

# $P(A|B) = P(A)$ Common univariate distributions

Discrete distributions: Negative Binomial

$X \sim \text{Neg-Binom}(r, p)$ ,  $0 < p < 1$

- pmf given by

$$f_X(x) = f_X(x|r, p) = \binom{x-1}{r-1} p^r (1-p)^{x-r}, \quad x = r, r+1, \dots,$$

$r=1$

- Motivation: distribution for the number of independent Bernoulli( $p$ ) trials needed to obtain  $r$  successes

- $Y = X - r$  (number of failures prior to the  $r$ th success) also common

$$f_Y(y|r, p) = \binom{y+r-1}{y} p^r (1-p)^y, \quad y = 0, 1, \dots,$$

- Showing that these probabilities sum to 1 is not easy (next slide)
- Be careful: both r.v.s  $X$  and  $Y$  (different) are called “negative binomial”
- Mean:  $EY = \frac{r(1-p)}{p}$  and hence  $EX = EY + r = \frac{r}{p}$
- Variance:  $\text{Var}(Y) = \frac{r(1-p)}{p^2} = \text{Var}(X)$

$$\bullet M_Y(t) = Ee^{tY} = \left[ \frac{p}{1 - (1-p)e^t} \right]^r, \quad t < -\log(1-p),$$

$$M_X(t) = Ee^{t(Y+r)} = Ee^{rt} e^{tY} = e^{rt} M_Y(t)$$

## Common univariate distributions

Discrete distributions: Negative Binomial (cont'd)

To show probabilities sum to 1:

1. Newton's negative binomial formula : if  $\alpha < 0$  and  $|x| < 1$ ,

$$(1+x)^\alpha = \sum_{k=0}^{\infty} g^{(k)}(0) \frac{(x-0)^k}{k!} = \sum_{k=0}^{\infty} \binom{\alpha}{k}^* x^k, \quad \binom{\alpha}{k}^* \equiv \frac{\alpha(\alpha-1)\cdots(\alpha-k+1)}{k!}$$

Taylor expanding  $g(x) = (1+x)^\alpha$  around 0:  $g^{(0)}(0) = g(0) = 1$ ,  $g^{(1)}(0) = g'(0) = \alpha$

2. for integers  $r \geq 1$  and  $k \geq 0$ , note that

$$\binom{-r}{k}^* (-1)^k = (-1)^k \frac{(-r)(-r-1)\cdots(-r-k+1)}{k!} = \frac{(r)(r+1)\cdots(r+k-1)}{k!} = \binom{r+k-1}{k}$$

$$\begin{aligned} \sum_{y=0}^{\infty} f_Y(y) &= \sum_{y=0}^{\infty} \binom{y+r-1}{y} p^r (1-p)^y = \sum_{y=0}^{\infty} \binom{-r}{y}^* (-1)^y p^r (1-p)^y \\ &= p^r \sum_{y=0}^{\infty} \binom{-r}{y}^* (p-1)^y \\ &= p^r [1 + (p-1)]^{-r} = 1 \end{aligned}$$

Show  $M_Y(t) = Ee^{tY} = \left[ \frac{p}{1 - (1-p)e^t} \right]^r$  for  $t < -\log(1-p)$

# Common univariate distributions

Discrete distributions: Geometric

$$X \sim \text{Geom}(p), 0 < p < 1$$

- special case of Negative Binomial ( $\underline{r=1}, p$ )
- Motivation: distribution for the number of independent Bernoulli( $p$ ) trials needed to obtain 1st success
- pmf given by

$$\mathbb{P}(X=x) = f_X(x) = f_X(x|p) = p(1-p)^{x-1}, \quad x = 1, 2, 3, \dots,$$

$$\bullet \text{ Mean: } EX = \frac{1}{p}$$

$$\bullet \text{ Variance: } \text{Var}(X) = \frac{1-p}{p^2}$$

$$\bullet M_X(t) = Ee^{tX} = \frac{pe^t}{1 - (1-p)e^t} \text{ for } t < -\log(1-p)$$

$$\text{Recall: } X \sim \text{Neg-Bin}(r, p)$$

$$M_X(t) = \frac{e^{rt}}{(1 - (1-p)e^t)^r}$$

$$\xrightarrow{r=1} M_X(t) = \frac{e^t p}{1 - (1-p)e^t}$$

## Common univariate distributions

Discrete distributions: Geometric (cont'd)

1. We've seen the cdf of a  $\text{Geom}(p)$  random variable  $X$ ; it's relatively simple.

$$X \sim \text{Geo}(p) \Rightarrow \underline{\underline{F_X(x)}} \stackrel{\text{def}}{=} P(X \leq x)$$

$$1 - F_X(x) = \underline{\underline{P(X > x)}} = (1-p)^x \quad (*)$$

$$F_X(x) = 1 - (1-p)^x \text{ for } x=1, 2, \dots$$

2. The Geometric distribution has the famous "memoryless" property: for any integer  $x_0 \geq 0$ .

$$\underbrace{P(X = x_0 + x | X > x_0)}_{\text{Interpretation: The conditional distribution of the remaining waiting number of trials until a 1st success, given that I've already waited } x_0 \text{ trials, is the same as the original distribution of the number of trials until 1st success.}} = \underbrace{P(X = x)}$$

Given that each trial is an independent Bernoulli trial, this does make sense: whether I start counting trials at the beginning or I start counting trials after  $x_0$  trials without success, the distribution of the remaining number of trials needed until a 1st success should be the same.

$$\begin{aligned} P(X = x+x_0 | X > x_0) &= \frac{P(X = x+x_0 \cap X > x_0)}{P(X > x_0)} = \frac{P(X = x+x_0)}{P(X > x_0)} \\ &= \frac{P(1-p)^{x+x_0-1}}{(1-p)^{x_0}} = p(1-p)^{x-1} \\ &= \underline{\underline{P(X=x)}} \end{aligned}$$

3. Example: Testing newly manufactured widgets with  $p = 0.01$  probability that a given widget fails a functionality test, what's the probability of running at least 50 units without a test failure?

