

Variable Selection in Nonlinear Interactions with the Group Lasso

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Supervisors



Yi Yang



Celia Greenwood

Post-doc Opportunities with Celia in Statistical Genetics



Robert Platt @robertwplatt · 13h

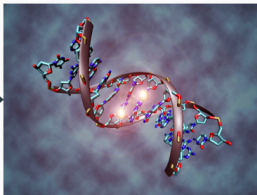
Setting up the McGill Biostat booth at #SSC2017WPG @McGillEBOH
pic.twitter.com/VSuSv7cBhZ

Motivation

Interactions with Environment



Environment
Gestational
Diabetes



Large Data
Child's epigenome
($p \approx 450k$)



Phenotype
Obesity measures

Variable Selection in Interaction Models

Classic Interaction Model

- $Y \rightarrow$ response
- $X_E \rightarrow$ environment
- $X_j \rightarrow$ predictors, $j = 1, \dots, p$

$$Y = \beta_0 \cdot \mathbf{1} + \sum_{j=1}^p \beta_j X_j + \beta_E X_E + \sum_{j=1}^p \alpha_j X_E X_j + \varepsilon$$

Heridity Property¹

$$Y = \beta_0 \cdot \mathbf{1} + \sum_{j=1}^p \beta_j X_j + \beta_E X_E + \sum_{j=1}^p \alpha_j X_E X_j + \varepsilon$$

Strong Hierarchy

$$\hat{\alpha}_j \neq 0 \quad \Rightarrow \quad \hat{\beta}_j \neq 0 \quad \text{and} \quad \hat{\beta}_E \neq 0$$

Weak Hierarchy

$$\hat{\alpha}_j \neq 0 \quad \Rightarrow \quad \hat{\beta}_j \neq 0 \quad \text{or} \quad \hat{\beta}_E \neq 0$$

¹Chipman, 1996, *Canadian Journal of Statistics*

Hierarchical Interactions: Current State of the Art

Type	Model	Software
Linear	CAP (Zhao et al. 2009, <i>Ann. Stat</i>)	X
	SHIM (Choi et al. 2009, <i>JASA</i>)	X
	hiernet (Bien et al. 2013, <i>Ann. Stat</i>)	hierNet(x, y)
	GRESH (She and Jiang 2014, <i>JASA</i>)	X
	FAMILY (Haris et al. 2014, <i>JCGS</i>)	FAMILY(x, z, y)
	glinternet (Lim and Hastie 2015, <i>JCGS</i>)	glinternet(x, y)
	RAMP (Hao et al. 2016, <i>JASA</i>)	RAMP(x, y)
Non-linear	VANISH (Radchenko and James 2010, <i>JASA</i>)	X
	funshim (Bhatnagar et al. 2017+)	funshim(x, e, y)

Lasso interaction model

$$Y = \beta_0 \cdot \mathbf{1} + \sum_{j=1}^p \beta_j X_j + \beta_E X_E + \sum_{j=1}^p \alpha_j X_E X_j + \varepsilon$$

$$\operatorname{argmin}_{\beta_0, \boldsymbol{\beta}, \boldsymbol{\alpha}} \mathcal{L}(Y; \Theta) + \lambda(\|\boldsymbol{\beta}\|_1 + \|\boldsymbol{\alpha}\|_1)$$

Reparametrization

$$\alpha_j = \gamma_j \beta_j \beta_E$$

Model

$$Y = \beta_0 \cdot \mathbf{1} + \sum_{j=1}^p \beta_j X_j + \beta_E X_E + \sum_{j=1}^p \gamma_j \beta_j \beta_E X_E X_j + \varepsilon$$

Objective Function

$$\operatorname{argmin}_{\beta_0, \beta, \gamma} \mathcal{L}(Y; \Theta) + \lambda_\beta \sum_{j=1}^p w_j |\beta_j| + \lambda_\gamma \sum_{j=1}^p w_{jE} |\gamma_{jE}|$$

