

In this case,

$$\mathbf{G} = \text{Var}(\mathbf{u}) = \sigma_s^2 \mathbf{I}_{n. \times n.} \text{ and}$$

$$\mathbf{R} = \text{Var}(\mathbf{e}) = \sigma_e^2 \mathbf{I}_{(7n.) \times (7n.)},$$

where $n. = n_1 + n_2 + n_3$ is the total number of subjects.

$$\Sigma = \text{Var}(\mathbf{y}) = \mathbf{ZGZ}^\top + \mathbf{R}$$

is a block diagonal matrix with one block of the form

$$\sigma_s^2 \mathbf{1} \mathbf{1}_{7 \times 7}^\top + \sigma_e^2 \mathbf{I}_{7 \times 7}$$

for each subject.

$$\mathbf{y} = X\beta + \mathbf{z} + \mathbf{e}$$

$\mathbf{G} = \text{Var}(\mathbf{u}) = \sigma_s^2 \mathbf{I}$, $\mathbf{R} = \text{Var}(\mathbf{e}) = \sigma_e^2 \mathbf{I}$, and

$$\begin{aligned}\Sigma &= \text{Var}(\mathbf{y}) = \mathbf{ZGZ}^\top + \mathbf{R} \\ &= \sigma_s^2 \begin{bmatrix} \mathbf{1}\mathbf{1}^\top & & & \\ & \mathbf{1}\mathbf{1}^\top & & \\ & & \ddots & \\ & & & \mathbf{1}\mathbf{1}^\top \end{bmatrix} + \sigma_e^2 \mathbf{I} \\ &= \begin{bmatrix} \underline{\sigma_e^2 \mathbf{I} + \sigma_s^2 \mathbf{1}\mathbf{1}^\top} & & & \\ & \ddots & & \\ & & \sigma_e^2 \mathbf{I} + \sigma_s^2 \mathbf{1}\mathbf{1}^\top & \end{bmatrix}.\end{aligned}$$

If predicting subject effects is not of interest and random subject effects are included only to introduce correlation among repeated measures on the same subject, we can work with an alternative expression of the same model by using the general linear model $\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}$, where

$$\text{Var}(\mathbf{y}) = \text{Var}(\boldsymbol{\epsilon}) = \begin{bmatrix} \sigma_e^2 \mathbf{I} + \sigma_s^2 \mathbf{1}\mathbf{1}^\top & & & \\ & \ddots & & \\ & & \ddots & \\ & & & \sigma_e^2 \mathbf{I} + \sigma_s^2 \mathbf{1}\mathbf{1}^\top \end{bmatrix}.$$

More generally, we can replace the mixed model

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \underbrace{\mathbf{Z}\mathbf{u} + \mathbf{e}}$$

with the model

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{\epsilon}$$

where

$$\text{Var}(\mathbf{y}) = \text{Var}(\mathbf{\epsilon}) = \begin{bmatrix} \mathbf{W} & \mathbf{0} & \cdots & \mathbf{0} \\ \mathbf{0} & \mathbf{W} & \cdots & \mathbf{0} \\ \vdots & \vdots & \ddots & \vdots \\ \mathbf{0} & \mathbf{0} & \cdots & \mathbf{W} \end{bmatrix}.$$

generally we assume \mathbf{W} are identical across
all subjects

- We can choose a structure for W that seems appropriate based on the design and the data.

requires estimation of σ_s^2 & σ_e^2

- One choice for W is a compound symmetric matrix like we have considered previously.

requires a much larger # of parameters to be estimated.

- Another choice for W is an unstructured positive definite matrix.

- A common choice for W when repeated measures are equally spaced in time is the first-order autoregressive structure known as AR(1).

