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On the use of genome-wide data to model and date the time of anthropogenic
hybridisation: an example from the Scottish wildcat

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Running title: Modelling hybridisation in Scottish wildcats

Abstract

While hybridisation has long been recognised as an important natural phenomenon in evolution, the conservation of taxa subject to introgressive hybridisation from domesticated forms is a subject of intense debate. Hybridisation of the Scottish wildcat, the UK's sole extant native felid, with the domestic cat is a good example in this regard. We develop a modelling framework to determine the timescale and mode of introgression using approximate Bayesian computation (ABC). Applying the model to ddRAD-seq data from 129 individuals, genotyped at 6,546 loci, we show that a population of wildcats genetically distant from domestic cats is still present in Scotland, though these individuals are found almost exclusively within the captive breeding program. Most wild-living cats sampled were introgressed to some extent. Additionally, we evaluate the effectiveness of current methods that are used to classify hybrids. We show that an optimised 35 SNP panel is a better predictor of the ddRAD-based hybrid score in comparison with a morphological method.

Keywords: hybridisation, wildcat, admixture, approximate Bayesian computation, introgression

Introduction

Hybridisation and introgression are important drivers of evolutionary change (Barton, 2001). Human-mediated hybridisation, however, is of increasing concern in conservation biology (Allendorf, Leary, Spruell, & Wenburg, 2001). Evolutionary processes may be disrupted by human activity, particularly when species distributions are altered by, for example, climate change, landscape use, or introduction of non-native species, leading to contact between populations that were originally allopatric. Whilst it is recognised this can generate a range of outcomes, some of which may be positive (e.g. genetic rescue; Johnson et al., 2010 or adaptive introgression; Pardo-Diaz et al., 2012), hybridisation and introgression are often considered threats to wild populations (Rhymer & Simberloff, 1996; Todesco et al., 2016). Loss of locally adaptive variation, reduction in fitness,

25 outbreeding depression or genetic swamping can all result in population or species extinction.
26 Furthermore, introgressive hybridisation between domesticated species and wild populations
27 increases the spread of potentially maladaptive, artificially selected variants in the wild (Randi,
28 2008).

29 The wildcat population in Scotland is an example of the threat of genetic extinction as a
30 result of hybridisation (Mathews et al., 2018). The wildcat, *Felis silvestris*, is Britain's most
31 endangered carnivoran and last remaining wild felid species. Wildcats have faced a long history of
32 persecution and habitat loss and can hybridise with domestic cats to produce fertile offspring.
33 Introgressive hybridisation is an increasingly serious threat to the dwindling population of this
34 species in the Britain, which is now at risk of complete genetic replacement by hybrids in the wild
35 (Breitenmoser, Lanz, & Breitenmoser-Würsten, 2019). Hybrids and feral domestic cats also compete
36 with wildcats for habitat and resources and pose a disease transmission risk.

37 Modern domestic cats are derived from the Near Eastern wildcat species *Felis lybica*. The
38 process of cat domestication was likely initiated as a result of their attraction to rodents, who
39 themselves were attracted to grain stores associated with settled agriculture ~9,500 years ago
40 (Driscoll et al., 2007). Though Driscoll *et al.* (2007) described just one wildcat species, *Felis silvestris*,
41 distributed across Europe, Asia, and Africa, a recently revised Felidae taxonomy recognises two
42 species of wildcat, *Felis silvestris* present in Europe, Caucasus and Turkey, and *Felis lybica* distributed
43 in Africa and Asia (Kitchener et al., 2017).

44 Artificial selection has altered the morphology, behaviour, and rate of reproduction of
45 domestic cats (Driscoll, Macdonald, & O'Brien, 2009). As a result, they are sufficiently diverged from
46 wildcats to be considered a separate species, *Felis catus* (International Commission on Zoological
47 Nomenclature, 2003). Domestic cats are widespread globally and found throughout the *Felis*
48 *silvestris* range. Hybridisation between domestic cats and wildcats is variable across the wildcat

range in Europe (Yamaguchi, Kitchener, Driscoll, & Nussberger, 2015) and is particularly acute in Scotland for reasons that remain poorly understood.

The remaining Scottish wildcat population is believed to be small, whereas hybrid cats are prevalent in certain areas; in a 2017/18 survey of wildcat conservation “Priority Areas” (Littlewood et al., 2014) the ratio of un-neutered hybrids to wildcats was estimated at 6:1 (Breitenmoser et al., 2019). The wild-living population in Scotland now resembles a ‘hybrid swarm’ - a continuum of genetic backgrounds as a result of repeated back-crossing and mating between hybrids (Beaumont et al., 2001; Senn et al., 2019). A recent review of wildcat conservation in Britain by the IUCN concluded the population was “too small, with hybridisation too far advanced and the population too fragmented” to be considered viable (Breitenmoser et al., 2019).

Introgressive hybridisation, by definition, results in the movement of genes between species. However, the consequences of the introduction of domestic cat genes into wildcat populations, or the fitness of hybrid offspring, is poorly understood. It is unknown whether introduced domestic cat genes confer any selective advantage or disadvantage in hybrid populations. This is especially interesting in the context of a changing environment for wildcats, specifically habitat loss or change, and increased competition with, and spread of diseases from, feral domestic cats (Breitenmoser et al., 2019).

Methods to detect signals of natural selection commonly rely on identifying large differences in allele frequencies between populations (Lewontin & Krakauer, 1973). This is challenging for genetically continuous populations, such as the hybrid swarm observed in Scottish wildcats (Waples & Gaggiotti, 2006). Here we apply the tool *pcadapt* to perform scans for selection (Luu, Bazin, & Blum, 2017). *Pcadapt* uses a PCA-based approach to detect variants which are outliers with respect to population structure; it is especially appropriate for admixed individuals as it does not require population information *a priori*.

Uncertainty also surrounds the temporal patterns of hybridisation in Scotland. Domestic cats are thought to have become widespread during the Roman occupation of Britain ~2,000 years ago (Serpell, 2014), though cat remains have been found at Iron Age sites, including sites on the Orkney islands off the north coast of Scotland (Macdonald et al., 2010; Smith, 1994). The wildcat population dramatically declined during the 18th and 19th centuries due to hunting and habitat loss, and by the start of the 20th century wildcat range in the UK was limited to north-west Scotland. Significant introgression is believed to have occurred within the last 100 years, when the wildcat population expanded, increasing contact between the small remaining population of wildcats and domestic cats (Breitenmoser et al., 2019). Historic samples, collected over the last c. 100 years, support an acceleration of hybridisation in Scotland over this period (Senn et al., 2019).

Without a comprehensive understanding of hybridisation history or dynamics, or the impact of introgressive hybridisation on fitness, conservation of this species in Britain is not straightforward. Accurate population estimates are difficult to obtain due to the elusive nature of the species and limited ability to distinguish hybrids in the field based on morphology (Breitenmoser et al., 2019). This problem is compounded by the lack of a baseline reference for Scottish wildcats. The difficulties inherent in distinguishing wildcat and hybrid phenotypes results in haphazard protection, impedes accurate monitoring, and undermines the Scottish wildcat's legal status as a protected species.

