

# Computational Methods For Case-Cohort Studies

Sahir Rai Bhatnagar

Queen's University

November 27, 2012

# Cohort studies

- All participants provide a wide range of information at time of recruitment e.g. detailed dietary questionnaires and blood and urine samples
- Because of large numbers and cost of analysing the biological specimens or genotyping, these resources are often not analysed in detail at the time but are stored for future use
- This design is expensive, inefficient for rare outcomes, long follow-up period needed, large sample size needed

# Cohort studies

- All participants provide a wide range of information at time of recruitment e.g. detailed dietary questionnaires and blood and urine samples
- Because of large numbers and cost of analysing the biological specimens or genotyping, these resources are often not analysed in detail at the time but are stored for future use
- This design is expensive, inefficient for rare outcomes, long follow-up period needed, large sample size needed

## Case-Cohort: A more efficient design

- A **random** sample of participants are selected from full cohort at baseline
- Detailed exposure information (covariates) can then be retrieved for

## Case-Cohort: A more efficient design

- A **random** sample of participants are selected from full cohort at baseline
- Detailed exposure information (covariates) can then be retrieved for
  - this subcohort

# Case-Cohort: A more efficient design

- A **random** sample of participants are selected from full cohort at baseline
- Detailed exposure information (covariates) can then be retrieved for
  - this subcohort
  - everyone in the full cohort who develop the disease of interest

## Case-Cohort: A more efficient design

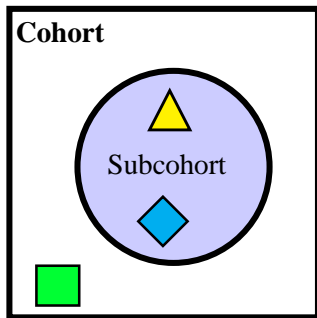
- A **random** sample of participants are selected from full cohort at baseline
- Detailed exposure information (covariates) can then be retrieved for
  - this subcohort
  - everyone in the full cohort who develop the disease of interest
- Key feature: inclusion of all cases that occur in the cohort

## Case-Cohort: A more efficient design

- A **random** sample of participants are selected from full cohort at baseline
- Detailed exposure information (covariates) can then be retrieved for
  - this subcohort
  - everyone in the full cohort who develop the disease of interest
- Key feature: inclusion of all cases that occur in the cohort



# Case-Cohort Design



Subcohort censored



Subcohort failure



Non-subcohort failure

# Objective

## Purpose of this presentation

- 1 Explain and promote the case-cohort design
- 2 Show that it's not as difficult as the literature says to compute accurate estimates

# Objective

## Purpose of this presentation

- 1 Explain and promote the case-cohort design
- 2 Show that it's not as difficult as the literature says to compute accurate estimates

# An Example

## Description of the analysed dataset

- Simple and age at first exposure stratified case-cohort samples drawn from a cohort of 1741 female patients who were discharged from two tuberculosis sanatoria in Massachusetts between 1930 and 1956 to investigate **breast cancer risk** and **radiation exposure** due to fluoroscopy

# An Example

## Description of the analysed dataset

- Simple and age at first exposure stratified case-cohort samples drawn from a cohort of 1741 female patients who were discharged from two tuberculosis sanatoria in Massachusetts between 1930 and 1956 to investigate **breast cancer risk** and **radiation exposure** due to fluoroscopy
- Radiation doses were estimated for those women who received radiation exposure to the chest from the X-ray fluoroscopy lung examination

# An Example

## Description of the analysed dataset

- Simple and age at first exposure stratified case-cohort samples drawn from a cohort of 1741 female patients who were discharged from two tuberculosis sanatoria in Massachusetts between 1930 and 1956 to investigate **breast cancer risk** and **radiation exposure** due to fluoroscopy
- Radiation doses were estimated for those women who received radiation exposure to the chest from the X-ray fluoroscopy lung examination
- The remaining women received treatments that did not require fluoroscopic monitoring and were radiation unexposed

# An Example

## Description of the analysed dataset

- Simple and age at first exposure stratified case-cohort samples drawn from a cohort of 1741 female patients who were discharged from two tuberculosis sanatoria in Massachusetts between 1930 and 1956 to investigate **breast cancer risk** and **radiation exposure** due to fluoroscopy
- Radiation doses were estimated for those women who received radiation exposure to the chest from the X-ray fluoroscopy lung examination
- The remaining women received treatments that did not require fluoroscopic monitoring and were radiation unexposed
- 75 breast cancer cases were identified with 54 exposed and 21 unexposed

