

Introduction to Data Science With Probability and Statistics

Lecture 21: Unbiased Estimators and Mean Squared Error

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Estimators

- One of the tasks is to use a dataset to estimate a quantity of interest.
- We should be able to deal with a situation where the dataset is modeled as one of the parameters of the model distribution or as a certain function of the parameters.

ESTIMATE. An *estimate* is a value t that only depends on the dataset x_1, x_2, \dots, x_n , i.e., t is some function of the dataset only:

$$t = h(x_1, x_2, \dots, x_n).$$

EG: $\bar{x} = \frac{\sum x_i}{n}$

Estimators

One can often think of several estimates for the parameter of interest. This raises questions like:-

- When is one estimate better than another?
- Does there exist a best possible estimate?

We can never say which of the two estimate values ' e_1 ' and ' e_2 ' computed from a dataset is closer to the 'true' parameter. This is because:-

- The measurements and the corresponding estimates are subject to randomness.
- One of the things we can say for each of them is how likely it is that they are within a given distance from the 'true' parameter.

Note that estimators are special cases of 'sample statistics'.



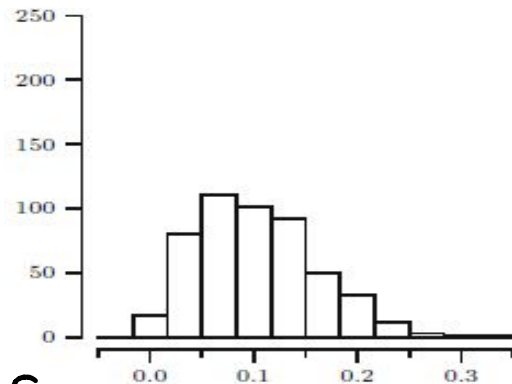
Estimators

ESTIMATOR. Let $t = h(x_1, x_2, \dots, x_n)$ be an estimate based on the dataset x_1, x_2, \dots, x_n . Then t is a realization of the random variable

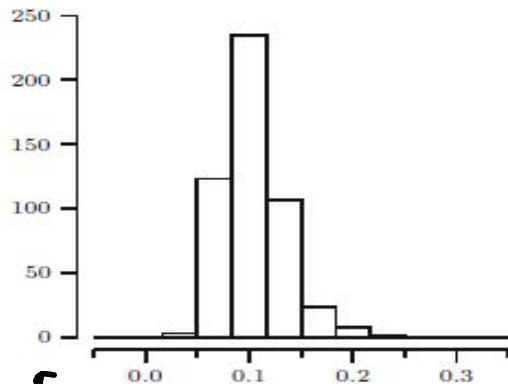
$$T = h(X_1, X_2, \dots, X_n).$$

The random variable T is called an *estimator*.

$$\text{eg: } \bar{X}_n = \frac{\sum x_i}{n}$$



E_1



E_2

sampling
distributions
of
estimators

The sampling distribution and unbiasedness

THE SAMPLING DISTRIBUTION. Let $T = h(X_1, X_2, \dots, X_n)$ be an estimator based on a random sample X_1, X_2, \dots, X_n . The probability distribution of T is called the *sampling distribution* of T .

DEFINITION. An estimator T is called an *unbiased* estimator for the parameter θ , if

$$E[T] = \theta$$

irrespective of the value of θ . The difference $E[T] - \theta$ is called the *bias* of T ; if this difference is nonzero, then T is called *biased*.

Unbiased estimators for expectation and variance

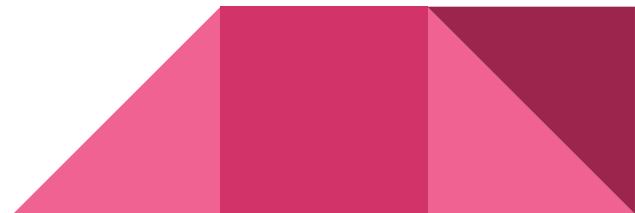
UNBIASED ESTIMATORS FOR EXPECTATION AND VARIANCE. Suppose X_1, X_2, \dots, X_n is a random sample from a distribution with finite expectation μ and finite variance σ^2 . Then

$$\bar{X}_n = \frac{X_1 + X_2 + \dots + X_n}{n}$$

is an *unbiased estimator* for μ and

$$S_n^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X}_n)^2$$

is an *unbiased estimator* for σ^2 .