The Scottish wildcat has served as a canonical example of domestic-wild hybridisation more generally. The aim of this study is, firstly, to clarify the population structure of wildcats in Scotland using a two-fold increase in the number of genetic markers compared to the most recent study (Senn et al., 2019). For this we use ddRAD-seq data; ddRAD-seq is an efficient way to sample thousands of markers for genome-wide estimates of hybridisation (Peterson, Weber, Kay, Fisher, & Hoekstra, 2012). Increasing the number of markers increases power to accurately identify complex hybrids and backcrosses (Boecklen & Howard, 1997), giving the greatest resolution to date of the hybrid swarm in Scotland.

Secondly, we use the expanded set of markers to evaluate the effectiveness of current tests to identify hybrid individuals. Finally, we obtain an estimate of the timescale of hybridisation using a model that predicts the observed pattern of population structure. A demographic model for Scottish wildcats was developed using an Approximate Bayesian Computational (ABC) framework (Beaumont, Zhang, & Balding, 2002), a model-based approach to parameter inference rooted in Bayesian statistics. We also apply the model to evaluating the performance of PCA-based methods to identify genes that are subject to natural selection in structured populations.

Methods

Data processing

ddRAD-seq data were generated for 129 individuals sampled between 1996 and 2017 (Senn et al., 2019). This included 71 individuals from the UK captive wildcat population (all sampled in 2017), 53 individuals from the wild in Scotland (22 Scottish Wildcat Action www.scottishwildcataction.org trapped cats, 31 roadkill samples) and five Scottish domestic cats, for full sample details see Supp. Table 1. Note that historical wildcat samples derived from museum specimens reported in Senn et al. (2019) could not be used for this study due to poor DNA quality.

Sequence reads were aligned using BWA (Li & Durbin, 2009) to the *Felis catus* reference genome v9.0 (GCF_000181335.3) (Pontius et al., 2007). Mapped reads were processed using STACKS v2.1 (Catchen, Hohenlohe, Bassham, Amores, & Cresko, 2013). In STACKS a minimum of three reads were required to form a 'stack'. Resulting variants were filtered using a minimum allele frequency of 0.05 and a maximum proportion of heterozygous individuals of 0.7, treating the three sample sources (domestic, wild-living, and captive) as separate populations.

PLINK v1.9 (Chang et al., 2015) and VCFtools v1.15 (Danecek et al., 2011) were used to filter the data from STACKS. Specifically, this led to the removal of individuals with >30% missing data and

stringent subsequent filtering of loci to remove all sites with missing data. Closely related individuals were identified using IBD estimates calculated by PLINK, corrected to account for admixture using the method described by Morrison (2013). Corrected IBD estimates were used as input for PRIMUS (Staples et al., 2014) which uses genetic data to reconstruct pedigrees up to third degree relatives. Individuals were then removed from the dataset to limit relatedness (for the full list of excluded individuals see Supp. Table 1). Population genetic summary statistics (observed and expected heterozygosity, inbreeding coefficient and pairwise F_{ST} ; Weir & Cockerham, 1984) were generated for the final dataset using PLINK and VCFtools.

Population structure

Principal component analysis (PCA) and ADMIXTURE (Alexander, Novembre, & Lange, 2009) were used to examine population structure. PCA was completed in R using *prcomp*. ADMIXTURE analyses were performed for seven values of K, ranging from two to eight, and included a calculation of cross-validation error to estimate the optimal value of K. All SNPs were included, the data were not considered dense enough to require thinning of markers (to minimise background linkage disequilibrium) prior to the analysis (Alexander et al., 2009).

Existing hybrid tests

Hybrid individuals are currently identified using a combination of genetic and morphological diagnostic tests: a seven-point pelage scoring system (Kitchener, Yamaguchi, Ward, & Macdonald, 2005) and a 35 SNP genetic test (Senn & Ogden, 2015). The pelage test (7PS) scores seven key morphological characteristics on an ordinal scale of 1,2,3 for domestic, hybrid or wildcat features, respectively. Putative wildcats score 19 or higher on this test (maximum score 21), though a lower threshold of 17 can be used to overcome possible recorder error, e.g., from poor quality camera-trap photos. The genetic test uses 35 SNPs that differentiate between wildcats and domestic cats (Nussberger, Greminger, Grossen, Keller, & Wandeler, 2013; Senn & Ogden, 2015). A 'hybrid score' is generated using STRUCTURE Q values between 0 and 1 (Pritchard, Stephens, & Donnelly, 2000);

higher values correspond to individuals with more wildcat ancestry. An LBQ score (i.e. the lower boundary of the Q value 90% CI) of 0.75 is proposed as the threshold to class individuals as putative wildcats, as distinct from hybrids (Senn & Ogden, 2015). Individuals with an $LBQ \geq 0.75$ are currently considered wildcats from a conservation management perspective.

We compared the performance of these hybrid tests using ADMIXTURE Q values from the ddRAD-seq data (6,546 SNPs) to determine hybrid status. None of the 35 SNPs from the genetic test were present in the ddRAD-seq data. Data were only included from individuals where both 35 SNP and pelage scores were available ($n=59$). The aim of this analysis was to compare the performance of these tests with diagnoses from a relatively dense marker set. Given the continuum of Q values observed in wild-living cats, a strict threshold ($Q \geq 0.9$) was used to select reference wildcat samples, but we recognise this threshold is somewhat arbitrary and does not necessarily denote 'true wildcat' status. Individuals with an ADMIXTURE Q score of 0.9 or more were classified as wildcat reference samples, and those below 0.9 as hybrids. Receiver operating characteristic (ROC) curves were then constructed to assess performance (Robin et al., 2011). Given the reference diagnosis, the true positive and false positive rates were calculated for both diagnostic tests at all possible threshold values. Plotting false positive rate against true positive rate (specificity vs sensitivity) for each classification threshold generated an ROC curve for each test. The area under the curve (AUC) is equivalent to the probability a test will rank a random positive instance higher than a random negative instance and is a useful metric to compare diagnostic tests. An AUC of 0.5 is essentially a random guess and an AUC of less than 0.5 is worse than random.

Outlier analysis

The data were screened for outliers using the R package *pcadapt* (Luu et al., 2017). The first three principal components were used in the analysis, following Cattell's Rule that eigenvalues relating to random variation lie on a straight line, and those relating to population structure depart from the

line (Cattell 1966). To reduce false positives, p-values $< 1 \times 10^{-6}$ were investigated as outliers (equivalent to 0.01 Bonferroni corrected).

To better understand the false positive rate of the outlying SNPs, simulated data (generated using a neutral model of evolution, described below) were also analysed using the same method in *pcadapt*. Ten simulated datasets were generated using a random sample of parameters values from the ABC posterior distribution (see below).

Demographic modelling

A demographic model for wildcats was developed within an ABC framework (Beaumont et al., 2002). ABC was developed as a rejection algorithm (Pritchard, Seielstad, Perez-Lezaun, & Feldman, 1999), in which simulated data are generated under a hypothesised model of evolution, with model parameters sampled from a known prior distribution. Summary statistics are taken from both the simulated data and observed data. An accepted sample of simulations (those with summaries closest to the observed data) are then used to estimate posterior distributions of the model parameters. Posterior estimates from the basic rejection algorithm can be improved with local linear (Beaumont et al., 2002) or non-linear regression (Blum & François, 2010).