# An Example

## Description of the analysed dataset

- Simple and age at first exposure stratified case-cohort samples drawn from a cohort of 1741 female patients who were discharged from two tuberculosis sanatoria in Massachusetts between 1930 and 1956 to investigate **breast cancer risk** and **radiation exposure** due to fluoroscopy
- Radiation doses were estimated for those women who received radiation exposure to the chest from the X-ray fluoroscopy lung examination
- The remaining women received treatments that did not require fluoroscopic monitoring and were radiation unexposed
- 75 breast cancer cases were identified with 54 exposed and 21 unexposed
- 100 subjects were randomly sampled without replacement



# An Example

## Description of the analysed dataset

- Simple and age at first exposure stratified case-cohort samples drawn from a cohort of 1741 female patients who were discharged from two tuberculosis sanatoria in Massachusetts between 1930 and 1956 to investigate **breast cancer risk** and **radiation exposure** due to fluoroscopy
- Radiation doses were estimated for those women who received radiation exposure to the chest from the X-ray fluoroscopy lung examination
- The remaining women received treatments that did not require fluoroscopic monitoring and were radiation unexposed
- 75 breast cancer cases were identified with 54 exposed and 21 unexposed
- 100 subjects were randomly sampled without replacement

# Advantages

- Exposure precedes outcome, while smaller scale reduces cost and effort
- In outbreak situations, multiple outcomes can be studied using only one sample of controls

# Advantages

- Exposure precedes outcome, while smaller scale reduces cost and effort
- In outbreak situations, multiple outcomes can be studied using only one sample of controls

# Advantages

- Exposure precedes outcome, while smaller scale reduces cost and effort
- In outbreak situations, multiple outcomes can be studied using only one sample of controls

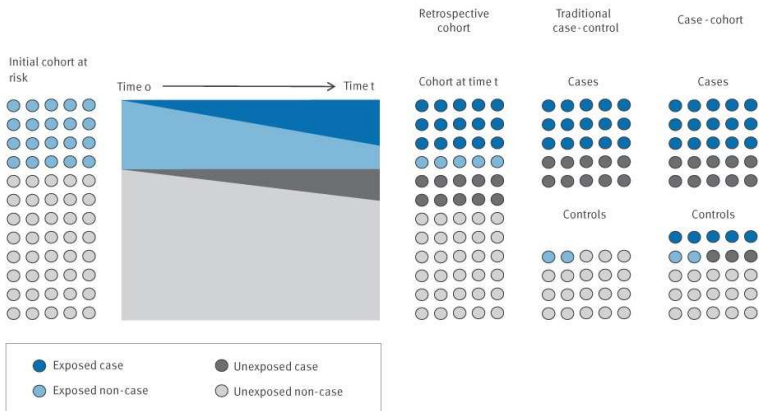
# Challenges

- Theoretically computationally difficult to compute variance estimates
- Because of such biased sampling with regard to case-status, risk estimation using the ordinary partial likelihood is not appropriate

# Challenges

- Theoretically computationally difficult to compute variance estimates
- Because of such biased sampling with regard to case-status, risk estimation using the ordinary partial likelihood is not appropriate

# Comparing three study designs



# Cox Proportional Hazards Model

First lets consider a relative risk regression model (*Cox, 1972*)

Cox PH Model

$$\lambda\{t; Z(u), 0 \leq u \leq t\} = \lambda_0(t) r\{X(t)\beta\}$$



# Cox Proportional Hazards Model

First lets consider a relative risk regression model (*Cox, 1972*)

## Cox PH Model

$$\lambda\{t; Z(u), 0 \leq u \leq t\} = \lambda_0(t)r\{X(t)\beta\}$$

- $\lambda(t)$ : failure rate of interest at time  $t$  for a subject

# Cox Proportional Hazards Model

First lets consider a relative risk regression model (*Cox, 1972*)

## Cox PH Model

$$\lambda\{t; Z(u), 0 \leq u \leq t\} = \lambda_0(t) r\{X(t)\beta\}$$

- $\lambda(t)$ : failure rate of interest at time  $t$  for a subject
- $\{Z(u); 0 \leq u < t\}$ : preceding covariate history

# Cox Proportional Hazards Model

First lets consider a relative risk regression model (*Cox, 1972*)

## Cox PH Model

$$\lambda\{t; Z(u), 0 \leq u \leq t\} = \lambda_0(t)r\{X(t)\beta\}$$

- $\lambda(t)$ : failure rate of interest at time  $t$  for a subject
- $\{Z(u); 0 \leq u < t\}$ : preceding covariate history
- $r(x)$ : is a fixed function with  $r(0) = 1$  e.g.  $r(x) = \exp\{x\}$

