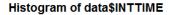
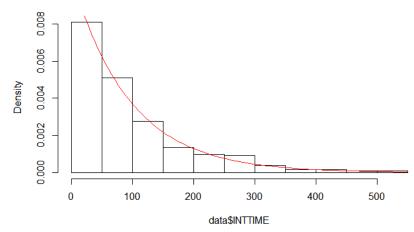
Q #1)

a) 
$$P(y > 2) = 1 - P(y < 2) = 1 - F(2)$$
  
 $F(y) = \int_{-\infty}^{+\infty} \frac{e^{-y/\beta}}{\beta} dy = -e^{-y/\beta} \Big]_{0}^{y} = 1 - e^{-y/\beta}$   
 $p(y > 2) = 1 - \left(1 - e^{-120/95}\right) = 0.2828$ 

b) The mean value and standard deviation for data are 95.52 and 91.54 respectively. It is consistent with the exponential distribution with beta = 95. The following plot shows the data hypothetical exponential distribution is same frame:





Q #2)

a) 
$$\mu = \alpha\beta = (3)(0.07) = 0.21$$
  
 $\sigma^2 = \alpha\beta^2 = (3)(0.07)^2 = 0.0147$ 

b) 
$$\mu \pm 3\sigma = 0.21 \pm \sqrt{0.0147} = [0, 0.57]$$

The value of 0.6 is beyond the 3 times of sd. The possibility of happening is not very much. We can also calculate the probability for values equal to or greater than this value which is 0.0078.

Q #3)

a) 
$$\mu_A = \alpha_A \beta_A = (2)(2) = 4$$
  
 $\mu_B = \alpha_B \beta_B = (1)(4) = 4$   
b)  $\mu_A^2 = \alpha_A \beta_A^2 = (2)(2)^2 = 8$   
 $\mu_B^2 = \alpha_B \beta_B^2 = (1)(4)^2 = 16$   
c)  $P_A(y < 1) = 0.0902$ 

c) 
$$P_A(y < 1) = 0.0902$$
  
 $P_B(y < 1) = 0.2212$   
The gas B has higher probability.

```
Q #4)
```

```
a) > a = 2; b = 4
             > pweibull(2, shape = a, scale = b)
             [1] 0.2211992
         b) > Mu = b^{(1/a)} * gamma((a+1)/a)
             > Var = b^{2/a} *(gamma((a+2)/a) - gamma((a+1)/a)^2)
             > Mu; Var
             [1] 1.772454
             [1] 0.8584073
         c) > SD = sqrt(Var)
             > Mu-2*SD
             [1] -0.08055165
             > Mu+2*SD
             [1] 3.625459
             > pweibull(Mu+2*SD, shape = a, scale = b) - pweibull(Mu-2*SD, shape = a, scale = b)
             [1] 0.5602273
         d) > 1 - pweibull(6, shape = a, scale = 2)
             [1] 0.0001234098
Q #5)
         a) > a = 2; b = 9
             > Mu = a / (a + b)
             > Var = a * b / ((a + b)^2 * (a + b + 1))
             > Mu; Var
             [1] 0.1818182
             [1] 0.01239669
         b) > 1 - pbeta(0.4, shape1 = a, shape2 = b)
             [1] 0.0463574
         c) > pbeta(0.1, shape1 = a, shape2 = b)
             [1] 0.2639011
```

Q #6)

a) Comparing the given distribution with Weibull distribution, alpha is 2 and beta is 16.

```
b) > a = 2; b = 16

> Mu = b^(1/a) * gamma((a+1)/a)

> Var = b^(2/a) *(gamma((a+2)/a) - (gamma((a+1)/a))^2)
```

- [1] 3.544908
- [1] 3.433629
- c) > 1 pweibull(6, shape = a, scale = b) [1] 0.8688151

Q #7)

a) 
$$p(x,y) = \frac{1}{6} \times \frac{1}{6} = \frac{1}{36}$$

b) 
$$p_1(x) = \sum_{y=1}^6 p(x, y) = 6 \times \frac{1}{36} = \frac{1}{6}$$

a) 
$$p(x,y) = \frac{1}{6} \times \frac{1}{6} = \frac{1}{36}$$
  
b)  $p_1(x) = \sum_{y=1}^6 p(x,y) = 6 \times \frac{1}{36} = \frac{1}{6}$   
 $p_2(y) = \sum_{x=1}^6 p(x,y) = 6 \times \frac{1}{36} = \frac{1}{6}$ 

c) 
$$p_1(x|y) = \frac{p(x,y)}{p_2(y)} = \frac{\frac{1}{36}}{\frac{1}{6}} = \frac{1}{6}$$

$$p_2(y|x) = \frac{p(x,y)}{p_1(x)} = \frac{\frac{1}{36}}{\frac{1}{6}} = \frac{1}{6}$$

d) 
$$p(x,y) = p_1(x)p_2(y)$$

Q #8)

		Х			P(y)
		1	2	3	
У	1	1/7	2/7	1/7	4/7
	2	0	0	2/7	2/7
	3	0	0	1/7	1/7
P(x)		1/7	2/7	4/7	1

