Generative Learning vs Discriminative Learning
Linear Discriminant Analysis
Quadratic Discriminant Analysis
GLAD and QDA, Another point of view
Naive Bayes
Lecture Summary

Statistical Pattern Recognition Lecture4 Bayesian Learning

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Fall2014

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Introduction

Classification based on the theory Bayesian Learning

$$P(y = 0|\mathbf{X}) \geqslant_{v=1}^{y=0} P(y = 1|\mathbf{X})$$

• Classification involves determining $P(y|\mathbf{X})$, from different perspectives.

Generative Learning and Discriminative Learning

- Discriminative Learning:
 - Direct learning of $P(y|\mathbf{X})$.
 - Modelling of decision boundary, to which side a new sample is assigned.
 - Logistic and softmax regression are called discriminative learners.
- Generative Learning
 - Explicit modelling of each class separately.
 - Compare new sample with each class probability, based on Bayesian rule:

$$P(y|\mathbf{X}) = \frac{\overbrace{P(\mathbf{X}|y)}^{P(\mathbf{X}|y)}\overbrace{P(y)}^{P(\mathbf{Y})}}{\underbrace{P(\mathbf{X})}_{normalizing factor}}$$

• Model P(y) and P(X|y) for each class y:

$$P(y) = \phi_1^{1\{y=1\}} \phi_2^{1\{y=2\}} \cdots \phi_c^{1\{y=c\}}$$

$$P(\mathbf{X}|y=i) = \frac{1}{(\sqrt{2\pi})^n |\Sigma|^{\frac{1}{2}}} \exp(\frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_i)^T \Sigma^{-1}(\mathbf{X} - \boldsymbol{\mu}_i))$$

• Parameter set: $\theta = \{\phi_1, \phi_2, ..., \phi_c, \mu_1, \mu_2, ..., \mu_c, \Sigma\}$

• Estimate parameters: $\theta = \{\phi_1, \phi_2, ..., \phi_c, \mu_1, \mu_2, ..., \mu_c, \Sigma\}$

$$I(\boldsymbol{\theta}) = \log \prod_{j=1}^{m} P(\mathbf{X}^{(j)}, y^{(j)}) = \log \prod_{j=1}^{m} P(\mathbf{X}^{(j)} | P(y^{(j)})) P(y^{(j)})$$

• Estimate parameters: $\theta = \{\phi_1, \phi_2, ..., \phi_c, \mu_1, \mu_2, ..., \mu_c, \Sigma\}$

$$I(\boldsymbol{\theta}) = \log \prod_{j=1}^{m} P(\mathbf{X}^{(j)}, y^{(j)}) = \log \prod_{j=1}^{m} P(\mathbf{X}^{(j)} | P(y^{(j)})) P(y^{(j)})$$

• Take partial derivative in terms of each individual parameter:

$$\phi_i^{MLE} = \frac{\sum_{j=1}^{m} 1\{y^{(j)} = i\}}{m}$$

• Estimate parameters: $\theta = \{\phi_1, \phi_2, ..., \phi_c, \mu_1, \mu_2, ..., \mu_c, \Sigma\}$

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• Take partial derivative in terms of each individual parameter:

$$\phi_i^{MLE} = \frac{\sum_{j=1}^{m} 1\{y^{(j)} = i\}}{m}$$

$$\mu_i^{MLE} = \frac{\sum_{j=1}^m 1\{y^{(j)} = i\} \mathbf{X}^{(j)}}{\sum_{j=1}^m 1\{y^{(j)} = i\}}$$

• Estimate parameters: $\theta = \{\phi_1, \phi_2, ..., \phi_c, \mu_1, \mu_2, ..., \mu_c, \Sigma\}$

$$I(\boldsymbol{\theta}) = \log \prod_{j=1}^{m} P(\mathbf{X}^{(j)}, y^{(j)}) = \log \prod_{j=1}^{m} P(\mathbf{X}^{(j)} | P(y^{(j)})) P(y^{(j)})$$

• Take partial derivative in terms of each individual parameter:

$$\phi_i^{MLE} = \frac{\sum_{j=1}^m 1\{y^{(j)} = i\}}{m}$$

$$\mu_i^{MLE} = \frac{\sum_{j=1}^{m} 1\{y^{(j)} = i\} \mathbf{X}^{(j)}}{\sum_{i=1}^{m} 1\{y^{(j)} = i\}}$$

$$\Sigma^{MLE} = rac{1}{m} \sum_{i=1}^{m} (\mathbf{X}^{(j)} - oldsymbol{\mu}_{y^{(j)}}) (\mathbf{X}^{(j)} - oldsymbol{\mu}_{y^{(j)}})^T$$

