

# Principal Component Analysis

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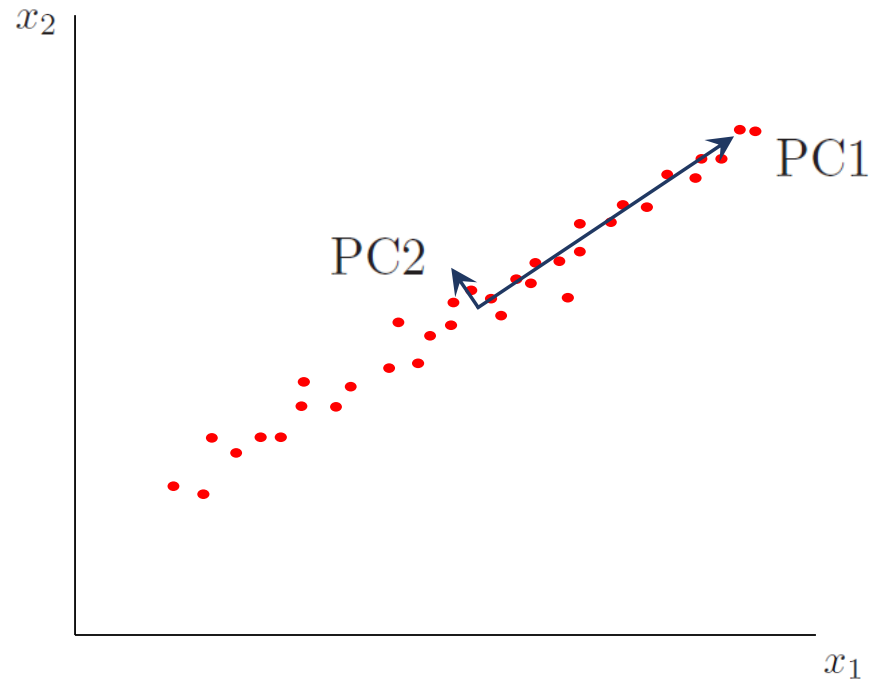
# Dimensionality reduction

- Dimensionality reduction is the process of reducing the number of features of a dataset.
- Types: Feature selection, Feature extraction.
- Feature selection: Selects a subset of features.
  - Removes irrelevant features from the dataset.
- Feature extraction: Selects a few combinations of input features that capture most of the variations of the data.
  - Creates new features (through transformation) using existing ones.

# Introduction to PCA

- Widely used method for dimensionality reduction.
- Original dataset – large number of interrelated input variables.
- Consider dataset:  $\{\mathbf{x}^{(1)}, \mathbf{x}^{(2)}, \dots, \mathbf{x}^{(N)}\}$ , where  $\mathbf{x}^{(n)}$  is a  $D$  dimensional variable.
- Goal: Represent the data in a lower dimension  $Q$  ( $< D$ ).
  - Transform the data to a new uncorrelated set of variables – the principal components.
  - Extraction of the most informative  $Q$  linear combinations which explains the data.
  - This is the projection of the data in  $D$  dimensions onto a lower-dimensional subspace.
- Orthogonal projection of data onto a lower dimensional (linear) space, such that the variance of the projected data is maximized.

# Principal components



- PC1: Direction of most variation
- PC2: Direction of second most variation orthogonal to PC1

# Dataset

- Consider dataset:  $\{\mathbf{x}^{(1)}, \mathbf{x}^{(2)}, \dots, \mathbf{x}^{(N)}\}$ , where  $\mathbf{x}^{(n)}$  is a  $D$  dimensional variable.

$$\mathbf{X} = \begin{bmatrix} x_1^{(1)} & x_1^{(2)} & \cdot & \cdot & x_1^{(N)} \\ x_2^{(1)} & x_2^{(2)} & \cdot & \cdot & x_2^{(N)} \\ \cdot & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot \\ x_D^{(1)} & \cdot & \cdot & \cdot & x_D^{(N)} \end{bmatrix} = \begin{bmatrix} \mathbf{x}_1 \\ \mathbf{x}_2 \\ \cdot \\ \cdot \\ \mathbf{x}_D \end{bmatrix}$$