- a) The table
- b)

c)

P(y)
4/7
2/7
1/7

d) 
$$P(y|1) = 1$$
,  $p(y|2) = 1$ ,  $p(y|3) = (1/4,2/4,1/4)$ 

Q #9)

a) 
$$f(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \int_{y}^{2y} \frac{e^{-\frac{y}{10}}}{10y} dx = \frac{e^{-\frac{y}{10}}}{10y} x \bigg|_{x=y}^{x=2y} = \frac{e^{-\frac{y}{10}}}{10}$$

The distribution is exponential distribution.

b) 
$$E(y) = \int_{-\infty}^{+\infty} f(x, y) y dy = \int_{0}^{+\infty} \frac{e^{-\frac{y}{10}}}{10y} y dy = -e^{-\frac{y}{10}} \Big|_{y=0}^{y=\infty} = 1$$

Q #10)

a) 
$$F(y) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f(x, y) dx dy = 1$$

$$\int \int \int \int ce^{-x^2} dy dx$$

$$= \int \int \int ce^{-x^2} dy dx$$

$$= \int \int ce^{-x^2} y \Big|_{y=0}^{y=x} dx$$

$$= \int \int cx e^{-x^2} dx$$

$$= -\frac{c}{2} e^{-x^2} \Big|_{x=0}^{x=\infty}$$

$$= \frac{c}{2} = 1$$

$$c = 2$$

b) 
$$f_1(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \int_0^x 2e^{-x^2} dy = 2ye^{-x^2} \Big]_{y=0}^{y=x} = 2xe^{-x^2}$$

$$\int_{-\infty}^{+\infty} f_1(x) dx = \int_0^{+\infty} 2xe^{-x^2} dx = -e^{-x^2} \Big]_{x=0}^{x=\infty} = 1$$

c) 
$$f_2(y|x) = \frac{f(x,y)}{f_1(x)} = \frac{2e^{-x^2}}{2xe^{-x^2}} = \frac{1}{x}$$

$$E(XY) = \sum xyp(x,y) = (-1)(-1)\left(\frac{1}{12}\right) + (-1)(1)\left(\frac{1}{12}\right) + (1)(-1)\left(\frac{1}{12}\right) + (1)(1)\left(\frac{1}{12}\right) = 0$$

$$E(X) = \sum xp_1(x) = (-1)\left(\frac{4}{12}\right) + (1)\left(\frac{4}{12}\right) = 0$$

$$E(Y) = \sum y p_2(y) = (-1) \left(\frac{4}{12}\right) + (1) \left(\frac{4}{12}\right) = 0$$

$$cov(X,Y) = E(XY) - E(X)E(Y) = 0$$

They are not independent. For example:

$$p(0,0) = 0 \neq p_1(0) \times p_2(0) = \frac{4}{12} \times \frac{4}{12}$$

Q #12)

- b) > Var [1] 0.005555556
- c) Base on CLT, the distribution for sample is a normal distribution.

Q #13)

```
b) > 1 - pnorm(11 - 0.5, mean = Mu, sd = SD)
    [1] 0.1269165
c) > pnorm(1 + .5, mean = Mu, sd = SD)
    [1] 0.001504434
    > pbinom(1, size = 20, prob = .40)
    [1] 0.0005240494
    > 1 - pnorm(11 - 0.5, mean = Mu, sd = SD)
    [1] 0.1269165
    > 1 - pbinom(10, size = 20, prob = .40)
    [1] 0.1275212
a) > Lead = c(1.32, 0, 13.1, .919, .657, 3.0, 1.32, 4.09, 4.45, 0)
    > Copper = c(.508, .279, .320, .904, .221, .283, .475, .130, .220, .743)
    > t.test(Lead, conf.level = 0.99)$conf.int
    [1] -1.147845 6.919045
    attr(,"conf.level")
    [1] 0.99
b) > t.test(Copper, conf.level = 0.99)$conf.int
    [1] 0.1518726 0.6647274
    attr(,"conf.level")
    [1] 0.99
c) The analysis shows that if we take a sample with probability of 99%, the lead content of water is les
    s than 6.92ppm and the Copper content is between 0.152 and 0.665 ppm.
d) 99% confident level means the if we take sample from the population, 99 sample out of 100 would
    have a value in interval range.
    Sdfjfls
    > data = read.csv("SOLARAD.csv")
    > IW = data$IOWA
    > ST = data$STJOS
    > diff = ST - IW
    > t.test(diff, conf.level = 0.95)$conf.int
    [1] 156.8193 239.1807
    attr(,"conf.level")
    [1] 0.95
```

Q #14)

Q #15)

Q #16)

a) > data = read.csv("DIAZINON.csv")

```
> Day = data$DAY
> Night = data$NIGHT
> diff = Day - Night
> t.test(diff, conf.level = 0.90)$conf.int
[1] -58.89917 -18.91901
attr(,"conf.level")
[1] 0.9
```

- b) The sample should be taken independently.
- c) From part a we infer that the mean values for day and night samples are not same. We can say wit h 90% confident level that the residual at night is higher than day's at least 18.9 and at most 58.9.