# Portfolio 3 - ReppArmadillo

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## **Local Polynomial Regression**

In this portfolio we'll use RcppArmadillo by looking at the following dataset on solar electricity production in Sidney, Australia<sup>1</sup>.

```
load("solarAU.RData")
head(solarAU)
```

```
## prod toy tod
## 8832 0.019 0.000000e+00 0
## 8833 0.032 5.708088e-05 1
## 8834 0.020 1.141618e-04 2
## 8835 0.038 1.712427e-04 3
## 8836 0.036 2.283235e-04 4
## 8837 0.012 2.854044e-04 5
```

Here, prod is the total production from 300 homes,  $toy \in [0,1]$  is the time-of-year and  $tod \in \{0,1,...,47\}$  is the time of day. We will be modelling the logarithm of prod, logprod, as a function of toy and tod, and to avoid problems when prod is 0, we'll add 0.01 to logprod.

```
solarAU$logprod <- log(solarAU$prod+0.01)</pre>
```

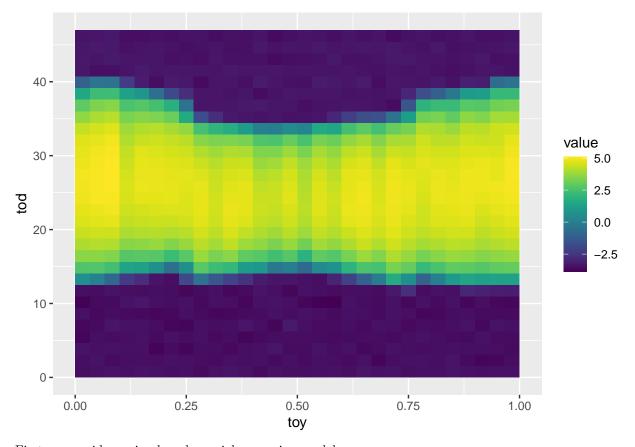
Having an initial look at the data, we see that production is, unsurprisingly, higher when there is more sunlight, i.e. in Australian winter (when tod is close to 0 or 1).

```
library(ggplot2)
library(viridis)
```

## Loading required package: viridisLite

```
ggplot(solarAU,
   aes(x = toy, y = tod, z = logprod)) +
   stat_summary_2d() +
   scale_fill_gradientn(colours = viridis(50))
```

<sup>&</sup>lt;sup>1</sup>https://www.ausgrid.com.au/Industry/Our-Research/Data-to-share/Solar-home-electricity-data



First we consider a simple polynomial regression model

$$\mathbb{E}(y|\mathbf{x}) = \beta_0, +\beta_1 \operatorname{tod} + \beta_2 \operatorname{tod}^2 + \beta_3 \operatorname{toy} + \beta_4 \operatorname{toy}^2 = \tilde{\mathbf{x}}^T \beta$$

where  $\tilde{\mathbf{x}} = \{ \text{tod}, \text{tod}^2, \text{toy}, \text{toy}^2 \}$ . This can easily be done in R as follows:

However, we want to solve this faster by using RcppArmadillo.

## Question 1

To find  $\hat{\beta} = \operatorname{argmin}_{\beta} ||\mathbf{y} - \mathbb{X}\beta||^2$  we need to compute  $\hat{\beta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$ , but the matrix inversion therein can be numerically unstable, we we'll write  $\mathbf{X}$  using the QR decomposition, that is:

$$\mathbf{X} = \mathbf{Q}\mathbf{R} = egin{bmatrix} \mathbf{Q}_1 & \mathbf{Q}_2 \end{bmatrix} egin{bmatrix} \mathbf{R}_1 \ \mathbf{0} \end{bmatrix} = \mathbf{Q}_1\mathbf{R}_1 \in \mathbb{R}^{m imes n}$$

where  $\mathbf{R}_1 \in \mathbb{R}^{n \times n}$  is upper triangular,  $\mathbf{0} \in \mathbb{R}^{(m-n) \times n}$  is a zero matrix, and both  $\mathbf{Q}_1 \in \mathbb{R}^{m \times n}$  and  $\mathbf{Q}_2 \in \mathbb{R}^{m \times (m-n)}$  have orthogonal columns. In this case our coefficients  $\hat{\beta}$  become:

$$\hat{\beta} = \mathbf{R}_1^{-1} \mathbf{Q}_1^T \mathbf{y}$$

.

