

STAT 5430

Lecture 14, F, Feb 21

- No new homework!
- Solutions to Homeworks 1-3 posted.
- Exam 1 is scheduled for W, Feb 26
6:15-8:15 PM (Sned seminar room) (two weeks)
3105
- No regular class on W, Feb 26
- See Canvas for study guide, practice exams
- Can bring 1 page formula sheet (front/back) with anything on it
- see Canvas for a "canned" sheet
- I'll provide table with STAT 5430 distributions (see Canvas)

STAT 5430: Summary to date

Where we have been & where we are headed

- Completed
 - Introduction to Statistical Inference
 - Point Estimation
 - * MME/MLE
 - Criteria for Evaluating Point Estimators
 - * bias, variance, UMVUE, MSE
 - Elements of Decision Theory
 - * Minimax, finding Bayes estimators
- Next: Sufficiency and Point Estimation
 - Sufficiency/Data Reduction
 - Factorization Theorem
 - Rao-Blackwell Theorem
 - Completeness/Lehman-Scheffe Theorem/UMVUE
 - Exponential Families

Sufficiency and Point Estimation (Chapter 6)

Sufficiency as Data Reduction

Definition: Let X_1, \dots, X_n be r.v.'s with joint pdf/pmf $f(\underline{x}|\theta)$, $\theta \in \Theta \subset \mathbb{R}^p$ and let $\underline{S} \equiv (S_1, \dots, S_k)$ be a vector of estimators. Then, \underline{S} is called (jointly) **sufficient** for θ if the conditional distribution of (X_1, \dots, X_n) given \underline{S} does *not* depend on θ .

Example: Let X_1, \dots, X_n be iid Geometric(θ), $0 < \theta < 1$. Show that $S \equiv X_1 + \dots + X_n$ is sufficient for θ .

Solution: conditional pmf of (X_1, \dots, X_n) given $S=s$ is

$$P_\theta(X_1=x_1, \dots, X_n=x_n | S=s) = \frac{P_\theta(X_1=x_1, \dots, X_n=x_n, S=s)}{P_\theta(S=s)}$$

$$= \begin{cases} \frac{P_\theta(X_1=x_1, \dots, X_n=x_n, S=s)}{P_\theta(S=s)} & \text{if } x_1+x_2+\dots+x_n=s \\ 0 & \text{O.W.} \end{cases}$$

↑ Neg-Binomial(n, θ)
* of trials until "n" successes

$$= \begin{cases} \frac{P_\theta(X_1=x_1, \dots, X_n=x_n)}{\binom{s-1}{n-1} \theta^n (1-\theta)^{s-n}} & \text{if } x_1+x_2+\dots+x_n=s \\ 0 & \text{O.W.} \end{cases}$$

$$= \begin{cases} \frac{P_\theta(X_1=x_1) \cdots P_\theta(X_n=x_n)}{\binom{s-1}{n-1} \theta^n (1-\theta)^{s-n}} & \text{if } x_1+\dots+x_n=s \\ 0 & \text{O.W.} \end{cases}$$

$$= \begin{cases} \frac{\prod_{i=1}^n [\phi(1-\theta)^{x_i-1}]}{\binom{s-1}{n-1} \phi^n (1-\theta)^{s-n}} & \text{if } x_1 + \dots + x_n = s \\ 0 & \text{o.w} \end{cases}$$

$$= \begin{cases} \frac{\phi^n (1-\theta)^{s-n}}{\binom{s-1}{n-1} \phi^n (1-\theta)^{s-n}} & \text{if } x_1 + \dots + x_n = s \\ 0 & \text{o.w} \end{cases}$$

$$= \begin{cases} \frac{1}{\binom{s-1}{n-1}} & \text{if } x_1 + \dots + x_n = s \\ 0 & \text{o.w} \end{cases}$$

free of ϕ ! $\Rightarrow S$ is sufficient for ϕ .

