



Using Machine Learning To Predict Antimicrobial MICs and Associated Genomic Features for Nontyphoidal *Salmonella*

Marcus Nguyen,^{a,b,*} S. Wesley Long,^{c,d} Patrick F. McDermott,^e Randall J. Olsen,^{c,d} Robert Olson,^{a,b,*} Rick L. Stevens,^{b,f,*} Gregory H. Tyson,^e Shaohua Zhao,^e James J. Davis^{a,b,*}

^aUniversity of Chicago Consortium for Advanced Science and Engineering, University of Chicago, Chicago, Illinois, USA

^bComputing, Environment and Life Sciences, Argonne National Laboratory, Argonne, Illinois, USA

^cCenter for Molecular and Translational Human Infectious Diseases Research, Department of Pathology and Genomic Medicine, Houston Methodist Research Institute and Houston Methodist Hospital, Houston, Texas, USA

^dDepartment of Pathology and Laboratory Medicine, Weill Cornell Medical College, New York, New York, USA

^eU.S. Food and Drug Administration, Center for Veterinary Medicine, Office of Research, Laurel, Maryland, USA

^fDepartment of Computer Science, University of Chicago, Chicago, Illinois, USA

ABSTRACT Nontyphoidal *Salmonella* species are the leading bacterial cause of foodborne disease in the United States. Whole-genome sequences and paired antimicrobial susceptibility data are available for *Salmonella* strains because of surveillance efforts from public health agencies. In this study, a collection of 5,278 nontyphoidal *Salmonella* genomes, collected over 15 years in the United States, was used to generate extreme gradient boosting (XGBoost)-based machine learning models for predicting MICs for 15 antibiotics. The MIC prediction models had an overall average accuracy of 95% within ± 1 2-fold dilution step (confidence interval, 95% to 95%), an average very major error rate of 2.7% (confidence interval, 2.4% to 3.0%), and an average major error rate of 0.1% (confidence interval, 0.1% to 0.2%). The model predicted MICs with no *a priori* information about the underlying gene content or resistance phenotypes of the strains. By selecting diverse genomes for the training sets, we show that highly accurate MIC prediction models can be generated with less than 500 genomes. We also show that our approach for predicting MICs is stable over time, despite annual fluctuations in antimicrobial resistance gene content in the sampled genomes. Finally, using feature selection, we explore the important genomic regions identified by the models for predicting MICs. To date, this is one of the largest MIC modeling studies to be published. Our strategy for developing whole-genome sequence-based models for surveillance and clinical diagnostics can be readily applied to other important human pathogens.

KEYWORDS antimicrobial susceptibility testing, deep learning, diagnostics, genome sequencing, machine learning

Nontyphoidal *Salmonella* species are the leading bacterial cause of foodborne disease in the United States (1, 2), causing over 1 million illnesses per year (3) and an estimated 80 million illnesses annually worldwide (4). Nontyphoidal *Salmonella* causes acute gastroenteritis and is usually contracted via fecal contamination of food sources (5). Although these infections are usually self-limiting and typically do not require antibiotic treatment (6), severe infections can occur (7). Antimicrobial resistance (AMR) is prevalent in *Salmonella* isolates, and infections caused by highly antimicrobial-resistant *Salmonella* strains result in worse outcomes than infections caused by susceptible strains (8–11).

In 1996, the National Antimicrobial Resistance Monitoring System (NARMS) was established as a collaboration between the U.S. Centers for Disease Control and

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Address correspondence to James J. Davis, jimdavis@uchicago.edu.

* Present address: Marcus Nguyen, Robert Olson, Rick L. Stevens, and James J. Davis, Argonne National Laboratory, Computing, Environment and Life Sciences, Argonne, Illinois, USA.

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Prevention (CDC), the U.S. Food and Drug Administration (FDA), the U.S. Department of Agriculture (USDA), and state and local health departments. A primary goal of NARMS is to monitor antimicrobial resistance in *Salmonella* and other foodborne bacteria, including *Campylobacter*, *Escherichia*, and *Enterococcus* (12). The data collected by NARMS are used to inform public health decisions aimed at identifying contaminated food sources and reducing the spread of AMR through enhanced stewardship. In recent years, NARMS has adopted whole-genome sequencing (WGS) as a routine monitoring tool. The WGS data are used to determine the source of outbreak strains and the virulence factor and AMR genes carried by each strain. As a result, a large collection of bacterial whole-genome sequences with extensive metadata is available for downstream research efforts (13).

Whole-genome sequencing is now routinely used for public health surveillance and to guide diagnostic and patient care decisions (14–18). For routine surveillance, WGS provides the highest possible resolution for individuating traits in bacteria, assessing phylogenetic relationships, conducting outbreak investigations, and predicting virulence and epidemicity. From the clinical perspective, rapid diagnostics are key to improving patient care. For a conventional microbiology laboratory diagnosis, the total time for organism growth, isolation, taxonomic identification, and antimicrobial MIC determination may exceed 36 h for relatively fast-growing bacteria and several days for slower-growing organisms (19–21). Since reducing the time to optimal antimicrobial therapy significantly improves patient outcomes (22–24), rapid sequencing-based approaches that predict MICs may have clinical utility. The extensive WGS data sets generated by health agencies and the scientific community, such as those for nontyphoidal *Salmonella* strains, provide the necessary training sets required for building predictive models.

Several investigations have recently built models for predicting AMR phenotypes from WGS data. To date, the most common approach has relied on using a curated reference set of genes and polymorphisms that are implicated in AMR (25–33). This reference-guided approach best predicts susceptibility and resistance when organisms are well studied and the AMR mechanisms are known. As larger collections of genomes have become available, several studies have used machine learning algorithms to predict susceptible and resistant phenotypes (27, 29, 31, 34–38). By using WGS and AMR phenotype data to train a machine learning model, predictions without *a priori* information about the underlying gene content of the genome or molecular mechanism for resistance to each agent are possible. Although this reference-free approach requires many genomes, it is unbiased and can potentially be used to enable the discovery of new genomic features that are involved in AMR (36, 37). These two general approaches have also been used to predict MICs for *Streptococcus*, *Neisseria*, and *Klebsiella* (35, 38–40). When a curated reference collection of genes and single nucleotide polymorphisms (SNPs) is used for predicting MICs, a rules-based or machine learning algorithm is required for determining how much a given feature contributes to the MIC. Thus, for MIC prediction, both reference-guided and reference-free approaches are expected to have similar advantages and disadvantages if the collection of genes and SNPs used by the reference-guided method is sufficient for predicting all MICs, including those that are in the susceptible range. For example, in previous work, we built a machine learning model to predict MICs for a comprehensive population-based collection of 1,668 *Klebsiella pneumoniae* clinical isolates (38). For each genome, we used nucleotide 10-mers and the MICs for each antibiotic as features to train the model. Extreme gradient boosting (XGBoost) was chosen as the machine learning algorithm (41). The model could rapidly predict the MICs for 20 antibiotics with an average accuracy of 92%. This demonstrated that it is possible to successfully predict MICs without using a precompiled set of AMR genes or polymorphisms.

In this study, we build models that use whole-genome sequence data to predict MICs for nontyphoidal *Salmonella* based on the strains collected and sequenced by NARMS from 2002 to 2016. Our strategy can be used to guide responses to outbreaks and inform antibiotic stewardship decisions.

MATERIALS AND METHODS

Genomes and metadata. A total of 5,278 nontyphoidal *Salmonella* genome sequences were used in this study. All strains were collected and sequenced as part of the NARMS program. The strains were recovered either from raw retail meat and poultry or directly from livestock animals at slaughter. Antimicrobial susceptibility testing was performed using broth microdilution on a Sensititre system (Thermo Scientific) for 15 antibiotics, ampicillin (AMP), amoxicillin-clavulanic acid (AUG), ceftriaxone (AXO), azithromycin (AZI), chloramphenicol (CHL), ciprofloxacin (CIP), trimethoprim-sulfamethoxazole (COT), sulfisoxazole (FIS), ceftiofur (FOX), gentamicin (GEN), kanamycin (KAN), nalidixic acid (NAL), streptomycin (STR), tetracycline (TET), and ceftiofur (TIO), at FDA and USDA NARMS laboratories (13). Clinical breakpoints are based on CLSI and FDA guidelines (42). Whole-genome sequencing was performed using the Illumina HiSeq and MiSeq platforms and standard methods (25). Accession numbers and MICs for each isolate are listed in Table S1 in the supplemental material. All non-AMR metadata, including the serotype, host, geographic location of isolation, and isolation year, were taken from the metadata associated with each NCBI Sequence Read Archive (SRA) entry.

Genomic analyses. The short-read sequence data for each strain were assembled with the PATRIC genome assembly service (43), using the full SPAdes pipeline, which uses the BayesHammer tool (44) for read correction and SPAdes software for assembly (45). All genomes were annotated using the PATRIC annotation service (43), which uses a variation of the RAST tool kit annotation pipeline (46). Annotated genomes are available on the PATRIC website (<https://patricbrc.org>). PATRIC genome identifiers are displayed in Table S1. Protein annotations, including those for proteins specifically asserted to be involved in AMR (47), were downloaded from the PATRIC work space and used for subsequent analyses. A phylogenetic tree was generated for the strains in the analysis by aligning the genes for the beta and beta prime subunits of the RNA polymerase using MAFFT software (48), concatenating the alignments, and computing a tree with the FastTree tool (49). The tree was rendered using the iTOL tool (50).

