

CS660 Homework 5

① Exercise 6.1

Consider the following bivariate distribution (P_{XY}) of two discrete random variables X and Y .

y_1	0.01	0.02	0.03	0.1	0.1
y_2	0.005	0.1	0.05	0.07	0.2
y_3	0.1	0.05	0.03	0.05	0.04

X

Compute:

- (a) The marginal distributions $p(n)$ and $p(y)$.
 → The marginal probability of $X = n_i$, $p(n_i)$ is the sum of the joint probabilities over all values of Y for each n_i :

$$p(n_i) = \sum_y p(n_i, y)$$

Similarly, the marginal probability of $Y = y_j$, $p(y_j)$

$$p(y_j) = \sum_n p(n, y_j)$$

Calculating $p(n)$.

$$p(n_1) = 0.01 + 0.05 + 0.1 = 0.16$$

$$p(n_2) = 0.02 + 0.1 + 0.05 = 0.17$$

$$p(n_3) = 0.03 + 0.05 + 0.03 = 0.11$$

$$p(n_4) = 0.1 + 0.07 + 0.05 = 0.22$$

$$p(n_5) = 0.1 + 0.2 + 0.04 = 0.34.$$

∴ the marginal distribution $p(n)$ is:

$$p(n) = \{0.16, 0.17, 0.11, 0.22, 0.34\}.$$

Calculating $p(y)$

$$p(y_1) = 0.01 + 0.02 + 0.03 + 0.1 + 0.1 = 0.26$$

$$p(y_2) = 0.05 + 0.1 + 0.05 + 0.07 + 0.2 = 0.47$$

$$p(y_3) = 0.1 + 0.05 + 0.03 + 0.05 + 0.04 = 0.27$$

\therefore the marginal distribution $p(y)$ is:

$$p(y) = \{0.26, 0.47, 0.27\}.$$

(b) The conditional distributions $p(n|y=y_1)$ and $p(y|x=n_3)$.

\rightarrow The conditional probability $p(n|y=y_1)$ is given by:

$$p(n|y=y_1) = \frac{p(n, y_1)}{p(y_1)}$$

As $p(y_1) = 0.26$; we calculate values as follows:

$$p(n_1|y=y_1) = \frac{0.01}{0.26} \approx 0.0385$$

$$p(n_2|y=y_1) = \frac{0.02}{0.26} \approx 0.0769$$

$$p(n_3|y=y_1) = \frac{0.03}{0.26} \approx 0.1154$$

$$p(n_4|y=y_1) = \frac{0.1}{0.26} \approx 0.3846$$

$$p(n_5|y=y_1) = \frac{0.1}{0.26} \approx 0.3846$$

\therefore the Conditional probability $p(n|y=y_1)$ is:

$$p(n|y=y_1) \approx \{0.0385, 0.0769, 0.1154, 0.3846, 0.3846\}$$

The conditional probability $p(y|x=n_3)$ is:

$$p(y|x=n_3) = \frac{p(n_3, y)}{p(n_3)}$$

As $p(n_3) = 0.11$, we calculate values as follows:
 $p(y_i | x=n_3) = \frac{0.03}{0.11} \approx 0.2727$

$$p(y_1 | x=n_3) = \frac{0.05}{0.11} \approx 0.4545$$

$$p(y_3 | x=n_3) = \frac{0.03}{0.11} \approx 0.2727$$

\therefore the conditional distribution $p(y|x=n_3)$ is:

$$p(y|x=n_3) \approx \{0.2727, 0.4545, 0.2727\}$$

- ② You have written a computer program that sometimes compiles and sometimes not. (code does not change). You decide to model the apparent stochasticity (success vs no success) n of the compiler using a Bernoulli distribution with parameter u :

$$p(n|u) = u^n (1-u)^{1-n}, n \in \{0, 1\}$$

- Choose a conjugate prior for the Bernoulli likelihood and compute the posterior distribution

$$p(u | n_1, \dots, n_N)$$

- ① The Model and likelihood:

$$p(n|u) = u^n (1-u)^{1-n}, n \in \{0, 1\}$$

$n=1$ indicates success,

$n=0$ indicates failure, u is probability of success.

(ii) Choosing a Conjugate Prior.

The conjugate prior for the Bernoulli likelihood is the Beta distribution. The Beta distribution is parameterized by two shape parameters α and β :

$$p(\mu) \propto \text{Beta}(\mu | \alpha, \beta) = \frac{\mu^{\alpha-1} (1-\mu)^{\beta-1}}{B(\alpha, \beta)}$$

$B(\alpha, \beta)$ is Beta function which normalizes the distribution.

(iii) Likelihood of Observing the Data.

