

# Calculation of Type I, II, and III Sums of Squares

SS (time | 1, temp, time x temp)

Every Type I, II, or III sum of squares is the error sums of squares for a reduced model minus the error sum of squares for a model that adds one term to the reduced model:

$$\mathbf{y}^\top (\mathbf{I} - \mathbf{P}_{\mathbf{X}_{\text{reduced}}}) \mathbf{y} - \mathbf{y}^\top (\mathbf{I} - \mathbf{P}_{\mathbf{X}_{\text{reduced}+\text{term}}}) \mathbf{y} = \mathbf{y}^\top (\mathbf{P}_{\mathbf{X}_{\text{reduced}+\text{term}}} - \mathbf{P}_{\mathbf{X}_{\text{reduced}}}) \mathbf{y},$$

where  $\mathcal{C}(\mathbf{X}_{\text{reduced}}) \subset \mathcal{C}(\mathbf{X}_{\text{reduced}+\text{term}}) \subseteq \mathcal{C}(\mathbf{X})$ .

As usual,  $\mathbf{X}$  represents the model matrix for the most complex model under consideration (a.k.a., the full model).

For all Type III sums of squares, the reduced+term model is the full model.

For example,  $SS(\text{temp}|1, \text{time})$  is

$$\mathbf{y}^\top (\mathbf{I} - \mathbf{P}_{\mathbf{X}_{\text{reduced}}}) \mathbf{y} - \mathbf{y}^\top (\mathbf{I} - \mathbf{P}_{\mathbf{X}_{\text{reduced+term}}}) \mathbf{y} = \mathbf{y}^\top (\mathbf{P}_{\mathbf{X}_{\text{reduced+term}}} - \mathbf{P}_{\mathbf{X}_{\text{reduced}}}) \mathbf{y},$$

where we can choose

$$\mathbf{X}_{\text{reduced}} = \begin{bmatrix} 1 & 0 \\ 1 & 0 \\ 1 & 0 \\ 1 & 0 \\ 1 & 0 \\ 1 & 1 \\ 1 & 1 \\ 1 & 1 \\ 1 & 1 \\ 1 & 1 \end{bmatrix} \quad \text{and} \quad \mathbf{X}_{\text{reduced+term}} = \begin{bmatrix} 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \\ 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 1 & 1 \end{bmatrix}.$$

## How does this apply to a Type III Sum of Squares like $SS(\text{time} | 1, \text{temp}, \text{time} \times \text{temp})$ ?

Here the reduced+term model is actually the full cell means model that includes an intercept, time and temperature main effects, and time  $\times$  temperature interaction.

The reduced model has an intercept, temperature main effect, and time  $\times$  temperature interaction but no time main effect.

How do we fit such a reduced model?

Code like `lm(y~b+a:b)` or `model y=b a*b`; does not fit the reduced model with no a main effects.

```
> model.matrix(~temp+time:temp)
```

	(Intercept)	temp30	temp20:time6	temp30:time6
1	1	0	0	0
2	1	0	0	0
3	1	1	0	0
4	1	1	0	0
5	1	1	0	0
6	1	0	1	0
7	1	0	1	0
8	1	0	1	0
9	1	0	1	0
10	1	1	0	1

*add*

This model matrix has **rank 4** and the same column space as  $X$ .

# Removing Time Main Effect from Cell Means Model

Time	Temp		Marginal Mean	
	20°	30°		
3 months	$\mu_{11}$	$\mu_{12}$	$\bar{\mu}_{1.} = \frac{(\mu_{11} + \mu_{12})}{2}$	$\bar{\mu}_{1.} - \bar{\mu}_{2.}$ $\mu_{11} + \mu_{12} = \mu_{21} + \mu_{22}$
6 months	$\mu_{21}$	$\mu_{22}$	$\bar{\mu}_{2.} = \frac{(\mu_{21} + \mu_{22})}{2}$	

The parameters  $\mu_{11}$ ,  $\mu_{12}$ ,  $\mu_{21}$ , and  $\mu_{22}$  can be anything as long as

$$\bar{\mu}_{1.} = \bar{\mu}_{2.} \iff \mu_{22} = \mu_{11} + \mu_{12} - \mu_{21}.$$

# A Model Matrix for Model with 1, temp, time $\times$ temp

$$\mathbf{X}_{\text{reduced}} \boldsymbol{\beta}_{\text{reduced}} = \begin{bmatrix} 1 & 0 & 0 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 1 & 1 & -1 \end{bmatrix} \begin{bmatrix} \mu_{11} \\ \mu_{12} \\ \mu_{21} \end{bmatrix} = \begin{bmatrix} \mu_{11} \\ \mu_{11} \\ \mu_{12} \\ \mu_{12} \\ \mu_{12} \\ \mu_{21} \\ \mu_{21} \\ \mu_{21} \\ \mu_{21} \\ \mu_{11} + \mu_{12} - \mu_{21} \end{bmatrix}$$

$\uparrow$   
 $\mu_{22} = \mu_{11} + \mu_{12} - \mu_{21}$

```
#Type III Sum of Squares for SS(time|1,temp,time x temp)
```

```
> x0=x[,1:3]
```

```
> x0[10,]=c(1,1,-1)
```

```
> x0
```

1	1	0	0
2	1	0	0
3	0	1	0
4	0	1	0
5	0	1	0
6	0	0	1
7	0	0	1
8	0	0	1
9	0	0	1
10	1	1	-1

```
> anova(lm(y~0+x0),lm(y~0+x))
```

Analysis of Variance Table

Model 1:  $y \sim 0 + x_0$

Model 2:  $y \sim 0 + x$

	Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
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1	7	24
2	6	12

1

12

6

0.04983 \*

this model matrix no longer accounts for a time main effect.

$$F = \frac{y^T (P_x - P_{x_0}) y / 1}{y^T (I - P_x) y / 6}$$

$$F = \frac{(24 - 12) / 1}{12 / 6} = \frac{12}{2} = 6$$

# Alternative Computation of Sums of Squares

Let  $SS = \mathbf{y}^\top (\mathbf{P}_{\mathbf{X}_{\text{reduced}+\text{term}}} - \mathbf{P}_{\mathbf{X}_{\text{reduced}}}) \mathbf{y}$  represent any Type I, II, or III sum of squares.

Let  $q = \text{rank}(\mathbf{X}_{\text{reduced}+\text{term}}) - \text{rank}(\mathbf{X}_{\text{reduced}})$  be the degrees of freedom associated with  $SS$ .

Let  $\mathbf{C}$  be any  $q \times p$  matrix whose rows are a basis for the row space of  $(\mathbf{P}_{\mathbf{X}_{\text{reduced}+\text{term}}} - \mathbf{P}_{\mathbf{X}_{\text{reduced}}}) \mathbf{X}$ .



## Alternative Computation of Sums of Squares

$$\underline{y^T (P_{j+1} - P_j) y} = \underline{(C\hat{\beta})^T \{ C(X^T X)^{-1} C^T \}^{-1} C\hat{\beta}}$$

Then the ANOVA  $F$  statistic

$$\frac{SS/q}{MSE} = \frac{\hat{\beta}^T C^T [C(X^T X)^{-1} C^T]^{-1} C\hat{\beta}/q}{\hat{\sigma}^2}.$$

Thus, any  $SS$  can be computed as  $\hat{\beta}^T C^T [C(X^T X)^{-1} C^T]^{-1} C\hat{\beta}$  for an appropriate matrix  $C$ .