Unbiased estimators for expectation and variance

$$\bar{X}_n = \frac{X_1 + X_2 + \dots + X_n}{n}$$

$$E[\bar{X}_n] = E\left[\frac{X_1 + \dots + X_n}{n}\right]$$

$$= \frac{1}{n} E[X_1 + \dots + X_n]$$

$$= \frac{1}{n} (E[X_1] + \dots + E[X_n])$$

$$= \frac{1}{n} \cdot n\mu = \mu$$

$$\therefore \boxed{E[\bar{X}_n] = \mu}$$

$$E[S_n^2] = \frac{1}{n-1} \sum_{i=1}^n E[(X_i - \bar{X}_n)^2]$$

$$E\left[\sum (X_i^2 - 2\bar{X}_n X_i + \bar{X}_n^2)\right]$$

$$= E\left[\sum X_i^2 - 2\bar{X}_n \sum X_i + n\bar{X}_n^2\right]$$

$$= E\left[\sum X_i^2 - 2\bar{X}_n \cdot n\bar{X}_n + n\bar{X}_n^2\right]$$

$$= E\left[\sum X_i^2 - n\bar{X}_n^2\right] = \sum (E[X_i^2] - E[n\bar{X}_n^2])$$

Now;

$$E[X_i^2] = \text{Var}(X_i) + \mu^2$$

Unbiased estimators for expectation and variance

$$\begin{aligned}\therefore \sum (E[x_i^2] - n(E[\bar{x}^2])) &= \sum [(\sigma^2 + \mu^2) - n \cdot (\frac{\sigma^2}{n} + \mu^2)] \\ &= n\sigma^2 + n\mu^2 - \sigma^2 - n\mu^2 \\ &= \sigma^2(n-1)\end{aligned}$$

- This explains why we divide by $n - 1$ in the formula for S_n^2 ; only in this case S_n^2 is an unbiased estimator for the “true” variance σ^2 .
- If we would divide by n instead of $n-1$, we would obtain an estimator with negative bias; it would systematically produce too-small estimates for σ^2 .

$$\begin{aligned}S_0, E[S_n^2] &= \frac{1}{n-1} E[\sum (x_i - \bar{x}_n)^2] \\ &= \frac{1}{\cancel{n-1}} \cdot \sigma^2(\cancel{n-1}) = \boxed{\sigma^2}\end{aligned}$$

Unbiased estimators for expectation and variance

Consider the following estimator for σ^2 :

$$V_n^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2.$$

Compute the bias $E[V_n^2] - \sigma^2$ for this estimator, where you can keep computations simple by realizing that $V_n^2 = (n-1)S_n^2/n$



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$$E[V_n^2] = E\left[\frac{n-1}{n} S_n^2\right] = \frac{n-1}{n} \cdot E[S_n^2] = \frac{n-1}{n} \sigma^2$$

$$\text{Bias: } E[V_n^2] - \sigma^2 = \frac{-\sigma^2}{n} < 0$$

Unbiased estimators

General Fact: Unbiasedness does not always carry over, i.e., if T is an unbiased estimator for a parameter θ , then $g(T)$ does not have to be an unbiased estimator for $g(\theta)$.

Exception: if $g(T) = aT + b$; & if T is unbiased for θ

$$\therefore E[aT+b] = aE[T] + b = a\theta + b$$

$\therefore aT+b$ is unbiased for $a\theta+b$ \Rightarrow



Unbiased estimators

Suppose our dataset is a realization of a random sample X_1, X_2, \dots, X_n from a uniform distribution on the interval $[-\theta, \theta]$, where θ is unknown.

- a. Show that $T = \frac{3}{n}(X_1^2 + X_2^2 + \dots + X_n^2)$ is an unbiased estimator for θ^2 .



Unbiased estimators

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We have to show $E[T] = \theta^2$;

$$\begin{aligned} \text{So, } E[T] &= \frac{3}{n} (E[X_1^2] + E[X_2^2] + \dots + E[X_n^2]) \\ &= \frac{3}{n} \left(\frac{\theta^2}{3} + \dots + \frac{\theta^2}{3} \right) \\ &= \frac{3}{n} \cdot n \cdot \frac{\theta^2}{3} = \boxed{\theta^2} \end{aligned}$$

$$\begin{aligned} E[X_i^2] &= \int_{-\theta}^{\theta} x^2 \cdot \frac{1}{2\theta} \cdot dx \\ &= \frac{1}{2\theta} \left[\frac{x^3}{3} \right]_{-\theta}^{\theta} \\ &= \frac{1}{2\theta} \cdot \left[\frac{\theta^3}{3} + \frac{\theta^3}{3} \right] = \frac{\theta^2}{3} \end{aligned}$$

Unbiased estimators

Suppose our dataset is a realization of a random sample X_1, X_2, \dots, X_n from a uniform distribution on the interval $[-\theta, \theta]$, where θ is unknown.

b. Is \sqrt{T} also an unbiased estimator for θ ? If not, argue whether it has positive or negative bias.

