MSMS 308: Practical 04

Ananda Biswas

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Question

Consider a lifetime variable that follows a **Weibull distribution** with shape parameter α and scale parameter λ . The probability density function of this distribution is given by:

$$f(t; \alpha, \lambda) = \frac{\alpha}{\lambda} \left(\frac{t}{\lambda}\right)^{\alpha - 1} e^{-(t/\lambda)^{\alpha}}, \quad t > 0.$$

Estimate the parameters by MLE and get respective Bias and Mean Squared Errors. Then plot the PDF, Survival Function, Hazard Function and Cumulative Hazard Function using the obtained estimates.

• Build-up for obtaining MLE

For a sample of size n, the likelihood function is given by

$$L(\alpha, \lambda) = \prod_{i=1}^{n} \frac{\alpha}{\lambda} \left(\frac{t_i}{\lambda}\right)^{\alpha - 1} e^{-(t_i/\lambda)^{\alpha}}$$
$$= \left(\frac{\alpha}{\lambda}\right)^n \left(\prod_{i=1}^{n} \frac{t_i}{\lambda}\right)^{\alpha - 1} \exp\left\{-\sum_{i=1}^{n} \left(\frac{t_i}{\lambda}\right)^{\alpha}\right\}.$$

The log-likelihood function is given by

$$l(\alpha, \lambda) = n \log \alpha - n \log \lambda + (\alpha - 1) \sum_{i=1}^{n} \log \left(\frac{t_i}{\lambda}\right) - \sum_{i=1}^{n} \left(\frac{t_i}{\lambda}\right)^{\alpha}.$$

Now,

$$\frac{\partial}{\partial \alpha} l(\alpha, \lambda) = \frac{n}{\alpha} + \sum_{i=1}^{n} \log \left(\frac{t_i}{\lambda} \right) - \sum_{i=1}^{n} \left(\frac{t_i}{\lambda} \right)^{\alpha} \log \left(\frac{t_i}{\lambda} \right) = u(\alpha, \lambda), \text{ say}$$
 (1)

and

$$\frac{\partial}{\partial \lambda} l(\alpha, \lambda) = -\frac{n}{\lambda} + (\alpha - 1) \sum_{i=1}^{n} \frac{\lambda}{t_i} \cdot \left(-\frac{t_i}{\lambda^2} \right) + \sum_{i=1}^{n} \frac{\alpha \cdot t_i^{\alpha}}{\lambda^{\alpha + 1}}$$

$$= -\frac{n}{\lambda} - (\alpha - 1) \frac{n}{\lambda} + \sum_{i=1}^{n} \frac{\alpha \cdot t_i^{\alpha}}{\lambda^{\alpha + 1}}$$

$$= -\frac{n\alpha}{\lambda} + \sum_{i=1}^{n} \frac{\alpha \cdot t_i^{\alpha}}{\lambda^{\alpha + 1}} = v(\alpha, \lambda), \text{ say.}$$

Setting $v(\alpha, \lambda) = 0$ we get,

$$-\frac{n\alpha}{\lambda} + \sum_{i=1}^{n} \frac{\alpha \cdot t_{i}^{\alpha}}{\lambda^{\alpha+1}} = 0$$

$$\Rightarrow \frac{n}{\lambda} = \sum_{i=1}^{n} \frac{t_{i}^{\alpha}}{\lambda^{\alpha+1}}$$

$$\Rightarrow \frac{n}{\lambda} = \frac{1}{\lambda^{\alpha+1}} \sum_{i=1}^{n} t_{i}^{\alpha}$$

$$\Rightarrow \lambda^{\alpha} = \frac{1}{n} \sum_{i=1}^{n} t_{i}^{\alpha}$$

$$\therefore \lambda = \left(\frac{1}{n} \sum_{i=1}^{n} t_{i}^{\alpha}\right)^{\frac{1}{\alpha}}$$
(2)

Setting $u(\alpha, \lambda) = 0$ does not yield any closed form solution. So for getting the ML estimate of α , we resort to numerical methods (here Newton-Raphson method).

Now,

$$u_{\alpha}(\alpha, \lambda) = \frac{\partial}{\partial \alpha} u(\alpha, \lambda) = -\frac{n}{\alpha^2} - \sum_{i=1}^{n} \left(\frac{t_i}{\lambda}\right)^{\alpha} \left[\log\left(\frac{t_i}{\lambda}\right)\right]^2; \tag{3}$$

At each iteration, with the present value of α we calculate λ by using (2); then we use the obtained value of λ in (1) and (3) to improve the estimate of α by Newton-Raphson method.

• R Program

```
estimate_lambda <- function(s, alpha) mean(s^alpha)^(1/alpha)
```

```
u <- function(alpha, lambda, s){
   a <- length(s) / alpha</pre>
```

```
b <- sum(log(s / lambda))

c <- sum((s / lambda)^alpha * log(s / lambda))

return(a + b - c)
}</pre>
```

```
u_alpha <- function(alpha, lambda, s){
    a <- - length(s) / alpha^2
    b <- sum((s / lambda)^alpha * log(s / lambda)^2)
    return(a - b)
}</pre>
```

```
estimate_alpha <- function(s, initial, epsilon = 0.0001, iterations = 100){
    alphas <- c(initial)
    for (i in 2:iterations) {
        1 <- estimate_lambda(s, alphas[i-1])
        alphas[i] <- alphas[i-1] - u(alphas[i-1], l, s) / u_alpha(alphas[i-1], l, s)
        if(abs((alphas[i] - alphas[i-1])) < epsilon) break
    }
    return(alphas[length(alphas)])
}</pre>
```

```
true_alpha <- 3; true_lambda <- 2</pre>
```

```
sample_size <- 100
```

```
alpha_hat = lambda_hat = c()
```

```
x <- rweibull(sample_size, shape = true_alpha, scale = true_lambda)
alpha_hat <- estimate_alpha(x, 1)
lambda_hat <- estimate_lambda(x, alpha_hat)</pre>
```