²Choi et al. 2010, JASA

funshim: An Extension to Nonlinear Effects

Basis Expansion

$$f_j(X_j) = \sum_{\ell=1}^{p_j} \psi_{j\ell}(X_j) \beta_{j\ell}$$

$$f(X_1) = \underbrace{\begin{bmatrix} \psi_{11}(X_{11}) & \psi_{12}(X_{12}) & \cdots & \psi_{11}(X_{15}) \\ \vdots & \vdots & \cdots & \vdots \\ \vdots & \vdots & \cdots & \vdots \\ \psi_{11}(X_{i1}) & \psi_{12}(X_{i2}) & \cdots & \psi_{11}(X_{i5}) \\ \vdots & \vdots & \cdots & \vdots \\ \vdots & \vdots & \cdots & \vdots \\ \psi_{11}(X_{N1}) & \psi_{12}(X_{N2}) & \cdots & \psi_{11}(X_{N5}) \end{bmatrix}}_{\Psi_1} \times \underbrace{\begin{bmatrix} \beta_{11} \\ \beta_{12} \\ \beta_{13} \\ \beta_{14} \\ \beta_{15} \end{bmatrix}}_{\theta_1}^{5 \times 1}$$

- $\theta_j = (\beta_{j1}, \dots, \beta_{jp_j}) \in \mathbb{R}^{p_j}$
- $\alpha_j = (\alpha_{j1}, \dots, \alpha_{jp_j}) \in \mathbb{R}^{p_j}$
- $\Psi_j \rightarrow n \times p_j$ matrix of evaluations of the $\psi_{j\ell}$

Model

$$Y = \beta_0 \cdot \mathbf{1} + \sum_{j=1}^p \Psi_j \theta_j + \beta_E X_E + \sum_{j=1}^p X_E \Psi_j \alpha_j + \varepsilon$$

Reparametrization

$$\alpha_j = \gamma_j \beta_E \theta_j$$

Model

$$Y = \beta_0 \cdot \mathbf{1} + \sum_{j=1}^p \Psi_j \theta_j + \beta_E X_E + \sum_{j=1}^p \gamma_j \beta_E X_E \Psi_j \theta_j + \varepsilon$$

Objective Function

$$\operatorname{argmin}_{\beta_E, \boldsymbol{\theta}, \boldsymbol{\gamma}} \mathcal{L}(Y; \boldsymbol{\Theta}) + \lambda_\beta \left(w_E |\beta_E| + \sum_{j=1}^p w_j \|\theta_j\|_2 \right) + \lambda_\gamma \sum_{j=1}^p w_{jE} |\gamma_j|$$

Algorithm

Block Relaxation (De Leeuw, 1994)

Algorithm 1: Block Relaxation Algorithm

Set the iteration counter $k \leftarrow 0$, initial values for the parameter vector $\Theta^{(0)}$;

for each pair $(\lambda_\beta, \lambda_\gamma)$ **do**

repeat

$$\gamma^{(k+1)} \leftarrow \operatorname{argmin}_{\gamma} Q_{\lambda_\beta, \lambda_\gamma}(\gamma, \beta_E^{(k)}, \theta^{(k)})$$

$$\theta^{(k+1)} \leftarrow \operatorname{argmin}_{\theta} Q_{\lambda_\beta, \lambda_\gamma}(\theta, \beta_E^{(k)}, \gamma^{(k+1)})$$

$$\beta_E^{(k+1)} \leftarrow \operatorname{argmin}_{\beta_E} Q_{\lambda_\beta, \lambda_\gamma}(\theta^{(k+1)}, \beta_E, \gamma^{(k+1)})$$

$k \leftarrow k + 1$

until convergence criterion is satisfied;

end

Objective Function

$$\operatorname{argmin}_{\beta_E, \theta, \gamma} \mathcal{L}(Y; \Theta) + \lambda_\beta \left(w_E |\beta_E| + \sum_{j=1}^p w_j \|\theta_j\|_2 \right) + \lambda_\gamma \sum_{j=1}^p w_{jE} |\gamma_j|$$

