



Machine Learning

Gaussian Mixture Model (GMM)

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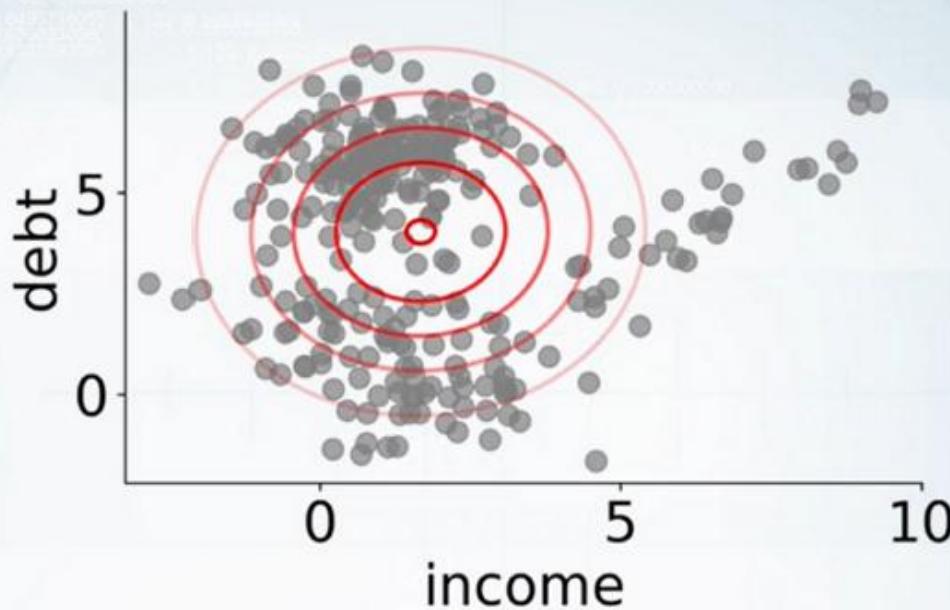
<https://www.aparat.com/mehran.safayani>



https://github.com/safayani/machine_learning_course



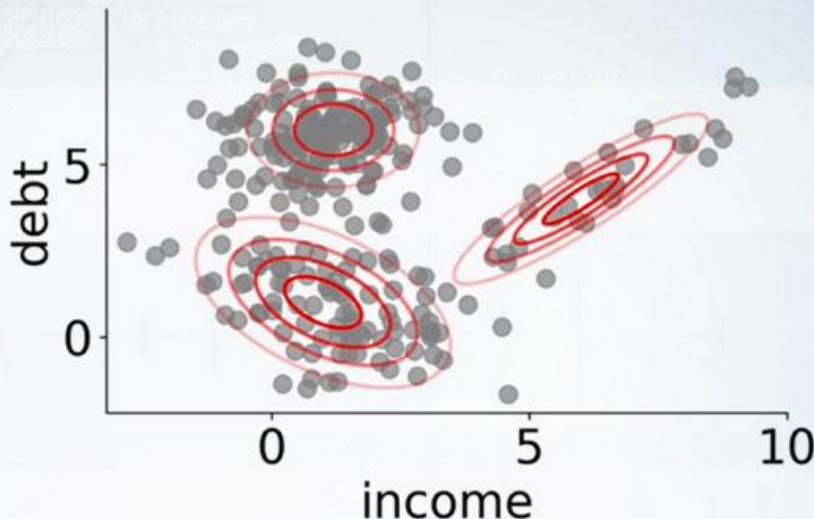
Probabilistic model of data



$$p(x \mid \theta) = \mathcal{N}(x \mid \mu, \Sigma)$$

$$\theta = \{\mu, \Sigma\}$$

Gaussian Mixture Model (GMM)



$$p(x | \theta) = \pi_1 \mathcal{N}(x | \mu_1, \Sigma_1) + \pi_2 \mathcal{N}(x | \mu_2, \Sigma_2) + \pi_3 \mathcal{N}(x | \mu_3, \Sigma_3)$$

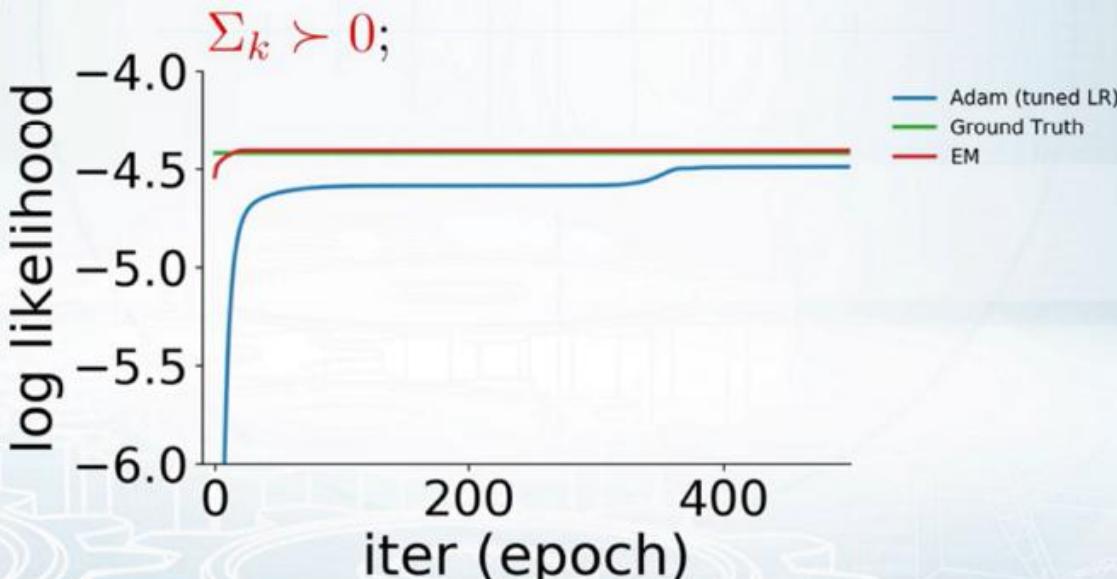
$$\theta = \{\pi_1, \pi_2, \pi_3, \mu_1, \mu_2, \mu_3, \Sigma_1, \Sigma_2, \Sigma_3\}$$

	Gaussian	GMM
Flexibility		
# of parameters		
Parameters	μ, Σ	$\{\pi_1, \pi_2, \pi_3\}$ $\{\mu_1, \mu_2, \mu_3\}$

Training GMM

$$\max_{\theta} \prod_{i=1}^N p(x_i | \theta) = \prod_{i=1}^N (\pi_1 \mathcal{N}(x_i | \mu_1, \Sigma_1) + \dots)$$

subject to $\pi_1 + \pi_2 + \pi_3 = 1; \pi_k \geq 0; k = 1, 2, 3.$



Introducing latent variable

$$p(x | \theta) = \pi_1 \mathcal{N}(x | \mu_1, \Sigma_1) + \pi_2 \mathcal{N}(x | \mu_2, \Sigma_2) + \pi_3 \mathcal{N}(x | \mu_3, \Sigma_3)$$



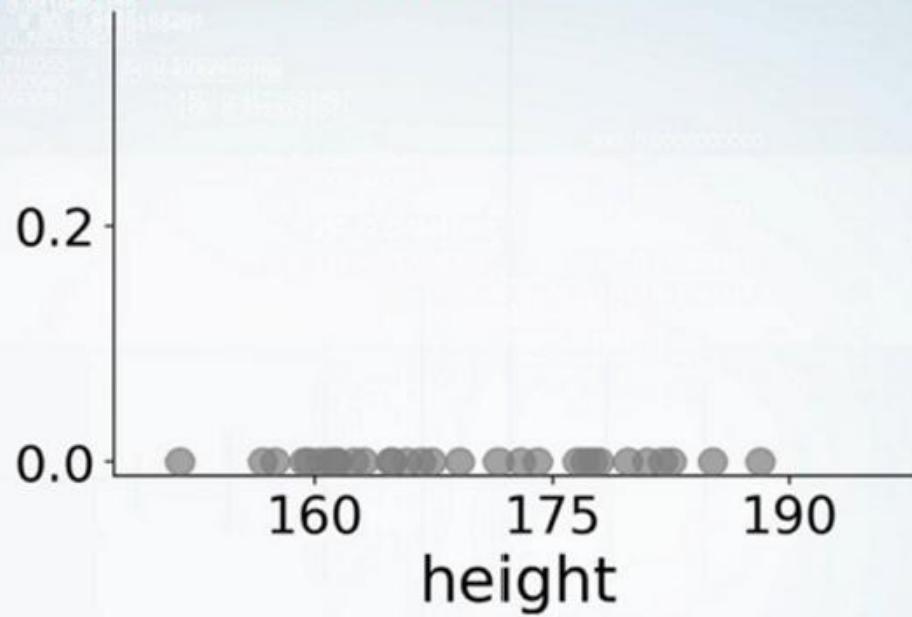
$$p(t = c | \theta) = \pi_c$$

$$p(x | t = c, \theta) = \mathcal{N}(x | \mu_c, \Sigma_c)$$

$$p(x | \theta) = \sum_{c=1}^3 p(x | t = c, \theta) p(t = c | \theta)$$

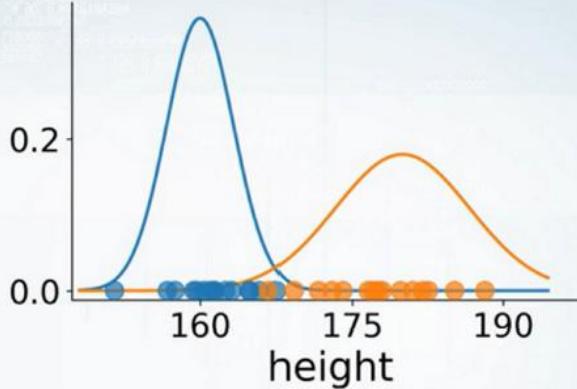
$p(n, t)$

Expectation Maximization



How to estimate parameter θ ?

Expectation Maximization



How to estimate parameter θ ?

