

# Mixture of $q$ -Weibull Distribution

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## Abstract

This study explores the use of  $q$ -Weibull distribution mixtures to model heterogeneous data. By integrating the flexibility of the  $q$ -Weibull distribution with mixture models, we develop a robust approach to capture complex data structures. The EM algorithm is employed for parameter estimation. Our results demonstrate significant improvements in modeling extreme events and non-Gaussian data, with potential applications in finance, healthcare, and reliability engineering.

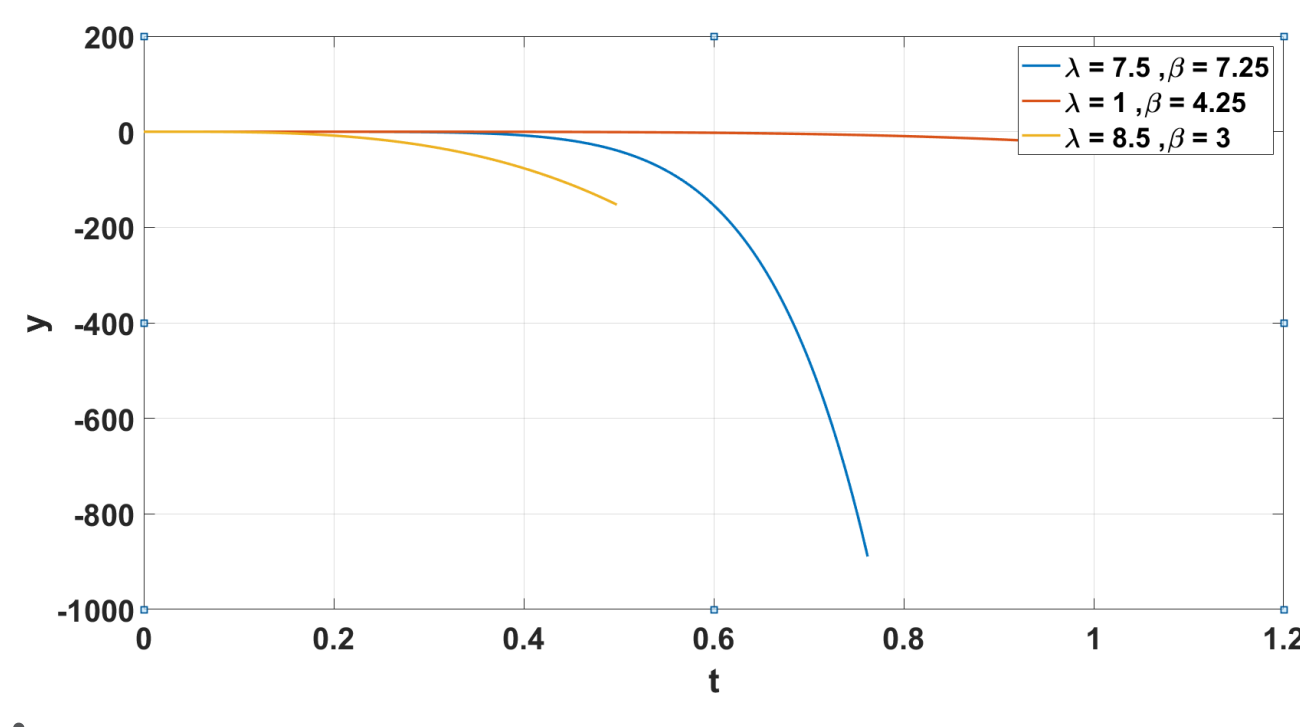
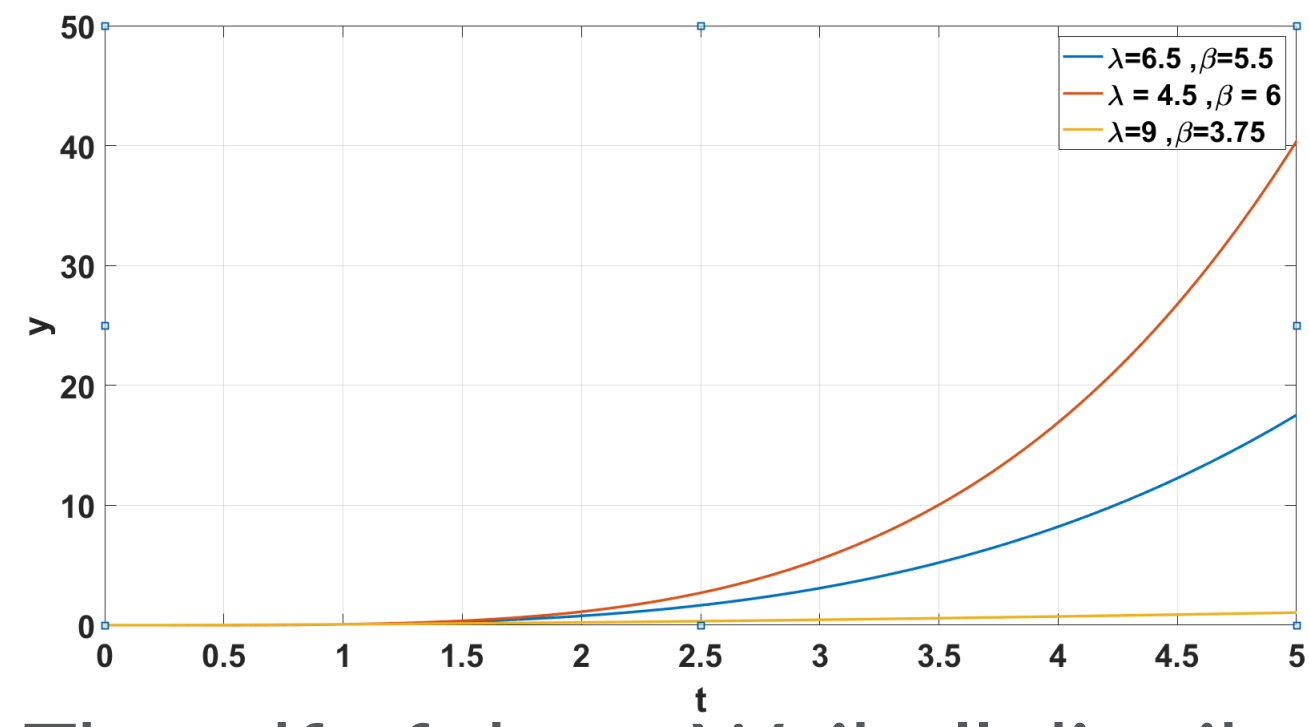
## 1) Definition

