



# Machine Learning

## Gaussian Mixture Model (GMM)

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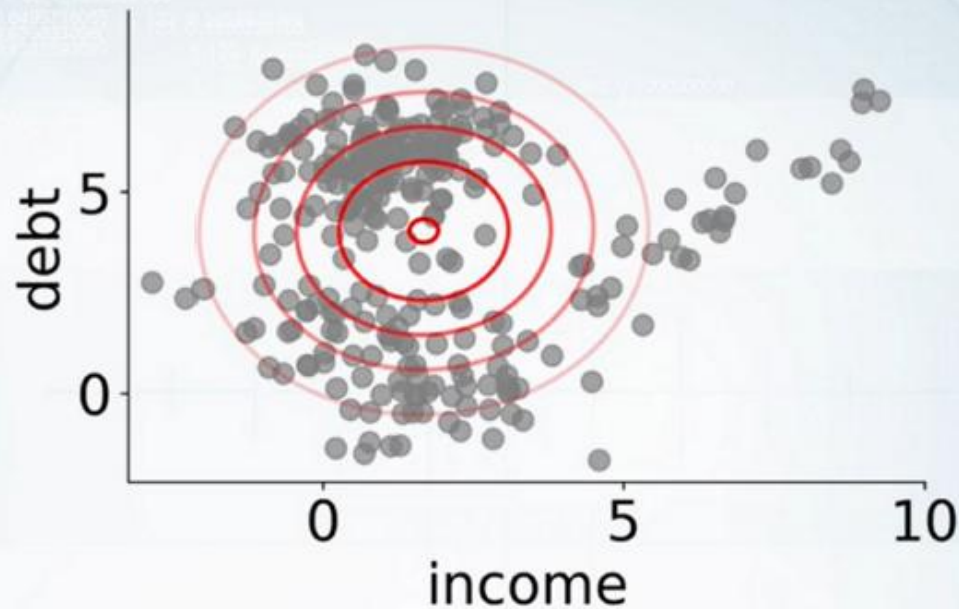
<https://www.aparat.com/mehran.safayani>



[https://github.com/safayani/machine\\_learning\\_course](https://github.com/safayani/machine_learning_course)



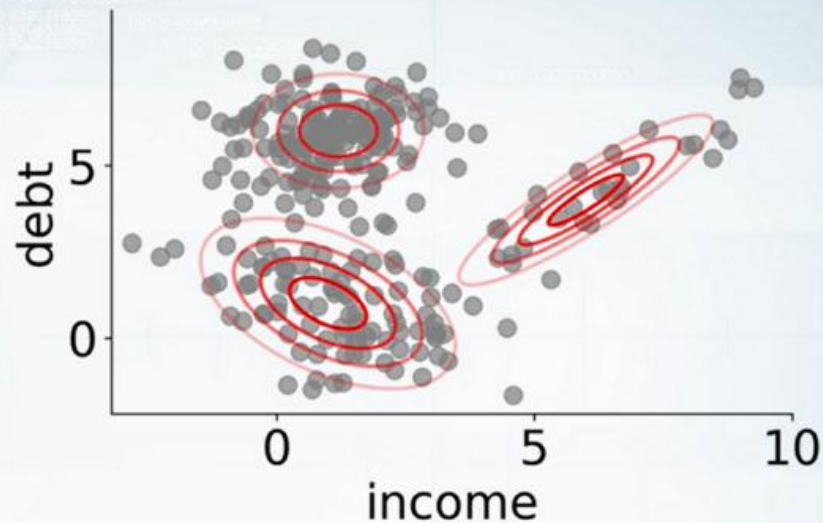
# Probabilistic model of data



$$p(x \mid \theta) = \mathcal{N}(x \mid \mu, \Sigma)$$

$$\theta = \{\mu, \Sigma\}$$

# Gaussian Mixture Model (GMM)



$$p(x \mid \theta) = \pi_1 \mathcal{N}(x \mid \mu_1, \Sigma_1) + \pi_2 \mathcal{N}(x \mid \mu_2, \Sigma_2) + \pi_3 \mathcal{N}(x \mid \mu_3, \Sigma_3)$$

$$\theta = \{\pi_1, \pi_2, \pi_3, \mu_1, \mu_2, \mu_3, \Sigma_1, \Sigma_2, \Sigma_3\}$$

Gaussian

GMM

Flexibility



# of parameters



Parameters

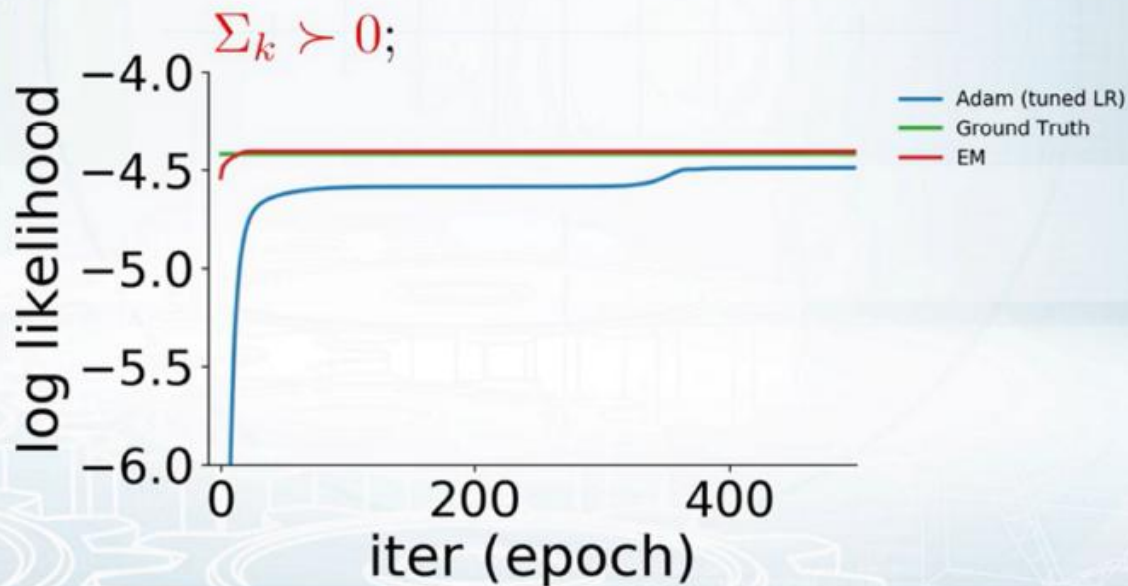
$\mu, \Sigma$

$\{\pi_1, \pi_2, \pi_3\}$   
 $\{\mu_1, \mu_2, \mu_3\}$

# Training GMM

$$\max_{\theta} \prod_{i=1}^N p(x_i | \theta) = \prod_{i=1}^N (\pi_1 \mathcal{N}(x_i | \mu_1, \Sigma_1) + \dots)$$

subject to  $\pi_1 + \pi_2 + \pi_3 = 1; \pi_k \geq 0; k = 1, 2, 3.$



# Introducing latent variable

$$p(x | \theta) = \pi_1 \mathcal{N}(x | \mu_1, \Sigma_1) + \pi_2 \mathcal{N}(x | \mu_2, \Sigma_2) + \pi_3 \mathcal{N}(x | \mu_3, \Sigma_3)$$



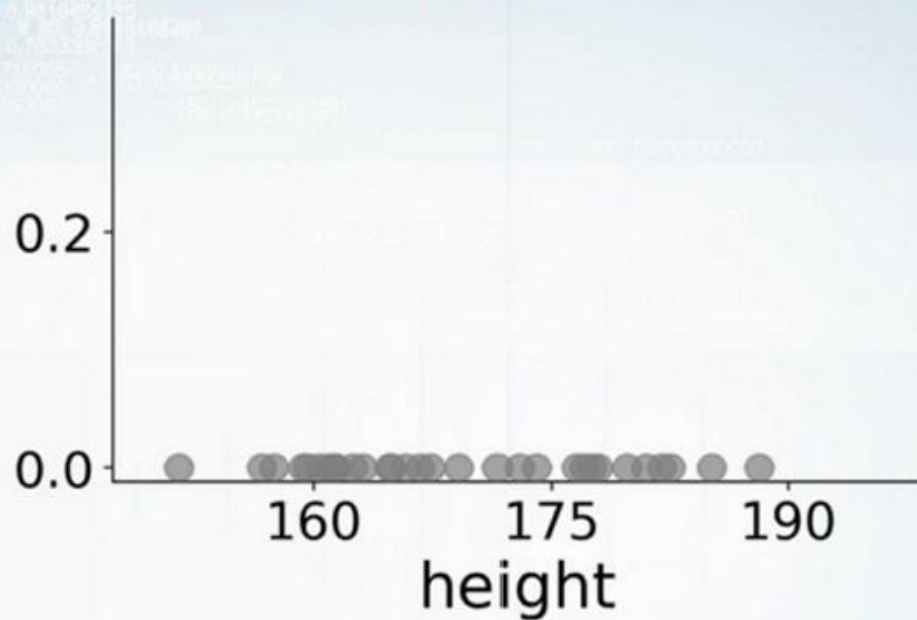
$$p(t = c | \theta) = \pi_c$$

$$p(x | t = c, \theta) = \mathcal{N}(x | \mu_c, \Sigma_c)$$

$$p(x | \theta) = \sum_{c=1}^3 p(x | t = c, \theta) p(t = c | \theta)$$

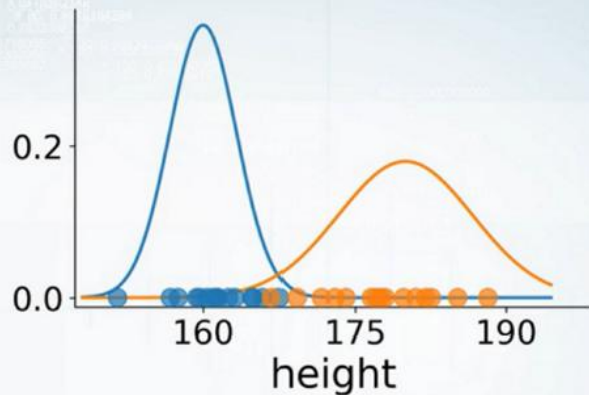
$p(x, t)$

# Expectation Maximization



How to estimate parameter  $\theta$ ?