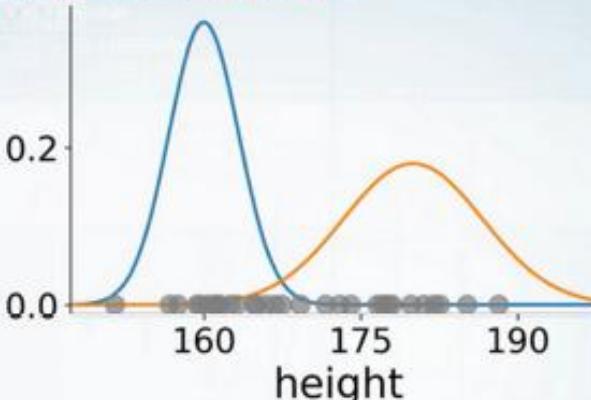
If sources t are known, easy:

$$p(x \mid t = 1, \theta) = \mathcal{N}(x \mid \mu_1, \sigma_1^2)$$

$$\mu_1 = \frac{\sum_{\text{blue } i} x_i}{\# \text{ of blue points}} \quad \sigma_1^2 = \frac{\sum_{\text{blue } i} (x_i - \mu_1)^2}{\# \text{ of blue points}}$$

$$\mu_1 = \frac{\sum_i p(t_i = 1 \mid x_i, \theta) x_i}{\sum_i p(t_i = 1 \mid x_i, \theta)} \quad \sigma_1^2 = \frac{\sum_i p(t_i = 1 \mid x_i, \theta) (x_i - \mu_1)^2}{\sum_i p(t_i = 1 \mid x_i, \theta)}$$

Expectation Maximization



What if we don't know the sources?

Given: $p(x | t = 1, \theta) = \mathcal{N}(-2, 1)$

Find: $p(t = 1 | x, \theta)$

$$p(t = 1 | x, \theta) = \frac{p(x | t = 1, \theta) p(t = 1 | \theta)}{Z}$$

Chicken and egg problem

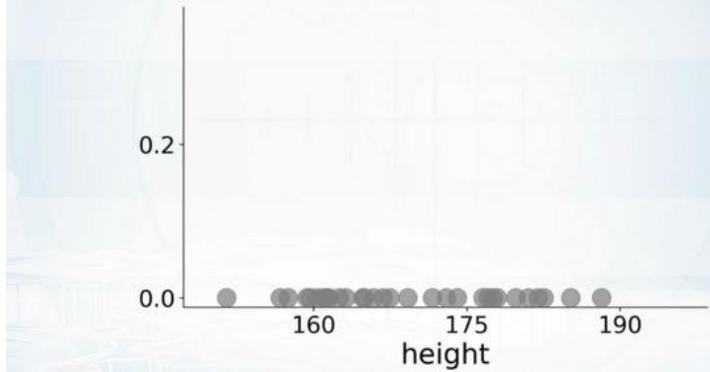
- Need Gaussian parameters to estimate sources
- Need sources to estimate Gaussian parameters



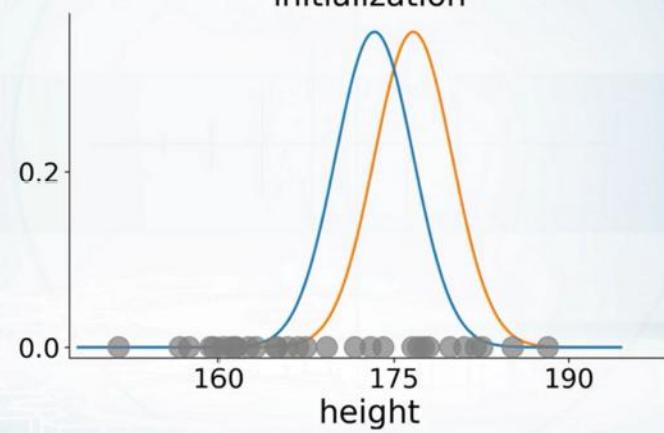
EM algorithm

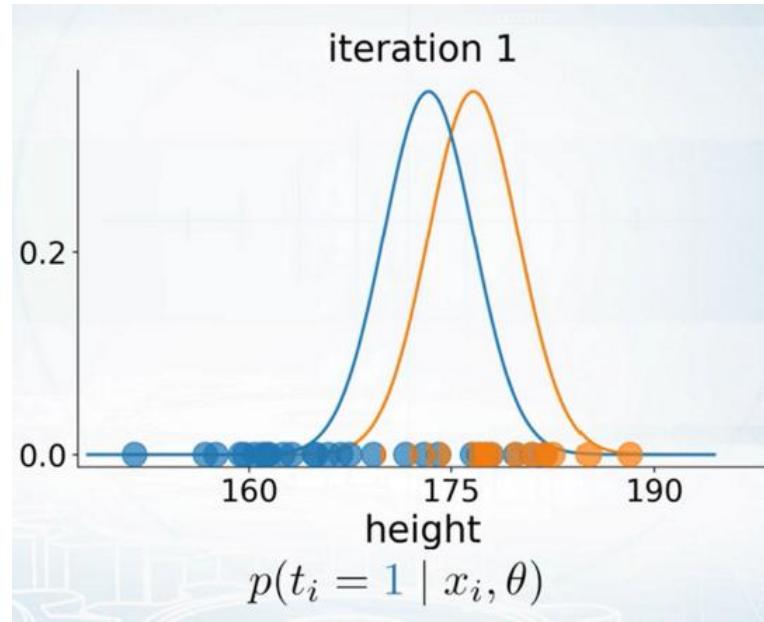
1. Start with 2 randomly placed Gaussians parameters θ
2. Until convergence repeat:
 - a) For each point compute $p(t = c \mid x_i, \theta)$: does x_i look like it came from cluster c ?
 - b) Update Gaussian parameters θ to fit points assigned to them

