

Common univariate distributions

Discrete distributions: Negative Binomial

$X \sim \text{Neg-Binom}(r, p), 0 < p < 1$

- pmf given by

$$f_X(x) = f_X(x|r, p) = \binom{x-1}{r-1} p^r (1-p)^{x-r}, \quad x = r, r+1, \dots,$$

- Motivation: distribution for the number of independent Bernoulli(p) trials needed to obtain r successes

- $Y = X - r$ (number of failures prior to the r th success) also common

$$f_Y(y|r, p) = \binom{y+r-1}{y} p^r (1-p)^y, \quad y = 0, 1, \dots,$$

- Showing that these probabilities sum to 1 is not easy (next slide)
- Be careful: both r.v.s X and Y (different) are called “negative binomial”

- Mean: $EY = \frac{r(1-p)}{p}$ and hence $EX = EY + r = \frac{r}{p}$

- Variance: $\text{Var}(Y) = \frac{r(1-p)}{p^2} = \text{Var}(X)$

- $M_Y(t) = Ee^{tY} = \left[\frac{p}{1 - (1-p)e^t} \right]^r, t < -\log(1-p),$

$$M_X(t) = Ee^{t(Y+r)} = Ee^{rt}e^{tY} = e^{rt}M_Y(t)$$

Common univariate distributions

Discrete distributions: Negative Binomial (cont'd)

To show probabilities sum to 1:

1. Newton's negative binomial formula : if $\alpha < 0$ and $|x| < 1$,

$$(1+x)^\alpha = \sum_{k=0}^{\infty} g^{(k)}(0) \frac{(x-0)^k}{k!} = \sum_{k=0}^{\infty} \binom{\alpha}{k}^* x^k, \quad \binom{\alpha}{k}^* \equiv \frac{\alpha(\alpha-1)\cdots(\alpha-k+1)}{k!}$$

Taylor expanding $g(x) = (1+x)^\alpha$ around 0: $g^{(0)}(0) = g(0) = 1$, $g^{(1)}(0) = g'(0) = \alpha$

2. for integers $r \geq 1$ and $k \geq 0$, note that

$$\binom{-r}{k}^* (-1)^k = (-1)^k \frac{(-r)(-r-1)\cdots(-r-k+1)}{k!} = \frac{(r)(r+1)\cdots(r+k-1)}{k!} = \binom{r+k-1}{k}$$

$$\begin{aligned} \sum_{y=0}^{\infty} f_Y(y) &= \sum_{y=0}^{\infty} \binom{y+r-1}{y} p^r (1-p)^y = \sum_{y=0}^{\infty} \binom{-r}{y}^* (-1)^y p^r (1-p)^y \\ &= p^r \sum_{y=0}^{\infty} \binom{-r}{y}^* (p-1)^y \\ &= p^r [1 + (p-1)]^{-r} = 1 \end{aligned}$$

Show $M_Y(t) = Ee^{tY} = \left[\frac{p}{1 - (1-p)e^t} \right]^r$ for $t < -\log(1-p)$

Common univariate distributions

Discrete distributions: Geometric

$$X \sim \text{Geom}(p), 0 < p < 1$$

- special case of Negative Binomial($r = 1, p$)
- Motivation: distribution for the number of independent Bernoulli(p) trials needed to obtain 1st success
- pmf given by

$$P(X=x) = f_X(x) = f_X(x|p) = p(1-p)^{x-1}, \quad x = 1, 2, 3, \dots,$$

- Mean: $EX = \frac{1}{p}$

- Variance: $\text{Var}(X) = \frac{1-p}{p^2}$

- $M_X(t) = Ee^{tX} = \frac{pe^t}{1 - (1-p)e^t}$ for $t < -\log(1-p)$

Recall: $X \sim \text{Neg-Bin}(r, p)$

$$M_X(t) = \frac{e^{rt} p^r}{(1 - (1-p)e^t)^r}$$

$r=1 \Rightarrow M_X(t) = \frac{e^t p}{1 - (1-p)e^t}$

Common univariate distributions

Discrete distributions: Geometric (cont'd)

1. We've seen the cdf of a $\text{Geom}(p)$ random variable X ; it's relatively simple.

$$X \sim \text{Geom}(p) \Rightarrow \underline{\underline{F_X(x)}} \stackrel{\text{def}}{=} \mathbb{P}(X \leq x)$$

$$1 - F_X(x) = \mathbb{P}(X > x) = (1-p)^x \quad (*)$$

$$F_X(x) = 1 - (1-p)^x \text{ for } x=1, 2, \dots$$

2. The Geometric distribution has the famous “memoryless” property: for any integer $x_0 \geq 0$.

$$\mathbb{P}(X = x_0 + x | X > x_0) = \mathbb{P}(X = x)$$

Interpretation: The conditional distribution of the remaining waiting number of trials until a 1st success, given that I've already waited x_0 trials, is the same as the original distribution of the number of trials until 1st success.

Given that each trial is an independent Bernoulli trial, this does make sense: whether I start counting trials at the beginning or I start counting trials after x_0 trials without success, the distribution of the remaining number of trials needed until a 1st success should be the same.

$$\begin{aligned} \mathbb{P}(X = x + x_0 | X > x_0) &= \frac{\mathbb{P}(X = x + x_0 \cap X > x_0)}{\mathbb{P}(X > x_0)} = \frac{\mathbb{P}(X = x + x_0)}{\mathbb{P}(X > x_0)} \\ &= \frac{p(1-p)^{x+x_0-1}}{(1-p)^{x_0}} = p(1-p)^{x-1} \\ &= \mathbb{P}(X = x) \quad \square \end{aligned}$$

3. Example: Testing newly manufactured widgets with $p = 0.01$ probability that a given widget fails a functionality test, what's the probability of running at least 50 units without a test failure?

