Ridge Regression, LASSO and Smoothing

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Task 1

For this first task we will be using the Communities and Crime dataset from the R package mogavs. This contains 1595 datapoints each with data on p = 122 variables and one target variable. We use a random 80%/20% split to obtain distinct training and test sets with 1595 and 399 datapoints respectively.

```
library(mogavs) # for crime & communities dataset
data("crimeData")
p = ncol(crimeData) - 1
propTrain = 0.8
trainIdx = sample(1:nrow(crimeData), nrow(crimeData)*0.8)
XTrain = crimeData[trainIdx, -(p+1)]
yTrain = crimeData[trainIdx, p+1]
XTest = crimeData[-trainIdx, -(p+1)]
yTest = crimeData[-trainIdx, p+1]
dim(XTrain); length(yTrain)
## [1] 1595 122
## [1] 1595
dim(XTest); length(yTest)
## [1] 399 122
## [1] 399
Next we center and scale our data.
# center the X data around the training mean
XTrain_mean = colMeans(XTrain)
XTrain = XTrain - rep(XTrain_mean, rep.int(nrow(XTrain), p))
XTest = XTest - rep(XTrain_mean, rep.int(nrow(XTest), p))
# scale the data so that variables have unit variance
D = diag(cov(XTrain))
XTrain = XTrain * (D^(-1/2))
Xest = XTest * (D^{-1/2})
```

Ridge Regression

For both ridge and LASSO regression we use the package glmnet, which imposes a penalty given by:

$$(1-\alpha)/2||\beta||_2^2 + \alpha||\beta||_1.$$

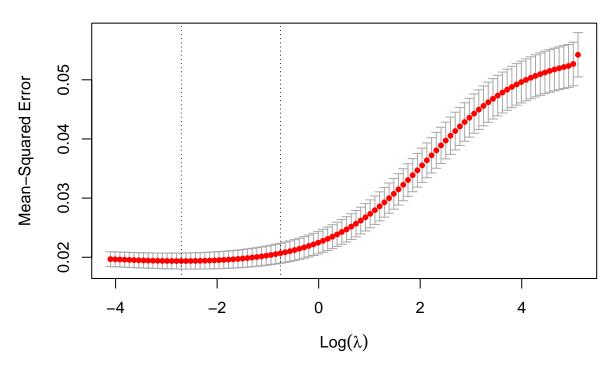
Clearly this gives us the ridge penalty for $\alpha = 0$ and the LASSO penalty for $\alpha = 1$, hence we proceed with $\alpha = 0$ and use the default 10-fold cross validation to find the best value for λ by which to multiply the penalty.

library(glmnet)

Loading required package: Matrix

Loaded glmnet 4.1-6

ridgeCV = cv.glmnet(as.matrix(XTrain), yTrain, alpha=0) #k=10-fold cv by default
plot(ridgeCV)



This leads to an optimum lambda of 0.06677693.

```
optLambda = ridgeCV$lambda.min
optLambda
```

[1] 0.06677693

Using this value of lambda we can then obtain predicted target values \hat{y} and see that the MSE of these compared the true target values y is roughly 0.0456.

```
ridge = glmnet(as.matrix(XTrain), yTrain, alpha=0, lambda=optLambda)

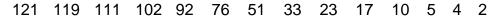
yRidge = predict(ridge, as.matrix(XTest))
sum((yTest - yRidge)^2)/length(yTest)
```

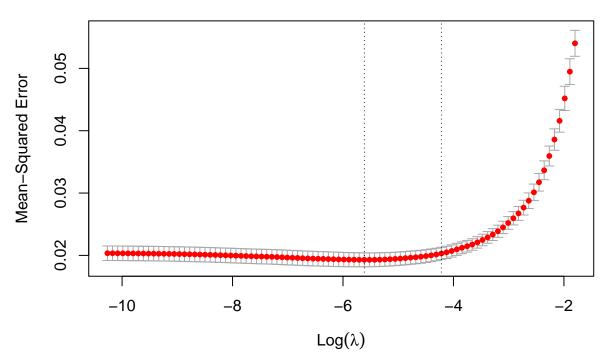
[1] 0.04559334

LASSO Regression

We now repeat the above but with $\alpha = 1$ to use LASSO regression instead of ridge regression.

lassoCV = cv.glmnet(as.matrix(XTrain), yTrain, alpha=1) # k=10-fold cv by default plot(lassoCV)





(The ticks along the top of the plot denotes the number of non-zero coefficients for the corresponding value of λ .)