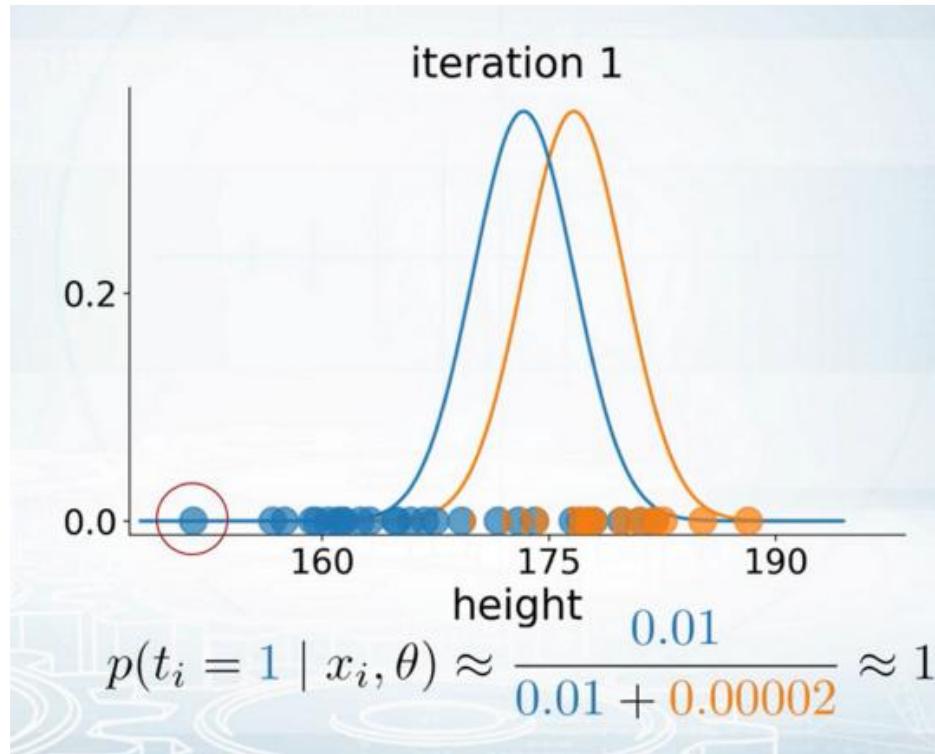
GMM EM example



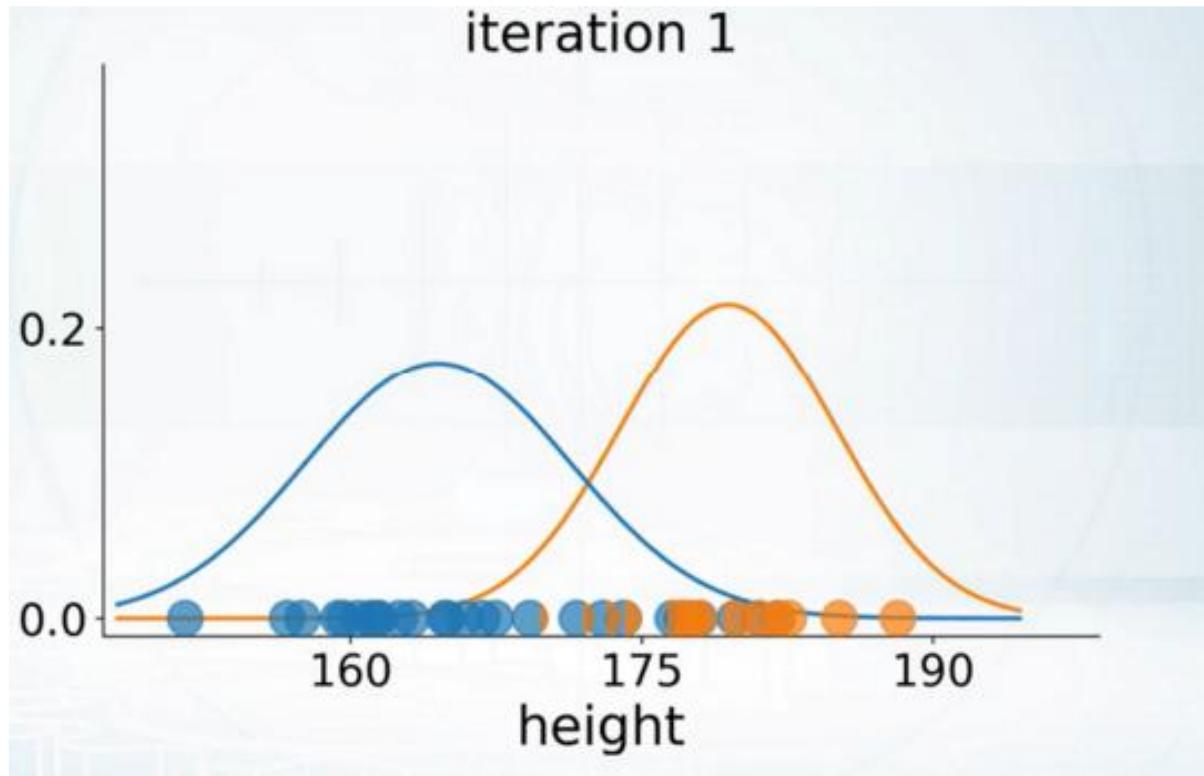
initialization



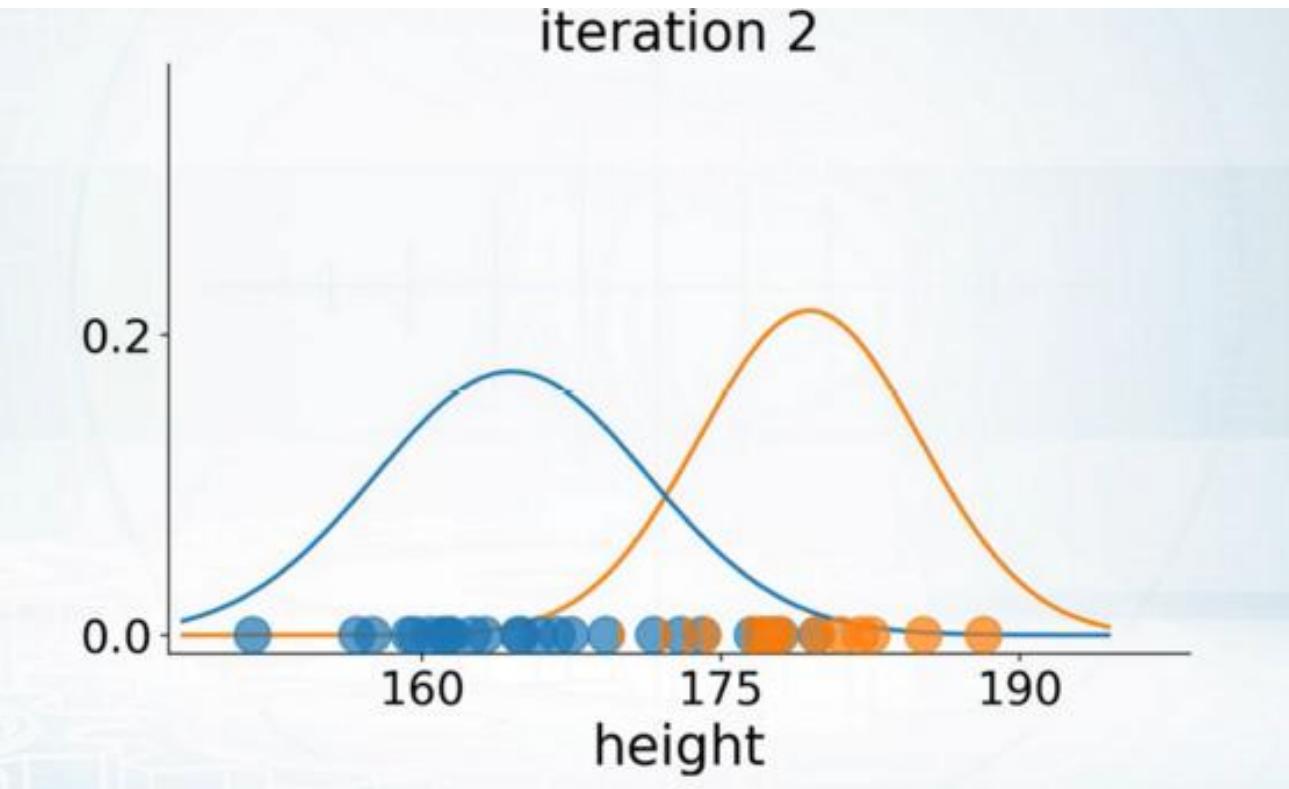




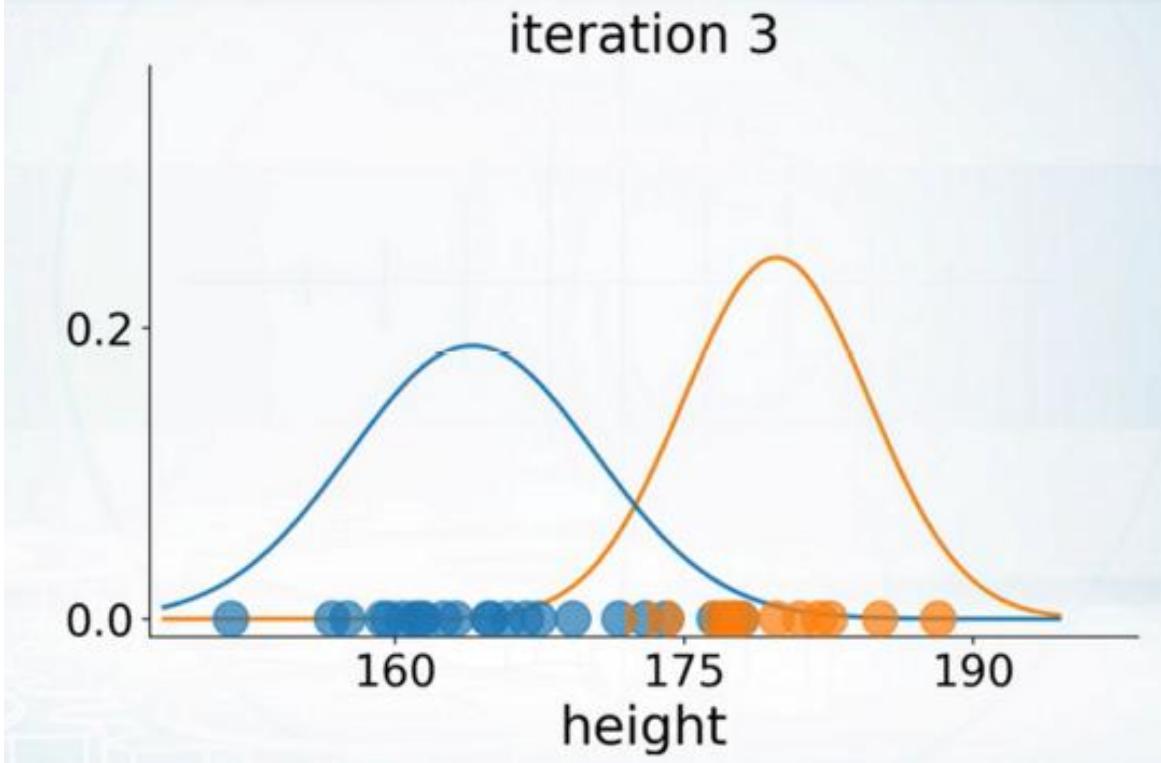
iteration 1



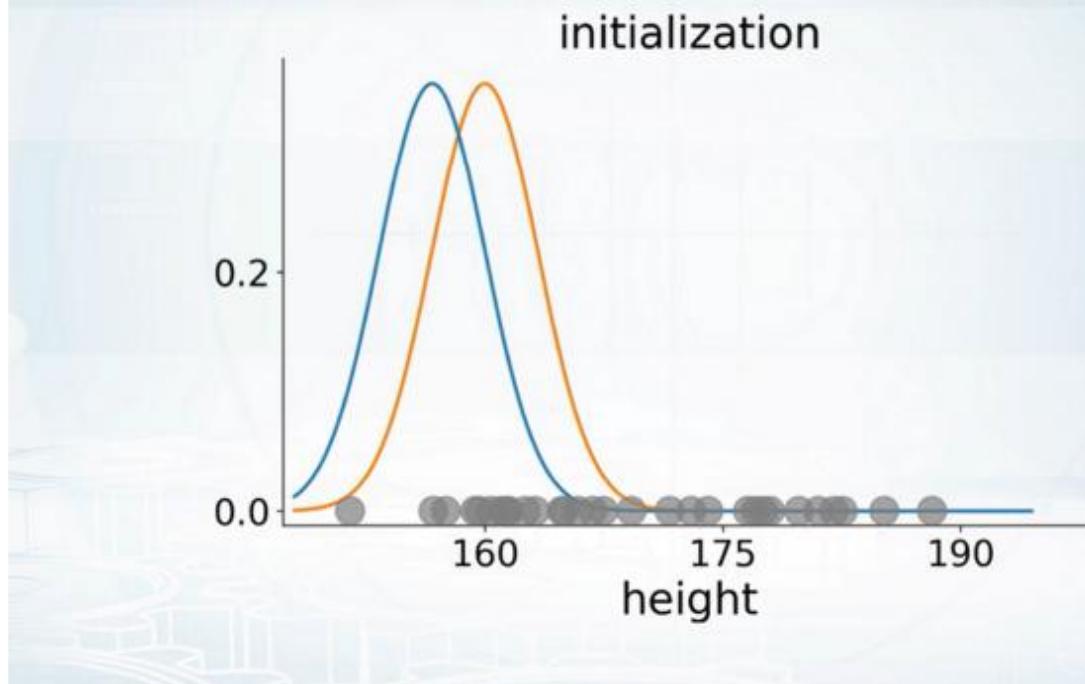
iteration 2



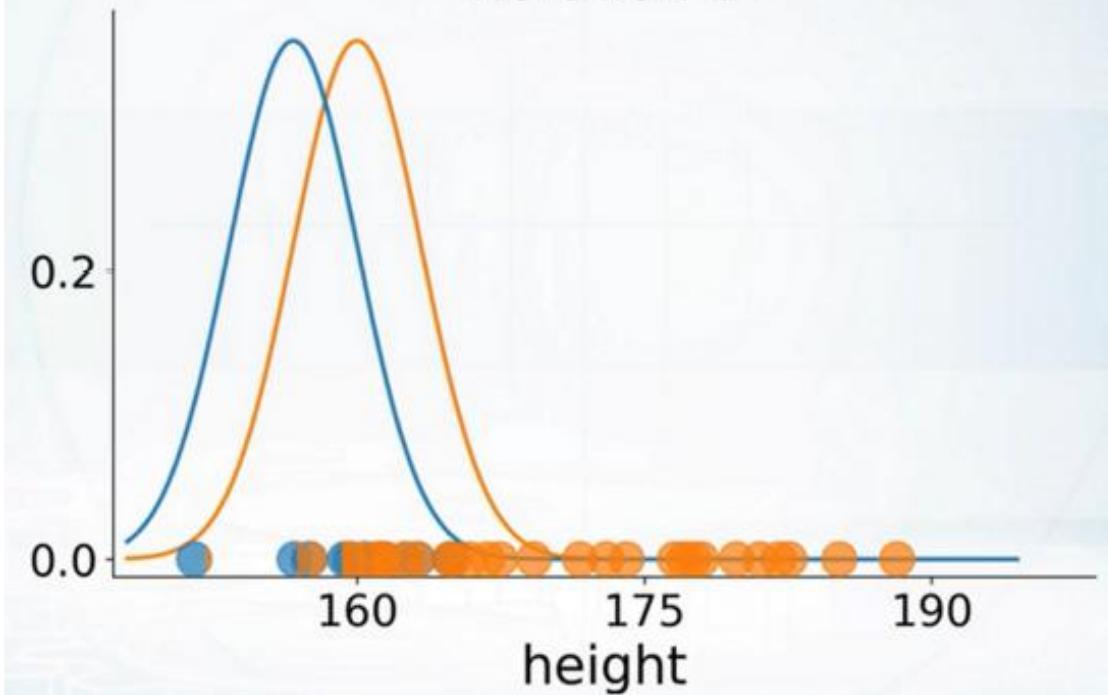
iteration 3



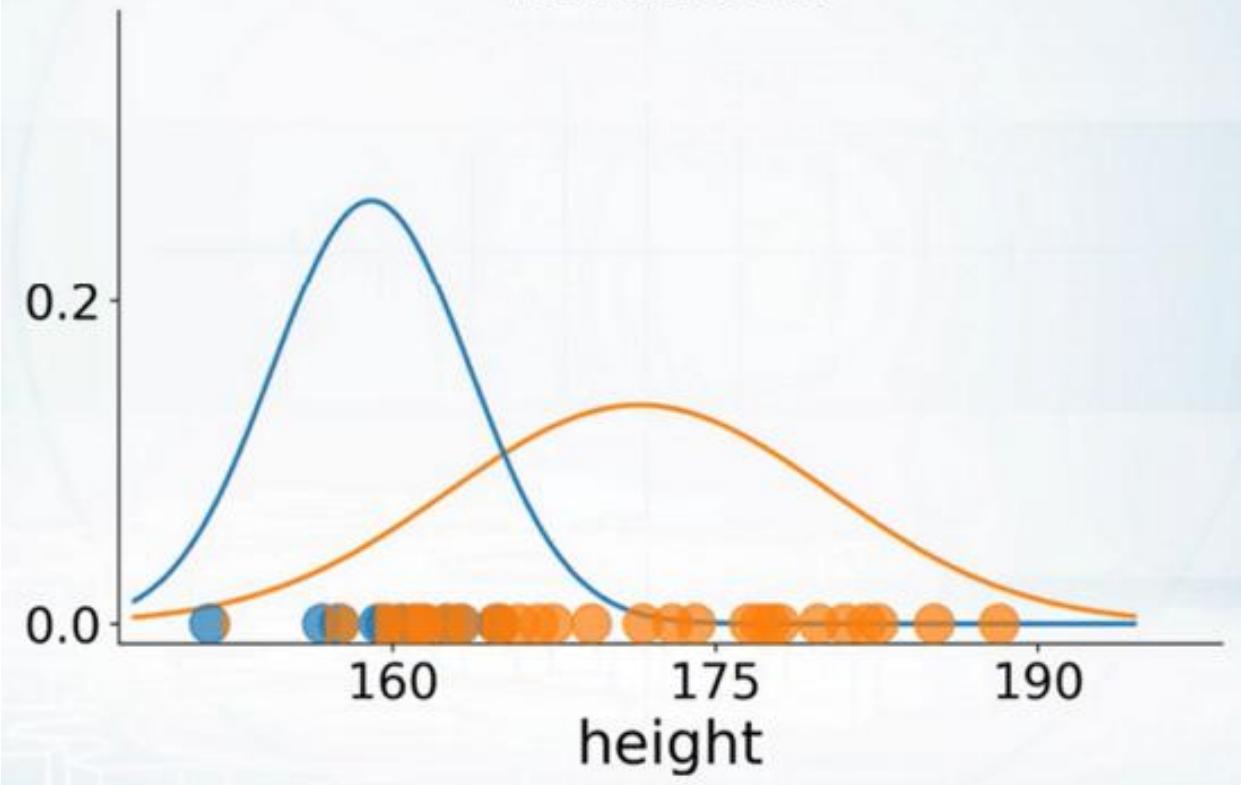
GMM EM local maximum example



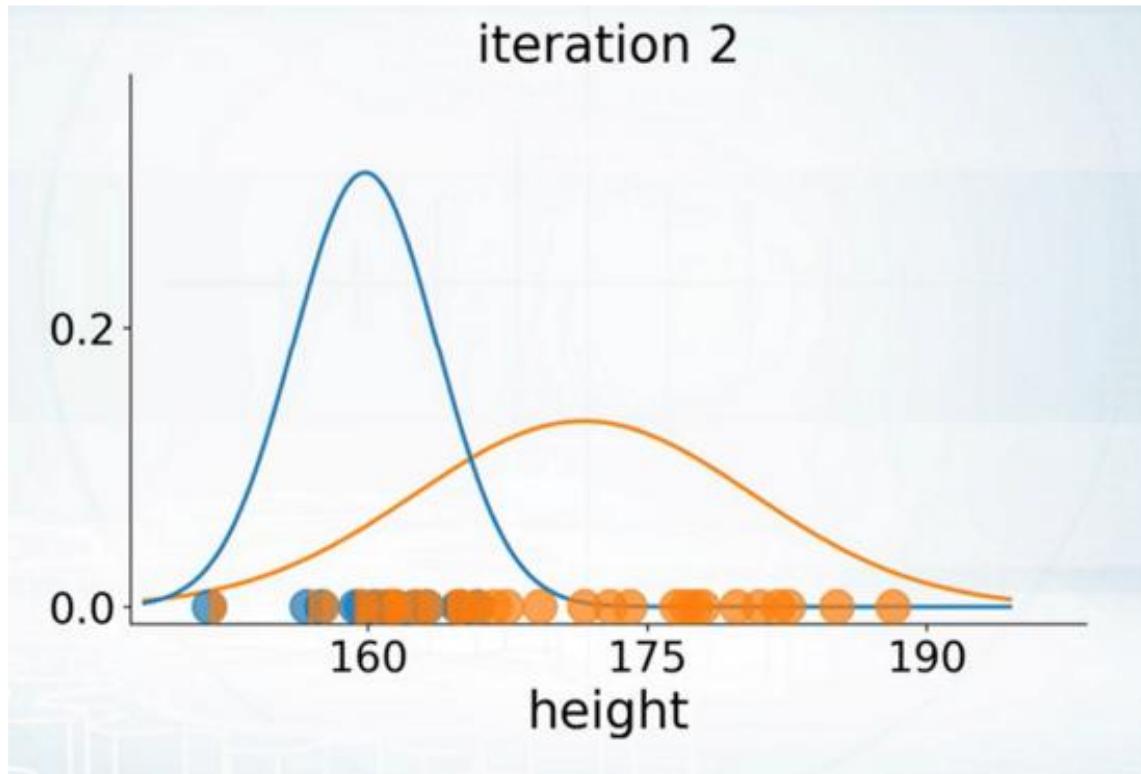
iteration 1



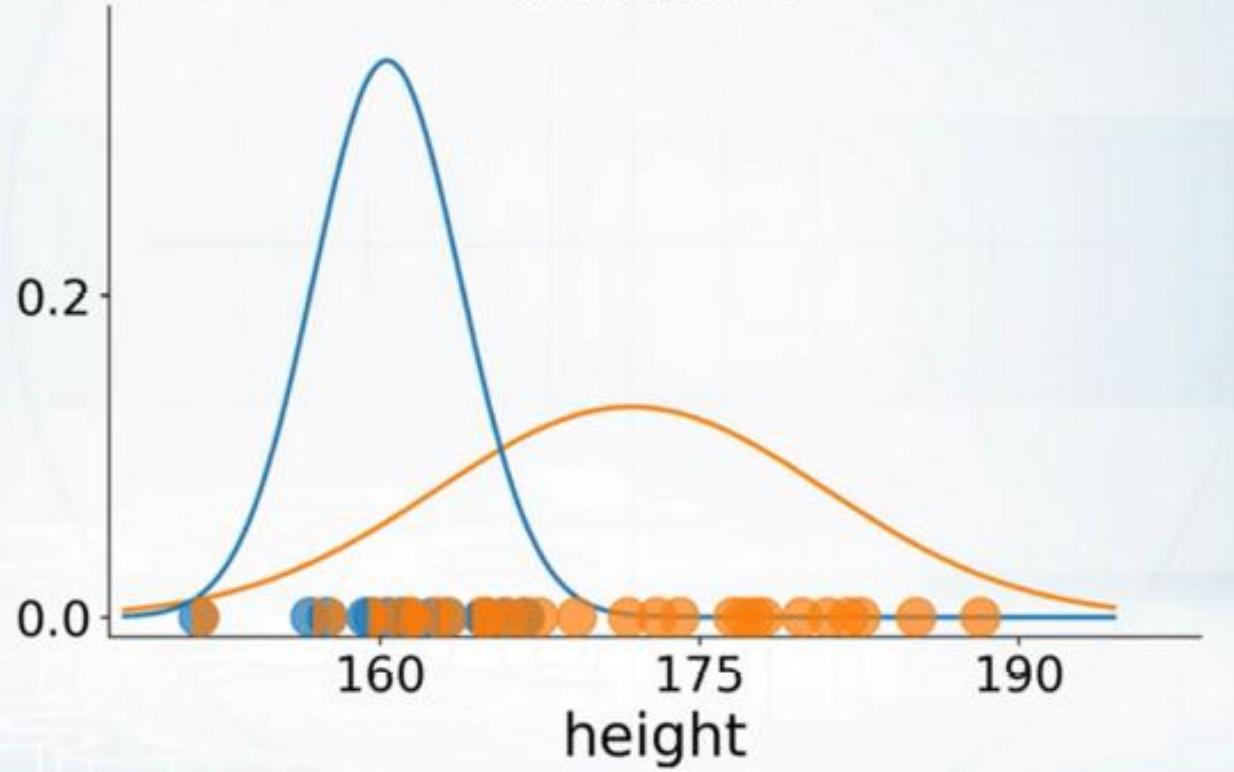
iteration 1



iteration 2

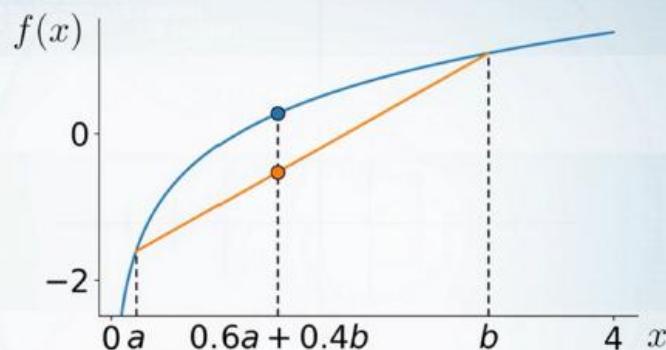


iteration 3



General form of Expectation Maximization

Concave functions



Def.: $f(x)$ is concave if

for any $a, b, \alpha : f(\alpha a + (1 - \alpha)b) \geq \alpha f(a) + (1 - \alpha)f(b)$
 $0 \leq \alpha \leq 1$

Jensen's inequality

If $f(\alpha a + (1 - \alpha)b) \geq \alpha f(a) + (1 - \alpha)f(b)$

Then $\alpha_1 + \alpha_2 + \alpha_3 = 1; \alpha_k \geq 0.$

$$f(\alpha_1 a_1 + \alpha_2 a_2 + \alpha_3 a_3) \geq \alpha_1 f(a_1) + \alpha_2 f(a_2) + \alpha_3 f(a_3)$$

$$f(\underbrace{\alpha_1 a_1 + \alpha_2 a_2 + \alpha_3 a_3}_{\mathbb{E}_{p(t)} t}) \geq \underbrace{\alpha_1 f(a_1) + \alpha_2 f(a_2) + \alpha_3 f(a_3)}_{\mathbb{E}_{p(t)} f(t)}$$

