

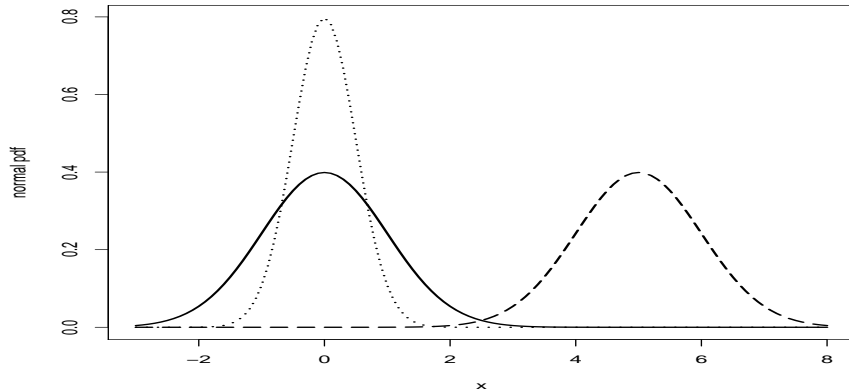
Common univariate distributions

Continuous distributions: Normal (Gaussian) distribution

$$X \sim \underline{N(\mu, \sigma^2)} \quad -\infty < \mu < \infty, \sigma > 0$$

- pdf given by

$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}, \quad -\infty < x < \infty$$



- Motivation: single most important distribution
 - widely used & analytically tractable
 - bell-shaped density seems to occur naturally
 - Central Limit Theorem (normal distribution is extremely relevant in large samples; more later)
- $\mu \in \mathbb{R}$ is the mean EX of the distribution
- $\sigma^2 = \text{Var}(X)$ is the variance of the distribution; σ is the standard deviation
- Many properties of the normal distribution can be most easily derived using the $N(0, 1)$ or **standard normal distribution**

$$\left(\begin{array}{l} 1. \text{ If } X \sim \underline{N(\mu, \sigma^2)} \text{ then } Z = (X - \mu)/\sigma \sim \underline{N(0, 1)}. \\ 2. \text{ If } Z \sim \underline{N(0, 1)}, \text{ then } \underline{X} = a + bZ \sim N(\mu = a, \sigma^2 = b^2) \text{ for } a, b \in \mathbb{R} \end{array} \right.$$

\uparrow \uparrow

Common univariate distributions

Standard normal distribution (cont'd)

In developing normal distributions, one actually *starts* with defining the standard normal distribution:

Definition: If a random variable Z has pdf $Z \sim N(0,1)$

$$\longrightarrow f(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2}, \quad -\infty < z < \infty$$

then we say Z has a standard normal distribution, denoted $N(0,1)$.

Notation:

1. it's fairly standard (at least in statistics) to use $\phi(z)$ to denote the standard normal pdf

$$\begin{array}{c} \phi(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2} \\ \downarrow \\ \text{pdf} \end{array}$$

2. it's also standard to use $\Phi(z)$ to denote the standard normal cdf

$$\begin{aligned} \Phi(z) &= \mathbb{P}(Z \leq z) = \int_{-\infty}^z \frac{1}{\sqrt{2\pi}} e^{-y^2/2} dy \\ \frac{d\Phi(z)}{dz} &= \phi(z) \end{aligned}$$

3. evaluating Φ is a numerical analysis problem because the pdf ϕ has no simple anti-derivative (the cdf Φ is often tabulated and can be computed with software too)

Common univariate distributions

Standard normal distribution (cont'd)

- Mean: $Z \sim N(0,1)$

$$EZ = \int_{-\infty}^{\infty} \underbrace{z}_{\text{odd}} \underbrace{\frac{1}{\sqrt{2\pi}} e^{-z^2/2}}_{\text{even}} dz = 0$$

- Variance:

$$1 = \text{Var}(Z) = EZ^2 = \int_{-\infty}^{\infty} \underbrace{z^2}_{\text{even}} \frac{1}{\sqrt{2\pi}} e^{-z^2/2} dz = \frac{2}{\sqrt{2\pi}} \int_0^{\infty} z^2 e^{-z^2/2} dz$$