Lasso problem (`glmnet`, Friedman, Hastie & Tibshirani 2010)

$$\operatorname{argmin}_{\gamma} \mathcal{L}(Y; \Theta) + \lambda_\beta \left(w_E |\beta_E| + \sum_{j=1}^p w_j \|\theta_j\|_2 \right) + \lambda_\gamma \sum_{j=1}^p w_{jE} |\gamma_j|$$

³<https://github.com/sahirbhatnagar/funshim>

Objective Function

$$\operatorname{argmin}_{\beta_E, \boldsymbol{\theta}, \gamma} \mathcal{L}(Y; \Theta) + \lambda_\beta \left(w_E |\beta_E| + \sum_{j=1}^p w_j \|\theta_j\|_2 \right) + \lambda_\gamma \sum_{j=1}^p w_{jE} |\gamma_j|$$

Group Lasso problem (gglasso, Yang and Zou 2015)

$$\operatorname{argmin}_{\beta_E, \boldsymbol{\theta}} \mathcal{L}(Y; \Theta) + \lambda_\beta \left(w_E |\beta_E| + \sum_{j=1}^p w_j \|\theta_j\|_2 \right) + \lambda_\gamma \sum_{j=1}^p w_{jE} |\gamma_j|$$

⁴<https://github.com/sahirbhatnagar/funshim>

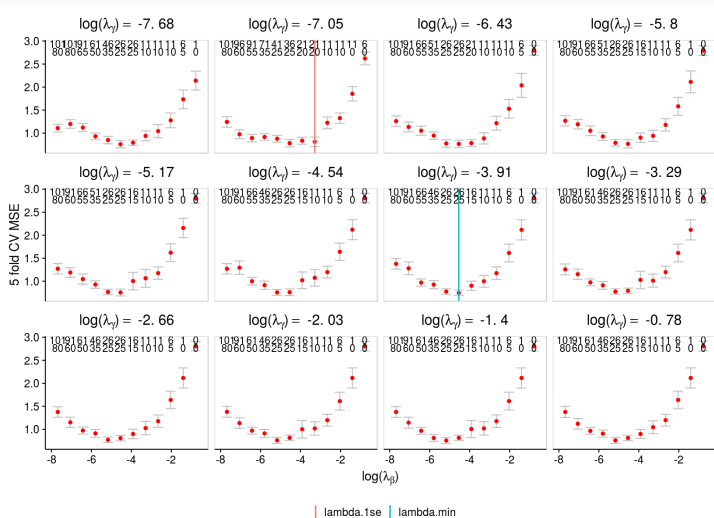
Simulations

Scenario 1

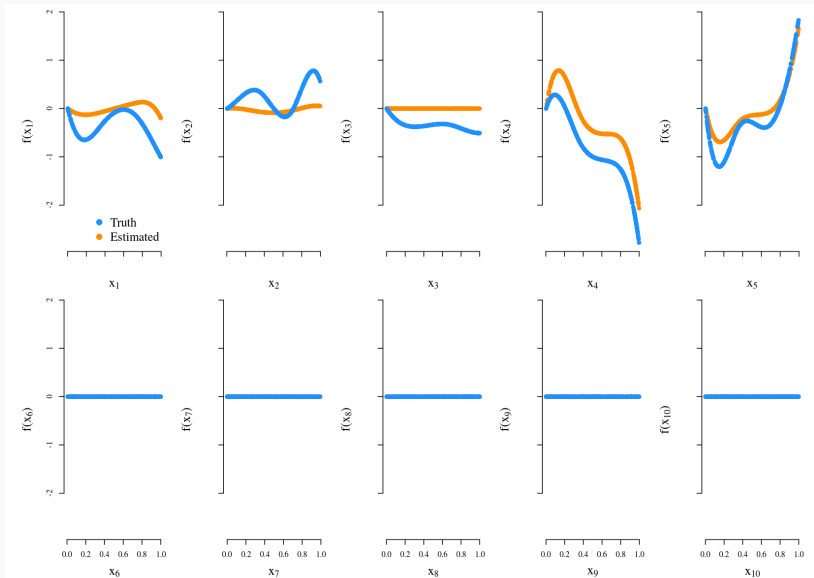
- $Y = \sum_{j=1}^5 f(X_j) + X_E + E(f(X_1) + f(X_2))$
- $f(\cdot) \rightarrow$ B-splines with 5 df
- $\theta_j \sim \mathcal{N}(0, 1)$
- $N = 400, p = 50$
- $50 \times 5 \times 2 + 1 = 501$ parameters to estimate