Below we create our design matrix X.

```
X = with(solarAU, cbind(1, tod, tod^2, toy, toy^2))
y = solarAU$logprod
head(X); head(y)
```

```
##
         tod
                          tov
## [1,] 1
           0 0.000000e+00 0.000000e+00
## [2,] 1
           1 1 5.708088e-05 3.258227e-09
## [3,] 1
           2 4 1.141618e-04 1.303291e-08
## [4,] 1
           3 9 1.712427e-04 2.932405e-08
            4 16 2.283235e-04 5.213164e-08
## [5,] 1
## [6,] 1
            5 25 2.854044e-04 8.145568e-08
## [1] -3.540459 -3.170086 -3.506558 -3.036554 -3.079114 -3.816713
```

Next we write our Rcpp version of lm using the QR decomposition function  $qr_{econ}$ . In this we also speed up copmutation by specifying that R is upper-triangular (with trimatu(R)) and setting  $solve_{opts}$ ::fast.

```
library(Rcpp)
library(RcppArmadillo)

sourceCpp(code = '
    // [[Rcpp::depends(RcppArmadillo)]]
#include <RcppArmadillo.h>
using namespace arma;

// [[Rcpp::export(name = "lm_Rcpp")]]
vec lm_Rcpp(mat& X, vec& y){
    mat Q, R;

    qr_econ(Q, R, X);

    return solve(trimatu(R), Q.t() * y, solve_opts::fast);
}
')

fit_Rcpp = lm_Rcpp(X, y)
t(fit_Rcpp); fit$coefficients
```

```
## [,1] [,2] [,3] [,4] [,5]
## [1,] -6.262757 0.8644039 -0.01757599 -5.918069 6.142989
## (Intercept) tod I(tod^2) toy I(toy^2)
## -6.26275685 0.86440391 -0.01757599 -5.91806924 6.14298863
```

Clearly the Rcpp implementation is reaching the same results, but we can check and see that it is in fact faster than using 1m:

```
library(microbenchmark)

lm_ <- function() lm(logprod ~ tod + I(tod^2) + toy + I(toy^2), data = solarAU)

lm_Rcpp_ <- function() lm_Rcpp(X, y)

microbenchmark(lm_(), lm_Rcpp_(), times = 1000)

## Unit: microseconds

## expr min lq mean median uq max neval</pre>
```

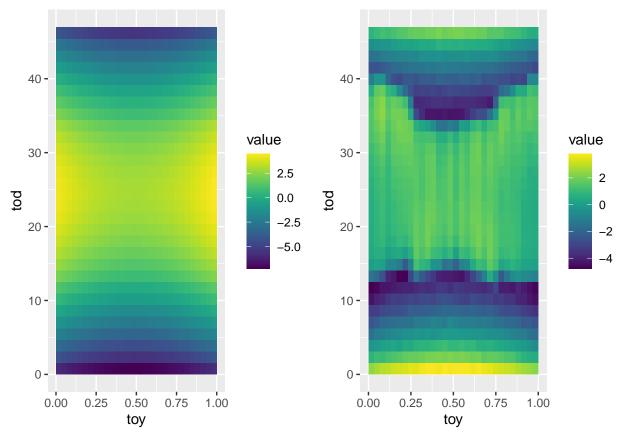
## Question 2

##

Unfortunately, looking at the polynomial fit we can see that there is a non-linear pattern left in the residuals:

1000

lm\_() 1635.450 1772.4745 2516.3336 1888.1300 2232.813 53066.581 tpp\_() 409.649 420.8865 476.1629 467.0285 503.654 1253.266