AR(1): First-Order Autoregressive Covariance Structure

same σ^2 across time

$$W = \sigma^2 \begin{bmatrix} 1 & \phi & \phi^2 & \phi^3 & \phi^4 & \phi^5 & \phi^6 \\ \phi & 1 & \phi & \phi^2 & \phi^3 & \phi^4 & \phi^5 \\ \phi^2 & \phi & 1 & \phi & \phi^2 & \phi^3 & \phi^4 \\ \phi^3 & \phi^2 & \phi & 1 & \phi & \phi^2 & \phi^3 \\ \phi^4 & \phi^3 & \phi^2 & \phi & 1 & \phi & \phi^2 \\ \phi^5 & \phi^4 & \phi^3 & \phi^2 & \phi & 1 & \phi \\ \phi^6 & \phi^5 & \phi^4 & \phi^3 & \phi^2 & \phi & 1 \end{bmatrix}$$

measurements further apart are less correlated

where $\sigma^2 \in (0, \infty)$ and $\phi \in (-1, 1)$ are unknown parameters.

- In the next slides, we will see how to fit a variety of general linear models that might be appropriate for repeated measures experiments.
- These slides illustrate a few example R commands for fitting general linear models to repeated measures data.
- We focus on the experiment designed to compare the effectiveness of three strength training programs.
- We will fit models that allows for a distinct mean for each of the $3 \times 7 = 21$ combinations of training program and time.

- We assume independence between subjects.
- The models differ in the choice for W , which is the variance-covariance structure assumed for the 7 observations from each subject.

```
#Read the data

d=read.delim(
  "http://dnett.github.io/S510/RepeatedMeasures.txt")

#Create Factors

d$Program = factor(d$Program)
d$Subj = factor(d$Subj)
d$Timef = factor(d$Time)

#Load the nlme package

library(nlme)
```

Compound Symmetry Structure for W

$$Y = X\beta + \epsilon_u + \epsilon_e$$

$$\begin{bmatrix} \sigma_e^2 + \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 \\ \sigma_s^2 & \sigma_e^2 + \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 \\ \sigma_s^2 & \sigma_s^2 & \sigma_e^2 + \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 \\ \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_e^2 + \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 \\ \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_e^2 + \sigma_s^2 & \sigma_s^2 & \sigma_s^2 \\ \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_e^2 + \sigma_s^2 & \sigma_s^2 \\ \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_e^2 + \sigma_s^2 & \sigma_s^2 \\ \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_s^2 & \sigma_e^2 + \sigma_s^2 \end{bmatrix}$$

```
o.lme = lme(Strength ~ Program * Time, data = d,  
random = ~ 1 | Subj)
```

```
> summary(o.lme)
```

Linear mixed-effects model fit by REML

Data: d

AIC	BIC	logLik
1466.82	1557.323	-710.4101

on main diagonal

Random effects:

Formula: ~1 | Subj

(Intercept) Residual

StdDev: 3.098924 1.094017

of W

$$(3.099)^2 + (1.094)^2 = 10.8$$

$\hat{\sigma}_s$

$\hat{\sigma}_e$

$$\hat{\sigma}_s^2 = (3.099)^2 \approx 9.6$$

```
> # Examine the estimated variance-covariance
> # matrix for the subvector of responses
> # from a single subject.
>
> getVarCov(o.lme, individuals = 1, type = "marginal")
```

Subj 1

Marginal variance covariance matrix

	1	2	3	4	5	6	7
1	10.8000	9.6033	9.6033	9.6033	9.6033	9.6033	9.6033
2	9.6033	10.8000	9.6033	9.6033	9.6033	9.6033	9.6033
3	9.6033	9.6033	10.8000	9.6033	9.6033	9.6033	9.6033
4	9.6033	9.6033	9.6033	10.8000	9.6033	9.6033	9.6033
5	9.6033	9.6033	9.6033	9.6033	10.8000	9.6033	9.6033
6	9.6033	9.6033	9.6033	9.6033	9.6033	10.8000	9.6033
7	9.6033	9.6033	9.6033	9.6033	9.6033	9.6033	10.8000