$$p(t = a_1) = \alpha_1,$$

$$p(t = a_2) = \alpha_2,$$

$$p(t = a_3) = \alpha_3$$

Jensen's inequality

If $f(\alpha a + (1 - \alpha)b) \geq \alpha f(a) + (1 - \alpha)f(b)$

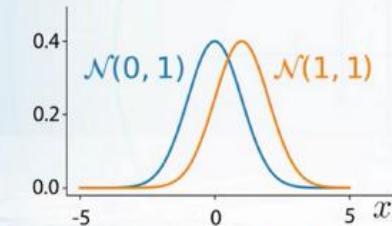
Then Jensen's inequality:

$$f(\mathbb{E}_{p(t)} t) \geq \mathbb{E}_{p(t)} f(t)$$

Kullback–Leibler divergence

Parameters difference: 1

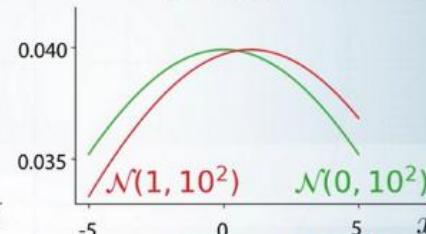
$$\mathcal{KL}(q_1 \parallel p_1) = 0.5$$



$$|\mu_1 - \mu_2| < 1$$

Parameters difference: 1

$$\mathcal{KL}(q_2 \parallel p_2) = 0.005$$



$$|\mu_1 - \mu_2| = 1$$

$$\mathcal{KL}(q \parallel p) = \int q(x) \log \frac{q(x)}{p(x)} dx$$

1. $\mathcal{KL}(q \parallel p) \neq \mathcal{KL}(p \parallel q)$
2. $\mathcal{KL}(q \parallel q) = 0$
3. $\mathcal{KL}(q \parallel p) \geq 0$



Proof: $-\mathcal{KL}(q \parallel p) = \mathbb{E}_q \left(-\log \frac{q}{p} \right) = \mathbb{E}_q \left(\log \frac{p}{q} \right)$

$$\leq \log(\mathbb{E}_q \frac{p}{q}) = \log \int q(x) \frac{p(x)}{q(x)} dx = 0$$

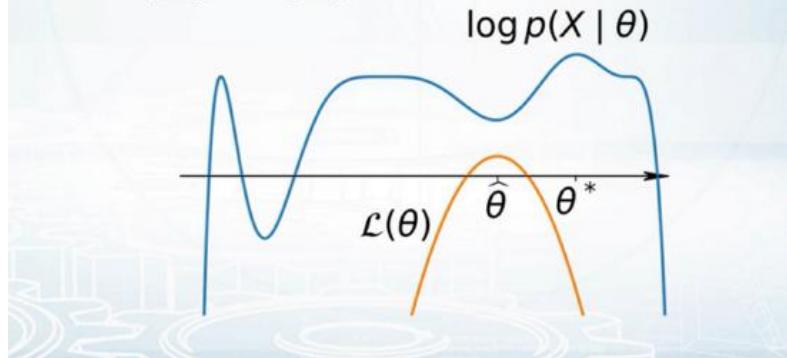
General form of Expectation Maximization