**$q$ -Weibull distribution:** The pdf of the  $q$ -Weibull distribution is obtained from the classical Weibull model by the substitution of  $q$ -exponential:

$$f_q(t) = (2 - q)\lambda\beta t^{\beta-1} [1 - (1 - q)\lambda t^\beta]^{1/1-q},$$

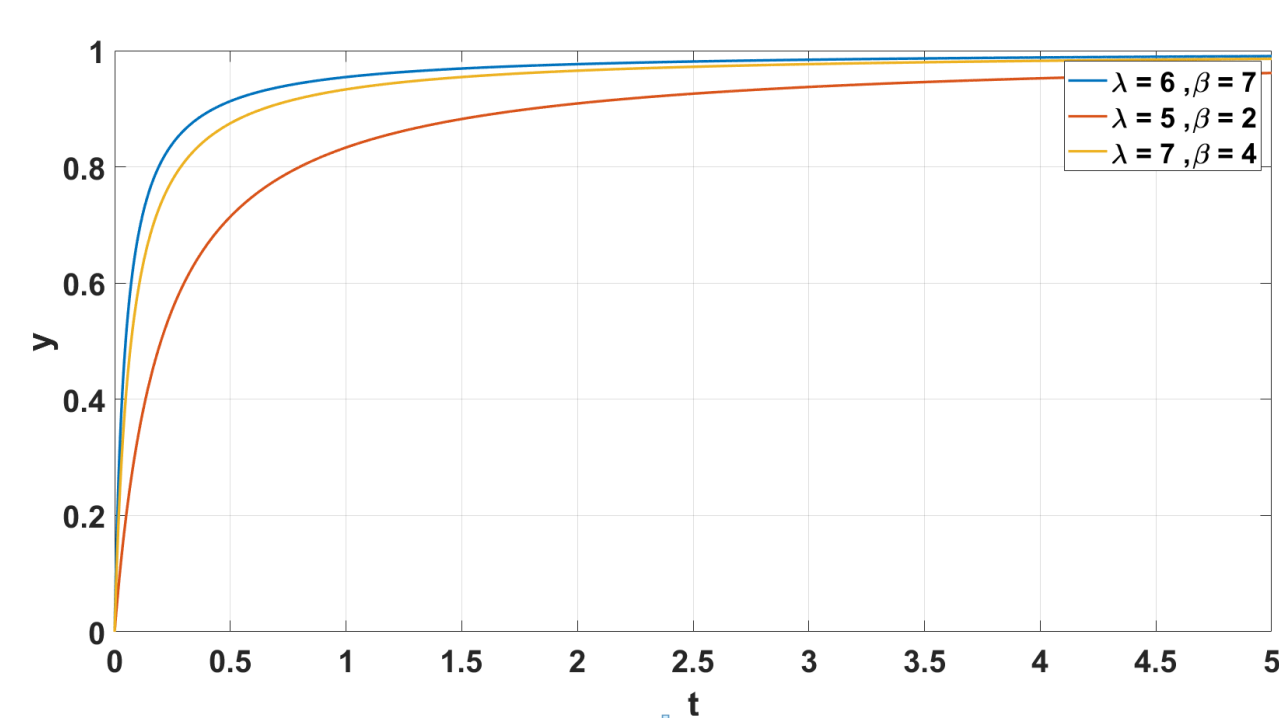
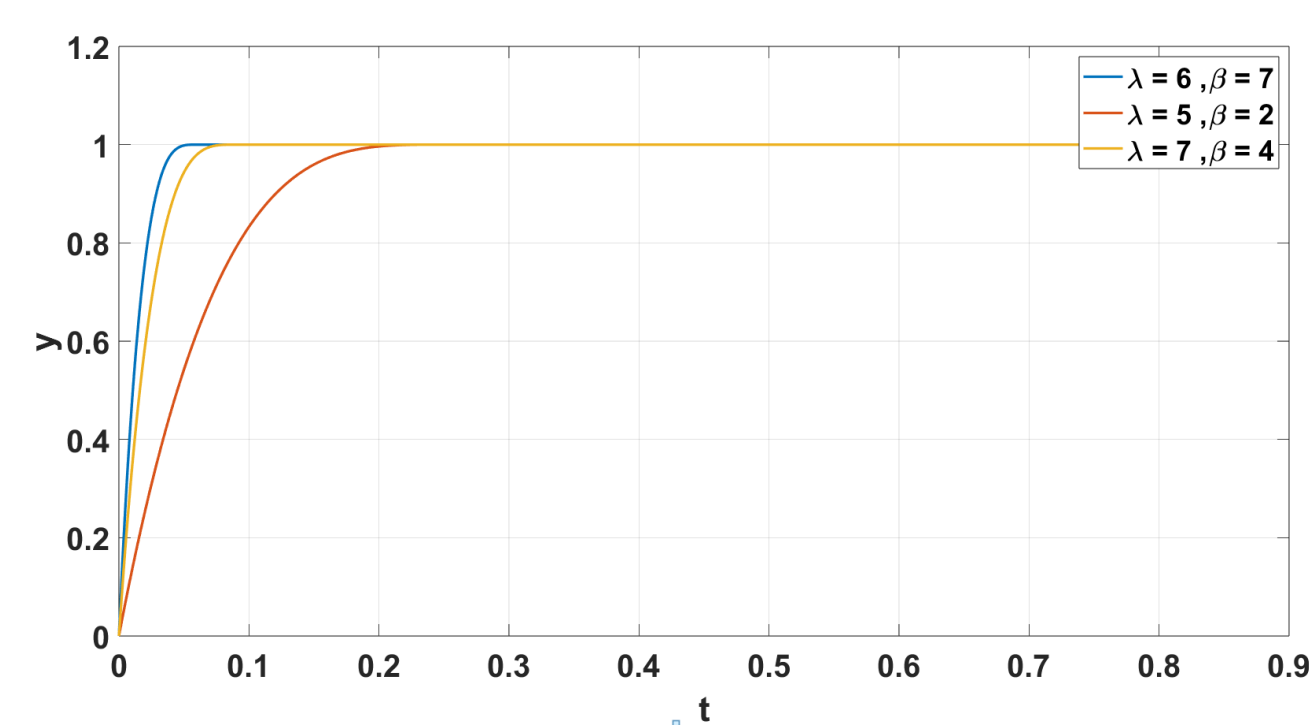
where