Alternative Parameterization for Compound Symmetry

$$\tilde{Y} = X\beta + \tilde{\epsilon}$$

$$\sigma^2 = \text{Var}(y) = \text{Var}(\tilde{\epsilon})$$

$$\begin{bmatrix} 1 & \rho \\ \rho & 1 & \rho & \rho & \rho & \rho & \rho & \rho \\ \rho & \rho & 1 & \rho & \rho & \rho & \rho & \rho \\ \rho & \rho & \rho & 1 & \rho & \rho & \rho & \rho \\ \rho & \rho & \rho & \rho & 1 & \rho & \rho & \rho \\ \rho & \rho & \rho & \rho & \rho & 1 & \rho & \rho \\ \rho & \rho & \rho & \rho & \rho & \rho & 1 & \rho \\ \rho & 1 \end{bmatrix} = W$$

```
o.cs = gls(Strength ~ Program * Timef, data = d,  
correlation = corCompSymm(form = ~ 1 | Subj))
```

```
> summary(o.cs)
```

Generalized least squares fit by REML

Model: Strength ~ Program * Timef

Data: d

AIC	BIC	logLik
1466.82	1557.323	-710.4101

Correlation Structure: Compound symmetry

Formula: ~1 | Subj

Parameter estimate(s):

Rho
0.8891805 - $\hat{\rho}$

.

.

.

Residual standard error: 3.286366

Degrees of freedom: 399 total; 378 residual

$\hat{\sigma}_e$

```
> getVarCov(o.cs)
```

Marginal variance covariance matrix

	[,1]	[,2]	[,3]	[,4]	[,5]	[,6]	[,7]
[1,]	10.8000	9.6033	9.6033	9.6033	9.6033	9.6033	9.6033
[2,]	9.6033	10.8000	9.6033	9.6033	9.6033	9.6033	9.6033
[3,]	9.6033	9.6033	10.8000	9.6033	9.6033	9.6033	9.6033
[4,]	9.6033	9.6033	9.6033	10.8000	9.6033	9.6033	9.6033
[5,]	9.6033	9.6033	9.6033	9.6033	10.8000	9.6033	9.6033
[6,]	9.6033	9.6033	9.6033	9.6033	9.6033	10.8000	9.6033
[7,]	9.6033	9.6033	9.6033	9.6033	9.6033	9.6033	10.8000

AR(1) Structure for W

$$\sigma^2 \begin{bmatrix} 1 & \phi & \phi^2 & \phi^3 & \phi^4 & \phi^5 & \phi^6 \\ \phi & 1 & \phi & \phi^2 & \phi^3 & \phi^4 & \phi^5 \\ \phi^2 & \phi & 1 & \phi & \phi^2 & \phi^3 & \phi^4 \\ \phi^3 & \phi^2 & \phi & 1 & \phi & \phi^2 & \phi^3 \\ \phi^4 & \phi^3 & \phi^2 & \phi & 1 & \phi & \phi^2 \\ \phi^5 & \phi^4 & \phi^3 & \phi^2 & \phi & 1 & \phi \\ \phi^6 & \phi^5 & \phi^4 & \phi^3 & \phi^2 & \phi & 1 \end{bmatrix}$$

```
o.ar1 = gls(Strength ~ Program * Timef, data = d,  
correlation = corAR1(form = ~ 1 | Subj))
```

```
> summary(o.ar1)
```

Generalized least squares fit by REML

Model: Strength ~ Program * Timef

Data: d

AIC	BIC	logLik
-----	-----	--------

1312.804	1403.306	-633.4018
----------	----------	-----------

Correlation Structure: AR(1)

Formula: ~1 | Subj

Parameter estimate(s):

Phi

$\hat{\varphi}$

0.9517769

.

.

.