# Expectation Maximization



How to estimate parameter  $\theta$ ?

If sources  $t$  are known, easy:

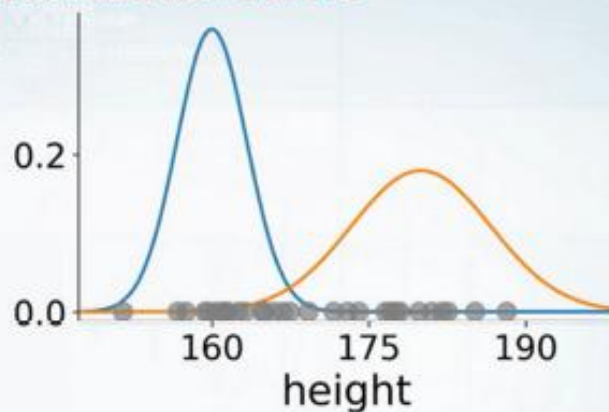
$$p(x \mid t = 1, \theta) = \mathcal{N}(x \mid \mu_1, \sigma_1^2)$$

$$\mu_1 = \frac{\sum_{\text{blue } i} x_i}{\# \text{ of blue points}} \quad \sigma_1^2 = \frac{\sum_{\text{blue } i} (x_i - \mu_1)^2}{\# \text{ of blue points}}$$

$$\mu_1 = \frac{\sum_i p(t_i = 1 \mid x_i, \theta) x_i}{\sum_i p(t_i = 1 \mid x_i, \theta)} \quad \sigma_1^2 = \frac{\sum_i p(t_i = 1 \mid x_i, \theta) (x_i - \mu_1)^2}{\sum_i p(t_i = 1 \mid x_i, \theta)}$$



## Expectation Maximization



What if we don't know the sources?

Given:  $p(x \mid t = 1, \theta) = \mathcal{N}(-2, 1)$

Find:  $p(t = 1 \mid x, \theta)$

$$p(t = 1 \mid x, \theta) = \frac{p(x \mid t = 1, \theta) p(t = 1 \mid \theta)}{Z}$$

## Chicken and egg problem

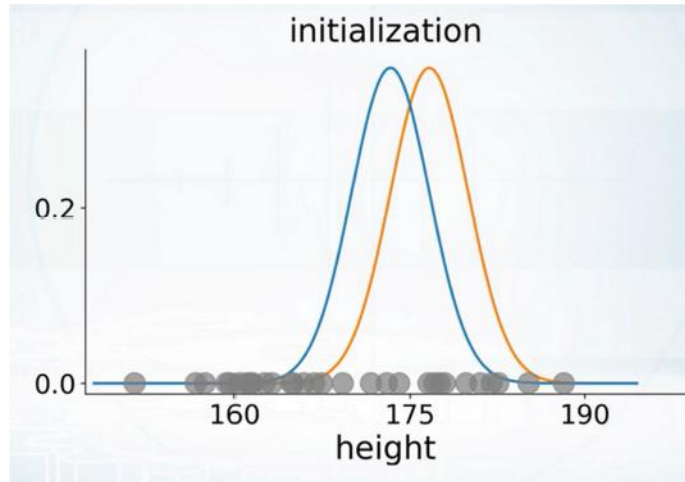
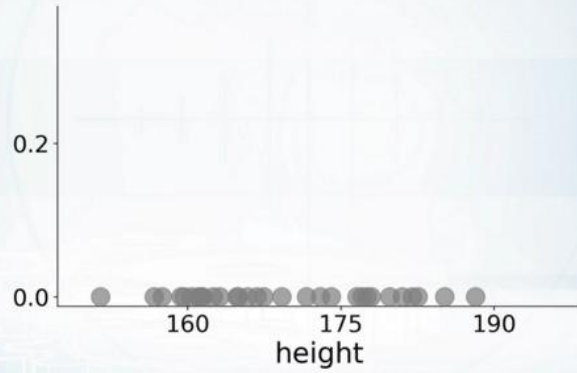
- Need Gaussian parameters to estimate sources
- Need sources to estimate Gaussian parameters

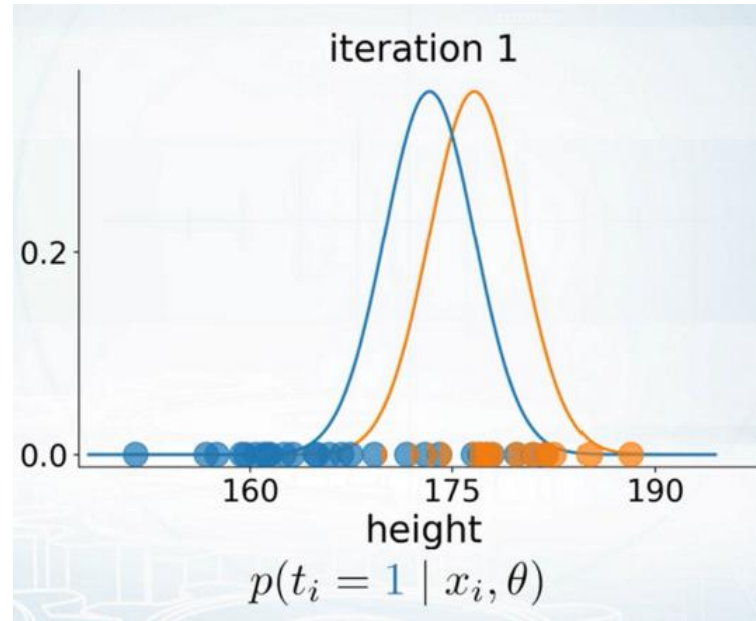


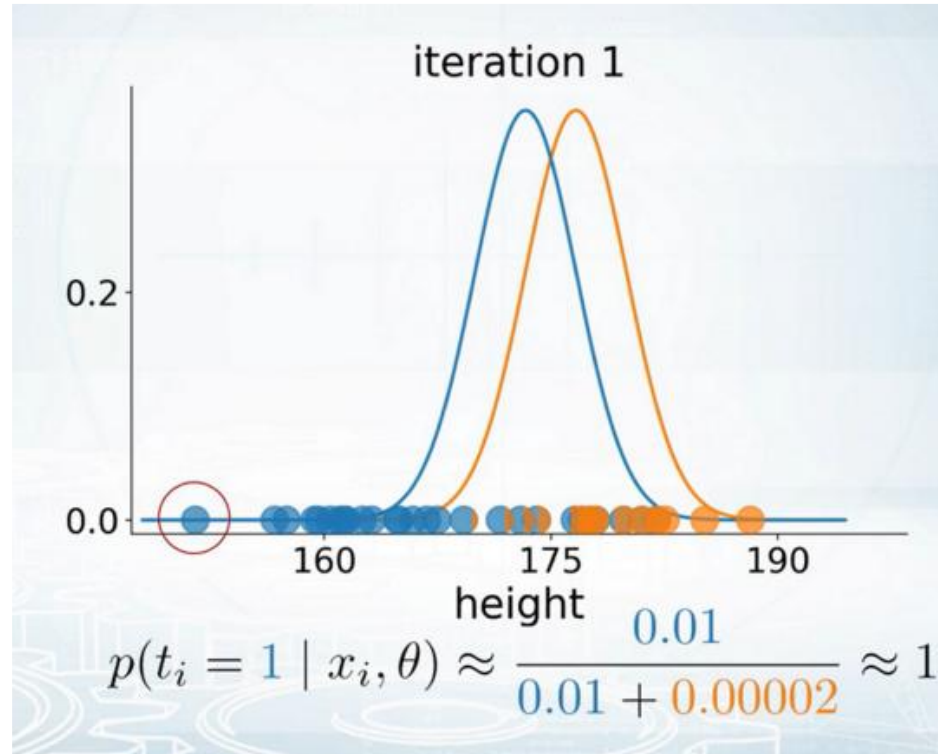
## EM algorithm

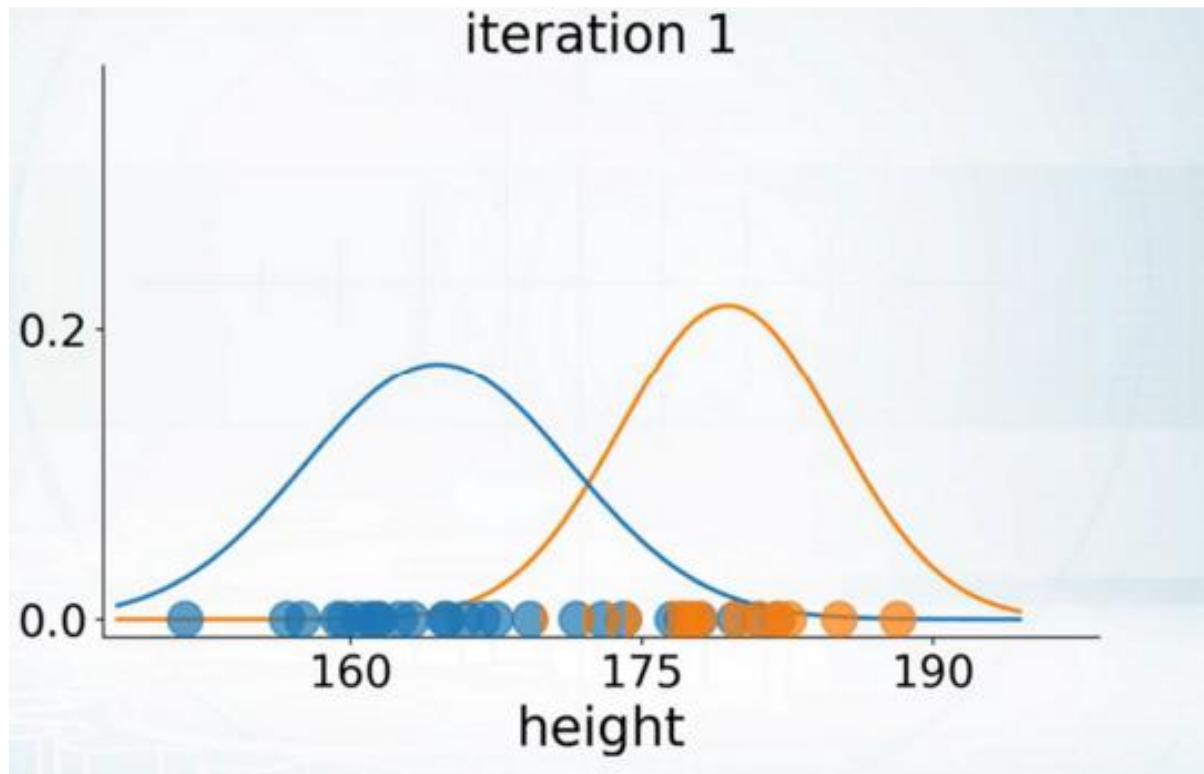
1. Start with 2 randomly placed Gaussians parameters  $\theta$
2. Until convergence repeat:
  - a) For each point compute  $p(t = c \mid x_i, \theta)$ : does  $x_i$  look like it came from cluster  $c$ ?
  - b) Update Gaussian parameters  $\theta$  to fit points assigned to them