$$\sqrt{y} = \frac{z}{\sqrt{2}}$$

$$y = \frac{z^2}{2}$$

$$dy = z dz$$

$$\frac{dy}{\sqrt{2y}} dz = \frac{dy}{z}$$

$$= \frac{dy}{\sqrt{2y}}$$

$$\Gamma(\alpha) = \int_0^{\infty} y^{\alpha-1} e^{-y} dy$$

$$= \frac{2}{\sqrt{\pi}} \Gamma\left(\frac{3}{2}\right)$$

$$= \frac{2}{\sqrt{\pi}} \cdot \frac{1}{2} \sqrt{\pi} = 1$$

$$= \frac{2}{\sqrt{\pi}} \int_0^{\infty} y^{1/2} e^{-y} dy$$

$$= \frac{2}{\sqrt{\pi}} \int_0^{\infty} y^{1/2+1-1} e^{-y} dy$$

$$\begin{aligned} \Gamma(z+1) &= z \Gamma(z) \\ \Gamma\left(\frac{3}{2}\right) &= \Gamma\left(\frac{1}{2}+1\right) = \frac{1}{2} \Gamma\left(\frac{1}{2}\right) \\ &= \frac{1}{2} \sqrt{\pi} \end{aligned}$$

- mgf:

$$M_Z(t) = Ee^{tZ} = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{tz} e^{-z^2/2} dz = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}(z^2-2tz)} dz$$

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{\frac{1}{2}[(z-t)^2 - t^2]} dz$$

$$= \frac{e^{-t^2/2}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{\frac{1}{2}(z-t)^2} dz = 1$$

$$M_Z(t) = e^{t^2/2}$$

Note:
Let $Y \sim N(t, 1)$
 $f_Y(y) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(y-t)^2}$

$$\begin{aligned} (a-b)^2 &= a^2 + b^2 - 2ab \\ z^2 - 2tz &= z^2 + t^2 - t^2 - 2tz \\ &= (z^2 + t^2 - 2tz) - t^2 \\ &= (z-t)^2 - t^2 \end{aligned}$$

Common univariate distributions

From standard normal to normal

Definition: If, for $\underline{\underline{\mu}} \in \mathbb{R}$ and $\underline{\underline{\sigma}} > 0$, X has the same distribution as

$$\mu + \sigma Z$$

for $Z \sim N(0, 1)$ (standard normal), then we say X has a normal $\underline{\underline{N(\mu, \sigma^2)}}$ distribution.

Step 1: Take any $\mu + \sigma^2$

Step 2: get a $Z \sim N(0, 1)$

Step 3: Construct

$$X = \mu + \sigma Z \sim N(\mu, \sigma^2)$$

Some facts/properties of $\underline{\underline{X \sim N(\mu, \sigma^2)}}$, $\mu \in \mathbb{R}$, $\sigma > 0$:

1. cdf: for $-\infty < x < \infty$,

$$F_X(x) = P(\underline{\underline{X}} \leq x) = P(\mu + \sigma Z \leq x) = P\left(Z \leq \frac{x - \mu}{\sigma}\right) = F_Z\left(\frac{x - \mu}{\sigma}\right)$$

$$\stackrel{(*)}{=} \Phi\left(\frac{x - \mu}{\sigma}\right)$$

CDF of standard $N(0, 1)$

2. pdf: for $-\infty < x < \infty$,

$$f_X(x) = \frac{d}{dx} F_X(x) \stackrel{(*)}{=} \frac{d}{dx} \Phi\left(\frac{x - \mu}{\sigma}\right) = \frac{1}{\sigma} \phi\left(\frac{x - \mu}{\sigma}\right)$$

$$= \frac{1}{\sigma} \frac{1}{\sqrt{2\pi}} e^{-\frac{(x - \mu)^2}{2\sigma^2}}$$

$$\phi(y) = \frac{1}{\sqrt{2\pi}} e^{-y^2/2}$$

3. Mean: $EX = E(\underline{\underline{\mu + \sigma Z}}) = \mu + \sigma EZ \neq \mu$

4. Variance: $\text{Var}(X) = \text{Var}(\mu + \sigma Z) = \sigma^2 \text{Var}(Z) = \sigma^2$

Common univariate distributions

From standard normal to normal (cont'd)

Given $\mu, \sigma^2, Z \Rightarrow X = \mu + \sigma Z \sim N(\mu, \sigma^2)$

5. mgf: for $-\infty < t < \infty$,

$$M_X(t) = \mathbb{E}e^{tX} = \mathbb{E}e^{t(\mu + \sigma Z)} = \mathbb{E}e^{t\mu}e^{t\sigma Z} = e^{t\mu} \mathbb{E}e^{t\sigma Z} = e^{t\mu} M_Z(t\sigma)$$

$$= e^{t\mu} e^{\frac{t^2 \sigma^2}{2}} = e^{t\mu + \frac{t^2 \sigma^2}{2}} \Rightarrow M_X(t)$$