$\longrightarrow X = \# \text{ of trials (tests) until the 1st failure}$

$$X \sim \text{Geom}(0.01)$$

$$\mathbb{P}(X > 50) = (1-p)^x = (1-0.01)^{50} \approx 0.605$$

$$\mathbb{E}(X) = \frac{1}{p} = \underline{\underline{100}}$$

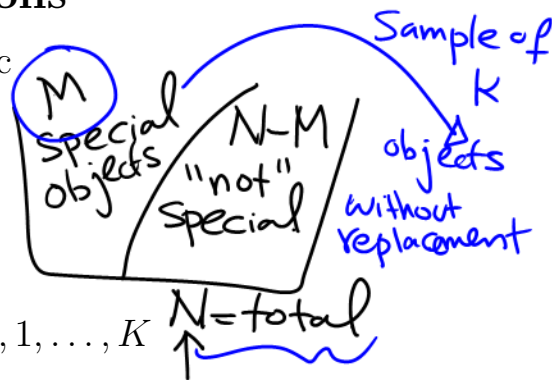
Common univariate distributions

Discrete distributions: Hypergeometric

$X \sim \text{Hypergeometric}(N, M, K)$ (integers N, M, K)

- pmf given by

$$P(X=x) = f_X(x) = \frac{\binom{M}{x} \binom{N-M}{K-x}}{\binom{N}{K}}, \quad x = 0, 1, \dots, K$$



- Motivation: Choose K objects without replacement from a total population of size N which contains M "special" objects. X is the number of "special" objects among the K chosen.

$$P(X=x) = \frac{M!}{x! (M-x)!} \frac{(N-M)!}{(K-x)! (N-M-K+x)!} \frac{K! (N-K)!}{N!}$$

pmf of $X \sim \text{Hyper}(N, M, K)$
for $x = 0, 1, 2, \dots, K$

total \uparrow special \uparrow

- Must have $0 \leq x \leq K$, $x \leq M$, and $K-x \leq N-M$ in $f_X(x)$

Typically, $N > 2M$ and $M > K$ so only the condition $0 \leq x \leq K$ matters

- Mean: $EX = KM/N$

$$P(X=x) = \frac{\binom{M}{x} \binom{N-M}{K-x}}{\binom{N}{K}} \text{ for } x = 0, 1, 2, \dots, K$$

Size of your sample x \uparrow special
% of special object

- Variance: $\text{Var}(X) = \frac{KM(N-M)(N-K)}{N^2(N-1)}$

To derive mean:

$$EX = \sum_{x=0}^K x \binom{M}{x} \binom{N-M}{K-x} / \binom{N}{K} = \sum_{x=1}^K x \frac{M!}{x! (M-x)!} \binom{N-M}{K-x} \frac{K! (N-K)!}{N!}$$

def $\sum_{x=0}^K x p(X=x)$

$$= \sum_{x=1}^K \frac{x M!}{x! (M-x)!} \binom{N-M}{K-x} \frac{K! (N-K)!}{N!}$$

$$= \frac{KM}{N} \sum_{x=1}^K \frac{(M-1)!}{(M-x)!} \binom{N-M}{K-x} \frac{(K-1)! (N-K)!}{(N-1)!}$$

79

$$Y \sim \text{HyperGeo}(N-1, M-1, K-1) \quad \underline{y=x-1} \quad \frac{KM}{N} \quad \sum_{y=0}^{K-1} \frac{(M-1)!}{(M-1-y)!(N-1)!} \binom{N-1-(M-1)}{K-1-(x-1)} \frac{(K-1)!(N-1-(K-0))!}{(N-1)!} = 1$$

Common univariate distributions

Discrete distributions: Hypergeometric and Binomial

1. Hypergeometric: sampling without replacement
 i.e., choose x special objects in a size K sample from a collection where M objects are “special” & $N - M$ are not
2. Binomial: sampling with replacement
 i.e., choose x special objects in a size n sample, where each selected item of the sample has probability p of being a special object

$$\text{Bin}(n=K, p=\frac{M}{N})$$

Suppose $X \sim \text{Hypergeometric}(\underline{N}, \underline{M}, \underline{K})$ and let $\underline{M/N = p}$ be the proportion of “special” objects and let $\underline{K = n}$ be the sample size. Then,

- $\underline{EX = K(M/N) = np}$
- $\text{Var}(X) = \frac{KM(N-M)(N-K)}{N^2(N-1)} = \underline{np(1-p) \frac{N-n}{N-1}}$
- the factor $(N-n)/(N-1)$ in $\text{Var}(X)$ is the finite sample correction factor

Note: $N \geq 20K$
 \uparrow sample
 $N \geq 20n$

Now fix $K = n$ and $p = M/N$ as above, and let $N \rightarrow \infty$. Then,

$$\lim_{N \rightarrow \infty} P(X = x) = \lim_{N \rightarrow \infty} \binom{M}{x} \binom{N-M}{K-x} / \binom{N}{K} = \binom{n}{x} p^x (1-p)^{n-x}, \quad x = 0, \dots, n$$

Result: hypergeometric tends to binomial distribution in large population size N