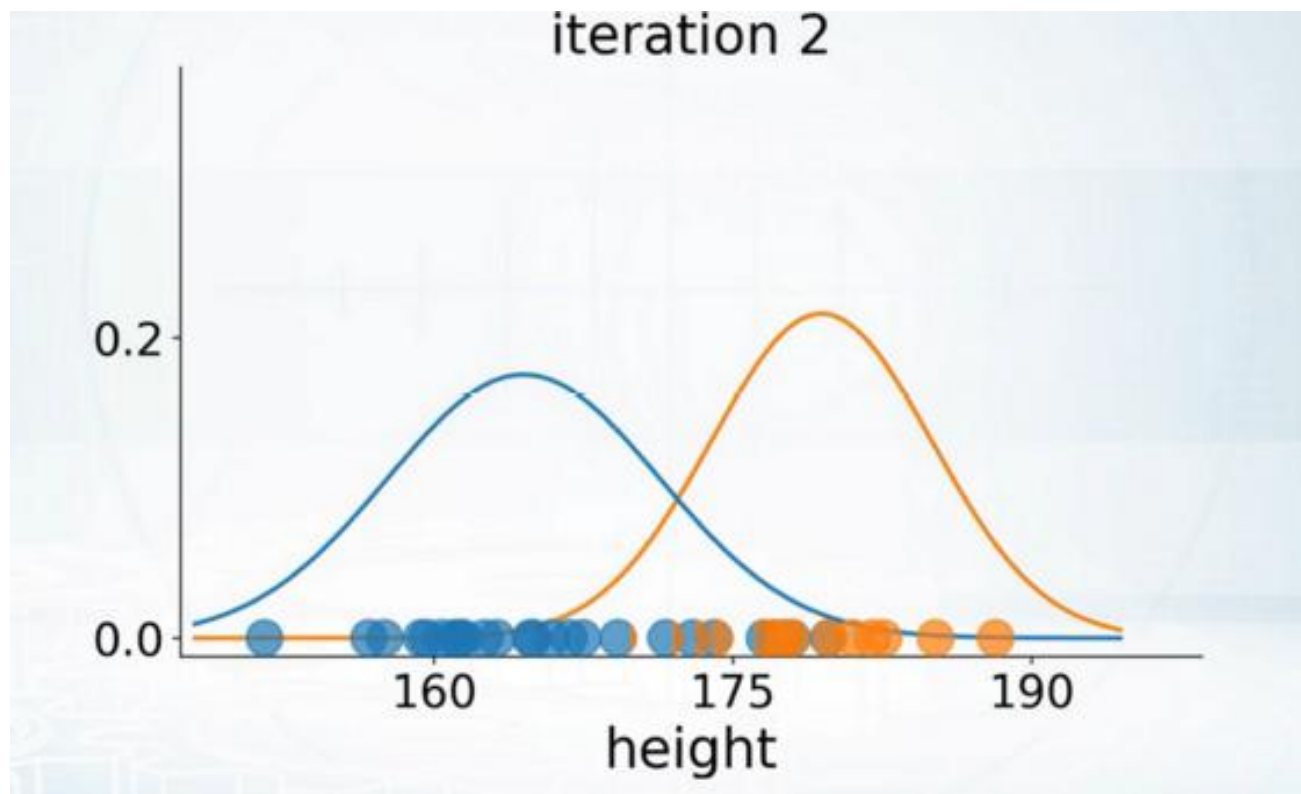
## GMM EM example



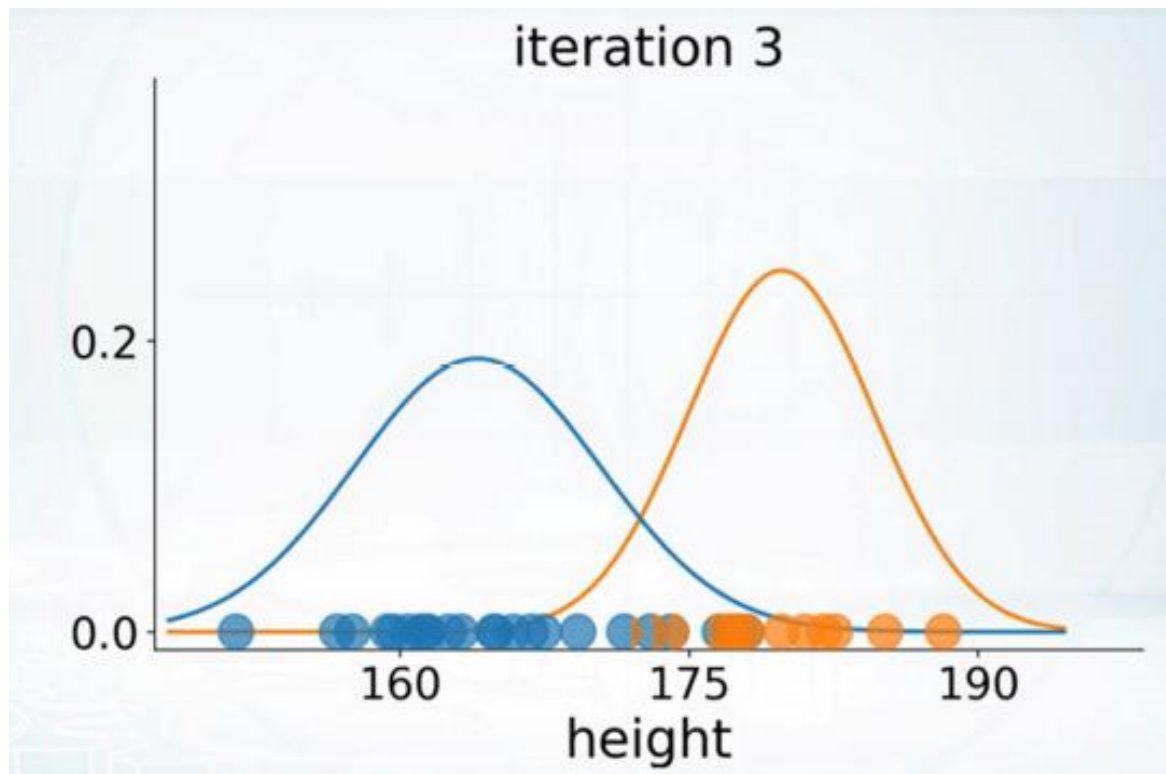




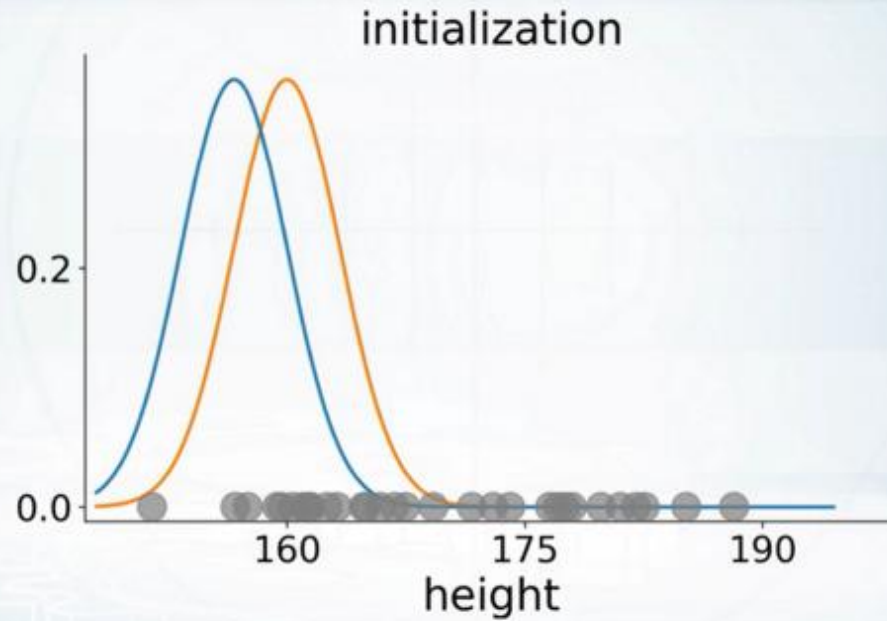


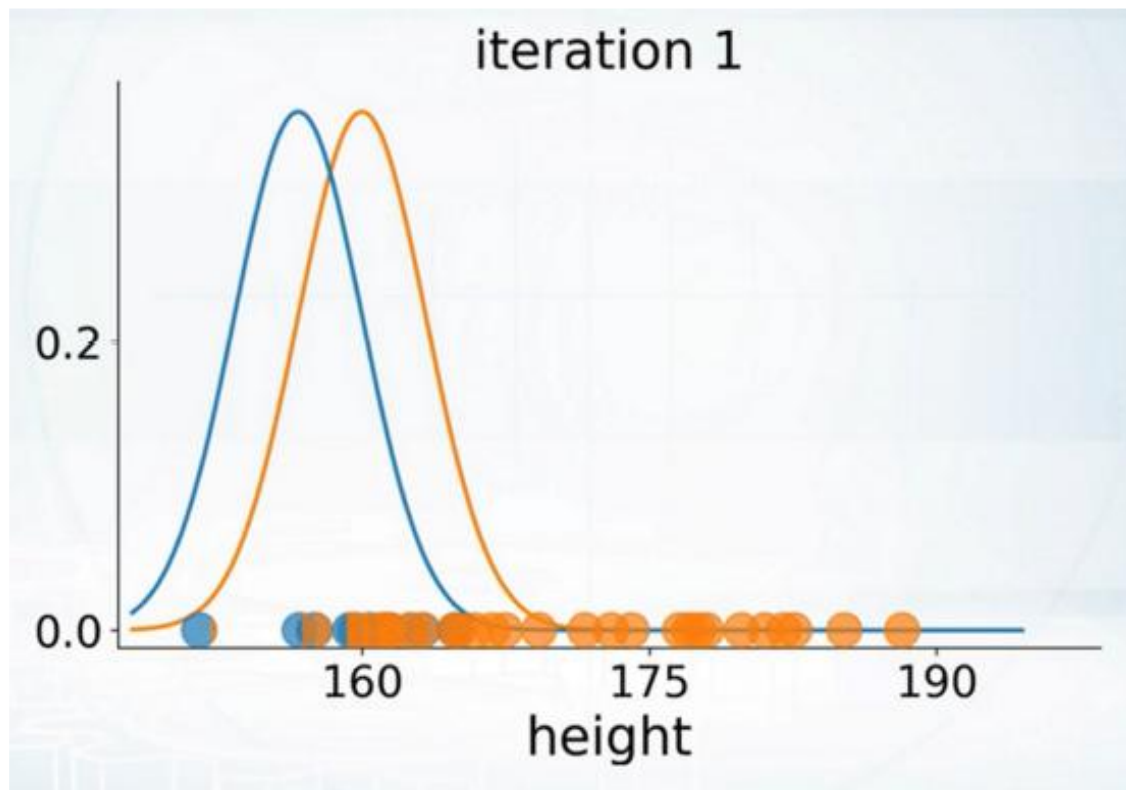


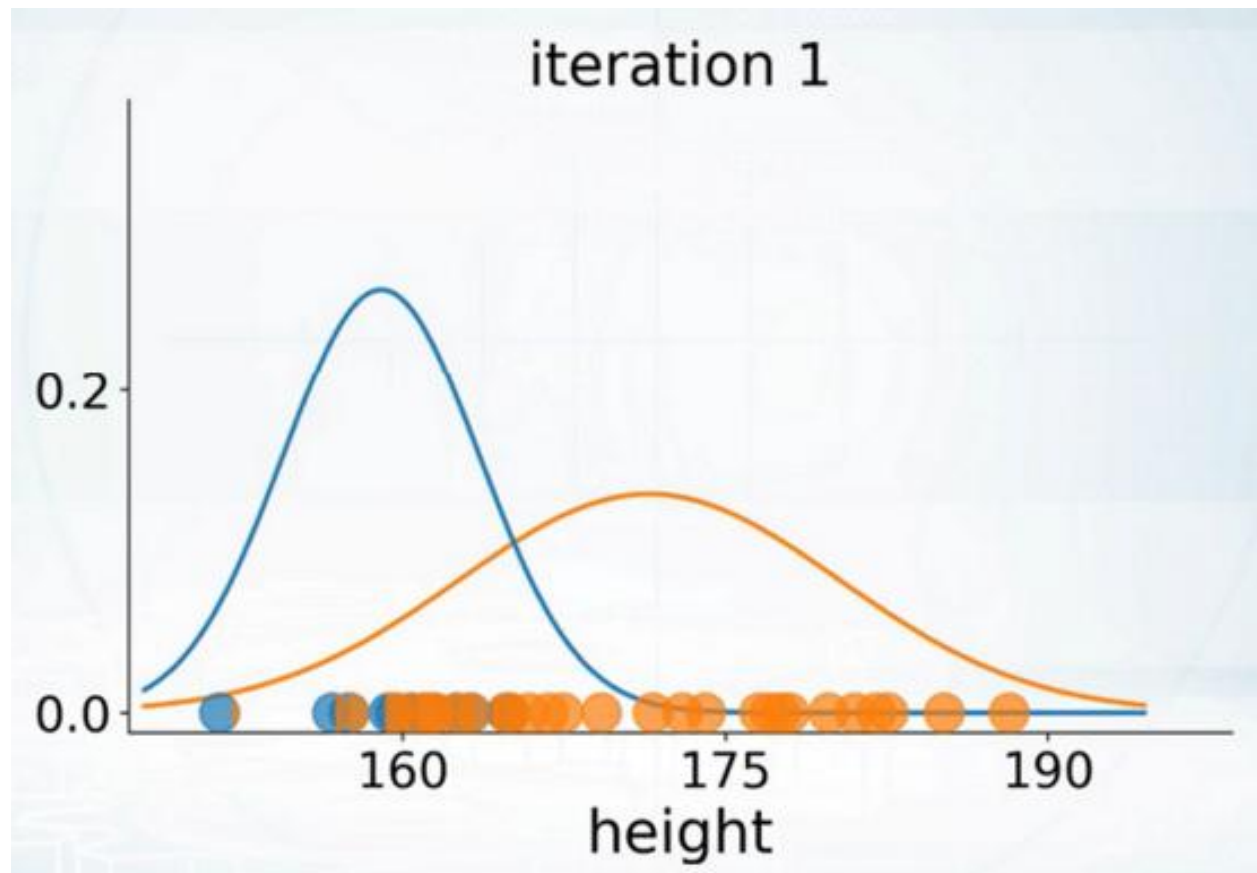


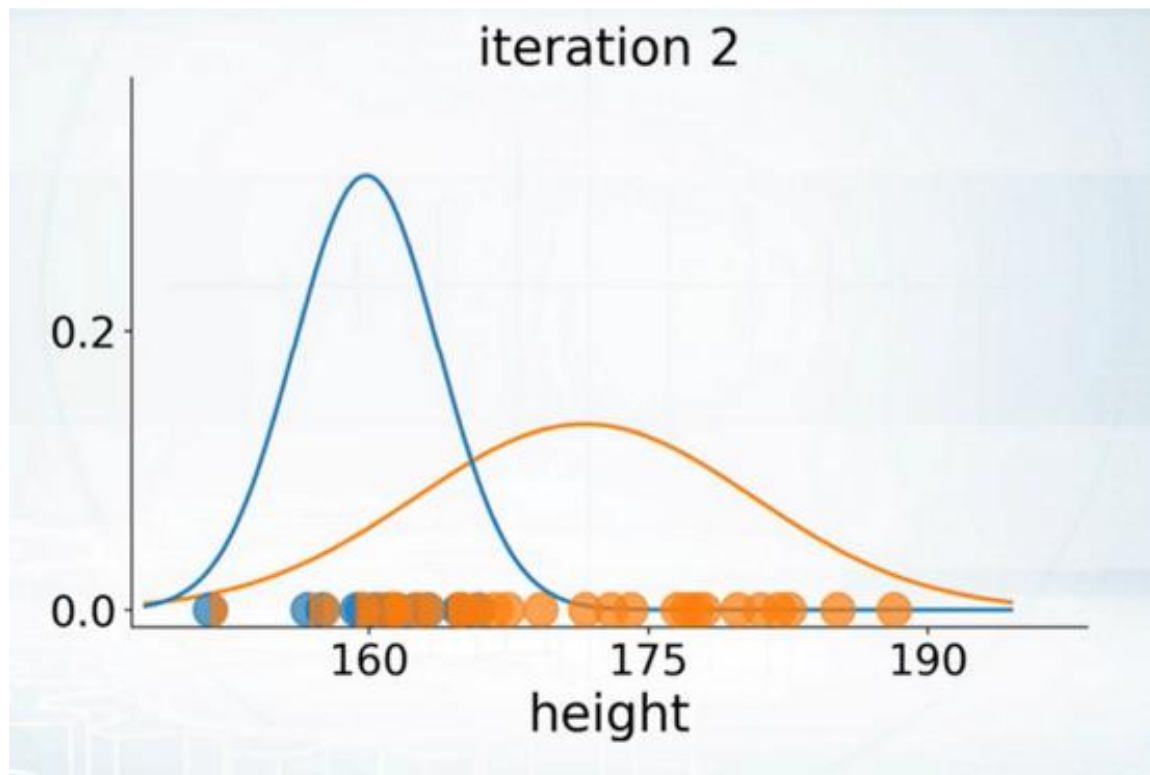


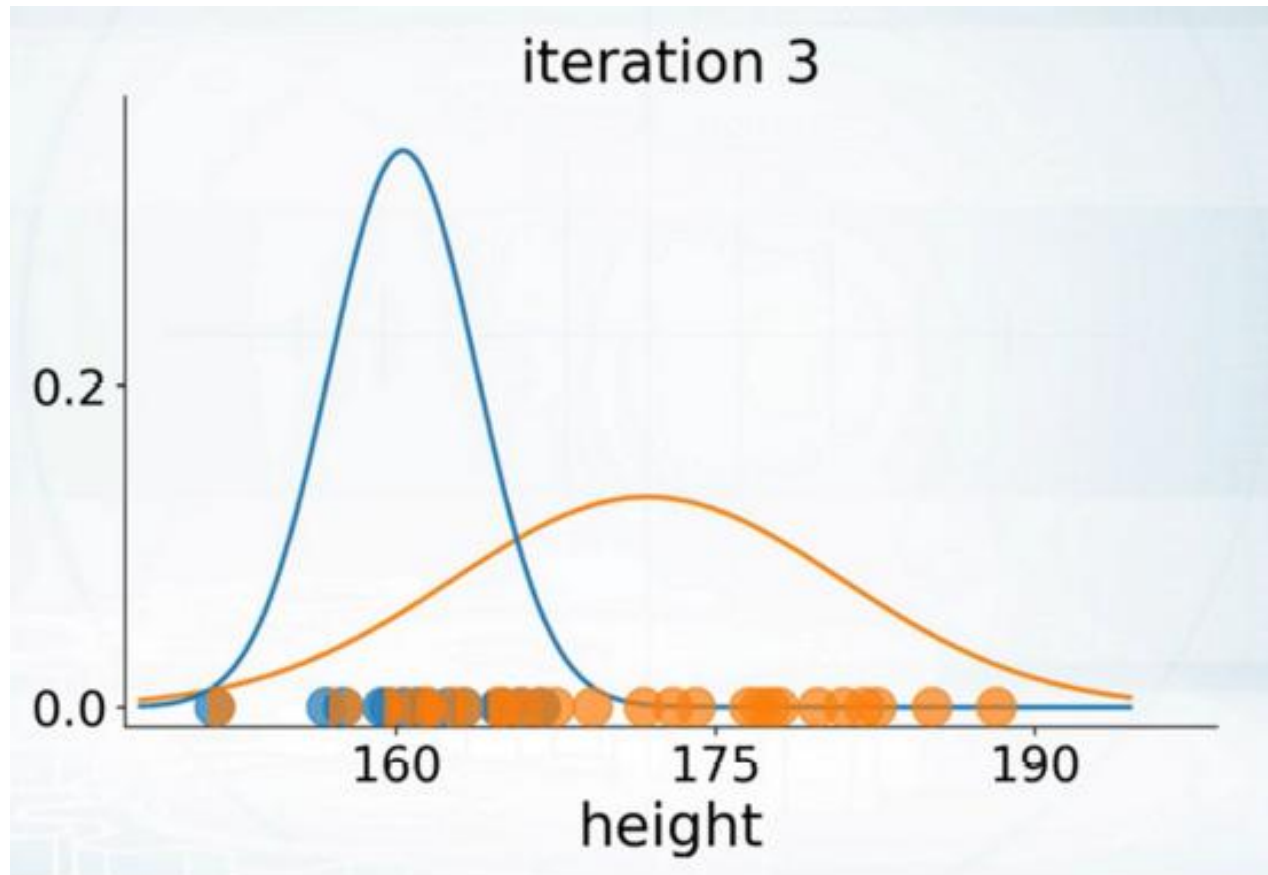
## GMM EM local maximum example





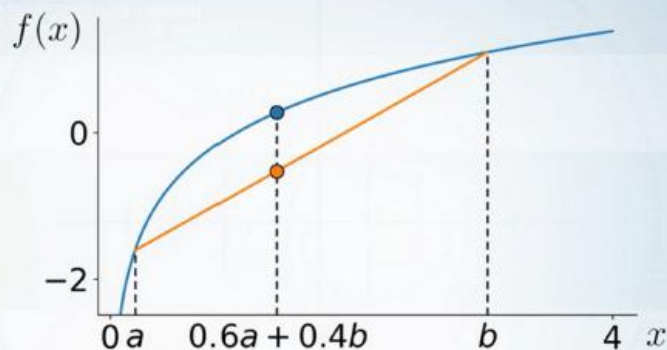






# General form of Expectation Maximization

## Concave functions



Def.:  $f(x)$  is concave if

$$\text{for any } a, b, \alpha : f(\alpha a + (1 - \alpha)b) \geq \alpha f(a) + (1 - \alpha)f(b) \\ 0 \leq \alpha \leq 1$$

## Jensen's inequality

If  $f(\alpha a + (1 - \alpha)b) \geq \alpha f(a) + (1 - \alpha)f(b)$

Then  $\alpha_1 + \alpha_2 + \alpha_3 = 1; \alpha_k \geq 0.$

$$f(\alpha_1 a_1 + \alpha_2 a_2 + \alpha_3 a_3) \geq \alpha_1 f(a_1) + \alpha_2 f(a_2) + \alpha_3 f(a_3)$$

$$f(\underbrace{\alpha_1 a_1 + \alpha_2 a_2 + \alpha_3 a_3}_{\mathbb{E}_{p(t)} t}) \geq \underbrace{\alpha_1 f(a_1) + \alpha_2 f(a_2) + \alpha_3 f(a_3)}_{\mathbb{E}_{p(t)} f(t)}$$

$$p(t = a_1) = \alpha_1,$$