MIC prediction. (i) Model generation. A model for predicting the MICs for the 15 antibiotics was built following the methods previously described by Nguyen and colleagues (38). Briefly, each genome was divided into the set of nonredundant overlapping nucleotide 10-mers using the k-mer counting program KMC (51). A matrix in which the k-mers, antibiotics, and MICs are treated as features for each genome was built. Each row in the matrix contains the k-mers for a genome as well as the MIC for a single antibiotic. The MIC prediction model was built using an XGBoost (41) regressor predicting linearized MICs. All model parameters were identical to those used by Nguyen et al. (38). Briefly, XGBoost is a computationally scalable method for generating gradient-boosted models. Gradient boosting is an ensemble method by which decision trees are generated to minimize an error function. We chose this method because of its scalability and its built-in ability to perform feature selection. The sensitivity and accuracy of the models generated in this study were tested by performing 10-fold cross validations. The cross validation partitions the matrix into 10 equal parts, such that each part has an equal (or nearly equal) number of antibiotic-MIC combinations. Ten rounds (folds) of training are performed: one part is used for testing, one part is used for validation, and eight parts are used for training. In this way, each part is used once for testing purposes, and any biases can be observed by tracking the average accuracy and confidence interval (CI) size over each fold. The validation set was used to monitor each model to prevent overfitting.

Unless otherwise stated, the accuracy of a model is reported as the ability to predict the correct MIC within ± 1 2-fold dilution step of the laboratory-derived MIC. Defining an accuracy to be within ± 1 2-fold dilution step is consistent with FDA requirements for automated MIC measuring device standards and is consistent with established clinical microbiology practices (20, 52–54). A comparison of raw accuracies and accuracies within ± 1 2-fold dilution step is shown in Table S2. To assess the accuracy of a model over various metadata categories, including date, serotype source, and location, the training set genomes are used to make the model. The test set genomes are used to assess the model accuracy for a given fold. For models based on date ranges, all parameters are identical and the accuracy over the genomes is reported from the held-out dates.

(ii) Subsampling. All models were built and tested on a server with 4 Intel E5-4669v4 central processing units (Xeon at 2.2 GHz, 22 cores hyperthreaded to 44) with 1.585 terabytes (TB) of random-access memory (RAM). In order to perform the model building on this machine (machines with more memory are currently somewhat uncommon), we reduced the matrix size to sets of size n , where n is $\leq 250, 500, 1,000, 2,000, 3,000, 4,000$, and 4,500 genomes. To create a diverse subset of size n , a hierarchical clustering method (55) was used to create n clusters by using the 10-mer distribution of each genome as the input features. To avoid the curse of dimensionality (56, 57), the taxicab/Manhattan distance (L_1 norm) rather than the Euclidean distance (L_2 norm) was used, since previous research has shown it to be both computationally fast and more accurate for high-dimensional data (58). From the resulting n clusters, one genome from each cluster was randomly selected from a uniform distribution to create the subset containing n genomes. For each subset of genomes, a matrix was generated, and models were generated as described above.

(iii) Feature identification. In order to unambiguously identify k-mers that are important to MIC prediction, we built separate models for each individual antibiotic using the method described above, except that we increased the k-mer length to 15 nucleotides in order to reduce the number of redundant k-mers within each genome and to enable analyses with the BLAST program (59). We also measured k-mer hits as presence versus absence, rather than counts, in order to simplify the analysis. Each model was built using the set of 1,000 diverse genomes from the subsampling experiment described above, and 10-fold cross validations were performed on each model.

The XGBoost feature importance score was computed using the internal XGBoost function `xgboost.Booster.get_score()`, using gain as the metric. This was computed for each fold within the 10-fold cross validation. This results in an importance score per feature (15-mer) from each fold. We summed the feature importance scores from each fold for the top 10 features in order to generate an overall importance score across all folds. This overall importance score captures both the importance of the 15-mer to a given fold and the number of times that the 15-mer was chosen as a top feature within each of the 10 folds.

XGBoost's internal feature importance is unable to provide correlations between features and label values and thus does not provide an indication of whether a k-mer is related to antibiotic resistance or susceptibility. This is partially due to the fact that many nonlinear correlations that may use multiple features exist. In order to see if the high-scoring k-mers correlate with resistance or susceptibility, we computed the distribution of MICs for the genomes containing each high-scoring k-mer. For example, a k-mer conferring susceptibility should be found in more genomes with lower MICs, while a k-mer conferring resistance should exist in genomes with higher MICs. Each high-scoring k-mer was also compared to the set of protein-encoding genes within each *Salmonella* genome. If a k-mer was found within a known AMR gene, that gene is reported. Otherwise, we report the distance (in nucleotides) between the center of the k-mer and the center of nearby AMR genes. This search was scoped to 3 kb to reduce the search space.

To find k-mers that are being used by the individual antibiotic models to predict susceptible MICs, we computed the presence or absence of each k-mer with high XGBoost feature importance values (described above) for the entire data set of 5,278 genomes. The k-mers with the largest difference in occurrence between the susceptible and resistant genomes are the ones that are chosen by the models for predicting susceptible MICs. To determine if there were significant SNPs in these k-mers, we found the genomic features containing the k-mer—protein-encoding gene, RNA gene, or intergenic region—using the BLASTn program (59). The corresponding feature or region was then found for all genomes in the collection. The features were aligned using MAFFT (48) and manually curated using the Jalview program (60). Poor-quality sequence was removed, all duplicates and paralogs were removed, and the subalignment covering the k-mer was extracted. To prevent possible biases due to clonality that may exist in the full set of genomes, the analysis was repeated on the diverse subset of 1,000 genomes (described above). We report an SNP in a k-mer region as being significant if the susceptible and resistant sets were significantly different ($P < 0.001$) for a given nucleotide position, based on a chi-square test, for both the full set of 5,278 genomes and the set of 1,000 diverse genomes. Sequence logos for k-mers containing significant SNPs were generated using the WebLogo application (61). k-mers from the azithromycin and ciprofloxacin models were excluded from this analysis because they each had seven resistant genomes. Comparisons of codon usage versus the genome average, genome mode, and high-expression gene sets were computed as described previously (62, 63).

Software availability. The *Salmonella* MIC prediction model based on 4,500 genomes—including the software and documentation for running the model—is available at our GitHub page (https://github.com/PATRIC3/mic_prediction).

Accession number(s). Data are available under BioProject accession numbers PRJNA292661 and PRJNA292666. The SRA run accession number for each genome is displayed in Table S1 in the supplemental material.

RESULTS

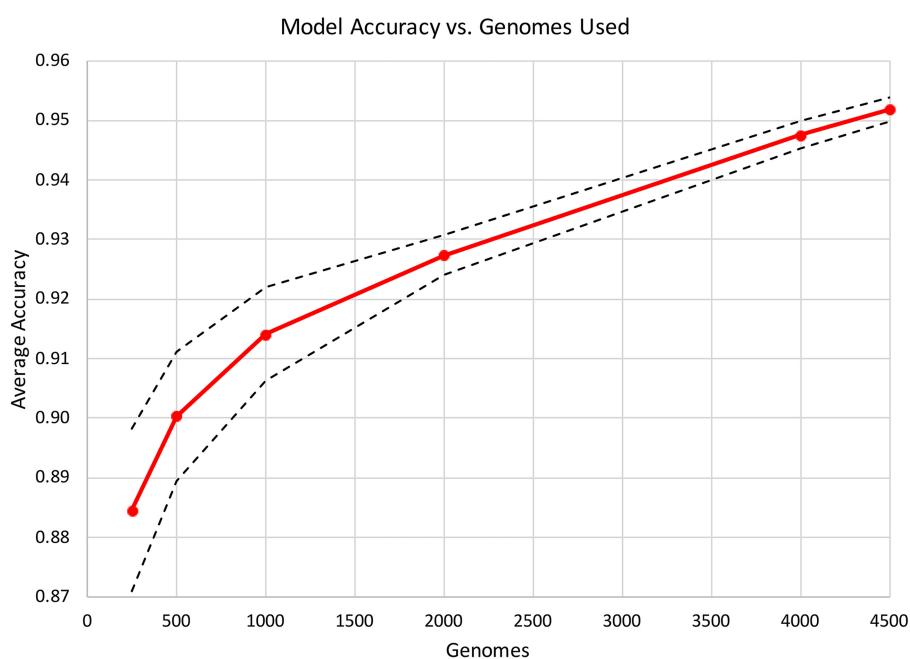
Model construction. For this study, we used a publicly available collection of 5,278 *Salmonella* whole-genome sequences generated by the NARMS project between 2002 and 2016. The strains were isolated from retail meat and food animal sources in the United States. The collection included 98 different serotypes, including *Salmonella enterica* serotypes Heidelberg (678 genomes), Kentucky (618 genomes), and Typhimurium var. 5— (588 genomes), from 41 states (see Table S1 in the supplemental material). Isolates were tested for resistance to up to 15 antimicrobial agents using the broth microdilution method. Many of the strains were randomly selected for sequencing as part of a compulsory nationwide collection program (Table 1).

