

Multivariate distributions

Cumulative distribution functions

The joint cdf of X and Y is: $F_{X,Y}(x,y) = F(x,y) = P(X \leq x, Y \leq y)$, for $x, y \in \mathbb{R}$

1. rarely used in the discrete case but is computed as

$$F(x,y) = P(X \leq x, Y \leq y) = \sum_{x_1 \leq x, y_1 \leq y} f(x_1, y_1)$$

2. may be useful in the continuous case and is computed as

$$\textcircled{X} \quad F(x,y) = P(X \leq x, Y \leq y) = \int_{-\infty}^y \int_{-\infty}^x f(s,t) ds dt$$

where also in the continuous case

$$\frac{\partial^2 F(x,y)}{\partial x \partial y} = \frac{d^2 F(x,y)}{dx dy} = f(x,y).$$

Properties: A function $F(x,y)$ is a cdf for some r.v. (X, Y) if and only if

- ✓ 1. $\lim_{x \rightarrow -\infty} F(x,y) = \lim_{y \rightarrow -\infty} F(x,y) = 0$ $\lim_{x \rightarrow -\infty} F(x,y) = \lim_{x \rightarrow -\infty} \int_{-\infty}^x \int_{-\infty}^y f(s,t) ds dt = 0$
- ✓ 2. $\lim_{x,y \rightarrow \infty} F(x,y) = 1$ $F(x,y) = \lim_{(x,y) \rightarrow \infty} \int_{-\infty}^x \int_{-\infty}^y f(s,t) ds dt = 1$
- 3. Right continuous in each argument: $\lim_{h \downarrow 0} F(x+h, y) = \lim_{h \downarrow 0} F(x, y+h) = F(x, y)$
- 4. “nondecreasing” (the probability assigned to any rectangle is ≥ 0)

$$\left\{ \begin{array}{l} P(x < X \leq x + \Delta_1, y < Y \leq y + \Delta_2) \\ = F(x + \Delta_1, y + \Delta_2) - F(x + \Delta_1, y) - F(x, y + \Delta_2) + F(x, y) \geq 0 \end{array} \right.$$

for all $x, y \in \mathbb{R}$ and for any $\Delta_1, \Delta_2 > 0$

$$\begin{aligned} P(x < X \leq x + \Delta_1, y < Y \leq y + \Delta_2) &\stackrel{\text{def}}{=} \int_x^{x+\Delta_1} \int_y^{y+\Delta_2} f(s,t) ds dt \\ &= \int_x^{x+\Delta_1} \left[\int_{-\infty}^{y+\Delta_2} f(s,t) ds - \int_{-\infty}^y f(s,t) ds \right] dt = \dots \end{aligned}$$

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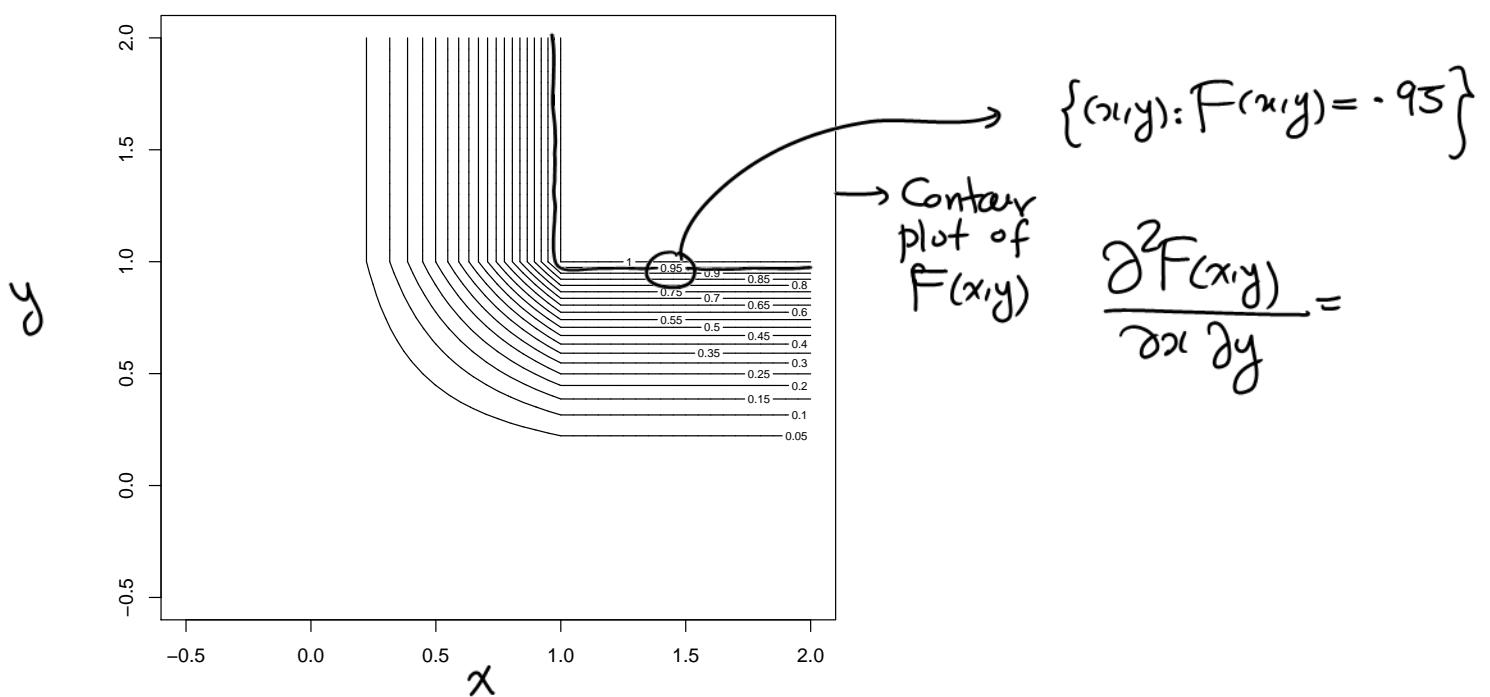
Cumulative distribution functions (cont'd)

