

1 First Exponential to Die

$$X \sim \text{Exp}(\lambda) \quad f_X(x) = \lambda e^{-\lambda x} \quad F_X(x) = 1 - e^{-\lambda x} = \Pr[X \leq x]$$

Let X and Y be $\text{Exponential}(\lambda_1)$ and $\text{Exponential}(\lambda_2)$ respectively, independent. What is



λ
arrivals
time

$$\Pr(\min(X, Y) = X),$$

the probability that the first of the two to die is X ?

$$\Pr[\min(X, Y) = X] = \Pr[X < Y]$$

$$\rightarrow = \int_0^{\infty} \Pr[Y > x | X = x] f_X(x) dx$$

$$= \int_0^{\infty} (1 - F_Y(x)) f_X(x) dx$$

$$= \int_0^{\infty} e^{-\lambda_2 x} \lambda_1 e^{-\lambda_1 x} dx$$

$$= \int_0^{\infty} \lambda_1 e^{-(\lambda_1 + \lambda_2)x} dx = \frac{\lambda_1}{\lambda_1 + \lambda_2}$$

$$\min(X, Y) = X$$

$$\min(5, 3) = 3$$

3 vs 5?

$$\Pr[Y > X]$$

$$= \sum_{x \in X} \Pr[Y > x \cap X = x]$$

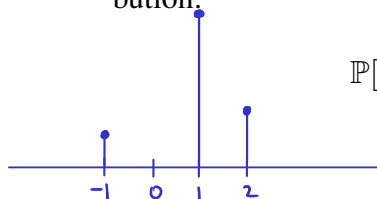
$$= \sum_{x \in X} \Pr[Y > x | X = x] \Pr[X = x]$$

$$= \int_0^{\infty} \Pr[Y > x] f_X(x) dx$$

2 Chebyshev's Inequality vs. Central Limit Theorem

Let n be a positive integer. Let X_1, X_2, \dots, X_n be i.i.d. random variables with the following distribution:

$$\Pr[X_i = -1] = \frac{1}{12}; \quad \Pr[X_i = 1] = \frac{9}{12}; \quad \Pr[X_i = 2] = \frac{2}{12}.$$



(a) Calculate the expectations and variances of X_1 , $\sum_{i=1}^n X_i$, $\sum_{i=1}^n (X_i - \mathbb{E}[X_i])$, and

$$\mathbb{E}[X_i] = -1 \cdot \frac{1}{12} + 1 \cdot \frac{9}{12} + 2 \cdot \frac{2}{12} = 1$$

$$\text{Var}(X_i) = \mathbb{E}[(X_i - \mathbb{E}[X_i])^2] = (-2)^2 \cdot \frac{1}{12} + 0^2 \cdot \frac{9}{12} + 1^2 \cdot \frac{2}{12} = \frac{1}{2}$$

$$Z_n = \frac{\sum_{i=1}^n (X_i - \mathbb{E}[X_i])}{\sqrt{n/2}}$$

$$\mathbb{E}[Z_n] = 0$$

$$\text{Var}(Z_n) = \text{Var}\left(\frac{1}{\sqrt{n/2}} \sum_{i=1}^n (X_i - \mathbb{E}[X_i])\right) = \frac{2}{n} \cdot \frac{n}{2} = 1$$

$$\mathbb{E}\left[\sum_{i=1}^n X_i\right] = n$$

$$\mathbb{E}\left[\sum_{i=1}^n (X_i - \mathbb{E}[X_i])\right] = \mathbb{E}\left[\sum_{i=1}^n (X_i - 1)\right] = \left(\sum_{i=1}^n \mathbb{E}[X_i] - n\right) = 0$$

$$\text{Var}\left(\sum_{i=1}^n X_i\right) = \sum_{i=1}^n \text{Var}(X_i) = \frac{n}{2}$$

$$\text{Var}\left(\sum_{i=1}^n (X_i - \mathbb{E}[X_i])\right) = \sum_{i=1}^n \text{Var}(X_i - \mathbb{E}[X_i]) = \sum_{i=1}^n \text{Var}(X_i) = \frac{n}{2}$$

