

Sheet 1problem 1

Let y_t , $t = -4, \dots, 0, \dots, 4$ be a TS with realizations $y = (y_{-4}, \dots, y_4)$. We assume $y_t = \mu_t + s_t + \varepsilon_t$

1) We assume $\mu_t = \beta_0 + \beta_1 \cdot t$ & want to estimate β_0 & β_1

$$\text{Model: } y_t = \beta_0 + \beta_1 \cdot t + u_t = (t+1) \begin{pmatrix} \beta_0 \\ \beta_1 \end{pmatrix} + u_t$$

$$\begin{pmatrix} y_{-4} \\ \vdots \\ y_0 \\ \vdots \\ y_4 \end{pmatrix} = \begin{pmatrix} -4 \\ \vdots \\ 0 \\ \vdots \\ 4 \end{pmatrix} = \underbrace{\begin{pmatrix} \beta_0 \\ \beta_1 \end{pmatrix}}_{=\beta} + \begin{pmatrix} u_{-4} \\ \vdots \\ u_0 \\ \vdots \\ u_4 \end{pmatrix}$$

$$\Rightarrow \hat{\beta} = (A^T A)^{-1} A^T y \quad (\text{I})$$

Here we use

$$\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{t} \quad \text{and} \quad \hat{\beta}_1 = \frac{\bar{ty} - \bar{t}\bar{y}}{\bar{t^2} - \bar{t}^2} \quad (\text{#}) \quad \text{from the lecture}$$

$$\bar{y} = 10.7, \bar{t} = 0, \bar{t^2} = \sum_{t=-4}^4 t^2 = \frac{2}{3} \sum_{t=1}^4 t^2 = \frac{60}{3}$$

$$\bar{ty} = \frac{1}{3} \sum_{t=-4}^4 t \cdot y = \frac{65 \cdot 1}{9} \Rightarrow \hat{\beta}_0 = 10.7, \hat{\beta}_1 \approx 1.1$$

$$\hat{M}_t = \hat{\beta}_0 + \hat{\beta}_1 \cdot t \rightarrow (6.3, 7.4, 8.5, 9.6, 10.7, 11.8, 12.9, 14.0, 15.1)$$

$$y_t^{\text{demand}} = y_t - \hat{M}_t \rightarrow \left(\frac{3.6}{S_1}, \frac{-4.6}{S_2}, \frac{1.1}{S_3}, \frac{0.8}{S_1}, \frac{-5.0}{S_2}, \frac{3.9}{S_3}, \frac{3.0}{S_1}, \frac{-5.2}{S_2}, \frac{2.1}{S_3} \right)$$

$$2) \text{ Model: } Y_t = \delta_1 \cdot S_{1,t} + \dots + \delta_3 \cdot S_{3,t} + \varepsilon_t, \quad \sum_{i=1}^3 \delta_i = 0, \hat{\delta} = \bar{y} - \bar{M}_{t=1}$$

S_1 means season 1, S_2 means season 2 and so on

Season	1	2	3
	3.6	-4.6	1.1
	0.8	-5.0	3.9
	3.0	-5.2	2.1
average	2.5	-4.9	2.4

$$\Rightarrow \hat{\delta}_i = \begin{cases} -2.5 & i=1 \\ 4.3 & i=2 \\ -2.4 & i=3 \end{cases}$$

problem 2

Let $\{X_t\}_{t \in \{1, \dots, T\}}$ be a TS with $X_t = \beta_0 + \beta_1 \cdot t + \varepsilon_t$

$$\text{cov}(\varepsilon_t, \varepsilon_{t+h}) = \begin{cases} \sigma^2 & h=0 \\ \rho \sigma^2 & h=\pm 1 \\ 0 & |h| \geq 2 \end{cases} \quad E(\varepsilon_t) = 0$$

a) $E(\hat{\beta})$ using (I)

$$\begin{aligned}
 & E((A^T A)^{-1} A^T y) \\
 &= E((A^T A)^{-1} A^T (A\beta + \varepsilon)) \\
 &= E(\underbrace{(A^T A)^{-1} A^T}_{\text{I}} \cdot A \beta + (A^T A)^{-1} A^T \varepsilon) \\
 &= E(\beta + (A^T A)^{-1} A^T \varepsilon) \\
 &= \beta + E((A^T A)^{-1} A^T \varepsilon) \\
 &\quad \underbrace{= (A^T A)^{-1} A^T \cdot E(\varepsilon)}_{=0}
 \end{aligned}$$

$$= \beta$$

$$\begin{aligned}
 b) \text{Var}(\hat{\beta}) & \left[= \begin{pmatrix} \text{Var}(\hat{\beta}_0) & \text{Cov}(\hat{\beta}_0, \hat{\beta}_1) \\ \cdot & \text{Var}(\hat{\beta}_1) \end{pmatrix} \right] = \text{Cov}(\hat{\beta}, \hat{\beta}) = \text{Cov}\left((A^T A)^{-1} A^T (A\beta + \varepsilon), (A^T A)^{-1} A^T (A\beta + \varepsilon)\right) \\
 &= \text{Cov}\left((A^T A)^{-1} A^T \varepsilon, \varepsilon^T A (A^T A)^{-1}\right) \\
 &= (A^T A)^{-1} A^T \text{Cov}(\varepsilon, \varepsilon^T) A (A^T A)^{-1}
 \end{aligned}$$