$$p(x_i \mid \theta) = \sum_{c=1}^3 p(x_i \mid t_i = c, \theta) p(t_i = c \mid \theta)$$

$$\begin{aligned}\max_{\theta} \log p(X \mid \theta) &= \log \prod_{i=1}^N p(x_i \mid \theta) \\ &= \sum_{i=1}^N \log p(x_i \mid \theta)\end{aligned}$$

$$\begin{aligned}
 \log p(X \mid \theta) &= \sum_{i=1}^N \log p(x_i \mid \theta) \\
 &= \sum_{i=1}^N \log \sum_{c=1}^3 p(x_i, t_i = c \mid \theta) \geq \mathcal{L}(\theta)
 \end{aligned}$$



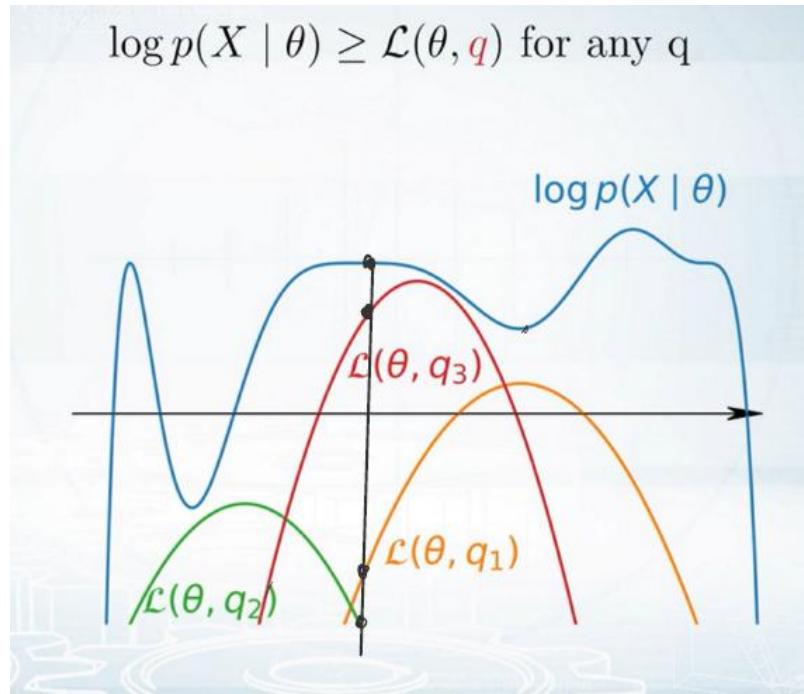
$$\begin{aligned}
\log p(X \mid \theta) &= \sum_{i=1}^N \log p(x_i \mid \theta) \\
&= \sum_{i=1}^N \log \sum_{c=1}^3 \frac{q(t_i = c)}{q(t_i = c)} p(x_i, t_i = c \mid \theta) \\
&\geq \sum_{i=1}^N \sum_{c=1}^3 q(t_i = c) \log \frac{p(x_i, t_i = c \mid \theta)}{q(t_i = c)}
\end{aligned}$$

Jensen's inequality

$$\log \left(\sum_c \alpha_c v_c \right) \geq \sum_c \alpha_c \log(v_c)$$

$$\begin{aligned}
\log p(X \mid \theta) &= \sum_{i=1}^N \log p(x_i \mid \theta) \\
&= \sum_{i=1}^N \log \sum_{c=1}^3 \frac{q(t_i = c)}{q(t_i = c)} p(x_i, t_i = c \mid \theta) \\
&\geq \sum_{i=1}^N \sum_{c=1}^3 q(t_i = c) \log \frac{p(x_i, t_i = c \mid \theta)}{q(t_i = c)} \\
&= \mathcal{L}(\theta, q)
\end{aligned}$$