• Determine class label of a new sample **X**^{new}:

$$y^{new} = argmax_y \ P(y|\mathbf{X})$$
 $= argmax_y \ \frac{P(\mathbf{X}|y)P(y)}{P(X)}$
 $= argmax_y \ P(\mathbf{X}|y)P(y)$

• Note that $P(\mathbf{X}|y)$ is a class dependent density.

Decision boundary is a line, a plane, or a hyper-plane. Why?

$$\frac{P(y=i|\mathbf{X}) = P(y=j|\mathbf{X})}{P(\mathbf{X}|y=i)P(y=i)} = \frac{P(\mathbf{X}|y=j)P(y=j)}{P(\mathbf{X})}$$

Decision boundary is a line, a plane, or a hyper-plane. Why?

$$\frac{P(y=i|\mathbf{X}) = P(y=j|\mathbf{X})}{P(\mathbf{X}|y=i)P(y=i)} = \frac{P(\mathbf{X}|y=j)P(y=j)}{P(\mathbf{X})}$$

$$P(X|y = i)P(y = i) = P(X|y = j)P(y = j)$$

Decision boundary is a line, a plane, or a hyper-plane. Why?

$$\frac{P(y=i|\mathbf{X}) = P(y=j|\mathbf{X})}{P(\mathbf{X}|y=i)P(y=i)} = \frac{P(\mathbf{X}|y=j)P(y=j)}{P(\mathbf{X})}$$

$$P(\mathbf{X}|y=i)P(y=i) = P(\mathbf{X}|y=j)P(y=j)$$

$$\begin{split} &\frac{1}{(2\pi)^{\frac{n}{2}}|\Sigma|^{\frac{1}{2}}} \; exp(\frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_i)^T \Sigma^{-1}(\mathbf{X} - \boldsymbol{\mu}_i)) P(y = i) \\ &= \frac{1}{(2\pi)^{\frac{n}{2}}|\Sigma|^{\frac{1}{2}}} \; exp(\frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_j)^T \Sigma^{-1}(\mathbf{X} - \boldsymbol{\mu}_j)) P(y = j) \end{split}$$

$$\exp(\frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_i)^T \boldsymbol{\Sigma}^{-1}(\mathbf{X} - \boldsymbol{\mu}_i)) P(y = i) = \exp(\frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_j)^T \boldsymbol{\Sigma}^{-1}(\mathbf{X} - \boldsymbol{\mu}_j)) P(y = j)$$

$$\exp(\frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_i)^T \boldsymbol{\Sigma}^{-1}(\mathbf{X} - \boldsymbol{\mu}_i)) P(y = i) = \exp(\frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_j)^T \boldsymbol{\Sigma}^{-1}(\mathbf{X} - \boldsymbol{\mu}_j)) P(y = j)$$

$$\frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_i)^T \Sigma^{-1}(\mathbf{X} - \boldsymbol{\mu}_i) + \log P(y = i) = \frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_j)^T \Sigma^{-1}(\mathbf{X} - \boldsymbol{\mu}_j) + \log P(y = j)$$

$$\begin{split} \exp(\frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_i)^T \boldsymbol{\Sigma}^{-1}(\mathbf{X} - \boldsymbol{\mu}_i)) P(y = i) &= \exp(\frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_j)^T \boldsymbol{\Sigma}^{-1}(\mathbf{X} - \boldsymbol{\mu}_j)) P(y = j) \\ \frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_i)^T \boldsymbol{\Sigma}^{-1}(\mathbf{X} - \boldsymbol{\mu}_i) + \log P(y = i) &= \frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_j)^T \boldsymbol{\Sigma}^{-1}(\mathbf{X} - \boldsymbol{\mu}_j) + \log P(y = j) \\ &\Rightarrow \log \frac{P(y = i)}{P(y = j)} - \frac{1}{2}[\mathbf{X}^T \boldsymbol{\Sigma}^{-1} \mathbf{X} - 2\mathbf{X}^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_i + \boldsymbol{\mu}_i^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_i] \\ &+ \frac{1}{2}[\mathbf{X}^T \boldsymbol{\Sigma}^{-1} \mathbf{X} - 2\mathbf{X}^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_j + \boldsymbol{\mu}_j^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_j] = 0 \end{split}$$

