

Sheet 11

problem 1:

a-) We have to show that the acf  
is absolutely summable

$$\sum_{i=0}^{\infty} |\gamma_i| \stackrel{\text{Insert ACF}}{=} \sum_{i=0}^{\infty} \left| \frac{\alpha^i \cdot \sigma^2}{1-\alpha^2} \right|$$

$$\stackrel{\text{geom series}}{=} \left| \frac{\sigma^2}{(1-\alpha^2)(1-\alpha)} \right| < \infty$$

b-) Confidence Interval for  $\mu$  AR(1) - process  
is stationary  $\rightarrow$  Wold decomposition

$$X_t = \mu + \alpha X_{t-1} + \varepsilon_t$$

$$\Leftrightarrow (L - \alpha B) X_t = \mu + \varepsilon_t$$

Chapter 7

$$\Leftrightarrow X_t = \sum_{i=0}^{\infty} \alpha^i B^i \mu + \sum_{i=0}^{\infty} \alpha^i B^i \varepsilon_t$$

St. 33

$$= \sum_{i=0}^{\infty} M \alpha^i + \sum_{i=0}^{\infty} \alpha^i \varepsilon_{t-i}$$

$$\text{Series} = \frac{M}{1-\alpha} + \sum_{i=0}^{\infty} \alpha^i \varepsilon_{t-i} \quad (\text{W.O})$$

$$\rightarrow E \left[ \sum_{i=0}^{\infty} \alpha^i \varepsilon_{t-i} \right]$$

$$\text{Since the } \sum_{i=-\infty}^{+\infty} |a_i| \stackrel{\text{geom Series}}{=} \frac{1}{1-\alpha} < \infty$$

$$= \sum_{i=0}^{\infty} \alpha^i E[\varepsilon_{t-i}]$$

$$= 0$$

$$\text{Theorem 8.5} \Rightarrow \sqrt{T} \left( \bar{X}_t - \frac{\mu}{1-\alpha} \right) \xrightarrow{D} N\left(0, \frac{\sigma^2}{(1-\alpha)^2}\right)$$

$$\Rightarrow \sqrt{T} \left( \bar{X}_t - \frac{\mu}{1-\alpha} \right) \frac{1-\alpha}{\sigma} \xrightarrow{D} N(0, 1)$$

We know,

$$0.95 = \lim_{T \rightarrow \infty} P\left(-1.96 \leq \sqrt{T}\left(\bar{X}_t - \frac{\mu}{1-\alpha}\right) \leq 1.96\right)$$

$$= \lim_{T \rightarrow \infty} P\left(\frac{1.96}{\sqrt{T}} \cdot \sigma + \frac{\bar{X}_t}{1-\alpha} \geq \mu \geq -1.96 \cdot \frac{\sigma}{\sqrt{T}} + \frac{\bar{X}_t}{1-\alpha}\right)$$

$$0.95 \underset{\text{insert}}{\approx} P\left(\frac{1.96 \cdot \sqrt{4}}{\sqrt{100}} + \frac{\bar{X}_t}{0.3} \geq \mu \geq \frac{-1.96 \cdot \sqrt{4}}{\sqrt{100}} + \frac{\bar{X}_t}{0.3}\right)$$

$$\Rightarrow CI_{0.95} \approx \left[-\frac{1.96}{5} + \frac{\bar{X}_t}{0.3}, \frac{1.96}{5} + \frac{\bar{X}_t}{0.3}\right]$$

Realized Confidence Interval

$$CI_{0.95, \text{realized}} = \left[-\frac{1.95}{5} + \frac{0.157}{0.3}, \frac{1.96}{5} + \frac{0.157}{0.3}\right]$$

$$= [-0.35, 0.44]$$

0 lies in this interval  $\Rightarrow$  We don't reject the null

Question 2: b-) is resolved in next week

Does  $\sqrt{T}(\bar{X}_t - \frac{\mu}{1-\alpha})$  really converge in distribution

$$\text{to } N\left(0, \frac{\sigma^2}{(1-\alpha)^2}\right)$$

Claim: Yes.  $\sqrt{T}(\bar{X}_t - \frac{\mu}{1-\alpha}) \xrightarrow[\text{Theorem 8.5}]{D} N\left(0, \sum_{h=-\infty}^{\infty} g(h)\right)$

$$\sum_{h=-\infty}^{\infty} g(h) \stackrel{(a)}{=} \sum_{h=-\infty}^{\infty} \frac{a^{|h|}}{1-\alpha^2} = \frac{\sigma^2}{1-\alpha^2} \sum_{h=-\infty}^{\infty} a^h$$

$$= \frac{\sigma^2}{1-\alpha^2} \left[ \sum_{h=0}^{\infty} a^h + \underbrace{\sum_{h=-\infty}^{-1} a^{|h|}}_{\underbrace{\phantom{0}}_{\text{brace}} \text{brace}} \right]$$

$$\sum_{h=1}^{\infty} a^h = \sum_{h=0}^{\infty} a^{h+1} = \sum_{h=0}^{\infty} a^h \cdot a$$

$$= \frac{\sigma^2}{1-\alpha^2} \left[ \sum_{h=0}^{\infty} a^h + a \sum_{h=0}^{\infty} a^h \right] = a \sum_{h=0}^{\infty} a^h$$

$$\text{geom series} = \frac{\sigma^2}{1-\alpha^2} \left[ \frac{1}{1-a} + \frac{a}{1-a} \right] = \frac{\sigma^2}{1-a} \left[ \frac{1+a}{1-a} \right]$$

$$\begin{array}{l} \text{Binomial} \\ \hline \text{formula} \end{array} \quad \frac{\sigma^2}{(1-\alpha)(1+\alpha)} \left[ \frac{1+\alpha}{1-\alpha} \right] = \frac{\sigma^2}{(1-\alpha)^2}$$

(2) R exercise

problem 3

(a)  $X_t = aX_{t-1} + \epsilon_t \iff X_t - aX_{t-1} = \epsilon_t$

$\Rightarrow \epsilon_t = Wx_t$ , where  $W$  is the linear filter

$$L-a \quad B$$

$\Rightarrow X_t = W^{-1}\epsilon_t$ , since we know from cpt 7  
 that such an inverse filter exists and  
 is given by  $W^{-1}(B) = \sum_{s=0}^{\infty} a^s B^s$ , as  $|a| < L$

Applying this filter

$$X_t = \sum_{s=0}^{\infty} a^s \epsilon_{t-s} \Rightarrow \text{This is MA}(\infty) \text{ by definition}$$

b-) Chapter 9, Slides 46-47

$$c-) f(w) = \frac{1}{2\pi} \chi(e^{iw})$$

$$\stackrel{\text{insert}}{=} \frac{\sigma^2}{2\pi} W^{-1}(e^{iw}) W^{-1}(e^{-iw})$$

$$\stackrel{\text{hint}}{=} \frac{\sigma^2}{2\pi} \frac{1}{(1-\alpha e^{iw})(1-\alpha e^{-iw})}$$

$$= -\frac{\sigma^2}{2\pi(1+\alpha^2-\alpha(e^{-100}+e^{100}))} \stackrel{\text{cancel}}{=} \frac{\sigma^2}{2\pi (1+\alpha^2-2\alpha \cdot \cos(w))}$$

Important Note

$$\sum_{i=0}^{\infty} a^i; \begin{array}{l} \text{geom series} \\ \frac{1}{1-a} \end{array}$$

if  $|a| < 1$

$$\sum_{i=0}^{\infty} a^i B^i \neq \frac{1}{1-aB}$$