- Want a lower-dimensional ( $Q < D$ ) representation of the data:

$$\mathbf{Z} = \begin{bmatrix} z_1^{(1)} & z_1^{(2)} & \cdot & \cdot & z_1^{(N)} \\ \cdot & \cdot & \cdot & \cdot & \cdot \\ z_Q^{(1)} & \cdot & \cdot & \cdot & z_Q^{(N)} \end{bmatrix}$$

# Variance

- Consider a vector  $\mathbf{x} = [x_1, x_2, \dots, x_N]$  having a mean value of 0.
- The variance of the vector  $\mathbf{x}$  can be computed as

$$\begin{aligned}\sigma_{\mathbf{x}}^2 &= \frac{1}{N-1} \sum_{i=1}^N (x_i - 0)(x_i - 0) \\ &= \frac{1}{N-1} \mathbf{x} \mathbf{x}^T\end{aligned}$$

# Covariance

- Now consider two vectors:  $\mathbf{x} = [x_1, x_2, \dots, x_N]$  and  $\mathbf{z} = [z_1, z_2, \dots, z_N]$ , both having mean 0.
- The covariance between vectors  $\mathbf{x}$  and  $\mathbf{z}$  can be computed as

$$\sigma_{\mathbf{xz}}^2 = \frac{1}{N-1} \mathbf{xz}^T$$

– Covariance measures the correlation between variables.

- If  $\sigma_{\mathbf{xz}}^2 \approx 0$  then  $\mathbf{x}$  and  $\mathbf{z}$  are almost uncorrelated.



# Covariance matrix

- Assume data is centered.
- The covariance matrix  $\mathbf{S}$  can be obtained as:

$$\mathbf{S} = \frac{1}{N-1} \mathbf{X} \mathbf{X}^T.$$

- Can write the covariance matrix as

$$\mathbf{S} = \begin{bmatrix} \sigma_{\mathbf{x}_1}^2 & \sigma_{\mathbf{x}_1 \mathbf{x}_2} & \cdot & \cdot & \sigma_{\mathbf{x}_1 \mathbf{x}_D} \\ \sigma_{\mathbf{x}_2 \mathbf{x}_1} & \sigma_{\mathbf{x}_2}^2 & \cdot & \cdot & \sigma_{\mathbf{x}_2 \mathbf{x}_D} \\ \cdot & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot \\ \sigma_{\mathbf{x}_D \mathbf{x}_1} & \cdot & \cdot & \cdot & \sigma_{\mathbf{x}_D}^2 \end{bmatrix}$$

- The  $i$ -th diagonal term corresponds to the variance in the  $i$ -th dimension of the problem.
- The off-diagonal terms are the covariances.
- Small off-diagonal term indicates that the variables are almost uncorrelated.
- $\mathbf{S}$  is symmetric.



# Covariance matrix

- Want to transform the covariance matrix  $\mathbf{S}$  to  $\mathbf{S}_{\mathbf{Z}}$  that has the following form:

$$\mathbf{S}_{\mathbf{Z}} = \begin{bmatrix} \sigma_{\mathbf{Z}_1}^2 & 0 & \cdot & \cdot & 0 \\ 0 & \sigma_{\mathbf{Z}_2}^2 & \cdot & \cdot & 0 \\ \cdot & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot \\ 0 & \cdot & \cdot & \cdot & \sigma_{\mathbf{Z}_D}^2 \end{bmatrix}$$

- The transformed matrix  $\mathbf{S}_{\mathbf{Z}}$  has no correlation between the different dimensions.
- Can order the variances such that:  $\sigma_{\mathbf{Z}_1}^2 \geq \sigma_{\mathbf{Z}_2}^2 \geq \dots \geq \sigma_{\mathbf{Z}_D}^2$ .
- So  $\sigma_{\mathbf{Z}_1}^2$  is the largest variance, and the dimension corresponding to it is known as the first principal component.
- Similarly  $\sigma_{\mathbf{Z}_2}^2$  is the variance of the second principal component.