Skip



Unbiased estimators

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Unbiased estimators

Suppose the random variables X_1, X_2, \dots, X_n have the same expectation μ .

- a. Is $S = \frac{1}{2}X_1 + \frac{1}{3}X_2 + \frac{1}{6}X_3$ an unbiased estimator for μ ?



Unbiased estimators

Suppose the random variables X_1, X_2, \dots, X_n have the same expectation μ .

a. Is $S = \frac{1}{2}X_1 + \frac{1}{3}X_2 + \frac{1}{6}X_3$ an unbiased estimator for μ ? *Yes.*

$$E[S] = \frac{1}{2} E[X_1] + \frac{1}{3} E[X_2] + \frac{1}{6} E[X_3]$$

$$= \left(\frac{1}{2} + \frac{1}{3} + \frac{1}{6} \right) \mu = \underline{\underline{\mu}}$$

Unbiased estimators

Suppose the random variables X_1, X_2, \dots, X_n have the same expectation μ .

b. Under what conditions on constants a_1, a_2, \dots, a_n is

$$T = a_1X_1 + a_2X_2 + \dots + a_nX_n$$

an unbiased estimator for μ ?



Unbiased estimators

Suppose the random variables X_1, X_2, \dots, X_n have the same expectation μ .

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$$T = a_1X_1 + a_2X_2 + \dots + a_nX_n$$

an unbiased estimator for μ ?

$$\begin{aligned} E[T] &= E[a_1x_1 + a_2x_2 + \dots + a_nx_n] \\ &= a_1E[x_1] + a_2E[x_2] + \dots + a_nE[x_n] \\ &= a_1\mu + a_2\mu + \dots + a_n\mu \quad \text{--- ①} \end{aligned}$$

$$\therefore a_1 + a_2 + \dots + a_n = 1 \text{ to make ①} = \mu.$$

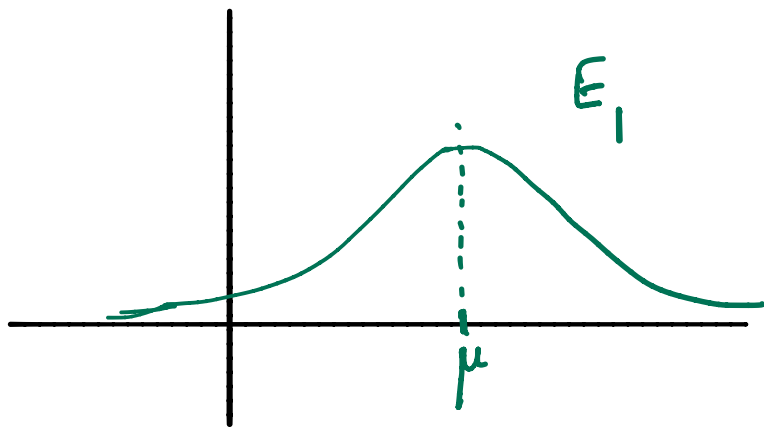
Efficiency and mean squared error

- If several unbiased estimators for the same parameter of interest exist, we need a criterion for comparison of these estimators.
- A natural criterion is some measure of spread of the estimators around the parameter of interest.
- For unbiased estimators we will use variance.
- For arbitrary estimators we introduce the notion of mean squared error (MSE), which combines variance and bias.



Variance of an estimator

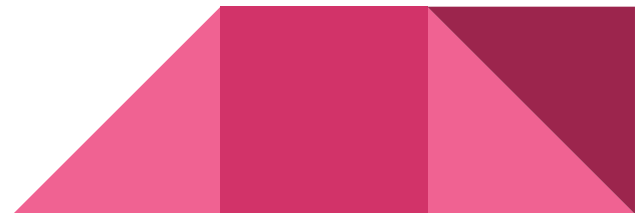
EFFICIENCY. Let T_1 and T_2 be two unbiased estimators for the same parameter θ . Then estimator T_2 is called *more efficient* than estimator T_1 if $\text{Var}(T_2) < \text{Var}(T_1)$, irrespective of the value of θ .