$$\exp(\frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_i)^T \boldsymbol{\Sigma}^{-1}(\mathbf{X} - \boldsymbol{\mu}_i)) P(y = i) = \exp(\frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_j)^T \boldsymbol{\Sigma}^{-1}(\mathbf{X} - \boldsymbol{\mu}_j)) P(y = j)$$

$$\frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_i)^T \boldsymbol{\Sigma}^{-1}(\mathbf{X} - \boldsymbol{\mu}_i) + \log P(y = i) = \frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_j)^T \boldsymbol{\Sigma}^{-1}(\mathbf{X} - \boldsymbol{\mu}_j) + \log P(y = j)$$

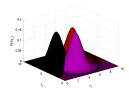
$$\Rightarrow \log \frac{P(y = i)}{P(y = j)} - \frac{1}{2}[\mathbf{X}^T \boldsymbol{\Sigma}^{-1} \mathbf{X} - 2\mathbf{X}^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_i + \boldsymbol{\mu}_i^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_i]$$

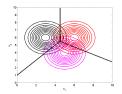
$$+ \frac{1}{2}[\mathbf{X}^T \boldsymbol{\Sigma}^{-1} \mathbf{X} - 2\mathbf{X}^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_j + \boldsymbol{\mu}_j^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_j] = 0$$

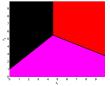
$$\Rightarrow \underline{\mathbf{X}^T \boldsymbol{\Sigma}^{-1}(\boldsymbol{\mu}_i - \boldsymbol{\mu}_j)}_{a\mathbf{X}} + \underline{\frac{1}{2}} \underline{\boldsymbol{\mu}_i^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_i - \frac{1}{2} \underline{\boldsymbol{\mu}_j^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_j + \log \frac{P(y = i)}{P(y = j)}} = 0$$

Analysis of GLDA when $\Sigma = \sigma^2 I$

- Classes are of identical distribution, but different means.
- Cross-section of classes distribution is spherical.
- Decision boundary is linear.
- Called classifier with nearest Euclidean distance to the class mean, when?



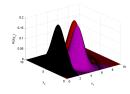


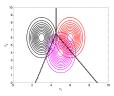


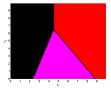
$$\Sigma = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

Analysis of GLDA when Σ is not identity

- Classes are of identical distribution, but different means.
- Cross-section of classes distribution is ellipsoidal.
- Decision boundary is linear.



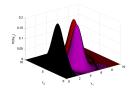


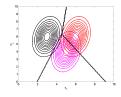


$$\Sigma = \begin{bmatrix} 0.5 & 0 \\ 0 & 1.5 \end{bmatrix}$$

Analysis of GLDA with arbitrary Σ

- Classes are of identical distribution, but different means.
- Cross-section of classes distribution is ellipsoidal. Linear decision boundary.
- Classes are aligned with direction of covariance eigenvectors.



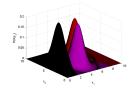


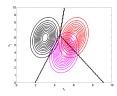


$$\Sigma = \begin{bmatrix} 0.5 & 0.2 \\ 0.2 & 1.5 \end{bmatrix}$$

Analysis of GLDA with arbitrary Σ

- Classes are of identical distribution, but different means.
- Cross-section of classes distribution is ellipsoidal. Linear decision boundary.
- Classes are aligned with direction of covariance eigenvectors.







$$\Sigma = \begin{bmatrix} 0.5 & 0.2 \\ 0.2 & 1.5 \end{bmatrix}$$

• Called classifier with nearest Mahalanobis distance to the class mean, when? $Dist(\mathbf{X}, \mu_i) = (\mathbf{X} - \mu_i)^T \Sigma^{-1} (\mathbf{X} - \mu_i)$

- Another generative learning model; a Bayesian classifier
- Here, classes are of multinomial distribution, and likelihoods are multivariate Gaussian with separate covariance Σ_i .
- Decision boundary becomes non-linear, Why?