The nontyphoidal *Salmonella* MIC prediction model was built by a strategy similar to our previously described strategy used to predict MICs for *K. pneumoniae* clinical isolates (38). Since the *Salmonella* data set has many more genomes and greater sampling in the range of susceptible MICs, it provides a critical test case for determining if the approach remains robust for other pathogens. In the *Klebsiella* study, we built individual models for each antibiotic, as well as a single large integrated model by combining the data from all antibiotics. We found that the combined model achieved slightly higher overall accuracies (by ~1% to 2%); however, the matrix that was necessary to train this model had a large memory footprint. Indeed, if we were to build a similar matrix for the current *Salmonella* data set using all 5,278 genomes, the model training would exceed 1.5 TB of RAM. RAM usage and training times are displayed in Fig. S1. Therefore, we first built models for all antibiotics using subsets of the genomes

TABLE 1 Number of genomes susceptible, intermediate, or resistant to the 15 antibiotics for the 5,278 *Salmonella* genomes used in this study

Antibiotic	No. of genomes:		
	Susceptible	Intermediate	Resistant
AMP	3,682	2	1,593
AUG	4,145	355	778
AXO	4,508	1	769
AZI	2,409	0	7
CHL	5,026	87	164
CIP	5,217	53	7
COT	5,219	0	58
FIS	3,356	0	1,573
FOX	4,501	98	679
GEN	4,577	68	633
KAN	837	3	84
NAL	5,233	0	45
STR	872	0	1,919
TET	2,364	28	2,885
TIO	4,517	8	753

ranging in size from 250 to 4,500 genomes, which were rationally selected to maximize genetic diversity (Fig. 1). A matrix built from 4,500 genomes is the largest that we can train on a 1.5-TB machine using this protocol. As the training set size increased from 250 to 1,000 genomes, the accuracy increased from 88.5% to 91.4%. Then, as the training set increased beyond 1,000 genomes, the accuracy continued to improve more modestly, with the 4,500-genome model having an average accuracy of 95.2%. The results indicated that the overall MIC prediction approach which was developed previously for *Klebsiella pneumoniae* also works for *Salmonella*, despite the differences in sampling, genetic diversity, and MICs. Also, we discovered that a smaller number of well-chosen diverse genomes can serve as a useful proxy for the entire set, since models built from ≥ 500 genomes have accuracies exceeding 90%.

**FIG 1** MIC prediction model accuracy for subsamples of genomes. Diverse subsamples of genomes were chosen, and the model accuracy within ± 1 2-fold dilution step based on a 10-fold cross validation is shown with the red plot line. The dashed lines represent the high and low values for the 95% confidence interval for the average accuracy at each given plot point.

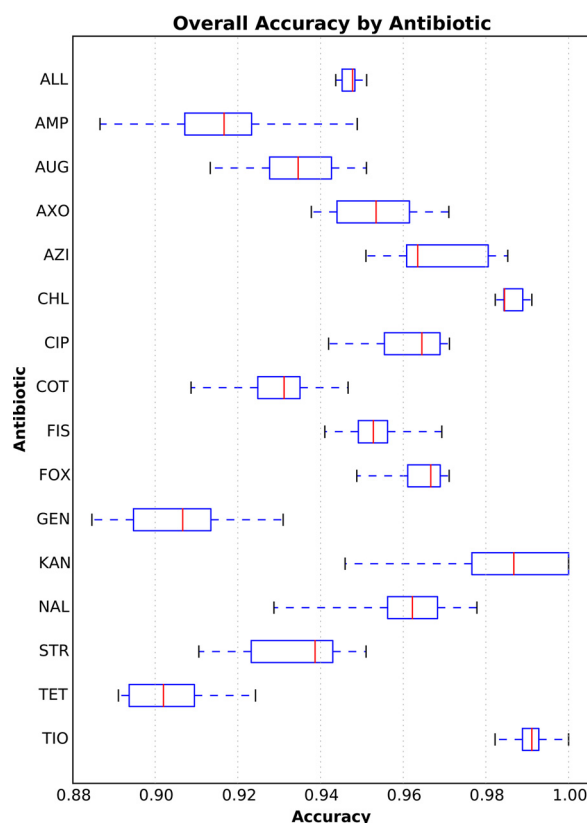


FIG 2 Box plot of the overall accuracies within ± 1 2-fold dilution step for each antibiotic in the 4,500-genome model. The y axis depicts each antibiotic (abbreviations are defined in Materials and Methods). The x axis depicts the accuracy. Each vertical red line represents the median accuracy over the holdout sets for each fold in the 10-fold cross validation. The blue box encompasses the data for the first and third quartiles. The dashed blue horizontal lines bounded by black vertical lines (or whiskers) depict the entire distribution of accuracies for each fold and antibiotic. The accuracy of the entire 4,500-genome model over all antibiotics and folds is depicted in the row marked "ALL."

Model accuracy. We computed the overall accuracy for each antibiotic using the model that was based on 4,500 genomes. For this model, all 15 antibiotics had average accuracies of $\geq 90\%$, with their first quartile being bound $\geq 89\%$ (Fig. 2). Chloramphenicol and ceftiofur had the highest accuracies (99%), and gentamicin and tetracycline had the lowest accuracies (91% and 90%, respectively) (Table S2). Since the model is robust to the various mechanisms of resistance for the 15 antibiotics, it is possible that the slightly lower accuracies for gentamicin and tetracycline could be due to the distribution of multiple AMR genes/mechanisms across the population of strains with resistant genomes (which will be analyzed in more detail below). Figure 3 depicts the accuracy of the 4,500-genome model for each MIC. Overall, the model was robust for both the resistant and susceptible MICs, and it tended to be more accurate when a MIC was represented by many genomes. The model tended to have lower accuracies for the highest and lowest MICs, perhaps because of underlying genetic differences between strains that have been reported with values greater than or equal to or less than or equal to a MIC, which represents a range of MICs rather than a discrete value.

The utility of AMR diagnostic devices is often described in terms of error rate. Major errors (MEs) are defined as susceptible genomes that have been incorrectly assigned resistant MICs by the model. Very major errors (VMEs) are defined as resistant genomes that have been incorrectly assigned susceptible MICs by the model. FDA standards for automated systems recommend a major error rate of $\leq 3\%$ (54). All antibiotics used in the model had ME rates within this range (Table 2). The FDA standards for VME rates indicate that the lower 95% confidence limit should be $\leq 1.5\%$ and that the upper limit

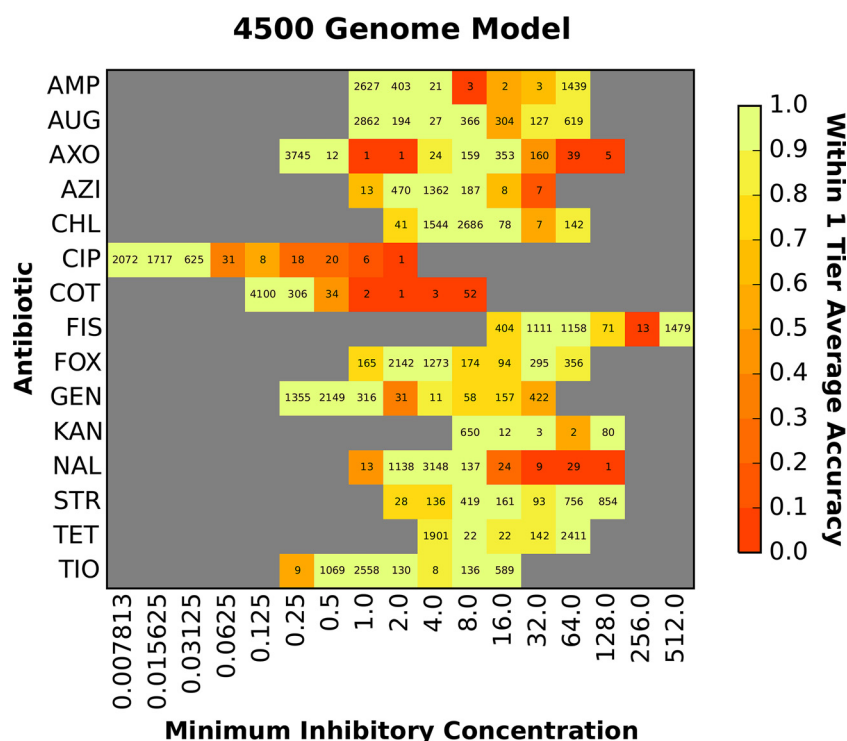


FIG 3 Accuracy of the MIC prediction model based on 4,500 diverse genomes. The heat map depicts the accuracy within ± 1 2-fold dilution step of the laboratory-derived MIC. The x axis shows the MIC (in micrograms per milliliter), and each antibiotic is shown on the y axis. The accuracy for each antibiotic-MIC combination is depicted by color, with bright yellow/green being the most accurate and red being the least accurate. The values shown in each cell are the number of genomes with that MIC for a given antibiotic.

should be $\leq 7.5\%$ (54). Seven of the 15 antibiotics—amoxicillin-clavulanic acid, ceftriaxone, chloramphenicol, ceftiofur, streptomycin, tetracycline, and ceftiofur—had acceptable VME rates by this measure. Ampicillin and sulfisoxazole had VME rates with 95% confidence intervals approaching the ranges of 0.022 to 0.033 and 0.026 to 0.053, respectively, failing only on the lower bound. The VME rates for some of the remaining

