

PML theory:

(a) By Bayes' Thm, we have:

$$p(c_1|x) = \frac{p(x|c_1)p(c_1)}{p(x|c_1)p(c_1) + p(x|c_2)p(c_2)}$$

$$= \frac{1}{1 + \frac{p(x|c_2)p(c_2)}{p(x|c_1)p(c_1)}}$$

$$= \frac{1}{1 + \frac{p(c_2|x)}{p(c_1|x)}} = \frac{1}{1 + e^{-\log \frac{p(c_1|x)}{p(c_2|x)}}}$$

$p(x)$ ~~are~~
dividing numerator
& denominator
of fraction



(b) We have

Bayes

$$a(x) = \log \frac{p(c_1|x)}{p(c_2|x)} = \log \frac{p(x|c_1)p(c_1)}{p(x|c_2)p(c_2)}$$

$$= \log \left[\exp \left[\varphi(x)^T (w_1 - w_2) \right] + \log \frac{z(w_2)}{z(w_1)} \right] + \log \frac{p(c_1)}{1 - p(c_1)}$$

$$= \varphi(x)^T (w_1 - w_2) + \log \frac{z(w_2)}{z(w_1)} + \log \frac{p(c_1)}{1 - p(c_1)}$$

which is of the form $\varphi(x)^T \theta + \theta_0$
with $\theta := w_1 - w_2$, $\theta_0 = \log \frac{z(w_2)p(c_1)}{z(w_1)(1 - p(c_1))}$

● (c) Revisiting Lecture 5, we see that

$$p(x|e_k) = \exp\left([x, -\frac{1}{2}x^2] \begin{bmatrix} \mu_k/\sigma_k^2 \\ 1/\sigma_k^2 \end{bmatrix} - \left(\frac{\mu_k^2}{2\sigma_k^2} + \log \sqrt{2\pi\sigma_k^2}\right)\right)$$

is the exponential-family-friendly expression for $x|e_k$

$$\bullet \Rightarrow a(x) = [x, -\frac{1}{2}x^2] \begin{bmatrix} \frac{\mu_1}{\sigma_1^2} - \frac{\mu_2}{\sigma_2^2} \\ \frac{1}{\sigma_1^2} - \frac{1}{\sigma_2^2} \end{bmatrix} + \left\{ \frac{\mu_2^2}{2\sigma_2^2} - \frac{\mu_1^2}{2\sigma_1^2} + \log \frac{\sqrt{2\pi\sigma_2^2}}{\sqrt{2\pi\sigma_1^2}} \right\} + \log \frac{p(e_1)}{1-p(e_1)}$$

i.e.

$$\bullet \theta = \begin{pmatrix} \theta_1 \\ \theta_2 \end{pmatrix} = \begin{pmatrix} \frac{\mu_1}{\sigma_1^2} - \frac{\mu_2}{\sigma_2^2} \\ \frac{1}{\sigma_1^2} - \frac{1}{\sigma_2^2} \end{pmatrix}$$

$$\theta_0 = \frac{\mu_2^2}{2\sigma_2^2} - \frac{\mu_1^2}{2\sigma_1^2} + \log \frac{\sigma_2}{\sigma_1} + \log \frac{p(e_1)}{1-p(e_1)}$$

if $\theta_0, \theta_1, \theta_2$ are given:

We have a system of 3 equations in 5 unknowns. ~~There~~;

$$\bullet \left\{ \begin{array}{l} \frac{\mu_1}{\sigma_1^2} - \frac{\mu_2}{\sigma_2^2} = \theta_1 \\ \frac{1}{\sigma_1^2} - \frac{1}{\sigma_2^2} = \theta_2 \\ \frac{\mu_2^2}{2\sigma_2^2} - \frac{\mu_1^2}{2\sigma_1^2} + \log \frac{\sigma_2}{\sigma_1} + \log \frac{p(e_1)}{1-p(e_1)} = \theta_0 \end{array} \right\}$$

- which clearly has no unique solution
so we cannot recover the parameters

if $p(\mathcal{C}_1) = 1/2$ & $\sigma_1, \sigma_2 = 1$ (assuming this is what $\sigma=1$ means?)

our system looks like:

- $$\begin{cases} \mu_1 - \mu_2 = \theta_1 \\ 0 = \theta_2 \dots \text{if } \theta_2 \neq 0 \text{ the variance assumption is wrong!} \\ \mu_2^2 - \mu_1^2 + 0 + 0 = \theta_0 \end{cases}$$

$$\Rightarrow \mu_2^2 - \mu_1^2 = (\mu_2 - \mu_1)(\mu_2 + \mu_1) = \theta_0$$

$$\Rightarrow -\theta_1 (\mu_2 + \mu_1) = \theta_0$$

$$\Rightarrow \mu_2 + \mu_1 = -\frac{\theta_0}{\theta_1}$$

$$\& \mu_1 - \mu_2 = \theta_1$$

$$\Rightarrow \begin{cases} \mu_1 = \frac{1}{2} \left(\theta_1 - \frac{\theta_0}{\theta_1} \right) \\ \mu_2 = -\frac{1}{2} \left(\theta_1 + \frac{\theta_0}{\theta_1} \right) \end{cases}$$

- So we recover the parameters in this case.