# Cox Proportional Hazards Model

First lets consider a relative risk regression model (*Cox, 1972*)

## Cox PH Model

$$\lambda\{t; Z(u), 0 \leq u \leq t\} = \lambda_0(t) r\{X(t)\beta\}$$

- $\lambda(t)$ : failure rate of interest at time  $t$  for a subject
- $\{Z(u); 0 \leq u < t\}$ : preceding covariate history
- $r(x)$ : is a fixed function with  $r(0) = 1$  e.g.  $r(x) = \exp\{x\}$
- $X(t)$ : row  $p$ -vector consisting of functions of  $Z(u)$

# Cox Proportional Hazards Model

First lets consider a relative risk regression model (Cox, 1972)

## Cox PH Model

$$\lambda\{t; Z(u), 0 \leq u \leq t\} = \lambda_0(t) r\{X(t)\beta\}$$

- $\lambda(t)$ : failure rate of interest at time  $t$  for a subject
- $\{Z(u); 0 \leq u < t\}$ : preceding covariate history
- $r(x)$ : is a fixed function with  $r(0) = 1$  e.g.  $r(x) = \exp\{x\}$
- $X(t)$ : row  $p$ -vector consisting of functions of  $Z(u)$
- $\beta$ : column  $p$ -vector of regression parameters to be estimated

# Cox Proportional Hazards Model

First lets consider a relative risk regression model (Cox, 1972)

## Cox PH Model

$$\lambda\{t; Z(u), 0 \leq u \leq t\} = \lambda_0(t) r\{X(t)\beta\}$$

- $\lambda(t)$ : failure rate of interest at time  $t$  for a subject
- $\{Z(u); 0 \leq u < t\}$ : preceding covariate history
- $r(x)$ : is a fixed function with  $r(0) = 1$  e.g.  $r(x) = \exp\{x\}$
- $X(t)$ : row  $p$ -vector consisting of functions of  $Z(u)$
- $\beta$ : column  $p$ -vector of regression parameters to be estimated
- $\lambda_0(t)$ : baseline hazard function

# Cox Proportional Hazards Model

First lets consider a relative risk regression model (*Cox, 1972*)

## Cox PH Model

$$\lambda\{t; Z(u), 0 \leq u \leq t\} = \lambda_0(t) r\{X(t)\beta\}$$

- $\lambda(t)$ : failure rate of interest at time  $t$  for a subject
- $\{Z(u); 0 \leq u < t\}$ : preceding covariate history
- $r(x)$ : is a fixed function with  $r(0) = 1$  e.g.  $r(x) = \exp\{x\}$
- $X(t)$ : row  $p$ -vector consisting of functions of  $Z(u)$
- $\beta$ : column  $p$ -vector of regression parameters to be estimated
- $\lambda_0(t)$ : baseline hazard function

# Exact and approximate pseudolikelihood estimators

- Indicator whether subject is at risk at time  $t$  ○

$$\tilde{\mathcal{L}}(\beta) = \prod_{i=1}^n \prod_t \left[ \frac{\exp\{\beta Z_i(t)\}}{\sum_{k \in \tilde{\mathcal{R}}_i(t)} \exp\{\beta Z_k(t)\}} \right] dN_i(t) \quad (1)$$

- The contribution of a failure by subject  $i$  at time  $t$  ○



# Exact and approximate pseudolikelihood estimators

- Indicator whether subject is at risk at time  $t$  ●

$$\tilde{\mathcal{L}}(\beta) = \prod_{i=1}^n \prod_t \left[ \frac{\exp \{ \beta Z_i(t) \}}{\sum_{k \in \tilde{\mathcal{R}}_i(t)} \exp \{ \beta Z_k(t) \}} \right] dN_i(t) \quad (1)$$

- The contribution of a failure by subject  $i$  at time  $t$  ●
- Sum of all subcohort nonfailures at risk at time  $t$  including the failure by subject  $i$  ●

# Exact and approximate pseudolikelihood estimators

- Indicator whether subject is at risk at time  $t$

$$\tilde{\mathcal{L}}(\beta) = \prod_{i=1}^n \prod_t \left[ \frac{\exp \{ \beta Z_i(t) \}}{\sum_{k \in \tilde{\mathcal{R}}_i(t)} \exp \{ \beta Z_k(t) \}} \right] dN_i(t) \quad (1)$$