TABLE 2 VME and ME rates for the 4,500-genome model^a

Antibiotic	VME rate		ME rate		No. of samples:	
	Avg	95% CI	Avg	95% CI	Resistant	Susceptible
All	0.027	0.024 to 0.030	0.001	0.001 to 0.002	10,979	47,366
AMP	0.028	0.022 to 0.033	0.000	0.000 to 0.001	1,442	3,054
AUG	0.012	0.000 to 0.025	0.000	0.000 to 0.000	746	3,449
AXO	0.022	0.011 to 0.032	0.000	0.000 to 0.001	740	3,758
AZI	0.857	0.508 to 1.207	0.000	0.000 to 0.000	7	2,040
CHL	0.000	0.000 to 0.000	0.000	0.000 to 0.001	149	4,271
CIP	0.417	0.099 to 0.933	0.000	0.000 to 0.000	7	4,445
COT	0.670	0.515 to 0.825	0.000	0.000 to 0.001	55	4,443
FIS	0.039	0.026 to 0.053	0.000	0.000 to 0.000	1,479	2,757
FOX	0.009	0.001 to 0.020	0.000	0.000 to 0.000	651	3,754
GEN	0.090	0.066 to 0.113	0.000	0.000 to 0.000	579	3,862
KAN	0.074	0.012 to 0.136	0.000	0.000 to 0.000	82	662
NAL	0.917	0.819 to 1.014	0.000	0.000 to 0.001	39	4,460
STR	0.014	0.008 to 0.020	0.027	0.013 to 0.040	1,703	744
TET	0.000	0.000 to 0.000	0.018	0.012 to 0.025	2,575	1,901
TIO	0.004	0.001 to 0.009	0.000	0.000 to 0.000	725	3,766

^aThe rates of very major errors (VME), defined as resistant genomes predicted to be susceptible, and major errors (ME), defined as susceptible genomes predicted to be resistant, are reported within ± 1 twofold dilution step. CI, 95% confidence interval.

TABLE 3 Model accuracy for the genomes from each sample collection year

Collection date (yr)	Accuracy	No. of genomes	No. of bins ^a
2002	0.97	55	624
2003	0.95	159	1,809
2004	0.96	235	2,850
2005	0.95	274	3,384
2006	0.95	313	3,880
2007	0.94	258	3,192
2008	0.95	388	4,821
2009	0.95	436	5,367
2010	0.94	230	2,820
2011	0.95	214	2,968
2012	0.96	257	3,694
2013	0.97	265	3,793
2014	0.95	506	7,100
2015	0.95	689	9,646
2016	0.96	83	1,161

^aThe total number of MICs available for the genomes isolated in that year.

antibiotics were higher because there were fewer resistant genomes. As more resistant genomes are collected and the data set becomes more balanced, we expect VME rates to be reduced.

In addition to the extensive MIC data, NARMS reports rich metadata, including isolation date, food or animal source, collection year, geographic location, and serotype. We computed the accuracy of the model over each available metadata category to determine if the model is robust to these differences and to ensure that no subset biased the model. The period of collection of the genomes spanned 15 years, with all the years except 2002 (the oldest) and 2016 (the most recent) having over 100 isolates. The model accuracy ranged from 94% to 97% over each collection year (Table 3). That is, the genetic factors that contribute to the MICs either remained stable over the 15-year period or were learned as the model was trained. Although the data set was mostly comprised of poultry meat or live animal isolates, the accuracy ranged from 94% to 96% over the four contamination sources: turkey, beef, pork, and chicken (Table 4). No obvious biases in the accuracies were detected on the basis of the state of isolation (an average of 95% accuracy over 41 states with a 95% CI equal to 0.95 to 0.96) (Fig. 4) or the serovar of each isolate (94% accuracy over 97 serovars with a 95% CI equal to 0.94 to 0.96) (Table S3). Since the traditional *Salmonella* serotyping scheme is based on the lipopolysaccharide O and flagellar H antigens, which are encoded by genes that influence the cell surface (64), we also constructed a phylogenetic tree for *Salmonella* genomes to observe the model accuracy over the various clades. Overall, no phylogenetic bias in the model accuracy was detected (Fig. S2).

One concern of using a model that is trained on the data from previous years and in some cases over 15 years old is that the training set is not representative of currently circulating strains. That is, the model may be inaccurate for predicting MICs for the genomes of strains that are currently circulating or will emerge in the future. For example, shifts in clonal groups, the evolution of AMR-associated genes, or the introduction of AMR genes by horizontal gene transfer is possible (65, 66). We evaluated this possibility by building models from subsets of the whole-genome sequence data using strains collected in earlier years and measuring the accuracy of the models on genomes collected in later years. Models were built for years prior to 2009 through 2014 and

TABLE 4 Model accuracy for the genomes isolated from various sources

Source	Accuracy	No. of genomes	No. of bins ^a
Chicken	0.96	1,981	25,869
Cow/beef	0.94	419	5,688
Pig/pork	0.95	448	6,144
Turkey	0.94	1,651	21,260

^aThe total number of MICs available for the genomes of each category.

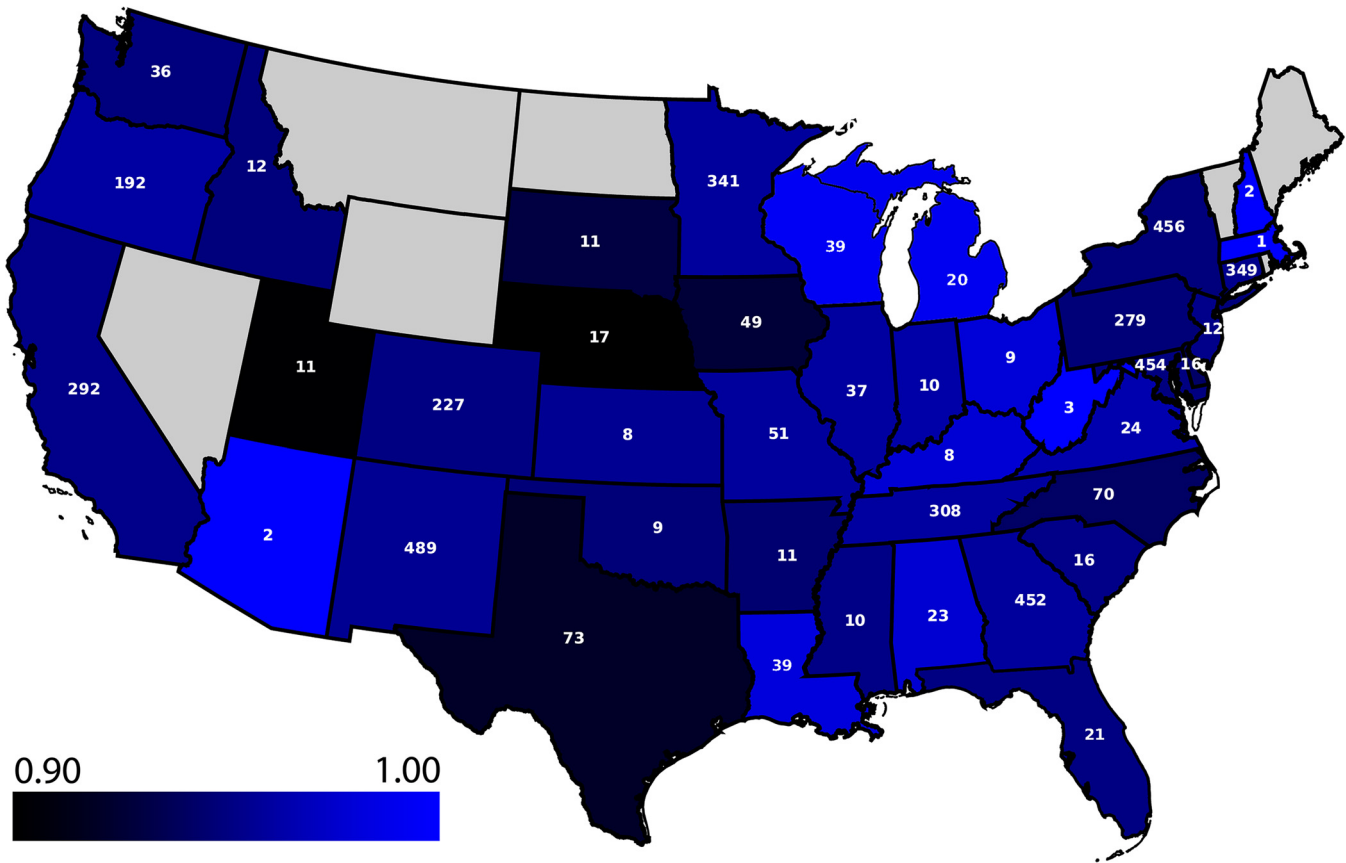


FIG 4 Average accuracy of the model based on 4,500 diverse genomes for predicting MICs for the *Salmonella* genomes from each state. Light blue is the most accurate, and dark blue/black is the least accurate. Note that the scale starts at an accuracy of 0.90. Each state is labeled with the number of genomes collected from that state. States without a label contain no samples and are colored in gray; no genomes from Alaska and Hawaii exist in the collection.

tested on the remaining genomes (Table 5). These models had accuracies ranging from 86% to 92%. As the number of years used for building the models increased, the number of genomes available for testing decreased, so we also tested each model on only the 462 genomes from 2015 and 2016. Similarly, the accuracy of each model on the genomes from 2015 and 2016 ranged from 87% to 90% (Table S4). The results indicate that within this data set, the models generated from genomes collected at earlier dates yielded stable MIC predictions for genomes collected at later dates. This finding is consistent with the pattern of AMR genes that was observed within the data set. Although the AMR gene content may vary from year to year, we did not observe any major sweeps or fixation events that drastically altered the AMR gene content of the collection between years, which would cause the MIC predictions to fail for a large fraction of the genomes (Table S5). Taken together, these data suggest that the MIC prediction models generated in this study are likely to be sustainable over time.