$$t \in \begin{cases} [0, +\infty[, & 1 < q < 2 \\ [0, [\lambda(1 - q)]^{-1/\beta}], & q < 1. \end{cases}$$



The cdf of the  $q$ -Weibull distribution is :

$$F_q(t) = 1 - [1 - (1 - q)\lambda t^\beta]^{2-q}.$$



## 2) Maximum likelihood estimation (MLE)

The MLE is a well known estimator. This estimator is important in statistics because of its asymptotic unbiasedness and minimal variance. Let  $T = t_1, \dots, t_n$  be a random sample from the  $q$ -Weibull distribution with set of parameters  $\theta = (q, \beta, \lambda)$ . The likelihood function  $L = (X|\theta)$  is :

$$L(X|\theta) = L = [(2 - q)\beta\lambda]^n \prod_{i=1}^n [t_i^{\beta-1} [1 - (1 - q)\lambda t_i^\beta]^{1/1-q}].$$

The first derivatives of the log-likelihood function with respect to  $q, \beta, \lambda$  are

$$\frac{\partial \ln(L)}{\partial q} = \frac{n}{q-2} + \frac{\lambda}{1-q} \sum_{i=1}^n \frac{t_i^\beta}{1 - (1 - q)\lambda t_i^\beta} + \frac{1}{(1-q)^2} \sum_{i=1}^n \ln [1 - (1 - q)\lambda t_i^\beta],$$

$$\frac{\partial \ln(L)}{\partial \beta} = \frac{n}{\beta} + \sum_{i=1}^n \ln(t_i) - \sum_{i=1}^n \frac{\lambda t_i^\beta \ln(t_i)}{[1 - (1 - q)\lambda t_i^\beta]},$$

$$\frac{\partial \ln(L)}{\partial \lambda} = \frac{n}{\lambda} - \sum_{i=1}^n \frac{t_i^\beta}{[1 - (1 - q)\lambda t_i^\beta]}.$$

## Bibliography

- [1] D. N. P. Murthy, M. Xie and R. Jiang (2004), Weibull Models, Wiley, New York, NY.
- [2] P. Bak, (1997), How Nature Works, Oxford University Press, Oxford.
- [3] KK. Jose, SR. Naik (2009) On the  $q$ -weibull distribution and its applications. Commun Stat-Theor Methods 38(6):912-926.
- [4] G. McLachlan and D. Peel. Finite mixture models : Wiley interscience, 2000.

