

STAT 5430

Lec 33, F, Apr 18

(No new homework)

Interval Estimation I

Pivotal Quantities

↓ vector of parameters $\underline{\theta}$

Definition: Let X_1, \dots, X_n be joint pdf/pmf $f(x|\underline{\theta})$, $\underline{\theta} \in \Theta \subset \mathbb{R}^p$. Then a random variable $Q(\underline{X}, \underline{\theta})$ is called a **pivot** or **pivotal quantity** if the distribution of $Q(\underline{X}, \underline{\theta})$ under $\underline{\theta}$ does not depend on $\underline{\theta}$.

Note: $Q(\underline{X}, \underline{\theta})$ is NOT a statistic (because can depend on $\underline{\theta}$)
 $P_{\underline{\theta}}(Q(\underline{X}, \underline{\theta}) \in A) = P(Q(\underline{X}, \underline{\theta}) \in A)$

Some examples: (pivots, unlike statistics, can be functions of parameters $\underline{\theta}$)

$\frac{X_1 - X_2}{X_3 - X_4}$ ancillary statistic

1. Let $X_1 \dots X_n$ be iid $N(\mu, \sigma^2)$ random variables.

$$\bar{X}_n \sim N(\mu, \frac{\sigma^2}{n})$$

pivot $\left\{ \frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}} \sim N(0,1) \right\}$ pivot $\left\{ \frac{\bar{X}_n - \mu}{s/\sqrt{n}} \sim T_{n-1} \text{ distribution} \right\}$

2. Let f_0 be a pdf on \mathbb{R} . Let $X_1 \dots X_n$ be iid with random variables common pdf $f(x|\underline{\theta})$ where

$$\underline{\theta} = (\theta_1, \theta_2) \quad f(x|\underline{\theta}) = \frac{1}{\theta_2} f_0\left(\frac{x - \theta_1}{\theta_2}\right), \quad x \in \mathbb{R},$$

for $\underline{\theta} = (\theta_1, \theta_2)$, $\theta_1 \in \mathbb{R}$ (location parameter) and $\theta_2 > 0$ (scale parameter).

Then, $Q(\underline{X}, \underline{\theta}) = \frac{\bar{X}_n - \theta_1}{\theta_2} \leftarrow \text{pivot}$

Why? Note: $Y_i \equiv \frac{X_i - \theta_1}{\theta_2} \sim f_0(y)$ (i.e. Y_1, \dots, Y_n iid $f_0(y)$)

and $Q(\underline{X}, \underline{\theta}) = \frac{1}{n} \sum_{i=1}^n Y_i \leftarrow \text{distribution doesn't depend on } \underline{\theta}$

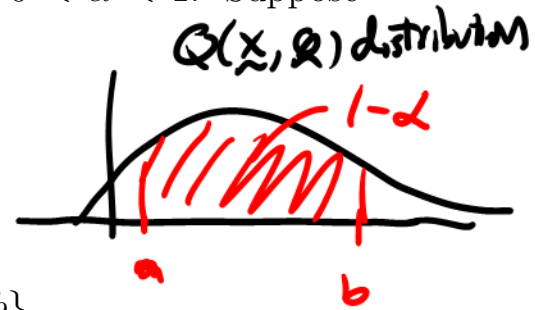
Interval Estimation I

Interval Estimation via Pivotal Quantities

Remarks:

1. Let $Q(\underline{X}, \underline{\theta})$ be a pivotal quantity ($\underline{\theta} \in \Theta \subset \mathbb{R}^p$) and $0 < \alpha < 1$. Suppose $-\infty \leq a \leq b \leq \infty$ are such that

$$P(a \leq Q(\underline{X}, \underline{\theta}) \leq b) = P_{\underline{\theta}}(a \leq Q(\underline{X}, \underline{\theta}) \leq b) = 1 - \alpha.$$



Then,

$$C_{\underline{X}} = \{\underline{\theta} : \underline{\theta} \in \Theta, a \leq Q(\underline{X}, \underline{\theta}) \leq b\}$$

is a **confidence region** for $\underline{\theta}$ with CC $(1 - \alpha)$

$$\text{That is, } \min_{\underline{\theta} \in \Theta} P_{\underline{\theta}}(\underline{\theta} \in C_{\underline{X}}) = \min_{\underline{\theta} \in \Theta} P_{\underline{\theta}}(a \leq Q(\underline{X}, \underline{\theta}) \leq b) = 1 - \alpha.$$

$$P(a \leq Q(\underline{X}, \underline{\theta}) \leq b)$$

2. If $\Theta \subset \mathbb{R}$ and $Q(\underline{X}, \underline{\theta})$ is monotone in $\underline{\theta} \in \mathbb{R}$, then the region $C_{\underline{X}}$ will be an interval.

