

Regression

$$f(x) = f(x_1, x_2, x_3) = E(Y|X_1 = x_1, X_2 = x_2, X_3 = x_3)$$

$$E[(Y - \hat{f}(X))^2|X = x] = \underbrace{[f(x) - \hat{f}(x)]^2}_{\text{Reducible}} + \underbrace{\text{Var}(\epsilon)}_{\text{Irreducible}}, \text{ when } \epsilon = Y - f(X)$$

Curse of dimensionality: a neighbor in high-dimension need no longer be local(**sparse**)-"neighborhood" becomes less meaningful (getting apart)

Bayes classifier has smallest error in population using true $P_k(x)$, Bayes error: lowest test error, but true probability is unknown, KNN: k-hyperparameter

Linear function with error (Y: dependent, output, response, target / X: independent, input, predictor, feature)

$$Y = \beta_0 + \beta_1 x_1 + \dots + \beta_r x_r + e / \beta_0: \text{intercept}, \beta_i: \text{slope}, e: \text{error}$$

Bias-Variance Trade-off

$$E(y_0 - \hat{f}(x_0))^2 = \text{Var}(\hat{f}(x_0)) + [\text{Bias}(\hat{f}(x_0))]^2 + \text{Var}(\epsilon) / \text{Bias}(\hat{f}(x_0)) = E(\hat{f}(x_0)) - f(x_0) / \text{Var}[\hat{f}(x)] = E[(\hat{f}(x) - E[\hat{f}(x)])^2]$$

더 많은 데이터를 포착하려 더 많이 움직인다면 **(Flexible) bias↓, variance↑**

Degree of Freedom: the number of independent pieces of information

sample variance will be what is called an **Unbiased estimator** of the population variance σ^2

Least Squares estimator

$$SS = \sum_{i=1}^n (Y_i - A - Bx_i)^2 / \frac{SS_R}{\sigma^2} \sim \chi_{n-2}^2 / \text{minimizing SS condition is } A = \bar{Y} - B\bar{x} / \begin{matrix} B = \frac{S_{xy}}{S_{xx}} \\ \sum_{i=1}^n \frac{(Y_i - E[Y_i])^2}{\text{Var}(Y_i)} = \sum_{i=1}^n \frac{(Y_i - \alpha - \beta x_i)^2}{\sigma^2} \sim \chi_n^2 \end{matrix}$$

Sample means

Variance

(normal distribution iid random variables(Unbiased estimator))

$$\bar{y} \equiv \frac{1}{n} \sum_{i=1}^n y_i \text{ and } \bar{x} \equiv \frac{1}{n} \sum_{i=1}^n x_i \quad SE(\hat{\beta}_1)^2 = \frac{\sigma^2}{\sum_{i=1}^n (x_i - \bar{x})^2}, \quad SE(\hat{\beta}_0)^2 = \sigma^2 \left[\frac{1}{n} + \frac{\bar{x}^2}{\sum_{i=1}^n (x_i - \bar{x})^2} \right] \text{ or } \begin{matrix} \text{Var}(B) = \frac{\sigma^2}{S_{xx}}, & \text{Var}(A) = \frac{\sigma^2}{n} + \frac{\sigma^2}{S_{xx}} \bar{x}^2 \end{matrix}$$

Example for H0(null hypothesis H0 : $\beta_1=0$)

$$t = \frac{\hat{\beta}_1 - 0}{SE(\hat{\beta}_1)} / H_0: \text{there's no relationship between X and Y}$$

Significance level γ test of H_0

$$\begin{matrix} \text{reject } H_0 & \text{if } \sqrt{\frac{(n-2)S_{xx}}{SS_R}} |B| > t_{\gamma/2, n-2} & \text{v value: } \sqrt{(n-2)S_{xx}/SS_R} |B| = \text{point} \\ \text{accept } H_0 & \text{otherwise} & \text{p-value} = P\{|T_{n-2}| > v\} = \text{면적} \\ & & = 2P\{T_{n-2} > v\} \end{matrix}$$

$$P\{t_{30} > 2.04\} = 0.025, \quad P\{t_{30} > 2.75\} = 0.005 / 95\%(2.04): [\hat{\beta}_1 - 2 \cdot SE(\hat{\beta}_1), \hat{\beta}_1 + 2 \cdot SE(\hat{\beta}_1)]$$