Scenario 1: Cross-validation results

`funshim::plot(cvfit)`

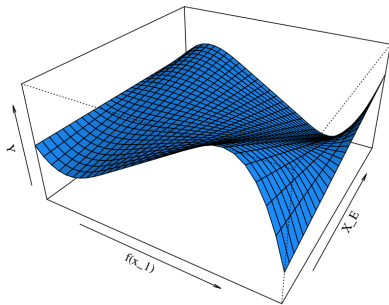


Scenario 1: Main Effects

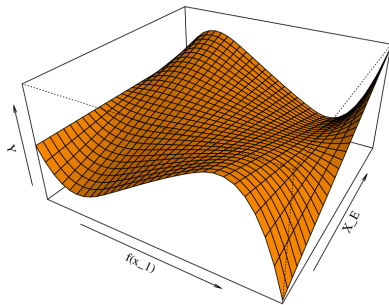


Scenario 1: Interaction Effects

Truth



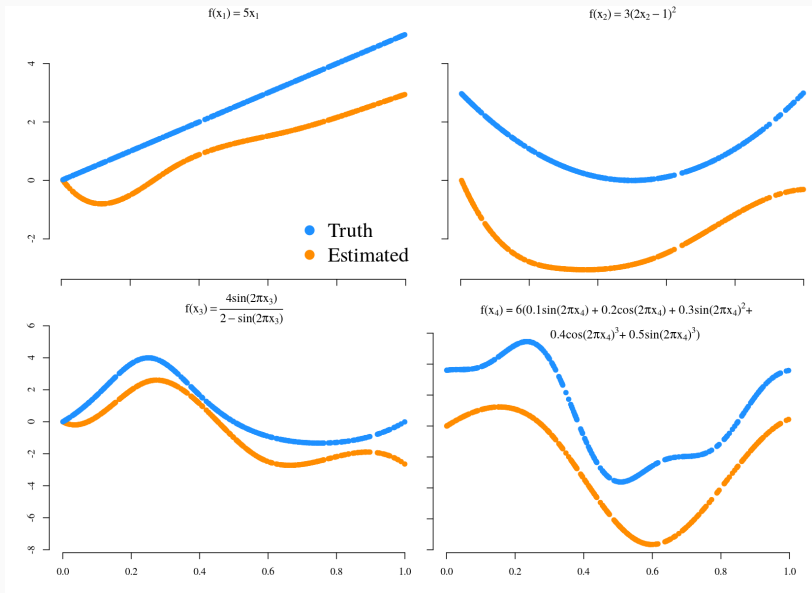
Estimated $X_E * f(X_1)$



Scenario 2

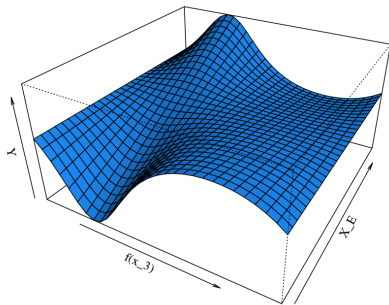
- $Y = \sum_{j=1}^4 f(X_j) + X_E + E(f(X_3) + f(X_4))$
- $f(X_1) \rightarrow \text{linear}$
- $f(X_2) \rightarrow \text{quadratic}$
- $f(X_3) \rightarrow \text{sinusoidal}$
- $f(X_4) \rightarrow \text{complicated sinusoidal}$

