

## Expected values

Some properties

**Theorem 2.2.5:** Suppose  $X$  is a r.v. such that  $E|g_1(X)| < \infty$  and  $E|g_2(X)| < \infty$  and let  $a, b, c \in \mathbb{R}$  be fixed constants. Then,

1.  $E[ag_1(X) + b] = aEg_1(X) + b$

2.  $E[ag_1(X) + bg_2(X) + c] = aEg_1(X) + bEg_2(X) + c$

3. If  $g_1(x) \geq a$  for all  $x$ , then  $Eg_1(X) \geq a$

4. If  $\underline{g_1(x)} \leq b$  for all  $x$ , then  $E\underline{g_1(X)} \leq b$

5. If  $g_1(x) \geq g_2(x)$  for all  $x$ , then  $Eg_1(X) \geq Eg_2(X)$

$$E[a_1g_1(X) + a_2g_2(X) + c] \stackrel{\text{def}}{=} \int_{-\infty}^{\infty} h(x)f_X(x)dx = \int_{-\infty}^{\infty} \{a_1g_1(x) + a_2g_2(x)\}f_X(x)dx + c$$

$$E(h(X)) = \int_R h(x)f(x)dx = a_1 \int_{-\infty}^{\infty} g_1(x)f_X(x)dx + a_2 \int_{-\infty}^{\infty} g_2(x)f_X(x)dx$$

$$E(X) = \int_R x f_X(x) dx$$

$$(5) \quad g_1(x) \geq g_2(x) \Rightarrow \int_{-\infty}^{\infty} g_1(x)f_X(x)dx \geq \int_{-\infty}^{\infty} g_2(x)f_X(x)dx \Rightarrow E(g_1(X)) \geq E(g_2(X))$$

Expectations are also invariant under transformation:

If  $\tilde{Y} = g(X)$ , then

$$\begin{aligned} \tilde{Y} &= g(x) \\ EY &\stackrel{\text{def}}{=} \sum_y y f_Y(y) dy = \sum_y y P(Y=j) \\ &= \sum_x g(x) f_X(x) dx = Eg(X) \end{aligned}$$

(In the continuous case, replace sums with integrals)

$$E(Y) = \int_{-\infty}^{\infty} y f_Y(y) dy = \int_{-\infty}^{\infty} g(x) f_X(x) dx \quad \begin{cases} Y = g(X) \\ E(Y) \end{cases}$$

## Expected values

### Variance

An important instance of this  $Eg(X)$  notion comes using  $g(X) = [X - EX]^2$

*Definition:* The variance of a random variable  $X$ , denoted  $\text{Var}(X)$  or  $\sigma_X^2$ , is

$$\text{Var}(X) = \sigma_X^2 = E[X - EX]^2, = E[(X - E(X))^2]$$

the expected squared distance between  $X$  and its mean  $EX$

Two important “variance” facts:

1.  $\text{Var}(a + bX) = b^2\text{Var}(X)$  for any real numbers  $a, b$

2.  $\text{Var}(X) = EX^2 - [EX]^2$

Proof:  $\text{Var}(X) \stackrel{\text{def}}{=} E[(X - E(X))^2] \stackrel{E(X) = \mu}{=} E[(X - \mu)^2]$

$$= E[X^2 + \mu^2 - 2\mu X] = E[X^2] + \mu^2 - 2\mu E(X)$$

$$= E[X^2] - \mu^2 = E[X^2] - (E(X))^2.$$

Example:  $X$  is uniform(0, 1). Find  $\text{Var}(X)$  and  $\text{Var}(Y)$  for  $Y = 1 + 3X^2$ .

## Expected values

Other moments and distributional summaries

Moments are an important summary of a distribution

1.  $\mu = \mu_X = EX$  is often called the mean
2.  $\mu'_n = EX^n$  is the  $n$ th moment provided  $EX^n$  exists, i.e.,

$$\sum_x |g(x)|f_X(x) < \infty \quad (\text{discrete case})$$
$$\int_{-\infty}^{\infty} |g(x)|f_X(x)dx < \infty \quad (\text{continuous case})$$

3.  $\mu_n = E[(X - \mu)^n]$  is the  $n$ th central moment provided  $EX^n$  exists
  - (a)  $\text{Var}(X) = \sigma_X^2 = E[(X - \mu)^2] = \mu_2$  is the variance
  - (b)  $\sigma_X = \sqrt{\text{Var}(X)}$  is the standard deviation
  - (c)  $\mu_3$  is skewness (i.e., measures distributional balance around  $\mu$ )
  - (d)  $\mu_4$  is kurtosis (i.e., measure of how long the distributional tails are)

Regarding moments:

1. If  $EX^r$  exists for some  $r > 0$  then  $EX^s$  exists for  $0 \leq s \leq r$
2. If  $EX^r$  does not exist for some  $r > 0$ , then  $EX^s$  will not exist for  $s > r$
3.  $EX^2$  exists if and only if  $\text{Var}(X)$  exists
4. For  $r > 0$ , the existence of  $EX^r$  is a matter of the distribution of  $X$  not having “heavy tails” (i.e.,  $X$  doesn’t assume “large” values with “large” probability)