## 3) Criteria for estimation

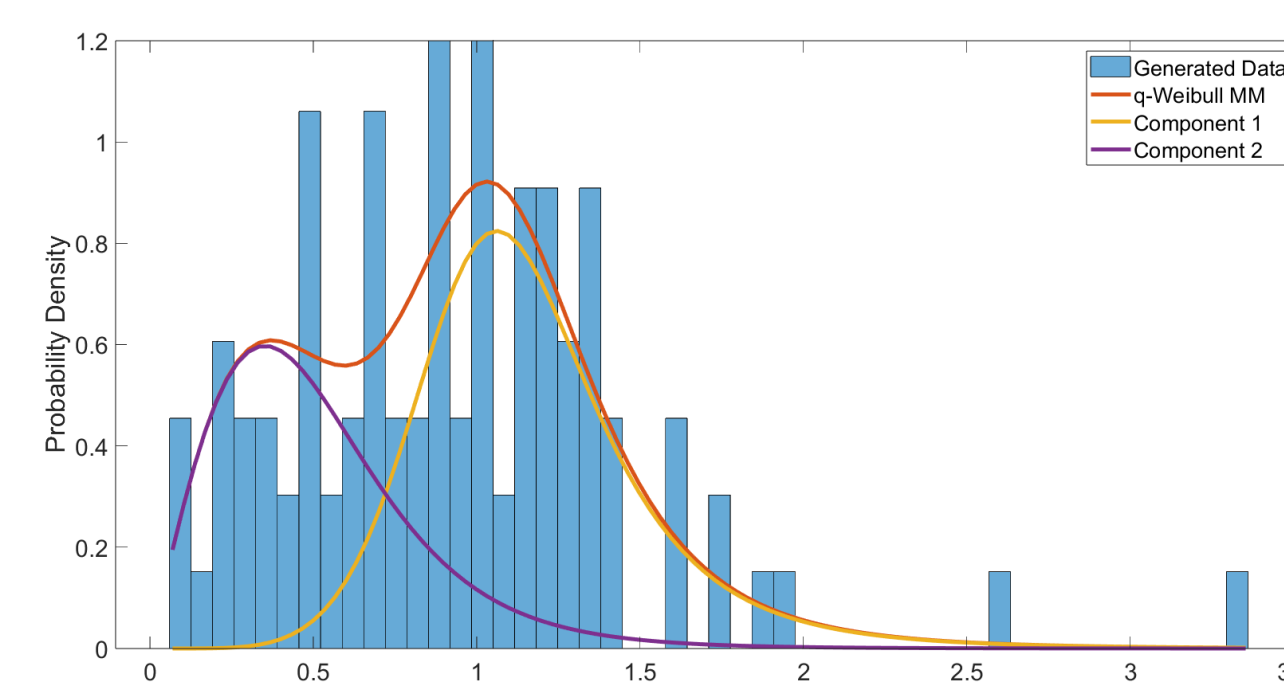
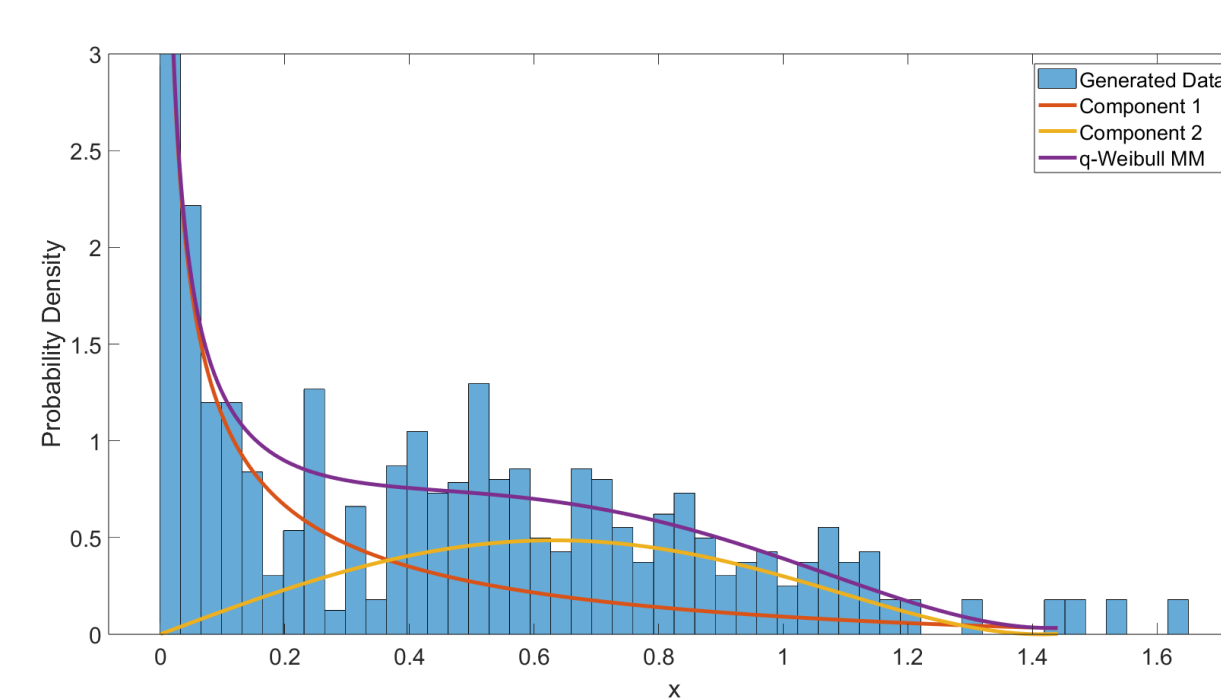
There are several selection criteria to determine which final model provides the best compromise between model fit and data complexity. The two most commonly used criteria are:  $AIC$ ,  $BIC$ .