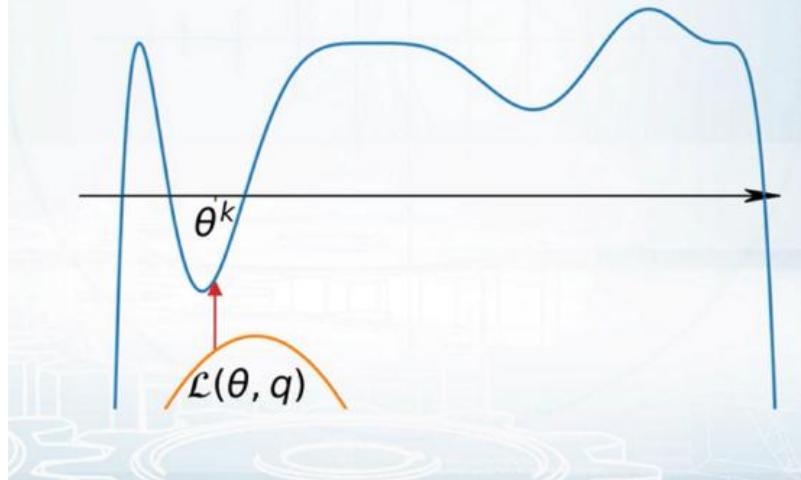
$$\log p(X \mid \theta) \geq \mathcal{L}(\theta, q) \text{ for any } q$$



$$\log p(X \mid \theta) \geq \mathcal{L}(\theta, q) \text{ for any } q$$

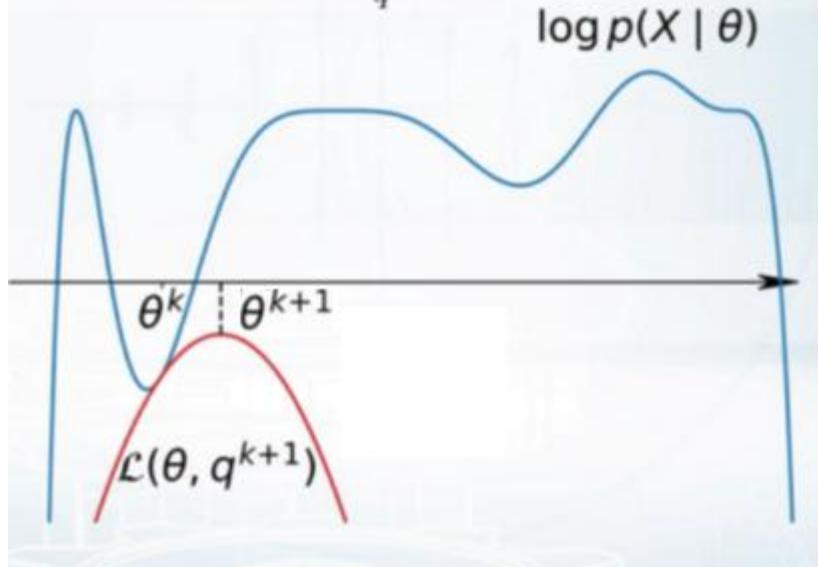
$$q^{k+1} = \arg \max_q \mathcal{L}(\theta^k, q)$$

$$\log p(X \mid \theta)$$



$$\log p(X \mid \theta) \geq \mathcal{L}(\theta, q) \text{ for any } q$$

$$q^{k+1} = \arg \max_q \mathcal{L}(\theta^k, q)$$

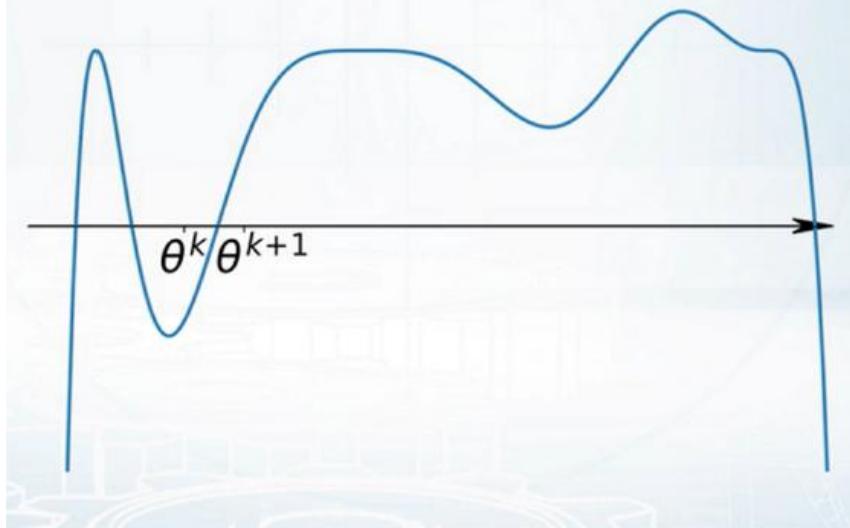


$$\theta^{k+1} = \arg \max_{\theta} \mathcal{L}(\theta, q^{k+1})$$

$\log p(X \mid \theta) \geq \mathcal{L}(\theta, q)$ for any q

$$q^{k+1} = \arg \max_q \mathcal{L}(\theta^k, q)$$

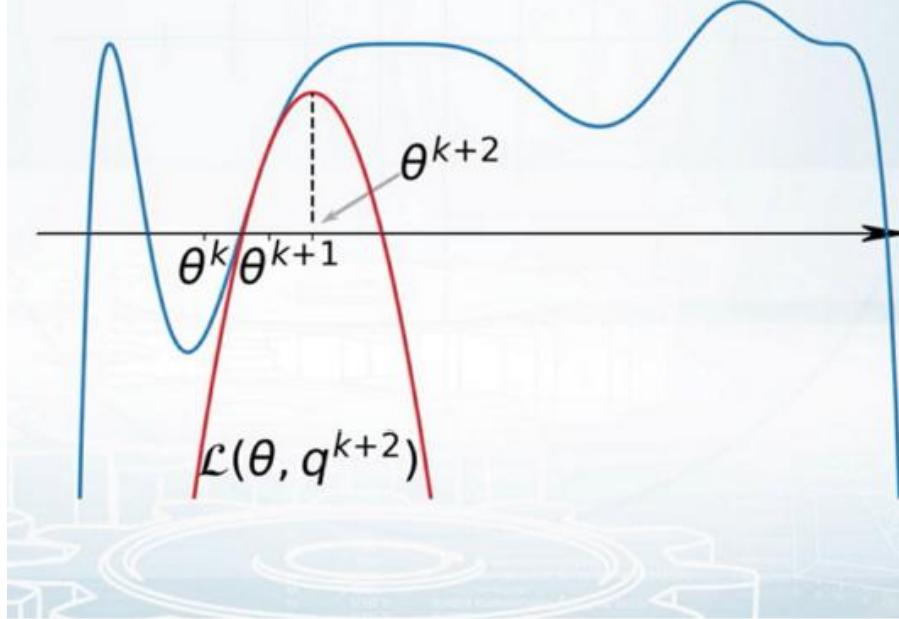
$\log p(X \mid \theta)$



$\log p(X \mid \theta) \geq \mathcal{L}(\theta, q)$ for any q

$$q^{k+1} = \arg \max_q \mathcal{L}(\theta^k, q)$$

$\log p(X \mid \theta)$



Summary of Expectation Maximization

$$\log p(X | \theta) \geq \mathcal{L}(\theta, q) \text{ for any } q$$

↑
Variational
lower bound

E-step

$$q^{k+1} = \arg \max_q \mathcal{L}(\theta^k, q)$$

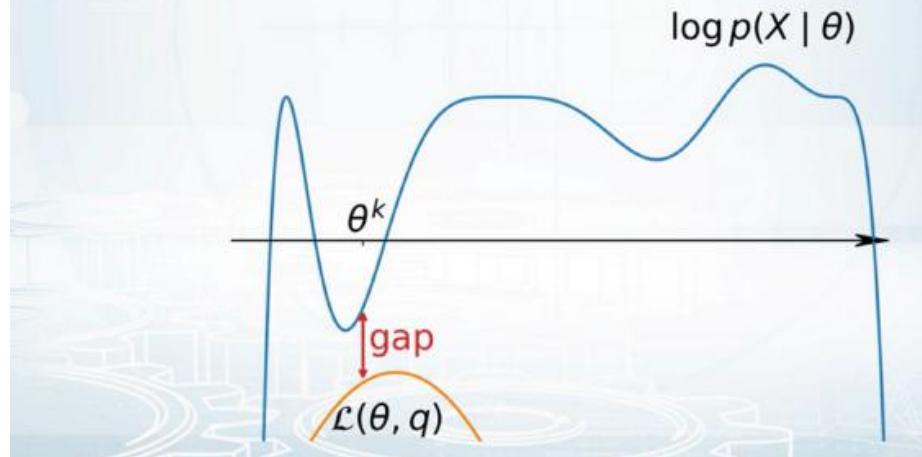
M-step

$$\theta^{k+1} = \arg \max_{\theta} \mathcal{L}(\theta, q^{k+1})$$

E-step details

$$\log p(X \mid \theta) \geq \mathcal{L}(\theta, q)$$

E-step: $\max_q \mathcal{L}(\theta^k, q)$



$$GAP = \log p(X|\theta) - L(\theta, q) = \sum_{i=1}^N \log p(x_i|\theta) - \sum_{i=1}^N \sum_{c=1}^3 q(t_i=c) \log \frac{p(\mathbf{x}_i, t_i=c|\theta)}{q(t_i=c)}$$

$$\sum_{i=1}^N \log p(x_i|\theta) \cdot \sum_{c=1}^3 q(t_i=c) - \sum_{i=1}^N \sum_{c=1}^3 q(t_i=c) \log \frac{p(x_i, t_i=c|\theta)}{q(t_i=c)}$$

$$= - \sum_{i=1}^N \sum_{c=1}^3 q(t_i=c) (\log p(x_i|\theta) - \log \frac{p(x_i, t_i=c|\theta)}{q(t_i=c)})$$

$$= \sum_{i=1}^N \sum_{c=1}^3 q(t_i=c) \log \frac{p(x_i|\theta)q(t_i=c)}{p(x_i, t_i=c|\theta)}$$

$$= \sum_{i=1}^N \sum_{c=1}^3 q(t_i=c) \log \frac{q(t_i=c)}{p(t_i=c|x_i, \theta)} = \sum_{i=1}^N \sum_{c=1}^3 KL(q(t_i=c)||p(t_i=c|x_i, \theta))$$