Residual standard error: 3.280242

Degrees of freedom: 399 total; 378 residual

```
> getVarCov(o.ar1, individual = 3)
Marginal variance covariance matrix
```

	[,1]	[,2]	[,3]	[,4]	[,5]	[,6]	[,7]
[1,]	10.7600	10.2410	9.7473	9.2772	8.8298	8.4040	7.9988
[2,]	10.2410	10.7600	10.2410	9.7473	9.2772	8.8298	8.4040
[3,]	9.7473	10.2410	10.7600	10.2410	9.7473	9.2772	8.8298
[4,]	9.2772	9.7473	10.2410	10.7600	10.2410	9.7473	9.2772
[5,]	8.8298	9.2772	9.7473	10.2410	10.7600	10.2410	9.7473
[6,]	8.4040	8.8298	9.2772	9.7473	10.2410	10.7600	10.2410
[7,]	7.9988	8.4040	8.8298	9.2772	9.7473	10.2410	10.7600



getting smaller

General Positive Definite Structure for W

With δ_1 set equal to 1 for identifiability purposes, a general 7×7 positive definite variance-covariance matrix is parameterized by R as follows:

$$\begin{bmatrix} 1 & \rho_{12} & \rho_{13} & \rho_{14} & \rho_{15} & \rho_{16} & \rho_{17} \\ \rho_{12} & 1 & \rho_{23} & \rho_{24} & \rho_{25} & \rho_{26} & \rho_{27} \\ \rho_{13} & \rho_{23} & 1 & \rho_{34} & \rho_{35} & \rho_{36} & \rho_{37} \\ \rho_{14} & \rho_{24} & \rho_{34} & 1 & \rho_{45} & \rho_{46} & \rho_{47} \\ \rho_{15} & \rho_{25} & \rho_{35} & \rho_{45} & 1 & \rho_{56} & \rho_{57} \\ \rho_{16} & \rho_{26} & \rho_{36} & \rho_{46} & \rho_{56} & 1 & \rho_{67} \\ \rho_{17} & \rho_{27} & \rho_{37} & \rho_{47} & \rho_{57} & \rho_{67} & 1 \end{bmatrix}$$

most flexibility in terms of correlation between any observations

$\sigma^2 \text{diag}(\delta_1, \dots, \delta_7)$

allow for different $\text{Var}(y)$ across time

over time within the same individual

The 7×7 case doesn't fit on one slide, but here is the 5×5 case.

$\sigma^2 \delta_1^2$ $\sigma^2 \rho_{12} \delta_1 \delta_2$ $\sigma^2 \rho_{13} \delta_1 \delta_3$ $\sigma^2 \rho_{14} \delta_1 \delta_4$ $\sigma^2 \rho_{15} \delta_1 \delta_5$

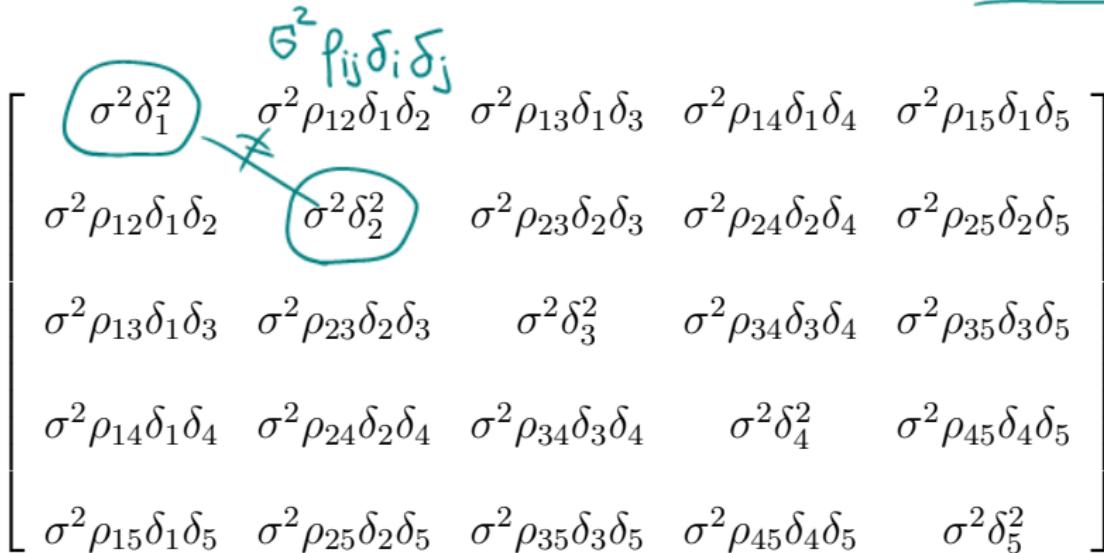
$\sigma^2 \rho_{12} \delta_1 \delta_2$ ~~$\sigma^2 \delta_2^2$~~ $\sigma^2 \rho_{23} \delta_2 \delta_3$ $\sigma^2 \rho_{24} \delta_2 \delta_4$ $\sigma^2 \rho_{25} \delta_2 \delta_5$