e.g., $F(x, y) = \begin{cases} 0 & \text{if } x + y < -1 \\ 1 & \text{if } x + y \geq -1 \end{cases}$ is *not* a valid cdf, though

Properties 1-3 hold, and $F(x, y)$ is nondecreasing in its arguments x or y (but “probabilities of rectangles” are not always non-negative). \equiv

Cdf example (continuous case): $f(x, y) = 4xy$, for $0 < x, y < 1$

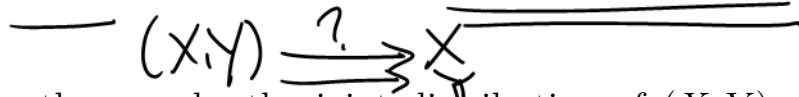
$$F(x, y) = \int_{-\infty}^x \int_{-\infty}^y f(s, t) ds dt = \begin{cases} 1 & x \geq 1, y \geq 1 \\ y^2 & x \geq 1, 0 < y < 1 \\ x^2 & 0 < x < 1, y \geq 1 \\ x^2 y^2 & 0 < x < 1, 0 < y < 1 \\ 0 & \text{otherwise} \end{cases}$$



Multivariate distributions

Marginal distributions

- Marginal distribution of X (or Y) describes the distribution of a single random variable that is implied by the joint distribution



In other words, the joint distribution of (X, Y) must determine individual distributions for \underline{X} and \underline{Y}

- For discrete and continuous distributions, we often get the marginal distribution from the joint pmf or pdf $f(x, y)$

$$P(A) = P(\{X=x, Y=y_1\} \cup \{X=x, Y=y_2\} \cup \{X=x, Y=y_3\} \cup \dots \cup \{X=x, Y=y_n\})$$

Marginal Distributions in Discrete Case:

$$\xrightarrow{\text{Pmf}} f_X(x) = P(X = x) = \sum_y P(X = x, Y = y) = \sum_y f(x, y), \quad x \in \mathbb{R}$$

↑
fix/given

similarly,

$$\xrightarrow{} f_Y(y) = P(Y = y) = \sum_x P(X = x, Y = y) = \sum_x f(x, y), \quad y \in \mathbb{R}$$

↑
fix/given

$$P(A) = P(A \cap B) + \dots + P(A \cap B_n) = \sum_i P(A \cap B_i)$$

Discrete Example (artificial): (X, Y) with joint pmf given in tabular form as

		x		
		1	2	3
y	3	1/12	1/12	1/6
	2	1/12	1/6	1/12
	1	1/6	1/12	1/12

$$f_X(x) = P(X = x)$$

$$= \begin{cases} 1/12 + 1/12 + 1/6 & x = 1 \\ 1/12 + 1/6 + 1/12 & x = 2 \\ 1/6 + 1/12 + 1/12 & x = 3 \end{cases}$$

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Discrete marginal distributions: another example

Recall previous discrete example:

- Suppose there are n independent trials, where each trial has three possible outcomes: outcome a with prob. p_1 , outcome b with prob. p_2 and outcome c with prob. $1 - p_1 - p_2$
- Let $X = \#$ of trials having outcome a & $Y = \#$ of trials having outcome b
- Then the joint pmf of (X, Y) is

$$f(x, y) = \frac{n!}{x!y!(n-x-y)!} p_1^x p_2^y (1-p_1-p_2)^{n-x-y}$$

for $0 \leq x, y, x+y \leq n$ (assuming $0 < p_1, p_2, p_1 + p_2 < 1$)

- Find the marginal pmf of X . Fix a x value in $\{0, 1, \dots, n\}$.

$$\begin{aligned}
 P(X=x) &= f_X(x) = \sum_{y=0}^{n-x} P(X=x, Y=y) \\
 &\stackrel{\text{fix}}{=} \sum_{y=0}^{n-x} \frac{n!}{x!y!(n-x-y)!} p_1^x p_2^y (1-p_1-p_2)^{n-x-y} \\
 &= \frac{n!}{x!(n-x)!} p_1^x (1-p_1)^{n-x} \sum_{y=0}^{n-x} \frac{(n-x)!}{y!(n-x-y)!} \frac{p_2^y (1-p_1-p_2)^{n-x-y}}{(1-p_1)^{n-x}} \\
 &= \frac{n!}{x!(n-x)!} p_1^x (1-p_1)^{n-x} \sum_{y=0}^{n-x} \frac{(n-x)!}{y!(n-x-y)!} \left(\frac{p_2}{1-p_1}\right)^y \left(\frac{1-p_1-p_2}{1-p_1}\right)^{n-x-y} \\
 \Rightarrow P(X=x) &= \binom{n}{x} p_1^x (1-p_1)^{n-x} \\
 \Rightarrow X &\sim \text{Bin}(n, p_1)
 \end{aligned}$$

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$$\begin{aligned}
 &= \sum_{y=0}^{n-x} \binom{n-x}{y} \tilde{p}^y (1-\tilde{p})^{n-x-y} \\
 &\quad \tilde{p} := \frac{p_2}{1-p_1} \\
 &\quad 1-\tilde{p} = 1 - \left[\frac{p_2}{1-p_1}\right] = \frac{1-p_1-p_2}{1-p_1} = 1
 \end{aligned}$$

→ $\textcircled{*}$ is the pmf of
 $\text{Bin}(n-x, \frac{P_2}{1-P_1})$

Multivariate distributions

Marginal distributions: continuous case

In the continuous (X, Y) case, note that, for $x \in \mathbb{R}$,

$$P(X \leq x) = P((X, Y) \in (-\infty, x] \times (-\infty, \infty)) = \int_{-\infty}^x \int_{-\infty}^{\infty} f(x, y) dy dx$$

$$\underbrace{\int_{-\infty}^x g(t) dt}_{\substack{g(t) \\ -\infty}} \equiv h(x)$$

Marginal Distributions in Continuous Case:

$$\overrightarrow{\underline{f_X(x)}} = \int_{-\infty}^{\infty} f(x, y) dy, \quad \text{for given } x \in \mathbb{R}$$

$$\underline{f_Y(y)} = \int_{-\infty}^{\infty} f(x, y) dx, \quad \text{for given } y \in \mathbb{R}$$

We often speak of “integrating out” x or y to obtain a marginal pdf

Continuous example: joint pdf $f(x, y) = 1/x$, for $0 < y < x < 1$

$$f_X(x) \stackrel{\text{def}}{=} \int_{-\infty}^{\infty} f(x, y) dy = \int_0^x \frac{1}{x} dy = \left[\frac{1}{x} y \right]_0^x = 1 \quad \forall x \in (0, 1)$$

$$\left\{ \begin{array}{l} f_Y(y) \stackrel{\text{def}}{=} \int_{-\infty}^{\infty} f(x, y) dx = \int_y^1 \frac{1}{x} dx = \left[\log x \right]_y^1 = \log 1 - \log y \\ f_Y(y) = \begin{cases} -\log y & y \in (0, 1) \\ 0 & \text{o.w.} \end{cases} \end{array} \right.$$

$$f_X(x) = \begin{cases} 1 & x \in (0, 1) \\ 0 & \text{o.w.} \end{cases} \quad \forall y \in (0, 1)$$

Multivariate distributions

From bivariate to multivariate

Now consider the n -dimensional case: \mathbb{R}^n -valued r.v. (X_1, \dots, X_n)

Cdf: $F(x_1, \dots, x_n) = P(\underline{X_1 \leq x_1, \dots, X_n \leq x_n})$ for $x_1, \dots, x_n \in \mathbb{R}$