$$p(t = a_2) = \alpha_2,$$

$$p(t = a_3) = \alpha_3$$

## Jensen's inequality

If  $f(\alpha a + (1 - \alpha)b) \geq \alpha f(a) + (1 - \alpha)f(b)$

Then

Jensen's inequality:

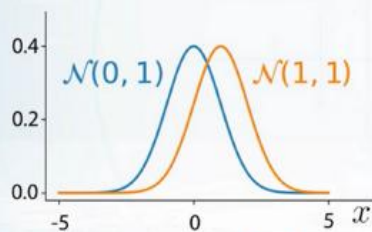
$$f(\mathbb{E}_{p(t)} t) \geq \mathbb{E}_{p(t)} f(t)$$



# Kullback-Leibler divergence

Parameters difference: 1

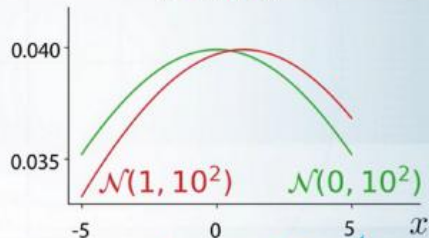
$$\mathcal{KL}(q_1 \parallel p_1) = 0.5$$



$$|\mu_1 - \mu_2| < 1$$

Parameters difference: 1

$$\mathcal{KL}(q_2 \parallel p_2) = 0.005$$



$$|\mu_1 - \mu_2| = 1$$

$$\mathcal{KL}(q \parallel p) = \int q(x) \log \frac{q(x)}{p(x)} dx$$

1.  $\mathcal{KL}(q \parallel p) \neq \mathcal{KL}(p \parallel q)$
2.  $\mathcal{KL}(q \parallel q) = 0$
3.  $\mathcal{KL}(q \parallel p) \geq 0$

**Proof:**  $-\mathcal{KL}(q \parallel p) = \mathbb{E}_q \left( -\log \frac{q}{p} \right) = \mathbb{E}_q \left( \log \frac{p}{q} \right)$

$$\leq \log(\mathbb{E}_q \frac{p}{q}) = \log \int q(x) \frac{p(x)}{q(x)} dx = 0$$

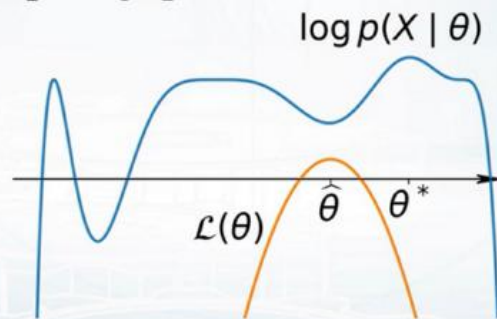
## General form of Expectation Maximization



$$p(x_i | \theta) = \sum_{c=1}^3 p(x_i | t_i = c, \theta) p(t_i = c | \theta)$$

