

ROC (7 marks)

Spring : 2020

6) a) For threshold 0.75 :

if score ≥ 0.75 , prediction = 'P'
 if score < 0.75 , prediction = 'N'

		Prediction	
		1	0
Actual	1	1	4
	0	0	4

$$\therefore TPR = \frac{TP}{TP+FN} = \frac{2}{2+4} = 0.33$$

$$\therefore FPR = \frac{FP}{FP+TN} = \frac{0}{0+4} = 0$$

$$\therefore TPR, FPR = (0.33, 0)$$

actual	score	prediction
P	0.9	P
P	0.8	P
n	0.7	n
P	0.6	n
P	0.55	n
P	0.54	n
n	0.53	n
n	0.52	n
P	0.51	n
n	0.505	n

For threshold 0.57 :

if score ≥ 0.57 , prediction = 'P'
 if score < 0.57 , prediction = 'n'

		Prediction	
		1	0
Actual	1	3	3
	0	1	3

$$\therefore TPR = \frac{3}{3+3} = 0.5$$

$$\therefore FPR = \frac{1}{1+3} = 0.25$$

$$\therefore TPR, FPR = (0.5, 0.25)$$

actual	score	prediction
P	0.9	P
P	0.8	P
n	0.7	P
P	0.6	P
P	0.55	n
P	0.54	n
n	0.53	n
n	0.52	n
P	0.51	n
n	0.505	n

For threshold 0.53 :

		Prediction	
		1	0
Actual	1	5	1
	0	2	2

P	P	$\therefore TPR = \frac{5}{5+1} = 0.83$
P	P	
n	P	
P	P	
P	P	
P	P	
P	P	
n	P	
n	n	
P	n	
n	n	

$$\therefore TPR, FPR = (0.83, 0.5)$$

For threshold 0.51 :

	1	0
1	6	0
0	4	1

P	P
P	P
n	P
P	P
P	P
P	P
n	P
n	P
P	P
n	n

$$\therefore TPR = \frac{6}{6+0} = 1$$

$$\therefore FPR = \frac{4}{4+1} = 0.8$$

$$\therefore TPR, FPR = (1, 0.8)$$

For threshold 0.50 :

	1	0
1	6	0
0	4	0

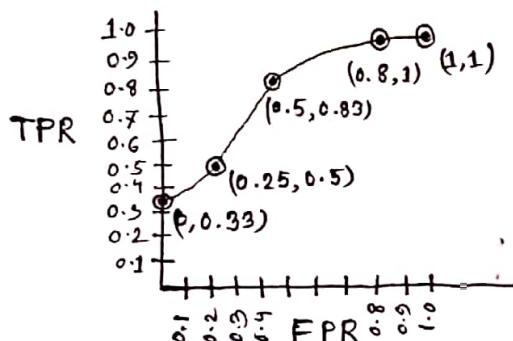
P	P
P	P
n	P
P	P
P	P
P	P
n	P
n	P
P	P
n	P

$$\therefore TPR = \frac{6}{6+0} = 1$$

$$\therefore FPR = \frac{4}{4+0} = 1$$

$$\therefore TPR, FPR = (1, 1)$$

Receiver Operating Characteristic (ROC)/ Receiver Operating Curve :



Fall: 2018 :

$$4) c) \text{ Accuracy} = \frac{TP+TN}{TP+FP+FN+TN} = \frac{100+300}{100+60+40+300} = 0.8$$

$$\text{Precision} = \frac{TP}{TP+FP} = \frac{100}{100+60} = 0.625$$

$$\text{Recall} = \frac{TP}{TP+FN} = \frac{100}{100+40} \approx 0.7143$$

$$\text{F measure} = \frac{2}{\frac{1}{\text{precision}} + \frac{1}{\text{recall}}} = \frac{2}{\frac{1}{0.625} + \frac{1}{0.7143}} \approx 0.67$$

$$TP = 100 ; FN = 40 ; FP = 60 ; TN = 300$$

(5-7 marks) PCA + LDA [can be a part of mandatory question 1]

Spring - 2020:

1) b) Step: 01: Get some data :

Consider, $X = \text{Dhaka}$ and $Y = \text{Rajshahi}$

Total $X = 363$ $[\# = 51]$

\therefore Average X , $\bar{X} = \frac{363}{9} \approx 40.33333333$

Total $Y = 276$

\therefore Average Y , $\bar{Y} = \frac{276}{9} = 30.6666667$

Step: 02: Subtract the mean :

$X_i - \bar{X}$	$Y_i - \bar{Y}$
10.6666667	9.3333333
19.6666667	-3.6666667
9.6666667	28.3933333
-10.3333333	-15.6666667
-38.3333333	-26.6666667
19.6666667	-9.6666667
-0.3333333	-2.6666667
21.6666667	40.3333333
-32.3333333	-19.6666667

Step: 03: calculate the co-variance matrix :

$$\text{cov}(x, x) = \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n-1} = \frac{4072}{8} = 509$$

$$\text{cov}(x, y) = \text{cov}(y, x) = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{n-1} = \frac{2806}{8} = 350.75$$

$$\text{cov}(y, y) = \frac{\sum_{i=1}^n (y_i - \bar{y})^2}{n-1} = \frac{3974}{8} = 496.75$$

$$\therefore \text{cov} = \begin{bmatrix} \text{cov}(x, x) & \text{cov}(x, y) \\ \text{cov}(y, x) & \text{cov}(y, y) \end{bmatrix} = \begin{bmatrix} 509 & 350.75 \\ 350.75 & 496.75 \end{bmatrix}$$

Step:04 : calculate eigenvector and eigenvalues:

$$(A - \lambda I) X = 0$$

$$\Rightarrow A - \lambda I = 0$$

$$\Rightarrow \begin{bmatrix} 509 & 350.75 \\ 350.75 & 496.75 \end{bmatrix} - \lambda \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} = 0$$

$$\Rightarrow \begin{bmatrix} 509 & 350.75 \\ 350.75 & 496.75 \end{bmatrix} - \begin{bmatrix} \lambda & 0 \\ 0 & \lambda \end{bmatrix} = 0$$

$$\Rightarrow \begin{bmatrix} 509 - \lambda & 350.75 \\ 350.75 & 496.75 - \lambda \end{bmatrix} = 0$$

$$\Rightarrow (509 - \lambda)(496.75 - \lambda) - 350.75^2 = 0$$

$$\Rightarrow 252845.75 - 509\lambda - 496.75\lambda + \lambda^2 - 123025.5625 = 0$$

$$\Rightarrow \lambda^2 - 1005.75\lambda + 120820.1875 = 0$$

$$\therefore \lambda = \frac{-(-1005.75) \pm \sqrt{(-1005.75)^2 - 4 \times 1 \times 120820.1875}}{2 \times 1}$$

$$\Rightarrow \lambda = \frac{1005.75 \pm 701.6069502}{2}$$

$$\therefore \lambda_1 = 853.6784751$$

$$\therefore \lambda_2 = 152.0715249$$

$$\therefore \text{Eigenvalues} = \begin{bmatrix} 853.6784751 \\ 152.0715249 \end{bmatrix}$$

Step:05: choosing components and forming a feature vector:

for $\lambda_1 = 853.6784751$:

$$\begin{bmatrix} 509 - 853.6784751 & 350.75 \\ 350.75 & 496.75 - 853.6784751 \end{bmatrix} \begin{bmatrix} x_1 \\ y_1 \end{bmatrix} = 0$$

$$\Rightarrow \begin{bmatrix} -344.6784751 & 350.75 \\ 350.75 & -356.9284751 \end{bmatrix} \begin{bmatrix} x_1 \\ y_1 \end{bmatrix} = 0$$

$$-344.6784751x_1 + 350.75y_1 = 0 \quad \dots \quad \dots \quad \textcircled{1}$$

$$350.75x_1 - 356.9284751y_1 = 0 \quad \dots \quad \dots \quad \textcircled{2}$$

From equation ①,

$$-344.6784751x_1 = -350.75y_1$$

$$\Rightarrow x_1 = \frac{-350.75}{-344.6784751} y_1$$

$$\therefore x_1 = 1.018 y_1$$

From equation ②,

$$350.75x_1 = 356.9284751y_1$$

$$\Rightarrow x_1 = \frac{356.9284751}{350.75} y_1$$

$$\therefore x_1 = 1.018 y_1$$

$$\therefore A = \sqrt{(1.018)^2 + (1)^2} = 1.427$$

$$\therefore e_1 \sim \begin{bmatrix} \frac{1.018}{A} \\ \frac{1}{A} \end{bmatrix} \sim \begin{bmatrix} 0.7134 \\ 0.7008 \end{bmatrix}$$

For $\lambda_2 = 152.0715249$,

$$\begin{bmatrix} 509 - 152.0715249 & 350.75 \\ 350.75 & 496.75 - 152.0715249 \end{bmatrix} \begin{bmatrix} x_1 \\ y_1 \end{bmatrix} = 0$$

$$\Rightarrow \begin{bmatrix} 356.9284751 & 350.75 \\ 350.75 & 344.6784751 \end{bmatrix} \begin{bmatrix} x_1 \\ y_1 \end{bmatrix} = 0$$

$$356.9284751x_1 + 350.75y_1 = 0 \quad \dots \quad \dots \quad \textcircled{3}$$

$$350.75x_1 + 344.6784751y_1 = 0 \quad \dots \quad \dots \quad \textcircled{4}$$

From equation (11),

$$356.9284751x_1 = -350.75y_1$$
$$\Rightarrow x_1 = \frac{-350.75}{356.9284751} y_1$$

$$\therefore x_1 = -0.98269 y_1$$

From equation (14),

$$350.75x_1 = -344.6784751y_1$$
$$\Rightarrow x_1 = \frac{-344.6784751}{350.75} y_1$$

$$\therefore x_1 = -0.98269 y_1$$

$$\therefore A = \sqrt{(-0.98269)^2 + 1^2} = 1.402$$

$$\therefore e_2 \sim \begin{bmatrix} \frac{-0.98269}{1.402} \\ \frac{1}{A} \end{bmatrix} \sim \begin{bmatrix} -0.7009 \\ 0.7133 \end{bmatrix}$$

$$\therefore \text{Eigenvector} = \begin{bmatrix} 0.7134 & -0.7009 \\ 0.7008 & 0.7133 \end{bmatrix}$$

$$\therefore \text{PCA (Principal Component Analysis)} = \begin{bmatrix} 0.7134 \\ 0.7008 \end{bmatrix} \text{ [maximum]}$$

$$\therefore \text{Eigen value} = 853.6784751$$

(ans:)

Fall-2020:

4) b) For class 1 (ω_1):

X_1	X_2	$A = X_1 - \bar{X}_1$	$B = X_2 - \bar{X}_2$	A^2	B^2	AB
5	1	5/3	-4/3	25/9	16/9	-20/9
1	3	-7/3	2/3	49/9	4/9	-14/9
4	3	2/3	2/3	4/9	4/9	4/9

$$\bar{X}_1 = \frac{5+1+4}{3} = \frac{10}{3} = 3.333333333$$

$$\bar{X}_2 = \frac{1+3+3}{3} = \frac{7}{3} = 2.333333333$$

$$\therefore \text{cov}(x, x) = \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n-1} = \frac{\sum A^2}{3-1} = \frac{\frac{26}{3}}{2} = 4.333333333$$

$$\therefore \text{cov}(x, y) = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{n-1} = \frac{\sum AB}{3-1} = \frac{-10/3}{2} = -1.666666667$$

$$\therefore \text{cov}(y, y) = \frac{\sum_{i=1}^n (y_i - \bar{y})^2}{n-1} = \frac{\sum B^2}{3-1} = \frac{\frac{8}{3}}{2} = 1.333333333$$

$$\therefore S_1 = \begin{bmatrix} \text{cov}(x, x) & \text{cov}(x, y) \\ \text{cov}(y, x) & \text{cov}(y, y) \end{bmatrix} = \begin{bmatrix} 4.333333333 & -1.666666667 \\ -1.666666667 & 1.333333333 \end{bmatrix}$$

for class 2 (ω_2):

X_1	X_2	$A = X_1 - \bar{X}_1$	$B = X_2 - \bar{X}_2$	A^2	B^2	AB
8	9	-2/3	7/3	4/9	49/9	-14/9
8	6	-2/3	-2/3	4/9	4/9	4/9
10	5	4/3	-5/3	16/9	25/9	-20/9

$$\bar{X}_1 = \frac{8+8+10}{3} = \frac{26}{3} = 8.666666667$$

$$\bar{X}_2 = \frac{9+6+5}{3} = \frac{20}{3} = 6.666666667$$

$$\therefore \text{cov}(x, x) = \frac{\sum A^2}{2} = \frac{\frac{8}{3}}{2} = 1.333333333$$

$$\therefore \text{cov}(x, y) = \frac{\sum AB}{2} = \frac{-10/3}{2} = -1.666666667$$

$$\therefore \text{cov}(y, y) = \frac{\sum B^2}{2} = \frac{26/3}{2} = 4.333333333$$

$$\therefore S_2 = \begin{bmatrix} 1.333333333 & -1.666666667 \\ -1.666666667 & 4.333333333 \end{bmatrix}$$

$$\therefore S_w = S_1 + S_2 = \begin{bmatrix} 5.666666666 & -3.333333334 \\ -3.333333334 & 5.666666666 \end{bmatrix}$$

$$\begin{aligned} \therefore S_B &= (\mu_1 - \mu_2)(\mu_1 - \mu_2)^T \quad \left[\mu_1 = \begin{bmatrix} \bar{x}_1 \\ \bar{x}_2 \end{bmatrix} \text{ for } \omega_1; \mu_2 = \begin{bmatrix} \bar{x}_1 \\ \bar{x}_2 \end{bmatrix} \text{ for } \omega_2 \right] \\ &= \left(\begin{bmatrix} 3.333333333 \\ 2.333333333 \end{bmatrix} - \begin{bmatrix} 8.666666667 \\ 6.666666667 \end{bmatrix} \right) \left(\begin{bmatrix} 3.333333333 \\ 2.333333333 \end{bmatrix} - \begin{bmatrix} 8.666666667 \\ 6.666666667 \end{bmatrix} \right)^T \\ &= \begin{bmatrix} -5.333333334 \\ -4.333333334 \end{bmatrix} \begin{bmatrix} -5.333333334 & -4.333333334 \\ 28.44444445 & 23.11111112 \\ 23.11111112 & 18.77777778 \end{bmatrix} \end{aligned}$$