- The contribution of a failure by subject  $i$  at time  $t$
- Sum of all subcohort nonfailures at risk at time  $t$  including the failure by subject  $i$
- Exact:  $\tilde{\mathcal{R}}_i(t) = (C \cup \{i\}) \cap \mathcal{R}(t)$
- Approximate:  $\tilde{\mathcal{R}}_i(t) = C \cap \mathcal{R}(t)$ , where  $C$  is the subcohort

# Exact and approximate pseudolikelihood estimators

- Indicator whether subject is at risk at time  $t$

$$\tilde{\mathcal{L}}(\beta) = \prod_{i=1}^n \prod_t \left[ \frac{\exp \{ \beta Z_i(t) \}}{\sum_{k \in \tilde{\mathcal{R}}_i(t)} \exp \{ \beta Z_k(t) \}} \right] dN_i(t) \quad (1)$$

- The contribution of a failure by subject  $i$  at time  $t$
- Sum of all subcohort nonfailures at risk at time  $t$  including the failure by subject  $i$
- Exact:  $\tilde{\mathcal{R}}_i(t) = (C \cup \{i\}) \cap \mathcal{R}(t)$
- Approximate:  $\tilde{\mathcal{R}}_i(t) = C \cap \mathcal{R}(t)$ , where  $C$  is the subcohort

# Exact and approximate pseudolikelihood estimators

- The unique sampling approach i.e. over selecting cases, leads to a **pseudolikelihood** rather than the usual partial likelihood
- Analysis must adjust for bias introduced in the distributions of covariates used in calculating the denominator of the pseudolikelihood

# Exact and approximate pseudolikelihood estimators

- The unique sampling approach i.e. over selecting cases, leads to a **pseudolikelihood** rather than the usual partial likelihood
- Analysis must adjust for bias introduced in the distributions of covariates used in calculating the denominator of the pseudolikelihood
- Bias incurred by including cases outside the subcohort is corrected by not allowing those cases to contribute to risk sets other than their own

# Exact and approximate pseudolikelihood estimators

- The unique sampling approach i.e. over selecting cases, leads to a **pseudolikelihood** rather than the usual partial likelihood
- Analysis must adjust for bias introduced in the distributions of covariates used in calculating the denominator of the pseudolikelihood
- Bias incurred by including cases outside the subcohort is corrected by not allowing those cases to contribute to risk sets other than their own
- We will focus our attention the **exact** approach rather than the approximate

# Exact and approximate pseudolikelihood estimators

- The unique sampling approach i.e. over selecting cases, leads to a **pseudolikelihood** rather than the usual partial likelihood
- Analysis must adjust for bias introduced in the distributions of covariates used in calculating the denominator of the pseudolikelihood
- Bias incurred by including cases outside the subcohort is corrected by not allowing those cases to contribute to risk sets other than their own
- We will focus our attention the **exact** approach rather than the approximate

# Approximate Variance of $\hat{\beta}$

Therneau and Li (1999) solved the variance estimation problem proposing the following approximation

$$\hat{I}^{-1} + \frac{m(n-m)}{n} \text{Cov } D_C \quad (2)$$

- $\hat{I}^{-1}$ : estimated covariance matrix of the parameter estimates  
(Inverse of Fisher Information matrix)



# Approximate Variance of $\hat{\beta}$

Therneau and Li (1999) solved the variance estimation problem proposing the following approximation

$$\hat{I}^{-1} + \frac{m(n-m)}{n} \text{Cov } D_C \quad (2)$$

- $\hat{I}^{-1}$ : estimated covariance matrix of the parameter estimates  
(Inverse of Fisher Information matrix)
- $n$ : size of full cohort

# Approximate Variance of $\hat{\beta}$

Therneau and Li (1999) solved the variance estimation problem proposing the following approximation

$$\hat{I}^{-1} + \frac{m(n-m)}{n} \text{Cov } D_C \quad (2)$$

- $\hat{I}^{-1}$ : estimated covariance matrix of the parameter estimates (Inverse of Fisher Information matrix)
- **n**: size of full cohort
- **m**: size of subcohort

# Approximate Variance of $\hat{\beta}$

Therneau and Li (1999) solved the variance estimation problem proposing the following approximation

$$\hat{I}^{-1} + \frac{m(n-m)}{n} \text{Cov } D_C \quad (2)$$

- $\hat{I}^{-1}$ : estimated covariance matrix of the parameter estimates (Inverse of Fisher Information matrix)
- **n**: size of full cohort
- **m**: size of subcohort
- $\text{Cov } D_C$ : empirical covariance matrix of *dfbeta* residuals from subcohort members

# Approximate Variance of $\hat{\beta}$

Therneau and Li (1999) solved the variance estimation problem proposing the following approximation

$$\hat{I}^{-1} + \frac{m(n-m)}{n} \text{Cov } D_C \quad (2)$$

- $\hat{I}^{-1}$ : estimated covariance matrix of the parameter estimates (Inverse of Fisher Information matrix)
- **n**: size of full cohort
- **m**: size of subcohort
- $\text{Cov } D_C$ : empirical covariance matrix of *dfbeta* residuals from subcohort members

## Dfbeta residuals

Are the approximate changes in the parameter estimates  $(\hat{\beta} - \hat{\beta}_{(j)})$  when the  $j^{th}$  observation is omitted. These variables are a weighted transform of the score residual variables and are useful in assessing local influence and in computing approximate and robust variance estimates.