$\sigma^2 \rho_{13} \delta_1 \delta_3$ $\sigma^2 \rho_{23} \delta_2 \delta_3$ $\sigma^2 \delta_3^2$ $\sigma^2 \rho_{34} \delta_3 \delta_4$ $\sigma^2 \rho_{35} \delta_3 \delta_5$

$\sigma^2 \rho_{14} \delta_1 \delta_4$ $\sigma^2 \rho_{24} \delta_2 \delta_4$ $\sigma^2 \rho_{34} \delta_3 \delta_4$ $\sigma^2 \delta_4^2$ $\sigma^2 \rho_{45} \delta_4 \delta_5$

$\sigma^2 \rho_{15} \delta_1 \delta_5$ $\sigma^2 \rho_{25} \delta_2 \delta_5$ $\sigma^2 \rho_{35} \delta_3 \delta_5$ $\sigma^2 \rho_{45} \delta_4 \delta_5$ $\sigma^2 \delta_5^2$

$\sigma^2 \rho_{ij} \delta_i \delta_j$



o.un = gls(Strength ~ Program * Timef, data = d,
correlation = corSymm(form = ~ 1 | Subj),
weight = varIdent(form = ~ 1 | Timef))

allows variance to change over time

```
> summary(o.un)
```

Generalized least squares fit by REML

Model: Strength ~ Program * Timef

Data: d

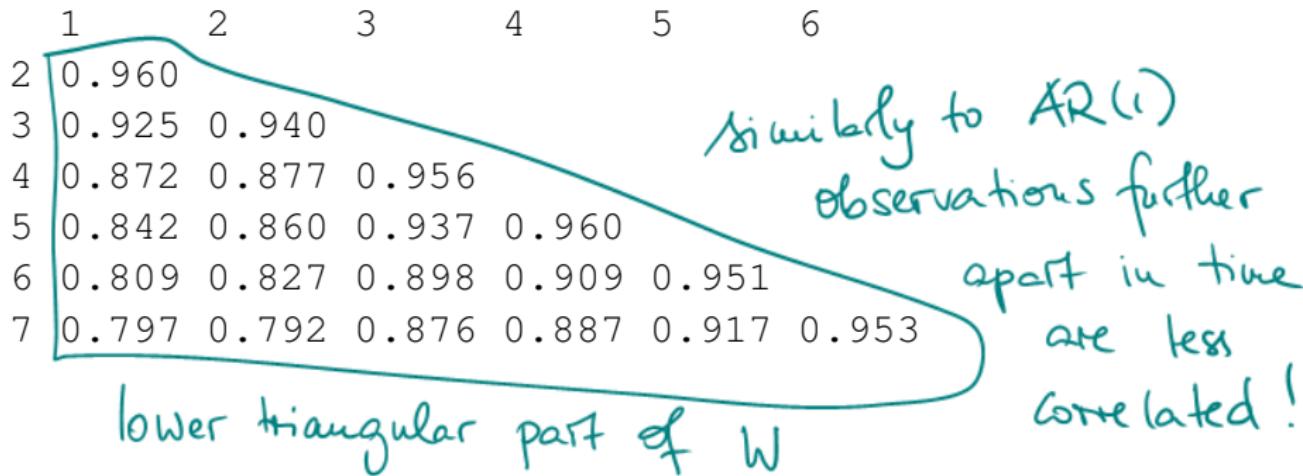
	AIC	BIC	logLik
	1332.896	1525.706	-617.4479