Confidence interval estimator for β of SSR

$$P \left\{ B - \sqrt{\frac{SS_R}{(n-2)S_{xx}}} t_{\alpha/2, n-2} < \beta < B + \sqrt{\frac{SS_R}{(n-2)S_{xx}}} t_{\alpha/2, n-2} \right\} = 1 - \alpha / \alpha(\text{error에 대한 분포})를 모르기 때문에 ~t distribution, 안다면 ~N$$

H0: $\alpha_1=0$ & Confidence interval estimator for α of SSR

$$\sqrt{\frac{n(n-2)S_{xx}}{SS_R \sum_i x_i^2}} (A - \alpha) \sim t_{n-2} / A \pm t_{\alpha/2, n-2} \sqrt{\frac{SS_R \sum_i x_i^2}{n(n-2)S_{xx}}} = 1-\alpha$$

Inferences About Use the Distributional Result

$$\begin{matrix} \beta & \sqrt{\frac{(n-2)S_{xx}}{SS_R}} (B - \beta) \sim t_{n-2} \\ \alpha & \sqrt{\frac{n(n-2)S_{xx}}{\sum_i x_i^2 SS_R}} (A - \alpha) \sim t_{n-2} \end{matrix}$$

Residual Standard Error(RSE) / (SSR: input 빼고 error만 고려하자)

$$RSE = \sqrt{\frac{1}{n-2} \text{RSS}} = \sqrt{\frac{1}{n-2} \sum_{i=1}^n (y_i - \hat{y}_i)^2} / RSE = \sqrt{\frac{1}{n-p-1} \text{RSS}} \quad (\text{RSS감소가 p에 비해 미미하면 RSE는 more variable higher}) \text{OR}$$

$$E \left[\frac{SS_R}{n-2} \right] = \sigma^2 \quad \frac{SS_R}{\sigma^2} \sim \chi_{n-2}^2$$

R-squared(R^2) and sample correlation coefficient(only for a single): 설명가능한 정도 / 0~1 클수록 better fit of the model to the data / more variable, always increase / 얼마나 학습을 잘했는가?

Coefficient of determination

$$R^2 = \frac{S_{YY} - SS_R}{S_{YY}} = 1 - \frac{SS_R}{S_{YY}}$$

$$R^2 = \frac{TSS - RSS}{TSS} = 1 - \frac{RSS}{TSS} \text{ or } SS_R = \sum_{i=1}^n (Y_i - A - Bx_i)^2 \quad S_{YY} = \sum_{i=1}^n (Y_i - \bar{Y})^2$$

sample correlation coefficient :

$$r = \frac{\sum_{i=1}^n (x_i - \bar{x})(Y_i - \bar{Y})}{\sqrt{\sum_{i=1}^n (x_i - \bar{x})^2} \sqrt{\sum_{i=1}^n (Y_i - \bar{Y})^2}} = \frac{S_{xy}}{\sqrt{S_{xx}S_{yy}}} \quad |r| = \sqrt{R^2} \quad (-1 \sim 1)$$

multiple linear regression: x 전부를 고려하며 나머지는 임의의 값 **vs simple linear regression:** x: 1개만 고려하고 나머지는 0

ideal scenario: the predictors are uncorrelated(i.i.d.) but claims of causality

	Coefficient	Std. Error	t-statistic	p-value
Intercept	2.939	0.3119	9.42	< 0.0001
TV	0.046	0.0014	32.81	< 0.0001
radio	0.189	0.0086	21.89	< 0.0001
newspaper	-0.001	0.0059	-0.18	0.8599

coefficient: β_i

std.error: standard deviation of estimated β_i

t-statistic: normalized random variables for H_0 , β_i assuming their zero means / normal dist / only can use sample mean&variance, cause they are unknown / 개별 변수의 영향력(t-statistic, q=1일때의 f dist), 전체 변수의 영향력(f-statistic for a large number of features even though p-values are small) / 나머지를 모두 0으로 만든게 아닌 특정값을 넣고 계산(single과의 차이) z-statistic(binomial일때) vs t-statistic(random)