# Procedure for creating analytic dataset

## Steps

- 1 Each subcohort non-failure contributes one line of data to the analytic data set as censored observations

# Procedure for creating analytic dataset

## Steps

- 1 Each subcohort non-failure contributes one line of data to the analytic data set as censored observations
- 2 A non-subcohort failure contributes no information prior to the failure time so one line of data is contributed to the analytic data set as a failure but only at the failure time

# Procedure for creating analytic dataset

## Steps

- 1 Each subcohort non-failure contributes one line of data to the analytic data set as censored observations
- 2 A non-subcohort failure contributes no information prior to the failure time so one line of data is contributed to the analytic data set as a failure but only at the failure time
- 3 A subcohort failure contributes two lines to the analytic data set:



# Procedure for creating analytic dataset

## Steps

- 1** Each subcohort non-failure contributes one line of data to the analytic data set as censored observations
- 2** A non-subcohort failure contributes no information prior to the failure time so one line of data is contributed to the analytic data set as a failure but only at the failure time
- 3** A subcohort failure contributes two lines to the analytic data set:
  - one line as a censored observation prior to the failure time

# Procedure for creating analytic dataset

## Steps

- 1 Each subcohort non-failure contributes one line of data to the analytic data set as censored observations
- 2 A non-subcohort failure contributes no information prior to the failure time so one line of data is contributed to the analytic data set as a failure but only at the failure time
- 3 A subcohort failure contributes two lines to the analytic data set:
  - one line as a censored observation prior to the failure time
  - and one line as a failure at the failure time

# Procedure for creating analytic dataset

## Steps

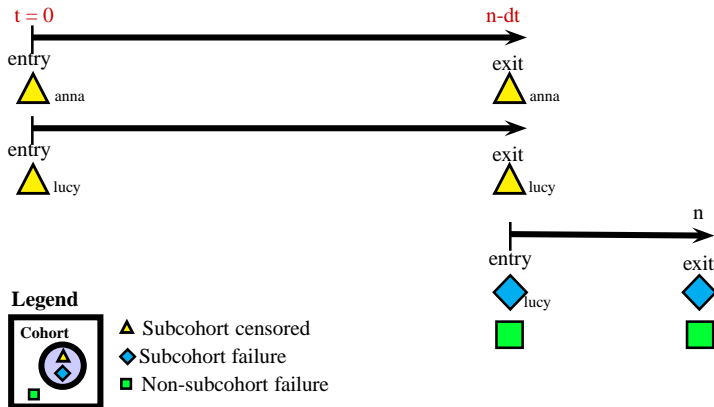
- 1 Each subcohort non-failure contributes one line of data to the analytic data set as censored observations
- 2 A non-subcohort failure contributes no information prior to the failure time so one line of data is contributed to the analytic data set as a failure but only at the failure time
- 3 A subcohort failure contributes two lines to the analytic data set:
  - one line as a censored observation prior to the failure time
  - and one line as a failure at the failure time
- 4 To create a time just before the exit time, an amount less than the precision of exit times given in the data is subtracted off from the actual failure time

# Procedure for creating analytic dataset

## Steps

- 1 Each subcohort non-failure contributes one line of data to the analytic data set as censored observations
- 2 A non-subcohort failure contributes no information prior to the failure time so one line of data is contributed to the analytic data set as a failure but only at the failure time
- 3 A subcohort failure contributes two lines to the analytic data set:
  - one line as a censored observation prior to the failure time
  - and one line as a failure at the failure time
- 4 To create a time just before the exit time, an amount less than the precision of exit times given in the data is subtracted off from the actual failure time

