

In this case,

$$\mathbf{G} = \text{Var}(\mathbf{u}) = \sigma_s^2 \mathbf{I}_{n. \times n.} \text{ and}$$

$$\mathbf{R} = \text{Var}(\mathbf{e}) = \sigma_e^2 \mathbf{I}_{(7n.) \times (7n.)},$$

where $n. = n_1 + n_2 + n_3$ is the total number of subjects.

$$\mathbf{\Sigma} = \text{Var}(\mathbf{y}) = \mathbf{ZGZ}^\top + \mathbf{R}$$

is a block diagonal matrix with one block of the form

$$\sigma_s^2 \mathbf{1}\mathbf{1}_{7 \times 7}^\top + \sigma_e^2 \mathbf{I}_{7 \times 7}$$

for each subject.

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{z}_u + \mathbf{e}$$

$$\mathbf{G} = \text{Var}(\mathbf{u}) = \sigma_s^2 \mathbf{I}, \quad \mathbf{R} = \text{Var}(\mathbf{e}) = \sigma_e^2 \mathbf{I}, \text{ and}$$

$$\begin{aligned} \boldsymbol{\Sigma} &= \text{Var}(\mathbf{y}) = \mathbf{Z}\mathbf{G}\mathbf{Z}^\top + \mathbf{R} \\ &= \sigma_s^2 \begin{bmatrix} \mathbf{1}\mathbf{1}^\top & & & \\ & \mathbf{1}\mathbf{1}^\top & & \\ & & \ddots & \\ & & & \mathbf{1}\mathbf{1}^\top \end{bmatrix} + \sigma_e^2 \mathbf{I} \\ &= \begin{bmatrix} \underline{\sigma_e^2 \mathbf{I} + \sigma_s^2 \mathbf{1}\mathbf{1}^\top} & & & \\ & \ddots & & \\ & & \ddots & \\ & & & \sigma_e^2 \mathbf{I} + \sigma_s^2 \mathbf{1}\mathbf{1}^\top \end{bmatrix}. \end{aligned}$$

If predicting subject effects is not of interest and random subject effects are included only to introduce correlation among repeated measures on the same subject, we can work with an alternative expression of the same model by using the general linear model $\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}$, where

$$\text{Var}(\mathbf{y}) = \text{Var}(\boldsymbol{\epsilon}) = \begin{bmatrix} \sigma_e^2 \mathbf{I} + \sigma_s^2 \mathbf{1}\mathbf{1}^\top & & \\ & \ddots & \\ & & \sigma_e^2 \mathbf{I} + \sigma_s^2 \mathbf{1}\mathbf{1}^\top \end{bmatrix}.$$

More generally, we can replace the mixed model

$$y = X\beta + \underbrace{Zu + e}_{\epsilon}$$

with the model

$$y = X\beta + \epsilon$$

where

$$\text{Var}(\mathbf{y}) = \text{Var}(\epsilon) = \begin{bmatrix} \mathbf{W} & \mathbf{0} & \cdots & \mathbf{0} \\ \mathbf{0} & \mathbf{W} & \cdots & \mathbf{0} \\ \vdots & \vdots & \ddots & \vdots \\ \mathbf{0} & \mathbf{0} & \cdots & \mathbf{W} \end{bmatrix}.$$

generally we assume \mathbf{W} are identical across all subjects

- We can choose a structure for W that seems appropriate based on the design and the data.

- One choice for W is a compound symmetric matrix like we have considered previously. *requires estimation of σ_s^2 & σ_e^2*

- Another choice for W is an unstructured positive definite matrix. *requires a much larger # of parameters to be estimated.*

- A common choice for W when repeated measures are equally spaced in time is the first-order autoregressive structure known as AR(1).

