# Statistics 520 - Assignment 3

#### Sam Olson

## Assignment 3

1. (10 pt.) Suppose that a random variable Y has a beta distribution with parameters  $\alpha$  and  $\beta$ . A standard form for the probability density function of Y is, for  $\alpha > 0$  and  $\beta > 0$ ,

$$f(y \mid \alpha, \beta) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} y^{\alpha - 1} (1 - y)^{\beta - 1}, \quad 0 < y < 1.$$

Put this density in canonical exponential family form.

Using properties of exponential families, find  $E\{\log(Y)\}$  and  $E\{\log(1-Y)\}$  expressed in terms of the original  $\alpha$  and  $\beta$  parameters.

*Note:* Use  $\Gamma'(x)$  to denote the derivative of the gamma function,

$$\frac{d}{dx}\Gamma(x)$$
.

#### Answer

The canonical exponential family form of the density is:

$$f(y \mid \alpha, \beta) = \exp\{(\alpha - 1)\log(y) + (\beta - 1)\log(1 - y) + \log\left(\frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)}\right)\} \mathbb{I}[y \in (0, 1)]$$

Where:

$$\theta_1 = \alpha - 1, \quad \theta_2 = \beta - 1.$$

$$T = (T_1, T_2)$$
 for  $T_1(y) = \log(y)$ ,  $T_2(y) = \log(1 - y)$ .

$$B(\theta) = -\log\left(\frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)}\right) = \log\Gamma(\alpha) + \log\Gamma(\beta) - \log\Gamma(\alpha+\beta).$$

 $c(y) = \mathbb{I}[y \in (0,1)],$  where  $\mathbb{I}$  denotes the indicator function

Note: Though  $B(\theta)$  is as given above, a simplified version which makes taking partial derivatives easier is the equivalent form:

Using properties of exponential families, and noting that the natural parameters  $(\theta_1, \theta_2)$  are linearly related to the parameters  $(\alpha, \beta)$ :

$$E\{\log(Y)\} = E\{T_1(Y)\} = \frac{\partial}{\partial \theta_1} B(\theta) = \frac{\partial}{\partial \alpha} (\log \Gamma(\alpha) + \log \Gamma(\beta) - \log \Gamma(\alpha + \beta))$$
$$= \frac{\Gamma'(\alpha)}{\Gamma(\alpha)} - \frac{\Gamma'(\alpha + \beta)}{\Gamma(\alpha + \beta)}$$

And

$$E\{\log(1-Y)\} = E\{T_2(Y)\} = \frac{\partial}{\partial \theta_2} B(\theta) = \frac{\partial}{\partial \beta} \left(\log \Gamma(\alpha) + \log \Gamma(\beta) - \log \Gamma(\alpha + \beta)\right)$$
$$= \frac{\Gamma'(\beta)}{\Gamma(\beta)} - \frac{\Gamma'(\alpha + \beta)}{\Gamma(\alpha + \beta)}$$

2. (5 pt.) Suppose that a random variable Y has a Poisson distribution with parameter  $\lambda$ . A standard form for the probability mass function of Y is, for  $\lambda > 0$ ,

$$f(y \mid \lambda) = \frac{1}{y!} \lambda^y \exp(-\lambda), \quad y = 0, 1, 2, \dots$$

Put this probability mass function in canonical exponential family form.

Using properties of exponential families, verify that  $E(Y) = \lambda$ .

### Answer

The canonical exponential family form of the density is:

$$f(y \mid \lambda) = \exp\{y \log(\lambda) - \lambda - \log(y!)\}\$$

 $\theta_1 = \log(\lambda).$ 

 $T_1(y) = y$ .

