# Stochastic Simulation CW1

Sumith Salluri (CID: 02021144)

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# 1 Question 1

```
[3]: import numpy as np
  import matplotlib.pyplot as plt
  import math
  from scipy import stats
  from scipy import integrate
  from scipy.special import gammaln
  from tqdm import trange
```

## 1.1 Proposed Density

To choose an appropriate proposal distribution q(x) for the Banana density, we need a distribution that captures the non-linear dependency of  $x_2$  on  $x_1$  and has a reasonable spread to ensure efficient sampling with a high acceptance rate.

The Banana density function is given by:

$$\overline{p}(x) = \exp\left(-\frac{x_1^2}{10} - \frac{x_2^2}{10} - 2(x_2 - x_1^2)^2\right)$$

This density has a non-linear structure where  $x_2$  is strongly dependent on  $x_2$ . We can use a product of two Gaussian distributions:  $x_1 \sim \mathcal{N}(0,5)$  and  $x_2 | x_1 \sim \mathcal{N}(x_1^2, (\frac{1}{2})^2)$ 

- $x_1 \sim \mathcal{N}(0,5)$  provides the right spread for  $x_1$  around zero, mirroring the term  $\exp\left(-\frac{x_1^2}{10}\right)$  in the Banana density.
- $x_2 | x_1 \sim \mathcal{N}(x_1^2, (\frac{1}{2})^2)$  makes  $x_2$  centered around  $x_1^2$ , mirroring the term  $\exp(-2(x_2 x_1^2)^2)$  in  $\overline{p}(x)$ .

Thus, we define the proposal distribution as:

$$q(x) = \frac{1}{\sqrt{2\pi \cdot 5}} \exp\left(-\frac{x_1^2}{2 \cdot 5}\right) \cdot \frac{1}{\sqrt{2\pi \cdot \left(\frac{1}{2}\right)^2}} \exp\left(-\frac{(x_2 - x_1^2)^2}{2 \cdot \left(\frac{1}{2}\right)^2}\right).$$

Simplifying, we get:

$$q(x) = \frac{1}{2\pi \cdot 5 \cdot \frac{1}{2}} \exp\left(-\frac{x_1^2}{2 \cdot 5^2} - \frac{(x_2 - x_1^2)^2}{2 \cdot (\frac{1}{2})^2}\right).$$

## 1.2 Calculating M

$$M = \sup \frac{p(x)}{q(x)}$$

$$= (2\pi\sqrt{5} \cdot \frac{1}{2})e^{\left(-\frac{x_1^2}{10} - \frac{x_2^2}{10} - 2(x_2 - x_1^2)^2 + \frac{x_1^2}{2\times 5} + \frac{(x_2 - x_1^2)^2}{2\times (x_2)^2}\right)}$$

$$= \pi\sqrt{5}e^{-\frac{x_2^2}{10}}$$

$$\log M = \log(\pi\sqrt{5}) - \frac{x_2^2}{10}$$

$$\frac{\partial}{\partial x_2}(\log M) = -\frac{x_2}{5} = 0$$

$$\Rightarrow x_2 = 0$$

M is maximized at  $x_2 = 0$ .

We can grid search  $x_1$  for the maximal M. This is implemented in the next code block.