Estep:

$$GAP = \log p(X|\theta) - L(\theta, q) = \log p(X|\theta) - ELBO$$

$$= \sum_{i=1}^N \sum_{c=1}^3 KL(q(t_i = c) || p(t_i = c | x_i, \theta)) >> 0$$

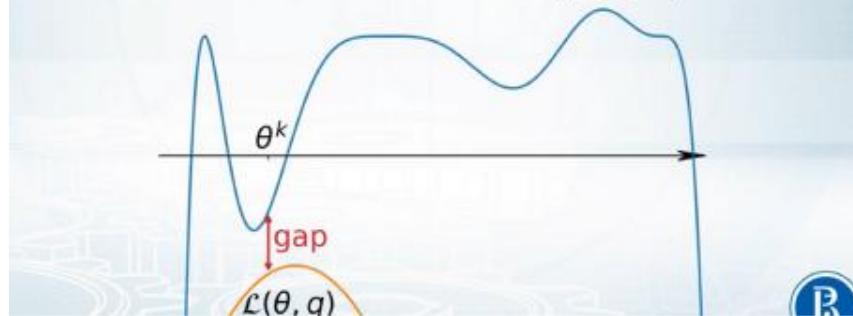
We maximize **ELBO** or minimize **KL**

$$\log p(X|\theta) = \sum_{i=1}^N \sum_{c=1}^3 KL(q(t_i = c) || p(t_i = c | x_i, \theta)) + ELBO = GAP + ELBO$$

Solution: $q(t_i = c) = p(t_i = c | x_i, \theta)$

$$\log p(X \mid \theta) - \mathcal{L}(\theta, q) = \sum_i \mathcal{KL}(q(t_i) \parallel p(t_i \mid x_i, \theta))$$

$$\text{E-step: } \arg \max_{q(t_i)} \mathcal{L}(\theta^k, q) = \frac{p(t_i \mid x_i, \theta)}{\log p(X \mid \theta)}$$



M-step details

$$\mathcal{L}(\theta, q) = \sum_i \sum_c q(t_i = c) \log \frac{p(x_i, t_i = c \mid \theta)}{q(t_i = c)}$$

$$= \sum_i \sum_c q(t_i = c) \log p(x_i, t_i = c \mid \theta)$$

$$\cancel{- \sum_i \sum_c q(t_i = c) \log q(t_i = c)}$$

$$= \mathbb{E}_q \log p(X, T \mid \theta) + \text{const}$$

Const w.r.t. θ

$$= \mathbb{E}_q \log p(X, T \mid \theta) + \text{const}$$



(Usually) concave function w.r.t. θ , easy to optimize

Expectation Maximization algorithm

For $k = 1, \dots$

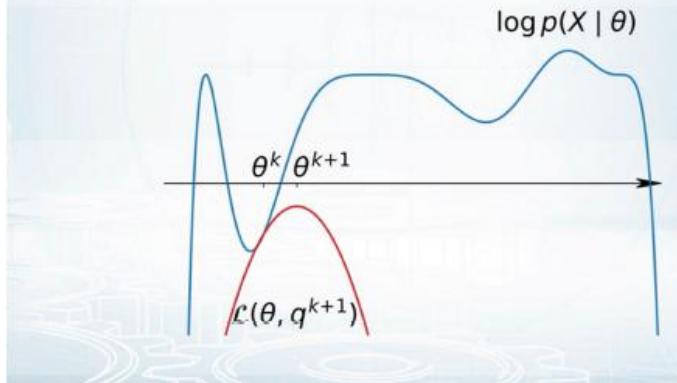
E-step

$$\begin{aligned} q^{k+1} &= \arg \min_q \mathcal{KL} [q(T) \parallel p(T \mid X, \theta^k)] \\ &\Leftrightarrow \\ q^{k+1}(t_i) &= p(t_i \mid x_i, \theta^k) \end{aligned}$$

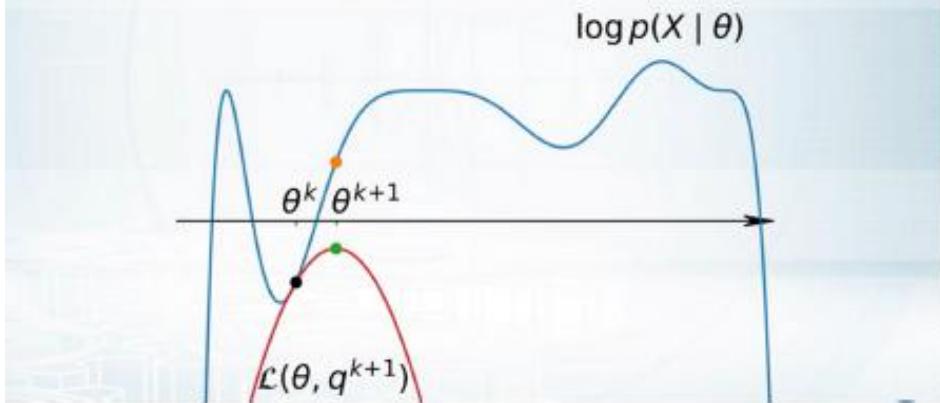
M-step

$$\theta^{k+1} = \arg \max_{\theta} \mathbb{E}_{q^{k+1}} \log p(X, T \mid \theta)$$

Convergence guarantees



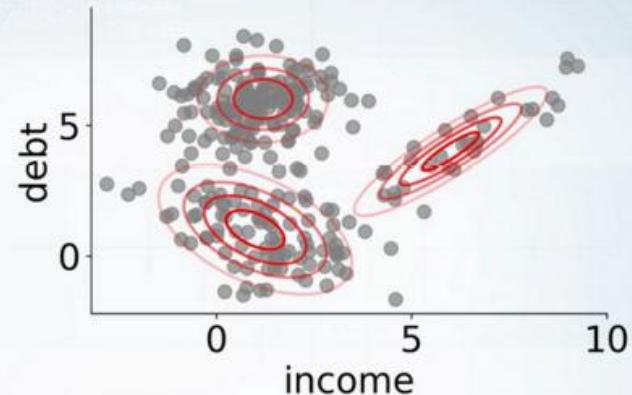
$$\log p(X \mid \theta^{k+1}) \geq \mathcal{L}(\theta^{k+1}, q^{k+1}) \geq \mathcal{L}(\theta^k, q^{k+1}) = \log p(X \mid \theta^k)$$



$$\log p(X \mid \theta^{k+1}) \geq \log p(X \mid \theta^k)$$

- On each iteration EM doesn't decrease the objective (good for debugging!)
- Guaranteed to converge to a local maximum (or saddle point)

Gaussian Mixture Model revisited



$$p(x | \theta) = \pi_1 \mathcal{N}(x | \mu_1, \Sigma_1) + \pi_2 \mathcal{N}(x | \mu_2, \Sigma_2) + \pi_3 \mathcal{N}(x | \mu_3, \Sigma_3)$$

$$\theta = \{\pi_1, \pi_2, \pi_3, \mu_1, \mu_2, \mu_3, \Sigma_1, \Sigma_2, \Sigma_3\}$$

E-step

EM: For each point compute
 $q(t_i) = p(t_i | x_i, \theta)$

GMM: For each point compute
 $p(t_i | x_i, \theta)$

M-step

EM: Update parameters to maximize
 $\max_{\theta} \mathbb{E}_q \log p(X, T | \theta)$

GMM: Update Gaussian parameters
to fit points assigned to them

$$\mu_1 = \frac{\sum_i p(t_i = 1 | x_i, \theta) x_i}{\sum_i p(t_i = 1 | x_i, \theta)}$$

Applying EM on Gaussian Mixtures

- Suppose we have some data $\mathbf{x} = x^{(1)}, \dots, x^{(m)}$
- which come from K different Gaussian distributions (K mixtures). We will use the following notations:
- μ_k : the mean of the k^{th} Gaussian component
- Σ_k : the covariance matrix of the k^{th} Gaussian component
- ϕ_k : the multinomial parameter of a specific datapoint belonging to the k^{th} component
- $z^{(i)}$: the latent variable (multinomial) for each $x^{(i)}$

We also assume that the dimension of each $x^{(i)}$ is n .

The goal is:

$$\max_{\mu, \Sigma, \phi} \ln p(\mathbf{x}; \mu, \Sigma, \phi)$$

- E-Step: $w_j^{(i)} = q_i(z^{(i)} = j) = p(z^{(i)} = j \mid x^{(i)}; \mu, \Sigma, \phi).$
- M-Step: We will write down the lower bound and get derivatives for each of the three parameters.

$$\sum_i^m \sum_j^K q_i(z^{(i)} = j) \ln \frac{p(x^{(i)}, z^{(i)} = j; \mu, \Sigma, \phi)}{q_i(z^{(i)} = j)}$$

$$= \sum_i^m \sum_j^K q_i(z^{(i)} = j) \ln \frac{p(x^{(i)} \mid z^{(i)} = j; \mu, \Sigma) \cdot p(z^{(i)} = j; \phi)}{q_i(z^{(i)} = j)}$$

- $x^{(i)} \mid z^{(i)} = j; \mu, \Sigma \sim \mathcal{N}(\mu_j, \Sigma_j)$
- $z^{(i)} = j; \phi \sim \text{Multi}(\phi)$

- We can then leverage these probability distributions and continue

$$ll := \sum_i^m \sum_j^K w_j^{(i)} \ln \frac{\frac{1}{\sqrt{(2\pi)^n |\Sigma_j|}} \exp\left(-\frac{1}{2}(x^{(i)} - \mu_j)^T \Sigma_j^{-1} (x^{(i)} - \mu_j)\right) \phi_j}{w_j^{(i)}}$$

- Now, we need to maximize this lower bound for each of the three parameters. Many of the derivatives on vector/matrix are based on the Matrix Cookbook.