Scenario 2: Main Effects

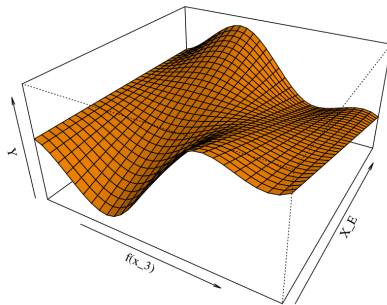


Scenario 2: Interaction Effects

Truth

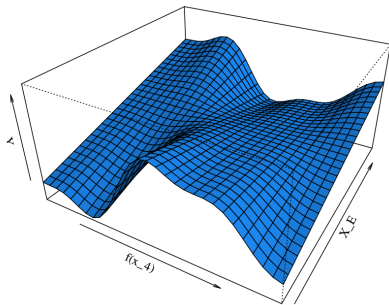


Estimated $X_E * f(X_3)$

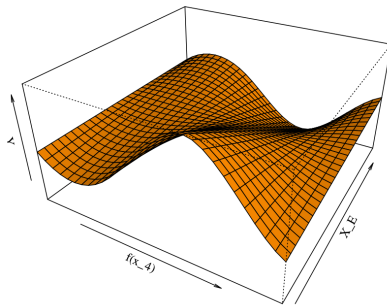


Scenario 2: Interaction Effects

Truth



Estimated $X_E * f(X_4)$



Discussion

Strengths and Limitations

Strengths

- Environment interactions with strong heredity property in $p \gg N$
- **funshim** allows for flexible modeling of input variables (imaging, gene expression, DNA methylation)
- R package provided with tuning parameter selection routines

Limitations

- Can only handle $E \cdot f(X)$ or $f(E) \cdot X$
- Does not allow for $f(X_1, E)$ or $f(X_1, X_2)$
- Current implementation is slow due to cross validation for 2 tuning parameters

- Are two tuning parameters really necessary ?

$$\lambda \left\{ (1 - \alpha) \left[w_E |\beta_E| + \sum_{j=1}^p w_j \|\theta_j\|_2 \right] + \alpha \sum_{j=1}^p w_{jE} |\gamma_j| \right\}$$

- Weak heredity property $\rightarrow \alpha_j = \gamma_j(|\beta_j| + |\beta_E|)$
- Numerical convergence and KKT checks
- Information Criterion instead of CV for tuning parameters
- Extension to GLM
- User defined interactions
- Non-parametric screening prior to model fitting
- Real data analysis

Acknowledgements



References

- Radchenko, P., & James, G. M. (2010). Variable selection using adaptive nonlinear interaction structures in high dimensions. *Journal of the American Statistical Association*, 105(492), 1541-1553.
- Choi, N. H., Li, W., & Zhu, J. (2010). Variable selection with the strong heredity constraint and its oracle property. *Journal of the American Statistical Association*, 105(489), 354-364.
- Chipman, H. (1996). Bayesian variable selection with related predictors. *Canadian Journal of Statistics*, 24(1), 17-36.
- Friedman, J., Hastie, T., & Tibshirani, R. (2010). Regularization paths for generalized linear models via coordinate descent. *Journal of statistical software*, 33(1)
- Yang, Y., & Zou, H. (2015). A fast unified algorithm for solving group-lasso penalize learning problems. *Statistics and Computing*, 25(6), 1129-1141
- De Leeuw, J. (1994). Block-relaxation algorithms in statistics. In *Information systems and data analysis* (pp. 308-324). Springer Berlin Heidelberg.

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