

MSMS-306 : Lifetime Data Analysis

Assignment: 02

August 3, 2025

Ananda Biswas

Exam Roll Number : 24419STC053

M.Sc. Statistics & Computing (Semester - III)

1. **Question :** Derive the estimators of the model parameters under the right censoring scheme when lifetimes are assumed to be coming from a Pareto distribution.

Ans: The probability density function of a lifetime T having pareto distribution with shape $\alpha > 0$ and minimum value $t_m > 0$ is given by

$$f(t) = \begin{cases} \frac{\alpha t_m^\alpha}{t^{\alpha+1}}, & t \geq t_m \\ 0, & t < t_m \end{cases}$$

The corresponding Cumulative Distribution Function and the Survival Function are

$$\begin{aligned} F(t) = P(T \leq t) &= \int_{t_m}^t f(x) dx = \int_{t_m}^t \frac{\alpha t_m^\alpha}{x^{\alpha+1}} dx \\ &= \alpha t_m^\alpha \int_{t_m}^t x^{-(\alpha+1)} dx \\ &= \alpha t_m^\alpha \left[\frac{x^{-\alpha}}{-\alpha} \right]_{t_m}^t \\ &= -t_m^\alpha (t^{-\alpha} - t_m^{-\alpha}) \\ &= 1 - \left(\frac{t_m}{t} \right)^\alpha, \quad t \geq t_m \end{aligned}$$

$$S(t) = \left(\frac{t_m}{t} \right)^\alpha, \quad t \geq t_m$$

 Type I Censoring : Let the observations are censored after time C .

$$\text{Define } \forall i, \delta_i = \begin{cases} 1, & i\text{-th observation is not censored} \\ 0, & i\text{-th observation is censored} \end{cases}$$

Then the likelihood function is given by

$$\begin{aligned}
L(\alpha, t_m) &= \prod_{i=1}^n f(t_i)^{\delta_i} \cdot S(C)^{1-\delta_i} \\
&= \prod_{i \in A} f(t_i) \cdot \prod_{i \in A^c} S(C) \quad \text{where } A = \{i | \delta_i = 1\} \\
&= \left[\left(\frac{t_m}{C} \right)^\alpha \right]^{n-r} \cdot \prod_{i \in A} \frac{\alpha t_m^\alpha}{t_i^{\alpha+1}}, \quad \text{on assuming the number of uncensored patients to be } r \\
&= \alpha^r t_m^{\alpha n} \cdot \left(\prod_{i \in A} \frac{1}{t_i^{\alpha+1}} \right) \cdot \frac{1}{C^{\alpha(n-r)}}
\end{aligned}$$

The log-likelihood function is given by

$$\ell(\alpha, t_m) = \log L(\alpha, t_m) = r \log \alpha + \alpha n \log t_m - (\alpha + 1) \sum_{i \in A} \log t_i - \alpha(n - r) \log C$$

The partial derivative of $\ell(\alpha, t_m)$ w.r.t. α is as follows.

$$\begin{aligned}
\frac{\partial \ell}{\partial \alpha} &= \frac{\partial}{\partial \alpha} (r \log \alpha) + \frac{\partial}{\partial \alpha} (\alpha n \log t_m) - \frac{\partial}{\partial \alpha} \left((\alpha + 1) \sum_{i \in A} \log t_i \right) - \frac{\partial}{\partial \alpha} (\alpha(n - r) \log C) \\
&= \frac{r}{\alpha} + n \log t_m - \sum_{i \in A} \log t_i - (n - r) \log C
\end{aligned}$$

We set the partial derivative to 0 and solve for α .

$$\frac{r}{\alpha} + n \log t_m - \sum_{i \in A} \log t_i - (n - r) \log C = 0$$

$$\Rightarrow \frac{r}{\alpha} = \sum_{i \in A} \log t_i + (n - r) \log C - n \log t_m$$

$$\therefore \hat{\alpha}_{MLE} = \frac{r}{\sum_{i \in A} \log t_i + (n - r) \log C - n \log \hat{t}_m}$$

$$\text{and } \hat{t}_m = t_{(1)} \quad \text{as } \hat{t}_m \leq t_i \quad \forall i$$