To improve our fit then, we'll use a local least regression model in which the estimated regression coefficients depend on  $\mathbf{x}$ , i.e.  $\hat{\beta} = \hat{\beta}(\mathbf{x})$ . In this case, we find  $\hat{\beta}(\mathbf{x}_0)$  as

$$\hat{\beta}(\mathbf{x}_0) = \underset{\beta}{\operatorname{argmin}} \sum_{i=1}^n \kappa_{\mathbf{H}}(\mathbf{x}_0 - \mathbf{x}_i)(y_i - \tilde{\mathbf{x}}_i^T \beta)^2$$

where  $\kappa_{\mathbf{H}}$  is a density kernel with positive definite bandwidth matrix  $\mathbf{H}$ . Since  $\kappa_{\mathbf{H}}(\mathbf{x}_0 - \mathbf{x}_i) \to 0$  as  $||\mathbf{x}_0 - \mathbf{x}_i|| \to \infty$ , we know that  $\hat{\beta}(\mathbf{x}_0)$  will depend more on data points closer to  $\mathbf{x}_0$  than those far away.

In R we can do this using the Gaussian kernel as follows:

```
library(mvtnorm)
lmLocal <- function(y, x0, X0, x, X, H){</pre>
```

```
w <- dmvnorm(x, x0, H)
fit <- lm(y ~ -1 + X, weights = w)
return( t(X0) %*% coef(fit) )
}</pre>
```

But rather than re-estimating the model for every  $\mathbf{x}_0$  (of which there are 17472), we will only use 2000 of the rows in our data, sampled uniformly at random without replacement.

nrow(solarAU)

```
## [1] 17472

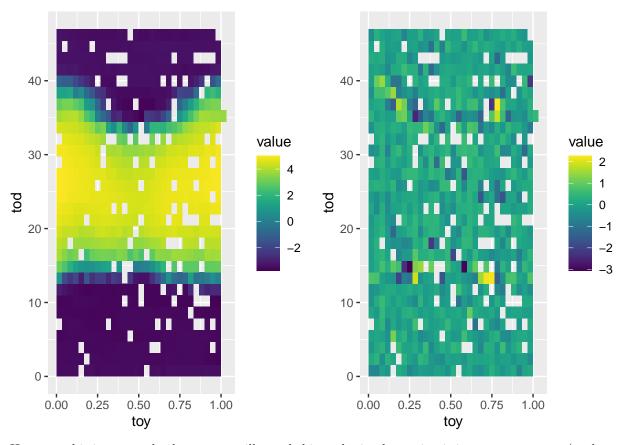
n <- nrow(X)
nsub <- 2e3
sub <- sample(1:n, nsub, replace = FALSE)

y <- solarAU$logprod
solarAU_sub <- solarAU[sub, ]
x <- as.matrix(solarAU[c("tod", "toy")])
x0 <- x[sub, ]
X0 <- X[sub, ]</pre>
```

And we obtain estimates at each of these locations as follows:

```
predLocal <- sapply(1:nsub, function(ii){
  lmLocal(y = y, x0 = x0[ii, ], X0 = X0[ii, ], x = x, X = X, H = diag(c(1, 0.1)^2))
})</pre>
```

This leads to a better fit with no clear pattern in the residuals left over.

