

STAT 5430

Lec 33, F , Apr 18

(No new homework)

## Interval Estimation I

### Pivotal Quantities

$\downarrow$  Vector of parameters  $\theta$

*Definition:* Let  $X_1, \dots, X_n$  be joint pdf/pmf  $f(x|\theta)$ ,  $\theta \in \Theta \subset \mathbb{R}^p$ . Then a random variable  $Q(X, \theta)$  is called a **pivot** or **pivotal quantity** if the distribution of  $Q(X, \theta)$  under  $\theta$  does not depend on  $\theta$ .

**Note:**  $Q(X, \theta)$  is NOT a statistic (because can depend on  $\theta$ )

$$P_{\theta}(Q(X, \theta) \in A) = P(Q(X, \theta) \in A)$$

*Some examples:* (pivots, unlike statistics, can be functions of parameters  $\theta$ )

$\frac{\bar{X}_1 - \bar{X}_2}{\bar{X}_3 - \bar{X}_4}$  ancillary statistic

1. Let  $X_1 \dots X_n$  be iid  $N(\mu, \sigma^2)$  random variables.

$$\bar{X}_n \sim N(\mu, \frac{\sigma^2}{n})$$

pivot  $\left\{ \frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}} \sim N(0, 1) \right.$  pivot  $\left\{ \frac{\bar{X}_n - \mu}{S/\sqrt{n}} \sim T_{n-1}$  distribution

2. Let  $f_0$  be a pdf on  $\mathbb{R}$ . Let  $X_1 \dots X_n$  be iid with random variables common pdf  $f(x|\theta)$  where

$$\theta = (\theta_1, \theta_2) \quad f(x|\theta) = \frac{1}{\theta_2} f_0\left(\frac{x - \theta_1}{\theta_2}\right), \quad x \in \mathbb{R},$$

for  $\theta = (\theta_1, \theta_2)$ ,  $\theta_1 \in \mathbb{R}$  (location parameter) and  $\theta_2 > 0$  (scale parameter).

Then,  $Q(X, \theta) = \frac{\bar{X}_n - \theta_1}{\theta_2}$   $\leftarrow$  pivot

Why? Note:  $Y_i \equiv \frac{X_i - \theta_1}{\theta_2} \sim f_0(y)$  (i.e.  $Y_1, \dots, Y_n$  iid  $f_0(y)$ )

and  $Q(X, \theta) = \frac{1}{n} \sum_{i=1}^n Y_i \leftarrow$  distribution doesn't depend on  $\theta$

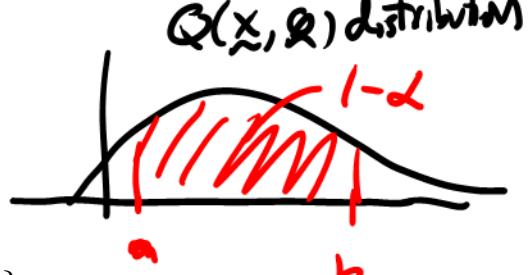
# Interval Estimation I

## Interval Estimation via Pivotal Quantities

**Remarks:**

- Let  $Q(\tilde{X}, \theta)$  be a pivotal quantity ( $\theta \in \Theta \subset \mathbb{R}^p$ ) and  $0 < \alpha < 1$ . Suppose  $-\infty \leq a \leq b \leq \infty$  are such that

$$P(a \leq Q(\tilde{X}, \theta) \leq b) = P_\theta(a \leq Q(\tilde{X}, \theta) \leq b) = 1 - \alpha.$$



Then,

$$C_{\tilde{X}} = \{\theta : \theta \in \Theta, a \leq Q(\tilde{X}, \theta) \leq b\}$$

is a **confidence region** for  $\theta$  with CC  $(1 - \alpha)$

That is,  $\min_{\theta \in \Theta} P_\theta(\theta \in C_{\tilde{X}}) = \min_{\theta \in \Theta} P_\theta(a \leq Q(\tilde{X}, \theta) \leq b) = 1 - \alpha.$

$$\underbrace{P(a \leq Q(\tilde{X}, \theta) \leq b)}_{1 - \alpha}$$

- If  $\Theta \subset \mathbb{R}$  and  $Q(\tilde{X}, \theta)$  is monotone in  $\theta \in \mathbb{R}$ , then the region  $C_{\tilde{X}}$  will be an interval.