Now,

$$S_w^{-1} S_B \omega = \lambda \omega$$

$$\Rightarrow |S_w^{-1} S_B - \lambda I| = 0$$

$$\Rightarrow \begin{bmatrix} 5.666666666 & -3.333333334 \\ -3.333333334 & 5.666666666 \end{bmatrix}^{-1} \begin{bmatrix} 28.44444445 & 23.11111112 \\ 23.11111112 & 18.77777778 \end{bmatrix} - \lambda I = 0$$

$$\Rightarrow \begin{bmatrix} 0.2698412699 & 0.1587301588 \\ 0.1587301588 & 0.2698419699 \end{bmatrix} \begin{bmatrix} 28.44444445 & 23.11111112 \\ 23.11111112 & 18.77777778 \end{bmatrix} - \lambda I = 0$$

$$\Rightarrow \begin{bmatrix} 11.34391535 & 9.216931222 \\ 10.75133894 & 8.735462885 \end{bmatrix} - \lambda \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} = 0$$

$$\Rightarrow \begin{bmatrix} 11.34391535 - \lambda & 9.216931222 \\ 10.75133894 & 8.735462885 - \lambda \end{bmatrix} = 0$$

$$\Rightarrow (11.34391535 - \lambda)(8.735462885 - \lambda) - 9.216931222 \times 10.75133894 = 0$$

$$\Rightarrow 99.09435151 - 20.07937824\lambda + \lambda^2 - 99.09435155 = 0$$

$$\Rightarrow \lambda^2 - 20.07937824\lambda = 0$$

$$\Rightarrow \lambda(\lambda - 20.07937824) = 0$$

$$\therefore \lambda = 0, 20.07937824$$

for $\lambda = 0$,

$$\begin{bmatrix} 11.34391535-0 & 9.216931222 \\ 10.75133894 & 8.735462885-0 \end{bmatrix} \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} = 0$$

$$\therefore 11.34391535x_1 + 9.216931222x_2 = 0 \dots \dots \textcircled{1}$$

$$\therefore 10.75133894x_1 + 8.735462885x_2 = 0 \dots \dots \textcircled{11}$$

from equation ①,

$$x_1 = -\frac{9.216931222x_2}{11.34391535} = -0.8125x_2$$

from equation ⑪,

$$x_1 = \frac{-8.735462885x_2}{10.75133894} = -0.8125x_2$$

$$\therefore A = \sqrt{(-0.8125)^2 + (1)^2} = 1.288470508$$

$$\therefore e_1 \sim \begin{bmatrix} -0.8125 \\ A \\ \frac{1}{A} \end{bmatrix} \sim \begin{bmatrix} -0.631 \\ 0.776 \end{bmatrix}$$

for $\lambda = 20.07937824$,

$$\begin{bmatrix} 11.34391535-20.07937824 & 9.216931222 \\ 10.75133894 & 8.735462885-20.07937824 \end{bmatrix} \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} = 0$$

$$\therefore -8.73546289x_1 + 9.216931222x_2 = 0 \dots \dots \textcircled{111}$$

$$\therefore 10.75133894x_1 - 11.34391536x_2 = 0 \dots \dots \textcircled{1111}$$

from equation ⑪,

$$x_1 = \frac{9.216931222}{8.73546289}x_2 = 1.055x_2$$

from equation ⑪1,

$$x_1 = \frac{11.34391536}{10.75133894}x_2 = 1.055x_2$$

$$\therefore A = \sqrt{(1.055)^2 + (1)^2} = 1.45362478$$

$$\therefore e_2 \sim \begin{bmatrix} 1.055 \\ 1.45362478 \\ \frac{1}{1.45362478} \end{bmatrix} \sim \begin{bmatrix} 0.726 \\ 0.688 \end{bmatrix}$$

\therefore we get maximum value for $\lambda = 20.08$ (ans:)

Clustering [6-7 marks]

Fall-2020 :

5) b) step:01: $k=2$

$$c_1 = \{1.0, 1.0\} \quad [\text{ID}=1]$$

$$c_2 = \{5.0, 7.0\} \quad [\text{ID}=4]$$

step:02:

$$d(c_1, 1) = \sqrt{|1.0 - 1.0|^2 + |1.0 - 1.0|^2} = 0$$

$$d(c_2, 1) = \sqrt{|5.0 - 1.0|^2 + |7.0 - 1.0|^2} \approx 7.21$$

$$d(c_1, 2) = \sqrt{|1.0 - 1.5|^2 + |1.0 - 2.0|^2} \approx 1.12$$

$$d(c_2, 2) = \sqrt{|5.0 - 1.5|^2 + |7.0 - 2.0|^2} \approx 6.10$$

$$d(c_1, 3) = \sqrt{|1.0 - 3.0|^2 + |1.0 - 4.0|^2} \approx 3.61$$

$$d(c_2, 3) = \sqrt{|5.0 - 3.0|^2 + |7.0 - 4.0|^2} \approx 3.61$$

$$d(c_1, 4) = \sqrt{|1.0 - 5.0|^2 + |1.0 - 7.0|^2} \approx 7.21$$

$$d(c_2, 4) = \sqrt{|5.0 - 5.0|^2 + |7.0 - 7.0|^2} = 0$$

$$d(c_1, 5) = \sqrt{|1.0 - 3.5|^2 + |1.0 - 5.0|^2} \approx 4.72$$

$$d(c_2, 5) = \sqrt{|5.0 - 3.5|^2 + |7.0 - 5.0|^2} = 2.5$$

$$d(c_1, 6) = \sqrt{|1.0 - 4.5|^2 + |1.0 - 5.0|^2} \approx 5.32$$

$$d(c_2, 6) = \sqrt{|5.0 - 4.5|^2 + |7.0 - 5.0|^2} \approx 2.06$$

\therefore ID 1, 2, 3 are under cluster 1

\therefore ID 4, 5, 6 are under cluster 2

\therefore New centroids are:

$$c_1 = \left(\frac{1}{3} (1+1.5+3), \frac{1}{3} (1+2+4) \right) = (1.83, 2.33)$$

$$c_2 = \left(\frac{1}{3} (5+3.5+4.5), \frac{1}{3} (7+5+5) \right) = (4.33, 5.67)$$

Step: 03:

$$d(c_1, 1) = \sqrt{|1.83-1.0|^2 + |2.33-1.0|^2} \approx 1.57$$

$$d(c_2, 1) = \sqrt{|4.33-1.0|^2 + |5.67-1.0|^2} \approx 5.74$$

$$d(c_1, 2) = \sqrt{|1.83-1.5|^2 + |2.33-2.0|^2} \approx 0.47$$

$$d(c_2, 2) = \sqrt{|4.33-1.5|^2 + |5.67-2.0|^2} \approx 4.63$$

$$d(c_1, 3) = \sqrt{|1.83-3.0|^2 + |2.33-4.0|^2} \approx 2.04$$

$$d(c_2, 3) = \sqrt{|4.33-3.0|^2 + |5.67-4.0|^2} = 2.14$$

$$d(c_1, 4) = \sqrt{|1.83-5.0|^2 + |2.33-7.0|^2} = 5.64$$

$$d(c_2, 4) = \sqrt{|4.33-5.0|^2 + |5.67-7.0|^2} = 1.49$$

$$d(c_1, 5) = \sqrt{|1.83-3.5|^2 + |2.33-5.0|^2} = 3.15$$

$$d(c_2, 5) = \sqrt{|4.33-3.5|^2 + |5.67-5.0|^2} = 1.07$$

$$d(c_1, 6) = \sqrt{|1.83-4.5|^2 + |2.33-5.0|^2} = 3.78$$

$$d(c_2, 6) = \sqrt{|4.33-4.5|^2 + |5.67-5.0|^2} = 0.69$$

\therefore ID 1, 2, 3 will be in cluster 1

\therefore ID 4, 5, 6 will be in cluster 2

Therefore, there is no change in the clusters.

\therefore Two final clusters are: $\{1, 2, 3\}, \{4, 5, 6\}$

Sprünge - 2020:

3) b) i) Manhattan distance = $|22-20| + |1-0| + |42-36| + |10-8|$
= $2 + 1 + 6 + 2$
= 11 (ans:)

ii) Minkowski distance = $\sqrt[3]{|22-20|^3 + |1-0|^3 + |42-36|^3 + |10-8|^3}$
= $\sqrt[3]{2^3 + 1^3 + 6^3 + 2^3}$
= $\sqrt[3]{233}$
 ≈ 6.15 (ans:)

iii) Supremum distance = $\max(|22-20|, |1-0|, |42-36|, |10-8|)$
= $\max(2, 1, 6, 2)$
= 6 (ans:)

Sprünge - 2020:

3) a) step: 01:

	P1	P2	P3	P4	P5
P1	0	2	6	10	9
P2	2	0	5	9	8
P3	6	5	0	4	5
P4	10	9	4	0	3
P5	9	8	5	3	0

step: 02: P1, P2 are closest

	P1, P2	P3	P4	P5
P1, P2	0	5.5	9.5	8.5
P3	5.5	0	4	5
P4	9.5	4	0	3
P5	8.5	5	3	0

Step :03 :

		P1, P2	P3	P4, P5
		0	5.5	9
		5.5	0	4.5
P1, P2	P3	9	4.5	0

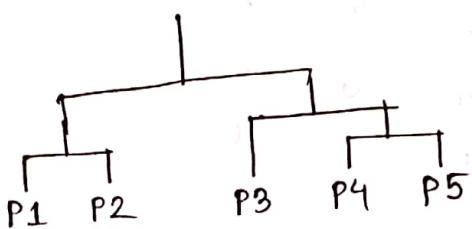
min

Step :04 :

		P1, P2	P3, (P4, P5)
		0	7.25
		7.25	0
P1, P2	P3, (P4, P5)		

\therefore Average distance = 7.25
 \therefore Clusters : $((P1, P2), (P3, (P4, P5)))$

Dendrogram:



Spring 2019:

1) a) Manhattan distance: $|x_1 - x_2| + |y_1 - y_2| + |z_1 - z_2|$

Distance calculation:

$$1, 2 = |0-1| + |0-1| + |0-0| = 1+1+0 = 2$$

$$1, 3 = |0-1| + |0-1| + |0-1| = 1+1+1 = 3$$

$$1, 4 = |0-1| + |0-1| + |0-0| = 1+1+0 = 2$$

$$2, 3 = |1-1| + |1-1| + |0-1| = 0+0+1 = 1$$

$$2, 4 = |1-1| + |1-1| + |0-0| = 0+0+0 = 0$$

$$3, 4 = |1-1| + |1-1| + |1-0| = 0+0+1 = 1$$

Step: 01:

	1	2	3	4
1	0	2	3	2
2	2	0	1	0
3	3	1	0	1
4	2	0	1	0

min

Step: 02:

	1	2, 4	3
1	0	2	3
2, 4	2	0	1
3	3	1	0
		min	

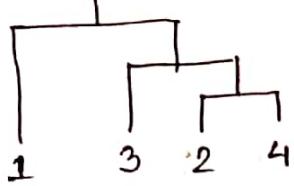
Step: 03:

	1	(2, 4), 3
1	0	2
(2, 4), 3	2	0

∴ Minimum distance = 2

∴ Clusters : $(1, ((2, 4), 3))$

Dendrogram:



Linear Discriminant Function [7-14 marks]

Fall- 2020:

1) b) Given, class 1, $\omega_1 = (-2)$
class 2, $\omega_2 = (51, 2)$

phi function, $\varphi = [x^2, x, 1]$

learning rate, $\eta = 0.45$

initial weight, $\omega(0) = [1 \ 1]^T$

for ω_1 : $\varphi = [4 \ -2 \ 1] \ [x_1 = -2]$

for ω_2 : $\varphi = [2601 \ 51 \ 1] \ [x_2 = 51]$

for ω_2 : $\varphi = [4 \ 2 \ 1] \ [x_3 = 2]$

Using normalization,

for ω_1 : $\varphi = [4 \ -2 \ 1]$

for ω_2 : $\varphi = [-2601 \ -51 \ -1]$
 $\varphi = [-4 \ -2 \ -1]$

We know: $\omega(t+1) = \omega(t) + \eta \sum_{\forall \varphi \text{ misclassified}} \varphi$

$$\begin{cases} \omega^T \varphi > 0 \rightarrow \text{cc} \\ \omega^T \varphi \leq 0 \rightarrow \text{mc} \end{cases}$$

φ	ω^T	$\omega^T \varphi$
4 -2 1	1 1 1	3 \rightarrow cc
-2601 -51 -1	1 1 1	-2653 \rightarrow mc
-4 -2 -1	-1169.45 -21.95 0.55	4699.2 \rightarrow cc
4 -2 1	-1169.45 -21.95 0.55	-4633.35 \rightarrow mc
-2601 -51 -1	-1167.65 -22.85 1	3038222 \rightarrow cc
-4 -2 -1	-1167.65 -22.85 1	4715.3 \rightarrow cc
4 -2 1	-1167.65 -22.85 1	-4623.9 \rightarrow mc
-2601 -51 -1	-1165.85 -23.75 1.45	3033585.65 \rightarrow cc
-4 -2 -1	-1165.85 -23.75 1.45	4709.45 \rightarrow cc

$$\therefore \omega^T = [-1165.85 \ -23.75 \ 1.45]$$

$$\therefore \varphi(\varphi) = \omega_1 \varphi_1 + \omega_2 \varphi_2 + \omega_0 = 0$$

$$\Rightarrow -1165.85 \varphi_1 - 23.75 \varphi_2 + 1.45 = 0$$

$$\therefore \text{Decision Boundary equation: } -1165.85 \varphi_1 - 23.75 \varphi_2 + 1.45 = 0 \quad (\text{ans.})$$

Spring-2020:

$$\begin{array}{c|c}
 \omega_1 & \omega_2 \\
 \hline
 \begin{bmatrix} 2 \\ 3 \end{bmatrix} \begin{bmatrix} 1 \\ 51 \end{bmatrix} & \begin{bmatrix} 3 \\ 1 \end{bmatrix} \begin{bmatrix} -2 \\ -4 \end{bmatrix} \\
 \hline
 \begin{bmatrix} 4 \\ 9 \\ -1 \\ 6 \end{bmatrix} \begin{bmatrix} 1 \\ 2601 \\ -50 \\ 51 \end{bmatrix} & \begin{bmatrix} 9 \\ 1 \\ 2 \\ 3 \end{bmatrix} \begin{bmatrix} 4 \\ 16 \\ 2 \\ 8 \end{bmatrix} \\
 \hline
 \begin{bmatrix} 4 \\ 9 \\ -1 \\ 6 \end{bmatrix} \begin{bmatrix} 1 \\ 2601 \\ -50 \\ 51 \end{bmatrix} & \begin{bmatrix} -9 \\ -1 \\ -2 \\ -3 \end{bmatrix} \begin{bmatrix} -4 \\ -16 \\ -2 \\ -8 \end{bmatrix}
 \end{array}$$

[normalization]

$$\begin{cases} \omega^T \gamma > 0 \rightarrow \text{cc} \\ \omega^T \gamma \leq 0 \rightarrow \text{mc} \end{cases}$$

We know: $\omega(t+1) = \omega(t) + \eta \sum_{\text{misclassified}} \gamma$

Here, $\eta = 2$, $\omega = \begin{bmatrix} 0.5 \\ 0.5 \\ 0.5 \\ 0.5 \end{bmatrix}$