# Eigenvalue decomposition

- Eigenvalue decomposition of the covariance matrix  $\mathbf{S}$ :

$$\mathbf{S} = \mathbf{V}\mathbf{\Lambda}\mathbf{V}^{-1}$$

where  $\mathbf{\Lambda}$  is a diagonal matrix,  $\mathbf{V}$  is a matrix of eigenvectors of  $\mathbf{S}$  with columns corresponding to right eigenvectors of  $\mathbf{S}$ .

- The diagonal elements of  $\mathbf{\Lambda}$  are the eigenvalues of  $\mathbf{S}$  for the corresponding eigenvectors.
- Since  $\mathbf{S}$  is symmetric, the eigenvalues are real and the eigenvectors are orthogonal to each other.
- The eigenvectors can be made orthonormal by taking  $\mathbf{V}\mathbf{V}^T = \mathbf{I}$ .

# Linear transformation

- Consider the following linear transformation of the original data  $\mathbf{X}$  into  $\mathbf{Z}$ :

$$\mathbf{Z} = \mathbf{V}^T \mathbf{X}$$

$$\begin{bmatrix} \uparrow & \uparrow & \dots & \uparrow \\ \mathbf{z}^{(1)} & \mathbf{z}^{(2)} & \dots & \mathbf{z}^{(N)} \\ \downarrow & \downarrow & \dots & \downarrow \end{bmatrix} = \begin{bmatrix} \uparrow & \uparrow & \dots & \uparrow \\ \mathbf{v}_1 & \mathbf{v}_2 & \dots & \mathbf{v}_D \\ \downarrow & \downarrow & \dots & \downarrow \end{bmatrix}^T \begin{bmatrix} \uparrow & \uparrow & \dots & \uparrow \\ \mathbf{x}^{(1)} & \mathbf{x}^{(2)} & \dots & \mathbf{x}^{(N)} \\ \downarrow & \downarrow & \dots & \downarrow \end{bmatrix}$$

# Linear transformation

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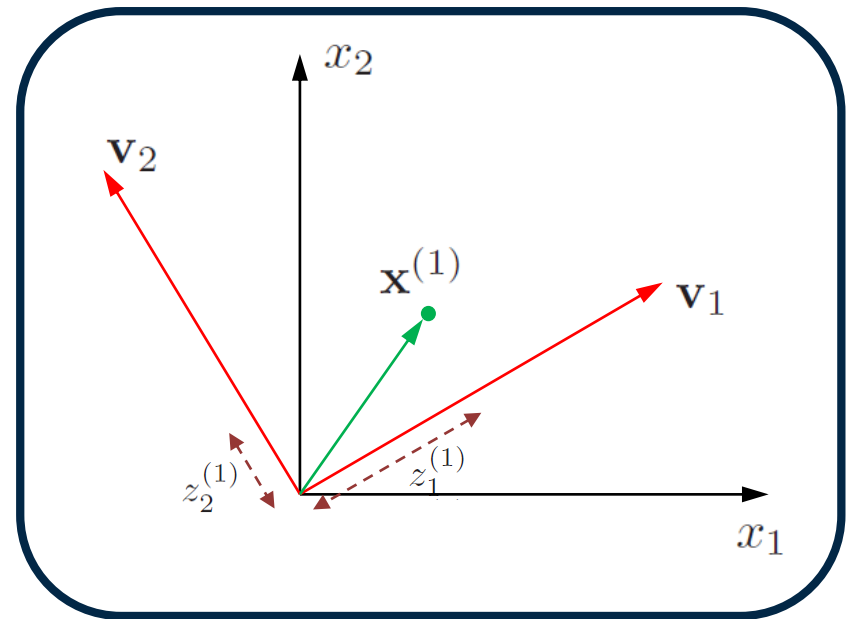
$$\mathbf{Z} = \mathbf{V}^T \mathbf{X}$$

