

# STAT 5430

Lec 38, W, Apr 30

- Homework 9 is assigned & due Sunday, May 4  
but you can submit this on Monday, May 5
- Exam 2 Solutions & grading key posted
- Final Exam on Tuesday, May 13, 7:30-9:30 PM
  - Comprehensive - but focus on material since Exam 2 (interval estimation)
  - Formula sheet for new material/interval & 2 formula sheets previous material
  - (3 sheets (front/back) total)
  - Practice Exams

see  
Canvas

check  $P_0(\theta \in [\alpha(T), \beta])$  for  $\theta < 0.1$   
numerically

```

M<-10000000
CC<-.90

theta<-0.01
M<-10000000
CC<-.90

T<-rgeom(M,theta)+1
alpha.2<-(1-CC)/2

LOWER<-1-(1-alpha.2)^{1/T}
UPPER<-1-(alpha.2)^{1/(T-1)}

COVER<-rep(1,M)
COVER[LOWER>theta]<-0
COVER[UPPER<theta]<-0
mean(COVER)

for(i in 1:n){
  theta<-t[i]
  print(i)
  T<-rgeom(M,theta)+1
  alpha.2<-(1-CC)/2

  LOWER<-1-(1-alpha.2)^{1/T}
  UPPER<-1-(alpha.2)^{1/(T-1)}

  COVER<-rep(1,M)
  COVER[LOWER>theta]<-0
  COVER[UPPER<theta]<-0

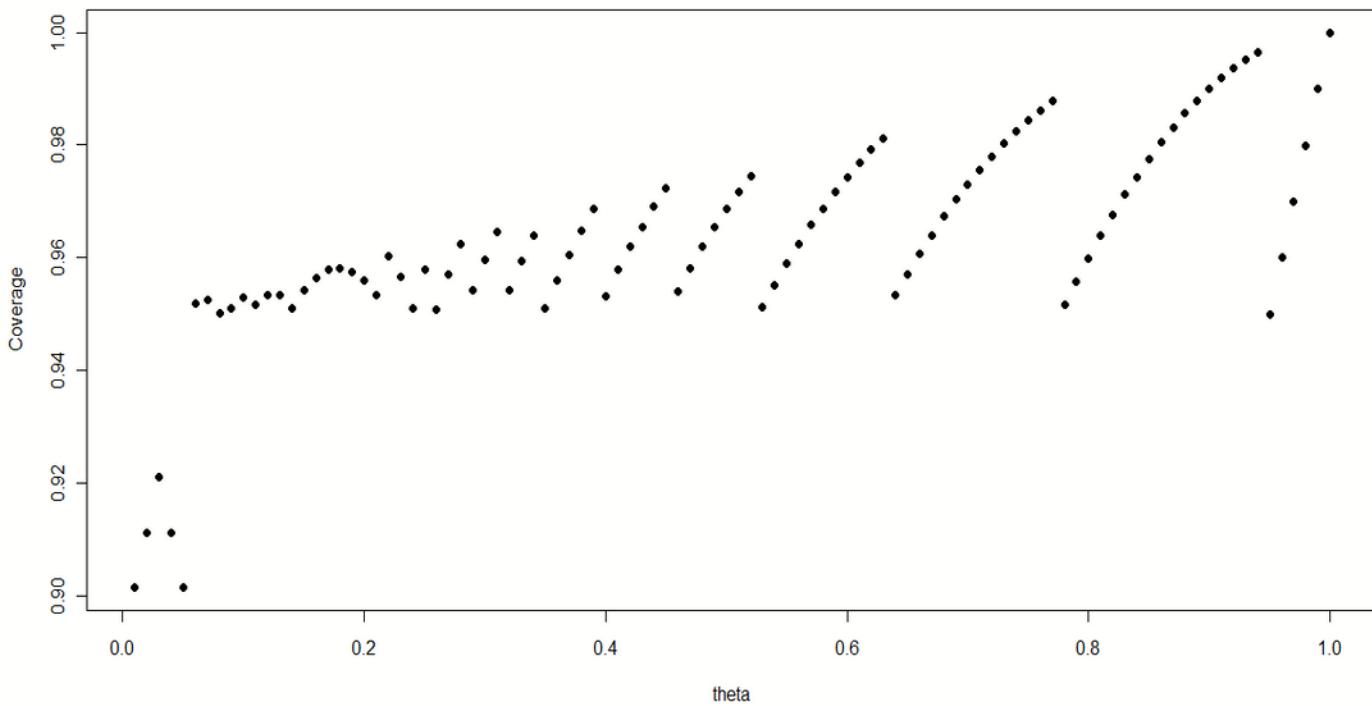
  t1[i]<-mean(COVER)
}

t1
  