Mean squared error

DEFINITION. Let T be an estimator for a parameter θ . The *mean squared error* of T is the number $\text{MSE}(T) = \text{E}[(T - \theta)^2]$.

$$\begin{aligned}\text{MSE}(T) &= \text{E}[(T - \theta)^2] \\ &= \text{E}[(T - \text{E}[T] + \text{E}[T] - \theta)^2] \\ &= \text{E}[(T - \text{E}[T])^2] + 2\text{E}[T - \text{E}[T]](\text{E}[T] - \theta) + (\text{E}[T] - \theta)^2 \\ &= \text{Var}(T) + (\text{E}[T] - \theta)^2.\end{aligned}$$



Mean squared error

Given are two estimators S and T for a parameter θ . Furthermore it is known that $\text{Var}(S) = 40$ and $\text{Var}(T) = 4$.

a. Suppose that we know that $E[S] = \theta$ and $E[T] = \theta + 3$. Which estimator would you prefer, and why?



Mean squared error

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T

$$\text{MSE}(S) = \text{Var}(S) + (E[S] - \theta)^2 = 40$$

$$\text{MSE}(T) = \text{Var}(T) + (E[T] - \theta)^2 = 4 + 9 = 13$$

Mean squared error

Given are two estimators S and T for a parameter θ . Furthermore it is known that $\text{Var}(S) = 40$ and $\text{Var}(T) = 4$.

b. Suppose that we know that $E[S] = \theta$ and $E[T] = \theta + a$ for some positive number a . For each a , which estimator would you prefer, and why?



Mean squared error

Given are two estimators S and T for a parameter θ . Furthermore it is known that $\text{Var}(S) = 40$ and $\text{Var}(T) = 4$.

b. Suppose that we know that $E[S] = \theta$ and $E[T] = \theta + a$ for some positive number a . For each a , which estimator would you prefer, and why?

$$\text{MSE}(S) = 40.$$

$$\text{MSE}(T) = 4 + a^2$$

$$\therefore 4 + a^2 < 40$$

$\therefore a < 6$ we will prefer T
else S .

Mean squared error

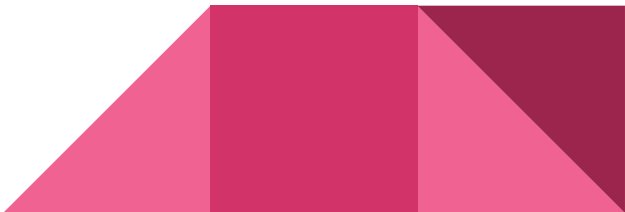
Suppose we have a random sample X_1, \dots, X_n from an $\text{Exp}(\lambda)$ distribution. Suppose we want to estimate the mean $1/\lambda$. Given an estimator:

$$T_1 = \bar{X}_n = \frac{1}{n}(X_1 + X_2 + \dots + X_n)$$

is an unbiased estimator of $1/\lambda$. Let M_n be the minimum of X_1, X_2, \dots, X_n . M_n has an $\text{Exp}(n\lambda)$ distribution. Given that:

$$T_2 = nM_n$$

is another unbiased estimator for $1/\lambda$. Which of the estimators T_1 and T_2 would you choose for estimating the mean $1/\lambda$? Substantiate your answer.



Imp: $\text{Var}(x_i) = \frac{1}{\lambda^2}$ if $x_i \sim \text{Exp}(\lambda)$

Now;

$$\begin{aligned}\text{var}(T_1) &= \text{var}\left(\frac{x_1}{n} + \frac{x_2}{n} + \dots + \frac{x_n}{n}\right) = \frac{1}{n^2} \left(\sum \text{Var}(x_i) \right) \\ &= \frac{1}{n^2} \left(\frac{1}{\lambda^2} + \frac{1}{\lambda^2} + \dots + \frac{1}{\lambda^2} \right) = \frac{1}{n^2} \cdot \frac{n}{\lambda^2} = \boxed{\frac{1}{n\lambda^2}}\end{aligned}$$

$$\text{Var}(M_n) = \frac{1}{\lambda^2 n^2} ; M_n \sim \text{Exp}(n\lambda)$$

$$\frac{1}{n\lambda^2} < \frac{1}{\lambda^2}$$

$$\therefore \text{var}(T_2) = \text{var}(nM_n) = n^2 \cdot \frac{1}{n^2 \lambda^2} = \boxed{\frac{1}{\lambda^2}}$$



Next: Linear Regression