Interval Estimation I

Interval Estimation via Pivotal Quantities

Example: Let $X_1 \dots X_n$ be iid $\text{Gamma}(\delta_0, \theta)$ where $\theta > 0$ (δ_0 fixed/known). Using a pivotal quantity based on $\sum_{i=1}^n X_i$, find a CI for θ with C.C. $1 - \alpha$.

Shape is known
Scale (unknown)

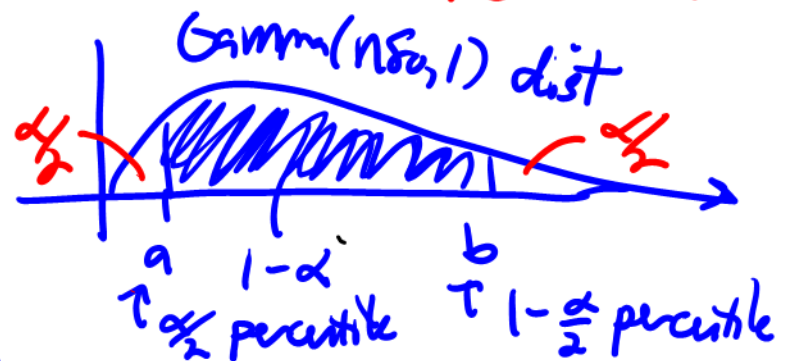
Solution: $\sum_{i=1}^n X_i \sim \text{Gamma}(n\delta_0, \theta)$

$Q(X, \theta) = \frac{\sum_{i=1}^n X_i}{\theta} \sim \text{Gamma}(n\delta_0, 1)$

↑ pivot

↑ dist. doesn't depend on θ

Find a & b such that



Confidence interval for $\theta > 0$

$$\begin{aligned} \text{is } \{ \theta : a \leq Q(\underline{X}, \theta) \leq b \} &= \{ \theta : a \leq \frac{\sum X_i}{\theta} \leq b \} \\ &= \{ \theta > 0 : \frac{\sum X_i}{b} \leq \theta \leq \frac{\sum X_i}{a} \} \\ &= \left[\frac{\sum X_i}{b}, \frac{\sum X_i}{a} \right] \end{aligned}$$

Interval Estimation I

Asymptotically Pivotal Quantities

Definition: A sequence of random variables $Q_n \equiv Q_n(X_1, \dots, X_n, \theta)$, $n \geq 1$, is called **asymptotically pivotal** or an **asymptotical pivot** if there exists a continuous random variable Q such that

$$Q_n \xrightarrow{d} Q \quad \text{as } n \rightarrow \infty$$

and Q has a distribution *not* depending on θ .

e.g. X_1, \dots, X_n iid $EX_i = \mu + \text{Var}(X_i) = \sigma^2$
 $\sqrt{n}(\bar{X}_n - \mu) \xrightarrow{d} N(0, \text{Var}(X_i) = \sigma^2)$ as $n \rightarrow \infty$ by CLT

Remark: $Q(X_1, \dots, X_n, \mu, \sigma^2) \equiv \frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \xrightarrow{d} \frac{1}{\sigma} N(0, \sigma^2) = N(0, \underbrace{\frac{1}{\sigma^2} \sigma^2}_1)$