AR(1): First-Order Autoregressive Covariance Structure

same σ^2 across time

$$W = \sigma^2 \begin{bmatrix} 1 & \phi & \phi^2 & \phi^3 & \phi^4 & \phi^5 & \phi^6 \\ \phi & 1 & \phi & \phi^2 & \phi^3 & \phi^4 & \phi^5 \\ \phi^2 & \phi & 1 & \phi & \phi^2 & \phi^3 & \phi^4 \\ \phi^3 & \phi^2 & \phi & 1 & \phi & \phi^2 & \phi^3 \\ \phi^4 & \phi^3 & \phi^2 & \phi & 1 & \phi & \phi^2 \\ \phi^5 & \phi^4 & \phi^3 & \phi^2 & \phi & 1 & \phi \\ \phi^6 & \phi^5 & \phi^4 & \phi^3 & \phi^2 & \phi & 1 \end{bmatrix}$$

measurements further apart are less correlated

where $\sigma^2 \in (0, \infty)$ and $\phi \in (-1, 1)$ are unknown parameters.

- In the next slides, we will see how to fit a variety of general linear models that might be appropriate for repeated measures experiments.
- These slides illustrate a few example R commands for fitting general linear models to repeated measures data.
- We focus on the experiment designed to compare the effectiveness of three strength training programs.
- We will fit models that allows for a distinct mean for each of the $3 \times 7 = 21$ combinations of training program and time.

- We assume independence between subjects.
- The models differ in the choice for \mathbf{W} , which is the variance-covariance structure assumed for the 7 observations from each subject.


```
#Read the data
```

```
d=read.delim(  
  "http://dnett.github.io/S510/RepeatedMeasures.txt")
```

```
#Create Factors
```

```
d$Program = factor(d$Program)  
d$Subj = factor(d$Subj)  
d$Timef = factor(d$Time)
```

```
#Load the nlme package
```

```
library(nlme)
```

Compound Symmetry Structure for W

$$X = X\beta + \underline{z}_u + \underline{e}$$

$$\begin{bmatrix} \sigma_e^2 + \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 \\ \sigma_s^2 & \sigma_e^2 + \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 \\ \sigma_s^2 & \sigma_s^2 & \sigma_e^2 + \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 \\ \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_e^2 + \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 \\ \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_e^2 + \sigma_s^2 & \sigma_s^2 & \sigma_s^2 \\ \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_e^2 + \sigma_s^2 & \sigma_s^2 \\ \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_e^2 + \sigma_s^2 \end{bmatrix}$$

```
o.lme = lme(Strength ~ Program * Timef, data = d,
            random = ~ 1 | Subj)
```

```
> summary(o.lme)
```

Linear mixed-effects model fit by REML

Data: d

AIC	BIC	logLik
1466.82	1557.323	-710.4101

Random effects:

Formula: ~1 | Subj

(Intercept) Residual

StdDev: 3.098924 1.094017

↑
 $\hat{\sigma}_s$

↑
 $\hat{\sigma}_e$

on main diagonal
of W

$$(3.099)^2 + (1.094)^2 \\ = 10.8$$

$$\hat{\sigma}_s^2 = (3.099)^2 \approx 9.6$$

```
> # Examine the estimated variance-covariance
> # matrix for the subvector of responses
> # from a single subject.
>
> getVarCov(o.lme, individuals = 1, type = "marginal")
```

Subj 1

Marginal variance covariance matrix

	1	2	3	4	5	6	7
1	10.8000	9.6033	9.6033	9.6033	9.6033	9.6033	9.6033
2	9.6033	10.8000	9.6033	9.6033	9.6033	9.6033	9.6033
3	9.6033	9.6033	10.8000	9.6033	9.6033	9.6033	9.6033
4	9.6033	9.6033	9.6033	10.8000	9.6033	9.6033	9.6033
5	9.6033	9.6033	9.6033	9.6033	10.8000	9.6033	9.6033
6	9.6033	9.6033	9.6033	9.6033	9.6033	10.8000	9.6033
7	9.6033	9.6033	9.6033	9.6033	9.6033	9.6033	10.8000