$$\begin{aligned} \max_{\theta} \log p(X | \theta) &= \log \prod_{i=1}^N p(x_i | \theta) \\ &= \sum_{i=1}^N \log p(x_i | \theta) \end{aligned}$$

$$\begin{aligned}\log p(X | \theta) &= \sum_{i=1}^N \log p(x_i | \theta) \\ &= \sum_{i=1}^N \log \sum_{c=1}^3 p(x_i, t_i = c | \theta) \geq \mathcal{L}(\theta)\end{aligned}$$



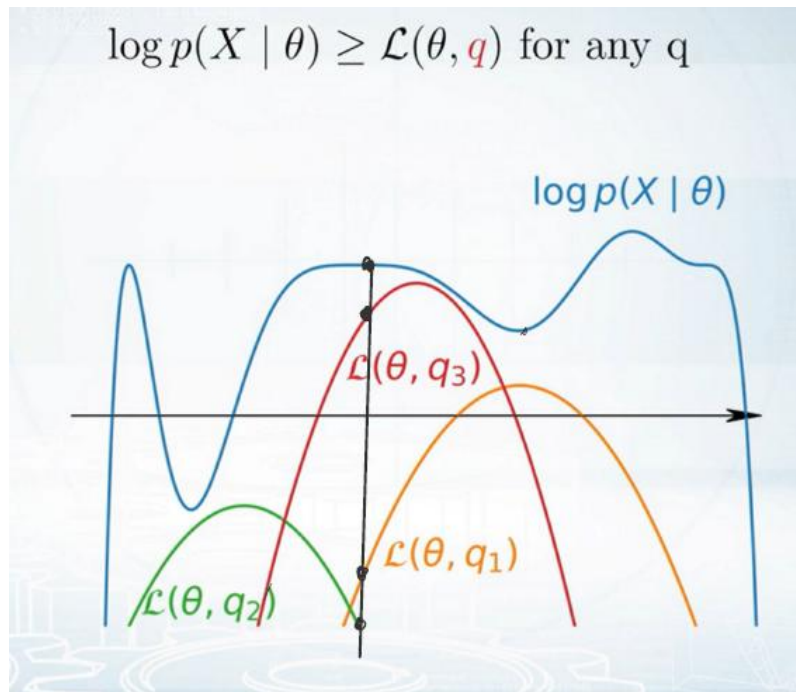
$$\begin{aligned}
 \log p(X \mid \theta) &= \sum_{i=1}^N \log p(x_i \mid \theta) \\
 &= \sum_{i=1}^N \log \sum_{c=1}^3 \frac{q(t_i = c)}{q(t_i = c)} p(x_i, t_i = c \mid \theta) \\
 &\geq \sum_{i=1}^N \sum_{c=1}^3 q(t_i = c) \log \frac{p(x_i, t_i = c \mid \theta)}{q(t_i = c)}
 \end{aligned}$$

Jensen's inequality

$$\log \left( \sum_c \alpha_c v_c \right) \geq \sum_c \alpha_c \log(v_c)$$

$$\begin{aligned}
\log p(X \mid \theta) &= \sum_{i=1}^N \log p(x_i \mid \theta) \\
&= \sum_{i=1}^N \log \sum_{c=1}^3 \frac{q(t_i = c)}{q(t_i = c)} p(x_i, t_i = c \mid \theta) \\
&\geq \sum_{i=1}^N \sum_{c=1}^3 q(t_i = c) \log \frac{p(x_i, t_i = c \mid \theta)}{q(t_i = c)} \\
&= \mathcal{L}(\theta, q)
\end{aligned}$$