TABLE 5 Ability of models trained on genomes from prior years to predict MICs for genomes collected in later years

Training set yr	Test set yr	Accuracy	95% CI for accuracy	No. of training bins ^a	No. of testing bins ^a	No. of training genomes	No. of testing genomes
2002–2008	2009–2016	0.88	0.88–0.89	36,563	22,412	1,819	2,681
2002–2009	2010–2016	0.88	0.88–0.89	31,196	27,779	2,255	2,245
2002–2010	2011–2016	0.88	0.88–0.88	28,376	30,599	2,485	2,015
2002–2011	2012–2016	0.88	0.88–0.89	25,408	33,567	2,699	1,801
2002–2012	2013–2016	0.88	0.87–0.88	21,714	37,261	2,956	1,544
2002–2013	2014–2016	0.86	0.86–0.87	17,921	41,054	3,221	1,279
2002–2014	2015–2016	0.92	0.92–0.92	10,807	48,168	3,728	772

^aThe total number of genome/antibiotic combinations.

Genomic regions important for MIC prediction. The 4,500-genome model described above contained data from all antibiotics and MICs, making feature extraction to determine which k-mers contribute to the MIC predictions for each antibiotic difficult. To address this limitation, we modified the protocol by building separate models for each antibiotic. Instead of using 10-mers, we increased the k-mer length to 15 nucleotides to reduce redundancy and make them identifiable using BLAST analysis (59). We also searched for the presence or absence of k-mers, rather than using k-mer counts, to simplify the analysis of the XGBoost decision trees. Since a 15-mer matrix can be 4^5 times larger than a 10-mer matrix, we used $\leq 1,000$ diverse genomes to reduce the memory footprint during training. Overall, the average accuracy for the individual models was nearly identical to the average accuracy for the combined 4,500-genome model (96% versus 95%, respectively), and in nearly all cases, the 95% confidence intervals overlapped between the combined and single antibiotic models (Table S6). Thus, for this data set, single antibiotic models with fewer genomes and larger k-mers performed as well as a combined model (Fig. S3).

During model training, XGBoost assigns an importance value to each k-mer used in a decision tree. When the model is used to predict the MICs for a new genome, the k-mers with the highest importance values are the most informative for the MIC prediction. Thus, by analyzing the feature importance values of each k-mer, we can use the models as a tool for understanding the genomic regions that differentiate MICs. For each antibiotic-specific model, we parsed the XGBoost decision trees from each fold of the 10-fold cross validation to extract the importance values for each k-mer. To understand the relationship between known AMR genes and the important k-mers that were chosen by each model, we then searched for k-mers that had high importance values within AMR genes or that occurred in close proximity to an AMR gene. Table 6 lists the highest-ranking k-mers from each model that occurred within or in close proximity to an AMR gene. In most cases, the k-mers corresponded to known AMR genes, including class A and C beta-lactamases for the beta-lactam antibiotics, aminoglycoside nucleotidyl- and acetyltransferases for the aminoglycosides, DNA gyrase and QnrB for the fluoroquinolones, TetA and TetR for tetracycline, and dihydrofolate reductase and dihydropteroate synthase for co-trimoxazole and sulfisoxazole. In the case of azithromycin, the collection mostly contained susceptible genomes (Table 1), so the first macrolide resistance gene observed corresponds to the eighth-ranking k-mer. The top 10 k-mers with the highest feature importance values from each of the 10 folds used in model training are listed in Tables S7 to S21. In addition to the top AMR k-mers displayed in Table 6, these tables show other highly ranking k-mers from the same AMR genes as well as k-mers from related genes that are known to confer resistance to the given antibiotic. In some cases, k-mers matching regions or genes from unrelated AMR mechanisms have high importance values, suggesting a pattern of co-occurrence on horizontally transferred genetic elements.

Since each model predicts the entire range of MICs, some of the highly ranking k-mers may be used to predict susceptible MICs. To assess this, we computed the fraction of susceptible and resistant genomes with each k-mer from Tables S7 to S21. The set of k-mers that were most enriched in the susceptible genomes is shown in Table 7. Overall, 7 of the top 10 k-mers represented significantly different SNPs ($P < 0.001$) both in the complete set of 5,278 genomes and in the set of 1,000 diverse genomes used to build the models (Fig. S4). The top k-mer associated with susceptibility was from the nalidixic acid model and occurred in the DNA gyrase *gyrA* gene. This is also the top k-mer that was found in an AMR gene for nalidixic acid from Table 6. In this case, the model relied more heavily on the wild-type version of the k-mer rather than any of the resistant versions (the remaining k-mers from Table 6 occurred almost exclusively in resistant genomes). The same *gyrA* k-mer was also found as a highly ranking k-mer in the case of ciprofloxacin (Table S12). Two significant *gyrA* SNPs were captured by this k-mer (Fig. S4). These are missense mutations in the resistant genomes occurring at Ser83 and Asp87, and changes at these positions have been shown to

TABLE 6 Highest-ranking AMR-related protein function (or genomic region) with a matching k-mer from the XGBoost models

Antibiotic	k-mer rank	Distance between k-mer and AMR gene ^a	k-mer sequence	PATRIC annotation(s)
AMP	1	Direct match	CTTAATCAGTGAGGC	Class A beta-lactamase (EC 3.5.2.6) => TEM family
AUG	1	Direct match	AAACGTCTTACTAAC	Class C beta-lactamase (EC 3.5.2.6) => CMY/CMY-2/CFE/LAT family
AXO ^b	1	566.0 ± 39.7	AAAGAGAAAGAAAGG	Class C beta-lactamase (EC 3.5.2.6) => CMY/CMY-2/CFE/LAT family
AZI	8	Direct match	CCCATTTCCGCCGCC	Macrolide 2'-phosphotransferase => Mph(A) family
CHL ^b	1	611.8 ± 5.1	AGACAAGTAAGCCGC	Chloramphenicol/florfenicol resistance, MFS efflux pump => FloR family
CIP	1	313.5 ± 70.5	ACAGTCCATCCAGGA	Pentapeptide repeat protein QnrB family => Quinolone resistance protein QnrB10
COT ^b	1	248.4 ± 7.0	AAAAACGATAGCTGC	Dihydrofolate reductase (EC 1.5.1.3)
FIS ^b	1	408.9 ± 11.3	CGCAACGGCTCAAGC	Dihydropteroate synthase type-2 (EC 2.5.1.15) @ Sulfonamide resistance protein
FOX	1	Direct match	AAAAAACCTTGGCA	Class C beta-lactamase (EC 3.5.2.6) => CMY/CMY-2/CFE/LAT family
GEN	1	439.0 ± 70.4, 451.2 ± 7.0	AGTTAAGCCGCGCCG	Aminoglycoside 3"-nucleotidyltransferase (EC 2.7.7.-) => APH(3'')-Ia (AadA family); Aminoglycoside N(3)-acetyltransferase (EC 2.3.1.81) => AAC(3)-II,III,IV,VI,VII,IX,X
KAN	1	Direct match	AAAAAGCCGTTTCTG	Aminoglycoside 3'-phosphotransferase (EC 2.7.1.95) => APH(3')-I
NAL	1	Direct match	ATTCGCGAGTGATG	DNA gyrase subunit A (EC 5.99.1.3)
STR	1	Direct match	ATTGTACGGCTCCG	Aminoglycoside 3"-nucleotidyltransferase (EC 2.7.7.-) => APH(3'')-Ia (AadA family)
TET	1	335.6 ± 12.6, 596.5 ± 22.5	CGTTCTGCTTGCGC	Tetracycline resistance regulatory protein TetR; Tetracycline resistance, MFS efflux pump => Tet(A)
TIO ^b	1	566 ± 39.7	AAAGAGAAAGAAAGG	Class C beta-lactamase (EC 3.5.2.6) => CMY/CMY-2/CFE/LAT family

^aA direct match means that the k-mer is an exact match to the protein-encoding gene. Average distances with standard deviations, measured as the number of nucleotides from the center of the k-mer to the center of the AMR gene, are shown for the set of 1,000 diverse genomes.

^bTIO has other AMR genes within 3 kb listed in Tables 57 to 521.