Fig. 4A outlines the model developed for wildcat demography. Wildcat and domestic cat populations diverge, under a neutral model of evolution, for 500 generations. Generation time for a wildcat is estimated to be three years (Beaumont et al., 2001; Nussberger, Currat, Quilodran, Ponta, & Keller, 2018). The divergence of the two populations from a common ancestor is modelled using a computationally efficient two-stage approach; firstly, starting SNP frequencies for each population were simulated from a beta-binomial distribution, parameterised by F_{ST} (Balding & Nichols, 1995). These initialise an individual-based model of genetic inheritance in which at time T_1 gene-flow from domestics begins at a rate of mig_1 per generation. Gene-flow occurs at the same rate in every subsequent generation. At time T_2 the captive wildcat population is established from a random sample of wildcat individuals (referred to as the wild-living population from this point forward).

There is (limited) gene-flow (m_{ig_2}) from the wild-living population to the captive wildcats (reflecting a number of wild-caught founders that have been incorporated into the captive population since it was established). Population sizes remain constant throughout the simulation; we do not model any fluctuations in wildcat population size (e.g., recent population expansion), and we do not model a decline in the wildcat population as a direct result of hybridisation. Furthermore, unlike Quilodr  n *et al.* (2020), we do not consider a spatial model for hybridisation. Previous analysis indicates a complex and patchy pattern of hybridisation, difficult to model on a large scale (Kilshaw *et al.*, 2016; Senn *et al.*, 2019).

Data were simulated under this model using SLiM (Haller & Messer 2017), a toolkit for evolutionary modelling. SLiM is individual-based, forward-simulating and, implements a Wright-Fisher model of evolution (amongst others) in which generations are non-overlapping, individuals are diploid, and offspring are generated through recombination and mutation of parental genotypes. 15,000 independent sites were modelled per individual (to replicate the observed SNP data from ddRAD-seq). After 500 generations the genotypes of 46 captive wildcats, 45 wild-living and four domestic cats were sampled at random, and summary statistics were calculated in R. Captive individuals with a Q35 score of <0.9 ($n=13$) were filtered from the observed data. This functioned as a proxy for the selection of putative wildcats for incorporation into the captive breeding programme, in the model migrants are selected at random. The total number of simulations used for ABC was 509,070.

Prior distributions were chosen based on existing knowledge of the model system (for details see Supp. Fig. 11). A wide prior was chosen for T_1 , allowing hybridisation to begin at any point in the simulation. A more informative prior was given to T_2 as we know the captive population was established in 1960.

Given the strong separation of domestic cats and wildcats across the first principal component (Fig. 1A), a set of PCA-based summaries was devised (measures of the distribution of

points across PC1 and PC2). Additional summaries included pairwise genetic distance (F_{ST}) and linkage disequilibrium measures, for full list see Supp. Table 2. The total number of summary statistics was 14. Owing to the correlation within and between parameters and summary statistics (Supp. Fig. 8), projection was used to reduce dimensionality and improve posterior estimates, following the approach of Fearnhead and Prangle (2012). Projection involves fitting a regression model between each parameter and the summary statistics. The regression model gives an estimate of the posterior mean for a given set of summary statistics. This prediction for each parameter can be viewed as a projection of the 14-dimensional summary statistics onto a 10-dimensional set of new summary statistics (Blum, Nunes, Prangle, & Sisson, 2013). To fit the regression model for the projection we chose 20% of simulated points that were closest to the observed set of summary statistics.

The final model parameters and summary statistics were decided via the process described in Supp. Figs. 5-7, which used goodness-of-fit test included in the R package *abc* (Csilléry, François, & Blum, 2012) and a novel method for dropping summary statistics (described in Supp. Box 1).

Parameter inference was carried out in R using the package *abc* (Csilléry et al., 2012). The closest 5,091 points (1%) were used to generate the posterior distributions, correcting for an imperfect match between the summary statistics and observed data using non-linear regression (neural network) (Blum et al., 2013; Raynal et al., 2019).

Results

The final dataset included 108 individuals: four Scottish domestic cats and 104 putative wildcats (45 wild individuals and 59 from the UK captive population), genotyped at 6,546 SNPs. 21 samples were excluded from the analysis to minimise relatedness in the dataset and/or as a result of stringent filtering of missing data. Population summary statistics are given in Table 1.

Population structure

Principal component analysis (Fig. 1A) showed a large proportion of the genotypic variation (23.9%) was explained by the first principal component (PC1). PC1 supports strong differentiation between domestic cats and a group of almost exclusively captive individuals, only two wild-living individuals are found at similarly extreme PC1 values. A large F_{ST} (0.446, Table 1) is observed between domestic cats and the captive wildcat population. The distinct PCA clustering and high F_{ST} values supports this as a cluster of putative wildcats. Most wild-living individuals are distributed across PC1, between these two groups, and are therefore considered putative hybrids. A much smaller proportion of the variance is explained by PC2 (2.8%) and PC3 (2.7%, Supp. Fig. 1).

An ADMIXTURE model with two ancestral populations (Fig. 1C, $K=2$) also supported distinct clustering of domestic cats and captive wildcats. The majority of wild individuals sampled had probable ancestry assigned to both groups, with varying amounts of 'domestic' ancestry. PC1 position was strongly correlated with ADMIXTURE Q values at $K=2$ (Spearman's $r = 0.998$, $p < 0.001$; Supp. Fig. 2). Fig. 1B shows sampling locations for the wild individuals (where available), coloured by ADMIXTURE proportions at $K=2$. Individuals with domestic ancestry appear geographically widespread, with no clear single point of introgression. At $K=3$ further clustering within the putative wildcats is observed, including within the captive population. Cross-validation error indicated the most likely value of K for the whole dataset is 5 (Supp. Fig. 4).

Existing hybrid tests

ROC curves showed that both diagnostic tests performed well, with AUC values of 0.984 and 0.854 (Fig. 2). The 35 SNP test ($LBQ \geq 0.75$) outperformed the morphology-based test, with a low rate of both false positives and false negatives. Using a threshold of 17 the 7PS test showed nine false negatives and six false positives for the individuals analysed (i.e., individuals with few wildcat markings or features, but a high proportion of probable wildcat ancestry, and vice versa). At the

higher threshold of 19 there was only one instance of a false positive, but 19 false negatives. The 35 SNP test showed two false negatives and four false positives.

Evidence for natural selection

Pcadapt found three outlying SNPs that were reported to be most correlated with PC1 (Fig. 3B, for details see Supp. Table 3). Fig. 3B shows the PCA plot for the first two principal components, as in Fig. 1, with individuals coloured by genotype at each of the three positions (i.e., heterozygous, homozygous for allele 1 or homozygous for allele 2). For each SNP there was a clear difference in allele frequency between the domestic cat and captive wildcat populations. Notably, wild-living individuals had a high frequency of the domestic-type allele at these loci. This pattern does not seem to be an artefact of captive breeding, for each SNP shown in Fig. 3B the 'domestic' allele is at low frequency in wild individuals at similar PC1 positions as captive individuals, and at least one of these individuals was homozygous for the wildcat-type allele.