# Graphic of how to create analytic dataset



# Original Case Cohort dataset

Basic case-cohort data

Subject ID	Dose in rad	Age at exit (in years)	Age at entry (in years)	0-cens,1-subc fail 2-non-subc fail	1-249 rad	250+ rad	age at first exposure group <sup>a</sup>
2866	0.4525	71.269	34.0014	0	1	0	3
2787	0.00984	69.0294	31.7454	0	1	0	4
2702	0.05486	47.5948	36.5065	0	1	0	3
34	0	55.4387	14.9377	1	0	0	1
3064	0.12788	35.4825	25.6838	0	1	0	3
2766	1.62311	64.3559	30.5161	0	1	0	3
2344	1.0624	69.692	25.4127	0	1	0	3
⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮
2698	0	42.3682	36.2026	2	0	0	4
2577	1.00338	50.9979	26.412	2	1	0	3
2348	1.30725	42.1246	24.1259	2	1	0	3
3106	0	55.2635	27.2635	2	0	0	3
2687	0	47.7563	23.2553	2	0	0	3
3018	1.6723	50.0014	38.8337	2	1	0	4

<sup>a</sup> 1 :< 15, 2 : 15 – 19, 3 : 20 – 29, 4 : 30+

# Comparison

## Original vs. Analytic dataset

Subject ID	Dose in rad	Age at exit (in years)	Age at entry (in years)	0-cens,1-subc fail 2-non-subc fail	1-249 rad	250+ rad	age at first exposure group	an_entry	an_exit	an_jnd
34	0	55.4387	14.9377	1	0	0	1			
2344	1.0624	69.692	25.4127	0	1	0	3			
2687	0	47.7563	23.2553	2	0	0	3			
34	0	55.4387	14.9377	1	0	0	1	14.9377	55.4386	0
34	0	55.4387	14.9377	1	0	0	1	55.4386	55.4387	1
2344	1.0624	69.692	25.4127	0	1	0	3	25.4127	69.6919	0
2687	0	47.7563	23.2553	2	0	0	3	47.7562	47.7563	1

# Comparison

## Original vs. Analytic dataset

Subject ID	Dose in rad	Age at exit (in years)	Age at entry (in years)	0-cens,1-subc fail 2-non-subc fail	1-249 rad	250+ rad	age at first exposure group	an_entry	an_exit	an_jnd
34	0	55.4387	14.9377	1	0	0	1			
2344	1.0624	69.692	25.4127	0	1	0	3			
2687	0	47.7563	23.2553	2	0	0	3			
34	0	55.4387	14.9377	1	0	0	1	14.9377	55.4386	0
34	0	55.4387	14.9377	1	0	0	1	55.4386	55.4387	1
2344	1.0624	69.692	25.4127	0	1	0	3	25.4127	69.6919	0
2687	0	47.7563	23.2553	2	0	0	3	47.7562	47.7563	1



# SAS Code

```
proc phreg data=analytic;
  model an_exit*an_ind(0) = dcat1 dcat2 /
    entry=an_entry covb;
  output out=dfbetas dfbeta= dfb_dcat1 dfb_dcat2;
  id id;
run;

proc corr data=dfbetas cov;
  var dfb_dcat1 dfb_dcat2;
  where an_ind eq 0;
run;
```

- **covb**: outputs the inverse information matrix  $\hat{I}^{-1}$
- **cov**: outputs the covariance matrix of *dfbeta* residuals from subcohort members (WHERE an\_ind = 0)

# SAS Code

```
proc phreg data=analytic;
  model an_exit*an_ind(0) = dcat1 dcat2 /
    entry=an_entry covb;
  output out=dfbetas dfbeta= dfb_dcat1 dfb_dcat2;
  id id;
run;

proc corr data=dfbetas cov;
  var dfb_dcat1 dfb_dcat2;
  where an_ind eq 0;
run;
```

- **covb**: outputs the inverse information matrix  $\hat{l}^{-1}$
- **cov**: outputs the covariance matrix of *dfbeta* residuals from subcohort members (WHERE an\_ind = 0)

# Exact pseudolikelihood and Asymptotic variance

## Analysis of Maximum Likelihood Estimates

Parameter	DF	Parameter Estimate	Standard Error	$\chi^2$	$\text{Pr} > \chi^2$	Hazard Ratio	Label
dcat1	1	0.6572	0.26117	6.332	0.0119	1.929	1-249 rad
dcat2	1	1.55325	0.50118	9.6051	0.0019	4.727	250+ rad

## Estimated Covariance Matrix ( $\times 10^{-2}$ )

Parameter		dcat1	dcat2
dcat1	1-249 rad	6.821	4.743
dcat2	250+ rad	4.743	25.118

## Estimated Covariance Matrix of the dfbeta residuals ( $\times 10^{-4}$ )

Parameter		dfb_dcat1	dfb_dcat2
dfb_dcat1	difference in the parameter for dcat1	5.487	2.998
dfb_dcat2	difference in the parameter for dcat2	2.998	47.878