```

```

plot(t,t1,xlab="theta",ylab="Coverage",main="Actual Coverage  
of 90% MBG Interval for Geometric Parameter theta",pch=20)
  
```

Actual Coverage of 90% MBG Interval for Geometric theta



## Interval Estimation II

### Bayes Intervals

*Definition:* Let  $\underline{X} = (X_1, \dots, X_n)$  have joint pdf/pmf  $f(\underline{x}|\theta)$ ,  $\theta \in \Theta \subset \mathbb{R}^p$  and let  $\pi(\theta)$  be a prior pdf for  $\theta$  on  $\Theta$ . Let  $f_{\theta|\underline{x}}(\theta)$  be the posterior pdf of  $\theta$  given  $\underline{X} = \underline{x}$ .

Then, a set  $C_{\underline{x}}$  is called a  $(1 - \alpha)$  **credible set** for  $\theta$  if

*posterior prob*  
or equivalently,  
*that*  
 $\theta \in C_{\underline{x}}$

$$\begin{array}{c} P(\theta \in C_{\underline{x}} | \underline{X} = \underline{x}) = 1 - \alpha, \quad \forall \underline{x} \\ \Downarrow \\ \int_{C_{\underline{x}}} f_{\theta|\underline{x}}(\theta) d\theta = 1 - \alpha, \quad \forall \underline{x}. \end{array}$$

*credible set*  
 $C_{\underline{x}} \subset \Theta$  is  
a guess for where  
 $\theta$  might be.

a credible set  $C_{\underline{x}}$  is NOT confidence region  
Not true that  $\min_{\theta \in C_{\underline{x}}} P_{\underline{x}}(\theta) = 1 - \alpha$

*Example:* Let  $X_1, \dots, X_n$  be iid Poisson( $\lambda$ ),  $\lambda > 0$ . Let the prior for  $\lambda$  on  $(0, \infty)$  be Exponential(1).  $\leftarrow \text{gamma}(1, 1)$

Then, the posterior dist. of  $\lambda$  given  $\underline{X}$  is  
 $\text{gamma}\left(1 + \sum_{i=1}^n x_i, (n+1)^{-1}\right)$  [check]

Hence, a  $(1-\alpha)$  credible set/interval for  $\lambda$  is given by  $[L, U]$  where

$$P(\lambda \in [L, U] | \underline{X}) = 1 - \alpha \quad (\text{posterior prob})$$

Note:  $\lambda | \underline{X} \sim \text{gamma}\left(1 + \sum_{i=1}^n x_i, (n+1)^{-1}\right)$

$$\Rightarrow \frac{2\lambda}{(n+1)^{-1}} | \underline{X} \sim \chi^2_{2(1 + \sum_{i=1}^n x_i)} \quad (\text{posterior dist})$$

$$\begin{aligned}
 \text{So, } 1-\alpha &= P(L \leq \lambda \leq U | X) \\
 &= P\left(\underbrace{\frac{2L}{(n+1)^2}}_{\chi^2_{2(\sum x_i)}; \frac{\alpha}{2}} \leq \underbrace{\frac{2\lambda}{(n+1)^2}}_{\chi^2_{2(\sum x_i + 1)}} \leq \underbrace{\frac{2U}{(n+1)^2}}_{\chi^2_{2(\sum x_i + 1); 1-\frac{\alpha}{2}}} | X\right)
 \end{aligned}$$

So, finally,  $[L, U]$ , where

$$L = \frac{\chi^2_{2(\sum x_i)}; \frac{\alpha}{2}}{2(n+1)} \quad \text{and} \quad U = \frac{\chi^2_{2(\sum x_i + 1); 1-\frac{\alpha}{2}}}{2(n+1)},$$

is a  $(1-\alpha)$  credible interval for  $\lambda$ .