## 4) The mixture of $q$ -Weibull distribution

The mixture of  $q$ -Weibull distribution is a method that combines several  $q$ -Weibull distribution to model complex data. Using the EM algorithm, the parameters of the mixture of  $q$ -Weibull distribution can be estimated from a set of observations.

$$f(x) = \sum_{i=1}^N \omega_i f_i(x, \theta_i),$$

where  $f_i(x, \theta_i)$  is the probability density function of the  $q$ -Weibull distribution  $i$  with parameters  $\theta_i$ , with  $\theta_i = (q_i, \beta_i, \lambda_i)$  for  $i \in \{1, \dots, N\}$ .



### Algorithm 1 EM Algorithm for a mixture of $q$ -Weibull distribution

Data  $q, \beta, \lambda, n$   
 $X = qwbvInd(q, \beta, \lambda, n)$  Estimation of parameters for the  $q$ -Weibull distribution  
 Maximum Likelihood  
 $\Theta = (\pi_1, \dots, \pi_N, q_1, \dots, q_N, \beta_1, \dots, \beta_N, \lambda_1, \dots, \lambda_N)$   
 Initialization  $\Theta$ ; repeat  
 Step E : For each  $x_t$ , calculate the conditional probabilities  $h_{tk}$  using the following formula :

$$h_{tk} = \frac{\pi_k f(x_t, q_k, \beta_k, \lambda_k)}{\sum_{j=1}^N \pi_j f(x_t, q_j, \beta_j, \lambda_j)}, \quad k = 1, \dots, N,$$

where  $f(x_t; q_k, \beta_k, \lambda_k)$  is the density of the  $q$ -Weibull distribution with parameters  $q_k, \beta_k, \lambda_k$ , evaluated at  $x_t$ .  
 Step M : Update the parameters as follows :

$$\pi_k = \frac{1}{n} \sum_{t=1}^n h_{tk}$$

$$q_k = \arg \max_{q_k} Q(\Theta, \Theta^{(l)}) = \arg \max_{q_k} \sum_{t=1}^n \sum_{k=1}^N h_{tk} \log f(x_t; q_k, \beta_k, \lambda_k)$$

$$\beta_k = \arg \max_{\beta_k} Q(\Theta, \Theta^{(l)}) = \arg \max_{\beta_k} \sum_{t=1}^n \sum_{k=1}^N h_{tk} \log f(x_t; q_k, \beta_k, \lambda_k)$$