Suppose we observe a set of N independent compilation results $n_1, n_2, n_3, \dots, n_N$, where each n_i is either 0 or 1.

$$\begin{aligned} p(n_1, \dots, n_N | \mu) &= \prod_{i=1}^N p(n_i | \mu) = \\ &= \prod_{i=1}^N \mu^{n_i} (1-\mu)^{1-n_i} \end{aligned}$$

let $S = \sum_{i=1}^N n_i$ be the total number of successful compilations. Then the likelihood simplifies to:

$$p(n_1, \dots, n_N | \mu) = \mu^S (1-\mu)^{N-S}.$$

(iv) Computing the Posterior Distribution.

Applying Baye's theorem,

$$p(\mu | n_1, \dots, n_N) \propto p(n_1, \dots, n_N | \mu) \cdot p(\mu)$$

$$\text{Using our prior } p(\mu) = \text{Beta}(\mu | \alpha, \beta) = \frac{\mu^{\alpha-1} (1-\mu)^{\beta-1}}{B(\alpha, \beta)};$$

$$p(\mu | n_1, \dots, n_N) \propto \mu^S (1-\mu)^{N-S} \cdot \mu^{\alpha-1} (1-\mu)^{\beta-1}$$

Combining the power of μ and $1-\mu$, we have:

$$p(\mu | n_1, \dots, n_N) \propto \mu^{S+\alpha-1} (1-\mu)^{N-S+\beta-1}$$

∴ The posterior distribution for μ is:

$$p(\mu | n_1, \dots, n_N) = \text{Beta}(\mu | \alpha+S, \beta+N-S)$$

③ Exercise 6.4.

→ Event B_1 : The fruit was picked from Bag 1.

Event B_2 : The fruit was picked from Bag 2.

Event M : The fruit picked was a mango.

To find - probability that the mango was picked from Bag 2; $P(B_2|M)$.

Baye's Theorem states:

$$P(B_2|M) = \frac{P(M|B_2) \cdot P(B_2)}{P(M)}$$

① Finding $P(B_1)$ and $P(B_2)$.

with probability 0.6 the coin lands heads and picked from bag 1.

$$P(B_1) = 0.6.$$

with probability 0.4 the coin lands tails and picked from Bag 2.

$$P(B_2) = 0.4.$$

② Finding $P(M|B_1)$ and $P(M|B_2)$.

Bag 1 contains 4 mangoes and 2 apples;

$$P(M|B_1) = \frac{\text{Number of mangoes in Bag 1}}{\text{Total fruits in Bag 1.}}$$

$$= \frac{4}{6}$$

$$= \frac{2}{3}.$$

$$P(M|B_2) = \frac{\text{No. of mangoes in Bag 2}}{\text{Total fruits in Bag 2}}$$

$$= \frac{1}{82} = \frac{1}{2}$$

$$(iii) P(M) = P(M|B_1) \cdot P(B_1) + P(M|B_2) \cdot P(B_2)$$

$$\therefore P(M) = \frac{2}{31} \times 0.6 + \frac{1}{21} \times 0.4$$

$$P(M) = 0.4 + 0.2$$

$$P(M) = 0.6.$$

Applying Baye's Theorem:

$$P(B_2|M) = \frac{P(M|B_2) \cdot P(B_2)}{P(M)}$$

$$= \frac{\frac{1}{82} \times 0.4}{0.6}$$

$$= \frac{0.2}{0.6} = \frac{1}{3}$$

\therefore The probability that the mango was picked from Bag 2 is $\frac{1}{3}$.

(4)

Exercise 6.11

The expectation of n with respect to its joint distribution $p(n, y)$ is:

$$E[n] = \sum_n \sum_y n p(n, y);$$

If n and y are discrete, or as:

$$E[n] = \sum_n \sum_y n p(n, y) dy dn;$$

If n and y are continuous.

We can express the joint distribution of (n, y) as:

$$p(n, y) = p(n|y) p(y)$$

$$\therefore E[n] = \sum_n \sum_y n p(n|y) p(y)$$

Or, in the continuous case:

$$E[n] = \int_n \int_y n f(n|y) p(y) dy dn$$

Since we are summing/integrating over both n and y , we can change the order of summation/integration:

$$E[n] = \sum_y p(y) \sum_n n p(n|y)$$

Or, in the continuous case:

$$E[n] = \int_y p(y) \int_n n p(n|y) dn dy.$$

We can see that the inner sum/integral $\sum_n n p(n|y)$ or $\int_n n p(n|y) dy$ is the definition of the conditional expectation of n given y , denoted $E[n|y]$.

$$E[n|y] = \sum_n p(n|y)$$

or,

$$E(n|y) = \int_n np(n|y) dx$$

$$\therefore E[n] = \sum_y p(y) E[n|y]$$

or,

$$E[n] = \int_y p(y) E[n|y] dy$$

The expression $\sum_y p(y) E[n|y]$ or $\int_y p(y) E[n|y] dy$ is simply the expectation of $E[n|y]$ with respect to y .

$$\therefore E[n] = E_y [E[n|y]]$$

⑤ Exercise 6.13.