The SNPs are located on three different chromosomes. At the corresponding positions in the domestic cat genome SNPs 5147 and 5885 are found within protein-coding regions. SNP 5147 is found within the *SLC31A2* gene (chromosome D4, $p = 1.991 \times 10^{-7}$). In humans and mice *SLC31A2* has been shown to have copper ion transmembrane transporter activity (Okazaki et al. 2002; van den Berghe et al. 2007). SNP 5885 (chromosome E3 $p = 1.794 \times 10^{-7}$) is found within *ITGAX*, *ITGAX* is predicted to encode integrin subunit alpha X, orthologues of which are found in many other mammals, including humans and mice. Integrins generally are adhesion receptors, linking the extracellular matrix and cell cytoskeleton (Schnapp et al. 1995). They also interact with growth factor receptors to promote cell cycle progression and cell migration. SNP 2022 (chromosome B2, $p = 1.403 \times 10^{-11}$) is located 383bp downstream from the *TRAM2* gene, which encodes translocation associated membrane protein 2. In humans, *TRAM2* has been identified to have roles in collagen synthesis, protein transport and protein insertion into the membrane of the endoplasmic reticulum (Stefanovic et al. 2004).

Outlier SNPs are candidates for loci under selection, though extreme outliers can also be generated via neutral processes. Fig. 3A shows that outlying SNPs were generated under a neutral model of wildcat demography, a result of pre-existing population structure, emphasised by genetic drift. Even using a conservative threshold to minimise the false discovery rate, nine out of the ten sets of simulated data contained at least one SNP found to be outlying with respect to population structure across PC1 (see Table 2).

Demographic modelling

Our demographic model is capable of simulating data within the range of the observed data and the model fits these data well (Supp. Figs. 9 & 10). The first two axes of the posterior predictive PCA plots (Fig. 4C) show broadly the same patterns as the observed data, particularly with respect to the distribution of wild-living individuals across PC1. Prior and posterior distributions for the three parameters of interest (T_1 , T_2 and mig_1) are shown in Fig 4B. The posterior mean for T_1 , the time of onset of gene flow from domestics to wildcats, is 3.3 generations (95% HPD: 1.21– 5.). For T_2 , the time the captive population was established, the mean is 19.3 generations (95% HPD: 9.4 – 30), respectively. Note that the estimate for T_1 is not constrained by the prior to any marked degree, whereas the historically informed prior for T_2 has a stronger effect. The migration rate of domestic cats into the wild-living population was estimated to be 0.13 (95% HPD: 0.076 – 0.19) i.e., for an individual selected at random from the wild-living population there is a 13% chance it is a domestic cat.

Discussion

Current status of the wildcat in Scotland

PCA and ADMIXTURE analysis (Fig. 1) demonstrated that a group of individuals genetically distinct from domestic cats (putative wildcats) persists in Scotland. Genetic differentiation between these

groups was supported by a high F_{ST} , as would be anticipated between two species (Hartl & Clark, 2007), and comparable to that between dogs and wolves (Cronin et al., 2015) or red and sika deer (McFarlane et al., 2020). This supports the findings of previous microsatellite (Beaumont et al., 2001) and SNP studies (Senn et al., 2019) that were able to differentiate between domestic cats and a group of putative wildcats in Scotland. Here we reanalyse the 76 samples used by Senn *et al.* (2019), with an additional 51 captive individuals and two additional wild individuals. We increase the resolution of this study with an additional 3,449 SNPs, and the data show the same broad patterns. Putative wildcats reported in this study were sampled almost exclusively from the UK captive population. Hybridisation in the wild appeared extensive. A continuum of genetic backgrounds is observed, the result of repeated hybridisation, backcrossing and mating between hybrids referred to as a 'hybrid swarm' (Mayr, 1963); almost all wild-living individuals sampled showed some evidence of introgression from domestic cats (Fig. 1). This supports the conclusion of Breitenmoser *et al.* (2019) that the wild population in Scotland is now too hybridised to be considered viable.

Demographic modelling supported a rapid emergence of the hybrid swarm effect and a recent crash in the Scottish wildcat population as a result of high geneflow from domestic cats. We take the generation time for wild-living cats to be around 3 years (Beaumont et al., 2001; Nussberger et al., 2018). The T_1 posterior mean (3.326 generations, or ~10 years) is implausibly recent, yet extensive model-checking (Fig. 4c, Supp. Figs. 5-10) suggests that the model generally fits well. The exact history of hybridisation in the Britain remains poorly understood (and is likely to show geographic variation) but hybridisation has been of increasing conservation concern since the 1980s (Hubbard et al., 1992, Kitchener et al. 1992, Easterbee et al. 1991) and is generally thought to be a consequence of wildcat range expansion in Scotland during the early 20th century coupled with continuing high levels of persecution, especially in eastern Scotland. This does not exclude the onset of significant introgression within the last few decades. Though no historical samples were included in this study, Senn *et al.* (2019) generated Q35 scores for 60 historic samples collected in Scotland

between 1895 and 1985. These are predominantly cats shot by gamekeepers and subsequently incorporated into museum collections, so there is potential bias towards individuals with wildcat features. Nonetheless, only five of the samples collected over this period are classified as hybrids, using the $LBQ < 0.75$ threshold, and one as a domestic cat. In another example of hybridising species, Galaverni *et al.* (2017) date recent admixture between wolves and dogs in Italy to the 1940s, peaking in the 1990s.

The wildcat model is limited, however, by the ability of unlinked SNPs to detect ancient or complex patterns of admixture. Results presented here suggest our model is unable to detect signals of admixture beyond 30 generations or in this case, c. 100 years (Supp. Fig. 10). Haplotype and linkage disequilibrium information (from sequence data) are needed for accurate dating of admixture events, especially to separate historical admixture from the very recent (Hellenthal *et al.*, 2014; Loh *et al.*, 2013); this work in whole genome sequenced individuals is now underway.

Mattucci *et al.* (2019) used SNP array data to date admixture in continental European wildcat populations. Individuals were sampled from all five main biogeographic groups: Iberia, Central Europe, Central Germany, Italy and the Dinaric Alps (Mattucci, Oliveira, Lyons, Alves, & Randi, 2016). The study found hybridisation across all populations, occurring between six and 22 generations before present. The most recent admixture time reported by this study was 3.15 generations (though this date depended on the approach used). Mattucci *et al.* (2019) reported admixture times for individuals previously classified as true wildcats using microsatellite data, highlighting the power of a sequence-based approach to detecting historic and/or complex patterns of admixture (Gärke *et al.*, 2012; Haas & Payseur, 2011).

A recent hybridisation time for Scottish wildcats only seems likely in the face of high geneflow from domestic cats. Our model estimates gene flow to be 13% (95% HPD: 7-19%). In comparison, Quilodrán *et al.* (2020), using a forward simulating approach to model introgression in

the Swiss Jura wildcat population, estimated the rate of introgression to be 6%. At this lower rate of introgression, it took 26 generations for the wildcat population to become 50% introgressed.

Quilodr  n *et al.* (2020) use a spatial model to quantify introgression. Although this would be challenging at the scale of the model presented here, especially considering the complex patterns of introgression observed in the wild (Fig. 1B), it may be helpful in a future study to apply the approach of Quilodr  n *et al.* (2020), in conjunction with parameter estimates from the current model, to focus on a geographical area of interest to better understand hybridisation dynamics in a priority area for conservation management.