## Interval Estimation I

Interval Estimation via Pivotal Quantities

*Example:* Let  $X_1 \dots X_n$  be iid  $\text{Gamma}(\delta_0, \theta)$  where  $\theta > 0$  ( $\delta_0$  fixed/known). Using a pivotal quantity based on  $\sum_{i=1}^n X_i$ , find a CI for  $\theta$  with C.C.  $1 - \alpha$ .

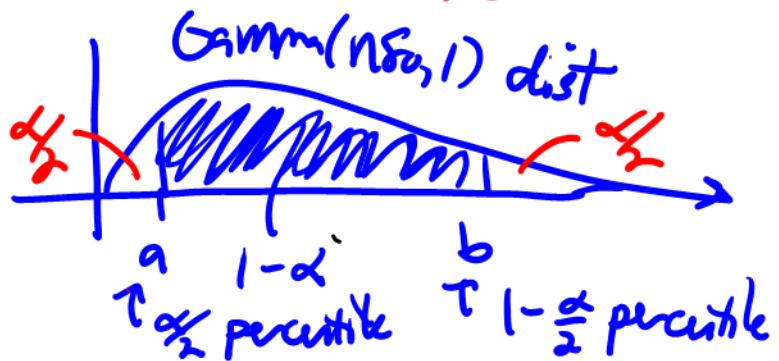
Solution:  $\sum_{i=1}^n X_i \sim \text{Gamma}(n\delta_0, \theta)$

$$Q(\underline{x}, \theta) = \frac{\sum_{i=1}^n X_i}{\theta} \sim \text{Gamma}(n\delta_0, 1)$$

$\tau_{\text{pivot}}$

$\tau$  dist. doesn't depend on  $\theta$

Find  $a$  &  $b$  such that



Confidence interval for  $\theta > 0$

$$\begin{aligned} \text{is } \{ \theta > 0 : a \leq Q(\underline{x}, \theta) \leq b \} &= \{ \theta : a \leq \frac{\sum X_i}{\theta} \leq b \} \\ &= \{ \theta > 0 : \frac{\sum X_i}{b} \leq \theta \leq \frac{\sum X_i}{a} \} \\ &= \left[ \frac{\sum X_i}{b}, \frac{\sum X_i}{a} \right] \end{aligned}$$

# Interval Estimation I

## Asymptotically Pivotal Quantities

*Definition:* A sequence of random variables  $Q_n \equiv Q_n(X_1, \dots, X_n, \theta)$ ,  $n \geq 1$ , is called **asymptotically pivotal** or an **asymptotical pivot** if there exists a continuous random variable  $Q$  such that

$$Q_n \xrightarrow{d} Q \quad \text{as } n \rightarrow \infty$$

and  $Q$  has a distribution *not* depending on  $\theta$ .

e.g.  $X_1, \dots, X_n$  iid  $E[X_i] = \mu$  &  $\text{Var}(X_i) = \sigma^2$   
 $\sqrt{n}(\bar{X}_n - \mu) \xrightarrow{d} N(0, \text{Var}(X_i) = \sigma^2)$  as  $n \rightarrow \infty$  by CLT  
 $Q(X_1, \dots, X_n, \mu, \sigma^2) \equiv \frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \xrightarrow{d} N(0, 1) = N(0, \underbrace{\sigma}_{\sqrt{\sigma^2}})$

Remark: Suppose  $Q_n(X_1, \dots, X_n, \theta)$  is asymptotically pivotal quantity and, for some  $-\infty \leq a \leq b \leq \infty$ ,  $P(a \leq Q_n \leq b) = 1 - \alpha$  holds for some  $0 < \alpha < 1$ . Then, for large  $n$ ,

*limit distribution*  $\xrightarrow{\text{red arrow}}$   
e.g.  $Q \equiv N(0, 1)$   
 $C_X = \left\{ \theta : \theta \in \Theta, a \leq Q_n(X_1, \dots, X_n, \theta) \leq b \right\}$   
is a **confidence region** for  $\theta$  with approximate CC  $(1 - \alpha)$