γ	ω^T	$\omega^T \gamma$
4 9 -1 6	0.5 0.5 0.5 0.5	9 → cc
1 2601 -50 51	0.5 0.5 0.5 0.5	1301.5 → cc
-9 -1 -2 -3	0.5 0.5 0.5 0.5	-7.5 → mc
-4 -16 -2 -8	-17.5 -1.5 -3.5 -5.5	145 → cc
4 9 -1 6	-17.5 -1.5 -3.5 -5.5	-113 → mc
1 2601 -50 51	-9.5 16.5 -5.5 6.5	43513.5 → cc
-9 -1 -2 -3	-9.5 16.5 -5.5 6.5	60.5 → cc
-4 -16 -2 -8	-9.5 16.5 -5.5 6.5	-267 → mc

$$\therefore \omega^T = [-9.5 \quad 16.5 \quad -5.5 \quad 6.5]$$

$$\therefore g(\gamma) = -9.5\gamma_1 + 16.5\gamma_2 - 5.5\gamma_3 + 6.5 = 0$$

$$\therefore \text{Decision boundary equation: } -9.5\gamma_1 + 16.5\gamma_2 - 5.5\gamma_3 + 6.5 = 0$$

(Ans:)

Fall-2019:

1) a) Hence,

$$\mu_1 = 0$$

$$\mu_2 = \begin{bmatrix} 1 \\ 1 \end{bmatrix}$$

$$\Sigma_1 = \begin{bmatrix} 3 & 0.5 \\ 0.5 & 3 \end{bmatrix}$$

$$\Sigma_2 = \begin{bmatrix} 6 \\ 5 \\ 6 \end{bmatrix}$$

decision boundary :

$$g_i(x) = g_{-i}(x)$$

$$\Rightarrow -\frac{1}{2} (x - \mu_1)^T \Sigma_1^{-1} (x - \mu_1) = -\frac{1}{2} (x - \mu_2)^T \Sigma_2^{-1} (x - \mu_2) \quad [P(\omega_1) = P(\omega_2)]$$

$$\Rightarrow \frac{-\|x - \mu_1\|^2}{2\Sigma_1} = \frac{-\|x - \mu_2\|^2}{2\Sigma_2}$$

$$\Rightarrow \frac{\|x\|^2 - 2 \cdot x \cdot \mu_1^t + \|\mu_1\|^2}{2 \Sigma_1} = \frac{\|x\|^2 - 2 \cdot x \cdot \mu_2^t + \|\mu_2\|^2}{2 \Sigma_2}$$

$$\Rightarrow \frac{x \cdot x}{2\Sigma_1} - \frac{2x\mu_1^t}{2\Sigma_1} + \frac{\mu_1^t \cdot \mu_1}{2\Sigma_1} = \frac{x \cdot x}{2\Sigma_2} - \frac{2x\mu_2^t}{2\Sigma_2} + \frac{\mu_2^t \cdot \mu_2}{2\Sigma_2}$$

$$\Rightarrow -\frac{2x\mu_1^t}{2\Sigma_1} + \frac{\mu_1^t \cdot \mu_1}{2\Sigma_1} = -\frac{2x\mu_2^t}{2\Sigma_2} + \frac{\mu_2^t \cdot \mu_2}{2\Sigma_2} \quad [x^t \cdot x = \text{constant}]$$

$$\Rightarrow \frac{2x\mu_1^t}{2\Sigma_1} - \frac{2x\mu_2^t}{2\Sigma_2} - \frac{\|\mu_1\|^2}{2\Sigma_1} + \frac{\|\mu_2\|^2}{2\Sigma_2} = 0$$

$$\Rightarrow \frac{2x \cdot 0}{\Sigma_1} - \frac{2x \mu_2^t}{2\Sigma_2} - \frac{\sqrt{\sigma^2}}{2\Sigma_1} + \frac{\mu_2^t \cdot \mu_2}{2\Sigma_2} = 0$$

$$\Rightarrow \Sigma_2^{-1} x \mu_2^t - \frac{\Sigma_2^{-1} \mu_2^t \cdot \mu_2^t}{2} = 0$$

if $\Sigma_2^{-1} \mu_2^t = w^t$; $\frac{\mu_2}{2} = x_0$ then from equation ①,

$w^t (x - x_0) = 0$
which is the equation of decision boundary.

The hyperplane won't be perpendicular with w^t
because there is Σ_2^{-1} . Σ_2^{-1} explains that the direction
isn't towards hyperplane.

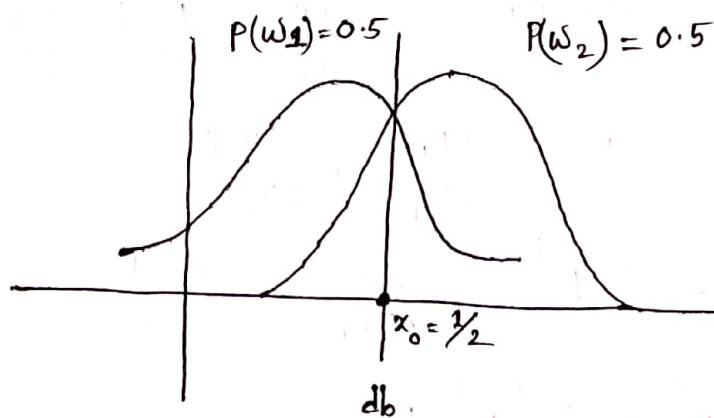
putting the values in equation ①,

$$\begin{bmatrix} 6 & 5 \\ 5 & 6 \end{bmatrix}^{-1} \begin{bmatrix} 1 & 1 \end{bmatrix} \left(x - \begin{bmatrix} \frac{1}{2} \\ \frac{1}{2} \end{bmatrix} \right) = 0$$

$$\Rightarrow \frac{1}{36-25} \begin{bmatrix} 6 & -5 \\ -5 & 6 \end{bmatrix} \begin{bmatrix} 1 & 1 \end{bmatrix} \left(x - \begin{bmatrix} \frac{1}{2} \\ \frac{1}{2} \end{bmatrix} \right) = 0$$

$$\Rightarrow \frac{1}{11} \begin{bmatrix} 6 & -5 \\ -5 & 6 \end{bmatrix} \begin{bmatrix} 1 & 1 \end{bmatrix} \left(x - \begin{bmatrix} \frac{1}{2} \\ \frac{1}{2} \end{bmatrix} \right) = 0$$

$$\Rightarrow \underbrace{\begin{bmatrix} \frac{6}{11} & \frac{-5}{11} \\ \frac{-5}{11} & \frac{6}{11} \end{bmatrix}}_{w^T} \begin{bmatrix} 1 & 1 \end{bmatrix} \left(x - \underbrace{\begin{bmatrix} \frac{1}{2} \\ \frac{1}{2} \end{bmatrix}}_{x_0} \right) = 0$$



Fall-2019:

Instance	F1	F2	Class
x_1	0.1	0.4	+
x_2	0.3	0.6	+
x_3	0.2	0.2	+
	$\mu = 0.2$	$\mu = 0.4$	
	$\sigma^2 = 0.01$	$\sigma^2 = 0.04$	

$$\therefore P(F_1 = 0.9 | +) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(0.2 - 0.9)^2}{2\sigma^2}}$$

$$= \frac{1}{\sqrt{2\pi \times 0.01}} e^{-\frac{0.49}{2 \times 0.01}}$$

$$= 9.133 \times 10^{-11} - \frac{(0.4 - 0.4)^2}{2 \times 0.04}$$

$$\therefore P(F_2 = 0.4 | +) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(0.4 - 0.4)^2}{2\sigma^2}}$$

$$= 1.994310088$$

$$\therefore P(F_1 = 0.9, F_2 = 0.4 | +) = 9.133 \times 10^{-11} \times 1.994310088$$

$$= 1.82 \times 10^{-10}$$

$$\therefore P(x_6 | +) P(+) = 1.82 \times 10^{-10} \times \frac{3}{5} = 1.092 \times 10^{-10}$$

Instance	F1	F2	Class
x_4	0.7	0.1	-
x_5	0.8	0.3	-
	$\mu = 0.75$	$\mu = 0.2$	
	$\sigma^2 = 0.005$	$\sigma^2 = 0.02$	

$$\therefore P(F_1 = 0.9 | -) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(0.75 - 0.9)^2}{2\sigma^2}} = 0.5945318088$$

$$\therefore P(F_2 = 0.4 | -) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(0.2 - 0.9)^2}{2\sigma^2}} = 1.35 \times 10^{-5}$$

$$\therefore P(F_1 = 0.9, F_2 = 0.4 | -) = 0.5945318088 \times 1.35 \times 10^{-5}$$

$$= 8.024 \times 10^{-6} \quad P(x_6 | -) P(-) = 3.21 \times 10^{-6}$$

Though $P(x_6 | +) P(+)$ $P(x_6 | -) P(-)$, class will be = “-”

Fall-2020:

$$\begin{aligned}4) \text{ a) i) } P(B) &= P(B, A) + P(B, \sim A) \\&= P(B|A) P(A) + P(B|\sim A) P(\sim A) \\&= 0.7 \times 0.3 + 0.5 \times (1-0.3) \\&= 0.21 + 0.35 \\&= 0.56 \\&\therefore P(A|B) = \frac{P(B|A) P(A)}{P(B)} = \frac{0.7 \times 0.3}{0.56} = 0.375\end{aligned}$$

$$\begin{aligned}\text{ii) } P(C|B) &= P(C|A, B) P(A|B) + P(C|\sim A, B) P(\sim A|B) \\&= P(C|A) P(A|B) + P(C|\sim A) P(\sim A|B) \\&= 0.2 \times 0.375 + 0.6 \times 0.625 \\&= 0.45\end{aligned}$$

$$P(\sim A|B) = \frac{P(B|\sim A) P(\sim A)}{P(B)} = \frac{0.5 \times (1-0.3)}{0.56} = 0.625$$

iii)

Spring-2020 :

$$\begin{aligned}
 5) b) P(HD=Yes) &= P(HD | E=Yes, D=Unhealthy) \times P(E=Yes) \times P(D=Unhealthy) \\
 &= 0.45 \times 0.7 \times (1-0.25) \\
 &= 0.23625
 \end{aligned}$$

$$\begin{aligned}
 P(HD=No) &= P(HD | E=Yes, D=Unhealthy) \times P(E=Yes) \times P(D=Unhealthy) \\
 &= (1-0.45) \times 0.7 \times (1-0.25) \\
 &= 0.28875
 \end{aligned}$$

$$\begin{aligned}
 P(Hb=Yes) &= P(Hb | D=Unhealthy) \times P(D=Unhealthy) \quad [Hb \text{ independent to } E] \\
 &= 0.85 \times (1-0.25) \\
 &= 0.6375
 \end{aligned}$$

$$\begin{aligned}
 P(Hb=No) &= P(Hb | D=Unhealthy) \times P(D=Unhealthy) \\
 &= (1-0.85) \times (1-0.25) \\
 &= 0.1125
 \end{aligned}$$

Spring-2020 :

7) b) i)

A	B	C	Class
0	0	0	+
0	0	1	+
1	0	1	+
1	1	1	+
1	0	1	+

A	B	C	Class
0	0	1	-
0	1	1	-
0	1	1	-
1	0	1	-
1	0	1	-

$$\therefore P(A|+) = P(A=0|+) \times P(A=1|+) \\
 = \frac{2}{5} \times \frac{3}{5} = 0.24$$

$$\therefore P(B|+) = P(B=0|+) \times P(B=1|+) \\
 = \frac{4}{5} \times \frac{1}{5} = 0.16$$

$$\therefore P(C|+) = P(C=0|+) \times P(C=1|+) \\
 = \frac{1}{5} \times \frac{4}{5} = 0.16$$

$$\therefore P(A|-) = P(A=0|-) \times P(A=1|-) \\
 = \frac{3}{5} \times \frac{2}{5} = 0.24$$

$$\therefore P(B|-) = P(B=0|-) \times P(B=1|-) \\
 = \frac{3}{5} \times \frac{2}{5} = 0.24$$

$$\therefore P(C|-) = P(C=0|-) \times P(C=1|-) \\
 = \frac{0}{5} \times \frac{5}{5} = 0$$

ii) For class = '+'

$$P(A=0, B=1, C=0) = \frac{2}{5} \times \frac{1}{5} \times \frac{1}{5} \\
 = 0.016$$

For class = '-'

$$P(A=0, B=1, C=0) = \frac{3}{5} \times \frac{2}{5} \times \frac{0}{5} \\
 = 0$$

$\therefore P(0,1,0|+) > P(0,1,0|-)$; class will be = '+'

Fall - 2020:

Size	Color	Target
XS	green	Yes
L	green	Yes
XL	green	Yes
XS	white	Yes
M	green	Yes

Size	Color	Target
XS	white	No
M	black	No
L	black	No

$$P(L, \text{white} | \text{Yes}) = P(L | \text{Yes}) \times P(\text{white} | \text{Yes})$$
$$= \frac{1}{5} \times \frac{1}{5}$$

$$= 0.04$$

$$P(L, \text{white} | \text{No}) = P(L | \text{No}) \times P(\text{white} | \text{No})$$
$$= \frac{1}{3} \times \frac{1}{3}$$

$$\approx 0.111111$$

$$P(L, \text{white} | \text{Yes}) \times P(\text{Yes}) = 0.04 \times \frac{5}{8} = 0.025$$

$$P(L, \text{white} | \text{No}) \times P(\text{No}) = 0.111 \times \frac{3}{8} = 0.042$$

∴ Target will be = 'No'

Adaboost (4-7 marks)

Algorithm :

Algorithm :

1) For $i=1$ to N initialize the data weight $w_1^{(i)} = \frac{1}{N}$

2) for $t = 1$ to T

a) Find a classifier $h_t(x)$ by minimizing the weighted error function:

$$\mathcal{L}_t = \sum_{i=1}^N w_t^{(i)} \times I(x^{(i)} \neq h_t(x^{(i)}))$$

b) Find the weighted errors of $h_t(x)$:

$$e_t = \frac{\sum_{i=1}^N \omega_t^{(i)} \times I(x^{(i)} \neq h_t(x^{(i)}))}{\sum_{i=1}^N \omega_t^{(i)}}$$

and the new component is assigned votes based on its error:

$$d_t = \ln \left(\frac{1 - e_t}{e_t} \right)$$

c) The normalized weights are updated:

$$\omega_{t+1}^{(i)} = \omega_t^{(i)} e^{\alpha_t I(\hat{y}^{(i)} \neq h_t(x^{(i)}))}$$

3) Combined classifier $\hat{y} = \text{sign}(H_T(x))$ where $H_T(x) = \sum_{t=1}^M \alpha_t h_t(x)$

Example :

Step: 01:

$$E_1 = \frac{3}{10} \quad \alpha_1 = \frac{1}{2} \ln \frac{1 - \frac{3}{10}}{\frac{3}{10}} = \ln \sqrt{\frac{7}{3}} \approx 0.42$$

$\alpha_1 = \frac{\pi}{2} \ln \frac{3}{10}$ upweighting misclassified circled points: $\ln \sqrt{3/2}$

$$W_1 = \frac{1}{10} e^{\ln \sqrt{7/3}} \approx 0.1528$$

$$W_1 = \frac{1}{10} e^{-0.1520} \approx 0.1520$$

downweighting correctly classified points:

$$W_1^{(i)} = \frac{1}{10} e^{-10\sqrt{3/3}} \approx 0.0655$$

Step:02:

+	+	-
+	⊖	-
+	⊖	-
+	-	-
h ₂		

$$\epsilon_2 = 3 \times 0.0655 = 0.1965$$

$$\alpha_2 = \frac{1}{2} \ln \frac{1 - 0.1965}{0.1965} = 0.7042$$

upweighting misclassified circled points:

$$W_2^{(i)} = 0.0655 e^{0.7042} \approx 0.1325$$

downweighting correctly classified (small +/-):

$$W_2^{(i)} = 0.0655 e^{-0.7042} \approx 0.0324$$

downweighting correctly classified (large +):

$$W_2^{(i)} = 0.1325 e^{-0.7042} \approx 0.0655$$

Step:03:

+	+	⊖
⊖	-	-
⊖	-	-
h ₃		

$$\epsilon_3 = 3 \times 0.0324 = 0.0972$$

$$\alpha_3 = \frac{1}{2} \ln \frac{1 - 0.0972}{0.0972} \approx 1.1144$$

upweighting misclassified circled points:

$$W_3^{(i)} = 0.0324 e^{1.1144} \approx 0.0988$$

Final:

$$H_{\text{final}} = \text{sign} (0.42 \boxed{} + 0.7 \boxed{} + 1.1 \boxed{})$$

	+	+	-
+	-	-	-
+	-	-	-

SVM [7 marks]

Fall-2020

$$\text{3) b) } L(x, y, \lambda) = f(x, y) - \lambda_1 g_1(x, y) - \lambda_2 g_2(x, y) \\ = x^2 + y^2 - \lambda_1(x+1) - \lambda_2(y+1) \\ = x^2 + y^2 - \lambda_1 x - \lambda_1 - \lambda_2 y - \lambda_2$$

We solve from the gradient of the lagrangian,

$$\nabla L(x, y, \lambda) = \nabla f(x, y) - \lambda_1 \nabla g_1(x, y) - \lambda_2 \nabla g_2(x, y) = 0$$

$$\frac{\partial}{\partial x}(L) = 2x - \lambda_1 = 0 \quad \dots \quad \dots \quad \dots \quad \dots \quad \textcircled{1}$$

$$\frac{\partial}{\partial y}(L) = 2y - \lambda_2 = 0 \quad \dots \quad \dots \quad \dots \quad \dots \quad \textcircled{2}$$

$$\frac{\partial}{\partial \lambda_1}(L) = -x - 1 = 0 \quad \dots \quad \dots \quad \dots \quad \dots \quad \textcircled{3}$$

$$\frac{\partial}{\partial \lambda_2}(L) = -y - 1 = 0 \quad \dots \quad \dots \quad \dots \quad \dots \quad \textcircled{4}$$

From equation $\textcircled{3}$,

$$-x = 1$$

$$\therefore x = -1$$

From equation $\textcircled{4}$,

$$-y = 1$$

$$\therefore y = -1$$

From equation $\textcircled{1}$,

$$2x(-1) = \lambda_1$$

$$\therefore \lambda_1 = -2$$

From equation $\textcircled{2}$,

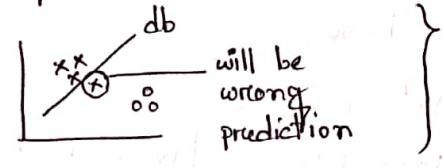
$$2x(-1) = \lambda_2$$

$$\therefore \lambda_2 = -2$$

$$\therefore f(x, y) = x^2 + y^2 = (-1)^2 + (-1)^2 = 1 + 1 = 2$$

(ans)

SVM (Support Vector Machine) : এটা মুক্ত classifier type algorithm.

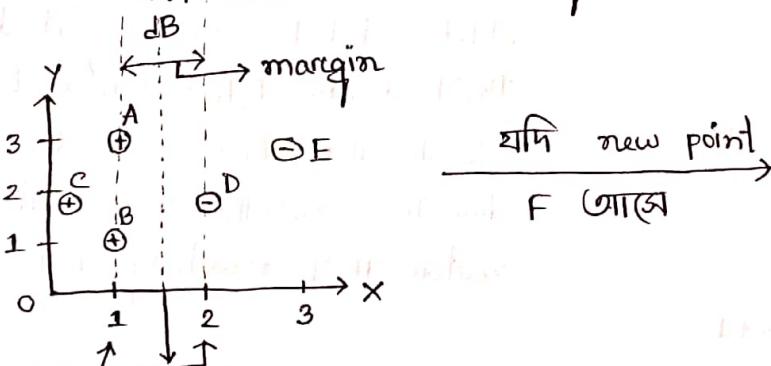


আই decision boundary (db) মুক্ত হওয়ার সময় কর্তৃত করা হবে যেন কোন ক্লাস থেকে same distance maintain করা। db এর দুই পার্শে মধ্যে gap consider করা হয়ে SVM এ যা perceptron classifier এ করা হয় না।

margin বাড়ানোর টাক্ষণ্য ব্যবহৃত SVM এর উপরে।

মানে classifier করাটা

শোকালাকা both class এর margin increase করা।



1 মধ্যে ক্লাস অংশ দ্বারা dotted line = margin line / gutter
2 " " " " " " = " "

margin = 1 এর 2 মধ্যের distance

A, B, C, D, E points are vector

যে pointগুলো margin এর উপর দিয়ে গিয়েছে আবু যে pointগুলো db করে decide করতে help করে এই pointগুলোই support vector।

So, A, B, D are support vectors whereas C, E are non-support vectors

gutter : • a line through closest + and - points near db

• parallel to db

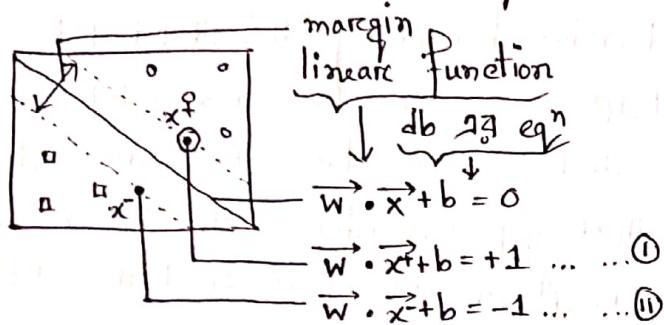
• not a region, it's a line

sv : • points that decide db

margin : • distance between the two gutter

* if we change the points and it won't effect on db then those points are not SV.

কোন db better? ans: যে hyperplane margin কে maximize করলে এটা।



imagine, $o = (+ve)$; $\square = (-ve)$

$$\therefore \text{margin} = \frac{2}{\|w\|^2}$$

from eqn ① and ④,

$$\vec{w} \cdot \vec{x}^+ + b = 0 \dots \dots \text{③}$$

$$\vec{w} \cdot \vec{x}^- + b = 0 \dots \dots \text{④}$$

from eqn ③ and ④,

$$\vec{w} \cdot \vec{x}^+ + b + 1 = \vec{w} \cdot \vec{x}^- + b + 1$$

$$\Rightarrow \vec{w} \cdot \vec{x}^+ - 1 = \vec{w} \cdot \vec{x}^- + 1$$

$$\Rightarrow \vec{w} \cdot \vec{x}^+ - \vec{w} \cdot \vec{x}^- = 1 + 1$$

$$\Rightarrow \vec{w} (\vec{x}^+ - \vec{x}^-) = 2 \dots \dots \text{⑤}$$

from figure,

$$\text{margin} = (x^+ - x^-) \cdot \frac{w}{\|w\|}$$

$$= \frac{2}{\|w\|} \quad [\text{from } \text{⑤}]$$

We know,

$$\begin{aligned} w^T x + w_0 &= 0 \\ \Rightarrow w \cdot x + (w_0) &= 0 \quad [w^T x = w \cdot x] \\ &\downarrow \text{bias} \end{aligned}$$

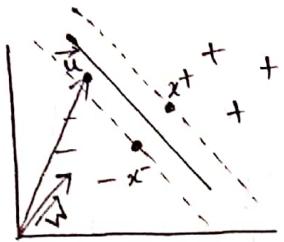
কোন point db এর উপর থাকলে জিতের eqn না হয়, তার db এর eqn না হয়।

কোন ভেক্টরের তার magnitude দিয়ে divide করলে unit ভেক্টর

পাওয়া যাব। কোন line এর direction পাওয়ার উপর unit vector দিলে multiply করা লাগে।

distance এর জন্য: $(x^+ - x^-)$

direction এর জন্য: $\frac{w}{\|w\|}$



যদি $\vec{w} \cdot \vec{x}_i + b \geq 0$ এবং class যের কাছাতে রাখ, তখন

$\vec{w} \cdot \vec{x}_i + b \geq 0$ হল (tve) [decision rule]

put additional constraints to calculate w and b

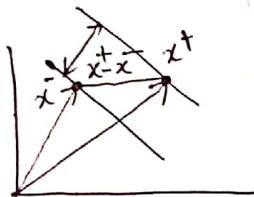
$$\vec{w} \cdot \vec{x}_i + b \geq 1 \rightarrow (\text{tve}) \dots \dots \text{VI}$$

$$\vec{w} \cdot \vec{x}_i + b \leq -1 \rightarrow (\text{rve}) \dots \dots \text{VII}$$

2 eqn are painful to carry. So, introducing new variable to make the math convenient.

y_i such that: $y_i = +$ for tve samples
 $y_i = -$ for rve samples

$$\therefore y_i (\vec{w} \cdot \vec{x}_i + b) \geq 1$$



$$\begin{aligned} \text{margin/width} &= (x^+ - x^-) \cdot \frac{\vec{w}}{\|\vec{w}\|} \\ &= \frac{x^+ \vec{w} - x^- \vec{w}}{\|\vec{w}\|} \\ &= \frac{(1-b) - (-1-b)}{\|\vec{w}\|} \quad [\text{from VI, VII}] \\ &= \frac{1-b+1+b}{\|\vec{w}\|} \\ &= \frac{2}{\|\vec{w}\|} \end{aligned}$$

$\max \frac{2}{\|\vec{w}\|}$ দ্বারা কারণ margin max করাই
 SVM এর target.

math কে easy করার জন্য:

$$\max \frac{2}{\|\vec{w}\|}$$

$$\approx \max \frac{1}{\|\vec{w}\|}$$

$$\approx \min \|\vec{w}\|$$

$$\approx \min \frac{\|\vec{w}\|^2}{2}$$

constraints: $y_i (\vec{w} \cdot \vec{x}_i + b) \geq 1$

optimization problem + কীর্তন
 তাই differentiation দিয়ে
 solve করা যাবেনা। একেন্দ্র
 lagrange multipliers লাগে,

$\boxed{\max / \min \text{ করা } \text{আস্তি}}$
 constraint থাকলে lagrange
 multipliers ব্যবহার কর

Lagrange multipliers:

Given, $F(x, y) = 2x^2 + 2y^2$ (function) find extreme value.
 $g(x, y) = x + y - 1 = 0$ (constraint) $\{ = \text{for equal constraint}$
 $>/< : \text{inequal constraint}$

Solution:

$$L(x, y, \lambda) = F(x, y) - \lambda g(x, y) \quad \left. \begin{array}{l} \lambda = \text{multiplier} \\ \text{Lagrange format transferred} \end{array} \right\}$$

$$= 2x^2 + 2y^2 - \lambda(x + y - 1)$$

We solve from the gradient of the lagrangian,

$$\nabla L(x, y, \lambda) = \nabla F(x, y) - \lambda \nabla g(x, y) = 0$$

$$\frac{\partial}{\partial x}(L) = -2x - \lambda = 0 \quad \dots \quad \dots \quad \dots \quad \dots \quad \textcircled{1}$$

$$\frac{\partial}{\partial y}(L) = -4y - \lambda = 0 \quad \dots \quad \dots \quad \dots \quad \dots \quad \textcircled{2}$$

$$\frac{\partial}{\partial \lambda}(L) = x + y - 1 = 0 \quad \dots \quad \dots \quad \dots \quad \dots \quad \textcircled{3}$$

From ① and ②, Putting $x = 2y$ in ③,

$$\lambda = -2x$$

$$2y + y - 1 = 0$$

$$\lambda = -4y$$

$$\Rightarrow 3y = 1$$

$$\therefore -2x = -4y$$

$$\therefore y = \frac{1}{3}$$

$$\Rightarrow x = \frac{-4y}{-2}$$

$$\therefore x = \frac{2}{3}$$

$$\therefore x = 2y$$

$$\therefore \lambda = \frac{-4}{3}$$

$$\therefore F(x, y) = 2 - \left(\frac{2}{3}\right)^2 - 2\left(\frac{1}{3}\right)^2 = 2 - \frac{4}{9} - \frac{2}{9} = \frac{18-4-2}{9} = \frac{12}{9} = \frac{4}{3}$$

For our topic:

$$L = \frac{1}{2} \|\vec{w}\|^2 - \sum \alpha_i [y_i (\vec{w} \cdot \vec{x}_i + b) - 1]$$

$$\therefore \frac{\partial L}{\partial \vec{w}} = \vec{w} - \sum \alpha_i y_i \vec{x}_i = 0 \quad \therefore \vec{w} = \sum \alpha_i y_i \vec{x}_i$$

$$\therefore \frac{\partial L}{\partial b} = -\sum \alpha_i y_i = 0 \quad \therefore \sum \alpha_i y_i = 0$$

$\alpha = \lambda$
 $\sum \text{কার্যন : বন্ড কোনো}$
 জো জানিনা

$y_i = +ve / -ve$

solve
dual optimization
द्वारा गणितीय
dot product वर्तन

putting the value of \vec{w} in L,

$$L = \frac{1}{2} \left(\sum_i \alpha_i y_i \vec{x}_i \right) \cdot \left(\sum_j \alpha_j y_j \vec{x}_j \right) - \sum_i \alpha_i y_i \left(\sum_j \alpha_j y_j x_j \right) - \sum_i \alpha_i y_i b + \sum_i \alpha_i$$

$$\rightarrow = \sum_i \alpha_i - \frac{1}{2} \sum_i \sum_j \alpha_i \alpha_j y_i y_j \vec{x}_i \cdot \vec{x}_j \quad [\because \sum_i \alpha_i y_i = 0] \quad \begin{array}{l} \text{quadratic eq^n} \\ \text{example} \\ \vec{a}x + b\vec{x} + c = 0 \end{array}$$

unknown = α ; feature = \vec{x} ; label = y $\{x_i^T x_j = k(x_i, x_j)\}$

$$\therefore f(x) = \sum_i \alpha_i y_i x_i^T x + b$$

Hence, parameters are multiplied to each other in L.

Q, it's a quadratic optimization problem. qP solver is used.

qP solver : quadratic problem solver [for finding α]

all the math depends on dot product.

decision rule : $\vec{w} \cdot \vec{u} + b \geq 0$

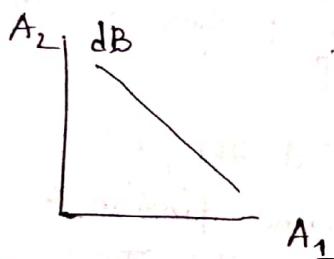
$$\Rightarrow \sum_i \alpha_i y_i \vec{x}_i \cdot \vec{u} + b = 0 \quad [\because \vec{w} = \sum_i \alpha_i \vec{x}_i]$$

ए एक अन्य अन्य $\alpha > 0$: support vector

" " " " " $\alpha = 0$: non-support vector

Problem

A_1	A_2	y	λ_i
0.38	0.47	+	65.52
0.49	0.61	-	65.52
0.92	0.91	-	0



$$b_1 = 1 - w_1 x_1 \\ = 1 - \{(-6.64) \times 0.38 + (-9.32) \times 0.47\}$$

$$= 7.9036 \\ b_2 = 1 - \{(-6.64) \times 0.49 + (-9.32) \times 0.61\} = 9.9388$$

Solⁿ : $\lambda_i = 0$ not considerable because $\lambda_i > 0$ are SV only.

$$\therefore w_1 = \sum \lambda_i y_i x_{i1} \\ = 65.52 \times 1 \times 0.38 + 65.52 \times (-1) \times 0.49 \\ = -6.64$$

$$\therefore w_2 = \sum \lambda_i y_i x_{i2} \\ = 65.52 \times 1 \times 0.47 + 65.52 \times (-1) \times 0.61 \\ = -9.32$$

$$\therefore \text{dB: } w_1 x_1 + w_2 x_2 + b = 0$$

$$\therefore 1 - 6.64 x_1 - 9.32 x_2 + 8.9212 = 0$$

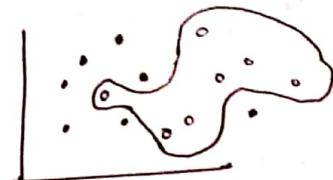
MMH = maximum margin hyperplane

$$\therefore \text{Average, } b = \frac{b_1 + b_2}{2} = 8.9212$$

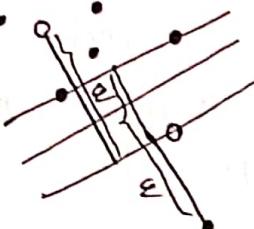
$x = (0.5, 0.5)$, କୌଣ କୌଣ class?

$$\begin{aligned}
 \hat{y}(x) &= w \cdot x + b \\
 &= -6.64 \times 0.5 - 9.32 \times 0.5 + 8.9212 \\
 &= 0.9412 \\
 &= +ve
 \end{aligned}$$

Hard margin : overfitting
sv



Soft margin:
sv



$$y_i (w^T x_i + b) \geq 1 - \epsilon_i, [\epsilon_i \geq 0]$$

$$w^2 = w^T w$$

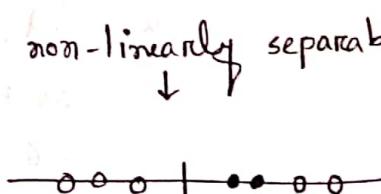
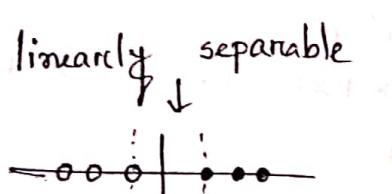
For hard : $\phi(w) = \frac{1}{2} w^T w$

$$y_i (w^T x_i + b) \geq 1$$

For soft : $\phi(w) = \frac{1}{2} w^T w + C \sum \epsilon_i$

$$y_i (w^T x_i + b) \geq 1 - \epsilon_i$$

misclassified ଶଳେ cost
বାଜୁଣେ : $C = \text{control}$
overfitting



high
dimensional
space ନିବି
(ଫାଇନ୍କ୍ରିଚନ୍ ପରି ଗାନ୍ଧୀର)

each point କେ high dimension ଏ ନିବି

ଅନେକ computation ଲାଗୁଇ, ϕ function ବେ କରେ ଏହି point

କେ ଆଜେ high dimension ଏ ଲାଗୁ ହସ୍ତ, ଏଥିର kernel function ଏ
ଏ ଏହି point ଦିଲେ ଦିଲେ ଫାଇନ୍କ୍ରିଚନ୍ ପରି କାର୍ଡିନେ ଏକାଏକେ explicitly କରସୁଥିବା,
ଫଳେ computation low ହସ୍ତ ଥାଏଇ

कार्ड इच्छ kernel function के parameters जागत्तु छाँट्ले एवं पूर्जनका dot product एवं relation के थाकत्तु इसे else kernel function का बहुत ही कार्ड बनता है।

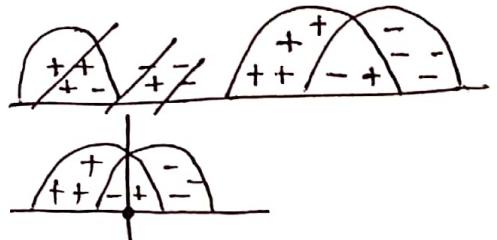
Bayes Theorem : Model दूरी वित्तनेर रूप : generative, discriminative/linear, perceptron, SVM → discriminative → {प्रतिक्रिया db draw करे एवं error or misclassified करे रूप

आगे data distribution करे करे, then जोड़कर classify करे : generative
 mean, variance/standard deviation एवं असंगति करे करे

Gaussian/Bernoulli/Poisson
normal

step-1 : mean + sd द्वारा एक class करे करे :

step-2 : threshold value set करे :



Bayes theorem, Naive bayes → generative

Random process : या accurately बता याएँ ना, guess करा लाएँ।

Example : flip a coin, rolling a dice, measuring the rain

Random process के output द्वारा आकर quantify करे random variable!

Capital letter द्वारा random variable represent करा रूप। ऐन : X, Y

For a coin toss : $X = \{h, t\}$ means coin toss is a random process

याएँ output h, t एवं इन्हें जीवानक बातें quantify कराहि 'X' द्वारा।

वही random process इत्यार probability कर जाएँ इच्छ event।

$P(x)$: जाने X पर value x आजाएँ probability कर।

$P(X=h)$ जाने X पर value h आजाएँ probability कर।

Discrete random variable : finite : single value एवं उपर काजू करे : countable

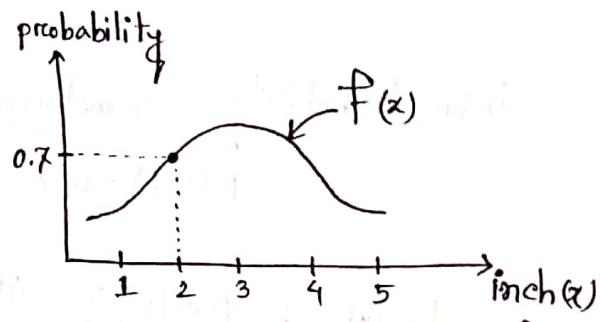
Continuous random variable : infinite : अजीब value एवं उपर काजू करे : uncountable

PDF (Probability Density Function) : निको area under the curve 3 बले।

continuous variable small letter द्वारा प्रकाश करे : $p(x|w_j)$

probability capital letter द्वारा प्रकाश करे : $P(z_j|x)$

$Y = \text{exact amount of rain tomorrow}$
 continuous random variable

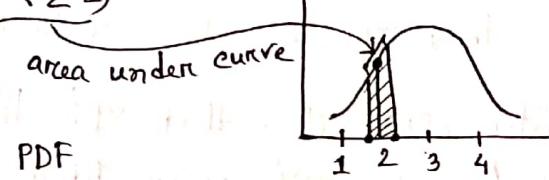


$\therefore P(Y=2) = 0.7$ यानि याघ्ना काढने प्रेत
 continuous random variable

तरीके ques ए 'Y=2' अद्यते यानि याघ्ना ना।

ques can be $\rightarrow P(1.9 < Y < 2.1)$

$$\therefore P(1.9 < Y < 2.1) = \int_{1.9}^{2.1} f(x) dx = \text{PDF}$$



काढकृत rain एवं
 assume काढनारा

area of line = 0

अज्ञात

area of range
 देव वाढा लागो

joint probability

$$P(A) = \frac{70}{100} = 0.7$$

$$P(B) = \frac{30}{100} = 0.3$$

$$\left\{ \begin{array}{l} P(A, C) = \frac{35}{100} = 0.35 \\ P(B, C) = \frac{15}{100} = 0.15 \end{array} \right\}$$

A	C	B
70	35	15
		30

Conditional Probability:

A एवं respect ए C एवं probability $P(C|A) = \frac{35}{70} = 0.5$

B एवं respect ए C एवं probability $P(C|B) = \frac{15}{30} = 0.5$

$$\therefore P(A, C) = P(C|A) P(A) \quad \left. \begin{array}{l} \text{Product Rule} \\ \Rightarrow 0.35 = 0.5 \times 0.7 \\ \Rightarrow 0.35 = 0.35 \quad [\text{proved}] \end{array} \right.$$

Likelihood Probability: $P(x|\omega_i)$ [x=feature ; ω =class]

ω_i class एवं x feature यांद्यारा probability

{if probability of x_1 event ω_1 greater than probability of x_1 event ω_2 तरीके x_1 belong कर्ये ω_1 class एवं

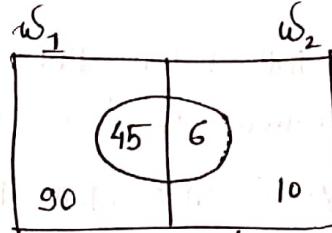
\rightarrow if $P(x_1|\omega_1) > P(x_1|\omega_2)$ then $x_1 \in \omega_1$

एवं kind of decision rule एवं गणे याज्ञ कर्ये

Prior Probability : something given in advance পৰিৱেজন : কোৱাৰি কিছি

$$P(\omega_1) = 0.7 ; P(\omega_2) = 0.3$$

likelihood : $P(x|\omega_1) = \frac{45}{90} = 0.5$
 $P(x|\omega_2) = \frac{6}{10} = 0.6$



যদি কৃতি কৃতি Likelihood think কৰি: ω_2 class এ যাবে ($0.6 > 0.5$)

যদি কৃতি prior think কৰি: ω_1 class এ যাবে ($0.7 > 0.3$)

এটা ambiguity হ'লি. ইচ্ছা, মানে কৃতি likelihood

নিলে এটা class আবাবে কৃতি prior নিলে আবাবে class

তাহে bayes theorem এ posterior probability আনা যাব।

$$P(\omega_i|x) = \frac{P(x|\omega_i) P(\omega_i)}{P(x)} \quad \left. \begin{array}{l} \text{likelihood} \\ \text{prior} \\ \text{evidence/normalization factor} \end{array} \right\} \rightarrow \text{এটা bayes theorem/rule}$$

decision rule: if $P(\omega_1|x) > P(\omega_2|x)$ then $x \in \omega_1$

Bayesian Decision Theory : এটা একটা classifier

decision make কৰে classifier design কৰাব উন্ত এই risk minimize
 কৰি। risk = classification error | ω = class/state of nature

conditional probability density = likelihood = class conditional probability

conditional probability = posterior

যদি ω_1 class এ আওয়াব কৰা হ'লে
 ω_2 class এ যাব / ω_2 class এ
 আওয়াব কৰা হ'লে ω_1 class এ যাব

$$P(\text{error}) = \begin{cases} P(\omega_1) & \text{if we decide } \omega_2 \\ P(\omega_2) & \text{if we decide } \omega_1 \end{cases}$$

এটা probability এই ইন্দ্র্য এই minimum $\left\{ P(\text{error}) = \min [P(\omega_1), P(\omega_2)] \right.$
 আবাবে আবাবে error

ω_1 if $P(\omega_1|x) > P(\omega_2|x)$; else ω_2

ω_1 if $P(x|\omega_1) P(\omega_1) > P(x|\omega_2) P(\omega_2)$; else ω_2

ω_1 if $\underbrace{P(x|\omega_1)/P(x|\omega_2)}_{\text{likelihood ratio}} > \underbrace{P(\omega_2)/P(\omega_1)}_{\text{threshold}}$; else ω_2 \rightarrow Bayes Rules

$$P(\text{errorc}|x) = \begin{cases} P(\omega_1|x) & \text{if we decide } \omega_2 \\ P(\omega_2|x) & \text{if we decide } \omega_1 \end{cases} \} \text{ probability of errorc}$$

$$\therefore P(\text{errorc}|x) = \min [P(\omega_1|x), P(\omega_2|x)]$$

$$P(\text{errorc}) = \int_{-\infty}^{\infty} P(\text{errorc}, x) dx = \int_{-\infty}^{\infty} P(\text{errorc}|x) P(x) dx \} \text{ average probability of errorc}$$

Minimum errorc classifier यले कर्त्ता } Bayes Theorem
errorc टाक्के यात्ते उक्तृप्ति minimum कर्त्ता

यदि एक अन्य विशिष्ट feature/class थाके तथा उनके general errorc function यानाना यात्ते यात्ते नाम है expected risk या cost अन्य जाते associate थाक्के आप्पे cost लाग्नि वास्तव्य loss function थाके। Loss function यसके correct रुपी '1' दिले, misclassified रुपी '0' दिले असले।

Feature vector x = Real number

$\omega_1, \omega_2, \omega_3, \dots, \omega_c = c$ क्षेत्रातील class

action = कर्त्ता class अथवा यात्ते तो = $\omega_1, \omega_2, \dots, \omega_c$ क्षेत्रातील class

loss function = $\lambda(x_i|\omega_j)$ [$x_i = \text{action}; \omega_j = \text{class}$]

{ माने decide कर्त्ता x_i ; class अथवा
किन्तु आजले यात्ते अथवा class ω_j

‘X’ কে লাগলাগ ; class এ কিন্তু আবলে রয়ে ; class

$$R(\alpha_i | x) = \sum_{j=1}^c \lambda(\alpha_i | w_j) \cdot P(w_j | x)$$

loss function

expected loss }
conditional risk }

actually j
class \neq i \rightarrow $\lambda(\alpha_i | w_j) = 0$

actually w_j \rightarrow $\lambda(\alpha_i | w_j) = 1$

actually w_j \rightarrow $\lambda(\alpha_i | w_j) = 1$

class \neq i \rightarrow $P(w_j | x) = 0$

class i \rightarrow $P(w_j | x) = 1$

যদিগুলো 'e'
অসংখ্যক class
তাই c times add হবে

$$\text{overall risk : } R = \int R(\alpha(x) | x) \cdot p(x) dx$$

अतः $\alpha(x) = \text{general decision rule} = \text{determine करें योनि action}$
 निये x पर जैसे गाने a_1 ?
 नाकि a_2 ? नाकि a_3 ?

Bayes rule overall risk एवं क्षमाय, loss एवं क्षमाय।

Step: 1 - अब $R(\alpha_i | x)$ calculate करें यह for given x

step: 2 - minimum $R(\alpha; x)$ choose यद्यपि

ये class में error करने feature की एवं class के बारे +
ये class में probability यानि feature की एवं class के बारे

$$\lambda_{ij} = \lambda(\alpha_i | \omega_j) \quad \left. \right\} \text{यात्रावाले जन्य ए loss टा लाई } \lambda_{ij}$$

माने λ_{ij} एवं i_j इले errors नाही याने predict कराउची;

but real \hat{A}_i ଏବଂ $i=j$ means ଅଣିପାଇଁ predict କରିଛି.

$$\begin{aligned}
 \text{conditional risks} \rightarrow & \left\{ \begin{array}{l} R(\alpha_i|x) = \sum_{j=1}^c \lambda(\alpha_i|\omega_j) P(\omega_j|x) \\ R(\alpha_1|x) = \lambda_{11} P(\omega_1|x) + \lambda_{12} P(\omega_2|x) \\ R(\alpha_2|x) = \lambda_{21} P(\omega_1|x) + \lambda_{22} P(\omega_2|x) \end{array} \right\} \begin{array}{l} \text{দুইটা class} \\ \text{চিন্তা করালে} \\ \text{Two category classification} \end{array} \\
 & \text{minimum error} \rightarrow \lambda_{11}, \lambda_{22} = \text{no error} ; \lambda_{12}, \lambda_{21} = \text{error}
 \end{aligned}$$

ω_1 if $R(\omega_1|x) < R(\omega_2|x)$; else ω_2 } minimum risk
 ज्ञान द्वारा class ए loss कम अंक class ए घाट्छ } decision rule

ω_1 if $(\lambda_{21} - \lambda_{11}) P(\omega_1|x) > (\lambda_{12} - \lambda_{22}) P(\omega_2|x)$; else ω_2

ω_1 if $\underbrace{\frac{P(x|\omega_1)}{P(x|\omega_2)}}_{\text{likelihood ratio}} > \underbrace{\frac{(\lambda_{12} - \lambda_{22})}{(\lambda_{21} - \lambda_{11})}}_{\text{threshold}} \frac{P(\omega_2)}{P(\omega_1)}$; else ω_2

Zero-one loss function

loss function : $i=j$ तो 0 } } तथा $\frac{\lambda_{12} - \lambda_{22}}{\lambda_{21} - \lambda_{11}} = \frac{1-0}{1-0} = 1$
 $i \neq j$ तो 1

Minimum risk classification rule

$$\begin{aligned}
 \lambda(\alpha_i|\omega_i) &= \begin{cases} 0 & i=j \\ 1 & i \neq j \end{cases} \\
 \therefore R(\alpha_i|x) &= \sum_{j=1}^c \underbrace{\lambda(\alpha_i|\omega_j)}_{i \neq j} P(\omega_j|x) = \sum_{i \neq j} P(\omega_j|x) = 1 - P(\omega_i|x)
 \end{aligned}$$

$i \neq j$ ज्ञान ना होने के लिए

So, R क्या हो चाहे $P(\omega_i|x)$ बड़ा हो या कम, R कम हो probability वाले।

ज्ञान एक ही जाते होते loss घटाओ + max probability निर्मित होते हैं।

So, Minimum risk classification एवं objective fulfilled होते हैं।

Discriminant function for Bayesian theory:

general loss function : $g_i(x) = -R(\alpha_i|x)$

zero-one loss function : $g_i(x) = P(\omega_i|x)$

$$g_i(x) = \frac{P(x|\omega_i) P(\omega_i)}{P(x)}$$

$= P(x|\omega_i) P(\omega_i)$ (एक class गुला आकर प्राप्त एवं evidence same होने पर $P(x)$ बहु दर्शा याया)

$$\begin{aligned}
 \text{new discriminant function} &\rightarrow = \ln P(x|\omega_i) + \ln P(\omega_i)
 \end{aligned}$$

dichotomizer

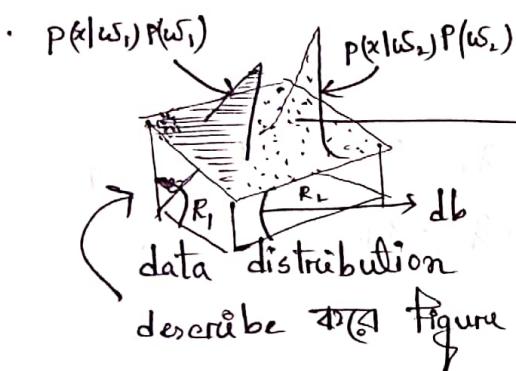
if $g(x) > 0 : g(x) = g_1(x) - g_2(x)$ (db equation)
 else ω_2

$$= P(\omega_1|x) - P(\omega_2|x)$$

$$= \ln \frac{P(x|\omega_1)}{P(x|\omega_2)} + \ln \frac{P(\omega_1)}{P(\omega_2)}$$

likelihood ratio

in applied ক্ষেত্র
math ক্ষেত্রে
easy রয়ে



উচ্চ প্রায়ানো রেটার: ঘনত্ব / density এবং জটিল
গ্রান ক্রমতি দিয়ে বুঝিয়েছে যেমন data.

এবং Gaussian/normal distribution এল

এটি feature এবং উপর base করে data distribution করব

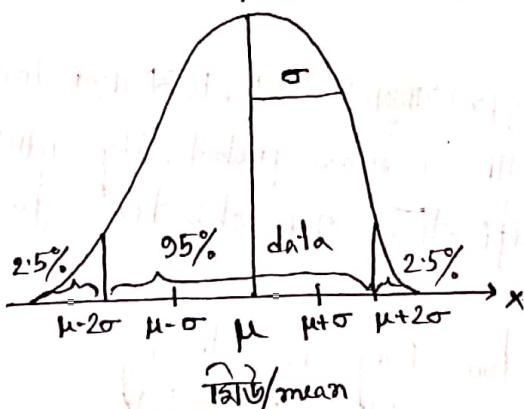
আস্টারে এল univariate normal distribution.

এটি feature এল আকে multivariate এল।

$p(x)$ — pdf

bell shape

normal distribution



σ^2 = variation / variance

σ = standard deviation

mean এ data থেকে থাকে

বড় mean এবং জ্ঞানোচ্চ
narrow থাকে।

variance = mean থেকে

data গুলো কতখানি

scattered.

mean এবং কাহু data
density maximum

only 'x' feature এবং উপর
base করে draw করা থাই
এটি univariate

normal distribution এর equation:

$$P(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

$$P(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right]$$

univariate normal density equation

μ আর x এর দুরত্ব = $x-\mu$

σ দিয়ে ভাগ = normalize এর জন্য

continuous density

probability density/distribution function

$$P(x) \sim N(\mu, \Sigma)$$

normal distribution function

convert করার সাহায্যে μ + Capital sigma (Σ)

co-variance matrix

এক এক স্টোরি feature

ব্যাকল

$$P(x) = \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp\left[-\frac{1}{2} (x-\mu)^T \cdot \Sigma^{-1} (x-\mu)\right]$$

$|\Sigma|$ = determinant

Σ = $d \times d$ covariance matrix

$$\Sigma = \begin{bmatrix} \sigma_1^2 & \dots & \sigma_d^2 \\ \sigma_d \dots & \dots & \sigma_d^2 \end{bmatrix}$$

variance

co-variance/co-correlation

$$(x-\mu)^T \cdot (x-\mu) = (x-\mu)^2$$

$[a^T a = a^2]$ formula

$$\text{normal distribution} \hookrightarrow N(x | \mu, \sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{1}{2\sigma^2} (x-\mu)^2\right]$$

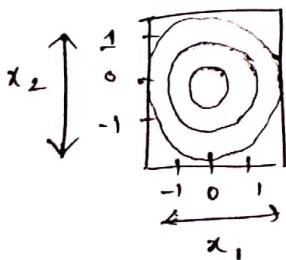
correlated শালে এটা বাড়লে আবেকাদা বাড়বে means x_1 আর x_2

correlated হবে যদি x_1 বাড়লে x_2 বাড়ে, x_1 কমলে x_2 কমি।

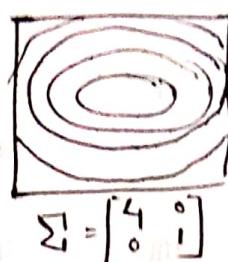
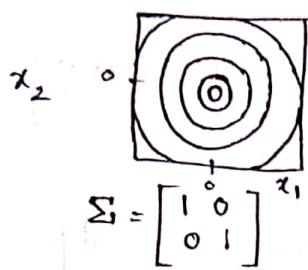
x_1, x_2 aren't correlated.

circle shape \Rightarrow correlated হয়না, oval shape \Rightarrow হয়।

top view দেখলে জোকে কন্টুর হবে।



2d multivariate normal density



आणे diagonal \rightarrow $\Sigma = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$

तरी variance = 1 आणे

$\Sigma = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$ covariance means independent (correlation नाही)

$\mu = [0, 0]$ means mean

0, 0 आहे

$\Sigma = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$

Σ like before

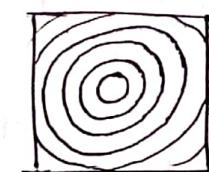
$\mu = [1, 1]$

आणे just mean changed

x_1 आहे
जट्यां variance changed 1 आण्या
shape $x_1 \neq x_2$

वर्गावरु लाग्या असेही।
means x_2 आहे
दिलें density असी

x_1 आहे
जट्यां गव इव्वे



$$\Sigma = \begin{bmatrix} 1 & 0.5 \\ 0.5 & 1 \end{bmatrix}$$

co-variance आडालात
आवृत्ती squeezed असेही



$$\Sigma = \begin{bmatrix} 1 & 0.9 \\ 0.9 & 1 \end{bmatrix}$$



$$\Sigma = \begin{bmatrix} 1 & -0.5 \\ -0.5 & 1 \end{bmatrix}$$

negative आहे
जट्यां figure घुसे
गिरावट



$$\Sigma = \begin{bmatrix} 1 & -0.9 \\ -0.9 & 1 \end{bmatrix}$$

co-variance देखी visualize करून याचा ये कोन feature आवडा
दृश्यगात्र आवृत्ती होताच दृश्यात नाही।

circular $\left\{ \begin{array}{l} \text{circular आणे variance आणार + co-variance नाही} \\ \text{circular दार्दी होणार न दिलें squeeze असेही आणे variance} \end{array} \right.$

non circular $\left\{ \begin{array}{l} \text{angle करू येत एल-variance + co-variance both आही,} \\ \text{diagonal वर्गावरु, याची वार्ता squeezed असेही co-variance} \end{array} \right.$

discriminant function: $g_i(x) = \ln p(x|\omega_i) + \ln P(\omega_i)$

$$\text{if } p(x|\omega_i) \sim N(\mu_i, \Sigma_i) \text{ then } q_i(x) = -\frac{1}{2}(x - \mu_i)^T \Sigma_i^{-1} (x - \mu_i) - \frac{d}{2} \ln 2\pi - \frac{1}{2} \ln |\Sigma_i| + \ln p(\omega_i)$$

multivariate gaussian density \rightarrow

case I : $\Sigma_i = \sigma^2 I$ (diagonal matrix)

(variance आहे बरु co-variance नाही)

যেহেতু covariance নাই, তাই statistically independent +
প্রিয়ের feature এর variance অন্যান

$$\therefore q_i(x) = -\frac{(x - \mu_i)^2}{2\sigma^2} + \ln P(\omega_i) \left[\text{disregarding } \frac{d}{2} \ln 2\pi \text{ and } \frac{1}{2} \ln |\Sigma_i| \right]$$

$$\|x - \mu_i\|^2 = (x - \mu_i)^T (x - \mu_i)$$

$$\therefore q_i(x) = -\frac{1}{2\sigma^2} [x^T x - 2\mu_i^T x + \mu_i^T \mu_i] + \ln P(\omega_i) [||x||^2 = x^T \cdot t]$$

disregarding $x^T x$ (constant),

$$= \frac{-2\mu_i^t x}{-2\sigma^2} - \frac{\mu_i^t \mu_i}{2\sigma^2} + \ln P(\omega_i)$$

$$= \underbrace{\frac{\mu_i^t}{\sigma^2} x_i}_{w_i^t} + \underbrace{\left[-\frac{\mu_i^t \mu_i}{2\sigma^2} + \ln P(\omega_i) \right]}_{w_{i0}}$$

$$= w_i^T x + w_{i0} \quad [\text{which is a linear discriminant}]$$

For $\mathbf{d}\mathbf{b}$: $g_i(x) = g_j(x)$

$$w^T(x - x_0) = 0$$

hyperplane \vec{w} ମେ ଏକ ଉଚ୍ଚତା କିମ୍ବା perpendicular

$$w = \mu_i - \mu_j \leftarrow \text{reason of orthogonal}$$

এবং x দিয়ে ঘোষণা করা হবে।

$$x_0 = \frac{1}{2} (\mu_i + \mu_j) - \frac{\sigma^2}{\|\mu_i - \mu_j\|^2} \ln \frac{P(\omega_i)}{P(\omega_j)} (\mu_i - \mu_j)$$

μ_1 এবং μ_2 রে line
draw কুঠালে আঁ mid
point ঢাকি x.

From equation ①,

$$g_i(x) = -\frac{1}{2\sigma^2} (-2\mu_i^T x + \mu_i^T \mu_i) + \ln P(\omega_i)$$

$$\therefore g_i(x) = g_j(x) \quad [\text{for db}]$$

$$\Rightarrow -\frac{1}{2\sigma^2} (-2\mu_i^T x + \mu_i^T \mu_i) + \ln P(\omega_i) = -\frac{1}{2\sigma^2} (-2\mu_j^T x + \mu_j^T \mu_j) + \ln P(\omega_j)$$

$$\Rightarrow \frac{-2\mu_i^T x}{-2\sigma^2} - \frac{\mu_i^T \mu_i}{2\sigma^2} + \ln P(\omega_i) = \frac{-2\mu_j^T x}{-2\sigma^2} - \frac{\mu_j^T \mu_j}{2\sigma^2} + \ln P(\omega_j)$$

$$\Rightarrow \frac{\mu_i^T x}{\sigma^2} - \frac{\mu_i^T \mu_i}{2\sigma^2} + \ln P(\omega_i) = \frac{\mu_j^T x}{\sigma^2} - \frac{\mu_j^T \mu_j}{2\sigma^2} + \ln P(\omega_j)$$

$$\Rightarrow \frac{\mu_i^T x}{\sigma^2} - \frac{\mu_j^T x}{\sigma^2} - \frac{\mu_i^T \mu_i}{2\sigma^2} + \frac{\mu_j^T \mu_j}{2\sigma^2} + \ln P(\omega_i) - \ln P(\omega_j) = 0$$

$$\Rightarrow \frac{x}{\sigma^2} (\mu_i^T - \mu_j^T) - \frac{1}{2\sigma^2} [\|\mu_i\|^2 - \|\mu_j\|^2] + \ln \frac{P(\omega_i)}{P(\omega_j)} = 0$$

$$\Rightarrow x(\mu_i^T - \mu_j^T) - \frac{(\mu_i + \mu_j)(\mu_i - \mu_j)}{2} + \sigma^2 \ln \frac{P(\omega_i)}{P(\omega_j)} = 0$$

$$\Rightarrow (\mu_i^T - \mu_j^T) \left[x - \frac{(\mu_i + \mu_j)}{2} + \frac{\sigma^2}{(\mu_i - \mu_j)} \ln \frac{P(\omega_i)}{P(\omega_j)} \right] = 0$$

$$\Rightarrow (\mu_i^T - \mu_j^T) \underbrace{\left[x - \left\{ \frac{1}{2}(\mu_i + \mu_j) - \frac{\sigma^2}{\|\mu_i - \mu_j\|^2} \ln \frac{P(\omega_i)}{P(\omega_j)} (\mu_i - \mu_j) \right\} \right]}_{w^T x_0} = 0$$

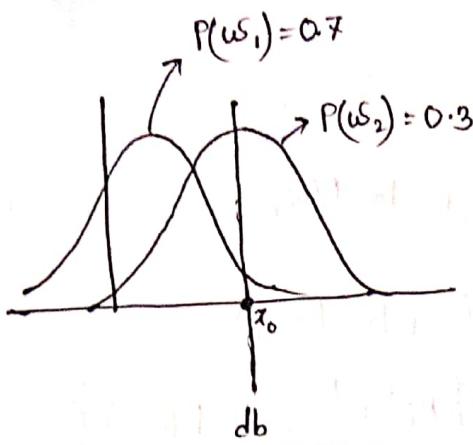
$$\therefore w^T (x - x_0) = 0$$

यदि 2nd class का prior probability equal है, तो तो $P(\omega_1) = P(\omega_2)$

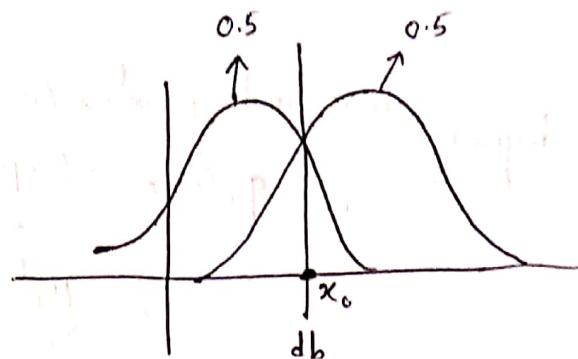
गाने $\ln(1)$ which is 0, तो $\frac{\sigma^2}{\|\mu_i - \mu_j\|^2} \ln \frac{P(\omega_i)}{P(\omega_j)} (\mu_i - \mu_j) = 0$ होगा।

तो तो $x_0 = \frac{1}{2}(\mu_i + \mu_j)$ होगा।

$P(\omega_1) \neq P(\omega_2)$ हले, याएं prior probability तोकि x_0 आदित्य नहीं होगा।



যেতে $P(\omega_1) > P(\omega_2)$
তবে x_0 টা $P(\omega_1)$ অন্তরে
দুই অন্তরে গিয়ে



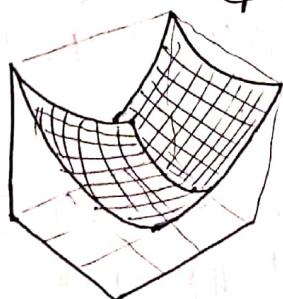
যখন ২টা class মধ্যে
priority অন্তরে অথবা
middle দিয়ে intersect করবে।

priority অন্তরে অথবা minimum distance classifier এর কাটো ধারা।

$$\text{discriminant becomes: } q_i(x) = -\frac{\|x - \mu_i\|^2}{2\sigma^2} + \ln P(\omega_i)$$

$$= -\|x - \mu_i\|^2 \text{ (euclidean distance)}$$

Case : II : ২টা class মধ্যে co-variance matrix অন্তরে : $\Sigma_i = \Sigma$
clusters have hyperellipsoidal shape and same size (centered at μ)



disregarding $\frac{d}{2} \ln 2\pi$ and $\frac{1}{2} \ln |\Sigma_i|$ constants :

$$q_i(x) = -\frac{1}{2} (x - \mu_i)^T \Sigma^{-1} (x - \mu_i) + \ln P(\omega_i)$$

$$= \frac{(x - \mu_i)^T (x - \mu_i)}{-2\Sigma} + \ln P(\omega_i)$$

$$= \frac{\|x - \mu_i\|^2}{-2\Sigma} + \ln P(\omega_i)$$

$$= \frac{\|x\|^2 - 2 \cdot x \cdot \mu_i^T + \|\mu_i\|^2}{-2\Sigma} + \ln P(\omega_i)$$

$$= \frac{x \cdot x - 2x\mu_i^T + \mu_i^T \cdot \mu_i}{-2\Sigma} + \ln P(\omega_i)$$

Priority rule
\$\omega_1\$ এর অন্তরে

disregarding $x^t \cdot x$ (constant) :

$$q_i(x) = -\frac{2x\mu_i^t + \mu_i^t \mu_i}{-2\Sigma} + \ln P(\omega_i)$$

$$= \frac{\mu_i^t}{\Sigma} x - \frac{\mu_i^t \mu_i}{2\Sigma} + \ln P(\omega_i)$$

$$= \underbrace{\Sigma^{-1} \mu_i^t x}_{w_i^t} + \underbrace{\left\{ -\frac{\mu_i^t \mu_i \Sigma^{-1}}{2} + \ln P(\omega_i) \right\}}_{w_{i0}}$$

$$= w_i^t x + w_{i0} \quad [\text{which is a linear discriminant}]$$

For db : $q_i(x) = q_j(x)$

$$\Rightarrow \frac{\mu_i^t}{\Sigma} x - \frac{\mu_i^t \mu_i}{2\Sigma} + \ln P(\omega_i) = \frac{\mu_j^t}{\Sigma} x - \frac{\mu_j^t \mu_j}{2\Sigma} + \ln P(\omega_j)$$

$$\Rightarrow \frac{\mu_i^t}{\Sigma} x - \frac{\mu_j^t}{\Sigma} x - \frac{\mu_i^t \mu_i}{2\Sigma} + \frac{\mu_j^t \mu_j}{2\Sigma} + \ln P(\omega_i) - \ln P(\omega_j) = 0$$

$$\Rightarrow x \{(\mu_i^t - \mu_j^t) \Sigma^{-1}\} - \left(\frac{\mu_i^t \mu_i - \mu_j^t \mu_j}{2} \right) \Sigma^{-1} + \ln \frac{P(\omega_i)}{P(\omega_j)} = 0$$

$$\Rightarrow x \{ \Sigma^{-1} (\mu_i^t - \mu_j^t) \} - \Sigma^{-1} \left\{ \frac{1}{2} (\|\mu_i\|^2 - \|\mu_j\|^2) \right\} + \frac{\|\mu_i - \mu_j\|^2}{\|\mu_i - \mu_j\|^2} \cdot \ln \frac{P(\omega_i)}{P(\omega_j)} = 0$$

$$\Rightarrow \Sigma^{-1} (\mu_i^t - \mu_j^t) \left[x - \frac{1}{2} (\mu_i + \mu_j) + \frac{\ln [P(\omega_i)/P(\omega_j)] (\mu_i - \mu_j)}{\Sigma^{-1} \|\mu_i - \mu_j\|^2} \right] = 0$$

$$\Rightarrow \underbrace{\Sigma^{-1} (\mu_i^t - \mu_j^t)}_{w^t} \left[x - \underbrace{\left\{ \frac{1}{2} (\mu_i + \mu_j) - \frac{\ln [P(\omega_i)/P(\omega_j)] (\mu_i - \mu_j)}{\Sigma^{-1} \|\mu_i - \mu_j\|^2} \right\}}_{x_0} \right] = 0$$

$$\therefore w^t (x - x_0) = 0$$

hyperplane \vec{w} perpendicular রেখে তা w^t নির্ণয় আর্থ।

বরং আজ্ঞানাত্তি জাতে ঘাতে কেন্তা w^t এর দ্বার্যে Σ^{-1} আছে।
x. point দিবারে ঘাতে।

reason of
not orthogonal
 Σ^{-1} বলছে এর direction দ্বারা
hyperplane নির্মাণ দিকে না।

prior equal अल्लि, $P(\omega_i) = P(\omega_j)$

$$\text{discriminant : } g_i(x) = -\frac{1}{2}(x - \mu_i)^t \Sigma_i^{-1} (x - \mu_i) + \ln P(\omega_i)$$

$$= -\frac{1}{2}(x - \mu_i)^t \Sigma_i^{-1} (x - \mu_i) \quad [\text{mahalanobis distance classifier}]$$

higher dimension में एक point के distance यह फूले

चाहिए mahalanobis distance classifier use हो।

Case III : Most General Case : Σ_i = arbitrary

Clusters have different shapes and sizes

disregarding $\frac{d}{2} \ln 2\pi$ (constant) :

$$g_i(x) = -\frac{1}{2}(x - \mu_i)^t \Sigma_i^{-1} (x - \mu_i) - \frac{1}{2} \ln |\Sigma_i| + \ln P(\omega_i)$$

$$= -\frac{1}{2} \Sigma_i^{-1} \|x - \mu_i\|^2 - \frac{1}{2} \ln |\Sigma_i| + \ln P(\omega_i)$$

$$= -\frac{1}{2} \Sigma_i^{-1} \{ \|x\|^2 - 2 \cdot x \cdot \mu_i + \|\mu_i\|^2 \} - \frac{1}{2} \ln |\Sigma_i| + \ln P(\omega_i)$$

$$= -\frac{1}{2} \Sigma_i^{-1} \|x\|^2 + x^t \Sigma_i^{-1} \mu_i - \frac{1}{2} \Sigma_i^{-1} \|\mu_i\|^2 - \frac{1}{2} \ln |\Sigma_i| + \ln P(\omega_i)$$

$$= x^t \underbrace{\left(-\frac{1}{2} \Sigma_i^{-1} \right)}_{W_i} x + \underbrace{\left(\Sigma_i^{-1} \mu_i \right)}_{w_i} \cdot x + \underbrace{\left(-\frac{1}{2} \mu_i^t \Sigma_i^{-1} \mu_i - \frac{1}{2} \ln |\Sigma_i| + \ln P(\omega_i) \right)}_{w_{i0}}$$

$$= x^t W_i x + w_i + w_{i0} \quad [\text{quadratic discriminant}] \quad [\text{not linear}]$$

db is determined by hyperquadrics : $g_i(x) = g_j(x)$

e.g. hyperplanes, pairs of hyperplanes, hyperspheres, hyperellipsoids, hyperparaboloids etc.

Classifiers based on:

MLE = maximum likelihood event

MAP = maximum a-posteriori probability

$$\text{MAP} : P(\omega_i | x) = \frac{P(x | \omega_i) \cdot P(\omega_i)}{P(x)} \quad (\text{Bayes theorem})$$

$$\text{MLE} : P(x | \omega_i) = \frac{1}{\sqrt{2\pi} \sigma} e^{-\frac{1}{2} \left(\frac{x-\mu}{\sigma}\right)^2} \quad (\text{Gaussian density function})$$

Parameter Estimation : (MLE)

class conditional densities = likelihood = $P(x | \omega_i)$

minimum error classifier / optimal classifier, বানানা

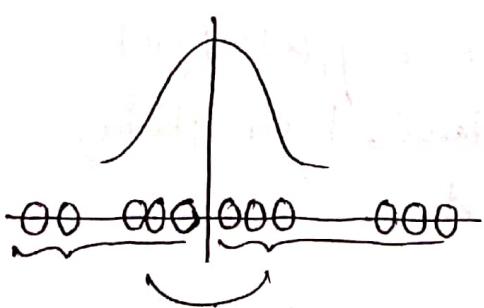
অন্তর্ভুক্ত যদি likelihood আবু priori জানা থাকে, কিন্তু real life এ priori পাওয়া গেলে likelihood পাওয়া difficult, So, complete knowledge না থাকা অন্তর্ভুক্ত probability

যেখানে ক্ষয় ক্ষেত্রে দৃষ্টিয়া, অসম্পর্কযুক্ত likelihood জানা থাকেনা,

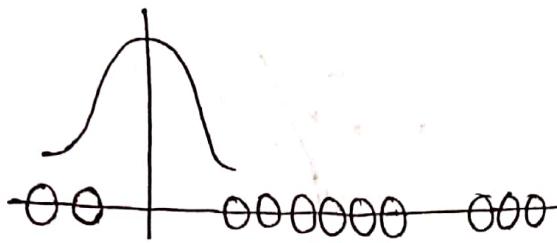
তাহে bayesian distribution best হওয়া অন্তর্ভুক্ত gaussian/ bernoulli/poisson/exponential distribution use করা লাগে, prior

মুন্তব্য knowledge 3 জানা না থাকতে পাই, high dimension এ

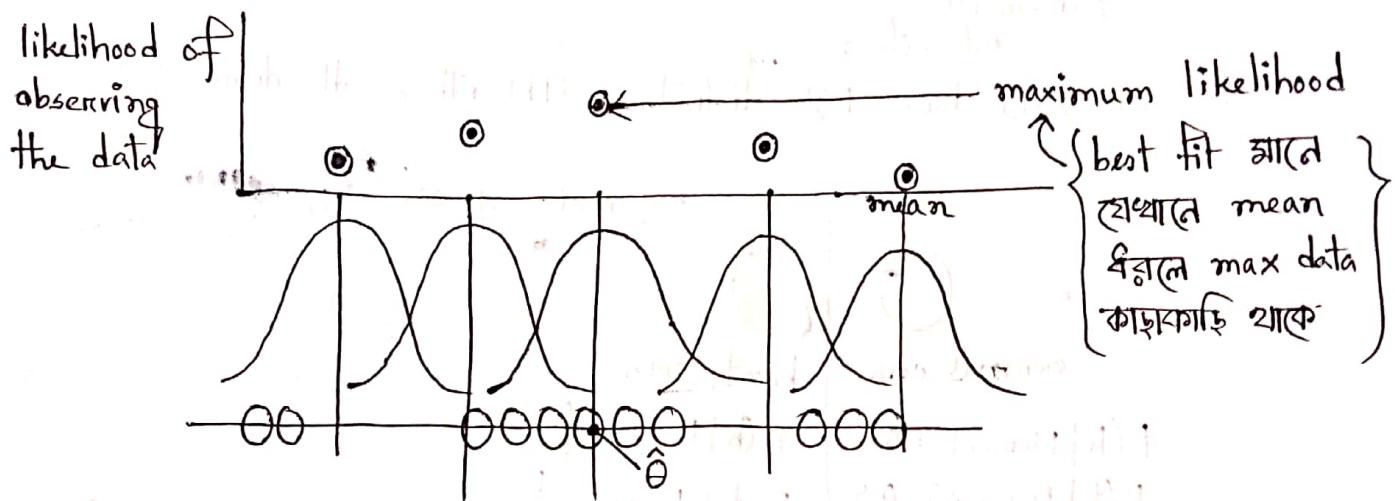
গেলে প্রমাণ problem আসে।



gaussian distribution করালে mean
যেখানে করা হ্যাত্বা, দুর্ধা যাত্রু max
data mean এবং কাছাকাছি, মধ্যান
একটা symmetric structure
maintain করবে।



distribution এর left এ সঞ্চাল
mean এর কাছাকাছি max value
থাকবেন



gaussian distribution এর parameters রয়ে : mean, variance

feature vector x যেকী ষষ্ঠি এল কোন distribution শব্দে জ্ঞান করিন।

MLE : Parameters জ্ঞান কি কিছি parameters রয়ে value জ্ঞান না।

MAP/Bayesian method : কৃতি ব্যৱত্তা prior দ্বাৰা থাকবে।

{যেকীন : mean (μ_i), covariance matrix (Σ_i) জ্ঞান থাকলেও μ_i, Σ_i এর
value জ্ঞান থামবে না।