## 1.3 Implementing Rejection Sampling

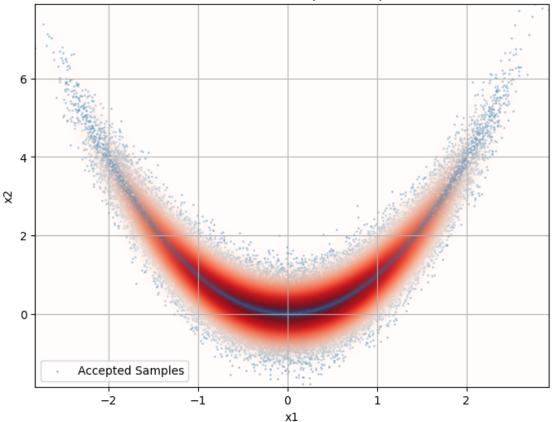
```
[65]: # Define the Banana density function
     def banana_density(x1, x2):
         return np.exp(-x1**2 / 10 - x2**2 / 10 - 2 * (x2 - x1**2)**2)
     # Define the proposal density q(x)
     def proposal_density(x1, x2, s):
         \rightarrow (x1**2))**2/(2*s[1]**2))
     # sample
     def sample_proposal(s):
         x1 = np.random.normal(0, s[0])
         x2 = np.random.normal(x1**2, s[1])
         return x1, x2
     # Function to calculate M by finding the maximum of p(x) / q(x)
     def calculate_M(s):
         x1_vals = np.linspace(-10, 10, 1000)
         max_ratio = 0
         for x1 in x1_vals:
             px = banana_density(x1, 0)
             qx = proposal_density(x1, 0, s=s)
```

```
if qx > 0:
                  ratio = px / qx
                  if ratio > max_ratio:
                      max_ratio = ratio
          return max_ratio
      # Rejection sampling function
      def rejection_sampler(N, s):
          M = calculate_M(s)
          accepted_samples = []
          total_samples = 0
          while len(accepted_samples) < N:</pre>
              # Sample from the proposal distribution q(x)
              x1, x2 = sample_proposal(s=s)
              px = banana_density(x1, x2)
              qx = proposal_density(x1, x2, s=s)
              # Acceptance condition in log-domain
              # print(px, qx)
              if px <= 0:
                  continue
              if math.log(np.random.rand()) < (math.log(px) - math.log(qx) - math.
       \rightarrow \log(M):
                  accepted_samples.append([x1, x2])
              total\_samples += 1
          acceptance_rate = N / total_samples
          return np.array(accepted_samples), M, acceptance_rate
[44]: # Run the rejection sampler and plot the results
      N = 50000
      samples, m, acceptance_rate = rejection_sampler(N, s=[math.sqrt(5), 0.5])
      print(f'Maximal M: {m}\nAcceptance Rate: {acceptance_rate}')
     Maximal M: 7.025634452704435
     Acceptance Rate: 0.5127416294928985
     Acceptance rate is > 0.5.
[35]: # Plot the scatter plot of accepted samples
      plt.figure(figsize=(8, 6))
      x1_vals = np.linspace(min(samples[:, 0]), max(samples[:, 0]), 300)
      x2_vals = np.linspace(min(samples[:, 1]), max(samples[:, 1]), 300)
      X1, X2 = np.meshgrid(x1_vals, x2_vals)
      Z = banana_density(X1, X2)
```

```
plt.scatter(samples[:, 0], samples[:, 1], s=1, alpha=0.4, label="Accepted_\( \) \( \sim \) Samples")
plt.contourf(X1, X2, Z, levels=500, cmap='Reds', alpha=0.2)

plt.title("Scatter Plot of Accepted Samples")
plt.xlabel("x1")
plt.ylabel("x2")
plt.legend()
plt.grid()
plt.show()
```

## Scatter Plot of Accepted Samples



As we can see the samples follow the banana distribution very well indicating the sampling is correct

## 1.4 Integral of Banana Density for theoretical acceptance rate

```
[36]: def integrate_function(f, a, b):
    return integrate.dblquad(f, *a, *b)
z, e = integrate_function(banana_density, a=(-12, 12), b=(-12, 12))
print(f'Integral of banana density: {z}')
```

Integral of banana density: 3.6013257789914386

## [38]: print(f'Theoretical acceptance rate: {z/m}')

Theoretical acceptance rate: 0.5125979444611085

Theoretical acceptance rate ( $\approx 0.5126$ ) and actual acceptance rate ( $\approx 0.5115$ ) are very close indicating correct implementation of proposal density and sampling. Over many runs the actual acceptance rate varies in and around the theoretical rate, sometimes above and sometimes below. This makes sense due to the randomness of the sampling algorithm.

#### 1.5 KS Scores

```
[52]: banana_samples = np.load('samples.npy')
```

KS Score p-values [0.5892151 0.95476728]

KS Scores are well above 0.05 indicating good sampling.

# 2 Question 2

### 2.1 SNIS estimator derivation for the marginal likelihood

$$p(y_{1:T}) = \int p(y_{1:T} \mid x) p(x) dx$$

$$= \frac{\int p(y_{1:T} \mid x) \overline{p}(x) dx}{Z}, \quad \text{where } Z = \int \overline{p}(x) dx$$

$$= \frac{\int \frac{p(y_{1:T} \mid x) \overline{p}(x)}{q(x)} q(x) dx}{\int \frac{\overline{p}(x)}{q(x)} q(x) dx}$$