## Interval Estimation II

### Bayes Intervals

*Definition:* A highest posterior density (HPD) credible set of level  $(1 - \alpha)$  is a set of the form

$$C_{\tilde{x}} = \{\theta \in \Theta : f_{\theta|\tilde{x}}(\theta) \geq C\}, \text{ for some } C > 0$$

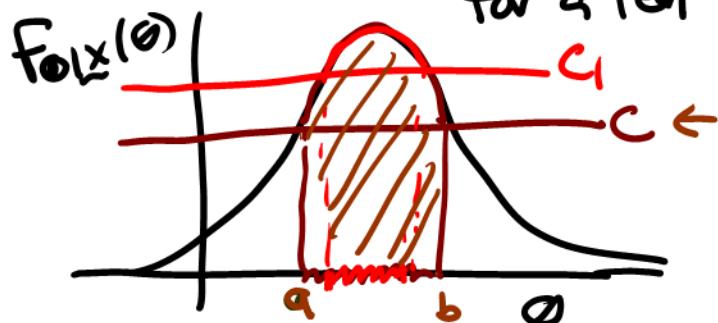
such that  $P(\theta \in C_{\tilde{x}} | X = \tilde{x}) = 1 - \alpha, \forall \tilde{x}$ .

posterior prob

has right post. prob  
to be a  $(1-\alpha)$  credibility set

*Discussion:* Why do this?

Consider posterior density  $f_{\theta|\tilde{x}}(\theta)$   
for a real-valued  $\theta$ , which is unimodal

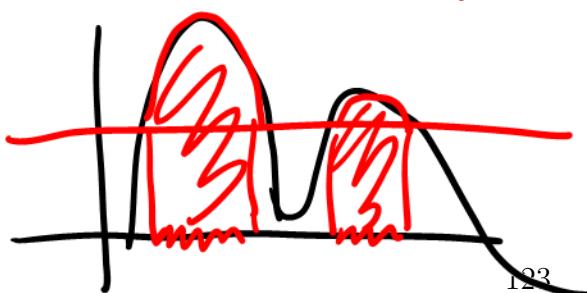


Pick  $C$  so that  $\int_a^b f_{\theta|\tilde{x}}(\theta) d\theta = 1 - \alpha \Rightarrow$  gives the HPD  
credible set as  $C_x = [a, b]$

Note: Where  $f_{\theta|\tilde{x}}(\theta)$  is "high", want to "pack in"  
an area of  $1 - \alpha$  over a short region of  $\theta$ .

So, HPD credible sets achieve  $(1 - \alpha)$  posterior coverage  
but tend to be small/informative sets for  $\theta$

(guesses for  $\theta$ )



## Interval Estimation II

Bayes HPD Intervals: Illustration

Example: Let  $X_1, \dots, X_n$  be iid  $N(\theta, \sigma^2)$  with  $\theta \in \mathbb{R}$  and known  $\sigma^2 > 0$ . Suppose a prior distribution for  $\theta$  is  $N(\mu, \tau^2)$  for some known  $\mu \in \mathbb{R}, \tau^2 > 0$ .

$\Rightarrow$  posterior distribution of  $\theta$  given  $\underline{x}$  is

$$\theta|\underline{x} \sim \text{Normal}(\underline{\mu}_{\theta|\underline{x}}, \underline{\sigma}_{\theta|\underline{x}}^2) \quad \text{unimodal}$$

$\uparrow \quad \rightarrow$   
depend on  $\underline{x}, n, \mu, \tau^2$   
(done this before)

Find  $\alpha(1-\alpha)$  HPD credible set for  $\theta$ :

$$C_{\underline{x}} = \left\{ \theta : f_{\theta|\underline{x}}(\theta) \geq C \right\} = \left\{ \theta : \frac{1}{\sqrt{2\pi} \underline{\sigma}_{\theta|\underline{x}}} e^{\frac{-(\theta - \underline{\mu}_{\theta|\underline{x}})^2}{2\underline{\sigma}_{\theta|\underline{x}}^2}} \geq C \right\}$$
$$= \left\{ \theta : \left| \frac{\theta - \underline{\mu}_{\theta|\underline{x}}}{\underline{\sigma}_{\theta|\underline{x}}} \right| \leq C_1 \right\}$$