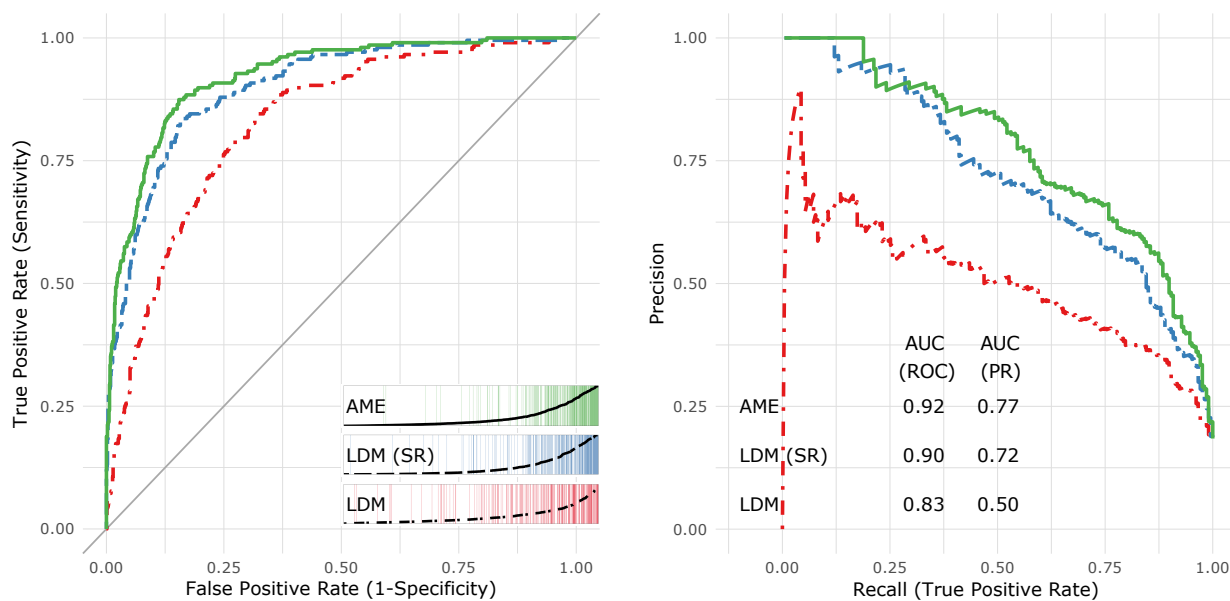
Figure 2: Goodness of fit statistics to assess how well the LDM, ERGM, and AME approaches account for network dependencies.

Comparison of amen & latentnet R Packages

Here we provide a comparison of the AME model we present in the paper with a variety of parameterizations from the **latentnet** package. The number of dimensions in the latent space in each of these cases is set to 2. LDM (SR) represents a model in which random sender and receiver effects are included.

	LDM	LDM (SR)	AME
Intercept/Edges	0.94*	0.60	-3.39*
	[0.09; 1.82]	[-1.10; 2.37]	[-4.38; -2.50]
Conflicting policy preferences			
Business vs. NGO	-1.37*	-3.07*	-1.37*
	[-2.42; -0.41]	[-4.77; -1.56]	[-2.44; -0.47]
Opposition/alliance	0.00	0.31	1.08*
	[-0.40; 0.39]	[-0.24; 0.86]	[0.72; 1.47]
Preference dissimilarity	-1.76*	-1.88*	-0.79*
	[-2.62; -0.90]	[-3.07; -0.68]	[-1.55; -0.08]
Transaction costs			
Joint forum participation	1.51*	1.56*	0.92*
	[0.86; 2.17]	[0.69; 2.41]	[0.40; 1.47]
Influence			
Influence attribution	0.08	0.30	1.09*
	[-0.40; 0.55]	[-0.37; 0.96]	[0.69; 1.53]
Alter's influence indegree	0.01	0.06	0.11*
	[-0.03; 0.04]	[-0.03; 0.14]	[0.07; 0.15]
Influence absolute diff.	0.04	-0.08*	-0.07*
	[-0.01; 0.09]	[-0.14; -0.02]	[-0.11; -0.03]
Alter = Government actor	-0.46	-0.11	0.55
	[-1.08; 0.14]	[-1.91; 1.76]	[-0.07; 1.15]
Functional requirements			
Ego = Environmental NGO	-0.60	-1.69	0.67
	[-1.32; 0.09]	[-3.74; 0.23]	[-0.38; 1.71]
Same actor type	1.17*	1.82*	1.04*
	[0.63; 1.71]	[1.10; 2.54]	[0.63; 1.50]

Table 3: * $p < 0.05$ (or 0 outside the 95% confidence interval).