$$\approx \frac{\frac{1}{N} \sum_{i=1}^{N} \frac{p(y_{1:T} \mid x_i) \overline{p}(x_i)}{q(x_i)}}{\frac{1}{N} \sum_{i=1}^{N} \frac{\overline{p}(x_i)}{\overline{q}(x_i)}}$$

$$= \sum_{i=1}^{N} p(y_{1:T} \mid x_i) \omega_i$$
where  $\omega_i = \frac{W_i}{\sum W_i}, \quad W_i = \frac{\overline{p}(x_i)}{q(x_i)}$ 
Now substituting:  $p(y_{1:T} \mid x) = \prod_{i=1}^{T} p(y_i \mid x)$ 

$$p^{\text{SNIS}}(y_{1:T}) = \sum_{i=1}^{N} \left( \prod_{j=1}^{T} p(y_j \mid x_i) \right) \omega_i$$

**Gaussian** For  $p_1(y_j \mid x) = \mathcal{N}(y_j; x_i, \sigma^2)$ , where  $\sigma^2 = 0.1$ :

$$p_{Gaussian}^{SNIS} = \sum_{i=1}^{N} \prod_{j=1}^{T} \frac{1}{\sqrt{2\pi \cdot 0.1}} \exp\left(-\frac{(y_j - x_i)^2}{2 \cdot 0.1}\right) \cdot \omega_i$$
$$= \sum_{i=1}^{N} \prod_{j=1}^{T} \sqrt{\frac{5}{\pi}} \exp\left(-5(y_j - x_i)^2\right) \cdot \omega_i$$
$$= \left(\frac{5}{\pi}\right)^{T/2} \sum_{i=1}^{N} \exp\left(-5\sum_{j=1}^{T} (y_j - x_i)^2 + \log \omega_i\right)$$

Student t For  $p_2(y_j \mid x) = \text{St}(y_j; x_i, \nu, \sigma^2)$ :

$$p_{Student-t}^{SNIS} = \sum_{i=1}^{N} \prod_{j=1}^{T} \frac{\Gamma\left(\frac{\nu+1}{2}\right)}{\Gamma\left(\frac{\nu}{2}\right)\sqrt{\nu\pi\sigma^{2}}} \left(1 + \frac{(y_{j} - x_{i})^{2}}{\nu\sigma^{2}}\right)^{-\frac{\nu+1}{2}} \cdot \omega_{i}$$

$$= \left(\frac{\Gamma\left(\frac{\nu+1}{2}\right)}{\Gamma\left(\frac{\nu}{2}\right)\sqrt{\frac{\nu\pi}{10}}}\right)^{T} \sum_{i=1}^{N} \prod_{j=1}^{T} \left(1 + \frac{10(y_{j} - x_{i})^{2}}{\nu}\right)^{-\frac{\nu+1}{2}} \cdot \omega_{i}$$

### 2.2 Implementing logsumexp

[15]: # naive implementation

```
def logsumexp_naive(nums):
    return np.log(np.sum(np.exp(nums)))

# stable implementation
def logsumexp_stable(values):
    max_val = np.max(values)
    return max_val + np.log(np.sum(np.exp(values - max_val)))

[16]: # Test with large values
test_arr = np.array([1000, 1010, 1002])

n = logsumexp_naive(test_arr)
print('naive logsumexp', n)
s = logsumexp_stable(test_arr)
print("stable logsumexp", s)
```

```
naive logsumexp inf stable logsumexp 1010.000380790048
```

C:\Users\sumit\AppData\Local\Temp\ipykernel\_16576\1801521852.py:3:
RuntimeWarning: overflow encountered in exp
return np.log(np.sum(np.exp(nums)))

Runtime overflow error proves naive logsumexp is numerically unstable for as values in input array get larger whilst the stable version runs fine

## 2.3 Implementing SNIS procedure for marginal likelihoods

### 2.3.1 Log-SNIS transformation for numerical stability:

#### Gaussian

$$\log p_{\text{Gaussian}}^{\text{SNIS}} = \frac{T}{2} \log \frac{5}{\pi} + \text{LSE} \left( -5 \sum_{j=1}^{T} (y_j - x_i)^2 + \log \omega_i \right)$$

Alternatively:

$$\log p_{\text{Gaussian}}^{\text{SNIS}} = \frac{T}{2} (\log 5 - \log \pi) + \text{LSE} \left( -5 \sum_{j=1}^{T} (y_j - x_i)^2 + \log W_i \right) - \text{LSE} \left( \log W_i \right)$$