The main idea is that, for large  $n$ ,

$$\begin{aligned} \min_{\theta \in \Theta} P_{\theta} \left( \theta \in C_X \right) &= \min_{\theta \in \Theta} P_{\theta} \left( a \leq Q_n(X_1, \dots, X_n, \theta) \leq b \right) \\ &\approx P(a \leq Q \leq b) \\ &= 1 - \alpha. \end{aligned}$$

$$L(\theta) = \prod_{i=1}^n \binom{m}{x_i} \theta^{x_i} (1-\theta)^{m-x_i}$$

## Interval Estimation I

Asymptotically Pivotal Quantities: Illustration

("large  $n$  + CLT + Variance estimation = asymptotic pivot")

Example 1: Let  $X_1, \dots, X_n$  be iid Binomial( $m, \theta$ ),  $0 < \theta < 1$ , with fixed  $m$ . Find an asymptotic pivot using the MLE  $\hat{\theta}_n$  and find a corresponding CI for  $\theta$  with approximate C.C.  $1 - \alpha$ .

Solution:  $\hat{\theta}_n = \bar{X}_n = \frac{1}{n} \sum_{i=1}^n Y_i$ ; where  $Y_i = \frac{X_i}{m}$  ( $E[Y_i] = \theta$ ,  $V[\theta](Y_i) = \frac{(1-\theta)\theta}{m}$ )

By CLT,

$\sqrt{n} (\bar{Y}_n - \theta) \xrightarrow{d} N(0, V[\theta](Y_i) = \frac{(1-\theta)\theta}{m})$  as  $n \rightarrow \infty$ .

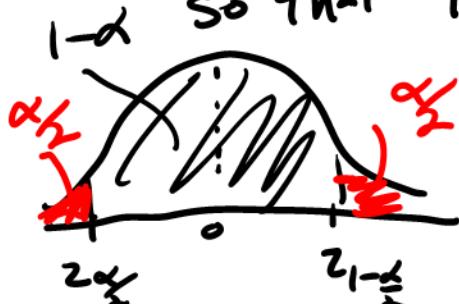
Letting  $g(\hat{\theta}_n) = \sqrt{\frac{(1-\hat{\theta}_n)\hat{\theta}_n}{m}}$   $\xrightarrow{P} g(\theta)$  as  $n \rightarrow \infty$  by WLLN ( $\bar{Y}_n \xrightarrow{P} E[\theta](Y_i)$ ) & continuous mapping theorem.

Hence,

$$\begin{aligned} \frac{\sqrt{n}(\hat{\theta}_n - \theta)}{g(\hat{\theta}_n)} &\equiv Q_n(X_1, \dots, X_n, \theta) \\ &\xrightarrow{d} Q \equiv \frac{1}{\sqrt{\frac{\theta(1-\theta)}{m}}} \cdot N(0, \frac{\theta(1-\theta)}{m}) \\ &\quad \text{using } g(\theta) = \sqrt{\frac{\theta(1-\theta)}{m}} \\ &= N(0, \frac{1}{\theta(1-\theta)} \frac{\theta(1-\theta)}{m}) \\ &= N(0, 1) \end{aligned}$$

Now use  $Q \sim N(0, 1)$  & find  $a, b$

so that  $P(a \leq Z \leq b) = 1 - \alpha$ .



$$C_X = \{ \theta < 1 : z_{\alpha/2} \leq \frac{\sqrt{n}(\hat{\theta}_n - \theta)}{g(\hat{\theta}_n)} \leq z_{1-\alpha/2} \}$$

$$= \left[ \hat{\theta}_n - z_{1-\alpha/2} \frac{g(\hat{\theta}_n)}{\sqrt{n}}, \hat{\theta}_n + z_{\alpha/2} \frac{g(\hat{\theta}_n)}{\sqrt{n}} \right]$$

is a CI for  $\theta$  with approximate C.C.  $1 - \alpha$ .

## Interval Estimation I

Asymptotically Pivotal Quantities: Illustration

(“large  $n + LRS + \chi^2$ -approximation = asymptotic pivot”)

Example 2: Let  $X_1, \dots, X_n, Y_1, \dots, Y_m$  be independent random variables where  $X_1, \dots, X_n$  are iid  $\text{Exponential}(\theta)$ ,  $\theta > 0$ , and  $Y_1, \dots, Y_m$  are iid  $\text{Exponential}(\lambda)$ ,  $\lambda > 0$ . Find a large-sample confidence region for  $(\theta, \lambda)$ , with approximate C.C.  $1 - \alpha$ , based on a likelihood ratio statistic. (LRS)