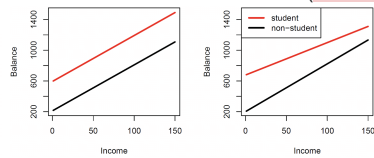
p-value: probability of t-statistics / H_0 의 accept reject 결정 / 클 수록 non-significant 작을수록 significant

	TV	radio	newspaper	sales
TV	1.0000	0.0548	0.0567	0.7822
radio		1.0000	0.3541	0.5762
newspaper			1.0000	0.2283
sales				1.0000

and, newspaper는 단독적으로는 의미가 매우 적지만, radio에 큰 영향 받고 sales에 영향 미친다 (correlation)

$$\text{balance}_i \approx \beta_0 + \beta_1 \times \text{income}_i + \begin{cases} \beta_2 + \beta_3 \times \text{income}_i & \text{if student} \\ 0 & \text{if not student} \end{cases}$$

$$= \begin{cases} (\beta_0 + \beta_2) + (\beta_1 + \beta_3) \times \text{income}_i & \text{if student} \\ \beta_0 + \beta_1 \times \text{income}_i & \text{if not student} \end{cases}$$



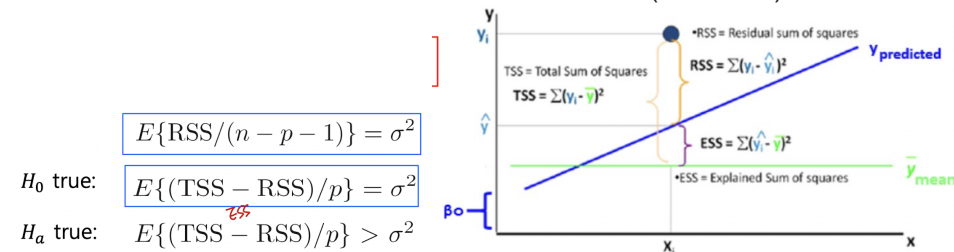
Is at Least one Predictor useful?

$$H_0 : \beta_1 = \beta_2 = \dots = \beta_p = 0$$

$$F = \frac{(TSS - RSS)/p}{RSS/(n-p-1)} \sim F_{p, n-p-1}$$

(n: sample 수, p: coefficient 수)

$$\text{Explained Sum of Squares (ESS)} = \sum (\hat{y}_i - \bar{y})^2 \rightarrow (TSS - RSS)$$



Deciding on important variables?

Forward selection: lowest RSS, stopping rule, greedy
Backward selection: largest p-value is removed

How well does model fit data?

R^2 or RSE

What Response value and how accurate?

confidence interval: average response / prediction interval: individual response

confidence interval은 평균을 이용하므로 sample의 수가 더 많을 때, Random error의 효과가 서로 상쇄된다.

즉, confidence interval이 prediction interval보다 항상 좁다 (=더 정확한 값을 얻을 수 있다는 이야기)

Interaction(Synergy)

한 축에 치우쳐서 평가하면 overestimate됨 / 두 변수의 곱으로 표현되는 변수 추가 ex) (radio * TV) or R^2: (96.8-89.7)/(100-89.7)=69% of variability in sales

hierarchy: hard to interpret in a model without main effects 즉, 매우 작은 p-value를 갖더라도 기존 variable을 유지해야 한다.

correlation of error terms: \rho가 커질수록 smooth / heteroscedasticity: 한쪽의 값이 커지면 variance가 커진다(키-몸무게) -> log 적용으로 해결