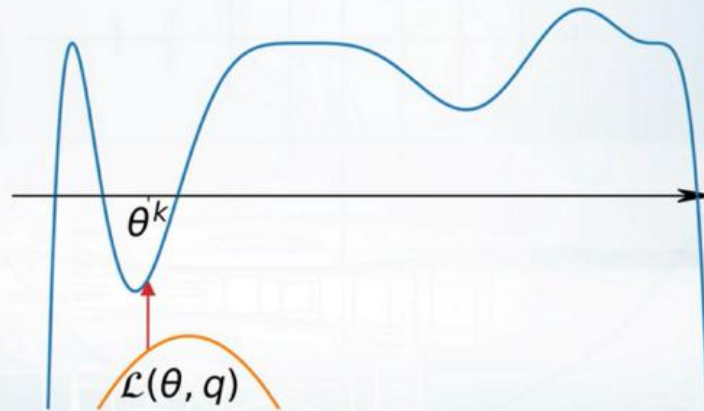
$$\log p(X | \theta) \geq \mathcal{L}(\theta, q) \text{ for any } q$$



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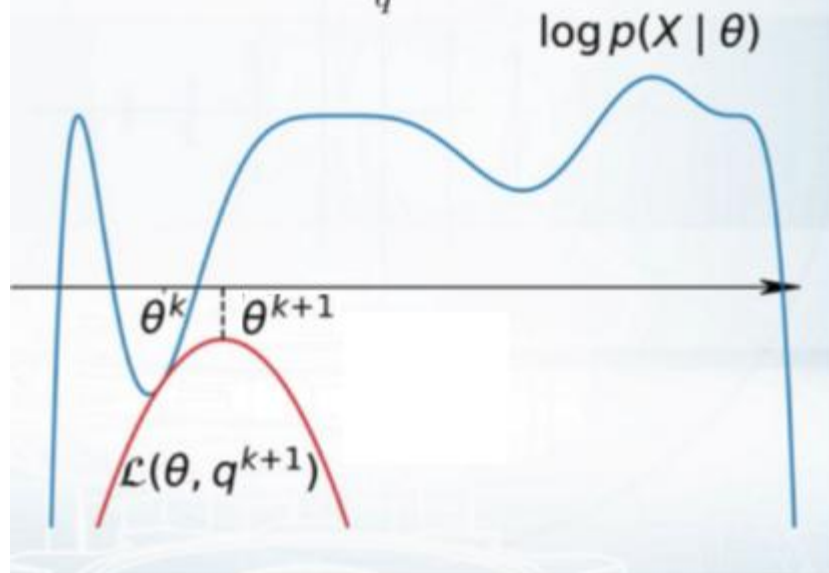
$$q^{k+1} = \arg \max_q \mathcal{L}(\theta^k, q)$$

$\log p(X | \theta)$



$$\log p(X | \theta) \geq \mathcal{L}(\theta, q) \text{ for any } q$$

$$q^{k+1} = \arg \max_q \mathcal{L}(\theta^k, q)$$



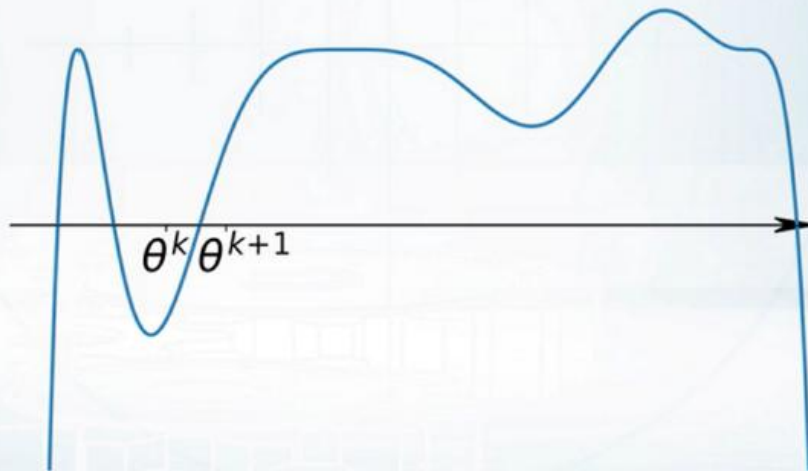
$$\theta^{k+1} = \arg \max_{\theta} \mathcal{L}(\theta, q^{k+1})$$



$$\log p(X | \theta) \geq \mathcal{L}(\theta, q) \text{ for any } q$$

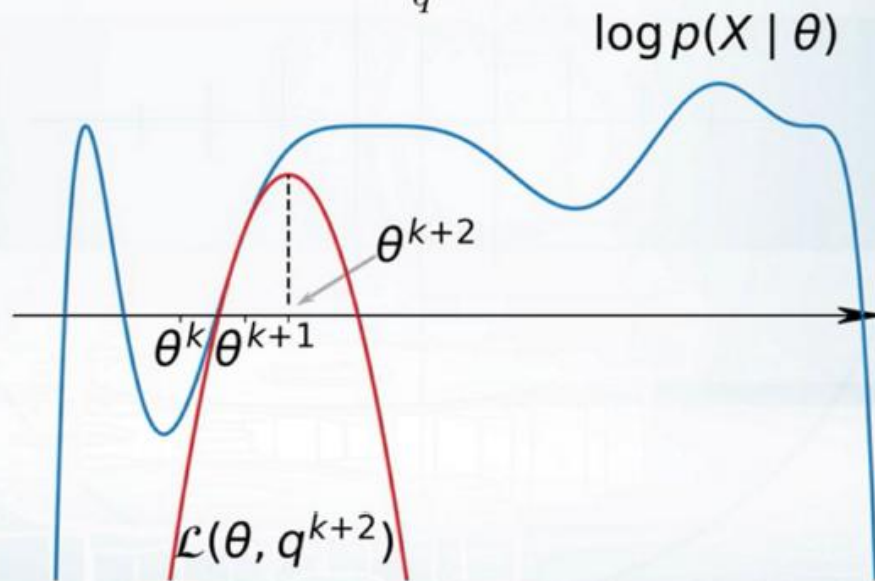
$$q^{k+1} = \arg \max_q \mathcal{L}(\theta^k, q)$$

$\log p(X | \theta)$



$$\log p(X | \theta) \geq \mathcal{L}(\theta, q) \text{ for any } q$$

$$q^{k+1} = \arg \max_q \mathcal{L}(\theta^k, q)$$



## Summary of Expectation Maximization

$$\log p(X | \theta) \geq \mathcal{L}(\theta, q) \text{ for any } q$$

Variational  
lower bound

E-step

$$q^{k+1} = \arg \max_q \mathcal{L}(\theta^k, q)$$

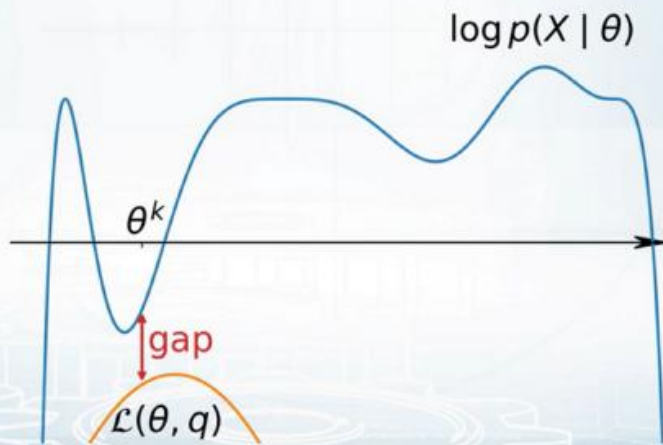
M-step

$$\theta^{k+1} = \arg \max_{\theta} \mathcal{L}(\theta, q^{k+1})$$

## E-step details

$$\log p(X | \theta) \geq \mathcal{L}(\theta, q)$$

E-step:  $\max_q \mathcal{L}(\theta^k, q)$



$$GAP = \log p(X|\theta) - L(\theta, q) = \sum_{i=1}^N \log p(x_i|\theta) - \sum_{i=1}^N \sum_{c=1}^3 q(t_i = c) \log \frac{p(\mathbf{x}_i, t_i = c|\theta)}{q(t_i = c)}$$

$$\sum_{i=1}^N \log p(x_i|\theta) \cdot \sum_{c=1}^3 q(t_i = c) - \sum_{i=1}^N \sum_{c=1}^3 q(t_i = c) \log \frac{p(x_i, t_i = c|\theta)}{q(t_i = c)}$$

$$= - \sum_{i=1}^N \sum_{c=1}^3 q(t_i = c) (\log p(x_i|\theta) - \log \frac{p(x_i, t_i = c|\theta)}{q(t_i = c)})$$

$$= \sum_{i=1}^N \sum_{c=1}^3 q(t_i = c) \log \frac{p(x_i|\theta)q(t_i = c)}{p(x_i, t_i = c|\theta)}$$

$$= \sum_{i=1}^N \sum_{c=1}^3 q(t_i = c) \log \frac{q(t_i = c)}{p(t_i = c|x_i, \theta)} = \sum_{i=1}^N \sum_{c=1}^3 KL(q(t_i = c) || p(t_i = c|x_i, \theta))$$

Estep:

$$GAP = \log p(X|\theta) - L(\theta, q) = \log p(X|\theta) - ELBO$$

$$= \sum_{i=1}^N \sum_{c=1}^3 KL(q(t_i = c) || p(t_i = c | x_i, \theta)) \gg 0$$

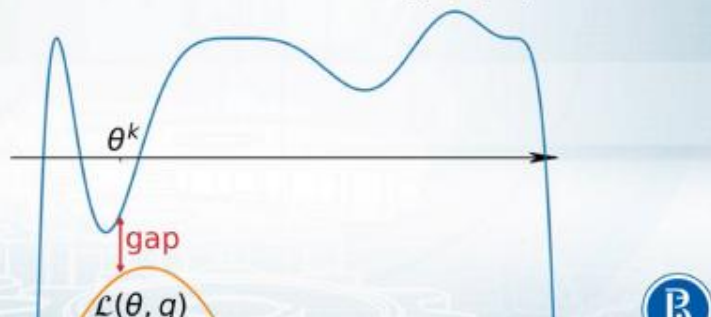
We maximize **ELBO** or minimize **KL**

$$\log p(X|\theta) = \sum_{i=1}^N \sum_{c=1}^3 KL(q(t_i = c) || p(t_i = c | x_i, \theta)) + ELBO = GAP + ELBO$$

