

Multivariate distributions

Moment generating functions

Definition: The **joint moment generating function** of (X_1, \dots, X_n) is

$$M_{X_1, \dots, X_n}(\underbrace{t_1, \dots, t_n}_{\text{def}}) = \mathbb{E} e^{t_1 X_1 + \dots + t_n X_n}, \quad t_1, \dots, t_n \in \mathbb{R}$$

if the expectation exists for all $-h < t_1, \dots, t_n < h$ for some $h > 0$

- Joint mgf can provide univariate mgfs

$$\underbrace{M_{X_i}(t_i)} = M_{X_1, \dots, X_n}(t_1 = 0, \dots, t_{i-1} = 0, t_i, t_{i+1} = 0, \dots, t_n = 0)$$

- Applications as before (more later):

- characterizes distributions

e.g., if (X_1, \dots, X_n) and (Y_1, \dots, Y_n) have the same mgfs, then these vectors have the same distribution

- transformations

e.g., mgf of $(a_1 X_1, \dots, a_n X_n)$ is $M_{X_1, \dots, X_n}(a_1 t_1, \dots, a_n t_n)$

- convergence (later)

- moments, e.g.,

$$\mathbb{E}(X_i^q X_j^r X_k^s) = \frac{\partial^{q+r+s}}{\partial t_i^q \partial t_j^r \partial t_k^s} M_{X_1, \dots, X_n}(t_1, \dots, t_n) \Big|_{(t_1, \dots, t_n) = (0, \dots, 0)}$$

$g=1, r=1 \quad \mathbb{E}(XY) = \frac{\partial^2}{\partial t_1 \partial t_2} M_{X,Y}(t_1, t_2) \Big|_{(t_1, t_2) = (0, 0)}$

$$(X, Y) \sim \text{Multinomial}(n, \underbrace{\frac{e^{t_1} p_1}{S}}_{\text{p}_1}, \underbrace{\frac{e^{t_2} p_2}{S}}_{\text{p}_2}) \quad , \quad f(x, y) = \frac{n!}{x! y! (n-x-y)!} \left(\frac{e^{t_1} p_1}{S} \right)^x \left(\frac{e^{t_2} p_2}{S} \right)^y \underbrace{\left(\frac{1 - (e^{t_1} p_1) - (e^{t_2} p_2)}{S} \right)^{n-x-y}}_{\text{p}_3}$$

Multivariate distributions

Moment generating functions: example

In n trials

$X := \#$ of outcomes "a"

$Y := \#$ of outcomes "b"

$n-x-y = \#$ of outcomes "c"

Let $(X, Y) \sim \text{Multinomial}(n, p_1, p_2)$, so the joint pmf of (X, Y) is

$$\sum_x \sum_y f(x, y) = 1$$

$$f(x, y) = \frac{n!}{x! y! (n-x-y)!} \underbrace{p_1^x p_2^y (1-p_1-p_2)^{n-x-y}}_{p_1 + p_2 + (1-p_1-p_2) = 1}$$

for $0 \leq x, y, x+y \leq n$

$$M_{X,Y}(t_1, t_2) \stackrel{\text{def}}{=} \mathbb{E} e^{t_1 X + t_2 Y} = \mathbb{E}(h(X, Y)) = \sum_x \sum_y h(x, y) f_{X,Y}(x, y)$$

$$x+y \leq n$$

$$= \sum_{x=0}^n \sum_{y=0}^{n-x} \underbrace{e^{t_1 x + t_2 y}}_{h(x, y)} \underbrace{\frac{n!}{x! y! (n-x-y)!} p_1^x p_2^y (1-p_1-p_2)^{n-x-y}}_{f_{X,Y}(x, y)}$$

$$= \sum_{x=0}^n \sum_{y=0}^{n-x} \frac{n!}{x! y! (n-x-y)!} \underbrace{(e^{t_1} p_1)^x (e^{t_2} p_2)^y (1-p_1-p_2)^{n-x-y}}_{S^n}$$