Alternative Parameterization for Compound Symmetry

$$Y = X\beta + \epsilon$$

$$\begin{aligned}\sigma^2 &= \text{Var}(y) \\ &= \text{Var}(\epsilon)\end{aligned}$$

$$\begin{bmatrix} 1 & \rho & \rho & \rho & \rho & \rho & \rho \\ \rho & 1 & \rho & \rho & \rho & \rho & \rho \\ \rho & \rho & 1 & \rho & \rho & \rho & \rho \\ \rho & \rho & \rho & 1 & \rho & \rho & \rho \\ \rho & \rho & \rho & \rho & 1 & \rho & \rho \\ \rho & \rho & \rho & \rho & \rho & 1 & \rho \\ \rho & \rho & \rho & \rho & \rho & \rho & 1 \end{bmatrix} = W$$

```
o.cs = gls(Strength ~ Program * Timef, data = d,  
correlation = corCompSymm(form = ~ 1 | Subj))
```

```
> summary(o.cs)
```

Generalized least squares fit by REML

Model: Strength ~ Program * Timef

Data: d

AIC	BIC	logLik
<u>1466.82</u>	<u>1557.323</u>	<u>-710.4101</u>

Correlation Structure: Compound symmetry

Formula: ~1 | Subj

Parameter estimate(s):

Rho
0.8891805 - $\hat{\rho}$

.
.
.

Residual standard error: 3.286366

$\hat{\sigma}_e$

Degrees of freedom: 399 total; 378 residual

```
> getVarCov(o.cs)
```

```
Marginal variance covariance matrix
```

	[,1]	[,2]	[,3]	[,4]	[,5]	[,6]	[,7]
[1,]	10.8000	9.6033	9.6033	9.6033	9.6033	9.6033	9.6033
[2,]	9.6033	10.8000	9.6033	9.6033	9.6033	9.6033	9.6033
[3,]	9.6033	9.6033	10.8000	9.6033	9.6033	9.6033	9.6033
[4,]	9.6033	9.6033	9.6033	10.8000	9.6033	9.6033	9.6033
[5,]	9.6033	9.6033	9.6033	9.6033	10.8000	9.6033	9.6033
[6,]	9.6033	9.6033	9.6033	9.6033	9.6033	10.8000	9.6033
[7,]	9.6033	9.6033	9.6033	9.6033	9.6033	9.6033	10.8000

AR(1) Structure for W

$$\sigma^2 \begin{bmatrix} 1 & \phi & \phi^2 & \phi^3 & \phi^4 & \phi^5 & \phi^6 \\ \phi & 1 & \phi & \phi^2 & \phi^3 & \phi^4 & \phi^5 \\ \phi^2 & \phi & 1 & \phi & \phi^2 & \phi^3 & \phi^4 \\ \phi^3 & \phi^2 & \phi & 1 & \phi & \phi^2 & \phi^3 \\ \phi^4 & \phi^3 & \phi^2 & \phi & 1 & \phi & \phi^2 \\ \phi^5 & \phi^4 & \phi^3 & \phi^2 & \phi & 1 & \phi \\ \phi^6 & \phi^5 & \phi^4 & \phi^3 & \phi^2 & \phi & 1 \end{bmatrix}$$

```
o.ar1 = gls(Strength ~ Program * Timef, data = d,  
            correlation = corAR1(form = ~ 1 | Subj))
```



```
> summary(o.ar1)
```

Generalized least squares fit by REML

Model: Strength ~ Program * Timef

Data: d

AIC	BIC	logLik
1312.804	1403.306	-633.4018

Correlation Structure: AR(1)

Formula: ~1 | Subj

Parameter estimate(s):

Phi	$\hat{\Phi}$
0.9517769	

.
.
.