 $B(\theta) = \exp(\theta_1).$ 

 $c(y) = -\log(y!)$ 

Using properties of exponential families, and noting the natural parameter  $\theta_1$  is non-linearly related to the parameter  $\lambda$ :

$$E\{T_1(Y)\} = \frac{\partial}{\partial \theta_1} B(\theta) = \frac{\partial}{\partial \theta_1} e^{\theta_1} = e^{\log(\lambda)} = \lambda$$

3. Suppose that a random variable Y has a gamma distribution with parameters  $\alpha$  and  $\beta$ . A standard form for the probability density function of Y is, for  $\alpha > 0$  and  $\beta > 0$ ,

$$f(y \mid \alpha, \beta) = \frac{\beta^{\alpha}}{\Gamma(\alpha)} y^{\alpha - 1} \exp(-\beta y), \quad y > 0.$$

*Note:* You may have seen a gamma density written with a parameter that is equal to  $1/\beta$  in the above expression. Use the parameterization given above to answer this question (I think it will be easier).

- (a) (5 pts.) Write the gamma density in the form of a two-parameter exponential family. Using properties of exponential families, derive the expected values of Y and  $\log(Y)$ .
- (b) (5 pts.) Write the gamma density in the form of an exponential dispersion family with parameters  $\theta$  and  $\phi$ . Derive the expected value of Y.

#### Answer

(a)

The gamma density in the form of a (canonical) two-parameter exponential family is of the form:

$$f(y \mid \alpha, \beta) = \exp\{(\alpha - 1)\log(y) - \beta y + \alpha\log(\beta) - \log\Gamma(\alpha)\}\mathbb{I}[y > 0]$$

Where:

$$\theta_1 = \alpha - 1$$
  $\theta_2 = -\beta$ ,

$$T = (T_1(y), T_2(y)), T_1(y) = \log(y), T_2(y) = y,$$

$$B(\theta) = \log \Gamma(\alpha) - \alpha \log(\beta)$$

And

 $c(y) = \mathbb{I}[y > 0]$ , where  $\mathbb{I}$  denotes the indicator function

Using properties of exponential families, and noting the natural parameters  $(\theta_1, \theta_2)$  are linear functions of the parameters  $(\alpha, \beta)$ , then:

$$E(\log(Y)) = E\{T_1(Y)\} = \frac{\partial}{\partial \theta_1} B(\theta) = \frac{\partial}{\partial \alpha} \left(\log \Gamma(\alpha) - \alpha \log(\beta)\right) = \frac{\Gamma'(\alpha)}{\Gamma(\alpha)} - \log(\beta)$$

(Another Digamma function, in the flesh!)

Also:

$$E(Y) = E\{T_2(Y)\} = \frac{\partial}{\partial \theta_2} B(\theta) = \frac{\partial}{\partial (-\beta)} \left( \log(\Gamma(\alpha) - \alpha \log(\beta)) \right) = \frac{\alpha}{\beta}$$

(b)

Now, taking the canonical form, we may then write the exponential dispersion family form as:

$$\begin{split} f(y \mid \alpha, \beta) &= \exp \left( (\alpha - 1) \log y - \beta y + \alpha \log \beta - \log \Gamma(\alpha) \right) \\ &= \exp \left\{ \alpha \left( \log(\frac{\beta}{\alpha}) - y \frac{\beta}{\alpha} \right) + \left( (\alpha - 1) \log y + \alpha \log \alpha - \log \Gamma(\alpha) \right) \right\} \\ &= \exp \left\{ \phi \left( y \theta - b(\theta) \right) + c(y, \phi) \right\}, \qquad y > 0, \end{split}$$

where

$$\phi = \alpha, \quad \theta = -\frac{\beta}{\alpha}$$

And

$$b(\theta) = -\log(-\theta)$$
, and  $c(y, \phi) = (\alpha - 1)\log y + \alpha\log \alpha - \log \Gamma(\alpha)$ 

Using the properties of an exponential dispersion family, we may calculate expectation via:

$$E(Y) = \frac{d}{d\theta}b(\theta) = \frac{d}{d\theta}\left(-\log(-\theta)\right) = -\frac{1}{\theta} = \frac{\alpha}{\beta}$$