Tentative evidence is presented here that the ‘hybrid swarm’ effect can develop rapidly following the breakdown of isolating mechanisms between two species, as has been observed in other hybridising species such as deer (Smith, Carden, Coad, Birkitt, & Pemberton, 2014), loaches (Kwan, Ko, & Won, 2014) and honey-bees (Pinto, Rubink, Patton, Coulson, & Johnston, 2005). Our results may also support a recent acceleration of hybridisation in Britain. Though it is difficult to conclude using the current model whether historical admixture has occurred (and to what extent), it is clear there has been significant recent introgression within the last few decades.

An important feature of the model is the captive wildcat population. There is significant interest surrounding this population, which comprises individuals that are among the last putative wildcats in Britain, and especially regarding its value to continuing conservation efforts. It is therefore important to understand the extent to which hybridisation has impacted this population. It is clear from Fig. 1 that hybrids are present, though the number appears to be low. From the ABC posterior distribution, T_2 (the time the captive population is established) occurs consistently before gene-flow from domestic cats begins (T_1). This suggests the formation of the captive population in the 1960s and 1970s may have occurred prior to significant recent admixture, and that this population is an important reservoir of wildcat genes in Britain (probably aided in recent years by accurate tests for hybrids, see below). How closely modern captive animals resemble the British

post-glacial population of wildcats, especially considering sympatry with domestic cats over the last 2000 years, remains to be determined.

Captive individuals have a wide distribution across PC2 and PC3 (though this explains only a small proportion of the variation in the genetic data, 2.8% and 2.7%, respectively), and ADMIXTURE plots show clustering within the captive population (Fig. 1C, K=3). The distribution of captive individuals across PC2 was a difficult feature to replicate in the model (Fig. 4C). It is hard to disentangle the impacts of maintaining a (historically small) captive breeding population, e.g. inbreeding, genetic drift, or adaption to captivity (Frankham, 2018; Woodworth, Montgomery, Briscoe, & Frankham, 2002), from genuine population structure. The presence of family groups was limited following the identification of close relatives using PRIMUS. However, estimates of relatedness are complicated by potential admixture (Morrison, 2013). Our results (Supp. Fig. 3) imply the distribution of individuals across PC2 or PC3 is not a gradient of inbreeding across the population.

Patterns relating to geographical origin in the wild samples were unclear due to the high levels of introgression (Fig. 1B). In terms of introgression it seems clear there have been multiple admixture events, possibly due to the pervasiveness of domestic cats in wildcat habitat in Scotland and continuing high levels of persecution that maintained wildcat populations at low levels (Kitchener & O'Connor 2010). The evidence presented here does not rule out that the observed clustering in the captive population reflects biogeographic structure in the Scottish wildcat population. The Great Glen, for example, has been suggested as a barrier to gene flow in the Scottish red deer population (Pérez-Espona *et al.* 2008). The Great Glen is a ~100km long valley, running along part of the Great Glen fault that bisects the Scottish Highlands. In red deer, strong population differentiation is observed between the eastern and western sides of the Great Glen, and it is possible that this is also a barrier to wildcat dispersal. However, wild-living individuals belonging

to a single cluster at $K=3$ were sampled from both sides of the Great Glen, so other geographical barriers may need to be considered and tested with additional sampling and modelling.

A second possibility is that ADMIXTURE clustering at values of K greater than two reflect temporal patterns of hybridisation, i.e., snapshots of the genetic composition of the wild-living population at various points since the mid-20th century (a number of wild founders have been incorporated into the captive population since it was founded in 1960). The value of K with the lowest cross-validation error was five, this may be an effect of trying to break a continuum of hybridisation levels into discrete units. It is interesting to note that captive individuals with probable domestic ancestry at $K=2$ all belong to the same cluster at $K=3$.

Mattucci *et al.* (2016) suggest that strong population structure within wildcats in mainland Europe (for example, between eastern and western Germany, Hertwig *et al.*, 2009) represents population expansion from five major mid-Pleistocene glacial refugia. Interestingly, PCA of the microsatellite data collected for this study shows a similar ‘anvil’ shape, with *Felis silvestris* more dispersed across PC2 than *Felis catus*. Population structure and expansion perhaps make this a feature of wildcat genetics more generally (especially when compared to inbred domestic cats), and we should avoid over-interpretation in the Scottish population (Lawson, van Dorp, & Falush, 2018).

Evidence for natural selection

The major application of outlier analyses is to detect loci under natural selection. There has been some debate in the literature as to whether RAD-seq data are appropriate for this kind of analysis (Catchen *et al.* 2017; Lowry *et al.* 2017; McKinney *et al.* 2017). Lowry *et al.* (2017) argue that the sparsity of RAD-seq markers misses many candidate loci, especially in species where linkage disequilibrium is low. This does not necessarily invalidate the small number of loci identified using RAD-seq, though it would be useful to confirm these findings with sequence data when possible.

Confounding effects, such as population structure and demography, are more problematic for this study. Even at neutral loci the demographic history of a population can cause allele frequency to vary hugely in space due to genetic drift and/or migration (Hoban et al. 2016). For populations that are highly differentiated the variance in F_{ST} among neutral loci is large. Differences in allele frequencies between domestic cats and wildcats are therefore not surprising considering the genetic differentiation between the two populations, and do not necessarily correspond to deviations from neutrality. Population expansion can also produce the same signal as selection due to ‘allele surfing’, where populations at the leading edge of an expansion are small, and contribute disproportionately to the expanding population, accelerating the effects of drift. As discussed above, the wildcat population in Scotland is thought to have been expanding since the early 20th century (Breitenmoser et al., 2019).

Here we have applied *pcadapt* to detect selection, which is designed to be robust to demographic biases and handle genetically continuous, admixed populations (Luu et al., 2017). However, simulation results, based on our best-fitting demographic model for the wildcats, show evidence of a high number of false-positives in this setting (Table 2), even using the most conservative approach to controlling false discovery rate. Although simulation-based tests using *pcadapt* have often shown that it performs well (Luu et al., 2017), scenarios with high recent admixture have not been investigated.

Based on this finding it is difficult to make conclusive statements about natural selection in Scottish wildcats, or fitness consequences for hybrid populations. Mattucci *et al.* (2019) reported a number of genomic regions in wildcat x domestic hybrids with a high frequency of either wildcat or domestic alleles, and genes within these regions were found to be significantly enriched for specific gene ontology categories. A striking feature of Fig. 3B is the similarity in allele frequencies between domestic and hybrids cats, even in less introgressed individuals, which perhaps constitutes tentative evidence for adaptive introgression in Scotland. Adaptive introgression has been shown to occur in

other wild populations which hybridise with domesticates, such as goats and sheep (Barbato et al., 2017; Grossen et al., 2014). The SNP correlated with PC1 with the most extreme p-value reported by *pcadapt* (Table 2, Supp. Table 3) is found in the domestic cat genome near the TRAM2 gene. TRAM2 has also been identified in genome scans for loci linked to the severity of leukaemia virus infection in cattle (Carignano et al., 2018). This finding highlights disease transmission as a potential driver of selection in hybrid populations. Both wildcat and domestic-like regions identified by Mattucci *et al.* (2019) included genes involved in the immune system or associated with diseases or infection, including feline leukaemia virus. Feline leukaemia virus is potentially fatal to both wildcats and domestic cats, and has similar prevalence (~10%) in both species in Scotland (Daniels et al. 1999).