### Student t

$$\begin{split} \log p_{\text{Student-t}}^{\text{SNIS}} &= T \log \Gamma \left( \frac{\nu+1}{2} \right) - T \log \Gamma \left( \frac{\nu}{2} \right) - \frac{T}{2} (\log(\pi) + \log(\nu) - \log(10)) \\ + \text{LSE} \left( \log \omega_i - \frac{\nu+1}{2} \log \left( \sum_{j=1}^T \left( 1 + \frac{10(y_j - x_i)^2}{\nu} \right) \right) \right) \end{split}$$

Alternatively:

$$\log p_{\text{Student-t}}^{\text{SNIS}} = \dots \text{(prefix)} \dots + \text{LSE} \left( \log W_i - \frac{\nu+1}{2} \log \left( \sum_{j=1}^T \left( 1 + \frac{10(y_j - x_i)^2}{\nu} \right) \right) \right) - \text{LSE}(\log W_i)$$

These alternate versions allow us to build a list of weights while iterating samples of x and normalise at the end.

```
return - (x[0]**2)/(2*s[0]**2) - (x[1] - (x[0]**2))**2/(2*s[1]**2) - np. \rightarrow \log(2 * np.pi * s[0] * s[1])
```

```
[90]: # Load data
      x_samples = np.load("samples.npy")
      y_data = np.load("y.npy")
      N, T = len(x_samples), len(y_data)
      sigma_sq = 0.1
      sigma = np.sqrt(sigma_sq)
      proposal_s = [math.sqrt(5), 1/2]
      nu=5
      def calc_weight(x):
          log_prior = log_banana_density(x)
          log_proposal = log_proposal_density(x, proposal_s)
          log_weight = log_prior - log_proposal
          return log_weight
      def snis_estimator_gauss(y_data, x_samples, lse_func=logsumexp_stable):
          log_likelihoods = []
          weights = []
          prefix = T/2 * (np.log(5) - np.log(np.pi))
          for x in x_samples:
              log_weight = calc_weight(x)
              log_likelihoods.append(log_weight - 5 * np.sum((y_data - x[0])**2))
              weights.append(log_weight)
          lse = lse_func(log_likelihoods)
          lse_w = lse_func(weights)
          return prefix + lse - lse_w
      def snis_estimator_student_t(y_data, x_samples, v=nu, lse_func=logsumexp_stable):
          log_likelihoods = []
          weights = []
          T = len(y_data)
          # Parts outside the LSE
          log10 = np.log(10)
          logpi = np.log(np.pi)
          lognu = np.log(v)
          log_gamma_nu_plus = gammaln((v + 1) / 2)
          log_gamma_nu = gammaln(v / 2)
```

```
prefix = T * (log_gamma_nu_plus - log_gamma_nu - (1 / 2) * (logpi + lognu -u
       →log10))
          for x in x_samples:
              log_weight = calc_weight(x)
              log_{terms} = np.sum(np.log(1 + (10 * (y_data - x[0])**2) / v))
              log_likelihood = log_weight - (v + 1) / 2 * log_terms
              log_likelihoods.append(log_likelihood)
              weights.append(log_weight)
          lse = lse_func(log_likelihoods)
          lse_w = lse_func(weights)
          return prefix + lse - lse_w
      def snis_estimator(y_data, x_samples, model, v=nu, lse_func=logsumexp_stable):
         match model:
              case 'gaussian':
                  return snis_estimator_gauss(y_data, x_samples, lse_func=lse_func)
              case 'student':
                  return snis_estimator_student_t(y_data, x_samples, v=nu,_
       →lse_func=lse_func)
              case _:
                  print('unknown model')
[98]: # Test runs using samples.npy and y.npy provided
      print("Gaussian (naive):", snis_estimator(y_data, x_samples, model='gaussian',__
       →lse_func=logsumexp_naive))
      print("Student-t (naive):", snis_estimator(y_data, x_samples, model='student',__
      →lse_func=logsumexp_naive))
      print("Gaussian (stable):", snis_estimator(y_data, x_samples, model='gaussian'))
      print("Student-t (stable):", snis_estimator(y_data, x_samples, model='student'))
     C:\Users\sumit\AppData\Local\Temp\ipykernel_8868\1801521852.py:3:
     RuntimeWarning: divide by zero encountered in log
       return np.log(np.sum(np.exp(nums)))
     Gaussian (naive): -inf
     Student-t (naive): -inf
     Gaussian (stable): -24774.842477348193
     Student-t (stable): -7574.39923590361
```