Correlation Structure: General

Formula: ~1 | Subj

Parameter estimate(s):

Correlation:



Variance function:

Structure: Different standard deviations per stratum

Formula: $\sim 1 | \text{Timef}$

Parameter estimates:

	2	4	6	8	10	12	14
1.000	1.000	1.039	1.104	1.071	1.174	1.157	1.203
.	$\hat{\sigma}_1$	$\hat{\sigma}_2$	$\hat{\sigma}_3$
.							
.							
							$\hat{\sigma}_7$

Residual standard error: 2.963129

Degrees of freedom: 399 total; 378 residual

```
> getVarCov(o.un, individual = 3)
Marginal variance covariance matrix
     [,1]   [,2]   [,3]   [,4]   [,5]   [,6]   [,7]
[1,] 8.7801 8.7571 8.9656 8.1984 8.6781 8.2203 8.4169
[2,] 8.7571 9.4730 9.4631 8.5686 9.2012 8.7307 8.6875
[3,] 8.9656 9.4631 10.7080 9.9266 10.6660 10.0700 10.2140
[4,] 8.1984 8.5686 9.9266 10.0770 10.6000 9.8987 10.0430
[5,] 8.6781 9.2012 10.6660 10.6000 12.0950 11.3440 11.3640
[6,] 8.2203 8.7307 10.0700 9.8987 11.3440 11.7560 11.6500
[7,] 8.4169 8.6875 10.2140 10.0430 11.3640 11.6500 12.7100
```

- To understand the reason for an identifiability constraint, notice that an arbitrary positive definite 7×7 covariance matrix depends on only

$$7 + 6 + 5 + 4 + 3 + 2 + 1 = \frac{7(7+1)}{2} = \underline{\underline{28}}$$

parameters. However, we have σ^2 , $6 + 5 + 4 + 3 + 2 + 1 = 21$ ρ parameters, and $\delta_1, \dots, \delta_7$.

- That's 29 parameters for a symmetric positive definite matrix that depends on at most 28 parameters.

to ensure identifiability R set $\delta_1 = 1$

- Thus, R chooses to set δ_1 to 1.
- Without such a constraint, it is easy to use different values of the parameters to define the same matrix. For example,

*2 time
points*

$$\begin{bmatrix} 3 & -1 \\ -1 & 7 \end{bmatrix} = 3 \begin{bmatrix} 1 & -\frac{1}{3} \\ -\frac{1}{3} & \frac{7}{3} \end{bmatrix} = 1 \begin{bmatrix} 3 & -1 \\ -1 & 7 \end{bmatrix}$$

$$\begin{array}{l} \sigma^2 \\ \hline \delta_1 \\ \hline \delta_2 \\ \hline \rho_{12} \end{array} \quad \begin{array}{l} 3 \\ 1 \\ \sqrt{\frac{7}{3}} \\ \frac{-1}{3\sqrt{\frac{7}{3}}} \end{array} \quad \begin{array}{l} 1 \\ \cancel{\sqrt{3}} \\ \cancel{\sqrt{7}} \\ \cancel{\frac{-1}{\sqrt{21}}} \end{array}$$

constraint

05-02-25

*end
lecture*

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