**Solution:**  $q(t_i = c) = p(t_i = c | x_i, \theta)$

$$\log p(X | \theta) - \mathcal{L}(\theta, q) = \sum_i \mathcal{KL}(q(t_i) \parallel p(t_i | x_i, \theta))$$

$$\text{E-step: } \arg \max_{q(t_i)} \mathcal{L}(\theta^k, q) = \log p(X | \theta)$$



## M-step details

$$\mathcal{L}(\theta, q) = \sum_i \sum_c q(t_i = c) \log \frac{p(x_i, t_i = c \mid \theta)}{q(t_i = c)}$$

$$= \sum_i \sum_c q(t_i = c) \log p(x_i, t_i = c \mid \theta)$$

~~$$- \sum_i \sum_c q(t_i = c) \log q(t_i = c)$$~~

$$= \mathbb{E}_q \log p(X, T \mid \theta) + \text{const}$$

Const w.r.t.  $\theta$



$$= \mathbb{E}_q \log p(X, T \mid \theta) + \text{const}$$



(Usually) concave function w.r.t.  $\theta$ , easy to optimize

## Expectation Maximization algorithm

For  $k = 1, \dots$

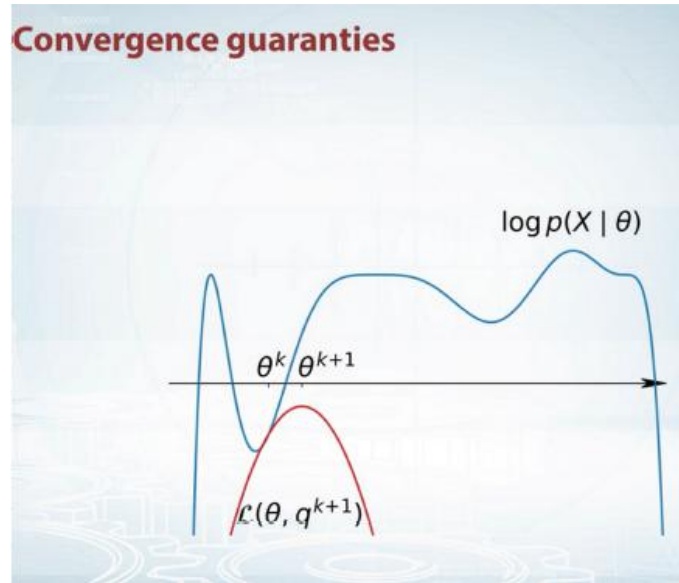
**E-step**

$$\begin{aligned} q^{k+1} &= \arg \min_q \mathcal{KL} [q(T) \parallel p(T \mid X, \theta^k)] \\ &\Leftrightarrow \\ q^{k+1}(t_i) &= p(t_i \mid x_i, \theta^k) \end{aligned}$$

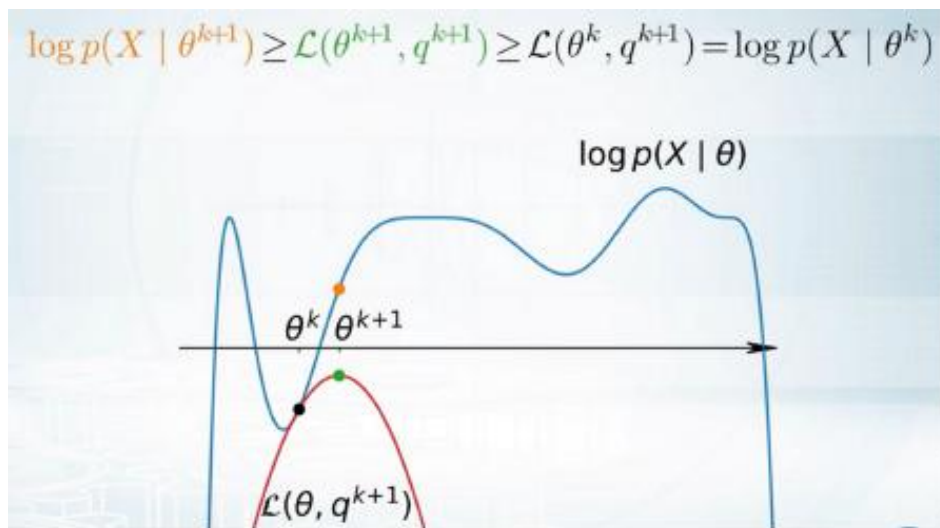
**M-step**

$$\theta^{k+1} = \arg \max_{\theta} \mathbb{E}_{q^{k+1}} \log p(X, T \mid \theta)$$

## Convergence guaranties



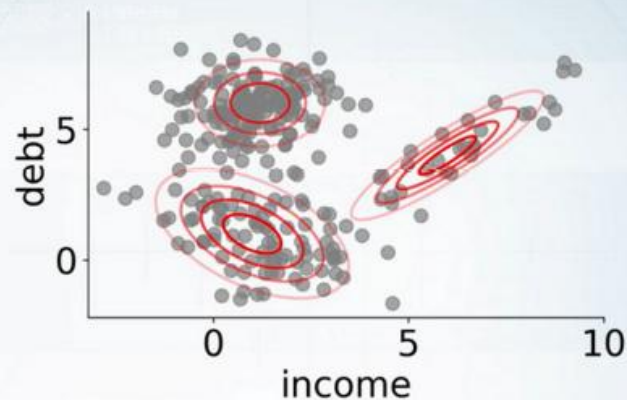
$$\log p(X \mid \theta^{k+1}) \geq \mathcal{L}(\theta^{k+1}, q^{k+1}) \geq \mathcal{L}(\theta^k, q^{k+1}) = \log p(X \mid \theta^k)$$



$$\log p(X \mid \theta^{k+1}) \geq \log p(X \mid \theta^k)$$

- On each iteration EM doesn't decrease the objective (good for debugging!)
- Guaranteed to converge to a local maximum (or saddle point)

## Gaussian Mixture Model revisited



$$p(x | \theta) = \pi_1 \mathcal{N}(x | \mu_1, \Sigma_1) + \pi_2 \mathcal{N}(x | \mu_2, \Sigma_2) \\ + \pi_3 \mathcal{N}(x | \mu_3, \Sigma_3)$$

$$\theta = \{\pi_1, \pi_2, \pi_3, \mu_1, \mu_2, \mu_3, \Sigma_1, \Sigma_2, \Sigma_3\}$$

### E-step

**EM:** For each point compute

$$q(t_i) = p(t_i \mid x_i, \theta)$$

**GMM:** For each point compute

$$p(t_i \mid x_i, \theta)$$

### M-step

**EM:** Update parameters to maximize

$$\max_{\theta} \mathbb{E}_q \log p(X, T \mid \theta)$$

**GMM:** Update Gaussian parameters  
to fit points assigned to them

$$\mu_1 = \frac{\sum_i p(t_i = 1 \mid x_i, \theta) x_i}{\sum_i p(t_i = 1 \mid x_i, \theta)}$$

# Applying EM on Gaussian Mixtures

- Suppose we have some data  $\mathbf{x} = x^{(1)}, \dots, x^{(m)}$
- which come from  $K$  different Gaussian distributions ( $K$  mixtures). We will use the following notations:
  - $\mu_k$ : the mean of the  $k^{th}$  Gaussian component
  - $\Sigma_k$ : the covariance matrix of the  $k^{th}$  Gaussian component
  - $\phi_k$ : the multinomial parameter of a specific datapoint belonging to the  $k^{th}$  component
  - $z^{(i)}$ : the latent variable (multinomial) for each  $x^{(i)}$

We also assume that the dimension of each  $x^{(i)}$  is  $n$ .

The goal is:

$$\max_{\mu, \Sigma, \phi} \ln p(\mathbf{x}; \mu, \Sigma, \phi)$$

- E-Step:  $w_j^{(i)} = q_i(z^{(i)} = j) = p(z^{(i)} = j \mid x^{(i)}; \mu, \Sigma, \phi).$
- M-Step: We will write down the lower bound and get derivatives for each of the three parameters.

$$\sum_i^m \sum_j^K q_i(z^{(i)} = j) \ln \frac{p(x^{(i)}, z^{(i)} = j; \mu, \Sigma, \phi)}{q_i(z^{(i)} = j)}$$

$$= \sum_i^m \sum_j^K q_i(z^{(i)} = j) \ln \frac{p(x^{(i)} \mid z^{(i)} = j; \mu, \Sigma) \cdot p(z^{(i)} = j; \phi)}{q_i(z^{(i)} = j)}$$

- $x^{(i)} \mid z^{(i)} = j; \mu, \Sigma \sim \mathcal{N}(\mu_j, \Sigma_j)$
- $z^{(i)} = j; \phi \sim \text{Multi}(\phi)$

- We can then leverage these probability distributions and continue

$$ll := \sum_i^m \sum_j^K w_j^{(i)} \ln \frac{\frac{1}{\sqrt{(2\pi)^n |\Sigma_j|}} \exp \left( -\frac{1}{2} (x^{(i)} - \mu_j)^T \Sigma_j^{-1} (x^{(i)} - \mu_j) \right) \phi_j}{w_j^{(i)}}$$

- Now, we need to maximize this lower bound for each of the three parameters. Many of the derivatives on vector/matrix are based on the Matrix Cookbook.