Given:

Let X be a continuous random variable with cumulative distribution function $F_X(u)$.
 $\therefore Y = F_X(u)$.

We want to prove that Y is uniformly distributed on $[0, 1]$. That is, we want to show that $Y \sim \text{Uniform}(0, 1)$.

① Find the CDF of Y .

To show that Y is uniformly distributed, we need to find its CDF, $F_Y(y) = P(Y \leq y)$.

By definition of Y , we have:

$$F_Y(y) = P(Y \leq y) = P(F_X(X) \leq y).$$

(ii) Rewrite in terms of X :

Since $F_X(n)$ is a continuous and strictly increasing function (as X has strictly monotonic CDF), we can take the inverse of F_X on both sides of the inequality $F_X(X) \leq y$ to rewrite this probability in terms of X :

$$P(F_X(X) \leq y) = P(X \leq F_X^{-1}(y)).$$

(iii) By the definition of the CDF $F_X(n)$, we know:

$$P(X \leq F_X^{-1}(y)) = F_X(F_X^{-1}(y))$$

Therefore,

$$F_Y(y) = P(Y \leq y) = y$$

The CDF of Y is $F_Y(y) = y$ for $y \in [0, 1]$. This is exactly the CDF of a uniform random variable on $[0, 1]$, since for $Y \sim \text{Uniform}(0, 1)$, we have

$$F_Y(y) = y \text{ for } y \in [0, 1]$$

\therefore CDF of $Y := F_X(X)$ is $F_Y(y) = y$ for $y \in [0, 1]$.

$\therefore Y := F_X(X)$ is uniformly distributed on $[0, 1]$

⑥ Let A be the event that a customer invests in tax-free bonds.

Let B be the event that a customer invests in mutual funds.

Given:

$$P(A) = 0.6 \text{ (tax-free bonds)}$$

$$P(B) = 0.3 \text{ (mutual funds)}$$

$P(A \cap B) = 0.15$ (the probability that the customer invests in both tax-free bonds and mutual funds).

(a) Probability that a customer will invest in either tax-free bonds or mutual funds:

$$\begin{aligned} P(A \cup B) &= P(A) + P(B) - P(A \cap B) \\ &= 0.6 + 0.3 - 0.15 \\ &= 0.9 - 0.15 \end{aligned}$$

$$P(A \cup B) = 0.75$$

(b) Probability that a customer will invest in neither tax-free bonds nor mutual funds:

$$\begin{aligned} P(\text{neither } A \text{ nor } B) &= 1 - P(A \cup B) \\ &= 1 - 0.75 \end{aligned}$$

$$\therefore P(\text{neither } A \text{ nor } B) = \underline{\underline{0.25}}$$

(1)

Let G represent the event that the suspect is guilty.
 Let I represent the event that the suspect is innocent.

Let T represent the event that the truth serum test indicates the suspect is guilty.

Given:

$$P(G) = 0.05 \text{ (guilty)}$$

$$P(I) = 1 - P(G) = 0.95 \text{ (95\% of suspects are innocent)}$$

The test (truth serum) has following properties:

- For guilty suspects: The test correctly identifies them as guilty 90% of the time
 $P(T|G) = 0.9$

This implies

$$P(\text{Not } T|G) = 0.1.$$

- For innocent suspects: The test incorrectly identifies them as guilty 1% of the time
 $P(T|I) = 0.01$

This implies

$$P(\text{Not } T|I) = 0.99.$$

We want to find the probability that the suspect is innocent given that the test indicates they are guilty ($P(I|T)$).

Apply Baye's Theorem,

$$P(I|T) = \frac{P(T|I) \times P(I)}{P(T)}.$$

We can find $P(T)$ by considering both ways in which the test could indicate guilt

$$P(T) = P(T|G) \cdot P(G) + P(T|I) \cdot P(I)$$

⑦ Calculating $P(T)$:

$$\begin{aligned} P(T) &= (0.97) \cdot (0.05) + (0.01) \cdot (0.95) \\ &= 0.045 + 0.0095 \\ &= 0.0545. \end{aligned}$$

⑧ Calculating $P(I|T)$:

$$P(I|T) = \frac{P(T|I) \cdot P(I)}{P(T)}.$$

$$= (0.01) \cdot (0.95)$$

$$0.0545$$

$$= \frac{0.0095}{0.0545}$$

$$\therefore P(I|T) \approx 0.1743$$

So, there is about a 17.43% chance that the suspect is innocent even though the truth serum indicates they are guilty.