TABLE 7 Important k-mers used by the individual antibiotic models for predicting susceptible MICs^a

Antibiotic	k-mer sequence	Sus	Res	Frac sus	Frac res	Genomic region	PATRIC annotation or genomic region
NAL	ATTCCGCAGTGTATG	5,233	45	1.00	0.38	PEG	DNA gyrase subunit A (EC 5.99.1.3)
AXO	TGGTATTCGCATCAA	4,508	769	0.78	0.48	PEG	Phosphoethanolamine transferase EptA
KAN	CTGGCTTTTTTTTT	837	84	0.30	0.00	RNA	RyhB RNA
STR	CCCTTATCCAACACG	872	1,919	0.85	0.55	PEG	Respiratory nitrate reductase delta chain (EC 1.7.99.4)
AXO	CAGAACCAGAATTG	4,508	769	0.74	0.46	PEGs	Formate-dependent nitrite reductase complex subunit Nrff, and Cytochrome c-type heme lyase subunit Nrfe, nitrite reductase complex assembly
TIO	AGAGAAGCCTGCCGC	4,517	753	0.68	0.40	PEG	Oxaloacetate decarboxylase alpha chain (EC 4.1.1.3)
AXO	ATCCCCGCCATTACA	4,508	769	0.73	0.46	PEG	Tagatose-1,6-bisphosphate aldolase GatY (EC 4.1.2.40)
AXO	TGCTGCAAACGCCA	4,508	769	0.69	0.45	PEG	Protein AraJ precursor
AXO	GAAAACAGGGTG TAG	4,508	769	0.47	0.23	INT	Upstream of IlvGMEDA operon leader peptide
FOX	GGATACCACGCCGGG	4,501	679	0.58	0.35	PEGs	Glucose dehydrogenase, PQQ-dependent (EC 1.1.5.2), and IncF plasmid conjugative transfer protein TraP

^aSus and Res, total number of susceptible and resistant genomes in the entire collection, respectively; Frac sus and Frac res, fraction of susceptible and resistant genomes with the k-mer in the entire collection, respectively; PEG, protein-encoding gene; RNA, RNA gene; INT, intergenic region.

confer quinolone resistance in *Escherichia coli* (67, 68). The remaining significant mutations from Fig. S4 that occurred in the protein-encoding genes were the same sense (not amino acid-changing) mutations. In the cases of *eptA* (Ser, TCG to TCA), *oadA* (Ala, GCC to GCA), the *AraJ* precursor gene (Leu, CTG to CTA), and the second *gcd* mutation (Thr, ACG to ACA), the codon changed from a commonly used codon in the susceptible genomes to the least preferred codon in the resistant genomes. In the cases of the *nrfe* and *nrff* mutations (Asn, AAT to AAC) and the first *gcd* mutation (Asp., GAC to GAT), the resistant genomes had the preferred codon of the pair. Whether these SNPs have a modulating effect on protein translation or contribute to the fitness of the resistant organisms requires further analysis.

DISCUSSION

In this study, we built, using XGBoost (41), machine learning-based MIC prediction models for nontyphoidal *Salmonella* genomes that achieved overall accuracies of 95% to 96% within a ± 1 2-fold dilution factor. To our knowledge, this is one of the largest and most accurate MIC prediction models to be published to date. Importantly, it provides a model strategy for performing MIC prediction directly from genome sequence data that could be applied to other human or veterinary pathogens.

The success of our MIC prediction model was dependent on the large, publicly available, population-based collection of genomes with associated metadata. Since researchers often focus on collecting highly resistant or otherwise unusual strains, the opportunities to generate balanced models are rare. We demonstrate the many benefits from comprehensive sampling for the entire range of possible MICs. First, the use of diverse and balanced data sets improved model accuracies because there was better sampling across all MIC dilutions. Second, having balanced data enabled us to achieve acceptable ME and VME rates for 7 of the 15 antibiotics used in the study. Third, compared with our recent model for *Klebsiella pneumoniae*, the larger and more balanced data set for nontyphoidal *Salmonella* enabled us to build models for individual antibiotics that had accuracies similar to those of the combined model. This enabled us to begin to disambiguate the important genomic regions relating to resistant and susceptible MICs. Finally, we show that MICs in the susceptible range can be accurately predicted with the algorithm using all genomic data rather than scoping it to known AMR genes or gene polymorphisms. This contrasts with prior work correlating MICs to known resistance mechanisms in *Salmonella* (69). In future studies, our strategy could be used as a starting point for identifying the subtle genomic changes that result in different MICs.

For each single-antibiotic model, we analyzed the k-mers that had high feature importance values and were important to the models for predicting MICs. The highly ranking k-mers that were enriched in the resistant genomes mainly occurred within or

in close proximity to well-known AMR genes. With the exception of the *gyrA* k-mer, the highly ranking k-mers that were enriched in the susceptible genomes were significant in several cases but were more difficult to interpret. Some of these susceptibility k-mers hint at a possible relationship between AMR and oxidative stress or electron transport, such as the k-mers matching components of the nitrate and nitrite reductases and *pqq*-dependent glucose dehydrogenase, which is consistent with the known link between antibiotics and oxidative stress (70, 71). Although XGBoost would be expected to yield similar results with other ensemble machine learning methods, it may be possible to generate a more interpretable set of features using support vector machines or a single perceptron. However, these methods are not ideal, given the size of the input data. Future machine learning work that eliminates the effect of the resistance k-mers and bench work specifically examining these mutations may shed light on the subtle effects that result in the gradient of susceptible MICs.

The genomes used in this study were collected over a 15-year period from 41 U.S. states. By building models encompassing genomes collected over ranges of earlier dates, we demonstrated stable and accurate MIC prediction for genomes collected at later dates. Presently, we are not aware of any large publicly available collections of *Salmonella* genomes with MIC data from other countries. Since the AMR gene content may vary across pathogen populations, validation of the *Salmonella* models using strains from other countries is important to its application in global health. Nevertheless, the present analysis clearly demonstrates that the current model provides accurate MIC predictions for United States isolates. Similarly, an analysis of this model on *Salmonella enterica* serotype Typhi strains would provide information about the utility of the model over broader phylogenetic distances.

SUPPLEMENTAL MATERIAL

Supplemental material for this article may be found at <https://doi.org/10.1128/JCM.01260-18>.

SUPPLEMENTAL FILE 1, PDF file, 0.5 MB.

SUPPLEMENTAL FILE 2, XLSX file, 3 MB.

SUPPLEMENTAL FILE 3, XLSX file, 0.1 MB.

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REFERENCES

- Centers for Disease Control and Prevention (CDC). 2017. Surveillance for foodborne disease outbreaks, United States, 2015. Annual report. U.S. Department of Health and Human Services, Centers for Disease Control and Prevention, Atlanta, GA. https://www.cdc.gov/foodsafety/pdfs/2015FoodBorneOutbreaks_508.pdf.
- Crim SM, Griffin PM, Tauxe R, Marder EP, Gilliss D, Cronquist AB, Cartter