$\rightarrow X = \# \text{ of trials (tests) until the 1st failure}$

$$X \sim \text{Geom}(0.01)$$

$$P(X > 50) = (1-p)^x = (1-0.01)^{50} \approx 0.605$$

$$E(X) = \frac{1}{p} = \underline{\underline{100}}$$

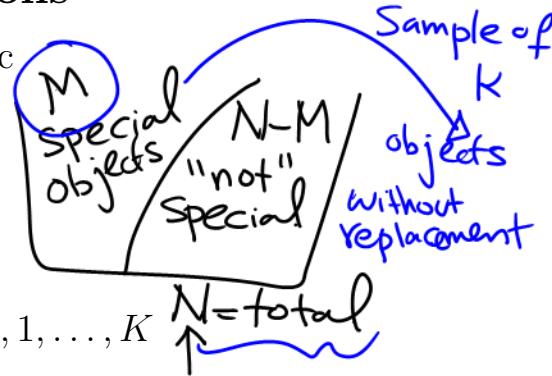
## Common univariate distributions

Discrete distributions: Hypergeometric

$X \sim \text{Hypergeometric}(N, M, K)$  (integers  $N, M, K$ )

- pmf given by

$$P(X=x) = f_X(x) = \binom{M}{x} \binom{N-M}{K-x} / \binom{N}{K}, \quad x = 0, 1, \dots, K$$



- Motivation: Choose  $K$  objects without replacement from a total population of size  $N$  which contains  $M$  "special" objects.  $X$  is the number of "special" objects among the  $K$  chosen.

$$\rightarrow P(X=x) = \frac{M!}{x!(M-x)!} \frac{(N-M)!}{(K-x)!(N-M-k+x)!} \frac{k!(N-k)!}{N!}$$

pmf of  $X \sim \text{Hyper}(N, M, K)$   
for  $x=0, 1, 2, \dots, K$  total

- Must have  $0 \leq x \leq K$ ,  $x \leq M$ , and  $K - x \leq N - M$  in  $f_X(x)$

Typically,  $N > 2M$  and  $M > K$  so only the condition  $0 \leq x \leq K$  matters

- Mean:  $EX = KM/N$
- Variance:  $\text{Var}(X) = \frac{KM(N-M)(N-K)}{N^2(N-1)}$

To derive mean:

$$\begin{aligned} EX &= \sum_{x=0}^K x \binom{M}{x} \binom{N-M}{K-x} / \binom{N}{K} = \sum_{x=1}^K x \frac{M!}{x!(M-x)!} \binom{N-M}{K-x} \frac{K!(N-K)!}{N!} \\ &\stackrel{\text{def}}{=} \sum_{x=1}^K \frac{x(M-1)!}{x!(x-1)!} \binom{N-M}{K-x} \frac{K!(K-1)!}{N!(N-1)!} \\ &= \frac{KM}{N} \sum_{x=1}^K \frac{(M-1)!}{(M-x)!(N-x)!} \binom{N-M}{K-x} \frac{(N-1)!}{(N-x)!(N-K)!} \end{aligned}$$

$$Y \sim \text{HyperGeo}(N-1, M-1, K-1) \quad y = x-1 \quad \frac{KM}{N}$$

~~Common univariate distributions~~

$$\sum_{y=0}^{K-1} \frac{(M-1)!}{(M-1-y)!(N-1)!} \binom{N-1-(M-1)}{K-1-(x-1)} = 1$$

Discrete distributions: Hypergeometric and Binomial

### 1. Hypergeometric: sampling without replacement

i.e., choose  $x$  special objects in a size  $K$  sample from a collection where  $M$  objects are “special” &  $N - M$  are not

### 2. Binomial: sampling with replacement

i.e., choose  $x$  special objects in a size  $n$  sample, where each selected item of the sample has probability  $p$  of being a special object

$$\text{Bin}(n=k, p=\frac{M}{N})$$

Suppose  $X \sim \text{Hypergeometric}(N, M, K)$  and let  $M/N = p$  be the proportion of “special” objects and let  $K = n$  be the sample size. Then,

- $EX = K(M/N) = np$

- $\text{Var}(X) = \frac{KM(N-M)(N-K)}{N^2(N-1)} = np(1-p)\frac{N-n}{N-1}$

- the factor  $(N-n)/(N-1)$  in  $\text{Var}(X)$  is the finite sample correction factor

Note:  $N \geq 2nk$   
 Sample  
 $N \geq 2n$

Now fix  $K = n$  and  $p = M/N$  as above, and let  $N \rightarrow \infty$ . Then,

$$\lim_{N \rightarrow \infty} P(X=x) = \lim_{N \rightarrow \infty} \binom{M}{x} \binom{N-M}{K-x} / \binom{N}{K} = \binom{n}{x} p^x (1-p)^{n-x}, \quad x = 0, \dots, n$$

Result: hypergeometric tends to binomial distribution in large population size  $N$