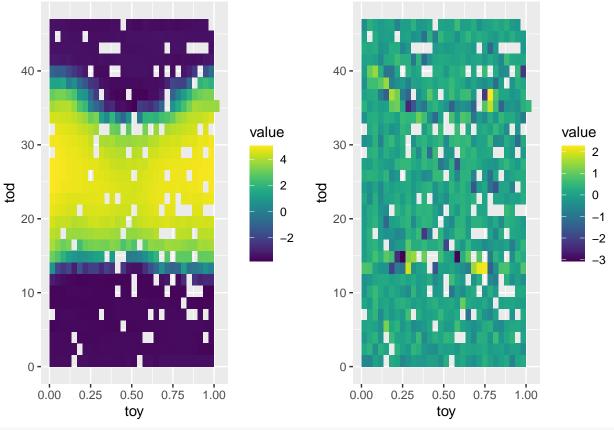


However, this is extremely slow, so we will speed this up by implementing it in RcppArmadillo (we have to implement the Gaussian kernel in here too).

```
sourceCpp(code = '
// [[Rcpp::depends(RcppArmadillo)]]
#include <RcppArmadillo.h>
using namespace arma;
vec dmvnorm_(mat & X, const rowvec & mu, mat & L){
  // Gaussian density kernel
  unsigned int m = X.n_rows;
  unsigned int n = X.n_cols;
  vec D = L.diag();
  vec out(m);
  vec z(n);
  double acc;
  unsigned int icol, irow, ii;
  for(icol = 0; icol < m; icol++){</pre>
    for(irow = 0; irow < n; irow++){</pre>
      acc = 0.0;
      for(ii = 0; ii < irow; ii++){</pre>
```

```
acc += z.at(ii) * L.at(irow, ii);
      z.at(irow) = ( X.at(icol, irow) - mu.at(irow) - acc ) / D.at(irow);
    out.at(icol) = sum(square(z));
  out = \exp(-0.5 * \text{out} - ((n / 2.0) * \log(2.0 * M_PI) + \sup(\log(D))));
 return out;
vec lm(mat X, vec y){
 // fit lm with QR decomp as before
 mat Q, R;
 qr_econ(Q, R, X);
 return solve(trimatu(R), Q.t() * y, solve_opts::fast);
// [[Rcpp::export(name = "lm_local_Rcpp")]]
vec lm_local(vec& y, mat& x0, mat& X0, mat& x, mat& X, mat& H){
  // find coefficients for each row and weight by the Gaussian density kernel
  int nsub = x0.n rows;
  vec out(nsub), weights;
  double fit;
 mat L = chol(H, "lower");
 for(int i=0; i<nsub; i++){</pre>
    weights = sqrt(dmvnorm_(x, x0.row(i), L));
   fit = as_scalar(X0.row(i) * lm(X.each_col() % weights, y % weights));
    out(i) = fit;
 return out;
}
```

Fitting this RcppArmadillo version we obtain the same results as before.



all.equal(as.vector(predLocal\_Rcpp), predLocal)

#### ## [1] TRUE

But, as expected, the RcppArmadillo version is much faster.

```
library(microbenchmark)

local_R <- function() {
  predLocal <- sapply(1:nsub, function(ii){
    lmLocal(y = y, x0 = x0[ii, ], X0 = X0[ii, ], x = x, X = X, H = diag(c(1, 0.1)^2))
    })
}

local_C <- function() lm_local_Rcpp(y, x0, X0, x, X, diag(c(1, 0.1)^2))
  microbenchmark(local_R(), local_C(), times=10)</pre>
```

```
## Unit: seconds
## expr min lq mean median uq max neval
## local_R() 11.196285 11.37293 11.495952 11.538341 11.656982 11.713108 10
## local_C() 1.738081 1.78699 2.754891 2.132787 4.121924 4.146339 10
```

## Question 3

Finally, we can implement a cross-validation routine to choose the bandwidth matrix  $\mathbf{H}$  (which we will keep as diagonal) by minimising the residual sum of squares. In particular, we'll test all combinations of the diagonal elements of  $\mathbf{H}$  being in  $\{0.01, 0.1, 1, 10, 100\}$ .

```
H_elems = c(0.01, 0.1, 1, 10, 100)
opt_RSS = Inf
opt_H_elems = c(NA, NA)

for (H_11 in H_elems){
    for (H_22 in H_elems){
        fit = lm_local_Rcpp(y, x0, X0, x, X, diag(c(H_11, H_22)))
        RSS = sum((fit - solarAU_sub$logprod)^2)
        if (RSS < opt_RSS){
            opt_RSS = RSS
            opt_H_elems = c(H_11, H_22)
        }
    }
}</pre>
```

opt\_H\_elems

```
## [1] 0.01 0.01
```

Here we find that the optimal bandwidth matrix we tried was  $\mathbf{H} = \begin{pmatrix} 0.01 & 0 \\ 0 & 0.01 \end{pmatrix}$ , which gives us the following fit and residuals:

