

Deconvolution in Nonparametric Statistics

K. De Brabanter¹

¹Department of Statistics

Department of Industrial and Manufacturing Systems Engineering
Iowa State University

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Outline

1 Introduction

- What is deconvolution?
- Where do such problems occur?
- Notations & assumptions

2 Density deconvolution

- Naive estimator
- Kernel density deconvolution estimator

3 Nonparametric regression with errors-in-variables

- Errors-in-variables problem formulation
- Kernel regression with errors-in-variables

What is deconvolution?

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Deconvolution

Goal: Estimate a function f while access is restricted to the quantity

$$z = f * G = \int f(x - y) dG(y),$$

i.e., the convolution of f and some probability distribution G .

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• SOLUTION

- 1 Estimate $z \Rightarrow \hat{z}$
- 2 Apply deconvolution procedure to $\hat{z} \Rightarrow \hat{f}$

Where do such problems occur?

- Density estimation based on contaminated data (Delaigle & Gijbels, 2006)
- Nonparametric regression with errors-in-variables (Delaigle & Meister, 2011)
- Image and signal deblurring (Qiu, 2005)
- Econometrics (Meister, 2009)

Notations & assumptions

Deconvolution

$$z = f * G = \int f(x - y) dG(y)$$

- For now, assume G is fully known
- G has a density function g in Lebesgue sense:

$$z = f * G = \int f(x - y)g(y) dy$$

- Main tool: Fourier transform (FT)

$$\mathcal{G}(t) = \int \exp(itx) dG(x), \quad t \in \mathbb{R}$$

- Result via FT: $\mathcal{Z} = \mathcal{F} \cdot \mathcal{G}$

Simple algorithm

$$\mathcal{Z} = \mathcal{F} \cdot \mathcal{G}$$

- ➊ Estimate \mathcal{Z} based on empirical information, denoted by $\hat{\mathcal{Z}}$.
- ➋ Calculate $\hat{\mathcal{Z}}(t)$ and divide it by $\mathcal{G}(t)$, leading to $\hat{\mathcal{F}}(t)$.
- ➌ Regularize $\hat{\mathcal{F}}(t)$ so that the inverse Fourier transform \hat{f} exists.

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Simple algorithm, theoretical effort must not be underestimated

Assumptions & general estimation procedure

- Direct data not always available in practice because of measurement error
- Observe contaminated data Y_1, \dots, Y_n instead of the true data X_1, \dots, X_n
- Assume the following model

$$Y_j = X_j + \varepsilon_j, \quad j \in \{1, \dots, n\}$$

- Assumptions
 - X_j and ε_j are real valued and independent
 - $\mathbf{E}(\varepsilon_j | X_j) = 0$
 - $\mathbf{Var}(\varepsilon_j | X_j) < \infty$
- **Goal:** estimate density f of the **UNOBSERVED** random variable X

Estimation procedure

- Basic statistics: density of the sum of two independent random variables \Rightarrow deconvolution of both addends
- **STEP 1:** estimate density z given observations Y_j

$$\mathcal{Z}(t) = \int \exp(itx) z(x) dx = \mathbf{E} \exp(itY) = \Psi_Y(t)$$

- **STEP 2:** define estimator of $\Psi_Y(t)$

$$\hat{\Psi}_Y(t) = \frac{1}{n} \sum_{j=1}^n \exp(itY_j)$$

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Estimation procedure (cont'd)

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- **STEP 3:** define estimator for $\mathcal{F}(t)$

$$\hat{\Psi}_X(t) = \frac{1}{n} \frac{\sum_{j=1}^n \exp(itY_j)}{\mathcal{G}(t)}$$

- This estimator is unbiased and consistent. A naive estimator for f is

$$\hat{f}_{\text{naive}}(x) = \frac{1}{2\pi} \int \exp(-itx) \hat{\Psi}_X(t) dt$$

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This estimator has some drawbacks \Rightarrow regularization needed!!

Kernel density deconvolution estimator

- Use the following estimator for the density z

$$\hat{z}(x) = \frac{1}{nh} \sum_{j=1}^n K\left(\frac{x - Y_j}{h}\right)$$

- $K : \mathbb{R} \rightarrow \mathbb{R}^+$
- bandwidth parameter $h > 0$

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Existence condition

If $K \in L_1(\mathbb{R}) \cap L_2(\mathbb{R})$, in Lebesgue sense, the estimator \hat{z} also lies $L_1(\mathbb{R}) \cap L_2(\mathbb{R})$ almost surely so that its Fourier transform exist.

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$$\mathcal{Z}(t) = \frac{1}{nh} \int \sum_{j=1}^n \exp(itx) K\left(\frac{x - Y_j}{h}\right) dx = \dots = \hat{\Psi}_Y(t) \cdot \mathcal{K}(th)$$

Kernel density deconvolution estimator (cont'd)

- $\mathcal{Z}(t) = \hat{\Psi}_Y(t) \cdot \mathcal{K}(th)$
- $\mathcal{Z}(t) = \mathcal{F}(t) \cdot \mathcal{G}(t)$



$$\hat{\Psi}_X(t) = \frac{\hat{\Psi}_Y(t)\mathcal{K}(th)}{\mathcal{G}(t)}$$



$$\boxed{\hat{f}(x) = \frac{1}{2\pi} \int \exp(-itx) \mathcal{K}(th) \frac{\frac{1}{n} \sum_{j=1}^n \exp(itY_j)}{\mathcal{G}(t)} dt}$$

Some examples

- Consider two densities

$$X \sim 0.5N(-3, 1^2) + 0.5N(2, 1^2) \text{ and } X \sim 0.5N(0, 1^2) + 0.5N(3, (1/2)^2)$$

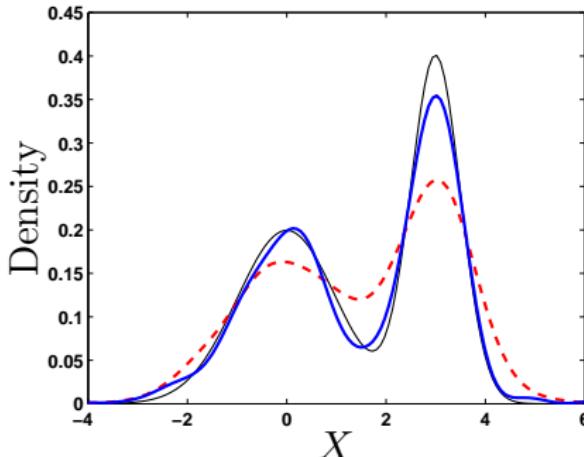
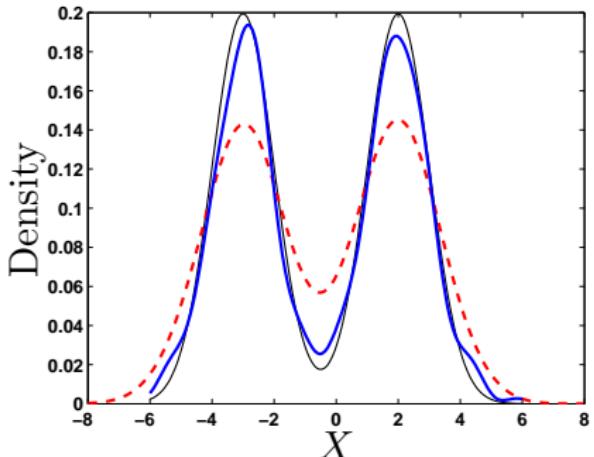
- Error density g is $\mathcal{L}(0, 0.5) \rightarrow \mathcal{G}(t) = 4/(4 + t^2)$
- Kernel: $K(u) = (1/2\pi)[\sin(\frac{1}{2}u)/\frac{1}{2}u]^2 \Rightarrow \mathcal{K}(t) = \max(1 - |t|, 0)$

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Errors-in-variables problem formulation

① Classical view

- Covariates X can only be observed with some additive independent noise
- We observe the i.i.d. data $(W_1, Y_1), \dots, (W_n, Y_n)$
- $W_j = X_j + \delta_j \quad \text{and} \quad Y_j = m(X_j) + \varepsilon_j \quad \text{for } j = 1, \dots, n$

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② Berkson regression (Berkson, 1950)

- Main difference: covariate is affected by additive noise **after** it was measured
- We observe the i.i.d. data $(X_1, Y_1), \dots, (X_n, Y_n)$
- $Y_j = m(X_j + \delta_j) + \varepsilon_j, \quad j = 1, \dots, n$

Kernel regression with errors-in-variables

- Adapt Nadaraya-Watson regression estimator for errors-in-variables

$$\hat{m}(x) = \sum_{j=1}^n \frac{K\left(\frac{x-X_j}{h}\right) Y_j}{\sum_{j=1}^n K\left(\frac{x-X_j}{h}\right)}$$

- Using similar techniques as for density estimation (Fan & Truong, 1993)