Sufficiency and Point Estimation

Factorization Theorem

Remarks on Sufficiency:

Recall in the definition: $p = \# \text{ of parameters}$, $k = \# \text{ of statistics}$

- $k = p$: e.g. last example Geometric(θ), $p=1=k$
- $k > p$: e.g. X_1, \dots, X_n iid UNIF(0, $\theta+1$) $\Rightarrow \underline{S} = (\min X_i, \max X_i)$ sufficient for θ
- $k < p$: e.g. $n=1, X_1 \sim N(\mu, \sigma^2)$ $\begin{matrix} p=1 \\ p=2 \end{matrix}$ but X_1 is sufficient $\begin{matrix} k=1 \end{matrix}$

Factorization Theorem: Let X_1, \dots, X_n be r.v.'s with joint pdf/pmf $f(x|\theta)$, $\theta \in \Theta \subset \mathbb{R}^p$ and let $\underline{S} = (S_1, \dots, S_k)$ be a vector of estimators. Then, \underline{S} is sufficient for θ if and only if there exist functions $g(\underline{S}, \theta)$ and $h(x)$ such that $h(x)$ does NOT depend on θ and

$$\begin{array}{c} \text{data pdf/pmf} \\ \xrightarrow{\quad X_1, \dots, X_n \quad} f(x|\theta) = g(\underline{S}, \theta)h(x) \quad \text{for all } x \text{ and all } \theta \\ \text{~~~~~} \underline{S} \text{ and } \theta \text{ are "linked" inside } f(x|\theta) \end{array}$$

Example: Let X_1, \dots, X_n be iid Negative-Binomial(r, θ), $0 < \theta < 1$ (known integer $r \geq 1$). Show that $S = X_1 + \dots + X_n$ is sufficient for θ . (last time Geo(θ) \sim Neg-Binom($r=1, \theta$))

Solution: joint pmf of X_1, \dots, X_n is

$$\begin{aligned} f(x|\theta) &= \prod_{i=1}^n f(X_i|\theta) = \prod_{i=1}^n \left[\binom{X_i-1}{r-1} \theta^r (1-\theta)^{X_i-r} I_{\{X_i \in A_r\}} \right] \\ &= \theta^{nr} (1-\theta)^{\sum x_i - nr} \underbrace{\prod_{i=1}^n \left[\binom{X_i-1}{r-1} I_{\{X_i \in A_r\}} \right]}_{\text{Indicator}} \\ &\quad \text{where } A_r = \{r, r+1, r+2, \dots\} \\ &\quad \underbrace{g\left(\sum_{i=1}^n x_i, \theta\right)}_{\text{if } x} \quad \underbrace{h(x)}_{A_r \subset \mathbb{N}} \end{aligned}$$

where $g(S, \theta) = \theta^{nr} (1-\theta)^{S-nr}$

Hence by factorization theorem, $S = \sum_{i=1}^n X_i$ is sufficient

Sufficiency and Point Estimation

Factorization Theorem, cont'd

$$\underline{z} = (1, \dots, 1)'$$

Example: Suppose $(X_1, \dots, X_n) \sim MVN(\mu \cdot \underline{1}, \sigma^2 \cdot A)$ where $\mu \in \mathbb{R}$, $\sigma^2 > 0$ and A is a known $n \times n$ positive definite matrix. Find a sufficient statistic for (μ, σ^2) .

Solution: joint pdf of (X_1, \dots, X_n) is

$$f(\underline{x} | \mu, \sigma^2) = \frac{1}{(\sigma^2 2\pi)^{\frac{n}{2}}} \frac{1}{[\det(A)]^{\frac{1}{2}}} \exp \left[-\frac{1}{2\sigma^2} (\underline{x} - \mu \cdot \underline{1})' A^{-1} (\underline{x} - \mu \cdot \underline{1}) \right]$$

$$= \frac{1}{\sigma^n} \exp \left[-\frac{1}{2\sigma^2} [\underline{x}' A^{-1} \underline{x} + 2\mu \underline{x}' A^{-1} \underline{1} + \mu^2 \underline{1}' A^{-1} \underline{1}] \right] \underbrace{\frac{1}{(2\pi)^{\frac{n}{2}} [\det(A)]^{\frac{1}{2}}} \mathcal{I}(\underline{x} \in \mathbb{R}^n)}_{h(\underline{x})}$$

$$g(\underline{x}' A^{-1} \underline{x}, \underline{x}' A^{-1} \underline{1}, \mu, \sigma^2)$$

Hence, by Factorization Theorem,

$\underline{S} = (\underline{x}' A^{-1} \underline{x}, \underline{x}' A^{-1} \underline{1})$ are sufficient for (μ, σ^2)

Remarks:

1. The choice of $g(\underline{S}, \theta)$ and $h(\underline{x})$ is not unique.
2. Any 1-to-1 function of a sufficient statistic is also sufficient.

Example: In last example, suppose $A = I_{n \times n}$.