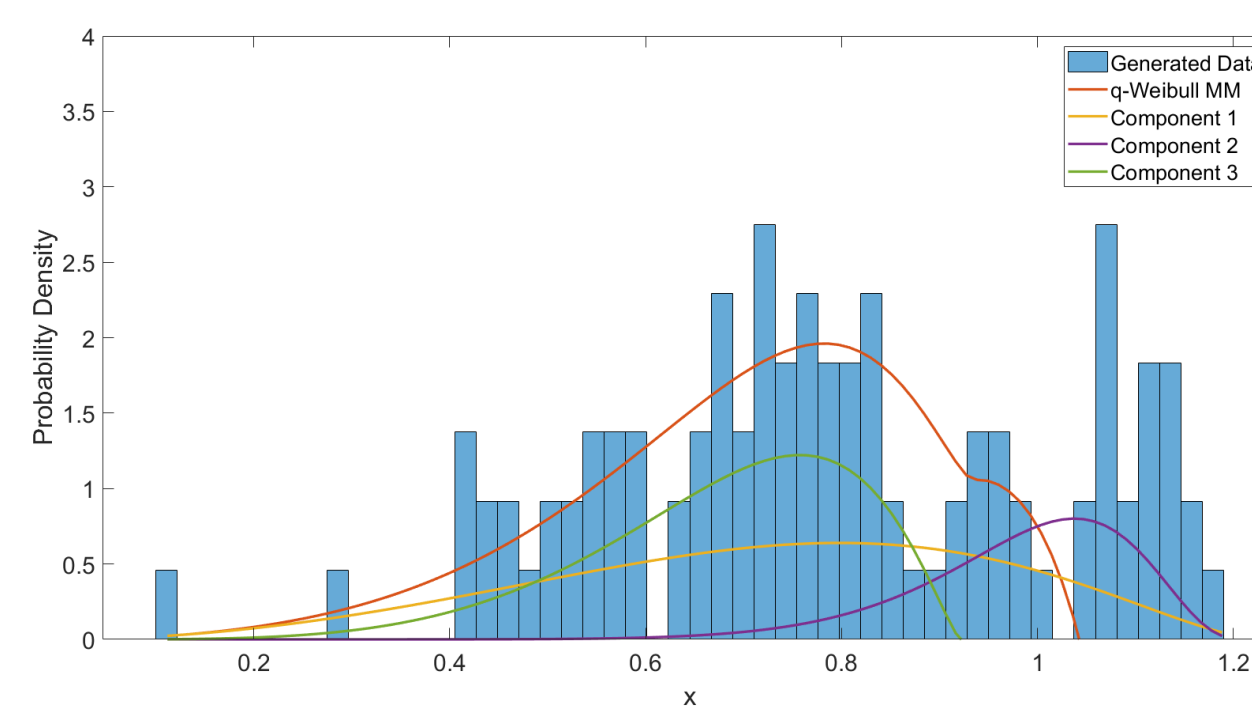
$$\lambda_k = \arg \max_{\lambda_k} Q(\Theta, \Theta^{(l)}) = \arg \max_{\lambda_k} \sum_{t=1}^n \sum_{k=1}^N h_{tk} \log f(x_t; q_k, \beta_k, \lambda_k)$$