- M, Tobin-D'Angelo M, Blythe D, Smith K. 2015. Preliminary incidence and trends of infection with pathogens transmitted commonly through food—Foodborne Diseases Active Surveillance Network, 10 US sites, 2006–2014. *MMWR Morb Mortal Wkly Rep* 64:495–499.
3. Scallan E, Hoekstra RM, Angulo FJ, Tauxe RV, Widdowson M-A, Roy SL, Jones JL, Griffin PM. 2011. Foodborne illness acquired in the United States. *Emerg Infect Dis* 17:7–40. <https://doi.org/10.3201/eid1701.091101p1>.
4. World Health Organization. 2015. WHO estimates of the global burden of foodborne diseases: foodborne disease burden epidemiology reference group 2007–2015. World Health Organization, Geneva, Switzerland.
5. Andino A, Hanning I. 2015. *Salmonella enterica*: survival, colonization, and virulence differences among serovars. *ScientificWorldJournal* 2015: 520179. <https://doi.org/10.1155/2015/520179>.
6. Aserkoff B, Bennett JV. 1969. Effect of antibiotic therapy in acute salmonellosis on the fecal excretion of salmonellae. *N Engl J Med* 281: 636–640. <https://doi.org/10.1056/NEJM196909182811202>.
7. Crump JA, Sjölund-Karlsson M, Gordon MA, Parry CM. 2015. Epidemiology, clinical presentation, laboratory diagnosis, antimicrobial resistance, and antimicrobial management of invasive *Salmonella* infections. *Clin Microbiol Rev* 28:901–937. <https://doi.org/10.1128/CMR.00002-15>.
8. Varma JK, Mølbak K, Barrett TJ, Beebe JL, Jones TF, Rabatsky-Ehr T, Smith KE, Vugia DJ, Chang HH, Angulo FJ. 2005. Antimicrobial-resistant nontyphoidal *Salmonella* is associated with excess bloodstream infections and hospitalizations. *J Infect Dis* 191:554–561. <https://doi.org/10.1086/427263>.
9. Varma JK, Greene KD, Ovitt J, Barrett TJ, Medalla F, Angulo FJ. 2005. Hospitalization and antimicrobial resistance in *Salmonella* outbreaks, 1984–2002. *Emerg Infect Dis* 11:943. <https://doi.org/10.3201/eid1106.041231>.
10. Krueger AL, Greene SA, Barzilay EJ, Henao O, Vugia D, Hanna S, Meyer S, Smith K, Pecic G, Hoefler D, Griffin PM. 2014. Clinical outcomes of nalidixic acid, ceftriaxone, and multidrug-resistant nontyphoidal *Salmonella* infections compared with pansusceptible infections in FoodNet sites, 2006–2008. *Foodborne Pathog Dis* 11:335–341. <https://doi.org/10.1089/fpd.2013.1642>.
11. Angulo FJ, Mølbak K. 2005. Human health consequences of antimicrobial drug-resistant *Salmonella* and other foodborne pathogens. *Clin Infect Dis* 41:1613–1620. <https://doi.org/10.1086/497599>.
12. Karp BE, Tate H, Plumblee JR, Dessai U, Whichard JM, Thacker EL, Hale KR, Wilson W, Friedman CR, Griffin PM, McDermott PF. 2017. National Antimicrobial Resistance Monitoring System: two decades of advancing public health through integrated surveillance of antimicrobial resistance. *Foodborne Pathog Dis* 14:545–557. <https://doi.org/10.1089/fpd.2017.2283>.
13. U.S. Food and Drug Administration (FDA). 2018. NARMS now. U.S. Food and Drug Administration, Rockville, MD. Updated 20 July 2018. <https://www.fda.gov/AnimalVeterinary/SafetyHealth/AntimicrobialResistance/NationalAntimicrobialResistanceMonitoringSystem/ucm416741.htm>.
14. Abrams AJ, Trees DL. 2017. Genomic sequencing of *Neisseria gonorrhoeae* to respond to the urgent threat of antimicrobial-resistant gonorrhea. *Pathog Dis* 75:ftx041. <https://doi.org/10.1093/femspd/ftx041>.
15. Goldberg B, Sichtig H, Geyer C, Ledebner N, Weinstock GM. 2015. Making the leap from research laboratory to clinic: challenges and opportunities for next-generation sequencing in infectious disease diagnostics. *mBio* 6(6):e01888–15. <https://doi.org/10.1128/mBio.01888-15>.
16. Didelot X, Bowden R, Wilson DJ, Peto TE, Crook DW. 2012. Transforming clinical microbiology with bacterial genome sequencing. *Nat Rev Genet* 13:601. <https://doi.org/10.1038/nrg3226>.
17. Brown EW, Gonzalez-Escalona N, Stones R, Timme R, Allard MW. 2017. The rise of genomics and the promise of whole genome sequencing for understanding microbial foodborne pathogens, p 333–351. In Gurtler JB, Doyle MP, Kornacki JL (ed), *Foodborne pathogens: virulence factors and host susceptibility*. Springer International Publishing, Cham, Switzerland.
18. McArthur AG, Tsang KK. 2017. Antimicrobial resistance surveillance in the genomic age. *Ann N Y Acad Sci* 1388:78–91. <https://doi.org/10.1111/nyas.13289>.
19. Opota O, Croxatto A, Prod'homme G, Greub G. 2015. Blood culture-based diagnosis of bacteraemia: state of the art. *Clin Microbiol Infect* 21: 313–322. <https://doi.org/10.1016/j.cmi.2015.01.003>.
20. Reller LB, Weinstein M, Jorgensen JH, Ferraro MJ. 2009. Antimicrobial susceptibility testing: a review of general principles and contemporary practices. *Clin Infect Dis* 49:1749–1755. <https://doi.org/10.1086/647952>.
21. Saha SK, Darmstadt GL, Baqui AH, Hanif M, Ruhulamin M, Santosham M, Nagatake T, Black RE. 2001. Rapid identification and antibiotic susceptibility testing of *Salmonella enterica* serovar Typhi isolated from blood: implications for therapy. *J Clin Microbiol* 39:3583–3585. <https://doi.org/10.1128/JCM.39.10.3583-3585.2001>.
22. Llor C, Bjerrum L. 2014. Antimicrobial resistance: risk associated with antibiotic overuse and initiatives to reduce the problem. *Ther Adv Drug Saf* 5:229–241. <https://doi.org/10.1177/2042098614554919>.
23. Kumar A, Roberts D, Wood KE, Light B, Parrillo JE, Sharma S, Suppes R, Feinstein D, Zanotti S, Taiberg L, Gurka D, Kumar A, Cheang M. 2006. Duration of hypotension before initiation of effective antimicrobial therapy is the critical determinant of survival in human septic shock. *Crit Care Med* 34:1589–1596. <https://doi.org/10.1097/01.CCM.0000217961.75225.E9>.
24. Palmer H, Palavecino E, Johnson J, Ohl C, Williamson J. 2013. Clinical and microbiological implications of time-to-positivity of blood cultures in patients with Gram-negative bacilli bacteremia. *Eur J Clin Microbiol Infect Dis* 32:955–959. <https://doi.org/10.1007/s10096-013-1833-9>.
25. McDermott PF, Tyson GH, Kabera C, Chen Y, Li C, Folster JP, Ayers SL, Lam C, Tate HP, Zhao S. 2016. Whole-genome sequencing for detecting antimicrobial resistance in nontyphoidal *Salmonella*. *Antimicrob Agents Chemother* 60:5515–5520. <https://doi.org/10.1128/AAC.01030-16>.
26. Hunt M, Mather AE, Sánchez-Busó L, Page AJ, Parkhill J, Keane JA, Harris SR. 2017. ARIBA: rapid antimicrobial resistance genotyping directly from sequencing reads. *Microb Genom* 3:e000131. <https://doi.org/10.1099/mgen.0.000131>.
27. Niehaus KE, Walker TM, Crook DW, Peto TE, Clifton DA. 2014. Machine learning for the prediction of antibacterial susceptibility in *Mycobacterium tuberculosis*, p 618–621. Abstr 2014 IEEE-EMBS Int Conf Biomed Health Informatics.
28. Stoesser N, Batty EM, Eyre DW, Morgan M, Wyllie DH, Del Ojo Elias C, Johnson JR, Walker AS, Peto TEA, Crook DW. 2013. Predicting antimicrobial susceptibilities for *Escherichia coli* and *Klebsiella pneumoniae* isolates using whole genomic sequence data. *J Antimicrob Chemother* 68: 2234–2244. <https://doi.org/10.1093/jac/dkt180>.
29. Pesesky MW, Hussain T, Wallace M, Patel S, Andleeb S, Burnham C-AD, Dantas G. 2016. Evaluation of machine learning and rules-based approaches for predicting antimicrobial resistance profiles in Gram-negative bacilli from whole genome sequence data. *Front Microbiol* 7:1887. <https://doi.org/10.3389/fmicb.2016.01887>.
30. Lipworth SIW, Hough N, Leach L, Morgan M, Jeffrey K, Andersson M, Robinson E, Smith G, Crook D, Peto T. 2018. Whole genome sequencing for predicting *Mycobacterium abscessus* drug susceptibility. *bioRxiv* 251918. <https://doi.org/10.1101/251918>.
31. Bradley P, Gordon NC, Walker TM, Dunn L, Heys S, Huang B, Earle S, Pankhurst LJ, Anson L, De Cesare M. 2015. Rapid antibiotic-resistance predictions from genome sequence data for *Staphylococcus aureus* and *Mycobacterium tuberculosis*. *Nat Commun* 6:10063. <https://doi.org/10.1038/ncomms10063>.
32. Harrison OB, Clemence M, Dillard JP, Tang CM, Trees D, Grad YH, Maiden MCJ. 2016. Genomic analyses of *Neisseria gonorrhoeae* reveal an association of the gonococcal genetic island with antimicrobial resistance. *J Infect* 73:578–587. <https://doi.org/10.1016/j.jinf.2016.08.010>.
33. Grad YH, Harris SR, Kirkcaldy RD, Green AG, Marks DS, Bentley SD, Trees D, Lipsitch M. 2016. Genomic epidemiology of gonococcal resistance to extended-spectrum cephalosporins, macrolides, and fluoroquinolones in the United States, 2000–2013. *J Infect Dis* 214:1579–1587. <https://doi.org/10.1093/infdis/jiw420>.
34. Coelho JR, Carriço JA, Knight D, Martínez J-L, Morrissey I, Oggioni MR, Freitas AT. 2013. The use of machine learning methodologies to analyse antibiotic and biocide susceptibility in *Staphylococcus aureus*. *PLoS One* 8:e55582. <https://doi.org/10.1371/journal.pone.0055582>.
35. Eyre DW, De Silva D, Cole K, Peters J, Cole MJ, Grad YH, Demczuk W, Martin I, Mulvey MR, Crook DW, Walker AS, Peto TEA, Paul J. 2017. WGS to predict antibiotic MICs for *Neisseria gonorrhoeae*. *J Antimicrob Chemother* 72:1937–1947. <https://doi.org/10.1093/jac/dkx067>.
36. Drouin A, Giguère S, Deraspe M, Marchand M, Tyers M, Loo VG, Bourgault A-M, Laviolette F, Corbeil J. 2016. Predictive computational phenotyping and biomarker discovery using reference-free genome comparisons. *BMC Genomics* 17:754. <https://doi.org/10.1186/s12864-016-2889-6>.
37. Davis JJ, Boisvert S, Brettin T, Kenyon RW, Mao C, Olson R, Overbeek R, Santerre J, Shukla M, Wattam AR, Will R, Xia F, Stevens R. 2016. Antimicrobial resistance prediction in PATRIC and RAST. *Sci Rep* 6:27930. <https://doi.org/10.1038/srep27930>.