Residual standard error: 3.280242

Degrees of freedom: 399 total; 378 residual

```
> getVarCov(o.ar1, individual = 3)
```

Marginal variance covariance matrix

	[,1]	[,2]	[,3]	[,4]	[,5]	[,6]	[,7]
[1,]	10.7600	10.2410	9.7473	9.2772	8.8298	8.4040	7.9988
[2,]	10.2410	10.7600	10.2410	9.7473	9.2772	8.8298	8.4040
[3,]	9.7473	10.2410	10.7600	10.2410	9.7473	9.2772	8.8298
[4,]	9.2772	9.7473	10.2410	10.7600	10.2410	9.7473	9.2772
[5,]	8.8298	9.2772	9.7473	10.2410	10.7600	10.2410	9.7473
[6,]	8.4040	8.8298	9.2772	9.7473	10.2410	10.7600	10.2410
[7,]	7.9988	8.4040	8.8298	9.2772	9.7473	10.2410	10.7600

← getting smaller

General Positive Definite Structure for W

With δ_1 set equal to 1 for identifiability purposes, a general 7×7 positive definite variance-covariance matrix is parameterized by R as follows:

$\sigma^2 \text{diag}(\delta_1, \dots, \delta_7)$

allow for different $\text{Var}(y)$ across time

1	ρ_{12}	ρ_{13}	ρ_{14}	ρ_{15}	ρ_{16}	ρ_{17}
ρ_{12}	1	ρ_{23}	ρ_{24}	ρ_{25}	ρ_{26}	ρ_{27}
ρ_{13}	ρ_{23}	1	ρ_{34}	ρ_{35}	ρ_{36}	ρ_{37}
ρ_{14}	ρ_{24}	ρ_{34}	1	ρ_{45}	ρ_{46}	ρ_{47}
ρ_{15}	ρ_{25}	ρ_{35}	ρ_{45}	1	ρ_{56}	ρ_{57}
ρ_{16}	ρ_{26}	ρ_{36}	ρ_{46}	ρ_{56}	1	ρ_{67}
ρ_{17}	ρ_{27}	ρ_{37}	ρ_{47}	ρ_{57}	ρ_{67}	1

most flexibility in terms correlation between any observations

$\text{diag}(\delta_1, \dots, \delta_7)$

over time within the same individual

The 7×7 case doesn't fit on one slide, but here is the 5×5 case.

$$\sigma^2 \rho_{ij} \delta_i \delta_j$$

$$\begin{bmatrix} \sigma^2 \delta_1^2 & \sigma^2 \rho_{12} \delta_1 \delta_2 & \sigma^2 \rho_{13} \delta_1 \delta_3 & \sigma^2 \rho_{14} \delta_1 \delta_4 & \sigma^2 \rho_{15} \delta_1 \delta_5 \\ \sigma^2 \rho_{12} \delta_1 \delta_2 & \sigma^2 \delta_2^2 & \sigma^2 \rho_{23} \delta_2 \delta_3 & \sigma^2 \rho_{24} \delta_2 \delta_4 & \sigma^2 \rho_{25} \delta_2 \delta_5 \\ \sigma^2 \rho_{13} \delta_1 \delta_3 & \sigma^2 \rho_{23} \delta_2 \delta_3 & \sigma^2 \delta_3^2 & \sigma^2 \rho_{34} \delta_3 \delta_4 & \sigma^2 \rho_{35} \delta_3 \delta_5 \\ \sigma^2 \rho_{14} \delta_1 \delta_4 & \sigma^2 \rho_{24} \delta_2 \delta_4 & \sigma^2 \rho_{34} \delta_3 \delta_4 & \sigma^2 \delta_4^2 & \sigma^2 \rho_{45} \delta_4 \delta_5 \\ \sigma^2 \rho_{15} \delta_1 \delta_5 & \sigma^2 \rho_{25} \delta_2 \delta_5 & \sigma^2 \rho_{35} \delta_3 \delta_5 & \sigma^2 \rho_{45} \delta_4 \delta_5 & \sigma^2 \delta_5^2 \end{bmatrix}$$

```
o.un = gls(Strength ~ Program * Timef, data = d,
           correlation = corSymm(form = ~ 1 | Subj),
           weight = varIdent(form = ~ 1 | Timef))
```

allows variance to change over time

```
> summary(o.un)
```

```
Generalized least squares fit by REML
```

```
Model: Strength ~ Program * Timef
```

```
Data: d
```