Estimates of the panameters for sample size = 100 are as follows:

```
alpha_hat; lambda_hat

## [1] 3.232285

## [1] 2.025045
```

Now we shall evaluate the bias and MSE of the estimates.

```
alpha_bias = lambda_bias = alpha_MSE = lambda_MSE = c()
```

```
alpha_estimates = lambda_estimates = c()

for (i in 1:100){
    x <- rweibull(sample_size, shape = true_alpha, scale = true_lambda)
    alpha_estimates[i] <- estimate_alpha(x, 1)
    lambda_estimates[i] <- estimate_lambda(x, alpha_estimates[i])
}

alpha_bias <- mean(alpha_estimates) - true_alpha

lambda_bias <- mean(lambda_estimates) - true_lambda

alpha_MSE <- mean( (alpha_estimates - true_alpha)^2 )

lambda_MSE <- mean( (lambda_estimates - true_lambda)^2 )</pre>
```

```
alpha_bias; lambda_bias; alpha_MSE; lambda_MSE

## [1] 0.02121329

## [1] 0.002625044

## [1] 0.06575218

## [1] 0.004191812
```

We take a mean of the different estimates to be our final estimate.

```
alpha <- mean(alpha_estimates); lambda <- mean(lambda_estimates)</pre>
```

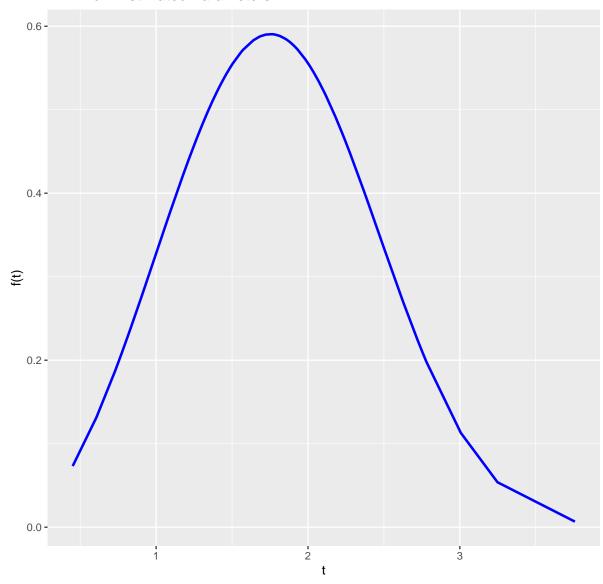
```
alpha; lambda

## [1] 3.021213

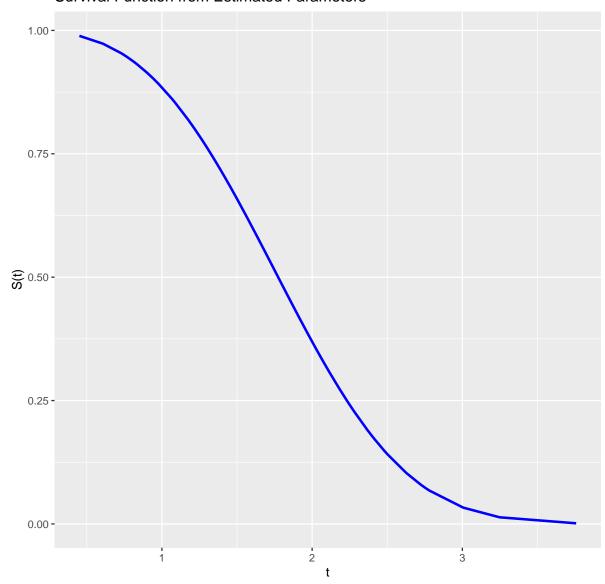
## [1] 2.002625
```

```
t_values <- rweibull(sample_size, scale = lambda, shape = alpha)
```

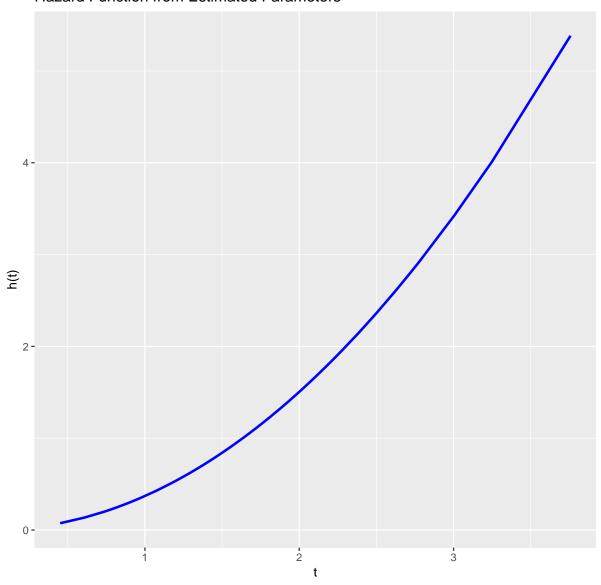
PDF from Estimated Parameters



Survival Function from Estimated Parameters



Hazard Function from Estimated Parameters



Cumulative Hazard Function from Estimated Parameters