📖 Type II Censoring : Let $t_{(1)}, t_{(2)}, \dots, t_{(r)}$ be the ordered observed lifetimes, and suppose the remaining $n - r$ units are right-censored at $t_{(r)}$. Then the likelihood function for the Pareto distribution is:

$$\begin{aligned}
 L(\alpha, t_m) &= \left[\prod_{i=1}^r f(t_i) \right] \cdot [S(t_{(r)})]^{n-r} \\
 &= \left[\prod_{i=1}^r \frac{\alpha t_m^\alpha}{t_{(i)}^{\alpha+1}} \right] \cdot \left[\left(\frac{t_m}{t_{(r)}} \right)^\alpha \right]^{n-r} \\
 &= \alpha^r t_m^{\alpha r} \cdot \left(\prod_{i=1}^r \frac{1}{t_{(i)}^{\alpha+1}} \right) \cdot \left(\frac{t_m}{t_{(r)}} \right)^{\alpha(n-r)} \\
 &= \alpha^r t_m^{\alpha n} \cdot \left(\prod_{i=1}^r \frac{1}{t_{(i)}^{\alpha+1}} \right) \cdot \frac{1}{t_{(r)}^{\alpha(n-r)}}
 \end{aligned}$$

The log-likelihood function is given by

$$\ell(\alpha, t_m) = \log L(\alpha, t_m) = r \log \alpha + \alpha n \log t_m - (\alpha + 1) \sum_{i=1}^r \log t_{(i)} - \alpha(n - r) \log t_{(r)}$$

The partial derivative of $\ell(\alpha, t_m)$ w.r.t. α is as follows.

$$\begin{aligned}
 \frac{\partial \ell}{\partial \alpha} &= \frac{\partial}{\partial \alpha} (r \log \alpha) + \frac{\partial}{\partial \alpha} (\alpha n \log t_m) - \frac{\partial}{\partial \alpha} \left((\alpha + 1) \sum_{i=1}^r \log t_{(i)} \right) - \frac{\partial}{\partial \alpha} (\alpha(n - r) \log t_{(r)}) \\
 &= \frac{r}{\alpha} + n \log t_m - \sum_{i=1}^r \log t_{(i)} - (n - r) \log t_{(r)}
 \end{aligned}$$

We set the partial derivative to 0 and solve for α .

$$\frac{r}{\alpha} + n \log t_m - \sum_{i=1}^r \log t_{(i)} - (n - r) \log t_{(r)} = 0$$

$$\Rightarrow \frac{r}{\alpha} = \sum_{i=1}^r \log t_{(i)} + (n - r) \log t_{(r)} - n \log t_m$$

$$\therefore \hat{\alpha}_{MLE} = \frac{r}{\sum_{i=1}^r \log t_{(i)} + (n - r) \log t_{(r)} - n \log \hat{t}_m} \quad \text{and } \hat{t}_m = t_{(1)} \text{ as } t_m \leq t_i \quad \forall i$$

✎ Random Censoring : The likelihood function under random censoring is:

$$\begin{aligned}
 L(\alpha, t_m) &= \prod_{i=1}^n [f(t_i)^{\delta_i} \cdot S(t_i)^{1-\delta_i}] \\
 &= \prod_{i=1}^n \left[\left(\frac{\alpha t_m^\alpha}{t_i^{\alpha+1}} \right)^{\delta_i} \cdot \left(\frac{t_m}{t_i} \right)^{\alpha(1-\delta_i)} \right] \\
 &= \alpha^{\sum \delta_i} \cdot t_m^{\alpha n} \cdot \prod_{i=1}^n \left(\frac{1}{t_i^{\delta_i(\alpha+1) + (1-\delta_i)\alpha}} \right)
 \end{aligned}$$

The log-likelihood function is given by

$$\ell(\alpha, t_m) = \log L(\alpha, t_m) = \left(\sum_{i=1}^n \delta_i \right) \log \alpha + \alpha n \log t_m - \sum_{i=1}^n [\delta_i(\alpha + 1) + (1 - \delta_i)\alpha] \log t_i$$

The partial derivative of $\ell(\alpha, t_m)$ w.r.t. α is as follows.

$$\begin{aligned}
 \frac{\partial \ell}{\partial \alpha} &= \left(\sum_{i=1}^n \delta_i \right) \frac{1}{\alpha} + n \log t_m - \sum_{i=1}^n [\delta_i \cdot \log t_i + (1 - \delta_i) \cdot \log t_i] \\
 &= \frac{1}{\alpha} \sum_{i=1}^n \delta_i + n \log t_m - \sum_{i=1}^n \log t_i
 \end{aligned}$$

We set the partial derivative to 0 and solve for α .

$$\begin{aligned}
 \frac{1}{\alpha} \sum_{i=1}^n \delta_i + n \log t_m - \sum_{i=1}^n \log t_i &= 0 \\
 \Rightarrow \frac{1}{\alpha} \sum_{i=1}^n \delta_i &= \sum_{i=1}^n \log t_i - n \log t_m \\
 \therefore \hat{\alpha}_{MLE} &= \frac{\sum_{i=1}^n \delta_i}{\sum_{i=1}^n \log t_i - n \log \hat{t}_m} \quad \text{and } \hat{t}_m = t_{(1)} \quad \text{as } t_m \leq t_i \quad \forall i
 \end{aligned}$$