Linear Discriminant Analysis

Generative Method
Bayesian Decision Rule



Model: $N(\mu, \Sigma)$
decision
 $g(x) = \omega^T x + \omega_0$

Discriminative method

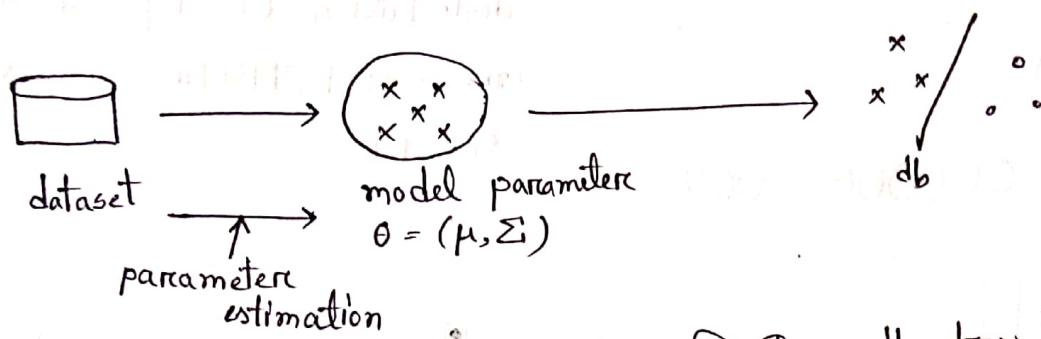
logistic regression
SVM
Deep networks



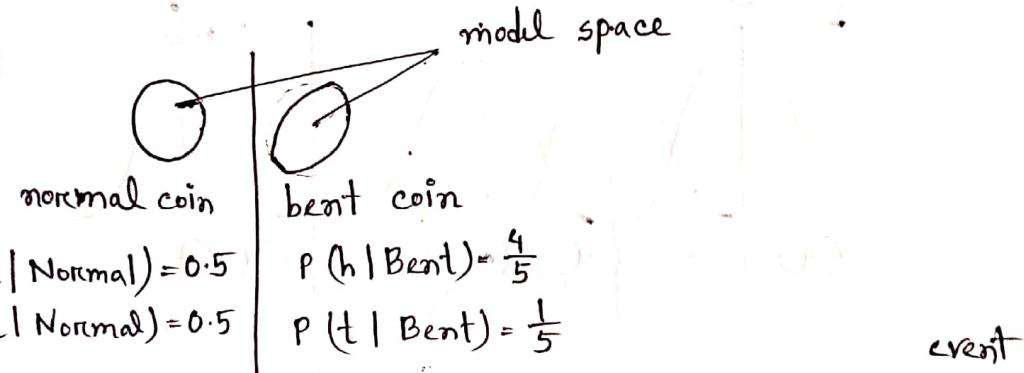
$$g(x) = \omega^T x + \omega_0$$

যারে distribution বেঁধু কৰি then
জেন্টেন্স উন্নয়ন base কৰে decision make কৰি
যা linear/quadratic হবে

(মাঝেজির max কৰে db draw কৰি)



θ এর উপর base করে distribute করি, Then db draw করি।



এখন output মূল হ্যাঁ: HTH HHT THH HTH
 আগলে normal coin কি flip করা ব্যবহৃত নাকি bent coin?
 ans: bent coin কর্তৃত output এ H কি এজাই আর
 bent coin এই H আজার probability কি।

MLE : $\arg \max_{\text{coin} \in \{\text{bent, normal}\}} P(\text{HTTHHTTHHTH} | \text{coin})$

$\arg \max_{\text{model} \in \text{model space}} P(\text{data} | \text{model})$

যানে given model এর respect এ এত data দ্বারা আই
 আদেশ কর্তৃত max করে করালৈ ans চলে আস।

গোচে feature, পাত্র class, এবং এর likelihood। So, the
 whole thing is maximum likelihood কর্তৃত max করাই

$$\therefore P(\text{data} | \text{bent}) = \frac{4}{5} \cdot \frac{1}{5} \cdot \frac{4}{5} \cdot \frac{4}{5} \cdot \frac{4}{5} \cdot \frac{1}{5} \cdot \frac{1}{5} \cdot \frac{4}{5} \cdot \frac{4}{5} \cdot \frac{4}{5} \cdot \frac{1}{5} \approx 0.000268$$

$$\therefore P(\text{data} | \text{normal}) = \frac{1}{2} \cdot \frac{1}{2} \approx 0.000244$$

$\therefore \max = \text{bent}$ (ans) (একজুড়েই করুন likelihood)

i.i.d = independent identically distributed

samples should be i.i.d in MLE

2. जन्में parameters value same

$p(x|\omega_j)$ के लिए यादे $p(x|\omega_j, \theta_j)$

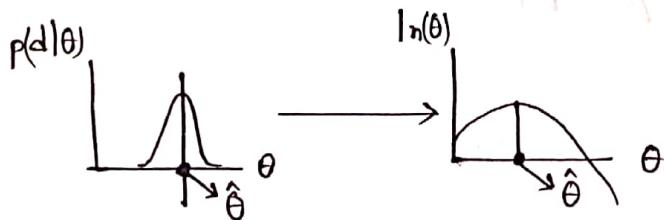
likelihood एवं value टिक है तो
कहु तो but जाए दिये दिये θ_j

θ एवं respect ए x sample गुला तो
कहु गुला कहु दिये

likelihood एवं
equation $\rightarrow \{ F(\theta) = p(D|\theta) = \prod_{k=1}^n p(x_k|\theta) \} \rightarrow (\text{if } D = \text{fixed}, p(D|\theta) \text{ isn't density})$

$p(D|\theta)$ = likelihood of θ with respect to dataset

θ का (ये value जन्में जन्में $p(D|\theta) = \max \ln \theta$ (mean value for gaussian distribution)



\ln निल
curve की spread कहु +
math एवं जूविंग
spread कहुलें max mean same
place है।

$p(x|\theta) \rightarrow \theta$ एवं जन्में x कहु दिये
जहि it's not conditional density

$$\text{MLE: } \hat{\theta} = \arg \max_{\theta} (p(D|\theta))$$

$$\ln (p(D|\theta))$$

$$= \arg \max_{\theta} \left(\ln \prod_{k=1}^n p(x_k|\theta) \right)$$

$$= \arg \max_{\theta} \left(\sum_{k=1}^n \ln p(x_k|\theta) \right)$$

\ln कहाय product (\prod) converted
into sum (Σ)

derivative বের করায় : $\nabla_{\theta} = \left[\frac{\partial}{\partial \theta_1}, \frac{\partial}{\partial \theta_2}, \dots, \frac{\partial}{\partial \theta_p} \right]^T$

θ এর gradient

যেই derivative এর value 0 হলে জিতে max value
 $\nabla_{\theta} I = 0$; $I(\theta) = \ln(p(D|\theta))$.

Suppose, $\mu = \text{unknown}$

$\sigma^2 = \text{variance} = \text{known}$

: unknown needs to be estimated. So, $\theta = \mu$

$$\begin{aligned} \hat{\mu} &= \arg \max_{\mu} I(\mu) = \arg \max_{\mu} \left(\sum_{k=1}^n \ln p(x_k | \mu) \right) \\ &= \arg \max_{\mu} \left(\sum_{k=1}^n \ln \left(\frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x_k-\mu)^2}{2\sigma^2}} \right) \right) \quad [\because p(x|\mu) = N(\mu, \sigma^2)] \\ &= \arg \max_{\mu} \sum_{k=1}^n \left(-\frac{(x_k-\mu)^2}{2\sigma^2} - \ln \sqrt{2\pi}\sigma \right) \end{aligned}$$

$$\therefore \frac{d}{d\mu} (I(\mu)) = 0$$

$$\Rightarrow \sum_{k=1}^n \frac{1}{\sigma^2} (x_k - \mu) = 0$$

$$\Rightarrow \sum_{k=1}^n x_k - n\mu = 0$$

$$\Rightarrow \hat{\mu} = \frac{1}{n} \sum_{k=1}^n x_k \quad (\text{mean এর equation})$$

MLE দিয়ে prove করলাগে যে, আর data sum করে data ক্র্যান্ত

দিয়ে তার মান mean আর এর normally আভিন্ন জানি।

if unknown = σ^2 then

$$\frac{d}{dt} (I(t)) = 0 \quad (t = \sigma^2 \text{ suppose})$$

$$\Rightarrow \frac{d}{dt} \sum_{k=1}^n \left(-\frac{(x_k-\mu)^2}{2t} - \ln \sqrt{2\pi t} \right) = 0$$

$$\Rightarrow \sum_{k=1}^n \left\{ \frac{(x_k - \mu)^2}{2} \cdot (-1) \cdot t^{-2} - \frac{1}{\sqrt{2\pi t}} \cdot \sqrt{2\pi} \cdot \frac{1}{2} \cdot t^{-\frac{1}{2}} \right\} = 0$$

$$\Rightarrow \sum_{k=1}^n \left(\frac{(x_k - \mu)^2}{2t^2} - \frac{1}{2\sqrt{t}} \right) = 0$$

$$\Rightarrow \sum_{k=1}^n \left(\frac{(x_k - \mu)^2}{2t^2} - \frac{1}{2t} \right) = 0$$

$$\Rightarrow \sum_{k=1}^n \left(\frac{(x_k - \mu)^2}{2t^2} - \frac{t}{2} \right) = 0$$

$$\Rightarrow \sum_{k=1}^n \left\{ (x_k - \mu)^2 - t \right\} = 0$$

$$\Rightarrow \sum_{k=1}^n (x_k - \mu)^2 - nt = 0$$

$$\Rightarrow nt = \sum_{k=1}^n (x_k - \mu)^2$$

$$\therefore \sigma^2 = \frac{1}{n} \sum_{k=1}^n (x_k - \hat{\mu})^2 \quad [t = \sigma^2]$$

Expected value \rightarrow mean

$$E(X) = \mu = \begin{cases} \sum_x x f(x) & \text{if } X \text{ is discrete} \\ \int_x x f(x) dx & \text{if } X \text{ is continuous} \end{cases} \quad \left\{ \begin{array}{l} f(x) = \text{probability} \\ \text{distribution} \end{array} \right.$$

$$\begin{array}{ccccccc} 1 & 2 & 3 & 4 & 5 & 6 & \rightarrow x \\ 0.1 & 0.1 & 0.1 & 0.1 & 0.5 & \rightarrow p \end{array}$$

$$\therefore \sum x p = 1 \cdot 0.1 + 2 \cdot 0.1 + 3 \cdot 0.1 + 4 \cdot 0.1 + 5 \cdot 0.1 + 6 \cdot 0.5$$

$$\therefore \mu = 4.5 \quad (\text{expected value})$$

probability distribution $f(x)$

The variance of a random variable X is defined as the expect squared deviation from the expected value.

$$\sigma^2 = \text{Var}(X)$$

$$= E[(X - \mu)^2]$$

$$= E[(X - E(X))^2]$$

$$= \left\{ \begin{array}{l} \sum_x (x - \mu)^2 f(x) \text{ if } X \text{ is discrete} \\ \int_{-\infty}^{\infty} (x - \mu)^2 f(x) dx \text{ if } X \text{ is continuous} \end{array} \right.$$

$$\int_{-\infty}^{\infty} (x - \mu)^2 f(x) dx \text{ if } X \text{ is continuous}$$

$$E[(X - \mu)^2] = E(X^2 - 2\mu X + \mu^2)$$

$$= E(X^2) - 2\mu E(X) + \mu^2$$

$$= E(X^2) - 2\mu \cdot \mu + \mu^2$$

$$= E(X^2) - 2\mu^2 + \mu^2$$

$$= E(X^2) - \mu^2$$

$$\therefore \sigma^2 = \text{Var}(X) = E(X^2) - [E(X)]^2$$

$$\text{Var}(X) = E[(X - \mu)^2]$$

$$= \sum_x (x - \mu)^2 p(x)$$

$$= \sum_x (x^2 - 2x\mu + \mu^2) p(x)$$

$$= \sum_x x^2 p(x) - 2\mu \sum_x x p(x) + \mu^2 \sum_x p(x)$$

$$= E(X^2) - 2\mu^2 + \mu^2$$

$$= E(X^2) - \mu^2$$

$$\therefore \text{Var}(X) = \sum x^2 p - \mu^2$$

$$\sum x^2 p = 0.1 + 0.4 + 0.9 + 1.6 + 2.5 + 1.8$$

$$= 23.5$$

$$= 23.5 - (4.5)^2$$

$$= 3.25$$

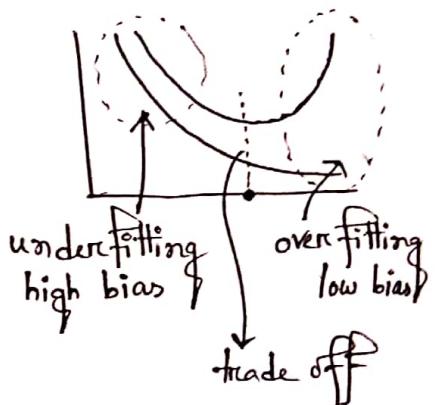
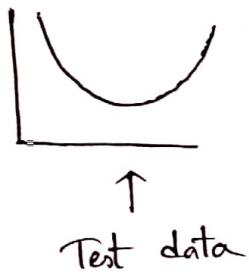
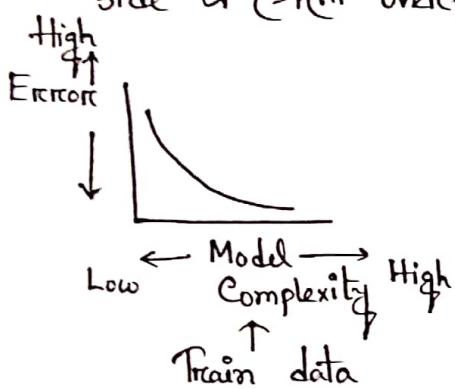
an estimator of a parameter is biased if the expected value of the estimate is different from the true value of the parameters

$$E[\hat{\sigma}^2] = \frac{n-1}{n} \sigma^2 ; E[\hat{\Sigma}] = \frac{n-1}{n} \Sigma$$

an estimator of a parameter is unbiased if the expected value of the estimate is same as the true value of the parameters.

$$E[\hat{\theta}] = \theta ; E[\hat{\mu}] = \mu$$

estimation কর্তৃক আলা রয়েছে যেটা কি কর্ণি bias + variance দিয়ে।
 true value থেকে আবাস estimated value কর্তৃক দূরে অটোই bias,
 bias high আলে distance যেকি, আবাস estimated value পুরু গর্যে
 data শুলো যেকি একজাতে আকে তালে low variance, দূরে ছড়িয়ে
 আকলে high variance, Trade off আলে balance, এ point এ bias +
 variance দুই জোই low লাব এটাই trade off. Trade off এর right
 side এ গোলে overfitting হয়, left side এ গোলে underfitting হয়।



loss curve = training error curve

loss function = error function