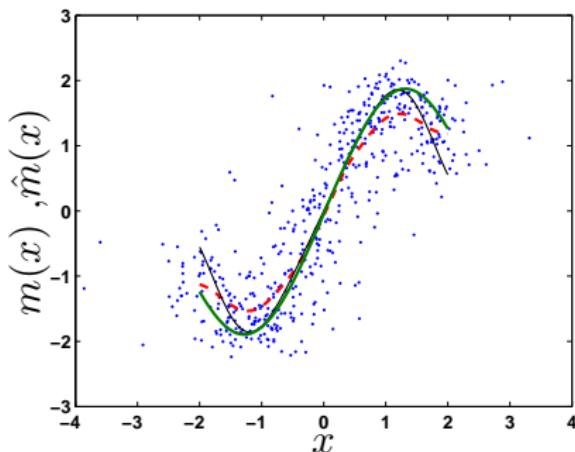
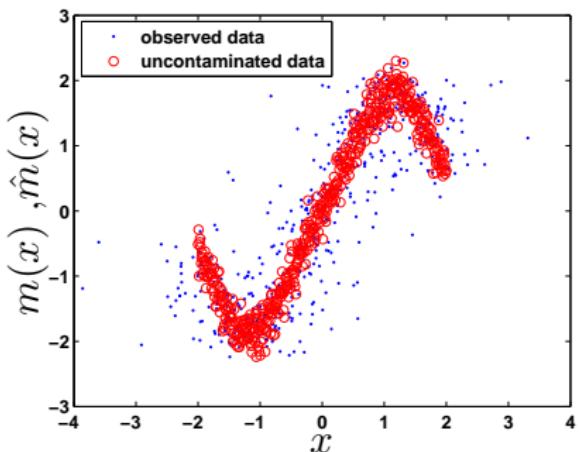
$$\hat{m}(x) = \frac{\frac{1}{nh} \sum_{j=1}^n \frac{1}{2\pi} \left[\int \exp\left(-i\left(\frac{x-W_j}{h}\right) u\right) \frac{\mathcal{K}(u)}{\mathcal{G}(\frac{u}{h})} du \right] Y_j}{\frac{1}{2n\pi} \sum_{j=1}^n \int \exp(-itx) \mathcal{K}(th) \frac{\exp(itW_j)}{\mathcal{G}(t)} dt}$$

Some examples

- $W_j = X_j + \delta_j$ and $Y_j = m(X_j) + \varepsilon_j$ for $j = 1, \dots, n$
- $m(x) = 2x \exp(-10x^4/81)$ with $x \in [-2, 2]$
- $\varepsilon \sim N(0, 0.2^2)$ and $\delta \sim \mathcal{L}(0, 0.2)$
- $\mathbf{E}[\varepsilon|X] = 0$ and the δ_j are independent of the (X_j, Y_j)

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Unknown noise distribution

- **Additional data:** error density is estimated from separate independent experiment (e.g. Golden standard - normal measurement)

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Unknown noise distribution

- **Additional data:** error density is estimated from separate independent experiment (e.g. Golden standard - normal measurement)
- **Replicated measurements:** independent measurements (affected by error) are done several times
- These are topics of research: see e.g. Efromovich (1997), Delaigle *et al.* (2009), Meister (2009), Delaigle & Meister (2011)



Thank you for your attention

Any questions...?

References

-  A. Delaigle & I. Gijbels. Data-driven boundary estimation in deconvolution problems. *Computational Statistics & Data Analysis*, 50(8):19651994, 2006.
-  A. Delaigle & A. Meister. Rate-optimal nonparametric estimation in classical and Berkson errors-in-variables. *Journal of Statistical Planning and Inference*, 141(1):102–114, 2011.
-  P. Qiu. *Image Processing and Jump Regression Analysis*. John Wiley & Sons, New York, 2005.
-  A. Meister. *Deconvolution Problems in Nonparametric Statistics*. Springer-Verlag, Berlin Heidelberg, 2009.
-  J. Berkson. Are there two regression problems? *Journal of the American Statistical Association*, 45:164180, 1950.
-  S. Efromovich. Density estimation for the case of supersmooth measurement error. *Journal of the American Statistical Association*, 92(438):526535, 1997.
-  A. Delaigle, J. Fan, & R.J. Carroll. A design-adaptive local polynomial estimator for the errors-in-variables problem. *Journal of the American Statistical Association*, 104(485):348359, 2009.
-  J. Fan & Y.K. Truong. Nonparametric regression with errors in variables. *The Annals of Statistics*, 21(4):19001925, 1993.