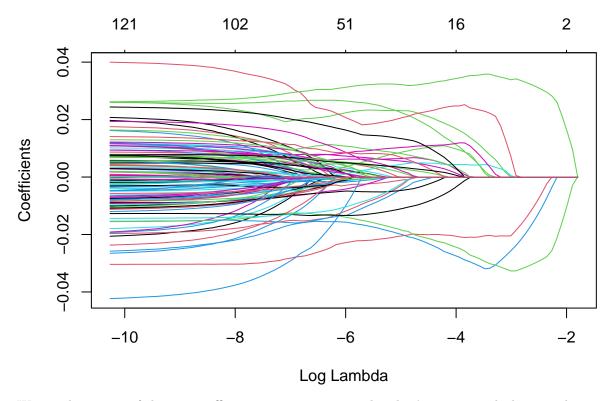
This then leads to the optimum lambda value of 0.003647541.

```
optLambda = lassoCV$lambda.min
optLambda
```

[1] 0.003647541

We can also plot the LASSO path. Note that we are actually plotting the *scaled* values of our β estimate (given as γ in the notes), which we will rescale to obtain our actual β estimates later.

```
lasso = glmnet(XTrain, yTrain, alpha=1, lambda=optLambda)
plot(glmnet(XTrain, yTrain, alpha=1), xvar='lambda')
```



We see that many of the 122 coefficients are sent to zero when $\log \lambda < -4$, at which point the remaining 16 nonzero coefficients require more significant increases in $\log \lambda$ to be sent to zero, culminating in just 2 nonzero coefficients when $\log \lambda = -2$.

As with ridge regression, we now calculate the MSE of the LASSO model's predictions. This leads to a slightly better MSE of 0.0435 compared to ridge's 0.0456.

```
yLasso = predict(lasso, as.matrix(XTest))
sum((yTest - yLasso)^2)/length(yTest)
```

[1] 0.04354691

Finally, we can rescale the coefficients to obtain our optimal β estimates for both ridge and LASSO regression as below. Here we see that ridge regression leads to no zero-valued coefficients, whilst LASSO leads to 83 zero-valued coefficients (this makes sense as an increased number of zero-valued coefficients is one of the main motivations behind LASSO regression).

```
ridgeBeta = D^(-1/2) * ridge$beta
lassoBeta = D^(-1/2) * lasso$beta
sum(ridgeBeta == 0); sum(lassoBeta == 0)
```

[1] 0

[1] 83

(But as mentioned before, the number of zero-coefficients would be the same if we checked the unscaled estimates too.)

```
sum(ridge$beta == 0); sum(lasso$beta == 0)
```

[1] 0

[1] 83

Task 2

In this task we will use the Bone Mineral Density dataset¹ and the R package mgcv to fit our model. This dataset contains 485 recordings of spinal bone mineral density for male and female subjects of varying ages.

```
X0 = read.table("boneMineralDensity.dat", header=T)
head(X0)
```

```
##
     idnum
             age gender
                              spnbmd
## 1
         1 11.70
                   male 0.018080670
## 2
         1 12.70
                   male 0.060109290
         1 13.75
                   male 0.005857545
## 3
## 4
         2 13.25
                   male 0.010263930
## 5
         2 14.30
                   male 0.210526300
## 6
         2 15.30
                   male 0.040843210
dim(XO)
```

```
## [1] 485 4
```

First we split our dataset into male and female observations since we will fit separate models for each.

```
maleX = X0[X0$gender=="male",]
femaleX = X0[X0$gender=="female",]
```

To fit out models we use the function gam with the formula spnbmd ~ s(age, bs="cr"): spnbmd gives us the variable of interest, the relative change in spinal BMD, whilst s(age, bs="cr") tells us that we want to fit a spline using the age variables over a basis (bs) of cubic regression splines (cr). This function uses generalised cross validation by default.

```
library(mgcv)
```

```
## Loading required package: nlme
## This is mgcv 1.8-42. For overview type 'help("mgcv-package")'.
maleSpline = gam(spnbmd ~ s(age, bs="cr"), data=maleX) # uses gcv by default
femaleSpline = gam(spnbmd ~ s(age, bs="cr"), data=femaleX)
summary(maleSpline)
```

```
## Family: gaussian
## Link function: identity
##
## Formula:
## spnbmd \sim s(age, bs = "cr")
##
## Parametric coefficients:
              Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 0.039625
                          0.002748
                                     14.42
                                             <2e-16 ***
##
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Approximate significance of smooth terms:
           edf Ref.df
##
                           F p-value
## s(age) 4.567 5.556 15.97 <2e-16 ***
## ---
```

¹https://hastie.su.domains/ElemStatLearn/

```
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
                       Deviance explained = 29.6%
## R-sq.(adj) = 0.282
## GCV = 0.0017492 Scale est. = 0.0017061 n = 226
summary(femaleSpline)
##
## Family: gaussian
## Link function: identity
## Formula:
## spnbmd ~ s(age, bs = "cr")
##
## Parametric coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 0.038926
                         0.002176
                                   17.89
                                           <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Approximate significance of smooth terms:
##
           edf Ref.df
                          F p-value
## s(age) 6.354
                7.47 37.29 <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## R-sq.(adj) = 0.519
                        Deviance explained = 53.1%
## GCV = 0.0012617 Scale est. = 0.0012258 n = 259
```

Below we then plot the two models along with the data points with female data in red and male in blue in order to recreate Figure 5.6 from The Elements of Statistical Learning. However, note that unlike Figure 5.6, we have different values of λ for the male and female splines, as found through generalised cross validation, rather than setting $\lambda \approx 0.00022$ for both models. This has led to a slightly different plot here, though one showing largely the same behaviour as Figure 5.6.