Here we can clearly see the naive versions failing and the stable versions giving us a useful result

## 2.4 Comparing each model and parameter over 100 Runs

```
[101]: # Experiment parameters
       num_runs = 100
       nu_values = np.linspace(1.5, 6.5, 20) # 20 evenly spaced values of nu
       sigma2 = 0.1
       rng = np.random.default_rng() # Random number generator
       num_x_samples = 100000
       rejection_sampling_s = [math.sqrt(5), 0.5]
[90]: # Collect results
       gaussian_results = []
       # Perform 100 runs for the Gaussian model
       for run in trange(num_runs):
           rng = np.random.default_rng(run)
           rejection_sampled_x, _, _ = rejection_sampler(num_x_samples,_
        →rejection_sampling_s)
           log_marginal = snis_estimator_gauss(y_data, rejection_sampled_x)
           gaussian_results.append(log_marginal)
      100%|| 100/100 [06:52<00:00, 4.12s/it]
[98]: np.save('log_marginal_saves/gaussian/gaussian.npy', gaussian_results)
[92]: # Box plot visualization
       plt.figure(figsize=(8, 3))
       plt.boxplot([gaussian_results], positions=[0], widths=0.5,
       →tick_labels=['Gaussian'], vert=False)
       plt.xticks(rotation=90)
       plt.ylabel("Log Marginal Likelihood")
       # plt.tight_layout()
       plt.show()
           Log Marginal Likelihood
             Gaussian
                                                -0.85
```

–2.4774e4

Here we can see the distribution of Gaussian log marginals over 100 runs, each with 100,000 samples of x. The mean is around -24774.82 and the IQR is about [-24774.87, -24774.83].

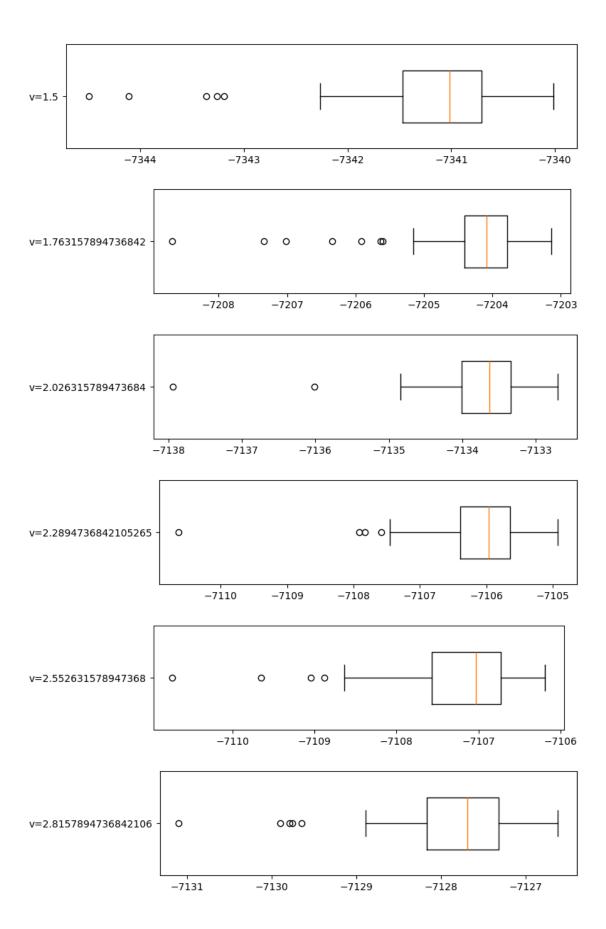
```
[104]: num_x_samples = 1000
       student_t_results = {nu: [] for nu in nu_values}
       # Perform 100 runs for each value of nu in the Student's t model
       for nu in nu_values:
          print(f'v={nu}')
           for run in trange(num_runs):
               rng = np.random.default_rng(run)
               rejection_sampled_x, _, _ = rejection_sampler(num_x_samples,_
       →rejection_sampling_s)
               log_marginal = snis_estimator_student_t(y_data, rejection_sampled_x,_
        بv=nu)
               student_t_results[nu].append(log_marginal)
           np.save(f'log_marginal_saves/student/student_v={nu}.npy',_
        ⇒student_t_results[nu])
      v=1.5
      100%|| 100/100 [00:15<00:00, 6.32it/s]
      v=1.763157894736842
      100%|| 100/100 [00:16<00:00, 6.24it/s]
      v=2.026315789473684
      100%|| 100/100 [00:14<00:00, 6.80it/s]
      v=2.2894736842105265
      100%|| 100/100 [00:15<00:00, 6.37it/s]
      v=2.552631578947368
      100%|| 100/100 [00:15<00:00, 6.64it/s]
      v=2.8157894736842106
      100%|| 100/100 [00:15<00:00, 6.52it/s]
      v=3.0789473684210527
      100%|| 100/100 [00:15<00:00, 6.44it/s]
      v=3.3421052631578947
      100%|| 100/100 [00:15<00:00, 6.45it/s]
      v=3.6052631578947367
      100%|| 100/100 [00:15<00:00, 6.54it/s]
      v=3.8684210526315788
```