$$\begin{aligned}
\nabla_{\mu_j} ll &= \nabla_{\mu_j} \sum_i^m w_j^{(i)} \ln \frac{\frac{1}{\sqrt{(2\pi)^n |\Sigma_j|}} \exp\left(-\frac{1}{2} (x^{(i)} - \mu_j)^T \Sigma_j^{-1} (x^{(i)} - \mu_j)\right) \phi_j}{w_j^{(i)}} \\
&= \nabla_{\mu_j} \sum_i^m w_j^{(i)} \left[ \ln \frac{\frac{1}{\sqrt{(2\pi)^n |\Sigma_j|}} \phi_j}{w_j^{(i)}} + \ln \exp\left(-\frac{1}{2} (x^{(i)} - \mu_j)^T \Sigma_j^{-1} (x^{(i)} - \mu_j)\right) \right] \\
&= \nabla_{\mu_j} \sum_i^m w_j^{(i)} \left[ \frac{1}{2} (x^{(i)} - \mu_j)^T \Sigma_j^{-1} (x^{(i)} - \mu_j) \right] \\
&= -\frac{1}{2} \sum_i^m w_j^{(i)} \nabla_{\mu_j} \left[ (x^{(i)} - \mu_j)^T \Sigma_j^{-1} (x^{(i)} - \mu_j) \right] \\
&[\text{For } f(x) = x^T A x : \nabla_x f(x) = (A + A^T)x] \\
&= \frac{1}{2} \sum_i^m w_j^{(i)} \nabla_{(x^i - \mu_j)} \left[ (x^{(i)} - \mu_j)^T \Sigma_j^{-1} (x^{(i)} - \mu_j) \right] \\
&= \frac{1}{2} \sum_i^m w_j^{(i)} \left[ \left( \Sigma_j^{-1} + (\Sigma_j^{-1})^T \right) (x^{(i)} - \mu_j) \right]
\end{aligned}$$

$$\nabla_{\mu_j} ll = 0$$

$$\sum_i^m w_j^{(i)} \left[ \Sigma_j^{-1} \left( x^{(i)} - \mu_j \right) \right] = 0$$

$$\sum_i^m w_j^{(i)} \left( x^{(i)} - \mu_j \right) = 0$$

$$\sum_i^m w_j^{(i)} x^{(i)} = \sum_i^m w_j^{(i)} \mu_j$$

$$\mu_j = \frac{\sum_i^m w_j^{(i)} x^{(i)}}{\sum_i^m w_j^{(i)}}$$

Derivative of  $\Sigma_j$

$$\begin{aligned}
 &= \sum_i^m w_j^{(i)} \nabla_{\Sigma_j} \left[ \ln \frac{1}{\sqrt{|\Sigma_j|}} - \frac{1}{2} (x^{(i)} - \mu_j)^T \Sigma_j^{-1} (x^{(i)} - \mu_j) \right] \\
 &= -\frac{1}{2} \sum_i^m w_j^{(i)} \left[ \frac{\partial \ln |\Sigma_j|}{\partial \Sigma_j} + \frac{\partial}{\partial \Sigma_j} (x^{(i)} - \mu_j)^T \Sigma_j^{-1} (x^{(i)} - \mu_j) \right]
 \end{aligned}$$

First, we consider the derivative of the first term in the square bracket:

$$\begin{aligned}
 \frac{\partial \ln |\Sigma_j|}{\partial \Sigma_j} &= \frac{1}{|\Sigma_j|} \frac{\partial |\Sigma_j|}{\partial \Sigma_j} \\
 &= \frac{1}{|\Sigma_j|} |\Sigma_j| (\Sigma_j^{-1})^Y \\
 &= \Sigma_j^{-1}
 \end{aligned}$$

Then, we do the second term

$$\frac{\partial}{\partial \Sigma_j} (x^{(i)} - \mu_j)^T \Sigma_j^{-1} (x^{(i)} - \mu_j) = -\Sigma_j^{-1} (x^{(i)} - \mu_j) (x^{(i)} - \mu_j)^T \Sigma_j^{-1}$$

Combined these results back and set it to zero, we have:

$$\begin{aligned}\nabla_{\Sigma_j} l &= -\frac{1}{2} \sum_i^m w_j^{(i)} \left[ \Sigma_j^{-1} - \Sigma_j^{-1} (x^{(i)} - \mu_j) (x^{(i)} - \mu_j)^T \Sigma_j^{-1} \right] \\ &= -\frac{1}{2} \sum_i^m w_j^{(i)} \left[ I - \Sigma_j^{-1} (x^{(i)} - \mu_j) (x^{(i)} - \mu_j)^T \right] \Sigma_j^{-1} \stackrel{\text{get}}{=} 0\end{aligned}$$

Rearrange the equation and we have:

$$\begin{aligned}\sum_i^m w_j^{(i)} \left[ \Sigma_j - (x^{(i)} - \mu_j) (x^{(i)} - \mu_j)^T \right] &= 0 \\ \sum_i^m w_j^{(i)} \Sigma_j &= \sum_i^m w_j^{(i)} (x^{(i)} - \mu_j) (x^{(i)} - \mu_j)^T \\ \frac{\partial A^{-1}}{\partial x} &= -A^{-1} \cdot \frac{\sum_i^m w_j^{(i)} (x^{(i)} - \mu_j) (x^{(i)} - \mu_j)^T}{\sum_i^m w_j^{(i)}}\end{aligned}$$

Derivative of  $\phi_j$

This is relatively simpler but we need to apply Lagrange multipliers because  $\sum_j \phi_j = 1$ .

$$\begin{aligned} ll &= \sum_i^m \sum_l^k w_l^{(i)} \ln \frac{\frac{1}{\sqrt{(2\pi)^n |\Sigma_l|}} \exp\left(-\frac{1}{2} (x^{(i)} - \mu_l)^T \Sigma_l^{-1} (x^{(i)} - \mu_l)\right) \phi_l}{w_l^{(i)}} \\ &= \sum_i^m \sum_l^k w_l^{(i)} \ln \phi_l \end{aligned}$$

We need to construct Lagrangian, with  $\lambda$  as the Lagrange multiplier:

$$\mathcal{L}(\phi) = ll + \lambda \left( \sum_l^k \phi_l - 1 \right)$$

We will take derivative on  $\mathcal{L}$  and set it to zero:

$$\begin{aligned}\frac{\partial \mathcal{L}(\phi)}{\partial \phi_j} &= \frac{\partial}{\partial \phi_j} \left[ l + \lambda \left( \sum_l^k \phi_l - 1 \right) \right] \\ &= \sum_i w_j^{(i)} \frac{1}{\phi_j} + \lambda \stackrel{\text{set}}{=} 0\end{aligned}$$

Rearrange and we will have  $\phi_j = -\frac{\sum_i w_j^{(i)}}{\lambda}$ . Recall that  $\sum_j \phi_j = 1$ , we have:

$$\begin{aligned}\sum_j \phi_j &= \sum_j -\frac{\sum_i w_j^{(i)}}{\lambda} = 1 \\ \lambda &= -\sum_j \sum_i w_j^{(i)} \\ &= -\sum_j \sum_i p(z^{(i)} = j \mid x^{(i)}) \\ &= -\sum_i 1 = -m\end{aligned}$$

Finally, we have:

$$\phi_j = \frac{\sum_i w_j^{(i)}}{m}$$

# K-Means from GMM perspective

From GMM to K-means:

- Fix covariances to be identical  $\Sigma_c = I$

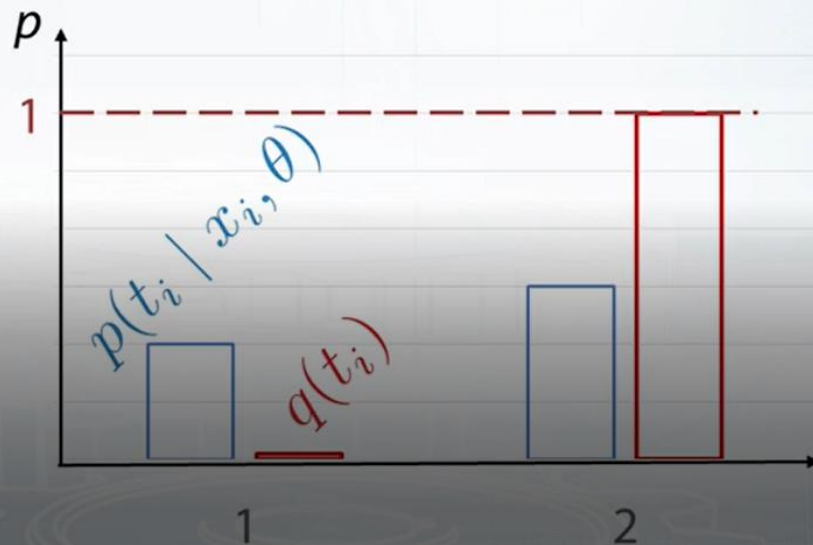
- Fix weights to be uniform  $\pi_c = \frac{1}{\# \text{ of Gaussians}}$

$$p(x_i \mid t_i = c, \theta) = \frac{1}{Z} \exp(-0.5 \|x_i - \mu_c\|^2)$$

## E-step

$$q^{k+1} = \arg \min_{q \in \mathcal{Q}} \mathcal{KL} [q(T) \parallel p(T \mid X, \theta^k)]$$

Where  $\mathcal{Q}$  is the set of delta-functions





## E-step

$$q^{k+1}(t_i) = \begin{cases} 1 & \text{if } t_i = c_i \\ 0 & \text{otherwise} \end{cases}$$

$$c_i = \arg \max_c p(t_i = c \mid x_i, \theta) = \arg \min_c \|x_i - \mu_c\|^2$$

$$\begin{aligned} p(t_i \mid x_i, \theta) &= \frac{1}{Z} p(x_i \mid t_i, \theta) p(t_i \mid \theta) \\ &= \frac{1}{Z} \exp(-0.5\|x_i - \mu_c\|^2) \pi_c \end{aligned}$$

## E-step

$$q^{k+1}(t_i) = \begin{cases} 1 & \text{if } t_i = c_i \\ 0 & \text{otherwise} \end{cases}$$

$$c_i = \arg \min_c \|x_i - \mu_c\|^2$$

Exactly like in K-Means!

# M-Step

$$\max . \sum_{i=1}^N \mathbb{E}_{q(t_i)} \log p(x_i, t_i \mid \mu)$$

$$\mu_c = \frac{\sum_{i=1}^N (q(t_i = c) \cdot x_i)}{\sum_{i=1}^N q(t_i = c)} \qquad \mu_c = \frac{\sum_{i=1}^N x_i}{\#\{i : c_i^* = 0\}}$$

$$q(t_i) = \begin{cases} 1 & \text{if } t_i = c_i^* \\ 0 & \text{if } t_i \neq c_i^* \end{cases}$$

# Reference

- Bayesian Methods for Machine Learning, HSE university, Coursera