Existing tests for hybrids

Accurately identifying hybrids in the field is crucial to effective conservation of the wildcat in Scotland. In the absence of uncontroversial reference samples, we have used a score based on 6,546 ddRAD SNPs and investigated the relative effectiveness of field-based tests in recovering this. An ROC analysis (Fig. 2) showed both diagnostic tests to be informative in identifying hybrid individuals as judged by scores from the ddRAD SNPs. The pelage score was a less reliable indicator of wildcat ancestry; this is unsurprising as the characteristics scored by this test are likely to be controlled by a limited number of genes (Cieslak, Reissmann, Hofreiter, & Ludwig, 2011; Eizirik et al., 2010), the transmission of which is still poorly understood. Devillard *et al.* (2014) and Kitchener *et al.* (2005) reported a greater degree of accuracy when using anatomical characteristics (skull size and shape and intestinal length) as opposed to than pelage in order to identify hybrids. Mattucci *et al.* (2019) found genomics regions in hybrid individuals with a high frequency of wildcat-type alleles contained (amongst others) genes relating to morphology. If selection is acting on key morphological features, as this result suggests, pelage may not give an accurate picture of hybridisation across the genome. Using a more lenient threshold (7PS ≥ 17 for putative wildcats) pelage scoring appeared to give a number of false negatives and false positives, i.e., individuals with

probable wildcat ancestry that did not necessarily score highly for wildcat features and vice versa. A more conservative threshold of $7PS \geq 19$ reduces the number of false positives but increases the false negative rate - a large number of individuals with high proportions of putatively wildcat ancestry are not classified as wildcats at this threshold.

We found the 35 SNP test to be a highly accurate predictor of the ddRAD SNP score; hybrids could be identified almost as well using the 35 SNPs as with a dense marker set of over 6000 SNPs. Four false positives and two false negatives were identified, though similar Q values were recovered using both marker sets for these individuals, so this may partly reflect the stringent threshold used to select reference wildcats from the ddRAD data.

Without accurate information on the history of hybridisation in Britain there is no uncontroversial baseline for Scottish wildcats with which to calibrate either diagnostic test. Therefore, we recommend the continued use of the pelage score and 35 SNP test in conjunction to identify hybrids, especially when considering individuals to be incorporated into the captive breeding programme.

Conclusion

We find a population of putative wildcats persists in Scotland. These individuals are almost exclusively found in the UK captive population, which appears to have been established prior to significant recent admixture and is supported by accurate tests for hybrids. It remains unclear to what extent historical admixture has affected the Scottish wildcat population, but divergence between domestic cats and putative wildcats remains high. The captive population is now an important resource for wildcat conservation in Britain. We find the wild-living population to be a hybrid swarm; almost all wild individuals sampled showed evidence of introgression from domestic cats. We predict a high rate of continuing gene-flow from domestic cats.

514

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Data Accessibility

All SNP data available from the Dryad Digital Repository [in progress]. Materials for demographic modelling available at [GitHub site, in progress]

Author Contributions

JHM designed the research, analysed the data, and wrote the paper. HS provided data for analysis. MB, DL and HS conceived the study and designed the research. DW and AK analysed the data. All authors critically reviewed the paper.

Tables and Figures

Table 1. Summary statistics for the three source populations: captive wildcats, wild individuals, and domestic cats. Weir & Cockerham (1984) estimates for population pairwise F_{ST} are shown on the right-hand side.

Summary	Population			Pairwise F_{ST}		
	Captive	Wild	Domestic		Captive	Wild
# Individuals	59	45	4	Captive		
# Loci	6546	6546	6546	Wild	0.130	
# Alleles	12258	13075	11448	Domestic	0.446	0.128
% missing data	0	0	0			
H_{Obs}	0.178	0.307	0.270			
H_{Exp}	0.285	0.285	0.285			
F	0.375	-0.077	0.055			

Table 2. *Pcadapt* using data simulated under a neutral model of evolution. The simulated data contain a number of outlying SNPs associated with PC1. For each of the 10 sets of simulated data the total number of SNPs is given, followed by the numbers of outlying SNPs associated with PC1 that are at least as small as the largest and smallest outlying p-values observed in the real data (unadjusted p-values). Following a Bonferroni correction (adjusted p-values), the number of outlying SNPs that were below a threshold of 0.01 is also reported.

Simulation No.	Total number of SNPs	Number of outlying SNPs associated with PC1		
		Unadjusted p-val $\leq 1.991 \times 10^{-7}$	Unadjusted p-val $\leq 1.403 \times 10^{-11}$	Adjusted p-val < 0.01
1	7492	8	0	14
2	6858	3	0	15
3	7542	0	0	2
4	7358	5	0	5
5	7101	17	1	24
6	8208	1	0	1
7	7286	0	0	1
8	7570	4	0	3
9	7296	0	0	0