$$P(\mathbf{X}|y=i) = \frac{1}{(\sqrt{2\pi})^n |\Sigma_i|^{\frac{1}{2}}} \exp(\frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_i)^T \Sigma_i^{-1} (\mathbf{X} - \boldsymbol{\mu}_i))$$

Decision boundary is parabolic:

$$P(y = i|\mathbf{X}) = P(y = j|\mathbf{X})$$

$$\frac{P(\mathbf{X}|y = i)P(y = i)}{P(\mathbf{X})} = \frac{P(\mathbf{X}|y = j)P(y = j)}{P(\mathbf{X})}$$

$$P(\mathbf{X}|y = i)P(y = i) = P(\mathbf{X}|y = j)P(y = j)$$

Decision boundary is parabolic:

$$P(y = i|\mathbf{X}) = P(y = j|\mathbf{X})$$

$$\frac{P(\mathbf{X}|y = i)P(y = i)}{P(\mathbf{X})} = \frac{P(\mathbf{X}|y = j)P(y = j)}{P(\mathbf{X})}$$

$$P(\mathbf{X}|y = i)P(y = i) = P(\mathbf{X}|y = j)P(y = j)$$

$$\begin{split} &\frac{1}{(2\pi)^{\frac{n}{2}}|\Sigma_{i}|^{\frac{1}{2}}} \exp(\frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_{i})^{T} \Sigma_{i}^{-1}(\mathbf{X} - \boldsymbol{\mu}_{i})) P(y = i) \\ &= \frac{1}{(2\pi)^{\frac{n}{2}}|\Sigma_{j}|^{\frac{1}{2}}} \exp(\frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_{j})^{T} \Sigma_{j}^{-1}(\mathbf{X} - \boldsymbol{\mu}_{j})) P(y = j) \end{split}$$

$$\begin{split} &\frac{1}{|\Sigma_i|^{\frac{1}{2}}} \exp(\frac{-1}{2} (\mathbf{X} - \boldsymbol{\mu}_i)^T \boldsymbol{\Sigma}_i^{-1} (\mathbf{X} - \boldsymbol{\mu}_i)) P(y = i) \\ &= \frac{1}{|\Sigma_j|^{\frac{1}{2}}} \exp(\frac{-1}{2} (\mathbf{X} - \boldsymbol{\mu}_j)^T \boldsymbol{\Sigma}_j^{-1} (\mathbf{X} - \boldsymbol{\mu}_j)) P(y = j) \end{split}$$

$$\begin{split} \frac{1}{|\Sigma_i|^{\frac{1}{2}}} \exp(\frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_i)^T \Sigma_i^{-1}(\mathbf{X} - \boldsymbol{\mu}_i)) P(y = i) \\ &= \frac{1}{|\Sigma_j|^{\frac{1}{2}}} \exp(\frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_j)^T \Sigma_j^{-1}(\mathbf{X} - \boldsymbol{\mu}_j)) P(y = j) \\ \frac{-1}{2} \log|\Sigma_i| &- \frac{1}{2} (\mathbf{X} - \boldsymbol{\mu}_i)^T \Sigma_i^{-1}(\mathbf{X} - \boldsymbol{\mu}_i) + \log P(y = i) \\ &= \frac{-1}{2} \log|\Sigma_j| &- \frac{1}{2} (\mathbf{X} - \boldsymbol{\mu}_j)^T \Sigma_j^{-1}(\mathbf{X} - \boldsymbol{\mu}_j) + \log P(y = j) \end{split}$$