Outlier: x는 잘 있는데 y가 튀 (Affecting RSE, confidence interval and p-values) (너 말고도 많아)

High Leverage point: y는 잘 있는데 x가 튀 / 개별 predictor만으로는 판단X, 전체 predictor를 고려(애 없음 안돼)

Logistic Regression(Classification): linear regression은 0보다 작거나 1보다 큰 값을 뱉음, estimate를 잘 못함 / nearest-neighbor averaging(if 0 prob)

$$p(X) = \frac{e^{\beta_0 + \beta_1 X}}{1 + e^{\beta_0 + \beta_1 X}} \quad \log\left(\frac{p(X)}{1-p(X)}\right) = \beta_0 + \beta_1 X$$

- 사실상 probability (0~1이므로) -> get qualitative response Y(classifying)

Maximum likelihood

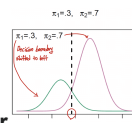
$$\ell(\beta_0, \beta) = \prod_{i: y_i=1} p(x_i) \prod_{i: y_i=0} (1-p(x_i))$$

: coefficient estimate, maximization

Discriminant Analysis(LDA, QDA, Naive bayes): bayes theorem기반 / normal (gaussian) distribution for each class

: more than 2 classes / small n / classes are well separated / two-classes일 경우 LDA=Logistic regression

$$\Pr(Y = k | X = x) = \frac{\Pr(X = x | Y = k) \cdot \Pr(Y = k)}{\Pr(X = x)} \rightarrow \Pr(Y = k | X = x) = \frac{\pi_k f_k(x)}{\sum_{l=1}^K \pi_l f_l(x)} \quad (\text{DA}) - \text{maximize posterior}$$



$$f_k(x) = \frac{1}{\sqrt{2\pi}\sigma_k} e^{-\frac{1}{2}\left(\frac{x-\mu_k}{\sigma_k}\right)^2}$$

(이때, \sigma=\sigma_k) 즉,

$$p_k(x) = \frac{\pi_k \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2}\left(\frac{x-\mu_k}{\sigma}\right)^2}}{\sum_{l=1}^K \pi_l \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2}\left(\frac{x-\mu_l}{\sigma}\right)^2}}$$

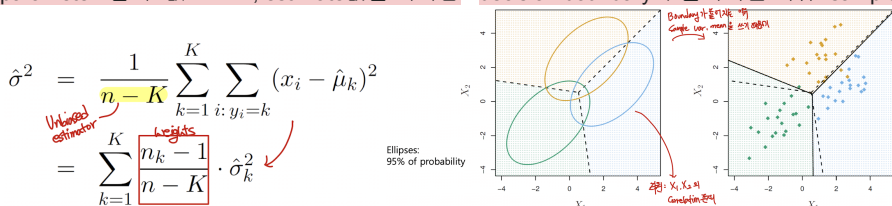
↑ max (즉, LDA 위)

Largest discriminant score

$$\delta_k(x) = x \cdot \frac{\mu_k}{\sigma^2} - \frac{\mu_k^2}{2\sigma^2} + \log(\pi_k)$$

- which is linear function / 만약 k=2, \pi_1=\pi_2=0.5일 때의 bayes decision boundary

parameter: 알 수 없으므로, estimate값을 써야 함 / decision boundary가 틀어지는 이유: sample variance, mean을 쓰기 때문에



softmax

$$\Pr(Y = k | X = x) = \frac{e^{\hat{\delta}_k(x)}}{\sum_{l=1}^K e^{\hat{\delta}_l(x)}}$$

/ False negative: true인데 negative로 판단 / Threshold를 수정해서 overfitting을 막을 수 있음

LDA vs QDA vs Naive Bayes

same covariance matrix \Sigma (LDA - 1차식) vs different own covariance matrix \Sigma (QDA - 2차식, overfitting in linear, flexible) vs diagonal matrix \Sigma (all independent)