- Consider a 2D example where the transformation is applied to a single data point  $\mathbf{x}^{(1)}$

$$\begin{bmatrix} \updownarrow \mathbf{z}^{(1)} \downarrow \end{bmatrix} = \begin{bmatrix} \leftarrow \mathbf{v}_1 \rightarrow \\ \leftarrow \mathbf{v}_2 \rightarrow \end{bmatrix} \begin{bmatrix} \updownarrow \mathbf{x}^{(1)} \downarrow \end{bmatrix}$$

↙

$$\begin{bmatrix} z_1^{(1)} \\ z_2^{(1)} \end{bmatrix}$$



# Linear transformation

- Consider the following linear transformation of the original data  $\mathbf{X}$  into  $\mathbf{Z}$ :

$$\mathbf{Z} = \mathbf{V}^T \mathbf{X}$$

- The covariance of  $\mathbf{Z}$  can be obtained as:

$$\begin{aligned} \mathbf{S}_Z &= \frac{1}{N-1} \mathbf{Z} \mathbf{Z}^T \\ &= \frac{1}{N-1} (\mathbf{V}^T \mathbf{X}) (\mathbf{V}^T \mathbf{X})^T \\ &= \frac{1}{N-1} (\mathbf{V}^T \mathbf{X}) (\mathbf{X}^T \mathbf{V}) \\ &= \frac{1}{N-1} \mathbf{V}^T (\mathbf{X} \mathbf{X}^T) \mathbf{V} \\ &= \mathbf{V}^T \left( \frac{1}{N-1} \mathbf{X} \mathbf{X}^T \right) \mathbf{V} \\ &= \mathbf{V}^T \mathbf{S} \mathbf{V} \end{aligned}$$

# Covariance matrix

- Consider the following linear transformation of the original data  $\mathbf{X}$  into  $\mathbf{Z}$ :

$$\mathbf{Z} = \mathbf{V}^T \mathbf{X}$$

- The covariance of  $\mathbf{Z}$  can be obtained as:

$$\begin{aligned}\mathbf{S}_Z &= \mathbf{V}^T \mathbf{V} \Lambda \mathbf{V}^{-1} \mathbf{V} \\ &= (\mathbf{V}^T \mathbf{V}) \Lambda (\mathbf{V}^T \mathbf{V}) \quad (\mathbf{V}^{-1} = \mathbf{V}^T \text{ as } \mathbf{V} \mathbf{V}^T = \mathbf{I}) \\ &= \Lambda\end{aligned}$$

- The covariance matrix  $\mathbf{S}_Z$  is diagonal as  $\Lambda$  is diagonal.

# Diagonal covariance matrix

- So we have

$$\mathbf{S}_{\mathbf{Z}} = \Lambda = \begin{bmatrix} \sigma_{\mathbf{Z}_1}^2 & 0 & \cdot & \cdot & 0 \\ 0 & \sigma_{\mathbf{Z}_2}^2 & \cdot & \cdot & 0 \\ \cdot & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot \\ 0 & \cdot & \cdot & \cdot & \sigma_{\mathbf{Z}_D}^2 \end{bmatrix}$$

- The diagonal terms of  $\mathbf{S}_{\mathbf{Z}}$  correspond to variances along the dimensions of the transformed vector space.
- Note, the diagonal matrix  $\Lambda$  comprise the eigenvalues of  $\mathbf{S}$ .
- The variances along the projected dimensions (eigenvectors of  $\mathbf{S}$ ) are the corresponding eigenvalues of  $\mathbf{S}$ .