until The convergence criterion is satisfied

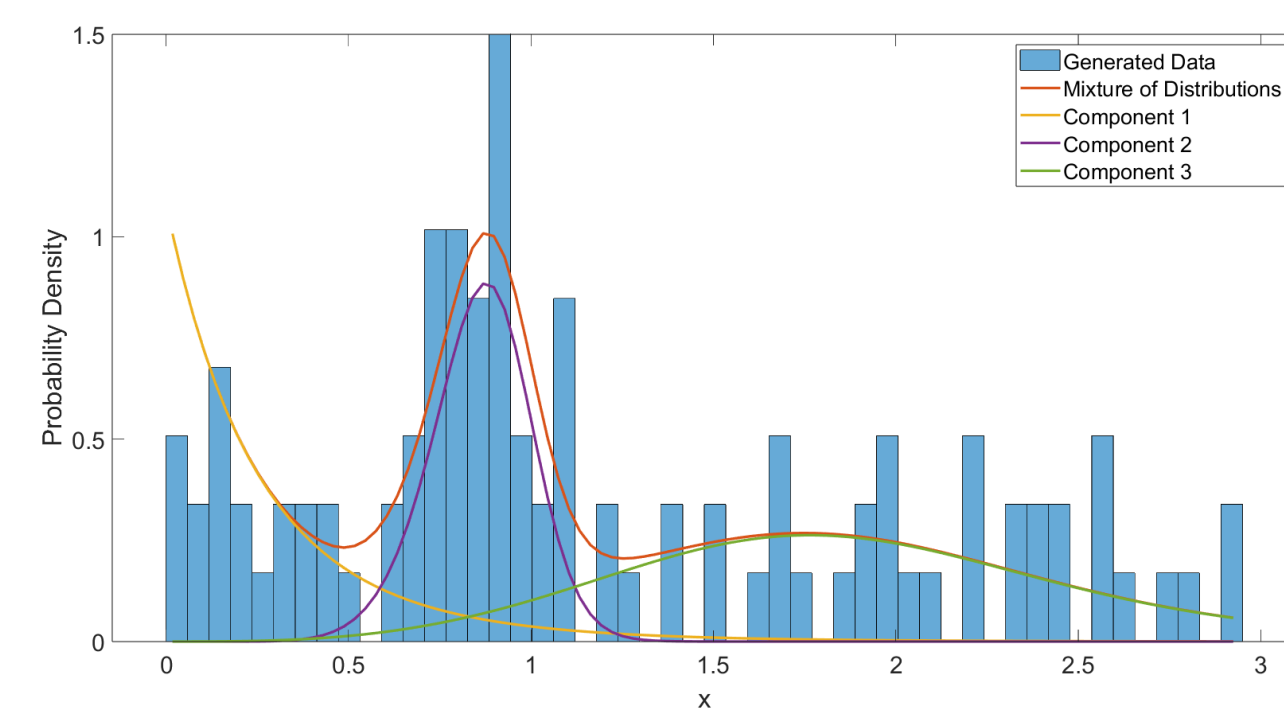
## 5) Result of simulation

To illustrate and highlight the performance of the EM algorithm described, some experimental results on simulated data, whose parameters are presented, are provided.

Parameter	True value	Estimated	std
$q_1$	0.4	0.3330	0.0474
$\beta_1$	2	1.9508	0.0348
$\lambda_1$	3	2.9470	0.0375
$q_2$	0.6	0.5396	0.0427
$\beta_2$	6	5.9702	0.0211
$\lambda_2$	7	6.9494	0.0358



Parameter	True value	Estimated	std
$q_1$	1.5	1.5041	0.0029
$\beta_1$	4	4.0162	0.0115
$\lambda_1$	3	2.9189	0.0573
$q_2$	1.3	1.3043	0.0030
$\beta_2$	1	1.0087	0.0062
$\lambda_2$	4.5	4.4027	0.0688



## Conclusion

This research investigates advanced statistical distribution  $q$ -Weibull and their applications in reliability engineering and risk management. It discusses their versatility, parameter estimation methods(MLE), model selection criteria, and practical applications using (EM) algorithm, serving as a valuable resource for improving decision-making and system performance across various domains.