Figure 3: Assessments of out-of-sample predictive performance using ROC curves, separation plots, and PR curves. AUC statistics are provided as well for both curves.
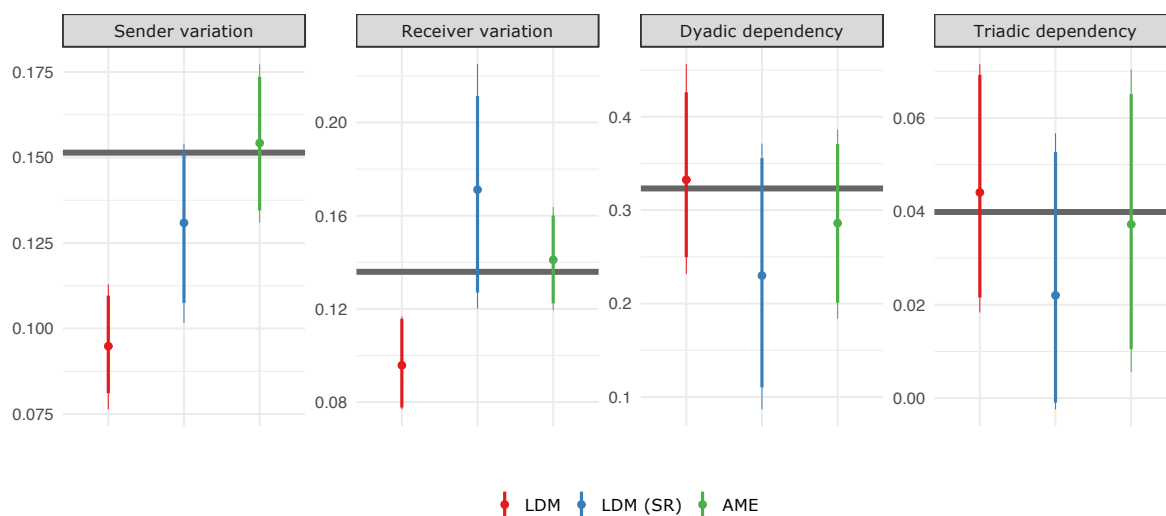
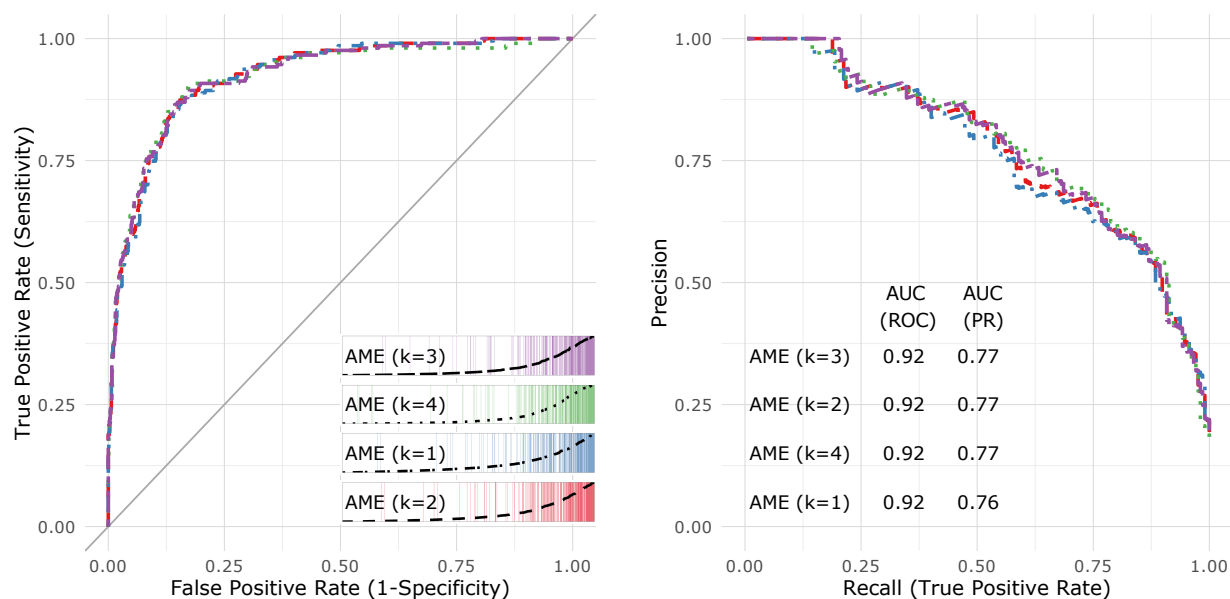**Figure 4:** Network goodness of fit summary using **amen**.

Comparison with other AME Parameterizations

Here we provide a comparison of the AME model we present in the paper that uses $K = 2$ for multiplicative effects and show how results change when we use $K = \{1, 3, 4\}$. Trace plots for $K = \{1, 3, 4\}$ are available upon request.