# **METHOD of LAGRANGE MULTIPLIERS**

# Method of Lagrange multipliers

- The first principal component can be written as linear combination of the original variables as

$$\begin{aligned}z_1 &= v_{11}x_1 + v_{12}x_2 + \dots + v_{1D}x_D \\ &= \mathbf{v}_1^T \mathbf{x}\end{aligned}$$

where  $\mathbf{v}_1^T = [v_{11}, v_{12}, \dots, v_{1D}]$ .

- For the  $N$  given data points, the corresponding vector in the first dimension is given as

$$\mathbf{z}_1 = \mathbf{v}_1^T \mathbf{X}.$$

- The variance in the first dimension is given as

$$\begin{aligned}\text{var}(\mathbf{z}_1) &= \text{var}(\mathbf{v}_1^T \mathbf{X}) \\ &= \frac{1}{N-1} \mathbf{v}_1^T \mathbf{X} \mathbf{X}^T \mathbf{v}_1 \\ &= \mathbf{v}_1^T \mathbf{S} \mathbf{v}_1\end{aligned}$$

and we want  $\text{var}(\mathbf{z}_1)$  to be maximized.

# 1<sup>st</sup> principal component

- Maximize the projected variance  $\mathbf{v}_1^T \mathbf{S} \mathbf{v}_1$  with respect to  $\mathbf{v}_1$  subject to normalization constraint:  $\mathbf{v}_1^T \mathbf{v}_1 = 1$ .
- Approach: Use the method of Lagrange multiplier to find the maximum of an objective function subject to a constraint.
- Consider the Lagrangian:  $\mathcal{L}_1 = \mathbf{v}_1^T \mathbf{S} \mathbf{v}_1 + \lambda_1 (1 - \mathbf{v}_1^T \mathbf{v}_1)$
- Objective:  $\max \mathcal{L}_1$ 
  - Differentiating  $\mathcal{L}_1$  with respect to  $\mathbf{v}_1$  and equating to 0:

$$\frac{d\mathcal{L}_1}{d\mathbf{v}_1} = \mathbf{S} \mathbf{v}_1 - \lambda_1 \mathbf{v}_1 = 0$$

# 1<sup>st</sup> principal component

- Therefore we have

$$\mathbf{S}\mathbf{v}_1 = \lambda_1 \mathbf{v}_1$$

- $\lambda_1$  is an eigenvalue of  $\mathbf{S}$ , and  $\mathbf{v}_1$  is the associated eigenvector.

- Multiplying both sides by  $\mathbf{v}_1^T$ , we have:

$$\begin{aligned}\mathbf{v}_1^T \mathbf{S} \mathbf{v}_1 &= \lambda_1 \mathbf{v}_1^T \mathbf{v}_1 \\ &= \lambda_1\end{aligned}$$

- Note that we want to maximize  $\mathbf{v}_1^T \mathbf{S} \mathbf{v}_1$ .
- Therefore  $\lambda_1$  is the largest eigenvalue of  $\mathbf{S}$ . This is called the 1st principal component.

## 2<sup>nd</sup> principal component

- The second principal component too can be written as linear combination of the original variables as

$$\begin{aligned} z_2 &= v_{21}x_1 + v_{22}x_2 + \dots + v_{2D}x_D \\ &= \mathbf{v}_2^T \mathbf{X} \end{aligned}$$

where  $\mathbf{v}_2^T = [v_{21}, v_{22}, \dots, v_{2D}]$ .

- The projection of the  $N$  data points in the second dimension can be given as

$$\mathbf{z}_2 = \mathbf{v}_2^T \mathbf{X}.$$

- Want  $\mathbf{z}_2$  to be uncorrelated to  $\mathbf{z}_1$  i.e.