1. $\underset{\text{all } x_i \rightarrow \infty}{\lim} F(x_1, \dots, x_n) = 1$
2. $\underset{\text{any } x_i \rightarrow -\infty}{\lim} F(x_1, \dots, x_n) = 0$
3. right continuous in each argument
4. “monotonicity” condition from bivariate case generalized to n -dim rectangle

Pdf or pmf: $f(x_1, \dots, x_n) \geq 0$ with integral (or sum) equal to 1

$$\underbrace{P((X_1, \dots, X_n) \in A)}_{\{(x_1, \dots, x_n) \in A\}} = \int \cdots \int f(x_1, \dots, x_n) dx_1 \dots, dx_n, \quad A \subset \mathbb{R}^n$$

$$P((X_1, \dots, X_n) \in A) = \sum_{(x_1, \dots, x_n) \in A} f(x_1, \dots, x_n)$$

Marginal pdf or pmf for *any* subset of $(X_1, \dots, X_n) \rightarrow n - \text{R.V.}$

e.g., if you want the marginal distribution of (X_1, \dots, X_k) from (X_1, \dots, X_n) , fix values of $x_1, \dots, x_k \in \mathbb{R}$, and sum or integrate out the remaining variables

$$\begin{aligned} \underbrace{f(x_1, \dots, x_k)}_{\text{f(x1, ..., xk)}} &= \int \cdots \int \underbrace{f(x_1, \dots, x_n)}_{\text{f(x1, ..., xn)}} dx_{k+1} \dots, dx_n \\ f(x_1, \dots, x_k) &= \sum_{(x_{k+1}, \dots, x_n)} f(x_1, \dots, x_n) \end{aligned}$$

→ We'll still mostly focus on the bivariate case, though the same ideas apply to more than 2 variables (though more computationally involved than we presently need)

Multivariate distributions

Expectations of several random variables

$$\text{Recall: } n=1, \mathbb{E}[g(X)] = \int g(x) f_X(x) dx$$

Provided $\int |g(x)| f_X(x) dx < \infty$

Definition: extension of the univariate definition

If (X_1, \dots, X_n) are jointly discrete with pmf f and $g : \mathbb{R}^n \rightarrow \mathbb{R}$, we define

$$\longrightarrow \mathbb{E}g(X_1, \dots, X_n) = \sum_{(x_1, \dots, x_n)} g(x_1, \dots, x_n) f(x_1, \dots, x_n)$$

$\sum_{x_1} \dots \sum_{x_n} g(x_1, \dots, x_n) \mathbb{P}(X_1=x_1, X_2=x_2, \dots, X_n=x_n)$

provided that

$$\sum_{(x_1, \dots, x_n)} |g(x_1, \dots, x_n)| f(x_1, \dots, x_n) < \infty$$

If (X_1, \dots, X_n) are jointly continuous with pmf f and $g : \mathbb{R}^n \rightarrow \mathbb{R}$, we define

$$\mathbb{E}g(X_1, \dots, X_n) = \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} g(x_1, \dots, x_n) f(x_1, \dots, x_n) dx_1 \dots dx_n$$

provided that

$$\int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} |g(x_1, \dots, x_n)| f(x_1, \dots, x_n) dx_1 \dots dx_n < \infty$$

In the bivariate case, these become

$n=2$

 $\mathbb{E}g(X, Y) = \begin{cases} \sum_{(x,y)} g(x, y) f(x, y) & \text{discrete case} \\ \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f(x, y) dx dy & \text{continuous case} \end{cases}$

Multivariate distributions

Expectations of several random variables (cont'd)

Theoretical properties as before

e.g., if $a_1, \dots, a_k, b \in \mathbb{R}$ and each $g_i : \mathbb{R}^n \rightarrow \mathbb{R}$

$$\mathbb{E} \left(\sum_{i=1}^k a_i g_i(X_1, \dots, X_n) + b \right) = b + \sum_{i=1}^k a_i \mathbb{E} g_i(X_1, \dots, X_n)$$

Marginal moments:

$$\mathbb{E}X_i = \mu_{X_i} = \begin{cases} \sum_{(x_1, \dots, x_n)} x_i f(x_1, \dots, x_n) & \text{discrete case} \\ \int \cdots \int \underbrace{x_i f(x_1, \dots, x_n)}_{\uparrow} dx_1 \dots, dx_n & \text{continuous case} \end{cases}$$

$$\mathbb{E}X = \iint x f(x,y) dx dy \quad \mathbb{E}Y = \iint y f(x,y) dx dy$$

Discrete Example: (X, Y) with joint pmf given in tabular form as

		x
		1 2 3
		1/12 1/12 1/6
y		1/12 1/6 1/12
2		1/6 1/12 1/12
1		1/12 1/12 1/6