# Time-dependent covariates

- The partial likelihood of Cox also allows time-dependent explanatory variables
- An explanatory variable is time-dependent if its value for any given individual can change over time

# Time-dependent covariates

- The partial likelihood of Cox also allows time-dependent explanatory variables
- An explanatory variable is time-dependent if its value for any given individual can change over time
- We introduce a latency variable **lat15** indicating 15 years since last fluoroscopy

# Time-dependent covariates

- The partial likelihood of Cox also allows time-dependent explanatory variables
- An explanatory variable is time-dependent if its value for any given individual can change over time
- We introduce a latency variable **lat15** indicating 15 years since last fluoroscopy

# Difficulties in programming

- Most software can account for time-dependent covariates for rate ratio estimation, however none can compute *dfbeta* residuals for these time-dependent covariates
- Thus it is not possible to compute the robust or asymptotic variance estimators for case-cohort data

# Difficulties in programming

- Most software can account for time-dependent covariates for rate ratio estimation, however none can compute *dfbeta* residuals for these time-dependent covariates
- Thus it is not possible to compute the robust or asymptotic variance estimators for case-cohort data



## Proposed solution

- Software can be “tricked” to accommodate time-dependent covariates by organizing the case-cohort data into risk sets
- Has the structure of individually matched case-control data with a risk set formed at each failure time

## Proposed solution

- Software can be “tricked” to accommodate time-dependent covariates by organizing the case-cohort data into risk sets
- Has the structure of individually matched case-control data with a risk set formed at each failure time
- **Case:** is the failure at a specific failure time
- **Controls:** are all those still at risk at the case failure time

## Proposed solution

- Software can be “tricked” to accommodate time-dependent covariates by organizing the case-cohort data into risk sets
- Has the structure of individually matched case-control data with a risk set formed at each failure time
- **Case:** is the failure at a specific failure time
- **Controls:** are all those still at risk at the case failure time

# Analytic dataset

## Analytic Dataset for time dependent covariates

Caseid	set_no	rstime	rsentry	Subject ID	Dose in rad	Age at exit (in years)	Age at entry (in years)	0-cens,1-subc fail 2-non-subc fail	1-249 rad	250+ rad	age at first exposure group	ccohentry	cc	latency	lat15
22	1	25.4292	25.4291	22	0.61714	25.4292	17.1773	1	1	0	1	17.1773	1	8.2519	0
22	1	25.4292	25.4291	2958	4.13045	33.6016	15.8303	0	0	1	1	15.8303	0	9.5989	0
22	1	25.4292	25.4291	295	0.58148	51.833	17.5496	0	1	0	2	17.5496	0	7.8795	0
22	1	25.4292	25.4291	261	0	52.8569	3.4771	0	0	0	1	3.4771	0	21.9521	1
⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮
22	1	25.4292	25.4291	34	0	55.4387	14.9377	1	0	0	1	14.9377	0	10.4914	0
22	1	25.4292	25.4291	334	1.15677	56.6543	18.3381	0	1	0	2	18.3381	0	7.091	0
22	1	25.4292	25.4291	2057	0	73.2402	20.8049	0	0	0	2	20.8049	0	4.6242	0
2350	33	47.7235	47.7234	2350	0.94436	47.7235	20.2218	2	1	0	2	47.7234	1	27.5017	1
2350	33	47.7235	47.7234	3043	0.92604	48.909	17.5414	0	1	0	1	17.5414	0	30.1821	1
2350	33	47.7235	47.7234	242	0.67137	49.1608	15.8795	0	1	0	1	15.8795	0	31.8439	1
2350	33	47.7235	47.7234	2244	0.00959	49.1828	16.9035	0	1	0	2	16.9035	0	30.82	1
2350	33	47.7235	47.7234	3150	0	50.4723	17.191	0	0	0	2	17.191	0	30.5325	1
⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮
2350	33	47.7235	47.7234	3317	1.28865	78.4559	27.7755	0	1	0	3	27.7755	0	19.948	1
2350	33	47.7235	47.7234	3182	0.95419	81.0951	35.05	0	1	0	4	35.05	0	12.6735	0
2350	33	47.7235	47.7234	3258	0.82631	86.642	38.36	0	1	0	4	38.36	0	9.3634	0
2350	33	47.7235	47.7234	3198	0	87.1157	42.5435	0	0	0	4	42.5435	0	5.18	0
3085	75	77.86	77.86	3085	0	77.8645	43.0335	2	0	0	4	77.8644	1	34.83	1
3085	75	77.86	77.86	3317	1.28865	78.4559	27.7755	0	1	0	3	27.7755	0	50.09	1
3085	75	77.86	77.86	3182	0.95419	81.0951	35.05	0	1	0	4	35.05	0	42.81	1
3085	75	77.86	77.86	3258	0.82631	86.642	38.36	0	1	0	4	38.36	0	39.5	1
3085	75	77.86	77.86	3198	0	87.1157	42.5435	0	0	0	4	42.5435	0	35.32	1
3085	75	77.86	77.86	2477	0	89.9849	53.9001	0	0	0	4	53.9001	0	23.96	1