10	7502	14	4	16
Total	74213	52	5	81

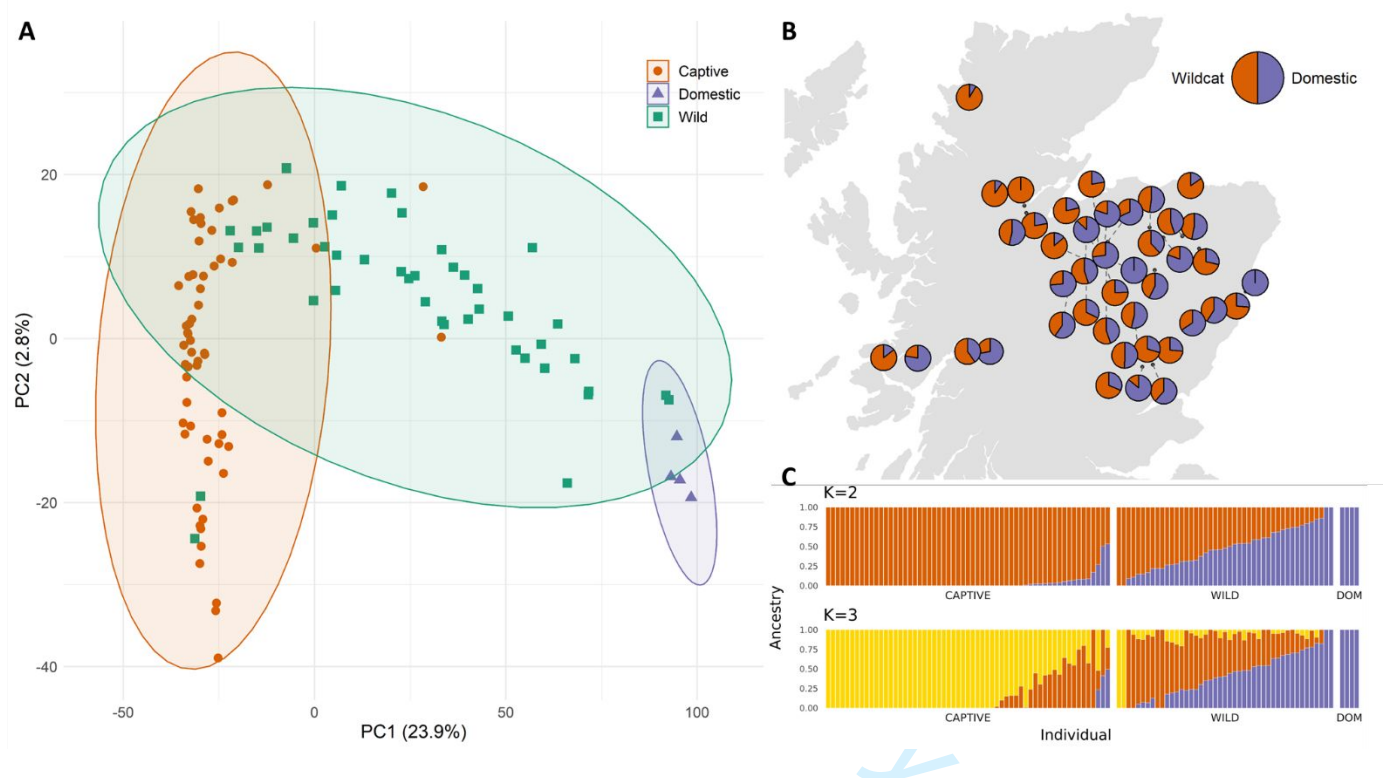


Figure 1. Population structure in the Scottish wildcat population. (A) Principal component analysis (PCA) shows a strong genetic differentiation between domestic cats and a group of putative wildcats across PC1. In the wild-living population a ‘hybrid swarm’ is observed, with a continuum of genetic backgrounds. (B) Sampling location of wild individuals (where known), pie charts show probable ancestry for each individual at K=2, as modelled using ADMIXTURE. (C) ADMIXTURE clustering (all individuals) at K=2 and K=3.

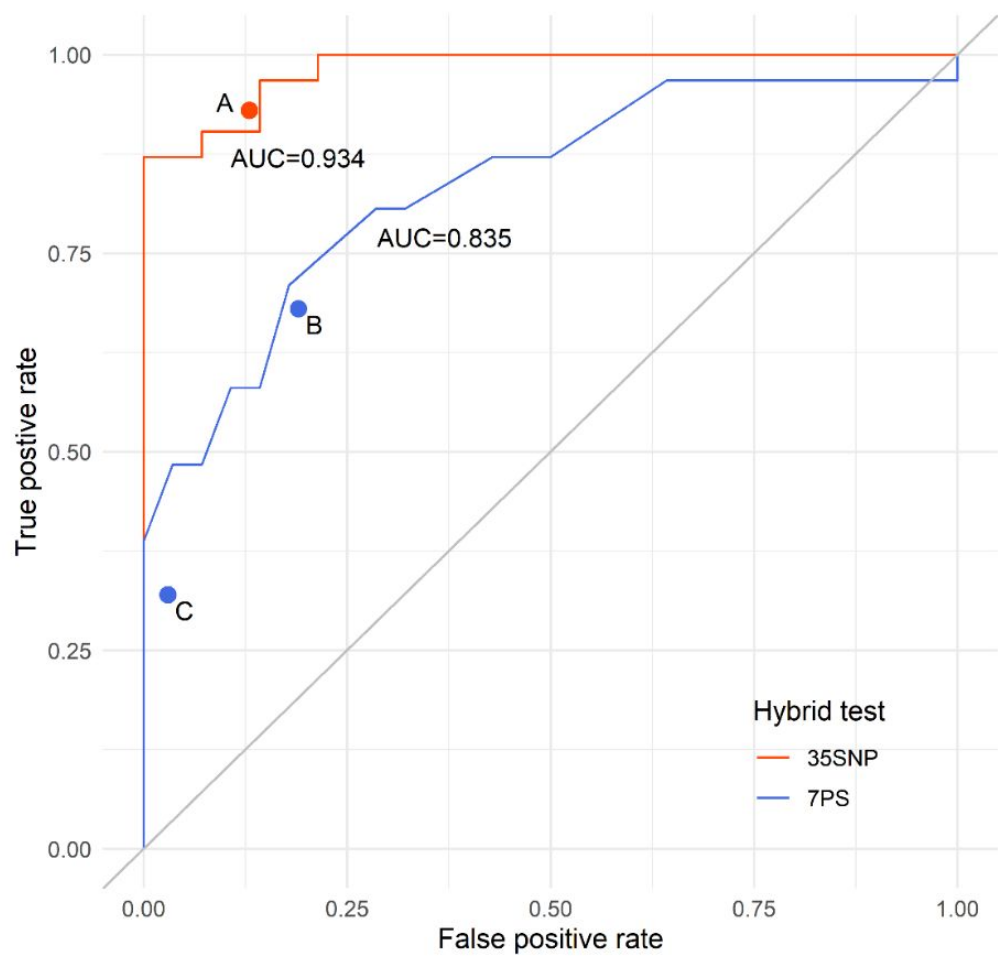


Figure 2. ROC curves for the current tests to identify wildcat/domestic hybrids: the 35 SNP genetic test (red) and seven-point pelage score (blue). True and false positive rates at the current thresholds for each test are shown using a point at the corresponding coordinate, (A) $LBQ \geq 0.75$, (B) $7PS \geq 17$, (C) $7PS \geq 19$.