Suppose $Q_n(X_1, \dots, X_n, \theta)$ is asymptotically pivotal quantity and, for some $-\infty \leq a \leq b \leq \infty$, $P(a \leq Q \leq b) = 1 - \alpha$ holds for some $0 < \alpha < 1$. Then, for large n ,

limit distribution
 e.g. $Q \equiv N(0, 1)$

$$C_X = \left\{ \theta : \theta \in \Theta, a \leq Q_n(X_1, \dots, X_n, \theta) \leq b \right\}$$

is a **confidence region** for θ with approximate CC $(1 - \alpha)$

The main idea is that, for large n ,

$$\begin{aligned} \min_{\theta \in \Theta} P_{\theta} \left(\theta \in C_X \right) &= \min_{\theta \in \Theta} P_{\theta} \left(a \leq Q_n(X_1, \dots, X_n, \theta) \leq b \right) \\ &\approx P(a \leq Q \leq b) \\ &= 1 - \alpha. \end{aligned}$$

$$L(\theta) = \prod_{i=1}^n \binom{m}{x_i} \theta^{x_i} (1-\theta)^{m-x_i}$$

Interval Estimation I

Asymptotically Pivotal Quantities: Illustration

("large n + CLT + Variance estimation = asymptotic pivot")

Example 1: Let X_1, \dots, X_n be iid Binomial(m, θ), $0 < \theta < 1$, with fixed m . Find an asymptotic pivot using the MLE $\hat{\theta}_n$ and find a corresponding CI for θ with approximate C.C. $1 - \alpha$.

Solution: $\hat{\theta}_n = \frac{\bar{X}_n}{m} = \frac{1}{n} \sum_{i=1}^n Y_i$ where $Y_i = \frac{X_i}{m}$ ($E Y_i = \theta$, $Var(Y_i) = \frac{(1-\theta)\theta}{m}$)

By CLT,

$$\sqrt{n}(\underbrace{\bar{Y}_n}_{\hat{\theta}_n} - \underbrace{\theta}_{E_0 Y_i}) \xrightarrow{d} N(0, Var_0(Y_i) = \frac{(1-\theta)\theta}{m}) \text{ as } n \rightarrow \infty.$$

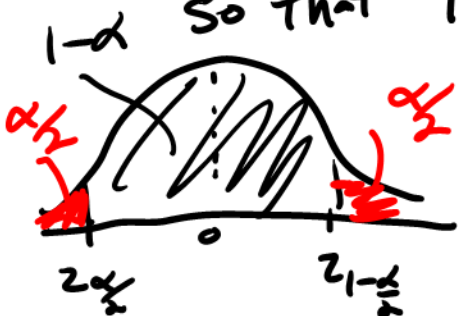
Letting $g(\hat{\theta}_n) = \sqrt{\frac{(1-\hat{\theta}_n)\hat{\theta}_n}{m}} \xrightarrow{P} g(\theta) \text{ as } n \rightarrow \infty$ by WLLN ($\bar{Y}_n \xrightarrow{P} E_0 Y_i$) & continuous mapping theorem.

Hence,

$$\frac{\sqrt{n}(\hat{\theta}_n - \theta)}{g(\hat{\theta}_n)} = Q_n(X_1, \dots, X_n, \theta) \xrightarrow{d} Q \equiv \frac{1}{\underbrace{\sqrt{\frac{\theta(1-\theta)}{m}}}_{g(\theta)}} \cdot N(0, \frac{\theta(1-\theta)}{m}) = N(0, \frac{1}{\frac{\theta(1-\theta)}{m}} \cdot \frac{\theta(1-\theta)}{m}) = N(0, 1)$$

Note use $Q \equiv Z \sim N(0, 1)$ & find a & b

so that $P(a \leq Z \leq b) = 1 - \alpha$.



$$C_X = \{ \alpha < 0 < 1: z_{\alpha/2} \leq \frac{\sqrt{n}(\hat{\theta}_n - \theta)}{g(\hat{\theta}_n)} \leq z_{1-\alpha/2} \}$$

$$= \left[\hat{\theta}_n - z_{1-\alpha/2} \frac{g(\hat{\theta}_n)}{\sqrt{n}}, \hat{\theta}_n + z_{\alpha/2} \frac{g(\hat{\theta}_n)}{\sqrt{n}} \right]$$

is a CI for θ with approximate C.C. $1 - \alpha$.

Interval Estimation I

Asymptotically Pivotal Quantities: Illustration

("large n + LRS + χ^2 -approximation = asymptotic pivot")

Example 2: Let $X_1, \dots, X_n, Y_1, \dots, Y_m$ be independent random variables where X_1, \dots, X_n are iid Exponential(θ), $\theta > 0$, and Y_1, \dots, Y_m are iid Exponential(λ), $\lambda > 0$. Find a large-sample confidence region for (θ, λ) , with approximate C.C. $1 - \alpha$, based on a likelihood ratio statistic. (LRS)