$$\text{Cov}(\varepsilon, \varepsilon^T) = \begin{pmatrix} \sigma^2 & \rho\sigma^2 & & \\ \rho\sigma^2 & \sigma^2 & \rho\sigma^2 & \\ & \rho\sigma^2 & \sigma^2 & \\ & & \rho\sigma^2 & \rho\sigma^2 \\ & & & \rho\sigma^2 \end{pmatrix}$$

$$= \sigma^2 \cdot I_{T \times T} + \rho \sigma^2 \begin{pmatrix} 0 & 1 & & & \\ 1 & 0 & & & \\ & & \ddots & & \\ & & & 0 & 1 \\ & & & 1 & 0 \end{pmatrix}$$

$$= (A^T A)^{-L} A^T \left(\sigma^{-2} I_{TxT} + \rho \sigma^{-2} \begin{pmatrix} 0 & 1 & & & \\ 1 & 0 & & & \\ 0 & & 0 & & \\ 0 & & 0 & -1 & 0 \\ 0 & & 0 & 0 & 0 \end{pmatrix} \right) A (A^T A)^{-L}$$

$$= \underbrace{\sigma^{-2} (A^T A)^{-L}}_{\text{Variance without Autocorrelation.}} + \underbrace{\rho \sigma^{-2} (A^T A)^{-L} A^T \left(\begin{pmatrix} 0 & 1 & & & \\ 1 & 0 & & & \\ 0 & & 0 & & \\ 0 & & 0 & -1 & 0 \\ 0 & & 0 & 0 & 0 \end{pmatrix} \right) A (A^T A)^{-L}}_{\text{Influence of Autocorrelation}}$$

\Rightarrow The variance of estimator $\hat{\beta}$ is influenced

$$A^T A = \begin{pmatrix} 1 & & & & \\ & \ddots & & & \\ & & 1 & & \\ & & & \ddots & \\ & & & & T \end{pmatrix} \begin{pmatrix} 1 & & & & \\ & \ddots & & & \\ & & 1 & & \\ & & & \ddots & \\ & & & & T \end{pmatrix} = \begin{pmatrix} T & & & & \\ & \sum_{t=1}^T t^2 = \frac{T(T+1)}{2} & & & \\ & & T & & \\ & & & \ddots & \\ & & & & \frac{T(T+1)}{2} \end{pmatrix}$$

$$\sum_{t=1}^T t^2 = \frac{T(T+1)(2T+1)}{6}$$

$$(A^T A)^{-1} = \frac{1}{\sum_{t=1}^T t^2} \begin{pmatrix} (T+1)(2T+1) & -(T+1) \\ -(T+1) & T \end{pmatrix}$$

$$\text{Var}(\hat{\beta}_0) = \frac{2\sigma^2}{T^2(T+1)} (T(2T+1) + \rho^4 (T^2 - 2T - 2)) \xrightarrow{T \rightarrow \infty} 0$$

$$\text{Var}(\hat{\beta}_1) = \frac{12\sigma^2}{(T-1)T(T+1)} \left(1 + \rho^2 \cdot \frac{T-2}{T+1} \right) \xrightarrow{T \rightarrow \infty} 0$$

\Rightarrow Together with unbiasedness we conclude that

$\hat{\beta}_0$ & $\hat{\beta}_1$ are consistent.

problem 4

\hat{p} : Autocorrelation
 $p: \mathbb{Z} \rightarrow \mathbb{R}$, st $\hat{p}(h) = \frac{\frac{1}{T} \sum_{t=1}^{T-h} (y_{t+h} - \bar{y})(y_t - \bar{y})}{\frac{1}{T} \sum_{t=1}^T (y_t - \bar{y})^2}$
 wlog $h > 0$

a-) $|\hat{p}(h)| \leq 1$, Cauchy-Swartz: Let x, y be vectors from vector space with an inner product $\langle \cdot, \cdot \rangle$
 then $|\langle x, y \rangle|^2 \leq \langle x, x \rangle \langle y, y \rangle$

$$|\hat{p}(h)|^2 \leq \frac{\langle \vec{y}_{t+h}^{T-h}, \vec{y}_t^{T-h} \rangle^2}{\langle y_t^T, y_t^T \rangle^2}$$

$$\vec{y}_{t+h} = \begin{pmatrix} y_{t+h} - \bar{y} \\ y_{t+h+1} - \bar{y} \\ \vdots \\ y_{T-h} - \bar{y} \end{pmatrix}, \quad y_t = \begin{pmatrix} y_t - \bar{y} \\ \vdots \\ y_{T-h} - \bar{y} \end{pmatrix}$$

$$|\hat{p}(h)|^2 \leq \frac{\langle \vec{y}_{t+h}^{T-h}, \vec{y}_t^{T-h} \rangle^2}{\langle y_t^T, y_t^T \rangle^2} \stackrel{CS}{\leq} \frac{\langle \vec{y}_{t+h}^{T-h}, \vec{y}_{t+h}^{T-h} \rangle \langle \vec{y}_t^{T-h}, \vec{y}_t^{T-h} \rangle}{\langle y_t^T, y_t^T \rangle \langle y_t^T, y_t^T \rangle} \leq$$

$$\frac{\langle y_t^T, y_t^T \rangle \langle y_t^T, y_t^T \rangle}{\underbrace{\langle y_t^T, y_t^T \rangle \langle y_t^T, y_t^T \rangle}_{=1}}$$