38. Nguyen M, Brettin T, Long SW, Musser JM, Olsen RJ, Olson R, Shukla M, Stevens RL, Xia F, Yoo H. 2018. Developing an in silico minimum inhibitory concentration panel test for *Klebsiella pneumoniae*. *Sci Rep* 8:421. <https://doi.org/10.1038/s41598-017-18972-w>.
39. Metcalf BJ, Chochua S, Gertz R, Li Z, Walker H, Tran T, Hawkins PA, Glennen A, Lynfield R, Li Y. 2016. Using whole genome sequencing to identify resistance determinants and predict antimicrobial resistance phenotypes for year 2015 invasive pneumococcal disease isolates recovered in the United States. *Clin Microbiol Infect* 22:1002.e1–1002.e8. <https://doi.org/10.1016/j.cmi.2016.08.001>.
40. Li Y, Metcalf BJ, Chochua S, Li Z, Gertz RE, Walker H, Hawkins PA, Tran T, Whitney CG, McGee L. 2016. Penicillin-binding protein transpeptidase signatures for tracking and predicting β -lactam resistance levels in *Streptococcus pneumoniae*. *mBio* 7(3):e00756-16. <https://doi.org/10.1128/mBio.00756-16>.
41. Chen T, Guestrin C. 2016. Xgboost: a scalable tree boosting system, p 785–794. *Proc 22nd ACM SIGKDD Int Conf Knowledge Discovery Data Mining ACM*.
42. U.S. Food and Drug Administration. 2011. National Antimicrobial Resistance Monitoring System—Enteric Bacteria (NARMS): executive report. U.S. Department of Health and Human Services. Food and Drug Administration, Rockville, MD.
43. Wattam AR, Davis JJ, Assaf R, Boisvert S, Brettin T, Bun C, Conrad N, Dietrich EM, Disz T, Gabbard JL, Gerdes S, Henry CS, Kenyon RW, Machi D, Mao C, Nordberg EK, Olsen GJ, Murphy-Olson DE, Olson R, Overbeek R, Parrello B, Pusch GD, Shukla M, Vonstein V, Warren A, Xia F, Yoo H, Stevens RL. 2017. Improvements to PATRIC, the all-bacterial bioinformatics database and analysis resource center. *Nucleic Acids Res* 45:D535–D542. <https://doi.org/10.1093/nar/gkw1017>.
44. Nikolenko SI, Korobeynikov AI, Alekseyev MA. 2013. BayesHammer: Bayesian clustering for error correction in single-cell sequencing. *BMC Genomics* 14:S7. <https://doi.org/10.1186/1471-2164-14-S1-S7>.
45. Bankevich A, Nurk S, Antipov D, Gurevich AA, Dvorkin M, Kulikov AS, Lesin VM, Nikolenko SI, Pham S, Pribelski AD, Pyshkin AV, Sirotkin AV, Vyahhi N, Tesler G, Alekseyev MA, Pevzner PA. 2012. SPAdes: a new genome assembly algorithm and its applications to single-cell sequencing. *J Comput Biol* 19:455–477. <https://doi.org/10.1089/cmb.2012.0021>.
46. Brettin T, Davis JJ, Disz T, Edwards RA, Gerdes S, Olsen GJ, Olson R, Overbeek R, Parrello B, Pusch GD, Shukla M, Thomason JA, Stevens R, Vonstein V, Wattam AR, Xia F. 2015. RASTtk: a modular and extensible implementation of the RAST algorithm for building custom annotation pipelines and annotating batches of genomes. *Sci Rep* 5:8365. <https://doi.org/10.1038/srep08365>.
47. Antonopoulos DA, Assaf R, Aziz RK, Brettin T, Bun C, Conrad N, Davis JJ, Dietrich EM, Disz T, Gerdes S, Kenyon RW, Machi D, Mao C, Murphy-Olson DE, Nordberg EK, Olsen GJ, Olson R, Overbeek R, Parrello B, Pusch GD, Santerre J, Shukla M, Stevens RL, VanOeffelen M, Vonstein V, Warren AS, Wattam AR, Xia F, Yoo H. 31 July 2017. PATRIC as a unique resource for studying antimicrobial resistance. *Brief Bioinform* <https://doi.org/10.1093/bib/bbx083>.
48. Katoh K, Standley DM. 2013. MAFFT multiple sequence alignment software version 7: improvements in performance and usability. *Mol Biol Evol* 30:772–780. <https://doi.org/10.1093/molbev/mst010>.
49. Price MN, Dehal PS, Arkin AP. 2010. FastTree 2—approximately maximum-likelihood trees for large alignments. *PLoS One* 5:e9490. <https://doi.org/10.1371/journal.pone.0009490>.
50. Letunic I, Bork P. 2007. Interactive Tree of Life (iTOL): an online tool for phylogenetic tree display and annotation. *Bioinformatics* 23:127–128. <https://doi.org/10.1093/bioinformatics/btl529>.
51. Deorowicz S, Kokot M, Grabowski S, Debudaj-Grabysz A. 2015. KMC 2: fast and resource-frugal k-mer counting. *Bioinformatics* 31:1569–1576. <https://doi.org/10.1093/bioinformatics/btv022>.
52. U.S. Food and Drug Administration. 2009. Guidance for industry and FDA. Class II special controls guidance document: antimicrobial susceptibility test (AST) systems. Center for Devices and Radiological Health, U.S. Food and Drug Administration, U.S. Department of Health and Human Services, Silver Spring, MD.
53. Jorgensen JH. 1993. Selection criteria for an antimicrobial susceptibility testing system. *J Clin Microbiol* 31:2841.
54. U.S. Food and Drug Administration. 2009. Class II special controls guidance document: antimicrobial susceptibility test (AST) systems. U.S. Food and Drug Administration, Rockville, MD.
55. Pedregosa F, Varoquaux G, Gramfort A, Michel V, Thirion B, Grisel OBM, Prettenhofer P, Weiss R, Dubourg V, Vanderplas J, Passos A, Cournapeau D, Brucher M, Perrot M, Duchesnay É. 2011. Scikit-learn: machine learning in Python. *J Machine Learning Res* 12:2825–2830.
56. Bellman R. 2013. Dynamic programming. Courier Corporation, New York, NY.
57. Shalev-Shwartz S, Ben-David S. 2014. Understanding machine learning: from theory to algorithms. Cambridge University Press, Cambridge, United Kingdom.
58. Aggarwal CC, Hinneburg A, Keim DA (ed). 2001. On the surprising behavior of distance metrics in high dimensional space. *In International conference on database theory*. Springer-Verlag, Berlin, Germany.
59. Camacho C, Coulouris G, Avagyan V, Ma N, Papadopoulos J, Bealer K, Madden TL. 2009. BLAST+: architecture and applications. *BMC Bioinformatics* 10:421. <https://doi.org/10.1186/1471-2105-10-421>.
60. Waterhouse AM, Procter JB, Martin DM, Clamp M, Barton GJ. 2009. Jalview version 2—a multiple sequence alignment editor and analysis workbench. *Bioinformatics* 25:1189–1191. <https://doi.org/10.1093/bioinformatics/btp033>.
61. Crooks GE, Hon G, Chandonia J-M, Brenner SE. 2004. WebLogo: a sequence logo generator. *Genome Res* 14:1188–1190. <https://doi.org/10.1101/gr.849004>.
62. Davis JJ, Olsen GJ. 2010. Modal codon usage: assessing the typical codon usage of a genome. *Mol Biol Evol* 27:800–810. <https://doi.org/10.1093/molbev/msp281>.
63. Davis JJ, Olsen GJ. 2011. Characterizing the native codon usages of a genome: an axis projection approach. *Mol Biol Evol* 28:211–221. <https://doi.org/10.1093/molbev/msq185>.
64. Ranieri ML, Shi C, Switt AIM, Den Bakker HC, Wiedmann M. 2013. Comparison of typing methods with a new procedure based on sequence characterization for *Salmonella* serovar prediction. *J Clin Microbiol* 51:1786–1797. <https://doi.org/10.1128/JCM.03201-12>.
65. Zhu L, Olsen RJ, Nasser W, Beres SB, Vuopio J, Kristinsson KG, Gottfredsson M, Porter AR, DeLeo FR, Musser JM. 2015. A molecular trigger for intercontinental epidemics of group A *Streptococcus*. *J Clin Invest* 125:3545–3559. <https://doi.org/10.1172/JCI82478>.
66. Nasser W, Beres SB, Olsen RJ, Dean MA, Rice KA, Long SW, Kristinsson KG, Gottfredsson M, Vuopio J, Raisanen K, Caugant DA, Steinbakk M, Low DE, McGeer A, Darenberg J, Henriques-Normark B, Van Beneden CA, Hoffmann S, Musser JM. 2014. Evolutionary pathway to increased virulence and epidemic group A *Streptococcus* disease derived from 3,615 genome sequences. *Proc Natl Acad Sci U S A* 111:E1768–E1776. <https://doi.org/10.1073/pnas.1403138111>.
67. Yoshida H, Kojima T, Yamagishi J-I, Nakamura S. 1988. Quinolone-resistant mutations of the gyrA gene of *Escherichia coli*. *Mol Gen Genet* 211:1–7. <https://doi.org/10.1007/BF00338386>.
68. Yoshida H, Bogaki M, Nakamura M, Nakamura S. 1990. Quinolone resistance-determining region in the DNA gyrase gyrA gene of *Escherichia coli*. *Antimicrob Agents Chemother* 34:1271–1272. <https://doi.org/10.1128/AAC.34.6.1271>.
69. Tyson GH, Zhao S, Li C, Ayers S, Sabo JL, Lam C, Miller RA, McDermott PF. 2017. Establishing genotypic cutoff values to measure antimicrobial resistance in *Salmonella*. *Antimicrob Agents Chemother* 61:e02140-16. <https://doi.org/10.1128/AAC.02140-16>.
70. Kohanski MA, Dwyer DJ, Hayete B, Lawrence CA, Collins JJ. 2007. A common mechanism of cellular death induced by bactericidal antibiotics. *Cell* 130:797–810. <https://doi.org/10.1016/j.cell.2007.06.049>.
71. Foti JJ, Devadoss B, Winkler JA, Collins JJ, Walker GC. 2012. Oxidation of the guanine nucleotide pool underlies cell death by bactericidal antibiotics. *Science* 336:315–319. <https://doi.org/10.1126/science.1219192>.