$$\frac{e^{t_1} p_1 + e^{t_2} p_2 + (1-p_1-p_2)}{S} \neq 1$$

$$\frac{e^{t_1} p_1}{S} + \frac{e^{t_2} p_2}{S} + \frac{1-p_1-p_2}{S} = 1$$

$$S = e^{t_1} p_1 + e^{t_2} p_2 + (1-p_1-p_2)$$

$$= S^n \sum_x \sum_y \frac{n!}{x! y! (n-x-y)!} \left(\frac{e^{t_1} p_1}{S} \right)^x \left(\frac{e^{t_2} p_2}{S} \right)^y \left(\frac{1-p_1-p_2}{S} \right)^{n-x-y}$$

pmf of $(n, \frac{e^{t_1} p_1}{S}, \frac{e^{t_2} p_2}{S})$

$$= S^n \left[e^{t_1} p_1 + e^{t_2} p_2 + (1-p_1-p_2) \right]^n$$

$$1 = \frac{e^{t_1} p_1}{S} + \frac{e^{t_2} p_2}{S} + \frac{1-p_1-p_2}{S} = \frac{e^{t_1} p_1 + e^{t_2} p_2 + (1-p_1-p_2)}{S}$$

$$M_{X,Y}(t_1, t_2) = S^n = \left[e^{t_1} p_1 + e^{t_2} p_2 + 1 - p_1 - p_2 \right]^n$$

Main idea was to consider

$$S^n \left(\frac{p_1 e^{t_1}}{S} \right)^x \left(\frac{p_2 e^{t_2}}{S} \right)^y \left(\frac{1 - p_1 - p_2}{S} \right)^{n-x-y}$$

$$S := p_1 e^{t_1} + p_2 e^{t_2} + 1 - p_1 - p_2$$

$$\begin{aligned} M_X(t_1) &= M_{X,Y}(t_1, 0) = \left(e^{t_1} p_1 + \cancel{e^{t_2} p_2} + 1 - p_1 - p_2 \right)^n \\ &= \left(e^{t_1} p_1 + \cancel{p_2} + 1 - p_1 - \cancel{p_2} \right)^n \\ &= \left(e^{t_1} p_1 + 1 - p_1 \right)^n \Rightarrow \end{aligned}$$

$$M_Y(t_2) = M_{X,Y}(0, t_2) = \left(e^{t_2} p_2 + 1 - p_2 \right)^n$$

$$\text{Cov}(X, Y) = \underbrace{\mathbb{E}XY}_{n p_1 p_2} - \underbrace{\mathbb{E}X}_{n p_1} \underbrace{\mathbb{E}Y}_{n p_2}$$

$$\mathbb{E}XY = \frac{\partial^2}{\partial t_1 \partial t_2} M_{X,Y}(t_1, t_2) \Big|_{(t_1, t_2) = (0, 0)} = n(n-1) p_1 p_2$$

$$\text{Cov}(X, Y) = n(n-1) p_1 p_2 - n^2 p_1 p_2 = -n p_1 p_2$$

Conditional distributions

Introduction

Recall $P(A|B)$ is the probability that A occurs given that B occurs:

$$\checkmark \quad P(A|B) = \frac{P(A \cap B)}{P(B)} \quad \text{for } P(B) > 0$$

We want to apply this idea to random variables

Already seen one example: truncated distributions, e.g, fix x_0

$$P(\underbrace{X \leq x}_A | \underbrace{X > x_0}_B) = \frac{P(\overbrace{X \leq x}^A, \overbrace{X > x_0}^B)}{P(\underbrace{X > x_0}_B)} = \frac{P(x_0 < X \leq x)}{P(X > x_0)} = \frac{F(x) - F(x_0)}{1 - F(x_0)} \quad x > x_0$$

$P(A|B) = \frac{P(A \cap B)}{P(B)}$

Definitions: Conditional distribution (given general event A)

Suppose we observe A with $P(A) > 0$

1. conditional cdf

$$F(x|A) = P(\underbrace{X \leq x}_A | A) = \frac{P(A, X \leq x)}{P(A)}$$

$$F_X(x) = P(X \leq x)$$

$$F_X(x|A) = P(X \leq x | A)$$

2. conditional pmf/pdf

$$f(x|A) = P(X = x|A) = \frac{P(A, X = x)}{P(A)}, \quad x \in \mathbb{R} \quad \text{pmf if } X \text{ is discrete}$$

$$f(x|A) = \frac{dF(x|A)}{dx}, \quad x \in \mathbb{R} \quad \text{pdf if } X \text{ is continuous}$$

For bivariate (X, Y) , we're interested in conditional pmf/pdf $f(x|Y = y)$