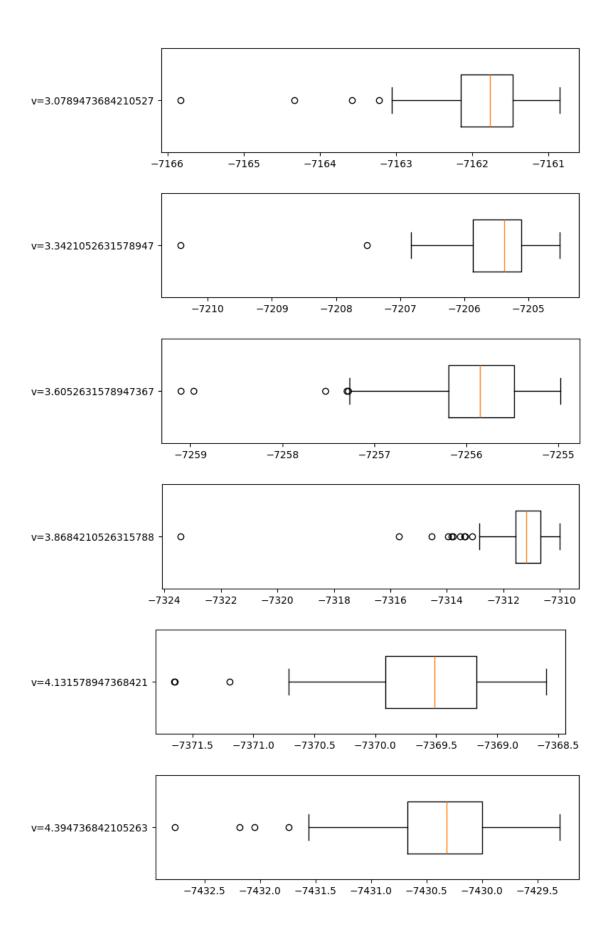
```
100%|| 100/100 [00:15<00:00, 6.56it/s]
v=4.131578947368421
100%|| 100/100 [00:15<00:00, 6.54it/s]
v=4.394736842105263
100%|| 100/100 [00:15<00:00, 6.49it/s]
v=4.657894736842105
100%|| 100/100 [00:15<00:00, 6.49it/s]
v=4.921052631578947
100%|| 100/100 [00:15<00:00, 6.46it/s]
v=5.184210526315789
100%|| 100/100 [00:15<00:00, 6.49it/s]
v=5.447368421052632
100%|| 100/100 [00:15<00:00, 6.51it/s]
v=5.7105263157894735
100%|| 100/100 [00:15<00:00, 6.35it/s]
v=5.973684210526316
100%|| 100/100 [00:16<00:00, 6.03it/s]
v=6.2368421052631575
100%|| 100/100 [00:15<00:00, 6.40it/s]
v = 6.5
100%|| 100/100 [00:15<00:00, 6.38it/s]
```