$$\text{covariance}(\mathbf{z}_1, \mathbf{z}_2) = 0$$

therefore we have

$$\frac{1}{N-1} \mathbf{v}_1^T \mathbf{X} \mathbf{X}^T \mathbf{v}_2 = 0 \quad \Rightarrow \quad \mathbf{v}_1^T \left( \frac{1}{N-1} \mathbf{X} \mathbf{X}^T \right) \mathbf{v}_2 = 0$$

$$\Rightarrow \mathbf{v}_1^T \mathbf{S} \mathbf{v}_2 = 0 \Rightarrow \mathbf{v}_2^T \mathbf{S} \mathbf{v}_1 = 0 \Rightarrow \mathbf{v}_2^T \lambda_1 \mathbf{v}_1 = 0 \Rightarrow \lambda_1 \mathbf{v}_2^T \mathbf{v}_1 = 0 \Rightarrow \mathbf{v}_2^T \mathbf{v}_1 = 0$$

- Objective:  $\max \mathbf{v}_2^T \mathbf{S} \mathbf{v}_2$  such that  $\mathbf{v}_2^T \mathbf{v}_2 = 1$  and  $\mathbf{v}_2^T \mathbf{v}_1 = 0$

# 2<sup>nd</sup> principal component

- Construct the Lagrangian:

$$\mathcal{L}_2 = \mathbf{v}_2^T \mathbf{S} \mathbf{v}_2 + \lambda_2(1 - \mathbf{v}_2^T \mathbf{v}_2) + \mu_1(\mathbf{v}_2^T \mathbf{v}_1)$$

- Objective:  $\max \mathcal{L}_2$

- Differentiating  $\mathcal{L}_2$  with respect to  $\mathbf{v}_2$  and equating to 0:

$$\frac{d\mathcal{L}_2}{d\mathbf{v}_2} = 2\mathbf{S}\mathbf{v}_2 - 2\lambda_2\mathbf{v}_2 + \mu_1\mathbf{v}_1 = 0 \quad \text{-----} \blacksquare$$

- Multiplying both sides by  $\mathbf{v}_1^T$ :

$$2\mathbf{v}_1^T \mathbf{S} \mathbf{v}_2 - 2\lambda_2 \mathbf{v}_1^T \mathbf{v}_2 + \mu_1 \mathbf{v}_1^T \mathbf{v}_1 = 0$$

- Using  $\mathbf{v}_1^T \mathbf{v}_1 = 1$ , and  $\mathbf{v}_1^T \mathbf{v}_2 = 0$  for  $\max \mathcal{L}_2$ , we have

$$2\mathbf{v}_1^T \mathbf{S} \mathbf{v}_2 + \mu_1 = 0 \quad \text{-----} \blacksquare$$

## 2<sup>nd</sup> principal component

- Now we have

$$\mathbf{v}_1^T \mathbf{S} \mathbf{v}_2 = \mathbf{v}_2^T (\mathbf{S} \mathbf{v}_1) = \mathbf{v}_2^T (\lambda_1 \mathbf{v}_1) = \lambda_1 (\mathbf{v}_2^T \mathbf{v}_1) = 0,$$

and on substitution in ■ gives

$$\mu_1 = 0.$$

- On substitution of  $\mu_1 = 0$  in ■ yields

$$\mathbf{S} \mathbf{v}_2 = \lambda_2 \mathbf{v}_2.$$

- Therefore  $\lambda_2$  is another eigenvalue of  $\mathbf{S}$ .
- Since we want to maximize  $\mathbf{v}_2^T \mathbf{S} \mathbf{v}_2$  ( $= \mathbf{v}_2^T \lambda_2 \mathbf{v}_2 = \lambda_2$ ) and also want  $\mathbf{v}_2$  to be uncorrelated to  $\mathbf{v}_1$ ,  $\lambda_2$  should be the second largest eigenvalue of  $\mathbf{S}$ .



# Percentage of variance

- The percentage of variance explained by the  $j$ th principal component:

$$PV_j = \frac{\lambda_j}{\sum_{i=1}^D \lambda_i} \times 100$$

- The percentage of variance accounted for by the first  $Q$  principal components is given by:

$$PV = \frac{\sum_{i=1}^Q \lambda_i}{\sum_{i=1}^D \lambda_i} \times 100$$