# SAS Code

```
proc phreg data=pclib.td_analytic nosummary;  
  model rstime*cc(0) = dcat1 dcat2 lat15  
    / entry=rscopy covb;  
  output out=dfbetas dfbeta= dfb_dcat1 dfb_dcat2 dfb_lat15;  
  id id;  
run;  
  
proc summary data=dfbetas sum;  
  class id;  
  var dfb_dcat1 dfb_dcat2 dfb_lat15;  
  output out=summed sum=dfb_dcat1 dfb_dcat2 dfb_lat15;  
  where cc eq 0;  
run;  
  
proc corr data=summed cov;  
  var dfb_dcat1 dfb_dcat2 dfb_lat15;  
run;
```

# Exact pseudolikelihood estimators

**Analysis of Maximum Likelihood Estimates**

Parameter	DF	Parameter Estimate	Standard Error	$\chi^2$	$\text{Pr} > \chi^2$	Hazard Ratio	Label
<b>dcat1</b>	1	0.65709	0.26112	6.3325	0.0119	1.929	1-249 rad
<b>dcat2</b>	1	1.68786	0.50750	11.0610	0.0009	4.727	250+ rad
<b>lat15</b>	1	0.61486	0.36062	2.9071	0.0882	1.849	

## Stratification by age at first exposure

- It is quite possible that age is confounding the main effects of the covariates
- To control for confounding we stratify by age at first exposure group
- Each stratum ( $s$ ) contributes independently to the pseudolikelihood
- the asymptotic variance is given by

$$\hat{I}^{-1} + \sum_s \frac{m_s(n_s - m_s)}{n_s} \text{Cov } D_{C_s} \quad (3)$$

## Stratification

<b>Age Stratified Groups</b>	
<b>Age</b>	<b>Group number</b>
<15	1
15-19	2
20-29	3
30+	4



# SAS Code

```
proc phreg data=analytic;
  model an_exit*an_ind(0) = dcat1 dcat2 / entry=an_entry covb;
  output out=dfbetas dfbeta= dfb_dcat1 dfb_dcat2;
  strata agefirstgr;
  id id;
run;

proc corr data=dfbetas cov;
  var dfb_dcat1 dfb_dcat2;
  by agefirstgr;
  where an_ind eq 0;
run;
```

# Exact pseudolikelihood estimators

## Stratified

Analysis of Maximum Likelihood Estimates							
Parameter	DF	Parameter Estimate	Standard Error	$\chi^2$	$\text{Pr} > \chi^2$	Hazard Ratio	Label
dcat1	1	0.5938	0.27148	4.7838	0.0287	1.811	1-249 rad
dcat2	1	0.9349	0.51737	3.2655	0.0708	2.547	250+ rad

## Unstratified

Analysis of Maximum Likelihood Estimates							
Parameter	DF	Parameter Estimate	Standard Error	$\chi^2$	$\text{Pr} > \chi^2$	Hazard Ratio	Label
dcat1	1	0.6572	0.26117	6.332	0.0119	1.929	1-249 rad
dcat2	1	1.55325	0.50118	9.6051	0.0019	4.727	250+ rad

# Summary

- 1 Efficiency and benefits of case-cohort design
- 2 Take advantage of available software

# Summary

- 1 Efficiency and benefits of case-cohort design
- 2 Take advantage of available software

# References I



Bryan Langholz and Jenny Jiao

Computational methods for case cohort studies

*Computational Statistics & Data Analysis*, 51:2007



J. Cologne et al.

Conventional case-cohort design and analysis for studies of interaction

*International Journal of Epidemiology*, 41:2012



Therneau, T and Li, H

Computing the Cox model for case cohort designs

*Lifetime data analysis*, 2(5):1999

## References II



Penny Webb and Chris Bain.

*Essential Epidemiology.*

Cambridge University Press, 2nd edition, 2011.



O Le Polain de Waroux

The case-cohort design in outbreak investigations

*Euro surveillance*, 17(25):2012