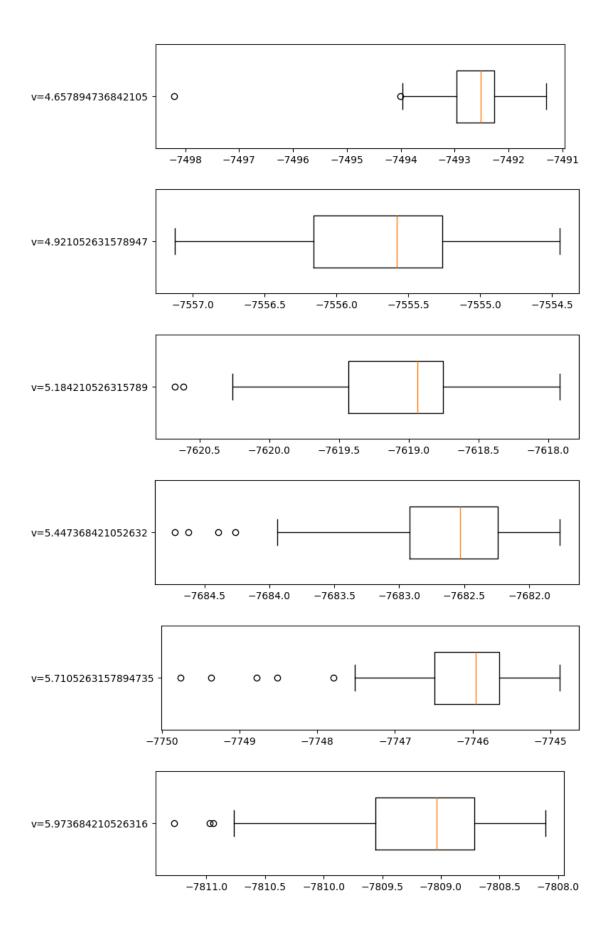
NOTE: Only 1000 samples of x were used for each of these student-t as it takes longer than the gaussian per run and we are iterating over many  $\nu$ s

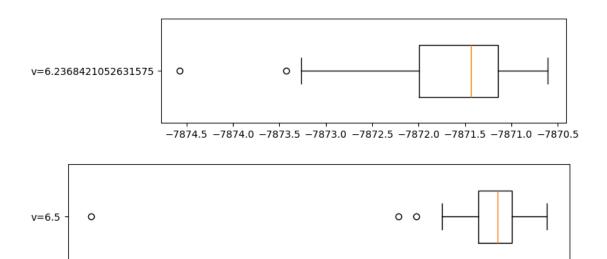
```
for i, (nu, vs) in enumerate(student_t_results.items()):
    plt.figure(figsize=(8, 2))
    plt.boxplot([vs], positions=[0], widths=0.5, tick_labels=[f'v={nu}'],
    vert=False)
    # ax[i % nr][i // nr].boxplot([vs], positions=[0], widths=0.5,
    vick_labels=[f'v={nu}'], vert=False)

plt.tight_layout()
    plt.show()
```









-7938

-7936

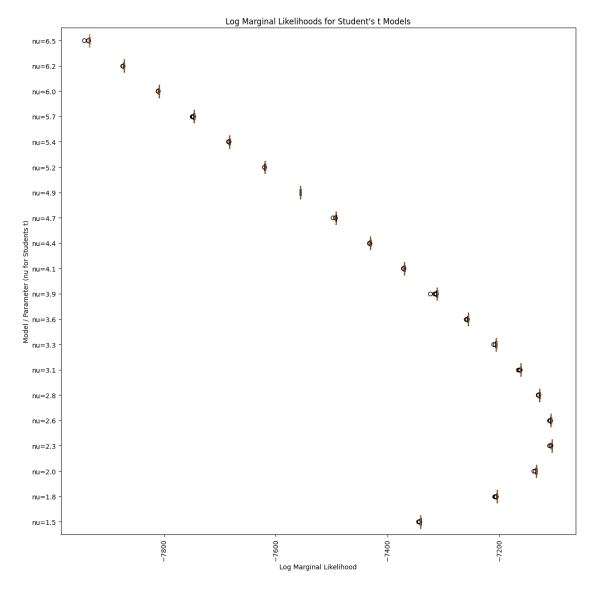
-7934

-7932

Here are the box plots of log marginal scores for 20  $\nu$ s within the range [1.5, 6.5]

-7940

-7942



Here we can see the log marginal means of each model over 100 runs given a  $\nu$  (y-axis). We can clearly see that as  $\nu$  approaches 2.6 from above, the log marginal likelihood increases, peaking at  $\nu = 2.3$  and then decreasing from 2.0 to 1.5. This indicates the optimal  $\nu$  in the given range is around 2.3 for the this Student-t model and fixed  $\sigma = 0.1^2$ .

The Gaussian model had a mean of -24774.82 and peaked at around -24774.68. Even the worst Student-t model in the given range is around 3x better (in terms of log marginal score), with a worst mean of roughly -7950 for  $\nu = 6.5$ . We can conclude that the Student-t model is far superior than the Gaussian.