Derivative of μ_j

$$\begin{aligned}
\nabla_{\mu_j} ll &= \nabla_{\mu_j} \sum_i^m w_j^{(i)} \ln \frac{\frac{1}{\sqrt{(2\pi)^n |\Sigma_j|}} \exp\left(-\frac{1}{2} (x^{(i)} - \mu_j)^T \Sigma_j^{-1} (x^{(i)} - \mu_j)\right) \phi_j}{w_j^{(i)}} \\
&= \nabla_{\mu_j} \sum_i^m w_j^{(i)} \left[\ln \frac{\frac{1}{\sqrt{(2\pi)^n |\Sigma_j|}} \phi_j}{w_j^{(i)}} + \ln \exp\left(-\frac{1}{2} (x^{(i)} - \mu_j)^T \Sigma_j^{-1} (x^{(i)} - \mu_j)\right) \right] \\
&= \nabla_{\mu_j} \sum_i^m w_j^{(i)} \left[\frac{1}{2} (x^{(i)} - \mu_j)^T \Sigma_j^{-1} (x^{(i)} - \mu_j) \right] \\
&= -\frac{1}{2} \sum_i^m w_j^{(i)} \nabla_{\mu_j} \left[(x^{(i)} - \mu_j)^T \Sigma_j^{-1} (x^{(i)} - \mu_j) \right] \\
&\quad [\text{For } f(x) = x^T A x : \nabla_x f(x) = (A + A^T)x] \\
&= \frac{1}{2} \sum_i^m w_j^{(i)} \nabla_{(x^{(i)} - \mu_j)} \left[(x^{(i)} - \mu_j)^T \Sigma_j^{-1} (x^{(i)} - \mu_j) \right] \\
&= \frac{1}{2} \sum_i^m w_j^{(i)} \left[\left(\Sigma_j^{-1} + (\Sigma_j^{-1})^T \right) (x^{(i)} - \mu_j) \right]
\end{aligned}$$

$$\nabla_{\mu_j} ll = 0$$

$$\sum_i^m w_j^{(i)} \left[\Sigma_j^{-1} \left(x^{(i)} - \mu_j \right) \right] = 0$$

$$\sum_i^m w_j^{(i)} \left(x^{(i)} - \mu_j \right) = 0$$

$$\sum_i^m w_j^{(i)} x^{(i)} = \sum_i^m w_j^{(i)} \mu_j$$

$$\mu_j = \frac{\sum_i^m w_j^{(i)} x^{(i)}}{\sum_i^m w_j^{(i)}}$$

Derivative of Σ_j

$$\begin{aligned}
 &= \sum_i^m w_j^{(i)} \nabla_{\Sigma_j} \left[\ln \frac{1}{\sqrt{|\Sigma_j|}} - \frac{1}{2} (x^{(i)} - \mu_j)^T \Sigma_j^{-1} (x^{(i)} - \mu_j) \right] \\
 &= -\frac{1}{2} \sum_i^m w_j^{(i)} \left[\frac{\partial \ln |\Sigma_j|}{\partial \Sigma_j} + \frac{\partial}{\partial \Sigma_j} (x^{(i)} - \mu_j)^T \Sigma_j^{-1} (x^{(i)} - \mu_j) \right]
 \end{aligned}$$

First, we consider the derivative of the first term in the square bracket:

$$\begin{aligned}
 \frac{\partial \ln |\Sigma_j|}{\partial \Sigma_j} &= \frac{1}{|\Sigma_j|} \frac{\partial |\Sigma_j|}{\partial \Sigma_j} \\
 &= \frac{1}{|\Sigma_j|} |\Sigma_j| \left(\Sigma_j^{-1} \right)^Y \\
 &= \Sigma_j^{-1}
 \end{aligned}$$

Then, we do the second term

$$\frac{\partial}{\partial \Sigma_j} (x^{(i)} - \mu_j)^T \Sigma_j^{-1} (x^{(i)} - \mu_j) = -\Sigma_j^{-1} (x^{(i)} - \mu_j) (x^{(i)} - \mu_j)^T \Sigma_j^{-1}$$

Combined these results back and set it to zero, we have:

$$\begin{aligned}\nabla_{\Sigma_j} ll &= -\frac{1}{2} \sum_i^m w_j^{(i)} \left[\Sigma_j^{-1} - \Sigma_j^{-1} (x^{(i)} - \mu_j) (x^{(i)} - \mu_j)^T \Sigma_j^{-1} \right] \\ &= -\frac{1}{2} \sum_i^m w_j^{(i)} \left[I - \Sigma_j^{-1} (x^{(i)} - \mu_j) (x^{(i)} - \mu_j)^T \right] \Sigma_j^{-1} \stackrel{\text{get}}{=} 0\end{aligned}$$

Rearrange the equation and we have:

$$\begin{aligned}\sum_i^m w_j^{(i)} \left[\Sigma_j - (x^{(i)} - \mu_j) (x^{(i)} - \mu_j)^T \right] &= 0 \\ \sum_i^m w_j^{(i)} \Sigma_j &= \sum_i^m w_j^{(i)} (x^{(i)} - \mu_j) (x^{(i)} - \mu_j)^T \\ \frac{\partial A^{-1}}{\partial x} &= -A^{-1} \underset{ox}{\cancel{A}} \\ \frac{\sum_i^m w_j^{(i)} (x^{(i)} - \mu_j) (x^{(i)} - \mu_j)^T}{\sum_i^m w_j^{(i)}}\end{aligned}$$

Derivative of ϕ_j

This is relatively simpler but we need to apply Lagrange multipliers because $\sum_j \phi_j = 1$.

$$\begin{aligned} ll &= \sum_i^m \sum_l^k w_l^{(i)} \ln \frac{\frac{1}{\sqrt{(2\pi)^n |\Sigma_l|}} \exp\left(-\frac{1}{2}(x^{(i)} - \mu_l)^T \Sigma_l^{-1} (x^{(i)} - \mu_l)\right) \phi_l}{w_l^{(i)}} \\ &= \sum_i^m \sum_l^k w_l^{(i)} \ln \phi_l \end{aligned}$$

We need to construct Lagrangian, with λ as the Lagrange multiplier:

$$\mathcal{L}(\phi) = ll + \lambda \left(\sum_l^k \phi_l - 1 \right)$$

We will take derivative on \mathcal{L} and set it to zero:

$$\begin{aligned}\frac{\partial \mathcal{L}(\phi)}{\partial \phi_j} &= \frac{\partial}{\partial \phi_j} \left[l + \lambda \left(\sum_l^k \phi_l - 1 \right) \right] \\ &= \sum_i w_j^{(i)} \frac{1}{\phi_j} + \lambda \stackrel{\text{set}}{=} 0\end{aligned}$$

Rearrange and we will have $\phi_j = -\frac{\sum_i w_j^{(i)}}{\lambda}$. Recall that $\sum_j \phi_j = 1$, we have:

$$\begin{aligned}\sum_j \phi_j &= \sum_j -\frac{\sum_i w_j^{(i)}}{\lambda} = 1 \\ \lambda &= -\sum_j \sum_i w_j^{(i)} \\ &= -\sum_j \sum_i p(z^{(i)} = j \mid x^{(i)}) \\ &= -\sum_i 1 = -m\end{aligned}$$

Finally, we have:

$$\phi_j = \frac{\sum_i w_j^{(i)}}{m}$$

K-Means from GMM perspective

From GMM to K-means:

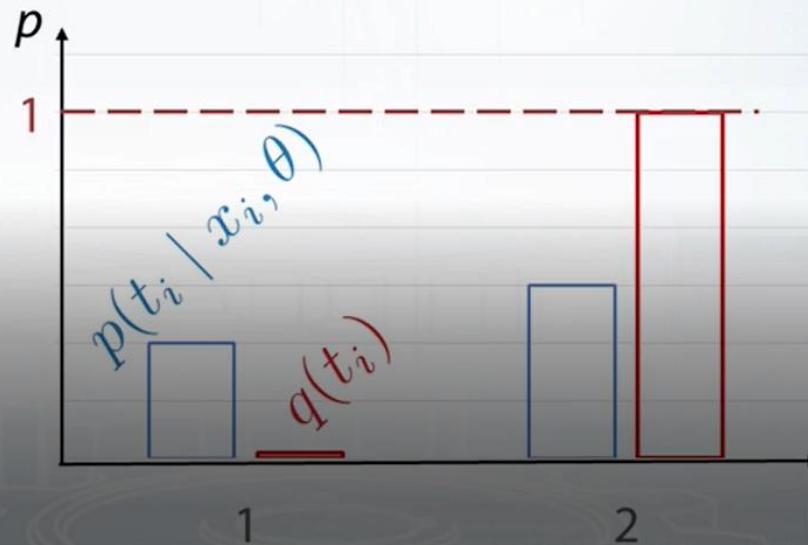
- Fix covariances to be identical $\Sigma_c = I$
- Fix weights to be uniform $\pi_c = \frac{1}{\# \text{ of Guassians}}$

$$p(x_i \mid t_i = c, \theta) = \frac{1}{Z} \exp(-0.5 \|x_i - \mu_c\|^2)$$

E-step

$$q^{k+1} = \arg \min_{q \in Q} \mathcal{KL} [q(T) \parallel p(T \mid X, \theta^k)]$$

Where Q is the set of delta-functions



E-step

$$q^{k+1}(t_i) = \begin{cases} 1 & \text{if } t_i = c_i \\ 0 & \text{otherwise} \end{cases}$$

$$c_i = \arg \max_c p(t_i = c \mid x_i, \theta) = \arg \min_c \|x_i - \mu_c\|^2$$

$$\begin{aligned} p(t_i \mid x_i, \theta) &= \frac{1}{Z} p(x_i \mid t_i, \theta) p(t_i \mid \theta) \\ &= \frac{1}{Z} \exp(-0.5 \|x_i - \mu_c\|^2) \pi_c \end{aligned}$$

E-step

$$q^{k+1}(t_i) = \begin{cases} 1 & \text{if } t_i = c_i \\ 0 & \text{otherwise} \end{cases}$$

$$c_i = \arg \min_c \|x_i - \mu_c\|^2$$

Exactly like in K-Means!

M-Step

$$\max . \sum_{i=1}^N \mathbb{E}_{q(t_i)} \log p(x_i, t_i \mid \mu)$$

$$\mu_c = \frac{\sum_{i=1}^N (q(t_i = c) \cdot x_i)}{\sum_{i=1}^N q(t_i = c)} \quad \mu_c = \frac{\sum_{i=1}^N x_i}{\#\{i : c_i^* = 0\}}$$

$$q(t_i) = \begin{cases} 1 & \text{if } t_i = c_i^* \\ 0 & \text{if } t_i \neq c_i^* \end{cases}$$

Reference

- Bayesian Methods for Machine Learning, HSE university, Coursera