$$\begin{split} \frac{1}{|\Sigma_i|^{\frac{1}{2}}} \exp(\frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_i)^T \boldsymbol{\Sigma}_i^{-1}(\mathbf{X} - \boldsymbol{\mu}_i)) P(y = i) \\ &= \frac{1}{|\Sigma_j|^{\frac{1}{2}}} \exp(\frac{-1}{2}(\mathbf{X} - \boldsymbol{\mu}_j)^T \boldsymbol{\Sigma}_j^{-1}(\mathbf{X} - \boldsymbol{\mu}_j)) P(y = j) \\ \frac{-1}{2} log |\Sigma_i| &- \frac{1}{2} (\mathbf{X} - \boldsymbol{\mu}_i)^T \boldsymbol{\Sigma}_i^{-1}(\mathbf{X} - \boldsymbol{\mu}_i) + log P(y = i) \\ &= \frac{-1}{2} log |\Sigma_j| &- \frac{1}{2} (\mathbf{X} - \boldsymbol{\mu}_j)^T \boldsymbol{\Sigma}_j^{-1}(\mathbf{X} - \boldsymbol{\mu}_j) + log P(y = j) \\ \Rightarrow &log \frac{P(y = i)}{P(y = j)} - \frac{1}{2} log \frac{|\Sigma_i|}{|\Sigma_j|} - \frac{1}{2} [\mathbf{X}^T \boldsymbol{\Sigma}_i^{-1} \mathbf{X} + \boldsymbol{\mu}_i^T \boldsymbol{\Sigma}_i^{-1} \boldsymbol{\mu}_i - 2\mathbf{X}^T \boldsymbol{\Sigma}_i^{-1} \boldsymbol{\mu}_i - \mathbf{X}^T \boldsymbol{\Sigma}_j^{-1} \mathbf{X} - \boldsymbol{\mu}_j^T \boldsymbol{\Sigma}_j^{-1} \boldsymbol{\mu}_j + 2\mathbf{X}^T \boldsymbol{\Sigma}_j^{-1} \boldsymbol{\mu}_j] = 0 \end{split}$$

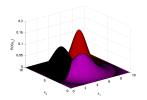
$$log \frac{P(y=i)}{P(y=j)} - \frac{1}{2}log \frac{\left|\Sigma_{i}\right|}{\left|\Sigma_{j}\right|} - \frac{1}{2}[\mathbf{X}^{T}(\Sigma_{i}^{-1} - \Sigma_{j}^{-1})\mathbf{X} + \boldsymbol{\mu}_{i}^{T}\Sigma_{i}^{-1}\boldsymbol{\mu}_{i}$$
$$-\boldsymbol{\mu}_{j}^{T}\Sigma_{j}^{-1}\boldsymbol{\mu}_{j} - 2\mathbf{X}^{T}(\Sigma_{i}^{-1}\boldsymbol{\mu}_{i} - \Sigma_{j}^{-1}\boldsymbol{\mu}_{j})] = 0$$

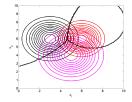
$$log \frac{P(y=i)}{P(y=j)} - \frac{1}{2}log \frac{\left|\Sigma_{i}\right|}{\left|\Sigma_{j}\right|} - \frac{1}{2}[\mathbf{X}^{T}(\Sigma_{i}^{-1} - \Sigma_{j}^{-1})\mathbf{X} + \boldsymbol{\mu}_{i}^{T}\Sigma_{i}^{-1}\boldsymbol{\mu}_{i}$$
$$-\boldsymbol{\mu}_{j}^{T}\Sigma_{j}^{-1}\boldsymbol{\mu}_{j} - 2\mathbf{X}^{T}(\Sigma_{i}^{-1}\boldsymbol{\mu}_{i} - \Sigma_{j}^{-1}\boldsymbol{\mu}_{j})] = 0$$

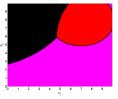
$$\Rightarrow \mathbf{X}^T a \mathbf{X} + b^T \mathbf{X} + c = 0$$

Analysis of QDA when $\Sigma = \sigma_i^2 I$

- Classes are of different distributions and different means.
- Cross-sections of classes distribution are spherical but of different sizes.



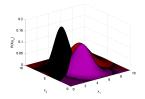


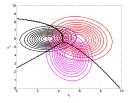


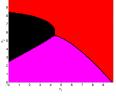
$$\Sigma_1 = \begin{bmatrix} 1.5 & 0 \\ 0 & 1.5 \end{bmatrix}, \ \Sigma_2 = \begin{bmatrix} 2 & 0 \\ 0 & 2 \end{bmatrix}, \ \Sigma_3 = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

Analysis of QDA with arbitrary $\Sigma_i \neq \Sigma_j$

- Classes are of different distributions and different means.
- Cross-sections of classes distribution are ellipsoidal and of different sizes.







$$\Sigma_1 = \begin{bmatrix} 1.5 & 0.1 \\ 0.1 & 0.5 \end{bmatrix}, \ \Sigma_2 = \begin{bmatrix} 1 & -0.2 \\ -0.2 & 2 \end{bmatrix}, \ \Sigma_3 = \begin{bmatrix} 2 & -0.25 \\ -0.25 & 1.5 \end{bmatrix}$$

GLAD and QDA, Another point of view

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Naive Bayes: An Example

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