	AIC	BIC	logLik
	1332.896	1525.706	-617.4479

Correlation Structure: General

Formula: ~1 | Subj

Parameter estimate(s):

Correlation:

	1	2	3	4	5	6
2	0.960					
3	0.925	0.940				
4	0.872	0.877	0.956			
5	0.842	0.860	0.937	0.960		
6	0.809	0.827	0.898	0.909	0.951	
7	0.797	0.792	0.876	0.887	0.917	0.953

lower triangular part of W

similarly to AR(1)
observations further
apart in time
are less
correlated!

Variance function:

Structure: Different standard deviations per stratum

Formula: ~1 | Timef

Parameter estimates:

	2	4	6	8	10	12	14
	1.000	1.039	1.104	1.071	1.174	1.157	1.203
·	$\hat{\sigma}_1$	$\hat{\sigma}_2$	$\hat{\sigma}_3$.	.	.	$\hat{\sigma}_7$
·				.	.	.	
·							

Residual standard error: 2.963129

Degrees of freedom: 399 total; 378 residual

```
> getVarCov(o.un, individual = 3)
```

Marginal variance covariance matrix

	[,1]	[,2]	[,3]	[,4]	[,5]	[,6]	[,7]
[1,]	8.7801	8.7571	8.9656	8.1984	8.6781	8.2203	8.4169
[2,]	8.7571	9.4730	9.4631	8.5686	9.2012	8.7307	8.6875
[3,]	8.9656	9.4631	10.7080	9.9266	10.6660	10.0700	10.2140
[4,]	8.1984	8.5686	9.9266	10.0770	10.6000	9.8987	10.0430
[5,]	8.6781	9.2012	10.6660	10.6000	12.0950	11.3440	11.3640
[6,]	8.2203	8.7307	10.0700	9.8987	11.3440	11.7560	11.6500
[7,]	8.4169	8.6875	10.2140	10.0430	11.3640	11.6500	12.7100

- To understand the reason for an identifiability constraint, notice that an arbitrary positive definite 7×7 covariance matrix depends on only

$$7 + 6 + 5 + 4 + 3 + 2 + 1 = \frac{7(7+1)}{2} = \underline{28}$$

parameters. However, we have σ^2 , $6 + 5 + 4 + 3 + 2 + 1 = 21$ ρ parameters, and $\delta_1, \dots, \delta_7$.

- That's 29 parameters for a symmetric positive definite matrix that depends on at most 28 parameters.

to ensure identifiability R set $\delta_i = 1$

- Thus, R chooses to set δ_1 to 1.
- Without such a constraint, it is easy to use different values of the parameters to define the same matrix. For example,

2 time
points

$$\begin{bmatrix} 3 & -1 \\ -1 & 7 \end{bmatrix} = 3 \begin{bmatrix} 1 & -\frac{1}{3} \\ -\frac{1}{3} & \frac{7}{3} \end{bmatrix} = 1 \begin{bmatrix} 3 & -1 \\ -1 & 7 \end{bmatrix}$$

$$\begin{array}{c} \underline{\sigma^2} \\ \underline{\delta_1} \\ \underline{\delta_2} \\ \underline{\rho_{12}} \end{array}$$

$$\begin{array}{c} 3 \\ 1 \\ \sqrt{\frac{7}{3}} \\ \frac{-1}{3\sqrt{\frac{7}{3}}} \end{array}$$

$$\begin{array}{c} 1 \\ \sqrt{3} \\ \sqrt{7} \\ \frac{-1}{\sqrt{21}} \end{array}$$

constraint

end
lecture

05-02-25

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