	AME (k=1)	AME (k=2)	AME (k=3)	AME (k=4)
Intercept/Edges	-3.08*	-3.39*	-3.72*	-3.93*
	[-3.91; -2.30]	[-4.38; -2.50]	[-4.84; -2.73]	[-5.12; -2.87]
Conflicting policy preferences				
Business vs. NGO	-1.28*	-1.37*	-1.48*	-1.51*
	[-2.20; -0.47]	[-2.44; -0.47]	[-2.63; -0.49]	[-2.69; -0.47]
Opposition/alliance	0.95*	1.08*	1.19*	1.28*
	[0.64; 1.27]	[0.72; 1.47]	[0.80; 1.64]	[0.86; 1.77]
Preference dissimilarity	-0.65*	-0.79*	-0.89*	-0.95*
	[-1.30; -0.03]	[-1.55; -0.08]	[-1.71; -0.12]	[-1.80; -0.14]
Transaction costs				
Joint forum participation	0.84*	0.92*	1.01*	1.06*
	[0.38; 1.31]	[0.40; 1.47]	[0.44; 1.62]	[0.43; 1.72]
Influence				
Influence attribution	1.00*	1.09*	1.21*	1.28*
	[0.63; 1.39]	[0.69; 1.53]	[0.75; 1.71]	[0.80; 1.84]
Alter's influence indegree	0.10*	0.11*	0.12*	0.13*
	[0.07; 0.14]	[0.07; 0.15]	[0.08; 0.17]	[0.09; 0.18]
Influence absolute diff.	-0.06*	-0.07*	-0.07*	-0.08*
	[-0.10; -0.03]	[-0.11; -0.03]	[-0.12; -0.04]	[-0.12; -0.04]
Alter = Government actor	0.52	0.55	0.60	0.64
	[-0.04; 1.07]	[-0.07; 1.15]	[-0.07; 1.27]	[-0.07; 1.35]
Functional requirements				
Ego = Environmental NGO	0.61	0.67	0.76	0.80
	[-0.31; 1.56]	[-0.38; 1.71]	[-0.38; 1.90]	[-0.40; 2.04]
Same actor type	0.97*	1.04*	1.11*	1.17*
	[0.60; 1.35]	[0.63; 1.50]	[0.64; 1.59]	[0.68; 1.68]

Table 4: * p < 0.05. 95% posterior credible intervals are provided in brackets.

Figure 5: Assessments of out-of-sample predictive performance using ROC curves, separation plots, and PR curves. AUC statistics are provided as well for both curves.
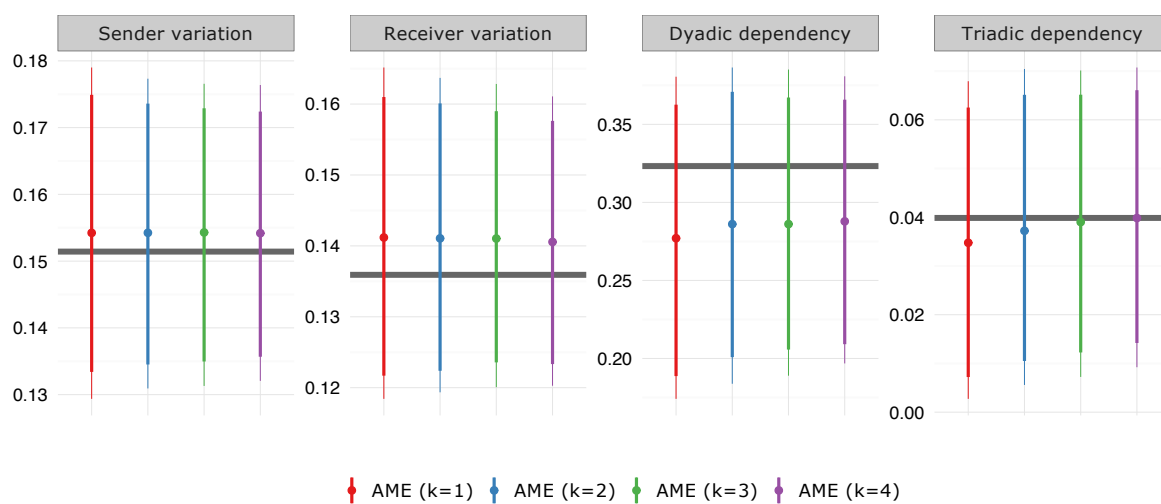
Figure 6: Network goodness of fit summary using **amen**.

Morris, M., Handcock, M. S., Hunter, D. R., 2008. Specification of exponential-family random graph models: Terms and computational aspects. Journal of Statistical Software 24 (4), 1554–7660.