⑨ Given:

$$f(n) = 2(1-n), \quad 0 < n < 1$$

and $f(n) = 0$, elsewhere.

The expected value $E(X)$ of X is defined by:

$$E[X] = \int_{-\infty}^{\infty} n f(n) dn$$

Since $f(n) = 0$ outside the interval $(0, 1)$, this simplifies to:

$$E[X] = \int_0^1 n \cdot 2(1-n) dn$$

Substitute $f(n) = 2(1-n)$ into the formula:

$$E(x) = \int_0^1 n \cdot 2(1-n) dn$$

$$= 2 \int_0^1 n(1-n) dn$$

$$E(x) = 2 \int_0^1 (n - n^2) dn$$

$$= 2 \left[\int_0^1 n dn - \int_0^1 n^2 dn \right]$$

$$= 2 \left[\left[\frac{n^2}{2} \right]_0^1 - \left[\frac{n^3}{3} \right]_0^1 \right]$$

$$= 2 \left(\frac{1}{2} - \frac{1}{3} \right)$$

$$= 2 \left(\frac{1}{6} \right)$$

$$\therefore E(x) = \frac{1}{3}$$

Since each unit represents \$5000, the average profit in dollars is:

$$\frac{1}{3} \times 5000$$

$$= \frac{5000}{3} \approx 1666.67.$$

\therefore The average profit per automobile is approximately 1666.67 dollars.

Q Given:

$$\alpha = 3 \text{ and } \beta = 2.$$

$$X \sim \text{Beta}(3, 2)$$

We want to find the probability that at least 80% of the new models will require service during the first year of operation. Mathematically, we need to calculate:

$$P(X \geq 0.8)$$

i) The PDF Beta (α, β) is given by:

$$f(n; \alpha, \beta) = \frac{n^{\alpha-1} (1-n)^{\beta-1}}{B(\alpha, \beta)}, \quad 0 < n < 1$$

As $\alpha = 3$ and $\beta = 2$.

$$\therefore f(n; 3, 2) = \frac{n^2 (1-n)^1}{B(3, 2)}, \quad 0 < n < 1.$$

ii) Calculating $P(X \leq 0.8)$ using Beta CDF.

using beta distribution table, the $P(X \leq 0.8) \approx 0.942$

$$P(X \geq 0.8) = 1 - P(X \leq 0.8)$$

$$\begin{aligned} P(X \geq 0.8) &= 1 - 0.942 \\ &= 0.058 \end{aligned}$$

This means there is about a 5.8% chance that at least 80% of these television sets will require service in their first year.

(10)

Given:

- Service calls arrive at a maintenance center according to a Poisson process with rate λ (calls per minute).
- We have a data set of 20 one-minute periods, yielding an average of 1.8 calls per minute.
- The prior distribution for λ is an exponential distribution with mean 2.

Let:

- Y : be the total number of calls observed over the $20 \frac{one}{minutes}$ periods.
- $Y \sim \text{Poisson}(20\lambda)$.

i) The probability mass function of Y for a Poisson random variable is:

$$P(Y|\lambda) = \frac{(20\lambda)^Y}{Y!} e^{-20\lambda}$$

The total number of calls observed over 20 minute is:
 $Y = 20 \times 1.8 = 36$

we have $Y = 36$ as our observed value.

(ii)

An exponential distribution with mean 2 has a rate parameter $\theta = \frac{1}{2}$, so:

$$p(\lambda) = \theta e^{-\theta\lambda} = \frac{1}{2} e^{-\lambda/2}$$

Substituting likelihood and prior expressions:

$$p(\lambda|Y) \propto (20\lambda)^{36} e^{-20\lambda} \times \frac{1}{2} e^{-\lambda/2}$$

Ignoring Constant terms; we get.

$$P(\lambda | y) \propto \lambda^{36} e^{-(20 + \frac{1}{2}\lambda)} = \lambda^{36} e^{-20 - 0.5\lambda}$$

(iii) The posterior distribution $p(\lambda | y)$ has the form:

$$p(\lambda | y) \propto \lambda^{36} e^{-20 - 0.5\lambda}$$

The Gamma distribution with shape parameter α and rate parameter β has the form:

$$\text{Gamma}(\lambda | \alpha, \beta) \propto \lambda^{\alpha-1} e^{-\beta\lambda}$$

Shape parameter $\alpha = 36 + 1 = 37$

Rate parameter $\beta = 0.5$.

∴ The posterior distribution of λ is:

$$\lambda | y \sim \text{Gamma}(37, 20.5)$$