(b) Use Chebyshev's Inequality to find an upper bound b for $\Pr[|Z_n| \geq 2]$.

$$\Pr[|Z_n| \geq 2] = \Pr[|Z_n - 0| \geq 2] \leq \frac{\text{Var}(Z_n)}{2^2} = \frac{1}{4}$$

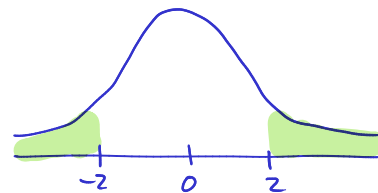
$$\frac{\sqrt{n} \left(\frac{1}{n} \sum_{i=1}^n X_i - \mathbb{E}[X_i] \right)}{\sqrt{1/2}}$$

$$Z_n = \frac{\sum_{i=1}^n (X_i - \mathbb{E}[X_i])}{\sqrt{n/2}} = \frac{\frac{1}{\sqrt{n}} \sum_{i=1}^n X_i - \mathbb{E}[X_i]}{\sqrt{1/2}}$$

(c) Can you use b to bound $\mathbb{P}[Z_n \geq 2]$ and $\mathbb{P}[Z_n \leq -2]$?

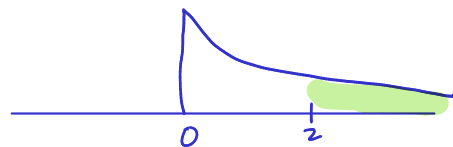
$$\Pr[Z_n \geq 2] \leq \Pr[|Z_n| \geq 2] \leq \frac{1}{4} = 0.25$$

$$\Pr[Z_n \leq -2] \leq \Pr[|Z_n| \geq 2] \leq \frac{1}{4}$$



(d) As $n \rightarrow \infty$, what is the distribution of Z_n ? CLT

$$\frac{\bar{X} - \mathbb{E}[X_i]}{\sqrt{\text{Var}(X_i)}} \rightarrow \mathcal{N}(0,1) \quad n \text{ iid } X_1, \dots, X_n \quad \frac{(\bar{X} - \mathbb{E}[X_i])}{\sqrt{\text{Var}(X_i)}} \rightarrow \mathcal{N}(0,1)$$



(e) We know that if $Z \sim \mathcal{N}(0,1)$, then $\mathbb{P}[|Z| \leq 2] = \Phi(2) - \Phi(-2) \approx 0.9545$. As $n \rightarrow \infty$, can you provide approximations for $\mathbb{P}[Z_n \geq 2]$ and $\mathbb{P}[Z_n \leq -2]$?

$$\frac{1}{n} \sum_{i=1}^n (X_i - \mathbb{E}[X_i])$$

$$\bar{X} = \frac{1}{n} \sum X_i$$

Can use symmetry!

$$\Pr[|Z| \geq 2] = 1 - 0.9545 = 0.0455$$

$$\Pr[Z_n \geq 2] = \Pr[Z_n \leq -2] = \frac{0.0455}{2} = 0.02275$$

3 Why Is It Gaussian?

Let X be a normally distributed random variable with mean μ and variance σ^2 . Let $Y = aX + b$, where $a > 0$ and b are non-zero real numbers. Show explicitly that Y is normally distributed with mean $a\mu + b$ and variance $a^2\sigma^2$. The PDF for the Gaussian Distribution is $\frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$. One approach is to start with the cumulative distribution function of Y and use it to derive the probability density function of Y .

[1. You can use without proof that the pdf for any gaussian with mean and sd is given by the formula

$\frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$ where μ is the mean value for X and σ^2 is the variance. 2. The derivative of CDF gives PDF.]

$$\mathbb{E}[Y] = a\mu + b \quad \text{Var}(Y) = a^2\sigma^2$$

$$\text{Var}(aX + b) = a^2 \text{Var}(X) = a^2\sigma^2$$