6. If $a, b \in \mathbb{R}$ with $b \neq 0$ and $X \sim N(\mu, \sigma^2)$, then $Y = a + bX \sim N(a + b\mu, b^2\sigma^2)$

$$\mathbb{E}(Y) = \mathbb{E}(a + bX) = a + b\mathbb{E}X = a + b\mu$$

$$\text{Var}(Y) = \text{Var}(a + bX) = b^2\sigma^2$$

$$M_Y(t) = \mathbb{E}[e^{tY}] = \mathbb{E}[e^{t(a+bX)}] = \mathbb{E}[e^{ta} e^{tbX}] = e^{ta} \mathbb{E}[e^{tbX}] = e^{ta} M_X(tb)$$

7. In particular, if $X \sim N(\mu, \sigma^2)$ then $Z = \frac{X - \mu}{\sigma} \sim N(0, 1)$.

$$b = \frac{1}{\sigma}, a = -\mu/\sigma$$

$$Z = \frac{X - \mu}{\sigma} \Rightarrow Z \sim N\left(\frac{-\mu}{\sigma} + \frac{1}{\sigma}\mu, \frac{1}{\sigma^2}\sigma^2\right) = N(0, 1)$$

8. 68-95-99% rule (prob that X lies within 1, 2, or 3 standard deviations of μ)

$$P(\mu - \sigma \leq X \leq \mu + \sigma) = F_X(\mu + \sigma) - F_X(\mu - \sigma) = \Phi(1) - \Phi(-1) = 0.6828$$

$$P(\mu - 2\sigma \leq X \leq \mu + 2\sigma) = F_X(\mu + 2\sigma) - F_X(\mu - 2\sigma) = \Phi(2) - \Phi(-2) = 0.9544$$

$$P(\mu - 3\sigma \leq X \leq \mu + 3\sigma) = F_X(\mu + 3\sigma) - F_X(\mu - 3\sigma) = \Phi(3) - \Phi(-3) = 0.9974$$

9. To recap:

(a) If $Z \sim N(0, 1)$ and $\mu \in \mathbb{R}$, $\sigma > 0$, then $X = \mu + \sigma Z \sim N(\mu, \sigma^2)$

(b) If $X \sim N(\mu, \sigma^2)$ then $Z = \frac{X - \mu}{\sigma} \sim N(0, 1)$

(c) In general, if $a, b \in \mathbb{R}$ with $b \neq 0$ and $X \sim N(\mu, \sigma^2)$, then $a + bX \sim N(a + b\mu, b^2\sigma^2)$

ASIDE

Common univariate distributions

Distributions related to normal

1. Log-normal: If $X \sim N(\mu, \sigma^2)$ then $Y = e^X \sim \text{LogNormal}(\mu, \sigma^2)$

- pdf given by

$$\rightarrow f_Y(y) = \frac{1}{y} f_X(\log y) = \frac{1}{y} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2}\left(\frac{\log y - \mu}{\sigma}\right)^2}, \quad 0 < y < \infty$$

- Mean: $EY = Ee^X = M_X(t=1) = e^{\mu + \frac{1}{2}\sigma^2}$

Note: $EY = Ee^X \geq e^{EX} = e^\mu$ by Jensen's inequality

- shape is like that of a gamma distribution with $\alpha > 1$

- common in economics (e.g., assume $\log(\text{income})$ is normal) and also as a failure time distribution

$$\begin{aligned} F_Y(y) &\stackrel{\text{def}}{=} P(Y \leq y) = P(e^X \leq y) \\ &= P(X \leq \ln y) \\ &= F_X(\ln y) \\ f_Y(y) &= \frac{d}{dy} F_Y(y) = \frac{d}{dy} F_X(\ln y) \\ &= \frac{1}{y} f_X(\ln y) \end{aligned}$$

2. If $Z \sim N(0, 1)$ then $Z^2 \sim \chi_1^2$ (chi-squared 1 df)

3. Standard Brownian motion $\{B(t) : t \geq 0\}$

- a stochastic process where $B(t)$ represents an object's position at time t
- $B(0) = 0$ and $B(t) \sim N(\mu = 0, \sigma^2 = t)$ for each $t > 0$
- increments $B(s)$ and $B(t+s) - B(s)$ are independent, any $t > s \geq 0$
- appears often in probability/statistics/economics