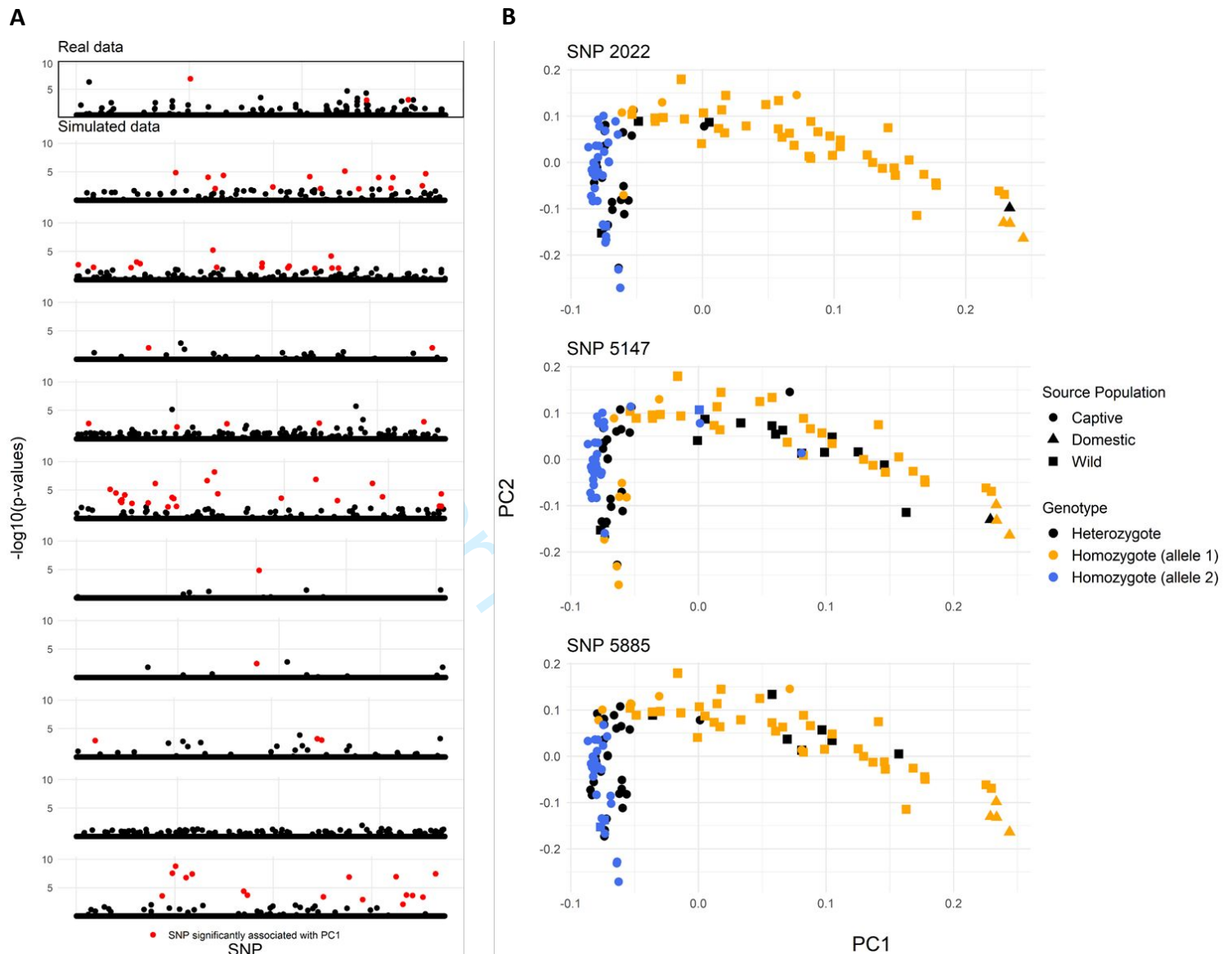


Figure 3. *Pcadapt* results for real and simulated data. (A) Manhattan plots for each set of SNPs analysed with *pcadapt*. The top row shows the real data, where these SNPs have been aligned to the domestic cat genome and are ordered by genomic position. The following rows are for simulated data. These data were simulated under a neutral model of evolution and generate a number of points classified as outliers by *pcadapt*. Red points correspond to outliers reported to be most correlated with PC1. (B) PCA plot coloured by genotype of the individual at each of the SNPs found to be significantly associated with PC1 in the real dataset.

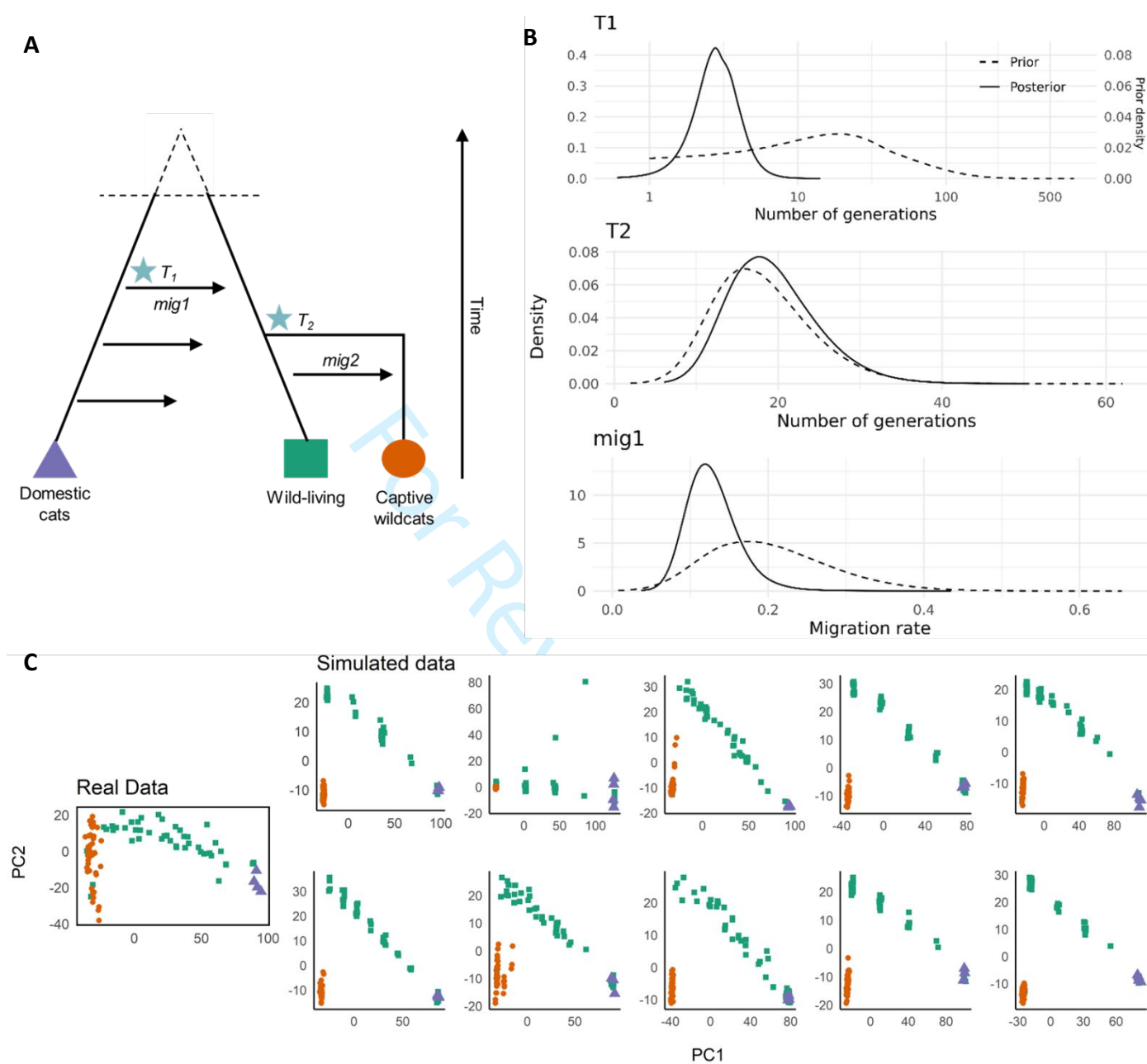


Figure 4. Modelling wildcat demography. (A) The model under which data were simulated; two parent populations (*F. catus* and *F. silvestris*) diverge under a neutral model of evolution. Gene-flow (introgression) from domestic cats begins at time T_1 , at a rate of mig_1 for every subsequent generation. At time T_2 the captive population is formed from a random sample of wild-living cats. Limited gene-flow from the wild population into the captive population occurs at a rate of mig_2 . (B) Prior and posterior distributions following ABC, dashed lines indicate the prior. Curves were fitted in R using *locfit* (Loader, 2013). The model supports recent introgression in the Scottish wildcat population following high gene-flow from domestics. (C) PCA plots for the real data (left) and for random sample of simulated data from the posterior distribution (right). The model is broadly able to simulate the same patterns as we observe in the real data.